# Mathematical Statistics

How to get information from the experiment?

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## Outline

- 1 Law of Large Numbers and CLT
- 2 Random Sample
- 3 Statistics and Estimators

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## Theorem (Markov's Inequality)

Let X be a random variable and let g(x) be a nonnegative function. Then, for any r > 0

$$P(g(X) \ge r) \le \frac{\mathbb{E}g(X)}{r}$$

### Theorem (Chebyshev's Inequality)

if the variance is small then X is unlikly to be too far from the mean

$$P(|X - \mu| \ge c) \le \frac{\sigma^2}{c^2}$$

### Theorem (Weak Law of Large Numbers)

Let  $X_1, \ldots, X_n$  be iid random variables with  $\mathbb{E} X_i = \mu$  and  $\operatorname{Var} X_i = \sigma^2 < \infty$ . Define  $\bar{X}_n = (1/n) \sum_{i=1}^n X_i$ . Then for every  $\epsilon > 0$ ,

$$\lim_{n\to\infty} P(|\bar{X}_n - \mu| < \epsilon) = 1$$

that is,  $X_n$  converges in probability to  $\mu$ .

The law of large numbers states that the sample mean converges to the distribution mean as the sample size increases, and is one of the fundamental theorems of probability. There are different versions of the law, depending on the mode of convergence.

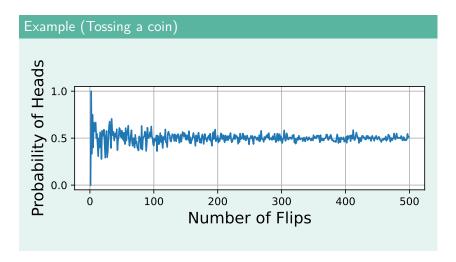
#### Proof.

$$P(|\bar{X}_n - \mu| \ge \epsilon) = P((\bar{X}_n - \mu)^2 \ge \epsilon^2) \le \frac{\mathbb{E}(\bar{X}_n - \mu)^2}{\epsilon^2} = \frac{\sigma^2}{n\epsilon^2}$$

Hence,

$$P(|\bar{X}_n - \mu| < \epsilon) = 1 - P(|\bar{X}_n - \mu| \ge \epsilon) \ge 1 - \frac{\sigma^2}{n\epsilon^2} \to 1$$
, as  $n \to \infty$ 



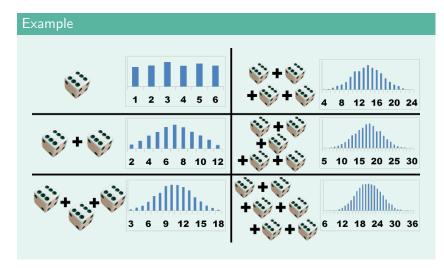


### Theorem (Central Limit Theorem)

Let  $X_1,\ldots,X_n$  be a sequence of iid random variables. Let  $\mathbb{E}X_i=\mu$  and  $\operatorname{Var}X_i=\sigma^2>0$  and finite. Define  $\bar{X}_n=(1/n)\sum\limits_{i=1}^n X_i$ . Let  $G_n(x)$  denote the cdf of  $\sqrt{n}(\bar{X}_n-\mu)/\sigma$ . Then for any  $-\infty < x < \infty$ ,

$$\lim_{n\to\infty} G_n(x) = \int_{-\infty}^x \frac{1}{\sqrt{2\pi}} e^{-y^2/2} dy$$

that is,  $\sqrt{n}(\bar{X}_n - \mu)/\sigma$  has a limiting standard normal distribution.

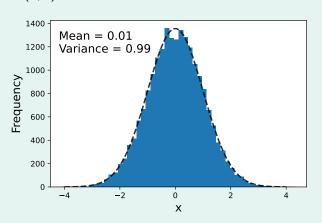


### Theorem

Roughly, the central limit theorem states that the distribution of the sum (or average) of a large number of independent, identically distributed variables will be approximately normal, regardless of the underlying distribution.

### Example

Let  $X_i \sim U(0,1)$ 



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## Random Sample

#### Definition

The random variables  $X_1, \ldots, X_n$  are called a random sample of size n from the population f(x) if  $X_1, \ldots, X_n$  are mutually independent random variables and the marginal pdf or pmf of each  $X_i$  is the same function f(x). Alternatively,  $X_1, \ldots, X_n$  are called independent and identically distributed (iid) random variables with pdf or pmf f(x). Another notation is  $X_{[n]} = [X_1, \ldots, X_n]$ 

#### **Definition**

The list of  $[x_1, \ldots, x_n] = x_{[n]}$  called a sample of size n of realization of **random element** X.

# Population of Sample

for sample  $x_{[n]}$  assume empirical distribution:

$$\mathbb{P} \quad \frac{X_1}{n} \quad \frac{X_2}{n} \quad \dots \quad \frac{X_n}{n}$$

This is discrete distribution. Let A be an event, then

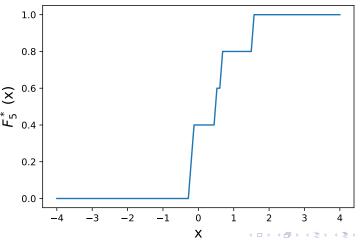
$$P(A) = \frac{1}{n} \sum_{i=1}^{n} [x_i \in A]$$

#### Definition

Empirical random variable  $X^*$  – r.v. that has distribution function  $F_n^*$ 

# Population of sample

For sample 
$$x_{[n]}$$
 assume  $F_n^*(x) = \frac{1}{n} \sum_{i=1}^n [x_i < x]$ 



# Why Sample?

### Theorem (Glivenko-Cantelli)

$$\sup_{x \in \mathbb{R}} |F_n^*(x) - F_X(x)| = ||F_n^* - F_X||_{\infty} \xrightarrow{a.s.} 0$$

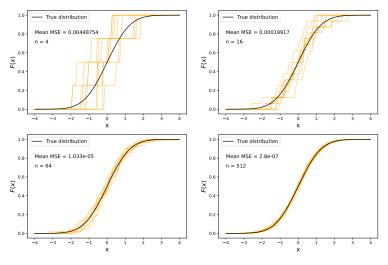
### Theorem (Kolmogorov)

If  $F_X$  is continuous, then

$$\sqrt{n} \|F_n^* - F_X\|_{\infty} \xrightarrow{d} K$$

Where K - Kolmogorov's distribution

# Why Sample?



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### Statistics and Estimators

#### Definition

Let  $X_1, \ldots, X_n$  be a random sample of size n from a population and let  $T = (x_1, \ldots x_n)$  be a real-valued or vector-valued function whose domain includes the sample space of  $(X_i, \ldots X_n)$ . Then the random variable or random vector  $Y = T(X_1, \ldots, X_n)$  is called a statistic. The probability function of a statistic Y is called the sampling distribution of Y.

#### Definition

A point estimator is any function  $W(X_1, ..., X_n)$  of a sample; that is, any statistic is a point estimator.

Estimate = 
$$W(x_1, ... x_n)$$



#### A Note About Statistic

Statistic is a function of sample! Statistic NOT a function of a unknown parameter.

$$T(x_1,\ldots,x_n)=x_1+x_2+\cdots+x_n$$
 Statistic 
$$T(x_1,\ldots,x_n)=\mathbb{E}X_1 - \text{ NOT statistic}$$
  $\mathbb{E}X_i=\mu$   $T(x_1,\ldots,x_n)=\sum (x_i-\mu)^2 - \text{depends on }\mu$  status

#### **Statistics**

#### Definition

The sample mean is the arithmetic average of the values in a random sample. It is usually denoted by

$$\bar{X} = \frac{X_1 + \dots + X_n}{n} = \frac{1}{n} \sum_{i=1}^n X_i$$

#### Definition

The sample variance (or corrected sample variance) is the statistic defined by

$$S^{2} = \frac{n}{n-1} \sum_{i=1}^{n} (X_{i} - \bar{X})^{2}$$

#### The Order Statistics

#### Definition

The order statistics of a random sample  $X_1, \ldots, X_n$  are the sample values placed in ascending order. They are denoted by  $X_{(1)}, \ldots X_{(n)}$ 

The order statistics are random variables that satisfy  $X_{(1)} \leq \cdots \leq X_{(n)}$ . In particular,

$$\begin{cases} X_{(1)} = \min_{1 \le 1 \le n} X_i, \\ X_{(2)} = \text{second smallest } X_i \\ \vdots \\ X_{(1)} = \max_{1 \le 1 \le n} X_i, \end{cases}$$

# Quantiles, Quartiles, Median

# Definition ( $\alpha$ -Quantile)

A number  $x_{\alpha}$  that  $\mathbb{P}(X \leq x_{\alpha}) \geq \alpha$  and  $\mathbb{P}(X \geq x_{\alpha}) \geq 1 - \alpha$ . If continuous we can write  $P(X < x_{\alpha}) = \alpha$ 

# Definition (Sample $\alpha$ -Quantile)

- if  $\frac{k}{n} < \alpha < \frac{k+1}{n}$ , then  $x_{\alpha}^* = x_{(k+1)}$
- if  $\alpha = \frac{k}{n}$ , then  $x_{\alpha}^*$  any number in  $[x_{(k)}, x_{(k+1)}]$

# Definition (Sample Quartile)

$$Q_1 = x_{0.25}^*$$
  $Q_2 = x_{0.5}^*$   $Q_3 = x_{0.75}^*$ 



# Quantiles, Quartiles, Median

# Definition (Sample Median $m^*$ )

- if n = 2k + 1, then  $m^* = x_{(k+1)}$
- if n = 2k, then  $m^* = \frac{1}{2}(x_{(k)} + x_{(k+1)})$

### Method of Moments

Let's remember definitions of the moments:

### Definition

$$\mu_1 = \frac{\left(\sum x\right)}{n}$$

$$\mu_2 = \frac{\sum (x-\mu)^2}{n}$$

$$\mu_3 = \frac{1}{n} \frac{\sum (x-\mu)^3}{\sigma^3}$$

$$\mu_4 = \frac{1}{n} \frac{\sum (x-\mu)^4}{\sigma^4}$$

#### Method of Moments

Let's assume that we have a  $X_1, \ldots, X_n$  sample form population with pdf  $f(x|\theta_1, \ldots, \theta_k)$ 

The idea of the method of moments is to equate the first k sample moments to corresponding k population moments and then solve the system of equation.

#### Method of Moments

#### Example

Suppose  $X_1, \ldots, X_n$  are iid  $n(\theta, \sigma^2)$ .  $\theta_1 = \theta$ ,  $\theta_2 = \sigma^2$ . We have  $m_1 = \bar{X}, m_2 = (1/n) \sum_i X_i^2, \mu_1 = \theta, \mu_2 = \theta^2 + \sigma^2$ , so we must solve

$$\bar{X} = \theta, \quad \frac{1}{n} \sum X_i^2 = \theta^2 + \sigma^2$$

Solving this system for  $heta, \sigma^2$  yields method of moments estimators

$$\tilde{\theta} = \bar{X}, \quad \sigma^2 = \frac{1}{n} \sum_{i} X_i^2 - \bar{X}^2 = \frac{1}{n} \sum_{i} (X_i - \bar{X})^2$$