

Data-Driven Design & Analyses of Structures & Materials (3dasm)

Lecture 6

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Outline for today

- Continuation of previous lecture: Bayesian inference for one hidden rv
 - Prior
 - Likelihood
 - Marginal likelihood
 - Posterior
 - Gaussian pdf's product

Reading material: This notebook + Chapter 3

Recap of Lecture 5: car stopping distance with known $oldsymbol{x}$

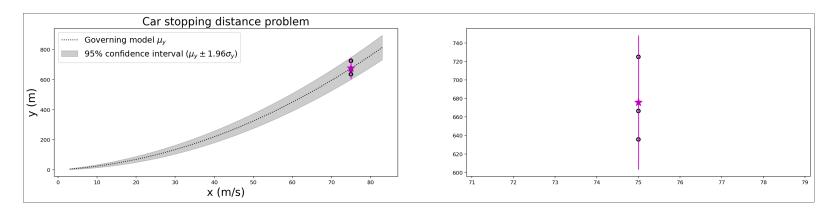
We focused on the car stopping distance problem under the following conditions:

- We kept x=75 m/s.
- ullet But the distribution of the rv z is not known: p(z)=?

Under these conditions, recall the "true" model by observing the following plot, including some data observations.

```
In [4]: # vvvvvvvvvv this is just a trick so that we can run this cell multiple times vvvvvvvvv
       fig car new, ax car new = plt.subplots(1,2); plt.close() # create figure and close it
if fig car new.get axes():
    del ax car new; del fig car new # delete figure and axes if they exist
    fig car new, ax car new = plt.subplots(1,2) # create them again
# ^^^^^^ end of the trick ^^^^^^^^^^^^^^^^^^^
N samples = 3 # CHANGE THIS NUMBER AND RE-RUN THE CELL
x = 75; empirical y = samples y(N samples, x); # Empirical measurements of N samples at x=75
empirical mu y = np.mean(empirical y); empirical sigma y = np.std(empirical y); # empirical mean and std
car fig(ax car new[0]) # a function I created to include the background plot of the governing model
for i in range(2): # create two plots (one is zooming in on the error bar)
    ax car new[i].errorbar(x , empirical mu y,yerr=1.96*empirical sigma y, fmt='m*', markersize=15);
    ax car new[i].scatter(x*np.ones like(empirical y),empirical y, s=40,
                         facecolors='none', edgecolors='k', linewidths=2.0)
print("Empirical mean[y] is",empirical mu y, "(real mean[y]=675)")
print("Empirical std[y] is",empirical sigma y,"(real std[y]=37.5)")
fig car new.set size inches(25, 5) # scale figure to be wider (since there are 2 subplots)
```

Empirical mean[y] is 675.7331498126623 (real mean[y]=675) Empirical std[y] is 37.00345355607556 (real std[y]=37.5)



Recap of Lecture 5: Summary of our model

1. The **observation distribution**:

$$p(y|z) = \mathcal{N}\left(y|\mu_{y|z} = wz + b, \sigma_{y|z}^2
ight) = rac{1}{C_{y|z}} \mathrm{exp}\left[-rac{1}{2\sigma_{y|z}^2}(y-\mu_{y|z})^2
ight]$$

where $C_{y|z}=\sqrt{2\pi\sigma_{y|z}^2}$ is the **normalization constant** of the Gaussian pdf, and where $\mu_{y|z}=wz+b$, with w, b and $\sigma_{y|z}^2$ being constants.

1. and the **prior distribution**: $p(z) = \frac{1}{C_z}$

where $C_z = z_{max} - z_{min}$ is the **normalization constant** of the Uniform pdf, i.e. the value that guarantees that p(z) integrates to one.

Recap of Lecture 5: Data

• Since we usually don't know the true process, we can only observe/collect data $y = \mathcal{D}_y$:

```
In [5]: print("Example of N=%1i data points for y at x=%1.1f m/s with :" % (N_samples,x), empirical_y)
```

Example of N=3 data points for y at x=75.0 m/s with : [635.74192878 724.95905357 666.49846709]

Recap of Lecture 5: Posterior from Bayes' rule applied to data Use Bayes' rule applied to data to determine the posterior:

$$p(z|y=\mathcal{D}_y) = rac{p(y=\mathcal{D}_y|z)p(z)}{p(y=\mathcal{D}_y)}$$

That requires calculating the likelihood (here, it results from a product of Gaussian densities):

$$p(y = \mathcal{D}_y|z) = rac{1}{\left|w
ight|^N} \cdot C \cdot rac{1}{\sqrt{2\pi\sigma^2}} \mathrm{exp}igg[-rac{1}{2\sigma^2} (z-\mu)^2 igg]$$

where
$$\mu = rac{w^2 \sigma^2}{\sigma_{y|z}^2} \sum_{i=1}^N \mu_i$$

$$\sigma^2=rac{\sigma_{y|z}^2}{w^2N}$$
 , and $C=rac{1}{2\pi^{(N-1)/2}}\sqrt{rac{\sigma^2}{\left(rac{\sigma_{y|z}^2}{w^2}
ight)^N}}$

After calculating the likelihood, we determined the marginal likelihood:

$$p(y=\mathcal{D}_y)=rac{C}{\left|w
ight|^NC_z}$$

From which we got the posterior:

$$p(z|y = \mathcal{D}_y) = \frac{p(y = \mathcal{D}_y|z)p(z)}{p(y = \mathcal{D}_y)}$$

$$= \frac{1}{p(y = \mathcal{D}_y)} \cdot \frac{1}{|w|^N} C \cdot \mathcal{N}(z|\mu, \sigma^2) \cdot \frac{1}{C_z}$$
(2)

$$= \frac{1}{p(y = \mathcal{D}_y)} \cdot \frac{1}{|w|^N} C \cdot \mathcal{N}(z|\mu, \sigma^2) \cdot \frac{1}{C_z}$$
(2)

$$= \mathcal{N}(z|\mu, \sigma^2) \tag{3}$$

which is a **normalized** Gaussian pdf in z with mean and variance as shown in the previous cell.

Determining the Posterior Predictive Distribution (PPD) from the posterior

However, as we mentioned, Bayes' rule is just a way to calculate the posterior:

$$p(z|y=\mathcal{D}_y) = rac{p(y=\mathcal{D}_y|z)p(z)}{p(y=\mathcal{D}_y)}$$

What we really want is the Posterior Predictive Distribution (PPD) . This comes after calculating the posterior given some data \mathcal{D}_{v} :

$$p(y|y=\mathcal{D}_y) = \int p(y|z)p(z|y=\mathcal{D}_y)dz$$

which is often written in simpler notation: $p(y|\mathcal{D}_y) = \int p(y|z)p(z|\mathcal{D}_y)dz$

$$p(y|\mathcal{D}_y) = \int \underbrace{p(y|z)}_{ ext{observation}} \underbrace{p(z|y=\mathcal{D}_y)}_{ ext{posterior}} dz$$

Considering the terms we found before, we get:

$$p(y|\mathcal{D}_y) = \int \underbrace{\frac{1}{|w|} \frac{1}{\sqrt{2\pi \left(\frac{\sigma_{y|z}}{w}\right)^2}} \exp\left\{-\frac{1}{2\left(\frac{\sigma_{y|z}}{w}\right)^2} \left[z - \left(\frac{y - b}{w}\right)\right]^2\right\}}_{\text{observation}} \underbrace{\mathcal{N}(z|\mu, \sigma^2)}_{\text{observation}} dz$$
(4)

$$egin{split} m{p}(y|\mathcal{D}_y) &= rac{1}{|w|} \int rac{1}{\sqrt{2\pi \left(rac{\sigma_{y|z}}{w}
ight)^2}} \mathrm{exp} \Bigg\{ -rac{1}{2\left(rac{\sigma_{y|z}}{w}
ight)^2} igg[z-\left(rac{y-b}{w}
ight)igg]^2 \Bigg\} \mathcal{N}(z|\mu,\sigma^2) dz \end{split}$$

$$p(y|\mathcal{D}_y) = rac{1}{|w|} \int \mathcal{N}\left(z\left|rac{y-b}{w}, \left(rac{\sigma_{y|z}}{w}
ight)^2
ight) \mathcal{N}(z|\mu, \sigma^2) dz$$

This is (again!) the product of two Gaussians!

In Lecture 5 (and the Homework!) you saw (and will demonstrate!) that the product of two or more univariate (and multivariate!) Gaussians is...

• Another Gaussian! Although it needs to be scaled by a constant...

So, we conclude that the PPD is an integral of a Gaussian:

$$rac{p(y|\mathcal{D}_y)}{|w|} = rac{1}{|w|} \int C^* \mathcal{N}\left(z|\mu^*, \left(\sigma^*
ight)^2
ight) dz$$

$$\begin{aligned} \text{where } \mu^* &= \left(\sigma^*\right)^2 \left(\frac{\mu}{\sigma^2} + \frac{(y-b)/w}{\left(\frac{\sigma_{y|z}}{w}\right)^2}\right) = \left(\sigma^*\right)^2 \left(\frac{\mu}{\sigma^2} + \frac{(y-b)\cdot w}{\sigma_{y|z}^2}\right) \\ \left(\sigma^*\right)^2 &= \frac{1}{\frac{1}{\sigma^2} + \frac{1}{\left(\frac{\sigma_{y|z}}{w}\right)^2}} = \frac{1}{\frac{1}{\sigma^2} + \frac{w^2}{\sigma_{y|z}^2}} \\ C^* &= \frac{1}{\sqrt{2\pi \left(\sigma^2 + \frac{\sigma_{y|z}^2}{w^2}\right)}} \exp\left[-\frac{\left(\mu - \frac{y-b}{w}\right)^2}{2\left(\sigma^2 + \frac{\sigma_{y|z}^2}{w^2}\right)}\right] \end{aligned}$$

This integral is simple to solve!

$$\frac{p(y|\mathcal{D}_y)}{|w|} = \frac{1}{|w|} \int C^* \mathcal{N}\left(z|\mu^*, (\sigma^*)^2\right) dz \qquad (5)$$

$$= \frac{C^*}{|w|} \int \mathcal{N}\left(z|\mu^*, (\sigma^*)^2\right) dz \qquad (6)$$

What's the result of integrating the blue term?

$$p(y|\mathcal{D}_y) = rac{C^*}{|w|}$$

In Homework 3 (exercise 2)

Rewrite the PPD to show that it becomes:

$$rac{p(y|\mathcal{D}_y)}{p(y|\mathcal{D}_y)} = \mathcal{N}\left(y|b+\mu w, w^2\sigma^2 + \sigma_{y|z}^2
ight)$$

a normalized univariate Gaussian!

A long way to show that the PPD is a simple Gaussian...

$$egin{equation} oldsymbol{p(y|\mathcal{D}_y)} = \mathcal{N}\left(y|b+\mu w, w^2\sigma^2 + \sigma_{y|z}^2
ight) \end{split}$$

where we recall that each constant is:

$$egin{aligned} b &= 0.1x^2 = 562.5 \ w &= x = 75 \ \sigma_{y|z}^2 &= s^2 = 80^2 \ \sigma^2 &= rac{\sigma_{y|z}^2}{w^2N} = rac{s^2}{w^2N} \ \mu &= rac{w^2\sigma^2}{\sigma_{y|z}^2} \sum_{i=1}^N \mu_i = \cdots = rac{\sum_{i=1}^N y_i}{wN} - rac{b}{w} \end{aligned}$$

A long way to show that the PPD is a simple Gaussian...

$$p(y|\mathcal{D}_y) = \mathcal{N}\left(y|b + \mu w, w^2\sigma^2 + \sigma_{y|z}^2\right)$$
(12)

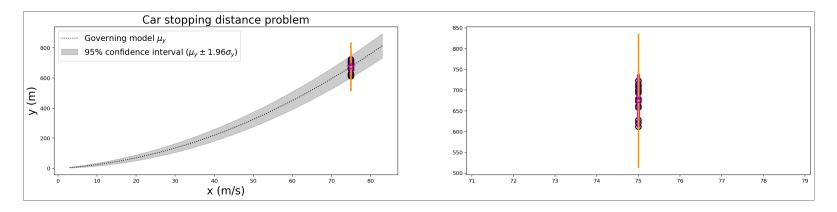
$$= \mathcal{N}\left(y \left| \left(\sum_{i=1}^{N} \frac{y_i}{N}\right), \sigma_{y|z}^2 \left(\frac{1}{N} + 1\right)\right.\right) \tag{13}$$

where y_i are each of the N data points of the observed data \mathcal{D}_y , and $\sigma_{y|z}^2 = s^2 = 80^2$ is the **variance we** assumed because we were not sure about the governing equation for y.

 Very Important Questions (VIQs): What does this result tell us? Did you expect this predicted distribution for y?

```
In [6]: fig car PPD, ax car PPD = plt.subplots(1,2); plt.close() # create figure and close it
       if fig car new.get axes():
    del ax car PPD; del fig car PPD; fig car PPD, ax car PPD = plt.subplots(1,2) # delete fig & axes & create them
N samples = 15 # CHANGE THIS NUMBER AND RE-RUN THE CELL
x = 75; empirical y = samples y(N samples, x); # Empirical measurements of N samples at x=75
empirical mu y = np.mean(empirical y); empirical sigma y = np.std(empirical y); # empirical mean and std
# Calculate PPD mean and standard deviation:
PPD mu y = np.mean(empirical y); sigma yGIVENz = 80; PPD sigma y = np.sqrt( sigma yGIVENz**2*(1/N samples + 1))
car fig(ax car PPD[0]) # a function I created to include the background plot of the governing model
for i in range(2): # create two plots (one is zooming in on the error bar)
    ax car PPD[i].errorbar(x , empirical mu y,yerr=1.96*empirical sigma y, fmt='m*', markersize=15, elinewidth=6);
    ax car PPD[i].errorbar(x , PPD mu y,yerr=1.96*PPD sigma y, color='#F39C12', fmt='*', markersize=5, elinewidth=3);
    ax car PPD[i].scatter(x*np.ones like(empirical y),empirical y, s=100,facecolors='none', edgecolors='k', linewidths=2.6
print("PPD & empirical mean[y] tend to the same value:",empirical mu y, "(real mean[y]=675)")
print("PPD std[y] we predict is", PPD sigma y, "& empirical std[y] is , empirical sigma y, (real std[y]=37.5)")
fig car PPD.set size inches(25, 5) # scale figure to be wider (since there are 2 subplots)
```

PPD & empirical mean[y] tend to the same value: 674.1592491169141 (real mean[y]=675) PPD std[y] we predict is 82.62364471909156 & empirical std[y] is 32.71332926951551 (real std[y]=37.5)



Reflection on what we are observing

- 1. Generally speaking, our PPD is quite reasonable and the result should be intuitive!
 - For few data points, the variance of the PPD is a bit larger than the variance we assumed for the conditional pdf p(y|z), i.e. $\sigma_{y|z}^2 = s^2 = 80^2$.
- 1. As the number of data points increases (see PPD as $N \to \infty$ or play with the figure above by increasing N), then the variance of the PPD $\sigma_y^2 \to s^2 = 80^2$ tends to the variance we assumed for our model: $\sigma_{y|z}^2 = s^2 = 80^2$.
 - This results from our choice of likelihood and prior... Our model was incorrect in both:
 - The hidden rv z is actually a Gaussian distribution, instead of a noninformative Uniform distribution
 - The real conditional pdf is the Dirac "distribution" around value $y = zx + 0.1x^2$, instead of the Gaussian distribution with a mean of $\mu_{y|z} = zx + 0.1x^2$ and a variance of $\sigma_{y|z} = s^2 = 80^2$.

Please keep this in your head:

- (Bayesian) ML is not magic. Every modeling choice you make affects the predictions you get.
- Of course, there are ways of getting "closer" to the truth! We'll take steps in that direction in the remainder of the course.

Next lecture we will redo everything but for a different prior distribution

Consider the same problem, but now starting from a different model:

1. Same observation distribution as before:

$$p(y|z) = \mathcal{N}\left(y|\mu_{y|z} = wz + b, \sigma_{y|z}^2
ight) = rac{1}{C_{y|z}}\mathrm{exp}\left[-rac{1}{2\sigma_{y|z}^2}(y-\mu_{y|z})^2
ight]$$

1. but now assuming a different **prior distribution**: $p(z) = \mathcal{N}\left(z|\tilde{\mu}_z=3, \tilde{\sigma}_z^2=2^2\right)$ In my notation, the superscript $\stackrel{<}{(\cdot)}$ indicates a parameter of the prior distribution.

Notes about the prior distribution we are assuming

- We would have to be very lucky if our "belief" coincided with the "true" distribution of z.
 - Usually, we have beliefs but they are not really true (this is not a comment about religion...).
 - Our hope is that our beliefs are at least reasonable!
- When defining a prior we are making a decision about two things:
 - 1. The distribution.
 - For example, in this exercise we are assuming that the prior is Gaussian (before we had assumed a "noninformative" Uniform prior). In this case we hit the jackpot! But remember that we are cheating here because we already know the actual distribution of *z* is a Gaussian!
 - 2. The parameters of the distribution.
 - For example, in this exercise we are assuming values that are not the true ones! This is normal! As I said, usually we don't know the truth about the "hidden" variable. Most times we don't even know how many hidden variables we have...

See you next class

Have fun!