## MA 413 - Statistics for Data Science

## Solutions to Exercise 2

1. The CDF of Y is

$$F_Y(y) = \begin{cases} 0 & -\infty < y \le 0 \\ y & 0 < y < 1 \\ 1 & 1 \le y < \infty, \end{cases}$$

and the CDF of X is given by  $\int_{-\infty}^{x} f_X(u) du$ , i.e.,

$$F_X(x) = \begin{cases} 0 & -\infty < x \le 0\\ 1 - e^{-\lambda x} & x > 0. \end{cases}$$

We would like  $F_X(x) = F_Y(y)$  for some function  $r: X \to Y$ , or X = r(Y). Particularly, when 0 < y < 1 and x > 0, we let  $y = 1 - e^{-\lambda x}$ , solving for x leading to

$$X = -\frac{1}{\lambda}\ln(1 - Y). \tag{1}$$

Eq. (1) is used to generate X from Y, and the resulting X satisfies the given distribution of  $f_X(x)$ .

**Remark.** As  $1-Y \sim U(0,1)$ , we can replace (1-Y) with Y for simplicity, and X can be alternatively generated with  $X = -\ln(Y)/\lambda$ .

2. (a) The second axiom of probability in the continuous case needs  $\int \int f_{X,Y}(x,y) dx dy = 1$ , which is equivalent to

$$C \int_0^1 \left( \int_0^y x \ y \ dx \right) dy = \frac{C}{8} = 1,$$

so we have C = 8.

(b) We calculate  $f_X(x)$  and  $f_Y(y)$ .

$$f_X(x) = \int f_{X,Y}(x,y) \ dy = 8x \int_x^1 y \ dy = 4x(1-x^2),$$
  
$$f_Y(y) = \int f_{X,Y}(x,y) \ dx = 8y \int_0^y y \ dx = 4y^3,$$

so  $f_{X,Y}(x,y) \neq f_X(x)f_Y(y)$  and X, Y are not independent.

3. (a) In the discrete case, the second axiom states  $\sum_{x=0}^{\infty} f_X(x) = 1$ , so we have

$$\sum_{x=0}^{\infty} f_X(x) = C \sum_{x=0}^{\infty} \frac{\lambda^x}{x!} = Ce^{\lambda} = 1,$$

yielding  $C = e^{-\lambda}$ . Note that the Taylor series for the exponential function  $e^{\lambda}$  is used. As a result,  $f_X(x)$  is PMF of the Poisson distribution.

(b) We compute the probability as follows

$$\Pr(X \le 5) = e^{-\lambda} \sum_{x=0}^{5} \frac{\lambda^x}{x!}.$$

When  $\lambda = 2$ , we compute

$$\Pr(X \le 5) = e^{-2} \left( \frac{2^0}{0!} + \frac{2^1}{1!} + \frac{2^2}{2!} + \frac{2^3}{3!} + \frac{2^4}{4!} + \frac{2^5}{5!} \right)$$
$$= e^{-2} \left( 1 + 2 + \frac{4}{2} + \frac{8}{6} + \frac{16}{24} + \frac{32}{120} \right)$$
$$= \frac{109}{15e^2} \approx 0.9834.$$

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(c) We compute the mean as follows.

$$\mathrm{E}(X) = e^{-\lambda} \sum_{x=0}^{\infty} x \frac{\lambda^x}{x!} = e^{-\lambda} \sum_{x=1}^{\infty} x \frac{\lambda^x}{x!} = = \lambda e^{-\lambda} \sum_{x=1}^{\infty} \frac{\lambda^{x-1}}{(x-1)!} = \lambda e^{-\lambda} \sum_{x=0}^{\infty} \frac{\lambda^x}{x!} = \lambda e^{-\lambda} e^{\lambda} = \lambda.$$

For Var(X), we use formula  $Var(X) = E(X^2) - (E(X))^2$ . Note that  $E(X^2) = E(X(X-1) + X) = E(X(X-1)) + E(X)$ . We compute E(X(X-1)) as follows,

$$E(X(X-1)) = e^{-\lambda} \sum_{x=0}^{\infty} x(x-1) \frac{\lambda^x}{x!} = \lambda^2 e^{-\lambda} \sum_{x=2}^{\infty} \frac{\lambda^{x-2}}{(x-2)!} = \lambda^2 e^{-\lambda} \sum_{x=0}^{\infty} \frac{\lambda^x}{x!} = \lambda^2.$$

To sum up, we obtain the variance,

$$Var(X) = E(X(X - 1)) + E(X) - (E(X))^2 = \lambda^2 + \lambda - \lambda^2 = \lambda.$$

The dispersion of X is Var(X)/E(X) = 1.

4. We compute its mean as

$$\begin{split} \mathbf{E}(X) &= \sum_{x=0}^{n} x \binom{n}{x} p^{x} (1-p)^{n-x} = \sum_{x=1}^{n} x \frac{n!}{x!(n-x)!} p^{x} (1-p)^{n-x} \\ &= np \sum_{x=1}^{n} \frac{(n-1)!}{(x-1)!(n-x)!} p^{x-1} (1-p)^{n-x} \\ &= np \sum_{x=1}^{n} \frac{(n-1)!}{(x-1)!((n-1)-(x-1))!} p^{x-1} (1-p)^{(n-1)-(x-1)} \\ &= np \sum_{x=0}^{n-1} \frac{(n-1)!}{x!(n-1-x)!} p^{x} (1-p)^{n-1-x} \\ &= np(p+(1-p))^{n-1} = np. \end{split}$$

**Remark.** Alternatively, as X is a Binomial random variable, it represents the number of heads in n tosses of a coin, so it can be written as  $X = \sum_{x=0}^{n} X_i$ , where  $X_i = 1$  if the i<sup>th</sup> toss is heads and  $X_i = 0$  otherwise. Then  $E(X_i) = p \times 1 + (1-p) \times 0 = p$ , and  $E(X) = E(\sum_{i=1}^{n} X_i) = \sum_{i=1}^{n} E(X_i) = np$ .

The following is the Matlab code to compare the mean and variance between Binomial and Poisson.

```
lambda = 1/50;
n_trial = 100;
p_success = lambda / n_trial;
pmf_B = makedist('Binomial','N',n_trial,'p',p_success);
E_B = mean(pmf_B)
Var_B = var(pmf_B)

pmf_P = makedist('Poisson','lambda',lambda);
E_P = mean(pmf_P)
Var_P = var(pmf_P)
```

5. Let  $X \sim \text{Poisson}(\lambda)$ . We would like the CDF of X matches that of Y, i.e.,  $F_X(x) = F_Y(y)$ , but we notice that X is discrete whereas Y is continuous, so given a particular  $y \in (0,1)$ , we choose a corresponding x ( $x \ge 1$ ) such that  $F_X(x-1) < F_Y(y) \le F_X(x)$ , or equivalently,

$$e^{-\lambda} \sum_{n=0}^{x-1} \frac{\lambda^n}{n!} < y \le e^{-\lambda} \sum_{n=0}^{x} \frac{\lambda^n}{n!}.$$
 (2)

In other words, X is given by Eq. (2) when the uniform random variable  $Y > F_X(0) = e^{-\lambda}$ , and X = 0 otherwise.

The following is the Matlab code to generate a Poisson random variable from a uniform one.

```
lambda = 2; % Poisson parameter, can be other value
Y = rand % uniform random variable in (0,1)
x = 0; % initialization of the target Poisson random variable
pmf = exp(-lambda);
FX = pmf;
while FX < Y
    x = x + 1;
    pmf = pmf * lambda / x;
    FX = FX + pmf;
end
x</pre>
```

6. We first compute  $\Gamma(2)$  as follows,

$$\Gamma(2) = \int_0^\infty y e^{-y} \ dy = -y e^{-y} \Big|_0^\infty + \int_0^\infty e^{-y} \ dy = 0 - e^{-y} \Big|_0^\infty = 1.$$

To find the distribution of Z = X/Y, we need a joint density function, so we assume X and Y are independent. As a result, we have

$$f_{X,Y}(x,y) = \begin{cases} f_X(x)f_Y(y) = \frac{y}{8}e^{-(x+y)/2} & x > 0, \ y > 0\\ 0 & \text{otherwise.} \end{cases}$$

Given z > 0, we get the CDF of Z as follows,

$$\begin{split} F_Z(z) &= \Pr(Z \leq z) = \Pr(X/Y \leq z) \\ &= \int \int_{\{(x,y): x \leq zy\}} f_{X,Y}(x,y) \ dx \ dy = \int_0^\infty \left( \int_0^{zy} \frac{y}{8} e^{-(x+y)/2} \ dx \right) \ dy \\ &= \frac{1}{8} \int_0^\infty y e^{-y/2} \left( \int_0^{zy} e^{-x/2} \ dx \right) dy = \frac{1}{8} \int_0^\infty y e^{-y/2} \left( -2e^{-x/2} \Big|_0^{zy} \right) dy \\ &= \frac{1}{4} \int_0^\infty y e^{-y/2} \left( -e^{-zy/2} + 1 \right) dy = \frac{1}{4} \int_0^\infty y \left( -e^{-(z+1)y/2} + e^{-y/2} \right) dy \\ &= \frac{y}{4} \left( \frac{2}{z+1} e^{-(z+1)y/2} - 2e^{-y/2} \right) \Big|_0^\infty - \frac{1}{4} \int_0^\infty \frac{2}{z+1} e^{-(z+1)y/2} - 2e^{-y/2} \ dy \\ &= -\frac{1}{4} \int_0^\infty \frac{2}{z+1} e^{-(z+1)y/2} - 2e^{-y/2} \ dy = -\frac{1}{4} \left( -\frac{4}{(z+1)^2} e^{-(z+1)y/2} + 4e^{-y/2} \right) \Big|_0^\infty \\ &= 1 - \frac{1}{(z+1)^2}, \end{split}$$

and  $F_Z(z) = 0$  when  $z \leq 0$ .

We compute the mean of X as

$$E(X) = \int_0^\infty x \frac{1}{2} e^{-x/2} \ dx = -x e^{-x/2} \Big|_0^\infty + \int_0^\infty e^{-x/2} \ dx = \int_0^\infty e^{-x/2} \ dx = -2 e^{-x/2} \Big|_0^\infty = 2,$$

from which we notice  $\int_0^\infty ue^{-u/2} du = 4$ . We will use it in the following computations. To compute the variance of X, we first compute  $\mathrm{E}(X^2)$ ,

$$E(X^{2}) = \int_{0}^{\infty} x^{2} \frac{1}{2} e^{-x/2} dx = -x^{2} e^{-x/2} \Big|_{0}^{\infty} + \int_{0}^{\infty} 2x e^{-x/2} dx = 2 \int_{0}^{\infty} x e^{-x/2} dx = 2 \times 4 = 8,$$

and thus

$$Var(X) = E(X^2) - (E(X))^2 = 8 - 2^2 = 4.$$

In addition, we have  $\int_0^\infty u^2 e^{-u/2} \ du = 16$ . We use it to compute the mean and variance of Y,

$$\begin{split} \mathrm{E}(Y) &= \int_0^\infty y \frac{y}{4} e^{-y/2} \ dy = \frac{1}{4} \int_0^\infty y^2 e^{-y/2} \ dy = \frac{1}{4} \times 16 = 4, \\ \mathrm{E}(Y^2) &= \frac{1}{4} \int_0^\infty y^3 e^{-y/2} \ dy = -\frac{1}{2} y^3 e^{-y/2} \Big|_0^\infty + \frac{1}{2} \int_0^\infty 3 y^2 e^{-y/2} \ dy = \frac{3}{2} \int_0^\infty y^2 e^{-y/2} \ dy = \frac{3}{2} \times 16 = 24, \end{split}$$

and

$$Var(Y) = E(Y^2) - (E(Y))^2 = 24 - 4^2 = 8.$$