

1 Log files from Stata (external links)

[SHARE-1-harmon \(reshape data\)](#)

[SHARE-2-harmon-Part5 \(generation of disease variables\)](#)

2 Sample Selection

```
. tab hacohort wave          if sfull,m

      hacohort: Sample |
      cohort | 2004 wave  2006/07 w  2011/12 w  2013 wave  2015 wave  2017 wave  2019/20 w |      Total
-----+-----
1.Original sample for |      16,430      16,430      16,430      16,430      16,430      16,430      16,430 | 115,010
2.2006 Refreshment sa |      2,448      2,448      2,448      2,448      2,448      2,448      2,448 | 17,136
-----+-----
      Total |      18,878      18,878      18,878      18,878      18,878      18,878      18,878 | 132,146

. tab iwstatr wave          if sfull,m

      r interview status |
      | 2004 wave  2006/07 w  2011/12 w  2013 wave  2015 wave  2017 wave  2019/20 w |      Total
-----+-----
      0.inap. |      8,617      3,754      1,463      1,673      758      244      19 | 16,528
      1.resp, alive |     10,261     13,580     14,961     14,708     14,447     13,759     10,389 | 92,105
      4.nr, alive |          0      1,104      1,233      749      1,137      1,026      2,209 | 7,458
      5.nr, died this wv |          0      279      468      593      603      958      1,061 | 3,962
      6.nr, died prev wv |          0          0      660     1,096     1,895     2,799     3,983 | 10,433
      9.nr, dk if alive or |          0      161      93      59      38      92      1,217 | 1,660
-----+-----
      Total |      18,878      18,878      18,878      18,878      18,878      18,878      18,878 | 132,146

. sum agemin                  if sfull,

      Variable |      Obs      Mean      Std. dev.      Min      Max
-----+-----
      agemin | 132,146  59.61256  6.309033      31      70

. qui log close log
```

3 Choice of Diseases

```
. codebook d_* timetonextdisease2, compact

Variable      Obs Unique      Mean  Min  Max  Label
-----+-----
d_hibp      169620      2 .4831859  0  1  ever had | taking meds for hibp
d_diab      169598      2 .1500018  0  1  ever had | taking meds for diab
d_heart      169596      2 .2217269  0  1  ever had | taking meds for heart
d_lung      169587      2 .0859913  0  1  ever had | taking meds for lung
d_psych      169856      2 .2651246  0  1  ever had | taking meds for psych
d_osteo      169583      2 .0953456  0  1  ever had | taking meds for osteo
d_cancr      169328      2 .082727  0  1  (only) ever had cancr
d_strok      169334      2 .0539053  0  1  (only) ever had strok
d_arthr      169406      2 .3204373  0  1  (only) ever had arthr
d_any      169929      2 .7559922  0  1  any disease
d_miss      169929      9 .0203203  0  8  # miss.diseases
d_count      169261     10 1.758054  0  9  # diseases
d_count_geq2  169261      2 .4889727  0  1  >=2 diseases
d_count_in~x  169261     10 .1953394  0  1  disease index (=count/total diseases)
d_anyatfir~s  399700      2 .6346585  0  1  already has disease at baseline
d_anyever    399700      2 .7874256  0  1  ever reports any disease
d_anyever_g2  400036      2 .4975852  0  1  ever reports having had any disease (g2aging)
timetonext~2  53524     179 54.78008  11 190  time (months) from C to C+1 (or more) diseases
-----+-----

. codebook diff_*,compact // assuming the first disease starts with hibp in the dataset

Variable      Obs Unique      Mean  Min  Max  Label
-----+-----
diff_d_count  85319      8 .3377091 -1  6  1st diff of # of diseases
diff_miss~nt 110881     11 .3521162 -3  7  1st diff of # of diseases: (L(t-2) used if L(t-1) missing) (=adj. for gaps)
diff_d_hibp   85632      2 .0674631  0  1  1st diff ('ever had' | medication) of d_hibp
diff_d_diab   85628      2 .0281684  0  1  1st diff ('ever had' | medication) of d_diab
diff_d_heart  85627      2 .0567344  0  1  1st diff ('ever had' | medication) of d_heart
diff_d_lung   85626      2 .0220494  0  1  1st diff ('ever had' | medication) of d_lung
diff_d_psych  85656      3 .0203488 -1  1  1st diff ('ever had' | medication) of d_psych
diff_d_osteo  85626      2 .0325485  0  1  1st diff ('ever had' | medication) of d_osteo
```

```

diff_d_cancr 85364 2 .0233002 0 1 1st diff ('ever had' | medication) of d_cancr
diff_d_strok 85371 2 .0167856 0 1 1st diff ('ever had' | medication) of d_strok
diff_d_arthr 85430 2 .0697881 0 1 1st diff ('ever had' | medication) of d_arthr
diff_hibper 85467 2 .0667392 0 1 1st diff ('ever had' - raw data) of hibper
diff_diaber 85407 2 .0273865 0 1 1st diff ('ever had' - raw data) of diaber
diff_hearter 85394 2 .0378949 0 1 1st diff ('ever had' - raw data) of hearter
diff_lunger 85371 2 .0207916 0 1 1st diff ('ever had' - raw data) of lunger
diff_psycher 64556 3 .0003563 -1 1 1st diff ('ever had' - raw data) of psycher
diff_osteoe 85355 2 .0087751 0 1 1st diff ('ever had' - raw data) of osteoe
diff_cancrer 85364 2 .0233002 0 1 1st diff ('ever had' - raw data) of cancrer
diff_stroker 85371 2 .0167856 0 1 1st diff ('ever had' - raw data) of stroker
diff_arthrer 85430 2 .0697881 0 1 1st diff ('ever had' - raw data) of arthrer
diff_miss~p 111212 3 .0775006 -1 1 1st diff ('ever had' | medication) (adj. for gaps) of d_hibp
diff_mis~per 111056 2 .0764119 0 1 1st diff (ever had - raw data) (adj. for gaps) of hibp
diff_miss~b 111194 3 .0328705 -1 1 1st diff ('ever had' | medication) (adj. for gaps) of d_diab
diff_mis~ber 110986 2 .0320221 0 1 1st diff (ever had - raw data) (adj. for gaps) of diab
diff_miss~rt 111188 3 .0550869 -1 1 1st diff ('ever had' | medication) (adj. for gaps) of d_heart
diff_mis~ter 110967 2 .0405346 0 1 1st diff (ever had - raw data) (adj. for gaps) of heart
diff_miss~g 111180 3 .0250135 -1 1 1st diff ('ever had' | medication) (adj. for gaps) of d_lung
diff_mis~ger 110943 2 .0242377 0 1 1st diff (ever had - raw data) (adj. for gaps) of lung
diff_miss~h 111299 3 .0148878 -1 1 1st diff ('ever had' | medication) (adj. for gaps) of d_psych
diff_mis~her 93648 3 .0053391 -1 1 1st diff (ever had - raw data) (adj. for gaps) of psych
diff_miss~o 111176 3 .0243488 -1 1 1st diff ('ever had' | medication) (adj. for gaps) of d_osteo
diff_mis~oer 110926 2 .0092494 0 1 1st diff (ever had - raw data) (adj. for gaps) of osteo
diff_miss~cr 110934 2 .0264752 0 1 1st diff ('ever had' | medication) (adj. for gaps) of d_cancr
diff_mi~cner 110934 2 .0264752 0 1 1st diff (ever had - raw data) (adj. for gaps) of cancr
diff_miss~k 110940 2 .0185686 0 1 1st diff ('ever had' | medication) (adj. for gaps) of d_strok
diff_mis~ker 110940 2 .0185686 0 1 1st diff (ever had - raw data) (adj. for gaps) of strok
diff_miss~hr 111016 2 .0769709 0 1 1st diff ('ever had' | medication) (adj. for gaps) of d_arthr
diff_mi~hrer 111016 2 .0769709 0 1 1st diff (ever had - raw data) (adj. for gaps) of arthr

```

```

. sum d_anyatfirstobs if sfull & wave==inw_first & agemin==50 // ppl w/ >1 conditions at baseline

```

Variable	Obs	Mean	Std. dev.	Min	Max
d_anyatfir~s	713	.4950912	.5003269	0	1

```

. sum d_anyatfirstobs if sfull & wave==inw_first & inrange(agemin,50,65) // ppl w/ >1 conditions at baseline

```

Variable	Obs	Mean	Std. dev.	Min	Max
d_anyatfir~s	14,087	.6508128	.4767301	0	1

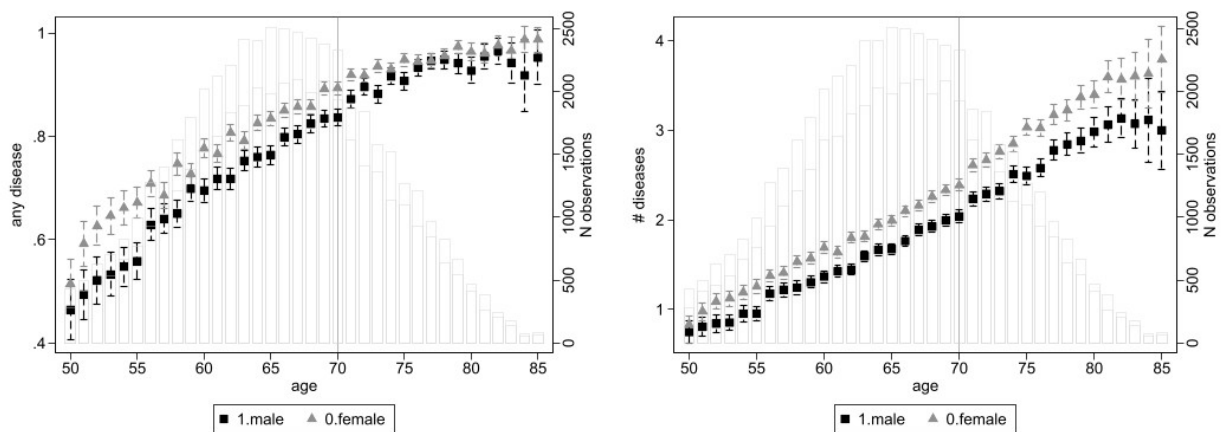
```

. qui log close log

```

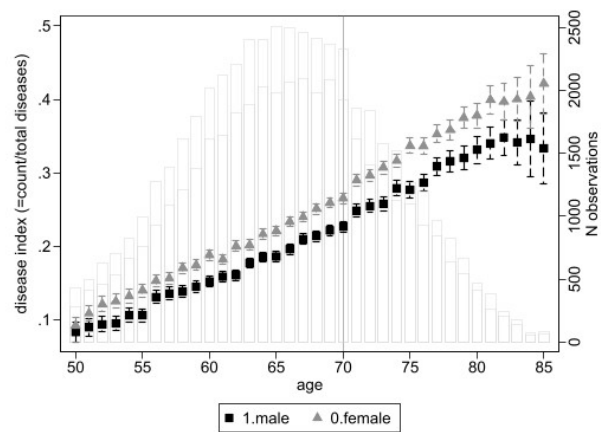
4 Figures and Tables

Figure 1: Variables by age (pooled sample) with 95% CI



(a) d_any (Any Disease)

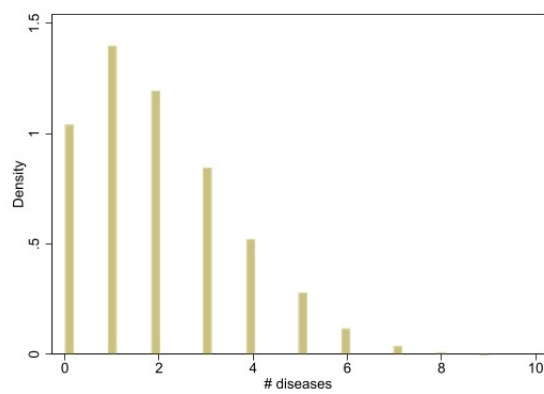
(b) d_count (number of diseases)



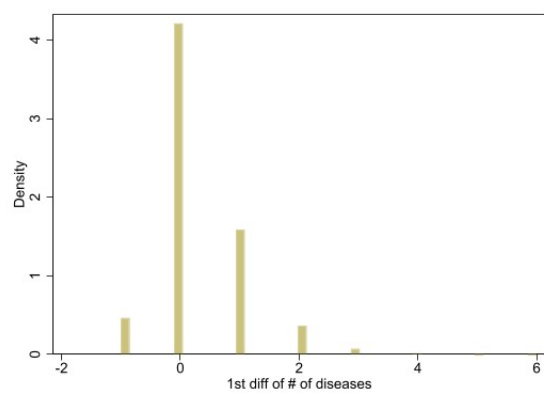
(c) d_count_index (count / considered diseases)

Note: ...

Figure 2: Distributions of variables



(a) d.count



(b) diff_d.count (1st difference in count)

```

-----
name: log
log: C:/Users/User/Documents/GitHub/2-projectMM-SHARE/files/logs/log-g_bytime-cohortmin5.txt
log type: text
opened on: 3 Feb 2024, 23:16:00

. ** xtline by age group **
.
. loc y "d_count"

. loc sample "sfull"

. loc timevar "timesincefirstobs" // timesincefirstobs_yr | time

. collapse (mean) 'y'_mean = 'y' (count) 'y'_freq = 'y' if 'sample'==1 , by(cohortmin5 'timevar')

. xtset cohortmin5 'timevar'

Panel variable: cohortmin5 (unbalanced)
Time variable: timesincefirstobs, 0 to 16
Delta: 1 unit

. loc y2 "'y'_mean"

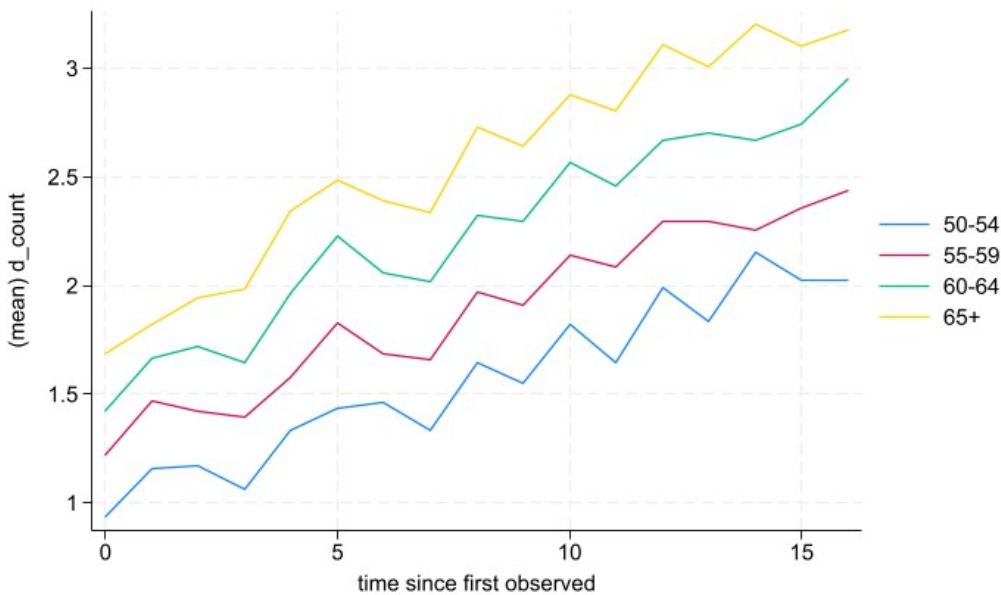
. xtline 'y2', overlay

. gr export "$outpath/fig/supplement/g_by'timevar'-cohortmin5_'sample'_d_count.jpg", replace
(file C:/Users/User/Documents/GitHub/2-projectMM-SHARE/files/fig/supplement/g_bytimesincefirstobs-cohortmin5_sfull_d_count.jpg
> not found)
file C:/Users/User/Documents/GitHub/2-projectMM-SHARE/files/fig/supplement/g_bytimesincefirstobs-cohortmin5_sfull_d_count.jpg
> written in JPEG format

. qui log close log

```

Figure 3: Count of Diseases by age at baseline



```

-----
name: log
log: C:/Users/User/Documents/GitHub/2-projectMM-SHARE/files/logs/log-g_bytime.txt
log type: text
opened on: 3 Feb 2024, 23:07:27

. **graph by time**
. *** Bookmark #1 Can I do this here with xtologit?
. loc timevar "timesincefirststobs"

. loc ctrls "age male"

. loc y "d_count"

. loc ylist "d_count d_count_index timetonextdisease2" //

. loc reg "xtreg" // xtreg | xtologit is very slow; choice may depend on distribution of dependent variable; u
> sing index may indeed be the most appropriate.

. loc sample "sfull"

. foreach sample in "sfull" "sneverdead" "shealthyatfirststobs" {
2. di "Sample is: 'sample' and variables are:"
3. sum 'ylist' 'ctrls' if 'sample'
4. foreach y of local ylist { /*repeat graph for each selected variable*/
5. loc ylabel : variable label 'y' /*uses variable label of y*/
6. qui {
7. **without controls**
. *qui log using "$outpath/logs/log-g_bytime.txt", text replace name(log) // put here if want to close it everytime re
> gardless of running loop
. *di "timevar: 'timevar' | ctrls 'ctrls' | y: 'y' | ylist 'ylist' | sample 'sample'"
. 'reg' 'y' i.'timevar' if 'sample' // without controls
8. margins 'timevar', noestimcheck // atmeans
9. marginsplot, xdimension('timevar') ytitle("ylabel'") note("Notes: This marginsplot uses the following sample: 'sample' a
> nd no controls." "The underlying regression is: 'reg'") // xla(, ang(45)) ytitle("ylabel'")
10. gr export "$outpath/fig/'saveloc'/g_bytime_'sample'_'y'.jpg", replace
11. **with controls**
. 'reg' 'y' i.'timevar' 'ctrls' if 'sample' // with controls
12. margins 'timevar', noestimcheck
13. marginsplot, xdimension('timevar') ytitle("ylabel'") note("Notes: This marginsplot uses the following sample: 'sample' a
> nd the following controls: 'ctrls'." "The underlying regression is: 'reg'") // xla(, ang(45))
14. gr export "$outpath/fig/'saveloc'/g_bytime_'sample'_'y'_withctrls.jpg", replace
15. *qui log close log
. }
16. }
17. }

Sample is: sfull and variables are:

Variable | Obs Mean Std. dev. Min Max
-----+-----
d_count | 88,666 1.99303 1.615312 0 9
d_count_in~x | 88,666 .2214478 .1794791 0 1
timetonext~2 | 35,824 57.22898 34.91639 11 190
age | 88,818 65.91678 7.248936 50 86
male | 127,946 .4714411 .4991857 0 1

--Break--
r(1);

end of do-file

--Break--
r(1);

. do "C:/Users/User/AppData/Local/Temp/STD4fb8_000000.tmp"

. pause on

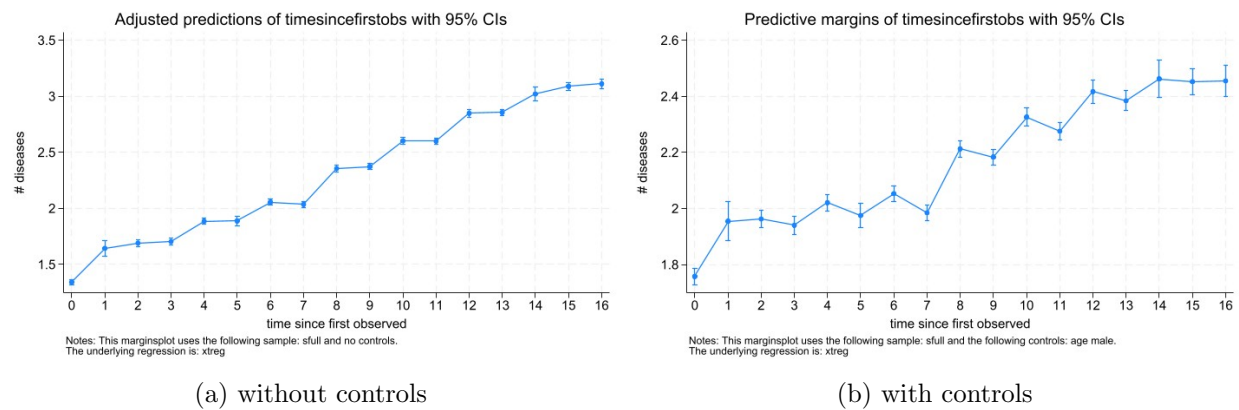
. pause off

. log close _all /*closes all open log files*/
name: log
log: C:/Users/User/Documents/GitHub/2-projectMM-SHARE/files/logs/log-g_bytime.txt
log type: text
closed on: 3 Feb 2024, 23:07:56
-----

```

Figure 4: Predictions of d_count adjusted by time without (left) and with (right - age, gender) control variables (number of diseases by time of followup)

Panel A: Full Sample



Notes: timesincefirstobs refers to year at interview rather than survey year/wave

5 Ordered Response Models

```
-----
name: log
log: C:/Users/User/Documents/GitHub/2-projectMM-SHARE/files/logs/log-t-regd_count-cohort.txt
log type: text
opened on: 4 Feb 2024, 15:41:05
```

```
. *sample 5 // select a ## % random subsample to speed up computation
. loc sample "sfull & d_count<8" // & d_count<8

. loc agectrls "age" // age c.age##c.age, leaving out cross term solves the issue

. loc ctrls "i.male married i.raeduc1 i.cohortmin5" // i.wave

. loc y "d_count"

. sum 'y' 'ctrls' if 'sample'==1 & wave==1
```

Variable	Obs	Mean	Std. dev.	Min	Max
d_count	10,231	1.202815	1.235598	0	7
male					
0.female	10,231	.5417848	.4982753	0	1
1.male	10,231	.4582152	.4982753	0	1
marriedr	10,230	.7985337	.4011151	0	1
raeduc1					
1.less th..	10,231	.4474636	.4972566	0	1
2.upper s..	10,231	.3211807	.4669529	0	1
3.tertiar..	10,231	.2313557	.4217198	0	1
cohortmin5					
50-54	9,468	.2369033	.4252049	0	1
55-59	9,468	.2730249	.4455371	0	1
60-64	9,468	.2554922	.4361605	0	1
65+	9,468	.2345796	.4237582	0	1

```
.
.
. *** Bookmark #1
. drop if agemin<50
(17,451 observations deleted)

. /*** ols (suitable only if assuming count approximates unobserved health reasonably well) ***
> xtreg 'y' 'ctrls' if 'sample'==1 , re
> STOP
> */
.
. *** ologit (wave 1 only) note: this is not considering the panel dimension ***
. eststo m1: ologit 'y' 'agectrls' 'ctrls' if 'sample'==1, vce(robust)
```

Iteration 0: Log pseudolikelihood = -150807.27
Iteration 1: Log pseudolikelihood = -143461.27
Iteration 2: Log pseudolikelihood = -143378.2
Iteration 3: Log pseudolikelihood = -143378.1
Iteration 4: Log pseudolikelihood = -143378.1

Ordered logistic regression	Number of obs = 84,923
	Wald chi2(8) = 14270.45
	Prob > chi2 = 0.0000
Log pseudolikelihood = -143378.1	Pseudo R2 = 0.0493

		Coefficient	Robust std. err.	z	P> z	[95% conf. interval]	
	d_count						
	age	.1149442	.0013085	87.84	0.000	.1123795	.1175089
	male						
	1.male	-.3286282	.0125839	-26.11	0.000	-.3532922	-.3039641

	marriedr		-.1887724	.0150957	-12.51	0.000	-.2183595	-.1591853
	raeduc1							
2.upper secondary or vocational			-.3127262	.0142281	-21.98	0.000	-.3406129	-.2848396
3.tertiary education			-.6689294	.0162643	-41.13	0.000	-.7008069	-.6370519
	cohortmin5							
	55-59		-.1472491	.0182118	-8.09	0.000	-.1829436	-.1115547
	60-64		-.3125041	.0212326	-14.72	0.000	-.3541192	-.270889
	65+		-.5451087	.0255699	-21.32	0.000	-.5952248	-.4949927

	/cut1		5.138158	.0790814			4.983161	5.293155
	/cut2		6.518929	.0799738			6.362183	6.675674
	/cut3		7.54489	.0809125			7.386304	7.703475
	/cut4		8.480449	.0817512			8.32022	8.640679
	/cut5		9.446869	.0825869			9.285002	9.608736
	/cut6		10.57117	.0840808			10.40638	10.73597
	/cut7		12.03472	.0914497			11.85549	12.21396

. *brant, detail // brant only works on ologit; ologit and xtologit are not identical when only 1 time period is used // brant
> does not work with d_count>=8 because of perfect prediction (too few observations fall into this category)

. *** gologit2 ***
. log using "\$outpath/logs/log-t-regd_count-cohort-gologit2.txt", text replace name(gologit2)

name: gologit2
log: C:/Users/User/Documents/GitHub/2-projectMM-SHARE/files/logs/log-t-regd_count-cohort-gologit2.txt
log type: text
opened on: 4 Feb 2024, 15:41:11

. eststo m2: gologit2 'y' 'agectrls' 'ctrls' if 'sample'==1, vce(cluster ID) autofit gamma // cutpoints (intercept) are identical to ologit (but not xtologit)

Testing parallel lines assumption using the .05 level of significance...

Step 1: Constraints for parallel lines imposed for 60.cohortmin5 (P Value = 0.1847)
Step 2: Constraints for parallel lines imposed for 65.cohortmin5 (P Value = 0.3571)
Step 3: Constraints for parallel lines imposed for 55.cohortmin5 (P Value = 0.2113)
Step 4: Constraints for parallel lines imposed for 1.male (P Value = 0.1111)
Step 5: Constraints for parallel lines imposed for 2.raeduc1 (P Value = 0.0523)
Step 6: Constraints for parallel lines are not imposed for
age (P Value = 0.00230)
marriedr (P Value = 0.01091)
3.raeduc1 (P Value = 0.00625)

Wald test of parallel lines assumption for the final model:

(1) [0]55.cohortmin5 - [1]55.cohortmin5 = 0
(2) [0]60.cohortmin5 - [1]60.cohortmin5 = 0
(3) [0]65.cohortmin5 - [1]65.cohortmin5 = 0
(4) [0]1.male - [1]1.male = 0
(5) [0]2.raeduc1 - [1]2.raeduc1 = 0
(6) [0]55.cohortmin5 - [2]55.cohortmin5 = 0
(7) [0]60.cohortmin5 - [2]60.cohortmin5 = 0
(8) [0]65.cohortmin5 - [2]65.cohortmin5 = 0
(9) [0]1.male - [2]1.male = 0
(10) [0]2.raeduc1 - [2]2.raeduc1 = 0
(11) [0]55.cohortmin5 - [3]55.cohortmin5 = 0
(12) [0]60.cohortmin5 - [3]60.cohortmin5 = 0
(13) [0]65.cohortmin5 - [3]65.cohortmin5 = 0
(14) [0]1.male - [3]1.male = 0
(15) [0]2.raeduc1 - [3]2.raeduc1 = 0
(16) [0]55.cohortmin5 - [4]55.cohortmin5 = 0
(17) [0]60.cohortmin5 - [4]60.cohortmin5 = 0
(18) [0]65.cohortmin5 - [4]65.cohortmin5 = 0
(19) [0]1.male - [4]1.male = 0
(20) [0]2.raeduc1 - [4]2.raeduc1 = 0
(21) [0]55.cohortmin5 - [5]55.cohortmin5 = 0
(22) [0]60.cohortmin5 - [5]60.cohortmin5 = 0
(23) [0]65.cohortmin5 - [5]65.cohortmin5 = 0
(24) [0]1.male - [5]1.male = 0
(25) [0]2.raeduc1 - [5]2.raeduc1 = 0
(26) [0]55.cohortmin5 - [6]55.cohortmin5 = 0
(27) [0]60.cohortmin5 - [6]60.cohortmin5 = 0

```
(28) [0]65.cohortmin5 - [6]65.cohortmin5 = 0
(29) [0]1.male - [6]1.male = 0
(30) [0]2.raeduc1 - [6]2.raeduc1 = 0
```

```
chi2( 30) = 46.85
Prob > chi2 = 0.0258
```

An insignificant test statistic indicates that the final model does not violate the proportional odds/ parallel lines assumption

If you re-estimate this exact same model with gologit2, instead of autofit you can save time by using the parameter

```
pl(50b.cohortmin5 55.cohortmin5 60.cohortmin5 65.cohortmin5 0b.male 1.male 1b.raeduc1 2.raeduc1)
```

Generalized Ordered Logit Estimates

```
Number of obs = 84,923
Wald chi2(26) = 6090.51
Prob > chi2 = 0.0000
Pseudo R2 = 0.0497
```

Log pseudolikelihood = -143307.41

```
( 1) [0]55.cohortmin5 - [1]55.cohortmin5 = 0
( 2) [0]60.cohortmin5 - [1]60.cohortmin5 = 0
( 3) [0]65.cohortmin5 - [1]65.cohortmin5 = 0
( 4) [0]1.male - [1]1.male = 0
( 5) [0]2.raeduc1 - [1]2.raeduc1 = 0
( 6) [1]55.cohortmin5 - [2]55.cohortmin5 = 0
( 7) [1]60.cohortmin5 - [2]60.cohortmin5 = 0
( 8) [1]65.cohortmin5 - [2]65.cohortmin5 = 0
( 9) [1]1.male - [2]1.male = 0
(10) [1]2.raeduc1 - [2]2.raeduc1 = 0
(11) [2]55.cohortmin5 - [3]55.cohortmin5 = 0
(12) [2]60.cohortmin5 - [3]60.cohortmin5 = 0
(13) [2]65.cohortmin5 - [3]65.cohortmin5 = 0
(14) [2]1.male - [3]1.male = 0
(15) [2]2.raeduc1 - [3]2.raeduc1 = 0
(16) [3]55.cohortmin5 - [4]55.cohortmin5 = 0
(17) [3]60.cohortmin5 - [4]60.cohortmin5 = 0
(18) [3]65.cohortmin5 - [4]65.cohortmin5 = 0
(19) [3]1.male - [4]1.male = 0
(20) [3]2.raeduc1 - [4]2.raeduc1 = 0
(21) [4]55.cohortmin5 - [5]55.cohortmin5 = 0
(22) [4]60.cohortmin5 - [5]60.cohortmin5 = 0
(23) [4]65.cohortmin5 - [5]65.cohortmin5 = 0
(24) [4]1.male - [5]1.male = 0
(25) [4]2.raeduc1 - [5]2.raeduc1 = 0
(26) [5]55.cohortmin5 - [6]55.cohortmin5 = 0
(27) [5]60.cohortmin5 - [6]60.cohortmin5 = 0
(28) [5]65.cohortmin5 - [6]65.cohortmin5 = 0
(29) [5]1.male - [6]1.male = 0
(30) [5]2.raeduc1 - [6]2.raeduc1 = 0
```

(Std. err. adjusted for 17,415 clusters in ID)

	d_count	Coefficient	Robust std. err.	z	P> z	[95% conf. interval]	
0							
	age	.1155052	.0022321	51.75	0.000	.1111303	.1198801
	male						
	1.male	-.3253926	.0257881	-12.62	0.000	-.3759363	-.274849
	marriedr	-.1266688	.0397224	-3.19	0.001	-.2045233	-.0488143
	raeduc1						
2.upper secondary or vocational		-.3077584	.029004	-10.61	0.000	-.3646052	-.2509116
3.tertiary education		-.6109739	.0414677	-14.73	0.000	-.6922491	-.5296987
	cohortmin5						
	55-59	-.1439795	.0366286	-3.93	0.000	-.2157702	-.0721887
	60-64	-.3077681	.0389323	-7.91	0.000	-.3840739	-.2314623
	65+	-.5442468	.0427478	-12.73	0.000	-.628031	-.4604626
	_cons	-5.245369	.1385054	-37.87	0.000	-5.516835	-4.973904

1							
	age	.112447	.0018557	60.60	0.000	.10881	.1160841
	male						
	1.male	-.3253926	.0257881	-12.62	0.000	-.3759363	-.274849
	marriedr	-.1665718	.0319044	-5.22	0.000	-.2291034	-.1040402
	raeduc1						
2.upper secondary or vocational		-.3077584	.029004	-10.61	0.000	-.3646052	-.2509116
3.tertiary education		-.6650196	.0362973	-18.32	0.000	-.736161	-.5938783
	cohortmin5						
	55-59	-.1439795	.0366286	-3.93	0.000	-.2157702	-.0721887
	60-64	-.3077681	.0389323	-7.91	0.000	-.3840739	-.2314623
	65+	-.5442468	.0427478	-12.73	0.000	-.628031	-.4604626
	_cons	-6.383653	.1178301	-54.18	0.000	-6.614596	-6.15271

2							
	age	.1142367	.0019311	59.16	0.000	.1104517	.1180216
	male						
	1.male	-.3253926	.0257881	-12.62	0.000	-.3759363	-.274849
	marriedr	-.1835188	.0326657	-5.62	0.000	-.2475425	-.1194951
	raeduc1						
2.upper secondary or vocational		-.3077584	.029004	-10.61	0.000	-.3646052	-.2509116
3.tertiary education		-.6986977	.0397992	-17.56	0.000	-.7767026	-.6206927
	cohortmin5						
	55-59	-.1439795	.0366286	-3.93	0.000	-.2157702	-.0721887
	60-64	-.3077681	.0389323	-7.91	0.000	-.3840739	-.2314623
	65+	-.5442468	.0427478	-12.73	0.000	-.628031	-.4604626
	_cons	-7.507797	.1287648	-58.31	0.000	-7.760171	-7.255422

3							
	age	.1166289	.0023114	50.46	0.000	.1120987	.1211591
	male						
	1.male	-.3253926	.0257881	-12.62	0.000	-.3759363	-.274849
	marriedr	-.2516604	.0372694	-6.75	0.000	-.3247071	-.1786138
	raeduc1						
2.upper secondary or vocational		-.3077584	.029004	-10.61	0.000	-.3646052	-.2509116
3.tertiary education		-.7200224	.0483888	-14.88	0.000	-.8148628	-.6251821
	cohortmin5						
	55-59	-.1439795	.0366286	-3.93	0.000	-.2157702	-.0721887
	60-64	-.3077681	.0389323	-7.91	0.000	-.3840739	-.2314623
	65+	-.5442468	.0427478	-12.73	0.000	-.628031	-.4604626
	_cons	-8.554732	.1607131	-53.23	0.000	-8.869724	-8.23974

4							
	age	.1200388	.0031195	38.48	0.000	.1139247	.1261529
	male						
	1.male	-.3253926	.0257881	-12.62	0.000	-.3759363	-.274849
	marriedr	-.3277274	.0473585	-6.92	0.000	-.4205483	-.2349065
	raeduc1						
2.upper secondary or vocational		-.3077584	.029004	-10.61	0.000	-.3646052	-.2509116
3.tertiary education		-.7917869	.0666359	-11.88	0.000	-.9223908	-.661183
	cohortmin5						
	55-59	-.1439795	.0366286	-3.93	0.000	-.2157702	-.0721887
	60-64	-.3077681	.0389323	-7.91	0.000	-.3840739	-.2314623
	65+	-.5442468	.0427478	-12.73	0.000	-.628031	-.4604626
	_cons	-9.698555	.2231467	-43.46	0.000	-10.13591	-9.261196

5							
	age	.1315769	.0047575	27.66	0.000	.1222525	.1409014
	male						
	1.male	-.3253926	.0257881	-12.62	0.000	-.3759363	-.274849
	marriedr	-.3735562	.0714011	-5.23	0.000	-.5134998	-.2336126
	raeduc1						
2.upper secondary or vocational		-.3077584	.029004	-10.61	0.000	-.3646052	-.2509116
3.tertiary education		-1.043973	.1066608	-9.79	0.000	-1.253024	-.8349215
	cohortmin5						
	55-59	-.1439795	.0366286	-3.93	0.000	-.2157702	-.0721887
	60-64	-.3077681	.0389323	-7.91	0.000	-.3840739	-.2314623
	65+	-.5442468	.0427478	-12.73	0.000	-.628031	-.4604626
	_cons	-11.58022	.3488322	-33.20	0.000	-12.26392	-10.89652

6							
	age	.1426234	.0090984	15.68	0.000	.1247909	.1604559
	male						
	1.male	-.3253926	.0257881	-12.62	0.000	-.3759363	-.274849
	marriedr	-.439421	.1250147	-3.51	0.000	-.6844452	-.1943968
	raeduc1						
2.upper secondary or vocational		-.3077584	.029004	-10.61	0.000	-.3646052	-.2509116
3.tertiary education		-1.260081	.2050079	-6.15	0.000	-1.661889	-.8582732
	cohortmin5						
	55-59	-.1439795	.0366286	-3.93	0.000	-.2157702	-.0721887
	60-64	-.3077681	.0389323	-7.91	0.000	-.3840739	-.2314623
	65+	-.5442468	.0427478	-12.73	0.000	-.628031	-.4604626
	_cons	-13.77573	.6728946	-20.47	0.000	-15.09458	-12.45688

Alternative parameterization: Gammas are deviations from proportionality

	d_count	Coefficient	Std. err.	z	P> z	[95% conf. interval]
Beta						
	age	.1155052	.0022321	51.75	0.000	.1111303 .1198801
	male					
	1.male	-.3253926	.0257881	-12.62	0.000	-.3759363 -.274849
	marriedr	-.1266688	.0397224	-3.19	0.001	-.2045233 -.0488143
	raeduc1					
2.upper secondary or vocational		-.3077584	.029004	-10.61	0.000	-.3646052 -.2509116
3.tertiary education		-.6109739	.0414677	-14.73	0.000	-.6922491 -.5296987
	cohortmin5					
	55-59	-.1439795	.0366286	-3.93	0.000	-.2157702 -.0721887
	60-64	-.3077681	.0389323	-7.91	0.000	-.3840739 -.2314623
	65+	-.5442468	.0427478	-12.73	0.000	-.628031 -.4604626
Gamma_2						
	age	-.0030582	.0017897	-1.71	0.087	-.0065659 .0004495
	marriedr	-.039903	.0310121	-1.29	0.198	-.1006855 .0208795
	raeduc1					
3.tertiary education		-.0540457	.0308482	-1.75	0.080	-.1145071 .0064156
Gamma_3						
	age	-.0012685	.0022521	-0.56	0.573	-.0056826 .0031456
	marriedr	-.05685	.0392856	-1.45	0.148	-.1338485 .0201484
	raeduc1					
3.tertiary education		-.0877238	.040665	-2.16	0.031	-.1674256 -.0080219
Gamma_4						

	age		.0011237	.0027249	0.41	0.680	-.004217	.0064644
	marriedr		-.1249917	.0457627	-2.73	0.006	-.2146849	-.0352984
	raeduc1							
	3.tertiary education		-.1090485	.0510318	-2.14	0.033	-.209069	-.009028

Gamma_5								
	age		.0045336	.0035035	1.29	0.196	-.0023332	.0114003
	marriedr		-.2010586	.0554595	-3.63	0.000	-.3097573	-.09236
	raeduc1							
	3.tertiary education		-.180813	.0693835	-2.61	0.009	-.3168021	-.044824

Gamma_6								
	age		.0160717	.0050633	3.17	0.002	.0061478	.0259957
	marriedr		-.2468874	.0773667	-3.19	0.001	-.3985234	-.0952514
	raeduc1							
	3.tertiary education		-.4329989	.1087106	-3.98	0.000	-.6460678	-.2199301

Gamma_7								
	age		.0271182	.0092786	2.92	0.003	.0089325	.0453039
	marriedr		-.3127522	.1286914	-2.43	0.015	-.5649827	-.0605217
	raeduc1							
	3.tertiary education		-.6491074	.2060489	-3.15	0.002	-1.052956	-.2452589

Alpha								
	_cons_1		-5.245369	.1385054	-37.87	0.000	-5.516835	-4.973904
	_cons_2		-6.383653	.1178301	-54.18	0.000	-6.614596	-6.15271
	_cons_3		-7.507797	.1287648	-58.31	0.000	-7.760171	-7.255422
	_cons_4		-8.554732	.1607131	-53.23	0.000	-8.869724	-8.23974
	_cons_5		-9.698555	.2231467	-43.46	0.000	-10.13591	-9.261196
	_cons_6		-11.58022	.3488322	-33.20	0.000	-12.26392	-10.89652
	_cons_7		-13.77573	.6728946	-20.47	0.000	-15.09458	-12.45688

```
. qui log close gologit2
```

```
.
. *** xt-ordered logit ***
. eststo m3: xtologit 'y' 'agectrls' 'ctrls' if 'sample'==1, // vce(cluster ID) // -vce(cl ID)- is equivalent to -robust
> -
```

Fitting comparison model:

```
Iteration 0: Log likelihood = -150807.27
Iteration 1: Log likelihood = -143461.27
Iteration 2: Log likelihood = -143378.2
Iteration 3: Log likelihood = -143378.1
Iteration 4: Log likelihood = -143378.1
```

Refining starting values:

```
Grid node 0: Log likelihood = -128397.9
```

Fitting full model:

```
Iteration 0: Log likelihood = -128397.9
Iteration 1: Log likelihood = -112403.36
Iteration 2: Log likelihood = -108174.54
Iteration 3: Log likelihood = -107487.08
Iteration 4: Log likelihood = -107414.46
Iteration 5: Log likelihood = -107413.69
Iteration 6: Log likelihood = -107413.69
```

```
Random-effects ordered logistic regression
Group variable: ID
```

```
Number of obs    = 84,923
Number of groups = 17,415
```

```
Random effects u_i ~ Gaussian
```

```
Obs per group:
    min = 1
    avg = 4.9
    max = 7
```

```
Integration method: mvaghermite
```

```
Integration pts. = 12
```

Log likelihood = -107413.69 Wald chi2(8) = 24482.47
 Prob > chi2 = 0.0000

	d_count	Coefficient	Std. err.	z	P> z	[95% conf. interval]	
	age	.327314	.0021525	152.06	0.000	.3230952	.3315329
	male						
	1.male	-.6405521	.0614861	-10.42	0.000	-.7610627	-.5200415
	marriedr	-.4390972	.0407436	-10.78	0.000	-.5189531	-.3592412
	raeduc1						
2.upper secondary or vocational		-.7380963	.0692194	-10.66	0.000	-.8737639	-.6024287
3.tertiary education		-1.706022	.0822345	-20.75	0.000	-1.867198	-1.544845
	cohortmin5						
	55-59	-.5382597	.0871279	-6.18	0.000	-.7090273	-.3674921
	60-64	-1.142239	.0887945	-12.86	0.000	-1.316273	-.9682053
	65+	-1.932319	.0927441	-20.83	0.000	-2.114094	-1.750543
	/cut1	15.2045	.14892			14.91262	15.49638
	/cut2	18.59412	.1554931			18.28936	18.89888
	/cut3	21.13106	.1606267			20.81624	21.44588
	/cut4	23.3672	.1652919			23.04323	23.69116
	/cut5	25.49624	.1699941			25.16306	25.82942
	/cut6	27.68228	.1757473			27.33782	28.02673
	/cut7	30.06939	.1862123			29.70442	30.43436
	/sigma2_u	14.7093	.2185309			14.28716	15.14391

LR test vs. ologit model: chibar2(01) = 71928.82 Prob >= chibar2 = 0.0000

```
. *margins, at(age=(50(2)80))
. *margins, dydx('marginsvar') // at(male = (1 0))
. *marginsplot
. *
. *      predict p0 p1 p2 p3 p4 p5 p6 p7, pr // p9
. *      sum      p?
. *mtable, dydx(raeduc1) // at(male = (0 1) raeduc1 = (1 2 3)) // at(male = (0 1) ) // raeduc1 = (0 1 2 ))
.
.
. log close log
.   name: log
.   log: C:/Users/User/Documents/GitHub/2-projectMM-SHARE/files/logs/log-t-regd_count-cohort.txt
.   log type: text
. closed on: 4 Feb 2024, 15:47:08
```

**6 THE PART BELOW CONTAINS PRELIMINARY RESULTS.
IT IS SYNCED TO GITHUB AND CAN BE DISREGARDED
UNTIL ADDED ABOVE. — For the team to edit/add notes
to the part above, use “7-graphsAndTables.tex”**