- oREV: an Item Response Theory based open receptive vocabulary task for 3 to 8-year-old
- children
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23 Abstract

Individual differences in early language abilities are an important predictor of later life outcomes. High-quality, easy-access measures of language abilities are rare, especially in 25 the preschool and primary school years. The present study describes the construction of a 26 new receptive vocabulary task for children between 3 and 8 years of age. The task was 27 implemented as a browser-based web application, allowing for in-person as well as remote 28 data collection via the internet. Based on data from N = 581 German-speaking children, we estimated the psychometric properties of each item in a larger initial item pool via Item Response Modeling. We then applied an automated item selection procedure to select an optimal subset of items based on item difficulty and discrimination. The so-constructed task has 22 items and shows excellent psychometric properties with respect to reliability, stability and convergent and discriminant validity. The construction, implementation, and item selection process described here makes it easy to extend the task or adapt it to different languages. All materials and code are freely accessible to interested researchers. The task can be used via the following website: https://ccp-odc.eva.mpg.de/orev-demo/. 37 Keywords: language development, vocabulary, individual differences, Item Response 38 Models

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Introduction

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Individual differences in language abilities are early emerging, stable across 43 development, and predictive of a wide range of psychological outcome variables including 44 cognitive abilities, academic achievement, and mental health (Bornstein, Hahn, Putnick, & Pearson, 2018; Marchman & Fernald, 2008; Morgan, Farkas, Hillemeier, Hammer, & Maczuga, 2015; Schoon, Parsons, Rush, & Law, 2010; Walker, Greenwood, Hart, & Carta, 1994). From a methodological perspective, high-quality, easy-access measures of language abilities are therefore central to both basic and applied research on individual differences in language abilities. Developing such measures is very time and resource intensive and, as a consequence, few exist. In this paper, we describe the construction of a new receptive vocabulary task for German-speaking children. Its theory-driven item generation process makes it linguistically credible. Its psychometric grounding in Item Response Theory (IRT) equips it with the advantages and properties of IRT models (Embretson & Reise, 2013). Its web-based design and implementation make the measure easy to adapt and administer in 55 different settings (in-person or remote) and thereby facilitates the scaling of data collection. Language has many domains and aspects that can be focused on when assessing 57 individual differences across children. One particular productive approach has been the study of children's vocabulary skills, that is, their knowledge of word-object mappings. This skill can be most effectively assessed, for example, by asking children to name an object (expressive vocabulary) or pick out an object that matches a word they just heard (receptive vocabulary). Children with larger vocabularies are taken to have advanced language skills more broadly. This assumption seems to be justified in light of strong correlations between vocabulary size and other language measures such as grammatical (Hoff, Quinn, & Giguere, 2018; e.g., Moyle, Weismer, Evans, & Lindstrom, 2007) or

- narrative skills (Bohnacker, Lindgren, & Öztekin, 2021; Fiani, Henry, & Prévost, 2021; Lindgren & Bohnacker, 2022; Tsimpli, Peristeri, & Andreou, 2016). Vocabulary skills have also been used as an indicator of developmental language disorders more broadly (Spaulding, Hosmer, & Schechtman, 2013). Finally, many of the predictive relations found 69 for early language skills mentioned above are based on vocabulary measures (Bleses, Makransky, Dale, Højen, & Ari, 2016; Golinkoff, Hoff, Rowe, Tamis-LeMonda, & Hirsh-Pasek, 2019; Pace, Alper, Burchinal, Golinkoff, & Hirsh-Pasek, 2019; Pace, Luo, Hirsh-Pasek, & Golinkoff, 2017). This set of findings underlines the importance of high-quality vocabulary measures. A range of measures exists to assess vocabulary skills in children. For very young 75 children (up to 3 years), a frequently used measure is the MacArthur–Bates Communicative Development Inventories (CDIs) (Fenson et al., 2007). Parents are provided with a list of 77 words and are asked to check those the child understands and/or produces. The CDI exists in different forms (e.g., Makransky, Dale, Havmose, & Bleses, 2016; Mayor & Mani, 2019), including an online version (DeMayo et al., 2021), and has been adapted to many different languages (see Frank, Braginsky, Yurovsky, & Marchman, 2021). Due to concentrated collaborative efforts, data from thousands of children learning dozens of languages has been pooled in centralized repositories (Frank, Braginsky, Yurovsky, & Marchman, 2017; Jørgensen, Dale, Bleses, & Fenson, 2010). As such, the CDI provides a positive example of a high-quality, easy-access measure that is heavily used in both basic and applied research. 85 However, the CDI is best suited for children in the first two years of life. From 2 86
- years onward, children are usually tested directly. Receptive vocabulary assessment is often part of standardized tests of cognitive abilities (e.g., Bayley, 2006; Gershon et al., 2013; Wechsler & Kodama, 1949). In addition, a range of dedicated measures exist for English (e.g., Dunn & Dunn, 1965; Dunn, Dunn, Whetton, & Burley, 1997; Golinkoff et al., 2017), German (Glück & Glück, 2011; Kauschke & Siegmüller, 2002; Kiese-Himmel, 2005; Lenhard, Lenhard, Segerer, & Suggate, 2015) and other languages.

Yet, from a researcher's perspective, these existing measures are often problematic for several reasons. Because they are standardized and normed instruments, using them comes with substantial licensing costs. For the same reasons, the corresponding materials are not openly available, which makes it difficult to expand or adapt them to different languages. Most measures also rely on in-person, paper-pencil testing, which makes large-scale data collection inefficient. Whenever more portable, computerized versions exist, they come with additional costs. As a consequence, nothing comparable to the collaborative research infrastructure built around the CDI exists for vocabulary measures for older children.

The development of so-called Cross-linguistic Lexical Tasks [CLTs; Haman, 101 Łuniewska, and Pomiechowska (2015)] constitutes a promising framework that might help 102 to overcome these issues. CLTs are picture-choice and picture-naming tasks aimed at 103 assessing receptive and expressive knowledge of nouns and verbs in children up to five 104 years. In a collaborative effort involving more than 25 institutions, versions for dozens of 105 different languages have been developed following the same guiding principles 106 (Armon-Lotem, 2015; Haman et al., 2017, 2015). In addition to cross-linguistic studies 107 with monolingual children, this procedure makes CLTs ideally suited to assess multilingual 108 preschool children. The tasks and the materials are not commercially licensed and can thus 109 be freely used for research purposes. 110

Despite these many positive characteristics, CLTs are limited in two important ways.

First, they were designed for children between 3 and 5 years and consequently show ceiling

effects for older children in this age range (Haman et al., 2017). This greatly limits their

usefulness in research across the preschool years. Second, and maybe more important,

CLTs have been developed following clear linguistic guidelines – but without a strict

psychometric framework<sup>1</sup>. As a consequence, it is unclear how the different items relate to

the underlying construct (e.g., vocabulary skills). We do not know which items

<sup>&</sup>lt;sup>1</sup> The same applies to most other vocabulary measures used in developmental research.

discriminate between varying ability levels and are therefore particularly diagnostic e.g., at different ages. Items could also be biased and show differential measurement properties in 119 relevant subgroups (e.g. girls and boys). In addition, some items might be simply 120 redundant in that they measure the underlying construct in the same way. Such 121 characteristics could make the task unnecessarily long. Modern psychometric approaches 122 like Item Response Theory (IRT) (Kubinger, 2006; Lord, 2012) allow researchers to 123 adequately model the probabilistic relationship between the items of a test and the 124 underlying latent trait. In addition, it can be empirically tested how well the individual 125 items are suited to capture a latent dimension and what psychometric properties are 126 associated with the specific test. This focus allows for evaluating the quality and usefulness 127 of each item and thereby provides a solid psychometric basis for constructing efficient and 128 high-quality tasks. In combination with a computerized implementation, IRT allows for adaptive testing during which participants are selectively presented with highly informative items given their (constantly updated) estimated level of ability. However, IRT-based task construction requires a higher initial investment: it takes a large item pool and large 132 sample sizes to estimate the item parameters that guide the selection of the best items. 133

## The current study

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Our goal was to develop a theoretically grounded, high-quality, easy-access measure
of receptive vocabulary skills for German-speaking children between 3 and 8 years of age.
For this purpose, we built on the existing CLT but substantially expanded the item pool.
We implemented the task as a browser-based web application, which made it highly
portable and allowed us to test a large sample of children online. Next, we used IRT to
estimate measurement characteristics of each item in the pool. We then developed an
algorithm that used these characteristics to automatically select a smaller subset of items
for the final task. The implementation infrastructure and construction process we describe
here make the task easy to share with interested researchers and practitioners and also

provide clear guidance on how to further adapt it to different languages.

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### Access to data and materials

The datasets generated during the current study as well as the analysis code are
available in the following repository: https://github.com/ccp-eva/orev. The task (after
item selection) can be accessed via the following link:
https://ccp-odc.eva.mpg.de/orev-demo/. Finally, the source code, pictures and sound files
used in the task can be accessed via the following repository:
https://github.com/ccp-eva/orev-demo.

## Item-pool generation

The initial item pool consisted of 32 items taken with permission from the German 153 CLT (Haman et al., 2017, 2015) and 20 new items. The addition of new items was 154 necessary due to ceiling effects for monolingual 5-year-olds in the previous version. New 155 items were generated in line with the construction of the original CLT in a stepwise 156 process. Each item consists of a target word and three distractors. To select target words, we first compiled a list of age-of-acquisition ratings for 3.928 German words from various sources (Birchenough, Davies, & Connelly, 2017; Łuniewska et al., 2019; A. Schröder, 159 Gemballa, Ruppin, & Wartenburger, 2012). From this list, we selected 20 words based on 160 the following criteria: words should refer to concepts that could easily and unambiguously 161 be depicted in a drawing, age-of-acquisition ratings should be spread equally between six 162 and ten years of age. We also computed (semantic) complexity indices for each word (see 163 Haman et al., 2017). This metric, however, did not reflect a dimension that was relevant 164 for item selection. 165

The so-selected 20 words served as additional target words in the item pool (total of 52 items). For each target word, we selected three distractors. The first distractor was

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unrelated to the target word but was chosen to have a comparable rated age-of-acquisition. 168 The second distractor was semantically related to the target word (e.g., ruin – fortress; elk 169 - mammoth). The third distractor was phonetically similar to the target. For example, the 170 initial part was substituted, while the rest of the word was kept similar (e.g., Gazelle [eng.: 171 gazelle] – Libelle [eng.: dragonfly]). The complete list of targets and distractors can be 172 found in the associate online repository. Finally, an artist (same as for the original CLT 173 items) drew pictures representing all target and distractor words. This procedure ensured 174 that the original CLT and the newly generated items formed a homogeneous item pool. 175

### Task design and implementation

The task was programmed in JavaScript, CSS, and HTML and presented as a website that could be opened in any modern web browser. In addition to participants' responses, we recorded webcam videos<sup>2</sup>. Both files were sent to a local server after the study was finished. The task started with several instruction pages that explained to parents the task and how they should assist their child if needed.

On each trial (see Figure 1), participants saw four pictures and heard a verbal 182 prompt (pre-recorded by a native German speaker) asking them to select one of the 183 pictures (prompt: "Zeige mir [target word]"; eng.: "Show me [target word]"). The verbal 184 prompt was automatically played at the beginning of each trial. The prompt could also be 185 replayed by clicking on a loudspeaker button if needed. Pictures could only be selected 186 once the verbal prompt finished playing. Selected pictures were marked via a blue frame. 187 Participants moved on to the next trial by clicking on a button at the bottom of the screen. 188 If children could not select the pictures themselves (via mouse click or tapping on the 189 touch screen), they were instructed to point to the screen and parents should select the 190 pointed-to picture. 191

<sup>&</sup>lt;sup>2</sup> Due to access rights issues, webcam recording was not possible when participants used iOS devices.

The positioning of the target was counterbalanced across four positions (upper/lower 192 and left/right corners) according to three rules: (1) the target picture appeared equally 193 often in each position; (2) the target picture could not appear in the same position in more 194 than three consecutive trials; (3) the target picture appeared in each position at least once 195 across seven subsequent trials. Distractors were distributed across the remaining three 196 positions so that each distractor type (i.e., unrelated, phonological, semantic) appeared 197 equally often in each position across trials. We generated two versions of the task with 198 different item orders. Each order was created so that trial number and age-of-acquisition 199 ratings were correlated with r = .85. This would make later trials more difficult, but not 200 perfectly so. 201

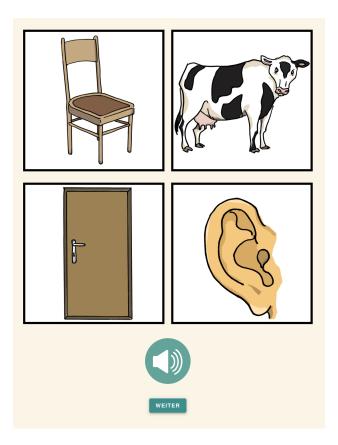


Figure 1. Screenshot of the task. On each trial, participants heard a word and were asked to pick out the corresponding picture. Verbal prompts could be replayed by pressing the loudspeaker button.

202 Item selection

The goal of the item selection process was to find a subset of high-quality items 203 necessary to measure vocabulary skills on an individual level. As a first step, we collected 204 data for the full 52-item task from more than 500 children in the target age range. Next, 205 we fit a Rasch (1PL) and a 2PL IRT model to the data to estimate parameters of interest 206 for each item which we used during the item selection process. We used a simulated 207 annealing process (Kirkpatrick, Gelatt Jr., & Vecchi, 1983) to simultaneously determine the 208 size of the reduced task and to select the best items. Our goal was to construct a reduced 200 task that a) included items of varying difficulty and b) fit the Rasch model so that an 210 individual's test score (number of solved items) is a sufficient statistic and the task is easy 211 to use. After selecting items and constructing the new task, we conducted visual model 212 checks, investigated differential item functioning (DIF) when the data was split either by 213 sex or by trial order, and assessed reliability. Data collection was pre-registered at https://osf.io/qzstk. The pre-registered sample size was based on recommendations found 215 in the literature (Morizot, Ainsworth, & Reise, 2007). However, these authors emphasize that the necessary sample size depends very much on the complexity of the model and that recommendations should be treated with caution. 218

# 219 Participants

Participants were recruited via a database of children living in Leipzig, Germany,
whose parents volunteered to in participate studies on child development and who
additionally indicated interest in participating in online studies. Leipzig is an
industrialized, urban Central-European city with approximately 600,000 inhabitants. The
city-wide median individual monthly net income in 2021 was ~ 1,600€. Children mostly
live in nuclear two-generational families. Socioeconomic status was not formally recorded,
although the majority of families come from mid to high socioeconomic backgrounds with

high levels of parental education.. Furthermore, it is very likely that selective responding skewed the sample towards highly motivated and interested families. Parents received an email with a short study description and a personalized link. After one week, parents received a reminder if they had not already taken part in the study. Response rate to invitations was  $\sim 50\%$ . The final sample included a total of 581 children (n = 307 girls) with a mean age of 5.63 (range: 3.01 - 7.99). Participants were randomly assigned to one of the two item orders. Data was collected between February and May 2022.

## Descriptive results

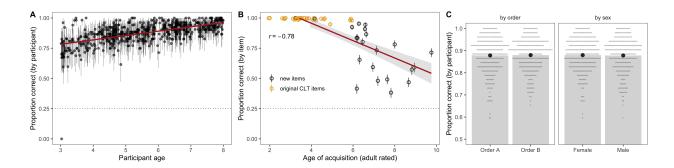


Figure 2. Descriptive results of the task. A: Proportion of correct responses (with 95% CI) for each participant by age. B: Proportion of correct responses (with 95% CI) for each item by rated age-of-acquisition of the target word. C: Proportion of correct responses (with 95% CI) by trial order (left) and sex (right).

On a participant level, performance in the full task (52 items) steadily increased with age (Figure 2A). On an item level, performance was above chance (25%) for all items. Furthermore, the average proportion of correct responses was negatively correlated with age-of-acquisition ratings (Figure 2B) and positively correlated (r = 0.31; 95% CI = 0.02 – 0.55) with the normalized frequency of the word in children's books reported in the childLex corpus (S. Schröder, Würzner, Heister, Geyken, & Kliegl, 2015). Figure 2B also shows the ceiling effect for the original CLT items found in Haman et al. (2017). These descriptive results replicate well-known results in the literature and emphasize the added

value of the newly developed items. Figure 2C shows that there were – on average – no 243 differences between participants who received order A and order B nor between female and 244 male participants. This result suggests that these grouping variables are suitable to 245 investigate differential item functioning (see below). 246

### Item response modeling

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IRT models were implemented in a Bayesian framework in R using the brms package 248 (Bürkner, 2017, 2019). Given the binary outcome of the data, we used logistic 249 mixed-effects models to predict the probability of a correct answer based on the 250 participant's latent ability and item characteristics. We fit two models: a Rasch (1PL) 251 model and a Birnbaum (2PL) model. The main difference between these two models lies in 252 their assumption about item discrimination, that is, how the probability of solving an item 253 changes with ability levels. While the Rasch model assumes that the rate of change (i.e. the slope of the logistic curve) is the same for all items, the 2PL model estimates a 255 separate discrimination parameter for each item. Only when items have similar 256 discrimination parameters is the sum score a sufficient statistic. Given the structure of the 257 task (selecting one out of four pictures), both models had a fixed guessing rate of 0.25<sup>3</sup>. All 258 models converged properly according to visual inspection (e.g., traceplots) and convergence 259 diagnostic measures (e.g. Rhat close to 1, Vehtari, Gelman, Simpson, Carpenter, & 260 Bürkner, 2021). For details about prior and MCMC settings, please see the analysis script 261 in the associated online repository. 262

For each item, we computed the following parameters to be used during the item 263 selection process: Difficulty (parameterized as easiness, i.e., the additive inverse of difficulty) according to the Rasch model, In- and Outfit based on the Rasch model and item discrimination according to the 2PL model. Difficulty estimates represent the level of

<sup>&</sup>lt;sup>3</sup> In the Rasch model, the number of solved items is still a sufficient statistic for an individual's ability when there is a fixed guessing rate (see Jiao, 2022).

ability (point on the latent dimension) for which the probability to solve an item is .5. Inand Outfit are calculated based on the deviation of a person's response to an item from the
response predicted by the model according to their level of ability and item difficulty. As
such, they reflect how well the Rasch model captured the responses to a particular item.
As noted above, item discrimination parameters in the 2PL model influence the rate at
which the probability of solving an item changes given ability levels. In the next section,
we describe how we used these parameters to select items.

#### 274 Automated item selection

The item selection process focused on selecting a smaller subset of items that fit the 275 Rasch model and allow for precise measurement at different levels of receptive vocabulary 276 ability. Only when items fit the Rasch model is the number of solved items a sufficient 277 statistic for an individual's ability. Being able to use the sum score – instead of estimating 278 person parameters via a model – makes the task very easy to use. For this purpose, we 279 defined an objective function that captured three important characteristics that the items 280 of any subset should have. First, items should be equally spaced across the latent ability 281 space. This characteristic ensures that the task is suited for different ability levels and thus 282 for a broader range of ages. We quantified the spread of any given subset as the standard 283 deviation of the distance (in easiness estimates) between adjacent items. Lower values 284 indicate smaller distances and thus an overall more equal spacing. Second (and third), 285 items should have In- and Outfit values close to 1. In- and outfit values that deviate from 1 286 indicate over- or under-fitting and suggest that the respective item is not in line with the Rasch model; conversely, the smaller the value, the better. Finally, items should have similar discrimination parameters according to the 2PL model. The Rasch model assumes that all items have the same discrimination, and thus selecting items with similar discrimination parameters in the 2PL model ensures a better fit of the Rasch model. We 291 quantified this aspect as the variance of discrimination parameters of a given subset of

items. Lower variances indicate more similar discrimination parameters and a better fit of
the Rasch model.

Next, we multiplied/divided these values by constants to put them on a similar numeric scale and to emphasize some aspects over others. We put special emphasis on Inand Outfit values (to select items that conform to the Rasch model) as well as on the equal spacing of items across the latent dimension (to select items suitable for different ages).

Details and data simulations can be found in the analysis script in the associated repository. Finally, we defined the objective function as the sum of the scaled parameters.

We used simulated annealing (Kirkpatrick et al., 1983) to find the optimal items for 301 any given subset size. This process randomly explores the large space of possible subsets, 302 starting from a randomly selected initial subset. Then, it proposes small random changes 303 by exchanging some items in the subset under consideration with others outside it. If such 304 a change increases the value of the objective function, the proposal is accepted, and the 305 improved subset is taken as the new starting point for subsequent proposals. However, to 306 avoid the process getting trapped in local optima, proposals that decrease the value of the 307 objective function may also be accepted, but probabilistically. The probability that a 308 proposal decreasing the objective function is accepted depends upon a parameter called 309 "temperature", which is gradually reduced from a high initial value to a lower value over 310 the course of the simulation. During the "hot" early phase, the process explores the space 311 relatively freely, accepting decreasing proposals often enough to allow it to move between local optima separated by less well-performing subsets, facilitating the discovery of global 313 optima. In the later "cool" phases, the process slowly converges to a strict "hill climbing" 314 search that accepts only increasing proposals, resulting in careful fine-tuning of the best 315 subset discovered in the hot phase. 316

We applied simulated annealing to subsets ranging from 10 to 40 items. For each (optimal) subset, we fit a Rasch model, a 2PL model and compared them using Bayesian

approximate leave-one-out cross-validation (Vehtari, Gelman, & Gabry, 2017) based on 319 differences in expected log posterior density (ELPD) estimates and the associated standard 320 error (SE). This method of comparison balances between fit to the data and out-of-sample 321 predictive accuracy and thereby adjusts for model complexity. Therefore, it favors neither 322 underfitting (because predictions would be too rigid) nor overfitting (because predictions 323 would only match the sample). We considered models to be equivalent up to a point when 324 the ELPD in favor of the 2PL model exceeded two times the standard error of the 325 difference. This rule of thumb is based on suggestions in the literature but is by no means 326 a hard and fast cut-off (Sivula, Magnusson, & Vehtari, 2020). In addition, we also 327 computed the correlation between performance based on the subset and the full task. 328

Our goal was to to find the optimal size and items for the subtest. Figure 3A 329 visualizes the model comparison ratio and shows that the fit of the Rasch model compared 330 to the 2PL model substantially decreases for subsets with more than 24 items. Figure 3B 331 visualizes the correlation between the subtest and the full task, both across all individuals 332 and separate by age and sex. Even though correlations were generally high, they reached a 333 plateau at around 20 items. Based on these results, we decided for 22 items as the size of 334 the subtest. For 22 items the ELPD difference (in favor of the 2PL model) ranged from 335 -0.86 (SE of difference = 2.81) to -2.77 (SE of difference = 2.92). This suggests that the two models were more or less equivalent at that point and that freely estimating the 337 discrimination parameters did not substantially improve the model fit. Thus, the Rasch 338 model provides a good absolute fit for the 22 selected items. Even though a smaller subtest 339 would have been justifiable (e.g. 20 or even 18), we decided to include more items to allow for more precise individual level measurements. We acknowledge, however, that this 341 decision is to some extent arbitrary. 342

When running the simulated annealing procedure for 22 items 100 times, it always returned the same item selection. We, therefore, chose this subset of items for the reduced task. The so-constructed task correlated highly with the full task, both across participants and when the data was split by age group and sex (see Figure 3B)<sup>4</sup>. The selection procedure via the simulated annealing algorithm ensured that the items were equally spread across the latent dimension (see Figure 3C for item characteristic curves).

### Differential item functioning

Next, we fit two additional Rasch models in which we estimated separate difficulty 350 parameters for two subgroups; one for sex (male and female) and one for the order in which 351 items were presented (order A or order B). This allowed us to assess the absolute fit of the 352 Rasch model and to assess differential item functioning (DIF, see Bürkner, 2019). DIF 353 refers to situations where items show differential characteristics for subgroups that 354 otherwise have the same overall score (Holland & Wainer, 2012). If the Rasch model fits 355 the data well and no item shows DIF, the estimates based on the two subgroups should be 356 very similar. Figure 4 shows that this was clearly the case for all items, no matter if the 357 data was split by test order or sex. As a consequence, we can say that the newly 358 constructed test was very well described by the Rasch model so that the number of solved 359 items represents a sufficient statistic for an individual's vocabulary skills.

### Psychometric properties of newly constructed task

## 52 Reliability

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We computed two reliability indices for the newly constructed oREV task: KR-20 (Kuder & Richardson, 1937) and Andrich Reliability (Andrich, 1982). Both indices

<sup>&</sup>lt;sup>4</sup> The high correlations are not surprising given that 17 of the 22 selected items were newly added items. The ceiling effect for most of the original CLT items meant that most of the variation between individuals was captured via the newly added items. Any test with many of the newly added items would have a high correlation with the full task. For this reason, we did not include the correlation with the full test in the objective function passed on to the simulated annealing algorithm.

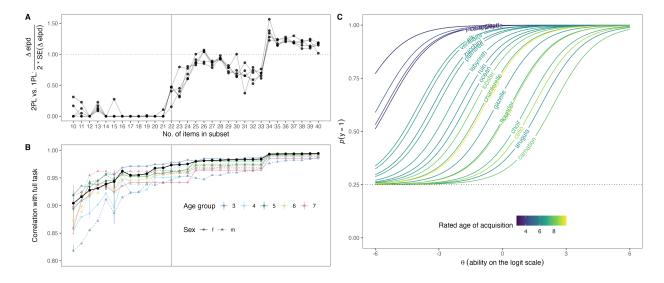


Figure 3. Item selection process. (A) Model comparison ratio comparing the fit of a Rasch model to the fit of a 2PL model for different sizes of the subtest. The y-axis label shows how the ratio is computed. Values of 0 indicate a better fit of the Rasch model compared to the 2PL model. The dashed line marks a ratio of 1, which we assumed to be the point when the 2PL model clearly provided a better fit. Points and lines show the results from five independent runs of the model comparison procedure. (B) Correlation between reduced and full task (52 items). Points show mean correlation based on 5 iterations. Vertical lines show the range of correlations in cases when they differed between iterations. Black lines and points show correlations for the full sample and colored points and lines show correlations by age group and sex. C) Item characteristic curves for the 22 colored by their rated age-of-acquisition.

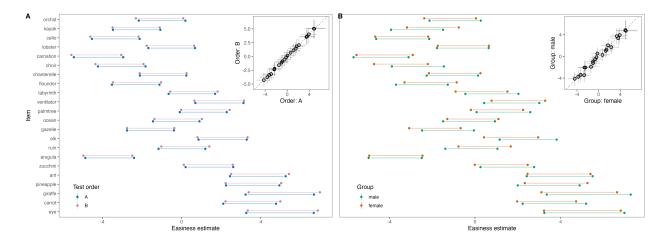


Figure 4. Easiness estimates for all items (ordered by rated age of acquisition) of the newly constructed subtest separate for each test order (A) and sex (B). Dots connected by lines show 95% CrI, color denotes the subgroup. Insets show correlations between the parameters for each subgroup based on the mode of the posterior distribution for each item.

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indicated good reliability for the subtest (KR-20 = 0.76; Andrich = 0.74) that were comparable to the full test (KR-20 = 0.78; Andrich = 0.77).
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## Stability

We were able to re-test 184 children (88 girls; mean age first testing = 4.96; range = 3.03 - 6.99) approximately one year (mean number of days between testings = 341; range = 302 - 369) after the initial data collection with the newly constructed task. Parents received personalized links and children were tested online. As expected, overall performance in the sample increased from 73% to 80% correct, showing developmental gains in receptive vocabulary with age. Nevertheless, individual differences were stable: performance was strongly correlated between the two time points (r = 0.67; 95% CI = 0.58 -0.74).

### 6 Convergent and discriminant validity

Finally, we assessed convergent and discriminant validity of our task. We used the 377 PPVT (Dunn & Dunn, 2007; Lenhard, Lenhard, Segerer, & Suggate, n.d.) as a convergent 378 measure of receptive vocabulary and the digit span task from the K-ABC (Lichtenberger, 379 Sotelo-Dynega, & Kaufman, 2009) as a discriminant measure of working memory. These 380 two tasks were unavailable as online versions, and we, therefore, turned to in-person data 381 collection. We tested 59 children in Kindergartens around Leipzig, Germany. We chose a 382 relatively narrow age range (mean = 5.54; range = 4.97 - 6.02) to avoid strong correlations 383 between tasks due to general developmental gains. Data was collected between January and May 2023.

oREV scores were highly correlated with PPVT scores (r = 0.65; 95% CI = 0.48 – 0.78), but not with digit span scores (r = 0.15; 95% CI = -0.11 – 0.40). Conversely, when we predicted the number of correctly solved items in the oREV by PPVT scores, digit span scores and age in a binomial model, PPVT scores had by far the strongest influence (Figure 5). Taken together, these results demonstrate the convergent and discriminant validity of the oREV task.

392 Discussion

Individual differences in language abilities in childhood are an important predictor of later life outcomes. Yet, high-quality, easy-access measures are rare, especially for pre- and primary school-aged children. Here we reported the construction of a new receptive vocabulary task for German-speaking children between 3 and 8 years of age. Building on earlier work (Haman et al., 2017), we first generated a larger initial pool with 52 items.

Next, we implemented the picture-selection task as a web application and collected data from over 500 children online. We used IRT models and an automated item selection algorithm to select a set of high-quality items that fit the Rasch model. The so-constructed

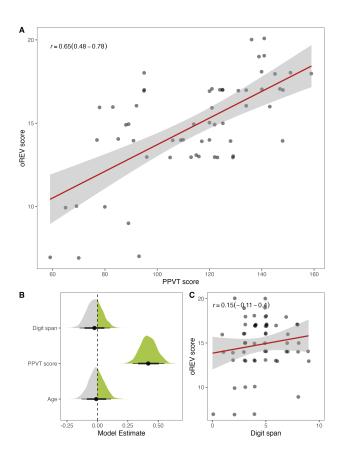


Figure 5. Convergent and discriminant validity. (A) Pearson correlation (with 95% CI) between PPVT and oREV scores. (B) Posterior model estimates for digit span, ppvt scores and age in a model predicting oREV scores. Points show posterior means with 95% CrI. (C) Pearson correlation (with 95% CI) between digit span and oREV scores. In (A) and (C): Grey points show individual data points with minimal horizontal and vertical noise added to avoid overplotting. Red lines show regression lines (with 95% CI) based on a simple linear model.

task has 22 items of varying difficulty, correlates with the full task at a rate of .97, shows 401 good reliability and stability. High correlation with a theoretically related task and low 402 correlation with an unrelated task illustrate convergent and discriminant validity. Its 403 browser-based implementation makes the task highly portable and facilitates large-scale 404 data collection. The construction and item selection process we described here makes it 405 easy to add additional items or adapt the task to different languages while retaining a high 406 psychometric quality of the end product. The task is freely accessible to all interested 407 researchers. 408

The task fills an important gap in the methods repertoire of developmental 409 researchers studying language development in early childhood. Existing measures show 410 ceiling effects, come with high licensing costs, and/or are not available in an electronic 411 format. Our task captures variation between children up until 8 years of age, is free to use, 412 and can be run on any modern web browser. However, the newly constructed task with 22 413 items is still relatively easy, that is, most 7-year-old children will solve the majority of 414 items (89% correct responses in the present sample). As a consequence, it does not 415 distinguish well between children with very strong vocabulary skills. Future extensions of 416 the task could thus focus on adding more difficult items. Figure 2B (see also Brysbaert & Biemiller, 2017) shows that target word age-of-acquisition ratings are a fairly good 418 predictor of item difficulty and could be used as a basis to generate new items. Extensions should focus on target words with rated age-of-acquisition above 10. Further extensions 420 could target other parts of speech, such as verbs and adjectives. 421

The sample we tested to construct the test was not representative (in terms of socioeconomic status) and likely skewed to children with higher language proficiency than average. As a consequence, the task and the dataset should not be used for diagnostic purposes but only as a research tool to capture variability in a population of interest.

Nevertheless, the good psychometric properties of the task make it an ideal candidate for future norming studies with representative samples.

The automated item selection process we implemented critically leveraged the 428 strengths of IRT modeling. For each item in the pool, we estimated its item difficulty and 429 item discrimination. The objective function we optimized via the simulated annealing 430 process was defined so that it would yield a subset in which items would a) be equally 431 spread out across the latent ability so that the task measured equally well at different skill 432 levels and b) have equal discrimination parameters so that the sum score is a sufficient 433 statistic for the ability parameter. In addition, we prioritized items with more precise 434 difficulty estimates (i.e., narrower CrIs). 435

This procedure presents a principled way of constructing a task with good 436 psychometric properties, which can easily be applied to any new set of items or versions of 437 the task in different languages. However, this approach does not make the careful, 438 principle-based construction of the initial item pool superfluous; it only selects the best of 430 the available items. Linguistic and psychometric considerations thus need to go hand in hand during task construction. For example, while nouns are more similar across 441 languages, verbs are more language-specific and might have different representations or 442 even be absent as a single word. For example, the German verb "wandern" (eng: "hiking") 443 can only be expressed by an analytical construction in the majority of Slavic languages. Furthermore, bilingual and monolingual lexicons might vary and background factors, such as length of exposure, the onset of second language acquisition, or birth order should be considered. Finally, language-specific morphosyntactic properties of grammar, such as perfective or imperfective aspect in verbs, should be taken into account.

A major advantage of the task presented here is its portability. Its implementation as
a web application makes it easy to administer both in-person and online and also reduces
the likelihood of experimenter error. In fact, we were able to collect data from more than
500 children online in just two months. It is also easy to add new items or to adapt the
existing task to a new language. Of course, extensions and new adaptations require a
renewed item evaluation and selection process. Nevertheless, the infrastructure and

materials developed here provide a good starting point for such an endeavor. The
computerized implementation of the task also allows for adaptive testing. Instead of all
participants completing the same set of items, each participant could be presented with –
potentially fewer – maximally informative items given their (continuously updated)
estimated skill level. However, this would require a more elaborate back-end – capable of
doing online parameter estimation – compared to the current version of the task.

461 Conclusion

468

We have described the construction of a new receptive vocabulary task for

German-speaking children between 3 and 8 years of age. The task has good psychometric

properties and shows convergent and discriminant validity. We believe it is an important

research instrument to measure individual differences in receptive vocabulary skills. The

task, and the materials it is constructed from, are openly available. As such, it closes a

prominent gap in the toolkit of developmental researchers.

## **Open Practices Statement**

The task can be accessed via the following website:

https://ccp-odc.eva.mpg.de/orev-demo/. The corresponding source code can be found in

the following repository: https://github.com/ccp-eva/orev-demo. The data sets generated

during and/or analysed during the current study are available in the following repository:

https://github.com/ccp-eva/orev/. Data collection was preregistered at:

https://osf.io/qzstk.

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