

A Bayesian mixed effects model to capture the effect of interseasonal weather on phenology

Dan Cunha

Boston University

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Harmonic regression for phenology modeling

Let h be the number of harmonics,

$$\epsilon_i \stackrel{i.i.d.}{\sim} N(0, \sigma^2), i = 1, \dots, n$$

$$y_i | t_i, \beta \sim \beta_0 + \beta_1 t_i + \sum_{l=1}^h \beta_{2l} \sin\left(\frac{2\pi t_i l}{365}\right) + \beta_{2l+1} \cos\left(\frac{2\pi t_i l}{365}\right) + \epsilon_i$$

And define X as the corresponding design matrix.

Harmonic regression is too rigid for phenological modeling

- 1 Harmonic regression assumes the same phenology each year
- 2 But each aspect of phenology, such as the green-up or senescence, naturally change year to year
- 3 The phenology model should incorporate that flexibility
- 4 A repeated measurement model can do just that

Harmonic regression is too rigid for phenological modeling

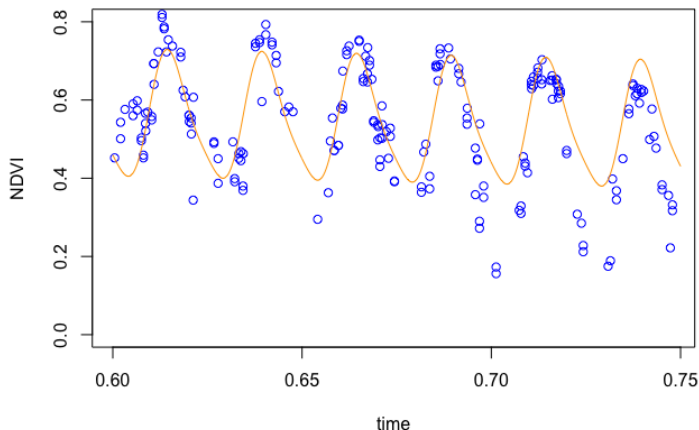


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Towards a Bayesian repeated measurement model

- 1 Repeated measurement models can provide flexibility to learn natural changes in phenology each year
- 2 Common repeated measurement models include mixed effects models
- 3 For example, a mixed effects model could learn a random effect for the harmonics each year.
- 4 This would enable scientists to infer, for example, a separate green-up time each year.

Towards a Bayesian repeated measurement model con't

- 1 The frequentist treatment of mixed effects models considers the random effects as nuisance parameters
- 2 But in phenology modeling, we are interested in plotting the mean predictive function given the data, as well as the series of random effects through time.
- 3 We can use a Bayesian approach instead, capturing the posterior distribution of the yearly random effects, in order to understand how phenology naturally changes each year.

A Bayesian repeated measurement model

Let J be the number of years in the study, d the number of random effects, and n the number of observations,

$$\epsilon_i \stackrel{i.i.d.}{\sim} N(0, \sigma^2), i = 1, \dots, n$$

$$y \sim X\beta + Z\gamma + \epsilon$$

$$\gamma \sim N(0_{Jd \times 1}, T)$$

Where $Z \in \mathbb{R}^{n \times Jd}$ is the design matrix for the random effects, and T is diagonal with parameters $\tau_{l=1, \dots, d}$, each having J repeated entries,

$$T = \text{diag}(\tau_1, \dots, \tau_1(\text{total}J), \dots, \tau_d, \dots, \tau_d)$$

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Approach for inference and computation

- 1 We are tasked with modeling trillions of pixels, so computation needs to be fast
- 2 We also wish to take a Bayesian approach to make an inference about the expectation of the random effects given the data
- 3 Expectation Maximization is suitable for this setting

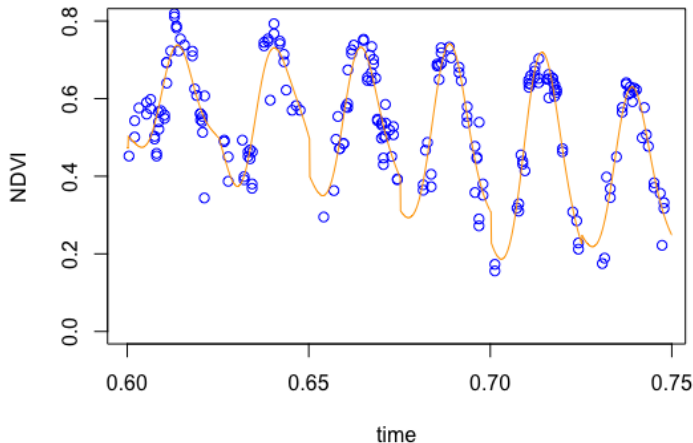
Posterior inference of the random effects

- ① During EM, the posterior expectation of γ is under the distribution $p(\gamma|y, \beta^{(t)})$.
- ② Since this distribution conditions on y and $\beta^{(t)}$, we can think of the likelihood as being a function of just the parameters γ , given the *working response* $(y - X\beta^{(t)})$. This likelihood is still normally distributed with respect to the working response, leading to an analytical posterior expectation calculation for γ .
- ③ Furthermore, we can use $(y - \mathbb{E}(Z\gamma))$ as the working response when performing the M-step for β .
- ④ The remaining M-step calculations for σ^2 and T are analytically tractable.

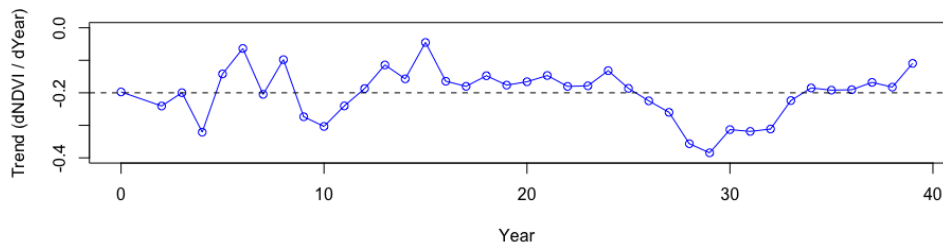
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Random effects on intercept and first harmonics fits well



The random effects provide insight into growth and drought patterns



Next steps

- 1 Include uncertainty of posterior estimates
- 2 Use labeled change point time series to cross validate model flexibility