Lecture 17 Maximum Likelihood Estimation (MLE)

BIO210 Biostatistics

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Spring, 2022

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Intuition over MLE

Experiment: A coin, with an unknown P(H) = p, was flipped 10 times. The outcome is HHHTHHHTHH.

Question: What is you best guess for p?

Thinking: Given the data/observation we have, what values should p take such that our data/observation is most likely to occurr?

Aim: find the value that maximise our chance of observing the data, and use that value as our best guess/estimate for p.

$$P(\mathsf{data}|P(H)=p)$$

Estimators of Parameters

• Parameter space Ω : the set of all possible values of a parameter θ or of a vector of parameters $(\theta_1, \theta_2, \theta_3, ..., \theta_k)$ is called the parameter space.

- Bernoulli: $\theta = p$, $\Omega = \{p \mid 0 \leqslant p \leqslant 1\}$
- Binomial: $\theta_1=n, \theta_2=p, \ \Omega=\{(n,p) \mid n=2,3,..., \text{a finite number}; 0\leqslant p\leqslant 1\}$
- Poisson: $\theta = \lambda$, $\Omega = \{\lambda \mid \lambda \geqslant 0\}$
- Normal (Gaussian): $\theta_1 = \mu, \theta_2 = \sigma^2$, $\Omega = \{(\mu, \sigma^2) \mid \mu \in \mathbb{R}, \sigma^2 \geqslant 0\}$
- We refer to an estimator of a parameter θ as $\hat{\theta}$. An estimator $\hat{\theta}$ of a parameter θ is unbiased if $E[\hat{\theta}] = \theta$. For example, $\hat{\mu} = \bar{X}$ is an unbiased estimator for μ .

Maximum Likelihood Estimation (MLE)

- Maximum likelihood estimation (MLE) is a technique used for estimating the parameters of a given distribution, using some observed data.
- Introduced by R.A. Fisher in 1912.
- MLE can be used to estimate parameters using a limited sample of the population, by finding particular values so that the observation is the most likely result to have occurred.

Maximum Likelihood Estimation (MLE)

Formal definition

Let $x_1, x_2, x_3, ..., x_n$ be observations from n i.i.d random variables $(X_1, X_2, X_3, ..., X_n)$ drawn from a probability distribution f_0 , where f_0 is known to be from a family of distributions f that depend on some parameters θ . For example, f_0 could be known to be from the family of normal distributions f, which depend on parameters μ and σ , and $x_1, x_2, x_3, ..., x_n$ would be observations from f_0 . The goal of MLE is to maximise the likelihood function:

$$\mathcal{L} = f(x_1, x_2, x_3, ..., x_n; \theta) = \prod_{i=1}^n f(x_i; \theta) = f(x_1; \theta) \cdot f(x_2; \theta) \cdot ... f(x_n; \theta)$$

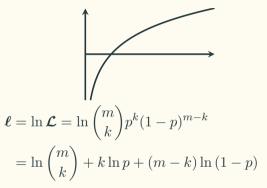
The log-likelihood function:

$$\ell = \ln \mathcal{L} = \sum_{i=1}^{n} \ln f(x_i; \theta)$$

- Other notation: $\mathcal{L} = f(x_1, x_2, x_3, ..., x_n | \theta) = \prod_{i=1}^n f(x_i | \theta)$
- Example 1: A (possibly unfair) coin is flipped 100 times, and 61 heads are observed. The coin either has probability of 1/3, 1/2 or 2/3 of obtaining a head each time it is flipped. Which of the three is the MLE?
- 1. $\theta : (n, p)$
- 2. $\Omega: \{(n,p) \mid n=100, p \in \{\frac{1}{3}, \frac{1}{2}, \frac{2}{3}\}\}$
- 3. $p_X(k) = \binom{n}{k} p^k (1-p)^{n-k}$
- 4. $\mathcal{L} = f(61; n, p)$

$$f(61; n, p) = {100 \choose 61} p^{61} (1-p)^{39} = \begin{cases} \approx 9.6 \times 10^{-9} & \text{, if } p = \frac{1}{3} \\ \approx 0.007 & \text{, if } p = \frac{1}{2} \\ \approx 0.04 & \text{, if } p = \frac{2}{3} \end{cases}$$

- Example 2 A more generalised case of coin flipping: A (possibly unfair) coin is flipped m times, and k heads are observed. Let P(H) = p. What is the MLE for p?
- 1. $\theta : (n, p)$
- 2. $\Omega : \{(n,p) \mid n=m, 0 \leqslant p \leqslant 1\}$
- 3. $p_X(k) = \binom{n}{k} p^k (1-p)^{n-k}$
- 4. $\mathcal{L} = f(k; n, p) = {m \choose k} p^k (1-p)^{m-k}$



What value should p take to maximise ℓ ?

Let
$$\frac{d\ell}{dp} = 0 \implies \hat{p} = \frac{k}{m}$$

- Example 3 DNA synthesis errors: The genetic material is copied and synthesised by DNA polymerase. One high-fidelity DNA polymerase, Pfu, originally isolated from the hyperthermophilic archae Pyrococcus furiosus, is believed to have very low error rate. Assume the errors generated by Pfu follow a Poisson distribution with λ mutations per 10^6 base pairs (Mb). We have examined n newly synthesised DNA fragments and observed that the nubmer of mutations per Mb is $k_1, k_2, k_3, ..., k_n$. What is the MLE for λ ?
- 1. $\theta:\lambda$
- 2. $\Omega : \{ \lambda \mid \lambda \geqslant 0 \}$
- 3. $p_x(k) = \frac{\lambda^k}{k!}e^{-\lambda}$

- 4.
$$\mathcal{L} = f(k_1, k_2, ..., k_n; \lambda) = \prod_{i=1}^n \frac{\lambda^{k_i}}{k_i!} e^{-\lambda}$$

Maximum Likelihood Estimation (MLE): Example 3 Solution

$$\mathcal{L} = f(k_1, k_2, ..., k_n; \lambda) = \frac{\lambda^{k_1}}{k_1!} e^{-\lambda} \cdot \frac{\lambda^{k_2}}{k_2!} e^{-\lambda} \cdot \frac{\lambda^{k_3}}{k_3!} e^{-\lambda} \cdot ... \cdot \frac{\lambda^{k_n}}{k_n!} e^{-\lambda} = \frac{\lambda^{k_1 + k_2 + \dots + k_n}}{k_1! k_2! \cdot ... k_n!} e^{-n\lambda}$$

$$\ell = \ln \frac{\lambda^{k_1 + k_2 + \dots + k_n}}{k_1! k_2! \cdot ... k_n!} e^{-n\lambda} = \ln \lambda^{k_1 + k_2 + \dots + k_n} - \ln (k_1! k_2! \cdot ... k_n!) + \ln e^{-n\lambda}$$

$$= (k_1 + k_2 + \dots + k_n) \ln \lambda - n\lambda - \ln (k_1! k_2! \cdot ... k_n!)$$

Now, we let:

$$\frac{d\ell}{d\lambda} = 0 \Rightarrow \frac{k_1 + k_2 + \dots + k_n}{\lambda} - n = 0 \Rightarrow \hat{\lambda} = \frac{1}{n} \sum_{i=1}^{n} k_i$$

- Example 4 Oct4 expression in embryonic stem cells: Let the random variable \boldsymbol{X} be the expression values of Oct4. We know $\boldsymbol{X} \sim \mathcal{N}(\mu, \sigma)$, but we μ and σ are unknown. Now we have sequenced 5 cells, and the expressions of Oct4 in those cells are $3.0,\ 3.5,\ 2.5,\ 3.2,\ 2.8$, respectively. What is the MLE for the parameters of this normal distribution?
- 1. $\theta : \mu, \sigma^2$
- 2. Ω : $\{(\mu, \sigma^2) \mid -\infty < \mu < +\infty, \ \sigma^2 \geqslant 0\}$

- 3.
$$f_X(x) = \frac{1}{\sqrt{2\pi}\sigma}e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

- 4.
$$\mathcal{L} = f(3.0, 3.5, 2.5, 3.2, 2.8; \mu, \sigma^2) = \prod_{i=1}^{5} \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x_i - \mu)^2}{2\sigma^2}}$$

Maximum Likelihood Estimation (MLE): Example 4 Solution

$$\mathcal{L} = f(3.0, 3.5, 2.5, 3.2, 2.8; \mu, \sigma^{2})$$

$$= \frac{1}{\sqrt{2\pi}} e^{-\frac{(3.0 - \mu)^{2}}{2\sigma^{2}}} \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{(3.5 - \mu)^{2}}{2\sigma^{2}}} \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{(2.5 - \mu)^{2}}{2\sigma^{2}}}$$

$$\cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{(3.2 - \mu)^{2}}{2\sigma^{2}}} \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{(2.8 - \mu)^{2}}{2\sigma^{2}}}$$

$$= \left(\frac{1}{\sqrt{2\pi}\sigma}\right)^{5} \cdot e^{-\frac{(3.0 - \mu)^{2} + (3.5 - \mu)^{2} + (2.5 - \mu)^{2} + (3.2 - \mu)^{2} + (2.8 - \mu)^{2}}{2\sigma^{2}}$$

Maximum Likelihood Estimation (MLE): Example 4 Solution

$$\ell = \ln\left(\frac{1}{\sqrt{2\pi}\sigma}\right)^5 + \ln e^{-\frac{5\mu^2 - 30\mu + 45.58}{2\sigma^2}}$$

$$= 5\ln\left(\frac{1}{\sqrt{2\pi}\sigma}\right) - \frac{5\mu^2 - 30\mu + 45.58}{2\sigma^2} = -\frac{5}{2\sigma^2} \cdot \mu^2 + \frac{15}{\sigma^2} \cdot \mu - \frac{45.58}{2\sigma^2} + 5\ln\left(\frac{1}{\sqrt{2\pi}\sigma}\right)$$

$$= -5 \cdot \ln \sigma - \frac{5\mu^2 - 30\mu + 45.58}{2} \cdot \sigma^{-2} - 5\ln\sqrt{2\pi}$$

Now, we can find out the values of μ and σ by letting:

$$\frac{\partial \ell}{\partial \mu} = 0$$
 and $\frac{\partial \ell}{\partial \sigma} = 0 \Rightarrow \hat{\mu} = 3$ and $\hat{\sigma^2} = 0.116$