

Lecture 17 Maximum Likelihood Estimation (MLE)

BIO210 Biostatistics

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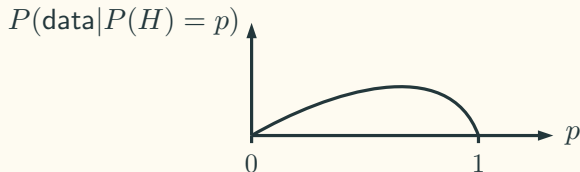
Intuition over MLE

Experiment: A coin, with an unknown $P(H) = p$, was flipped 10 times. The outcome is $HHHTHHHTHH$.

Question: What is your best guess for p ?

Thinking: Given the data/observation we have, what values should p take such that our data/observation is most likely to occur ?

Aim: find the value that **maximise our chance of observing the data**, and use that value as our best guess/estimate for p .



Estimators of Parameters

- **Parameter space Ω :** the set of all possible values of a parameter θ or of a vector of parameters $(\theta_1, \theta_2, \theta_3, \dots, \theta_k)$ is called the parameter space.
 - Bernoulli: $\theta = p$, $\Omega = \{0 \leq p \leq 1\}$
 - Binomial: $\theta_1 = n, \theta_2 = p$, $\Omega = \{(n, p) \mid n = 2, 3, \dots, \text{a finite number}; 0 \leq p \leq 1\}$
 - Poisson: $\theta = \lambda$, $\Omega = \{\lambda = 0, 1, 2, 3, 4, \dots\}$
 - Normal (Gaussian): $\theta_1 = \mu, \theta_2 = \sigma^2$, $\Omega = \{(\mu, \sigma^2) \mid \mu \in \mathbb{R}, \sigma^2 \geq 0\}$
- We refer to an estimator of a parameter θ as $\hat{\theta}$. An estimator $\hat{\theta}$ of a parameter θ is unbiased if $E[\hat{\theta}] = \theta$. For example, $\hat{\mu} = \bar{X}$ is an unbiased estimator for μ .

Maximum Likelihood Estimation (MLE)

- **Maximum likelihood estimation (MLE)** is a technique used for estimating the parameters of a given distribution, using some observed data.
- Introduced by R.A. Fisher in 1912.
- MLE can be used to estimate parameters using a limited sample of the population, by finding particular values so that the observation is the most likely result to have occurred.

Maximum Likelihood Estimation (MLE)

Formal definition

Let $x_1, x_2, x_3, \dots, x_n$ be observations from n **i.i.d** random variables $(X_1, X_2, X_3, \dots, X_n)$ drawn from a probability distribution f_0 , where f_0 is known to be from a family of distributions \mathcal{f} that depend on some parameters θ . For example, f_0 could be known to be from the family of normal distributions \mathcal{f} , which depend on parameters μ and σ , and $x_1, x_2, x_3, \dots, x_n$ would be observations from f_0 . The goal of MLE is to maximise the **likelihood function**:

$$\mathcal{L} = f(x_1, x_2, x_3, \dots, x_n; \theta) = \prod_{i=1}^n f(x_i; \theta) = f(x_1; \theta) \cdot f(x_2; \theta) \cdots f(x_n; \theta)$$

The log-likelihood function:

$$\ell = \ln \mathcal{L} = \sum_{i=1}^n \ln f(x_i; \theta)$$

Maximum Likelihood Estimation (MLE): Example 1

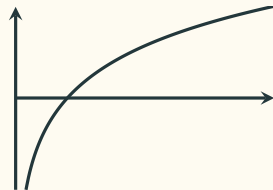
- Other notation: $\mathcal{L} = f(x_1, x_2, x_3, \dots, x_n | \theta) = \prod_{i=1}^n f(x_i | \theta)$
- **Example 1:** A (possibly unfair) coin is flipped 100 times, and 61 heads are observed. The coin either has probability of $1/3$, $1/2$ or $2/3$ of obtaining a head each time it is flipped. Which of the three is the MLE?
 - 1. $\theta : (n, p)$
 - 2. $\Omega : \{(n, p) \mid n = 100, p \in \{\frac{1}{3}, \frac{1}{2}, \frac{2}{3}\}\}$
 - 3. $p_X(k) = \binom{n}{k} p^k (1-p)^{n-k}$
 - 4. $\mathcal{L} = f(61; n, p)$

$$f(61; n, p) = \binom{100}{61} p^{61} (1-p)^{39} = \begin{cases} \approx 9.6 \times 10^{-9} & , \text{ if } p = \frac{1}{3} \\ \approx 0.007 & , \text{ if } p = \frac{1}{2} \\ \approx 0.04 & , \text{ if } p = \frac{2}{3} \end{cases}$$

Maximum Likelihood Estimation (MLE): Example 2

- **Example 2 A more generalised case of coin flipping:** A (possibly unfair) coin is flipped m times, and k heads are observed. Let $P(H) = p$. What is the MLE for p ?

- 1. $\theta : (n, p)$
- 2. $\Omega : \{(n, p) \mid n = m, 0 \leq p \leq 1\}$
- 3. $p_X(k) = \binom{n}{k} p^k (1-p)^{n-k}$
- 4. $\mathcal{L} = f(k; n, p) = \binom{m}{k} p^k (1-p)^{m-k}$



$$\begin{aligned}\ell = \ln \mathcal{L} &= \ln \binom{m}{k} p^k (1-p)^{m-k} \\ &= \ln \binom{m}{k} + k \ln p + (m-k) \ln (1-p)\end{aligned}$$

What value should p take to maximise ℓ ?

$$\text{Let } \frac{d\ell}{dp} = 0 \Rightarrow \hat{p} = \frac{k}{m}$$

Maximum Likelihood Estimation (MLE): Example 3

- **Example 3 DNA synthesis errors:** The genetic material is copied and synthesised by DNA polymerase. One high-fidelity DNA polymerase, *Pfu*, originally isolated from the hyperthermophilic archae *Pyrococcus furiosus*, is believed to have very low error rate. Assume the errors generated by *Pfu* follow a Poisson distribution with λ mutations per 10^6 base pairs (Mb). We have examined n newly synthesised DNA fragments and observed that the number of mutations per Mb is $k_1, k_2, k_3, \dots, k_n$. What is the MLE for λ ?
 - 1. $\theta : \lambda$
 - 2. $\Omega : \{\lambda \mid \lambda \geq 0\}$
 - 3. $p_x(k) = \frac{\lambda^k}{k!} e^{-\lambda}$
 - 4. $\mathcal{L} = f(k_1, k_2, \dots, k_n; \lambda) = \prod_{i=1}^n \frac{\lambda^{k_i}}{k_i!} e^{-\lambda}$

Maximum Likelihood Estimation (MLE): Example 3 Solution

$$\mathcal{L} = f(k_1, k_2, \dots, k_n; \lambda) = \frac{\lambda^{k_1}}{k_1!} e^{-\lambda} \cdot \frac{\lambda^{k_2}}{k_2!} e^{-\lambda} \cdot \frac{\lambda^{k_3}}{k_3!} e^{-\lambda} \cdots \frac{\lambda^{k_n}}{k_n!} e^{-\lambda} = \frac{\lambda^{k_1+k_2+\dots+k_n}}{k_1!k_2!\cdots k_n!} e^{-n\lambda}$$

$$\begin{aligned}\ell &= \ln \frac{\lambda^{k_1+k_2+\dots+k_n}}{k_1!k_2!\cdots k_n!} e^{-n\lambda} = \ln \lambda^{k_1+k_2+\dots+k_n} - \ln(k_1!k_2!\cdots k_n!) + \ln e^{-n\lambda} \\ &= (k_1 + k_2 + \dots + k_n) \ln \lambda - n\lambda - \ln(k_1!k_2!\cdots k_n!)\end{aligned}$$

Now, we let:

$$\frac{d\ell}{d\lambda} = 0 \Rightarrow \frac{k_1 + k_2 + \dots + k_n}{\lambda} - n = 0 \Rightarrow \hat{\lambda} = \frac{1}{n} \sum_{i=1}^n k_i$$

Maximum Likelihood Estimation (MLE): Example 4

- **Example 4 *Oct4* expression in embryonic stem cells:** Let the random variable X be the expression values of *Oct4*. We know $X \sim \mathcal{N}(\mu, \sigma)$, but we μ and σ are unknown. Now we have sequenced 5 cells, and the expressions of *Oct4* in those cells are 3.0, 3.5, 2.5, 3.2, 2.8, respectively. What is the MLE for the parameters of this normal distribution?

- 1. $\theta : \mu, \sigma$
- 2. $\Omega : \{(\mu, \sigma) \mid -\infty < \mu < +\infty, \sigma \geq 0\}$

- 3.
$$f_X(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x - \mu)^2}{2\sigma^2}}$$

- 4.
$$\mathcal{L} = f(3.0, 3.5, 2.5, 3.2, 2.8; \mu, \sigma) = \prod_{i=1}^5 \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x_i - \mu)^2}{2\sigma^2}}$$

Maximum Likelihood Estimation (MLE): Example 4 Solution

$$\begin{aligned}\mathcal{L} &= f(3.0, 3.5, 2.5, 3.2, 2.8; \mu, \sigma) \\&= \frac{1}{\sqrt{2\pi}} e^{-\frac{(3.0 - \mu)^2}{2\sigma^2}} \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{(3.5 - \mu)^2}{2\sigma^2}} \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{(2.5 - \mu)^2}{2\sigma^2}} \\&\quad \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{(3.2 - \mu)^2}{2\sigma^2}} \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{(2.8 - \mu)^2}{2\sigma^2}} \\&= \left(\frac{1}{\sqrt{2\pi}\sigma} \right)^5 \cdot e^{-\frac{(3.0 - \mu)^2 + (3.5 - \mu)^2 + (2.5 - \mu)^2 + (3.2 - \mu)^2 + (2.8 - \mu)^2}{2\sigma^2}}\end{aligned}$$

Maximum Likelihood Estimation (MLE): Example 4 Solution

$$\begin{aligned}\ell &= \ln \left(\frac{1}{\sqrt{2\pi}\sigma} \right)^5 + \ln e^{-\frac{5\mu^2 - 30\mu + 45.58}{2\sigma^2}} \\&= 5 \ln \left(\frac{1}{\sqrt{2\pi}\sigma} \right) - \frac{5\mu^2 - 30\mu + 45.58}{2\sigma^2} = -\frac{5}{2\sigma^2} \cdot \mu^2 + \frac{15}{\sigma^2} \cdot \mu - \frac{45.58}{2\sigma^2} + 5 \ln \left(\frac{1}{\sqrt{2\pi}\sigma} \right) \\&= -5 \cdot \ln \sigma - \frac{5\mu^2 - 30\mu + 45.58}{2} \cdot \sigma^{-2} - 5 \ln \sqrt{2\pi}\end{aligned}$$

Now, we can find out the values of μ and σ by letting:

$$\frac{\partial \ell}{\partial \mu} = 0 \text{ and } \frac{\partial \ell}{\partial \sigma} = 0 \Rightarrow \hat{\mu} = 3 \text{ and } \hat{\sigma} = 0.116$$