MULTINORMAL MAXIMUM LIKELIHOOD WITH A COVARIANCE BOUND

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In some consulting work the problem came up to find the maximum likelihood estimate of the covariance matrix of a multivariate normal distribution in *n* dimensions, under the constraint that the estimate is bounded below by a known matrix.

This means we have to solve

(1)
$$\min_{\Sigma \gtrsim \Sigma_0} \log |\Sigma| + \mathbf{tr} \, \Sigma^{-1} S,$$

where *S* is an observed covariance matrix and Σ_0 is a known positive definite bound. We use the symbol \gtrsim for the Loewner order, i.e. $A \gtrsim B$ means that A - B is positive semi-definite.

We can use an initial change of variables to simplify (1). Since Σ_0 is positive definite, we can define the new variable $\Theta = \Sigma_0^{-\frac{1}{2}} \Sigma \Sigma_0^{-\frac{1}{2}}$ and the new target $T = \Sigma_0^{-\frac{1}{2}} S \Sigma_0^{-\frac{1}{2}}$. Problem (1) becomes

(2)
$$\min_{\Theta \geq I} \log |\Theta| + \mathbf{tr} \, \Theta^{-1} T.$$

Observe that $\Theta \gtrsim I$ simply means that the smallest eigenvalue of Θ must be at least one.

Make an additional change of variables. Suppose $T = K\Lambda K'$ is any eigendecomposition of T. Define the new variables $\Xi = K\Theta K'$, and Problem (2) now is

(3)
$$\min_{\Xi \geqslant I} \log |\Xi| + \mathbf{tr} \,\Xi^{-1} \Lambda.$$

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For a final change of variables, suppose $\Xi = L\Omega L'$ is any eigen-decomposition of Ξ . Then

(4)
$$\min_{LL'=L'} \min_{\Omega \gtrsim I} \log |\Omega| + \operatorname{tr} L\Omega^{-1} L' \Lambda.$$

In scalar notation we can write for (4)

(5)
$$\min_{LL'=L'L=I} \min_{\omega \ge 1} \sum_{i=1}^{n} \log \omega_i + \sum_{i=1}^{n} \sum_{j=1}^{n} \lambda_i \omega_j^{-1} \ell_{ij}^2.$$

The matrix Π with elements $\pi_{ij} = \ell_{ij}^2$ is doubly stochastic, i.e. it is non-negative and its rows and columns add up to one. We write this as $\Pi \in \mathcal{D}$. Since not every doubly stochastic matrix in \mathcal{D} is the elementwise square of a rotation matrix we have

(6)
$$\min_{L'L=LL'=I} \sum_{i=1}^{n} \sum_{j=1}^{n} \lambda_i \omega_j^{-1} \ell_{ij}^2 \ge \min_{\Pi \in \mathcal{D}} \sum_{i=1}^{n} \sum_{j=1}^{n} \lambda_i \omega_j^{-1} \pi_{ij}.$$

In other words minimizing over \mathcal{D} is a convex relaxation of the original problem. We now show that actually the relaxed and the original problem have the same solution.

The set \mathcal{D} of doubly stochastic matrices is the convex hull of the set \mathcal{P} of permutation matrices. This is the famous Birkhoff-Von Neumann Theorem [Berge, 1997, page 182]. And, by an equally famous theorem due to Hardy, Littlewood, and Polya [Berge, 1997, page 184],

(7)
$$\min_{\Pi \in \mathcal{P}} \sum_{i=1}^{n} \sum_{j=1}^{n} \lambda_{i} \omega_{j}^{-1} \pi_{ij} = \sum_{i=1}^{n} \frac{\lambda_{[i]}}{\omega_{[i]}}$$

where the minimum is attained if the permutation puts ω in the same order as λ . Indices in square brackets are used for ordered vectors. Thus

(8a)
$$\lambda_{[1]} \geq \cdots \geq \lambda_{[n]},$$

(8b)
$$\omega_{[1]} \geq \cdots \geq \omega_{[n]}$$
.

Now it suffices to solve

(9)
$$\min_{\omega_{[n]} \ge 1} \sum_{i=1}^{n} \left\{ \log \omega_{[i]} + \frac{\lambda_{[i]}}{\omega_{[i]}} \right\},$$

for which the solution is simply $\omega_{[i]} = \max(\lambda_{[i]}, 1)$.

Thus if we define $\hat{\Omega}$ with diagonal elements $\max(\lambda_i, 1)$, then the solution to Problem (1) is $\hat{\Sigma} = \Sigma_0^{\frac{1}{2}} K' \hat{\Omega} K \Sigma_0^{\frac{1}{2}}$.

 $\hat{\Sigma}$ is a continuous function of S, and consequently a consistent estimate of Σ . But it is not differentiable at locations where S does not satisfy the constraint, and thus it is not asymptotically normal and the likelihood ratio test for the constraint is not asymptotically chi-squared. In fact, the results of Chernoff [1954] show the asymptotic distribution is a mixture of chi-squares, with mixture probabilities given by the asymptotic probabilities that one or more of the smaller eigenvalues violate the constraints. If the true Σ is strictly larger than Σ_0 , then $\hat{\Sigma}$ is almost surely equal to S, and the log likelihood ratio is almost surely equal to zero.

References

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