UNBOUNDED LIKELIHOODS IN ARRAY MULTINORMAL GROWTH CURVE MODELS

JAN DE LEEUW

ABSTRACT. Meet the abstract. This is the abstract.

1. Introduction

The negative log-likelihood for a matrix normal model with data $Y \in \mathbb{R}^{n \times m}$ is

$$\sigma(V,W,X,B,Z)=m\log\det(V)+n\log\det(W)+\operatorname{tr} V^{-1}(Y-XBZ')W^{-1}(Y-XBZ')'.$$
 Suppose

- \mathcal{V} is a set of positive definite matrices or order n,
- W is a set of positive definite matrices or order m,
- X is a set of $n \times p$ matrices,
- Z is a set of $m \times q$ matrices,
- \mathcal{B} is a set of $p \times q$ matrices.

Define

$$\sigma(\mathcal{V},\mathcal{W},\mathcal{X},\mathcal{B},\mathcal{Z}) = \inf_{V \in \mathcal{V}} \inf_{W \in \mathcal{W}} \inf_{X \in \mathcal{X}} \inf_{Z \in \mathcal{Z}} \inf_{B \in \mathcal{B}} \sigma(V,W,X,B,Z)$$

Theorem 1.1. If V and W contain the set of all diagonal matrices, and if

$$\mathcal{M} = \begin{bmatrix} \mu' \\ \mu' \\ \vdots \\ \mu' \end{bmatrix}$$

then $\sigma(Y \mid \mathcal{V}, \mathcal{W}, \mathcal{M}) = -\infty$, i.e. the maximum likelihood estimates of $(\mathcal{V}, \mathcal{W}, \mathcal{M})$ do not exist.

Proof. Take W = I and μ equal to the first row of Y. V is diagonal, with all elements equal to one, except the first which is equal to ϵ . Then

$$m \log \det(V) + n \log \det(W) + \operatorname{tr} V^{-1}(Y - M)W^{-1}(Y - M)' =$$

= $m \log \epsilon + \operatorname{tr}(Y - M)'(Y - M)$

which goes to
$$-\infty$$
 if $\epsilon \to 0$.

Obviously this result remains true if \mathcal{M} is defined any way that makes it possible to fit one row or one column of Y exactly. This implies the usual Anderson-Rubin result for fixed score ML factor analysis. The result is not true, at least not necessarily true, if \mathcal{V} is a set of correlation matrices, or if M = XBZ' with X and Z fixed.

Alternatively we can use residual maximum likelihood. If K is $n \times (n-1)$ and orthonormal, with columns that add up to zero, then K'M = 0, and thus we can compute REML estimates by minimizing

$$m \log \det(K'VK) + (n-1) \log \det(W) + \operatorname{tr} (K'VK)^{-1}K'YW^{-1}Y'K.$$

For M = XBZ' with X and Z known, we can compute the likelihood of $X'_{\perp}YZ_{\perp}$, using bases for the orthogonal complements of X and Z. This keep us in the domain of the matrix normal. In particular, if $V = \sum \theta_s A_s$ and $W = \sum \xi_s B_s$ have linear structure, then in REML they still have linear structure.

NB: If we apply this REML approach to fixed score factor analysis, we must define the likelihood of $Y - \Pi_r(Y)$, where P_r defines the best rank p approximation to Y. This can be done, at least in principle, by using the Jacobian of the singular value decomposition (see http://preprints.stat.ucla.edu/547/lowrank.pdf). This provides an alternative to the Anderson-Rubin non-central Wishart

method, and to MacDonald's maximum likelihood ratio method. That, however, is another project.

2. General

Suppose \mathcal{P}_n is the cone of all positive definite matrices of order n, and $\overline{\mathcal{P}}_n$ is its closure, i.e. the set of all positive semi-definite matrices. For $V \in \mathcal{P}_n$ we also write $V \succ 0$ and for $V \in \overline{\mathcal{P}}_n$ we write $V \succeq 0$.

Suppose $S \succeq 0$. Define $f_S : \mathcal{P}_n \Rightarrow \mathbb{R}$ by

$$f_S(V) = \log \det(V) + \operatorname{tr} V^{-1} S$$
,

and

$$f_{S}(\mathcal{V}) = \inf_{V \in \mathcal{V}} f_{S}(V).$$

Theorem 2.1. $f_S(\mathcal{P}_n) > -\infty$ if and only if S > 0.

Proof. If S > 0 then let $W = S^{-\frac{1}{2}}VS^{-\frac{1}{2}}$. Now

$$f_S(\mathcal{P}_n) = \log \det(S) + \inf_{W \in \mathcal{P}_n} \left\{ \log \det(W) + \operatorname{tr} W^{-1} \right\}.$$

If $\lambda_i > 0$ are the eigenvalues of W, then

$$\log \det(W) + \operatorname{tr} W^{-1} = \sum_{i=1}^{n} \left\{ \log \lambda_i + \frac{1}{\lambda_i} \right\},\,$$

which attains a minimum equal to n if $\lambda_i = 1$ for all i, which means that W = I and V = S.

Suppose $S \succeq 0$ and $\mathbf{rank}(S) = r < n$. Let K_{\perp} be an $n \times (n-r)$ matrix with an orthonormal basis for the null space of S. Define $V(\epsilon) = S + \epsilon K_{\perp} K'_{\perp}$. Then $\mathbf{tr} \ V^{-1}(\epsilon)S = r$ and $\log \det(V(\epsilon)) = \sum_{i=1}^{r} \log \sigma_i + (n-r) \log \epsilon$, where the σ_i are the non-zero eigenvalues of S. Thus $\lim_{\epsilon \to 0} f_S(V(\epsilon)) = -\infty$.

Corollary 2.2. *If* S > 0 *then* $f_S(V) \ge \log \det(S) + n$.

Proof.
$$f_S(\mathcal{V}) \geq f_S(\mathcal{P}_n)$$
. The proof of Theorem 2.1 shows $f_S(\mathcal{P}_n) = \log \det(S) + n$.

Thus if S > 0 the loss function is bounded below. Thus in the sequel we suppose $S \succeq 0$ is singular of rank r. Then $S = K\Lambda K'$, with K an $n \times r$ orthonormal, $K'K_{\perp} = 0$, and Λ diagonal with $\Lambda > 0$.

Theorem 2.3. *If there is a* $\delta > 0$ *such that*

$$\inf_{V \in \mathcal{V}} \min_{s} \lambda_{s}(K'_{\perp}VK_{\perp}) \geq \delta$$

then $f_S(\mathcal{V}) > -\infty$.

Proof. Define A = K'VK, $B = K'VK_{\perp}$, and $C = K'_{\perp}VK_{\perp}$. Then

$$f_S(V) = \log \det \begin{bmatrix} A & B \\ B' & C \end{bmatrix} + \operatorname{tr} \begin{bmatrix} A & B \\ B' & C \end{bmatrix}^{-1} \begin{bmatrix} \Lambda & 0 \\ 0 & 0 \end{bmatrix},$$

and, using the expression for inverses and determinants of partitioned matrices,

$$f_S(V) = \log \det(A - B'C^{-1}B) + \log \det(C) + \operatorname{tr} (A - B'C^{-1}B)^{-1}\Lambda.$$

Because $\Lambda > 0$ we have from Theorem 2.2

$$f_S(V) \ge \log \det(\Lambda) + \log \det(C) + r$$
,

and thus, from the condition in the theorem,

$$f_S(V) \ge \log \det(\Lambda) + (n-r) \log \delta + r$$

which implies $f_S(\mathcal{V}) > -\infty$.

Theorem 2.4. If the closure of V contains at least one singular matrix V_0 and there is a $\delta > 0$ such that

$$\sup_{V \in \mathcal{V}} \max_{s} \lambda_{s}(K'V^{-1}K) \leq \delta$$

then $f_S(\mathcal{V}) = -\infty$.

Proof. For the second result we use

$$f_S(V) \leq \log \det(V) + \delta \operatorname{tr}(\Lambda),$$

and thus
$$V_k \to V_0$$
 implies $f_S(V_k) \to -\infty$.

Department of Statistics, University of California, Los Angeles, CA 90095-1554

E-mail address, Jan de Leeuw: deleeuw@stat.ucla.edu

URL, Jan de Leeuw: http://gifi.stat.ucla.edu