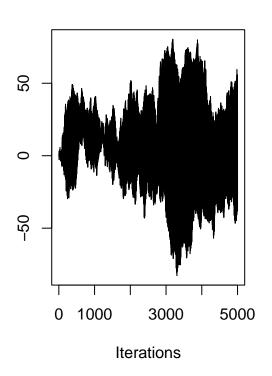
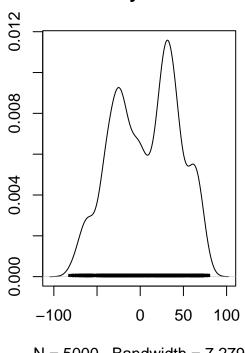
```
# There are three parameters of interest: theta1, theta2, theta3
# Define the likelihood function
likelihood <- function(theta) {</pre>
 return(dnorm(theta[1], mean = 0, sd = 1) *
         dnorm(theta[2], mean = theta[1], sd = 1) *
         dnorm(theta[3], mean = theta[2], sd = 1))
}
# Define the prior distribution
prior <- function(theta) {</pre>
 return(dnorm(theta[1], mean = 0, sd = 1) *
         dnorm(theta[2], mean = 0, sd = 1) *
         dnorm(theta[3], mean = 0, sd = 1))
}
# Define the posterior distribution
posterior <- function(theta) {</pre>
 return(likelihood(theta) * prior(theta))
}
# Define the Gibbs sampler
gibbs_sampler <- function(theta_0, iter) {</pre>
  theta1 <- theta_0[1]
 theta2 <- theta_0[2]</pre>
 theta3 <- theta_0[3]
  samples <- matrix(numeric(iter * 3), ncol = 3)</pre>
  for (i in 1:iter) {
    theta1 \leftarrow rnorm(1, mean = theta2, sd = 1)
    theta2 <- rnorm(1, mean = theta3, sd = 1)
    theta3 <- rnorm(1, mean = theta1, sd = 1)
    samples[i, ] <- c(theta1, theta2, theta3)</pre>
  }
 return(samples)
}
library(coda)
# Run the Gibbs sampler with two different starting points
# Note the second starting choice is a poor choice
#the chain does not converge within 5000 interations
set.seed(123)
samples1 \leftarrow as.mcmc(gibbs_sampler(theta_0 = c(0, 0, 0), iter = 5000))
set.seed(123)
samples2 \leftarrow as.mcmc(gibbs_sampler(theta_0 = c(6,7,8), iter = 5000))
# Combine the results and check for convergence
samples_list <- list(samples1, samples2)</pre>
plot(samples1[,1])
```



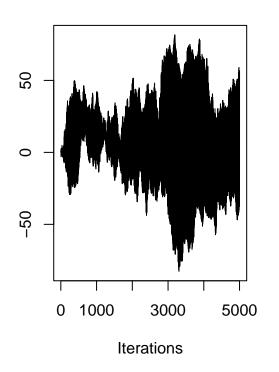
# Density of var1



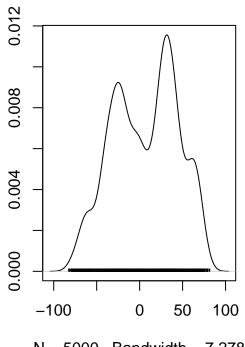
N = 5000 Bandwidth = 7.279

plot(samples1[,2])

#### Trace of var1

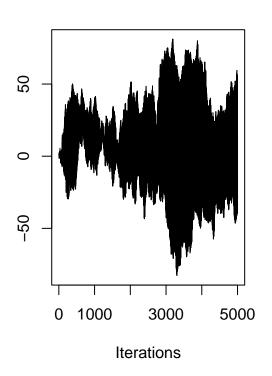


## Density of var1

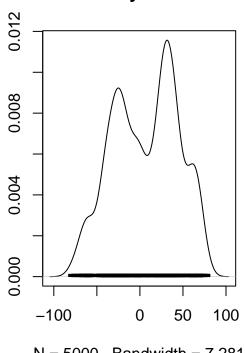


N = 5000 Bandwidth = 7.278

plot(samples1[,3])



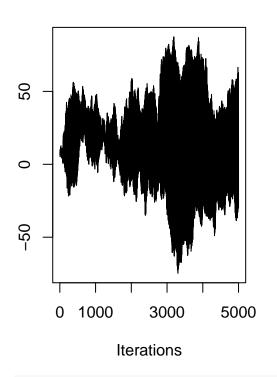
# Density of var1



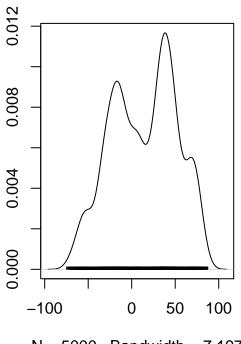
N = 5000 Bandwidth = 7.281

plot(samples2[,1])

## Trace of var1

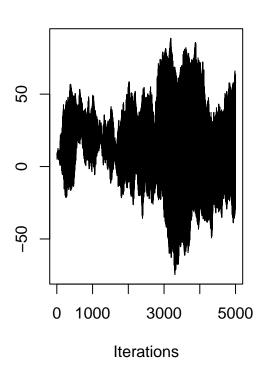


## Density of var1

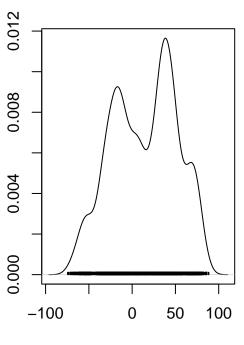


N = 5000 Bandwidth = 7.197

plot(samples2[,2])



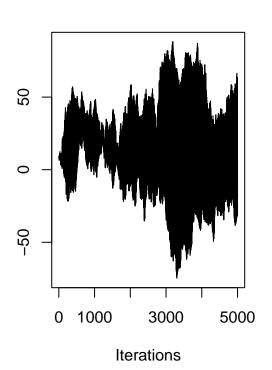
## Density of var1



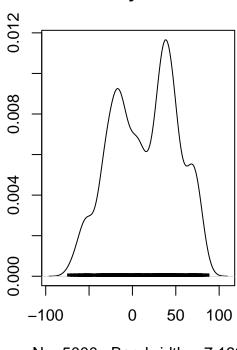
N = 5000 Bandwidth = 7.196

plot(samples2[,3])

#### Trace of var1



## Density of var1



N = 5000 Bandwidth = 7.199

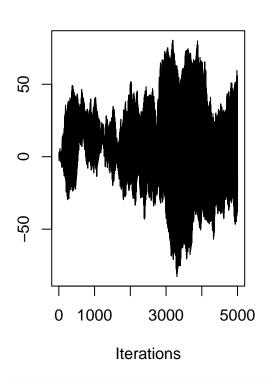
#Z-score for a test of equality that compares the means of the first and last parts of each chain geweke.diag(samples1[,])

```
##
## Fraction in 1st window = 0.1
## Fraction in 2nd window = 0.5
##
    var1 var2
## 0.7032 0.6947 0.7195
geweke.diag(samples2[,])
##
## Fraction in 1st window = 0.1
## Fraction in 2nd window = 0.5
##
   var1
          var2
                 var3
## 0.7020 0.6939 0.7184
#Gelman-Rubin diagnostic
(diagnostics <- gelman.diag(samples_list)) #need at least two chains to check for convergence
## Potential scale reduction factors:
##
       Point est. Upper C.I.
## [1,]
             1.01
                         1.05
## [2,]
              1.01
                         1.05
## [3,]
              1.01
                        1.05
##
## Multivariate psrf
##
## 1.24
#This will produce the potential scale reduction factor (PSRF) for each parameter,
#allowing you to check for convergence of the Metropolis-Hastings algorithm.
#If the PSRF is substantially greater than 1.1 for any parameter, it suggests
#that the chain has not converged.
#Effective Sample Size
effectiveSize(samples_list)
##
       var1
                var2
                         var3
## 420.6048 412.0096 379.6192
# Part A
##Using the gelman-rubin we see the value of point estimation and upper CI is approximately 1 which ind
## Using the Geweke, since the z-score is less than 2 we can say that the chains have converged.
## The effective sample size, we see that the chain has converged a bit as it is > 100
# There are three parameters of interest: theta1, theta2, theta3
# Define the likelihood function
```

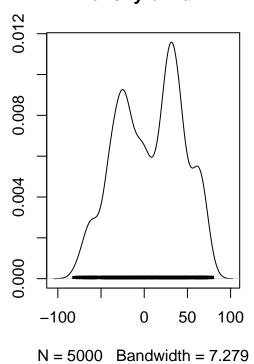
likelihood <- function(theta) {</pre>

return(dnorm(theta[1], mean = 0, sd = 1) \*

```
dnorm(theta[2], mean = theta[1], sd = 1) *
         dnorm(theta[3], mean = theta[2], sd = 1))
}
# Define the prior distribution
prior <- function(theta) {</pre>
  return(dnorm(theta[1], mean = 0, sd = 1) *
         dnorm(theta[2], mean = 0, sd = 1) *
         dnorm(theta[3], mean = 0, sd = 1))
}
# Define the posterior distribution
posterior <- function(theta) {</pre>
  return(likelihood(theta) * prior(theta))
# Define the Gibbs sampler
gibbs_sampler <- function(theta_0, iter) {</pre>
  theta1 <- theta_0[1]
  theta2 <- theta_0[2]</pre>
  theta3 <- theta_0[3]</pre>
  samples <- matrix(numeric(iter * 3), ncol = 3)</pre>
  for (i in 1:iter) {
    theta1 <- rnorm(1, mean = theta2, sd = 1)
    theta2 <- rnorm(1, mean = theta3, sd = 1)</pre>
    theta3 <- rnorm(1, mean = theta1, sd = 1)</pre>
    samples[i, ] <- c(theta1, theta2, theta3)</pre>
  return(samples)
library(coda)
# Run the Gibbs sampler with two different starting points
# Note the second starting choice is a poor choice
#the chain does not converge within 5000 interations
set.seed(123)
samples1 \leftarrow as.mcmc(gibbs_sampler(theta_0 = c(0, 0, 0), iter = 5000))
set.seed(123)
samples2 \leftarrow as.mcmc(gibbs_sampler(theta_0 = c(0.3,0.3,0.3), iter = 5000))
# Combine the results and check for convergence
samples_list <- list(samples1, samples2)</pre>
plot(samples1[,1])
```

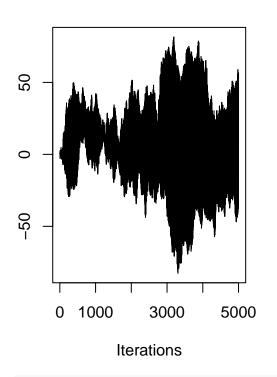


# Density of var1

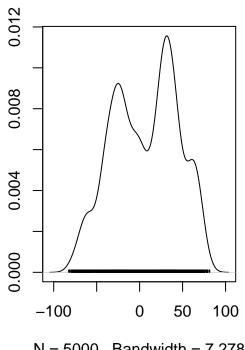


plot(samples1[,2])

Trace of var1

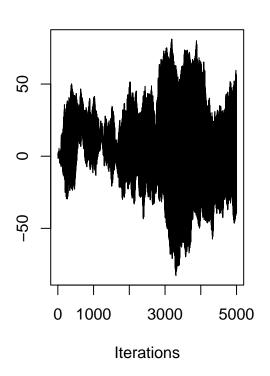


Density of var1

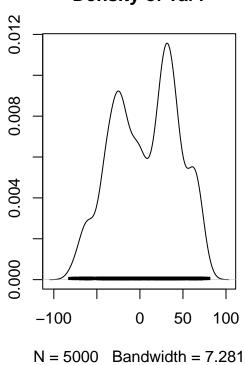


N = 5000 Bandwidth = 7.278

plot(samples1[,3])

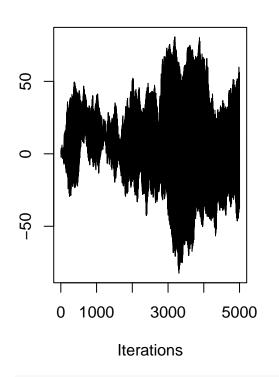


# Density of var1

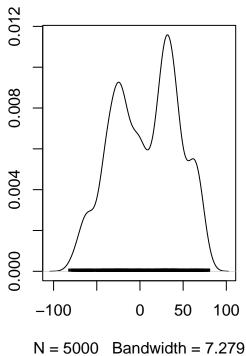


plot(samples2[,1])

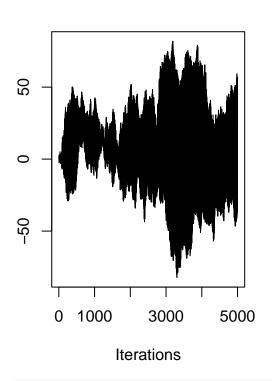
Trace of var1



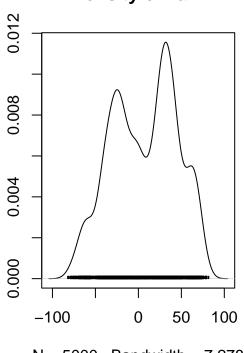
Density of var1



plot(samples2[,2])



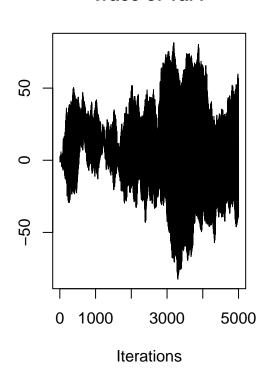
## Density of var1



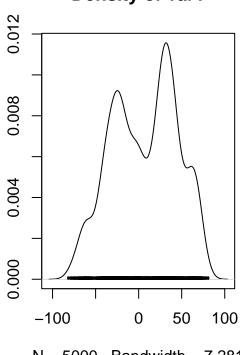
N = 5000 Bandwidth = 7.278

plot(samples2[,3])

#### Trace of var1



## Density of var1



N = 5000 Bandwidth = 7.281

#Z-score for a test of equality that compares the means of the first and last parts of each chain geweke.diag(samples1[,])

```
##
## Fraction in 1st window = 0.1
## Fraction in 2nd window = 0.5
##
    var1
            var2
## 0.7032 0.6947 0.7195
geweke.diag(samples2[,])
##
## Fraction in 1st window = 0.1
## Fraction in 2nd window = 0.5
##
##
   var1
           var2
                  var3
## 0.7032 0.6947 0.7195
#Gelman-Rubin diagnostic
(diagnostics <- gelman.diag(samples_list)) #need at least two chains to check for convergence
## Potential scale reduction factors:
##
##
        Point est. Upper C.I.
## [1,]
                 1
## [2,]
                 1
## [3,]
                 1
                            1
## Multivariate psrf
##
## 1
#This will produce the potential scale reduction factor (PSRF) for each parameter,
{\it \#allowing you to check for convergence of the {\it Metropolis-Hastings algorithm}}.
#If the PSRF is substantially greater than 1.1 for any parameter, it suggests
#that the chain has not converged.
#Effective Sample Size
effectiveSize(samples_list)
       var1
                var2
## 621.5855 610.7937 569.6906
# Part B
## On reducing the sample size of (0.3,0.3,0.3), the chains have converged properly as seen by the valu
```