Minimum Wage as a Place-Based Policy: Evidence from US Housing Rental Markets

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Minimum wage policies attempt to improve the livelihoods of low-wage workers.

- Increase wages with small effects on employment (e.g., Cengiz et al. 2019)
- Decrease inequality (Autor, Manning, and Smith 2016) and poverty (Dube 2019)

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However, a significant pass-through of MWs to rents may undermine the objectives of the policy.

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Large variation of MW levels in the US even within metropolitan areas.

- Divergence between MW levels at workplace and residence
- Expect spatially heterogeneous effects

This paper

What we do

- Accounting for spatial spillovers, estimate elasticity of rents in the local housing market to workplace MW and residence MW changes
- Estimate share of the extra dollar generated by MW increases pocketed by landlords in each local market

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How we do it

- Propose a novel measure of exposure to MW changes based on commuting shares
- Construct novel dataset of MW policies at ZIP code level
- Exploit high-frequency (month) high-resolution (ZIP code) rents data from Zillow
- Leverage timing and spatial variation in MW changes within metropolitan areas

An initial intuition

Think of a metropolitan area and a MW increase in the business district (CBD).

Partial equilibrium: short term

- Firms producing in the CBD will pay a higher wage. Income redistribution from CBD consumers to low-income workers.
- Income changes are heterogeneous across space because people work and reside in different locations.
- Housing is a normal good, so demand in some areas increases and landlords charge a higher rent.

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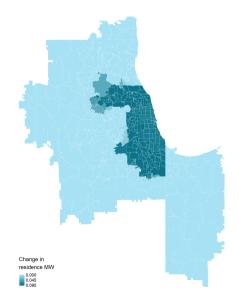
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General equilibrium: long term (Not this paper!)

- People change residence and workplace locations (sorting).
- Developers build more houses (supply response).

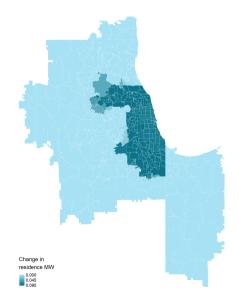
A motivating example



Cook County, IL

- Raised local MW from \$12 to \$13 in July 2019.
- State MW is \$8.25 since 2010, and federal MW is \$7.25 since 2009.

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- State MW is \$8.25 since 2010, and federal MW is \$7.25 since 2009.
- A model where only same-location MW affects rents would miss likely rents increases outside of Cook County

A novel model-based measure of exposure to minimum wages

For ZIP code i and month t we define the workplace MW as

$$\underline{w}_{it}^{\mathsf{wkp}} = \sum_{z \in \mathcal{Z}(i)} \pi_{iz} \ln \underline{W}_{zt} ,$$

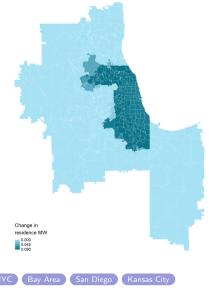
where

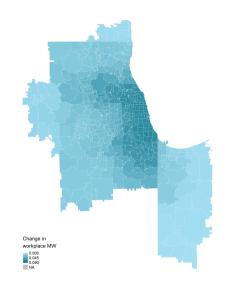
- W_{zt} is statutory MW in z at time t
- $\mathcal{Z}(i)$ are workplace locations of i's residents
- $\pi_{iz} = L_{iz}/L_i$ is the share of i's residents who work in z

The residence MW is simply

$$\underline{\mathbf{w}}_{it}^{\mathsf{res}} = \mathsf{In}\,\underline{W}_{it}$$

A motivating example (continuation)





Preview of findings

Main estimation results

- 10 percent \uparrow in workplace MW \implies 0.55 percent \uparrow in rents
- 10 percent \uparrow in residence MW \implies 0.21 percent \downarrow in rents
- 10 percent \uparrow in both measures \implies 0.34 percent \uparrow in rents

Preview of findings

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Counterfactual increase in federal MW from \$7.25 to \$9 in highly affected areas

- Rent changes vary between -0.4 to 0.75 percent (median 0.5 percent)
- ullet Share pocketed by landlords is between -15 and 17 cents (median 10 cents)

Outline for Today

Partial Equilibrium Model (intuition)

Data

Empirical Strategy and Results

Counterfactual: A federal MW increase

Concluding remarks

Partial Equilibrium Model (intuition)

Overview

Goals of the model

- Stylized answer to what is the effect of MW changes on rents
- Motivate and derive a new measure of exposure to MW

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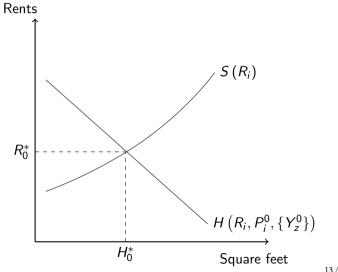
Assumptions

- A higher MW increases income, which increases housing demand
- A higher MW increases non-tradable, which decreases housing demand
- Static model, so residence and workplace locations of workers are fixed

These assumptions are consistent with the literature

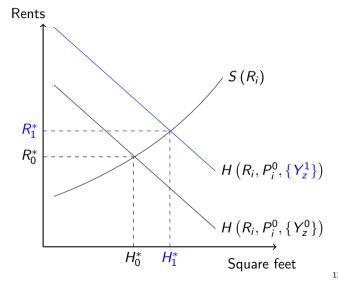
Comparative statics

1. Equilibrium in ZIP code *i*



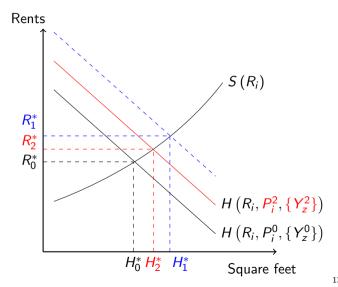
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Comparative statics

- 1. Equilibrium in ZIP code i
- 2. MW increases in some z
- 3. MW increases in i



Representation

In this model, assuming homogeneity across workplace locations of

- 1. elasticity of per-person housing demand to income, and
- 2. elasticity of income to the MW we obtain

$$\Delta \log \text{ rents} = \beta_i \times \Delta \text{workplace MW} + \gamma_i \times \Delta \text{residence MW}$$

Representation

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we obtain

$$\Delta$$
log rents = $\beta_i \times \Delta$ workplace MW + $\gamma_i \times \Delta$ residence MW

Discussion:

- Assumption (1) would hold for homothetic preferences
- In estimation can allow for heterogeneity as long as not correlated with MW changes

Data

Zillow Data

- Leader online real estate and rental platform in the U.S. (more than 110 million homes and 170 million unique monthly users in 2019).
- Provides *median* rents data at ZIP code, county, and state levels at a monthly frequency for several housing categories.

Zillow Data

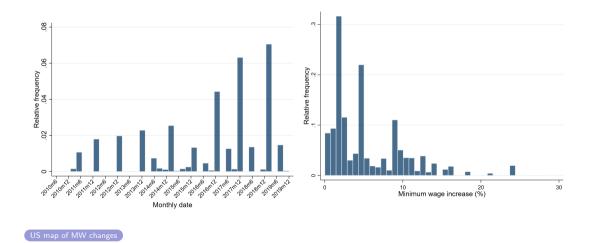
- Leader online real estate and rental platform in the U.S. (more than 110 million homes and 170 million unique monthly users in 2019).
- Provides median rents data at ZIP code, county, and state levels at a monthly frequency for several housing categories.
- We use category single-family, condominium, and cooperative houses (SFCC).
 - Most populated series in Zillow
 - We also estimate our models with other housing categories
- Limitation: Zillow sample is not random.

Zillow ZIP Codes and Population Density

The Statutory MW

- Collect MW data at state, county and city levels between Jan 2010 and Dec 2019.
- Spatial match:
 - Assign census blocks to USPS ZIP codes based on blocks' centroids
 - Add matching of places, counties, and states using census crosswalk
- Assign MWs to each block and define statutory MW as maximum between city, county, state, and federal level.
- Define statutory MW in ZIP code i and month t, \underline{W}_{it} , as weighted average of statutory MWs at block, using housing units as weights.

Distribution of (positive) MW changes



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Constructing the MW measures

Collect data from LEHD Origin-Destination Employment Statistics (LODES) for years 2009–18.

Origin-destination matrices at block level constructed from tax records

Construct **origin-destination matrix** at ZIP code level using spatial match.

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Construct origin-destination matrix at ZIP code level using spatial match.

We observe:

- Number of workers residing in a ZIP code and working in every other ZIP code
- Analogous matrix for number of workers aged less than 29 and earning less than \$1,251

In our baseline specification we use constant commuter shares from 2017.

Results are robust to using other years and groups

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Define the MW measures as

$$\underline{w}_{it}^{\mathsf{wkp}} = \sum_{z \in \mathcal{Z}(i)} \pi_{iz} \ln \underline{W}_{zt}$$
 and $\underline{w}_{it}^{\mathsf{res}} = \ln \underline{W}_{it}$

Empirical Strategy and Results

Empirical model

We estimate versions of the following empirical model:

$$\Delta r_{it} = \delta_t + \beta \Delta \underline{w}_{it}^{\text{wkp}} + \gamma \Delta \underline{w}_{it}^{\text{res}} + \Delta \mathbf{X}_{it}' \eta + \Delta \varepsilon_{it},$$

where

- $r_{it} = \ln R_{it}$
- \mathbf{X}'_{it} are time-varying controls at the county level

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For causal effect of β and γ we need:

- (Rank) Independent variation in MW measures after conditioning on controls
- (Parallel trends) Identification assumption:

$$E\left[\begin{pmatrix} \Delta \underline{\mathbf{w}}_{is}^{\text{wkp}} \\ \Delta \underline{\mathbf{w}}_{is}^{\text{res}} \end{pmatrix} \Delta \varepsilon_{it} \middle| \delta_{t}, \Delta \mathbf{X}_{it} \right] = \begin{pmatrix} 0 \\ 0 \end{pmatrix} \quad \forall s$$

Identification assumption concerns

MW policies are rarely set by considering differential dynamics of the rental housing market within metropolitan areas.

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We extend our model adding leads and lags to test for parallel trends

$$\Delta r_{it} = \delta_t + \sum_{k=-s}^{s} \beta \Delta \underline{w}_{i,t+k}^{\text{wkp}} + \sum_{k=-s}^{s} \gamma \Delta \underline{w}_{i,t+k}^{\text{res}} + \Delta \mathbf{X}_{it}' \eta + \Delta \varepsilon_{it},$$

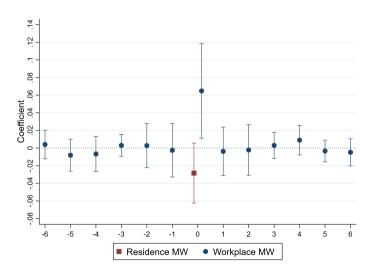
We discuss other robustness checks later.

Main results

	Change wkp. MW $\Delta \underline{w}_{it}^{\mathrm{wkp}}$	Change log rents Δr_{it}			
	(1)	(2)	(3)	(4)	
Change residence MW $\Delta \underline{w}_{it}^{res}$	0.8705 (0.0298)	0.0268 (0.0135)		-0.0207 (0.0171)	
Change workplace MW $\Delta \underline{w}_{it}^{\text{wkp}}$			0.0324 (0.0150)	0.0546 (0.0281)	
Sum of coefficients				0.0339 (0.0153)	
County-quarter economic controls P-value equality	Yes	Yes	Yes	Yes 0.0938	
R-squared Observations	0.9467 131,383	0.0209 131,383	0.0209 131,383	0.0209 131,383	

Note: Standard errors clustered at the state level throughout.

Including leads and lags of workplace MW





Robustness checks and Sample Selection

Concerns about migration

- Literature finds small effects along several years (e.g., Pérez Pérez 2021)
- Use different commuting shares, even allowing them to change yearly

Concerns about parallel trends assumption

- Alternative strategies: "stacked" model and Arellano and Bond (1991)
- Inclusion of non-parametric CBSA trends
- Inclusion of ZIP code-specific parametric trends

Concerns that results are particular to our sample or not generalizable

- Estimate on unbalanced and fully-balanced samples (instead of partially balanced)
- Reweight observations to match characteristics of urban ZIP codes

Sample selection concerns

Counterfactual: A federal MW increase

Overview

Entire commuting structure determines the incidence of MW policies.

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Consider an increase of the federal MW to \$9 in January 2020.

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How much out of each extra dollar is captured by landlords?

Share pocketed by landlords

Define the share pocketed by landlords as

$$\rho_i := \frac{\Delta H_i}{\Delta Y_i} = \frac{H_i^{\mathsf{Post}} R_i^{\mathsf{Post}} - H_i^{\mathsf{Pre}} R_i^{\mathsf{Pre}}}{\Delta Y_i}$$

where "Pre" and "Post" indicate moments before and after the increase.

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where "Pre" and "Post" indicate moments before and after the increase.

Change in rented space are unobserved. We assume $H_i^{\mathsf{Pre}} = H_i^{\mathsf{Post}} = H_i$ so

$$\rho_i = \frac{H_i^{\text{Post}} R_i^{\text{Post}} - h_i^{\text{Pre}} R_i^{\text{Pre}}}{\Delta Y_i} = h_i \frac{\Delta R_i}{\Delta Y_i}$$

If $\Delta H_i > 0$ then our estimate of ρ_i is a lower bound.

Share pocketed under the model

According to the model,

$$\Delta r_i = \beta \Delta \underline{\mathbf{w}}_i^{\mathsf{wkp}} + \gamma \underline{\mathbf{w}}_i^{\mathsf{res}}$$

We also define, for $y_i = \ln Y_i$,

$$\Delta y_i = \varepsilon \Delta \underline{\mathbf{w}}_i^{\mathsf{wkp}}$$

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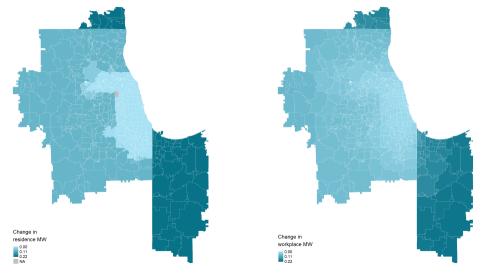
Algebra implies

$$\rho_{i} = \alpha_{i} \left[\frac{\exp\left(\beta \underline{\mathcal{W}}_{i}^{\mathsf{wkr}} + \gamma \underline{\mathcal{W}}_{i}^{\mathsf{res}}\right) - 1}{\exp\left(\varepsilon \underline{\mathcal{W}}_{i}^{\mathsf{wkr}}\right) - 1} \right]$$

where $\alpha_i = (H_i R_i) / Y_i$ is the share of *i*'s expenditure in housing.

We use our estimates to compute ρ_i for urban ZIP codes that are located in affected CBSAs.

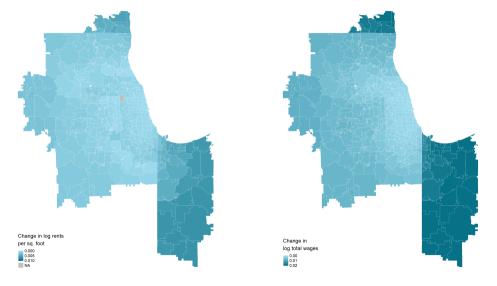
Changes in residence and workplace MWs



Residence MW Workplace MW

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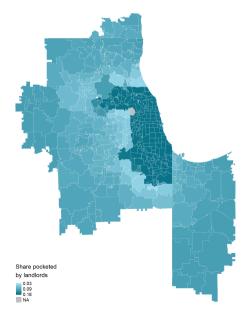
Estimated changes in per-square-foot rents and total wages



Changes in log rents per sqft.

Changes in log total wages

Share pocketed by landlords



The incidence of MW changes on average

		Average change in		Avg. share pocketed		
	N	Res. MW	Wkp. MW	s = 0.25	s = 0.45	
Effect in ZIP codes with						
previous MW \leq \$9	5,882	0.161	0.153	0.075	0.136	
previous MW > \$9	1,070	0.000	0.017	0.126	0.227	

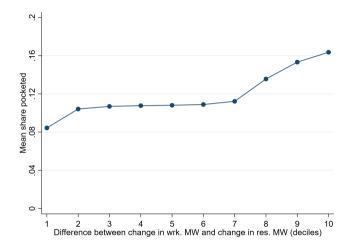
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More generally, one can think of the effect for different values of

$$\Delta \underline{w}_i^{\mathsf{wkp}} - \Delta \underline{w}_i^{\mathsf{res}}$$

The incidence of MW changes according to intensity of treatment



Notes: The figure shows computations of the share pocketed for the following parameters: $\beta=0.0546$, $\gamma=-0.0207$. $\varepsilon=0.1083$, and $\alpha=0.35$.

Concluding remarks

Conclusion

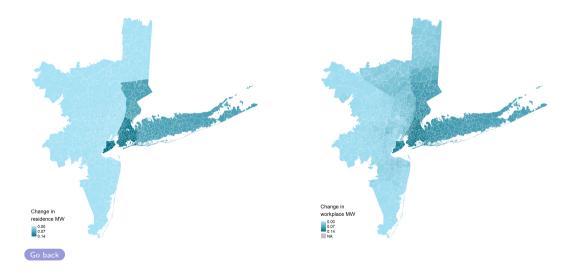
- When studying effects of place-based policies on housing market must account for divergence between workplace and residence locations
- In the case of the MW, hikes in workplace locations *increase* rents whereas hikes in residence locations *decrease* rents
- Even with a two-parameter model we are able to describe and predict rich spatial patterns in rent changes
- Landlords pocket a non-negligible fraction of the income increase generated by the MW

Thank You!

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Appendix

New York (MW changes in January 2019)



Bay area (MW changes in January 2019)



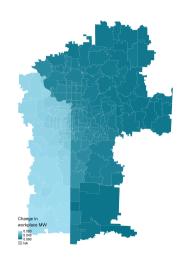
San Diego (MW changes in January 2019)



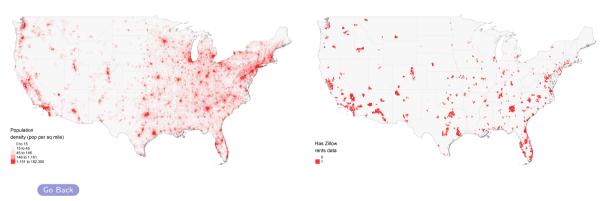


Kansas City (MW changes in January 2019)

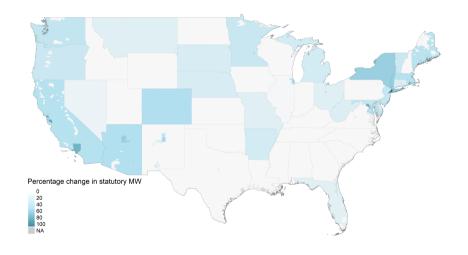




Comparison between Zillow Sample and Population Density

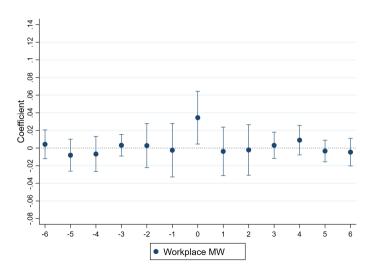


MW changes between Jan 2010 and Dec 2019, mainland US



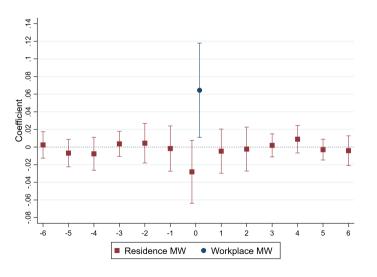


Excluding residence MW



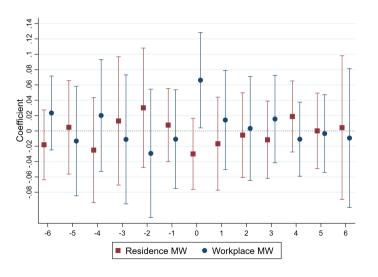


Including leads and lags of residence MW





Including leads and lags of both MW measures





Sample selection concerns

	Change log rents Δr_{it}						
	Baseline (1)	Baseline Reweighted (2)	Fully-balanced (3)	Fully-balanced Reweighted (4)	Unbalanced (5)	Unbalanced Reweighted (6)	
Change residence MW $\Delta \underline{w}_{it}^{\text{res}}$	-0.0207 (0.0171)	-0.0186 (0.0309)	-0.0201 (0.0200)	-0.0223 (0.0307)	-0.0254 (0.0210)	-0.0168 (0.0204)	
Change workplace MW $\Delta \underline{w}_{it}^{\mathrm{wkp}}$	0.0546 (0.0281)	0.0581 (0.0427)	0.0682 (0.0306)	0.0868 (0.0389)	0.0471 (0.0309)	0.0393 (0.0369)	
P-value equality	0.0938	0.2863	0.0792	0.1101	0.1559	0.3146	
R-squared	0.0209	0.0185	0.0216	0.0180	0.0160	0.0127	
Observations	131,383	130,533	78,912	78,381	193,292	192,177	



Estimates of the effect of the MW on total wages in a ZIP code

		Log total wages					
	(1)	(2)	(3)	(4)	(5)		
Workplace MW	0.1488 (0.0704)	0.1112 (0.0405)	0.1083 (0.0390)	0.1310 (0.0917)	0.0262 (0.0841)		
Sample	All	All	All	Baseline	All		
Economic controls	No	Yes	Yes	Yes	Yes		
CBSA $ imes$ year FE	No	No	Yes	Yes	Yes		
Within R-squared	0.0165	0.1395	0.0266	0.0376	0.0018		
Observations	274,271	247,962	247,852	12,943	235,193		

Notes: unit of observation is ZIP code by year pairs. All regressions include ZIP code FE and year FE. Workplace MW measure is yearly average of monthly 2017 variable.

