



# CHANGE POINT DETECTION IN END-TO-END MEASUREMENTS TIME SERIES

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# Contents

<b>List of Figures</b>	<b>vii</b>
<b>List of Tables</b>	<b>viii</b>
<b>1 Introduction</b>	<b>1</b>
1.1 Difficulties . . . . .	1
1.2 Contributions . . . . .	1
1.3 Dissertation Outline . . . . .	1
<b>2 Change Point Detection</b>	<b>2</b>
2.1 Problem Definition . . . . .	2
2.2 Notation . . . . .	3
2.3 Sliding Windows . . . . .	4
2.4 Optimization Model . . . . .	5
2.5 HMM (Hidden Markov Model) . . . . .	6
2.6 Bayesian Inference . . . . .	7
<b>3 Dataset</b>	<b>9</b>
3.1 Description of End-to-End Packet Loss Measurements Time Series . .	9
3.2 Change Points Classification Survey . . . . .	9
3.3 Description of Change Points Dataset . . . . .	9
3.4 Performance Evaluation . . . . .	9
<b>4 Conclusions</b>	<b>10</b>
<b>Bibliography</b>	<b>11</b>

# List of Figures

2.1	Toy example of a sliding windows method. . . . .	5
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# List of Tables



# Chapter 1

## Introduction

### 1.1 Difficulties

### 1.2 Contributions

### 1.3 Dissertation Outline

# Chapter 2

## Change Point Detection

A change point detection algorithm is concerned to identify points in time where the statistical properties of a time series have changed. This problem have a broad application in different knowledge fields, and in general, the algorithms performance are closely related with the input characteristics. Also, when the latent information of the procedures that generated a time series is missing the target statistical properties can be considered subjective, bringing difficulties not only in the detection phase but also in the problem formalization.

In this context this chapter specifies the problem and briefly discusses several change point detection algorithms. The literature of this area is extensive, and it is common to find methods that presents a poor performance due to a variety of reasons, such as they are too specific or because the mechanisms were only analyzed through theoretical aspects. Therefore, it was chosen a set of methods that can provide a good practical and theoretical perspective, and also flexibility to insert adaptations that can better handle some input peculiarities. Furthermore, through this chapter is exposed several obstacles when dealing with real data, and some adopted solutions which are not described in literature.

### 2.1 Problem Definition

The problem can be categorized in offline or online. In the offline version, to decide if a specific point at time  $t$  is a change point the solver has available the whole time series, including past and future information on  $t$ . In the other hand, in the online version the information is available up to time  $t$ . The choice between these options is defined by the application domain, in some cases data are processed in real time and the change points should be detected as soon as possible, but in others the changes are identified by historical purposes and offline algorithms can be used.

It is intuitive that the offline case is more robust, since there are more information to make a classification. In practice, to increase the statistical confidence of a

decision the online definition is relaxed, and to decide if a point at time  $t$  is a change point it is possible to use data up to a small window in the future of  $t$ , which in real time processing means that the application should wait until more data are available. This trick plays a trade-off between minimizing the time to detect a change and correctly classify a point. Therefore, in some cases, the online version can be transformed in the offline case by only modifying the input availability.

In this work it is considered the following input and change points characteristics, which were defined considering the final application scenario:

- Univariate time series. However, it is possible to extend several methods presented here to deal with multivariate data.
- Unevenly time series, that is, data is not regularly sampled in time.
- Time series with different lengths.
- Unknown number of change points.
- Different number of points between change points.
- Focus on changes in the underlying mean and distribution, disregarding other kinds of changes such as in periodicity.
- Outliers are not considered statistical changes.
- There is no latent information of the time series.
- It is considered the online and offline options.

## 2.2 Notation

An univariate time series composed of  $n$  points is defined by two vectors,  $\mathbf{x} = (x_1, \dots, x_n)$  and  $\mathbf{y} = (y_1, \dots, y_n)$ . The value  $y_i$  indicates the  $i$ -th sampled value and  $x_i$  indicates the associated sample time. It is assumed that the points are sorted by time, that is,  $x_{i-1} < x_i$  for  $i = 2, \dots, n$ . Since unevenly time series is considered,  $x_i - x_{i-1}$  can be different for different  $i$  values. For  $s \geq t$  the following convention is adopted:  $\mathbf{y}_{s:t} = (y_s, \dots, y_t)$ .

The presence of  $k$  change points implies that data is split into  $k + 1$  segments, also called windows. Let  $\tau_i$  indicates the  $i$ -th change point for  $i = 1, \dots, k$ . Also let  $\tau_0 = 0$ ,  $\tau_{k+1} = n$  and  $\boldsymbol{\tau} = (\tau_0, \dots, \tau_{k+1})$ . Then, the  $i$ -th segment is defined by  $\mathbf{y}_{\tau_{i-1}+1:\tau_i}$ , assuming that  $\tau_{i-1} < \tau_i$  for  $i = 1, \dots, k + 1$ .

Through the previous definitions, change point detection algorithms mainly aim to find both  $k$  and  $\boldsymbol{\tau}$ .

## 2.3 Sliding Windows

Sliding windows techniques use two sliding windows over the time series, and reduce the problem of detecting change points to the problem of testing whether data from the segments were generated by different distributions. One approach is to consider a distance metric between two empirical distributions as the base to infer the change points. Letting  $d(\mathbf{a}, \mathbf{b})$  be the distance between two empirical distributions defined by the windows  $\mathbf{a}$  and  $\mathbf{b}$ , and considering windows of length  $m$ , the Algorithm 1 presents a simple sliding windows method.

---

**Algorithm 1** Sliding Windows

---

```

1:  $i \leftarrow 1$ 
2: while  $i + 2m - 1 \leq n$  do
3:   if  $d(\mathbf{y}_{i:i+m-1}, \mathbf{y}_{i+m:i+2m-1}) > \alpha$  then
4:     Report  $i + m - 1$  as a change point
5:      $i \leftarrow i + m$ 
6:   else
7:      $i \leftarrow i + 1$ 
8:   end if
9: end while

```

---

In this method, when the distance between the distributions is above some threshold  $\alpha$  a change point is reported. This is a common approach for an online application, however, it is possible to increase the classification accuracy in offline cases. As an example, the top plot of figure 2.1 presents a simulated time series, the segment  $\mathbf{y}_{1:1000}$  was generated sampling a  $N(1, 0.2)$  distribution, and  $\mathbf{y}_{1001:2000}$  was sampled through  $N(5, 0.2)$ . The distribution of a window was constructed binning the data with bins of size 0.02. The bottom plot of the same figure presents the associated Hellinger distance [1] between two sliding windows, where the point  $(i, H_i)$  represents the distance between the windows  $\mathbf{y}_{i-100:i-1}$  and  $\mathbf{y}_{i:i+99}$ .

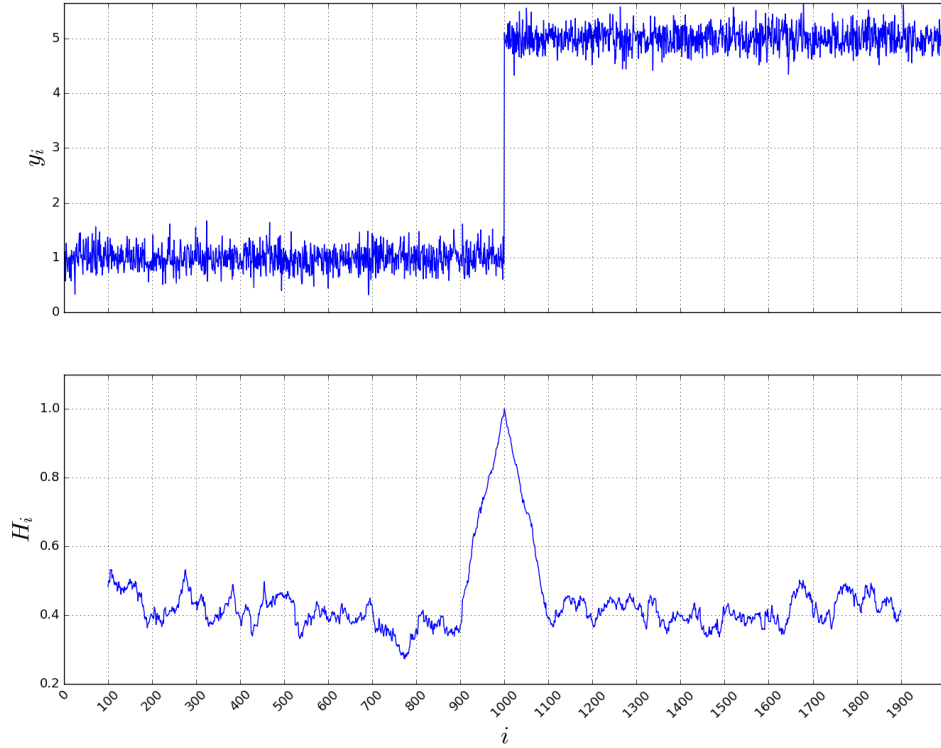


Figure 2.1: Toy example of a sliding windows method.

It can be observed that there is a peak on the distance in the exact location where the distribution changed. However, using only the threshold method it is possible to prematurely infer the position of the change point, therefore, an alternative is to also use a peak detection algorithm. Besides, the distance function choice has a direct impact on the classification accuracy.

As stated in [2], a performance improvement can be achieved concurrently executing the same sliding windows algorithm, however, with different windows lengths, which facilitates the detection of segments with distinct number of points.

## 2.4 Optimization Model

Given a fixed value of  $k$ , one approach is to define a cost function that measures the homogeneity of a window, and therefore, choose the change points that globally optimize this homogeneity. Let the cost of the  $i$ -th segment be defined as  $C(\mathbf{y}_{\tau_{i-1}+1:\tau_i})$ , then the cost of a segmentation is the sum of all segments costs.

A common choice for the function  $C$  is the MSE (Mean Squared Error), which can capture changes in the mean. Another usual approach is to consider distribution

changes through negative maximum log-likelihood functions, considering that data within a window is iid.

Therefore, given a fixed  $k$ , the optimal segmentation is obtained through the following optimization problem, which is called the constrained case [3]:

$$\min_{\tau_{1:k}} \sum_{i=1}^{k+1} C(\mathbf{y}_{\tau_{i-1}+1:\tau_i}) \quad (2.1)$$

This problem can be solved using dynamic programming with  $O(kn^2f(n))$  time complexity, where  $f(n)$  is related with  $C$  evaluation. Several segment cost functions can be evaluated in  $O(1)$  after a  $O(n)$  preprocessing phase, implying in an overall  $O(kn^2)$  complexity. It is possible to proof that MSE, negative maximum log-likelihood functions of normal, exponential, poisson and binomial distributions have this characteristic. Also, the formulation can consider a minimum value of a window length.

Modeling segments with continuous distributions can lead to practical difficulties. One of them is the fact that segments can form degenerate distributions, that is, the data of a window can have zero variance, which is always the case of unitary length segments. In these scenarios the negative maximum log-likelihood is undefined. Two approaches can be used to overcome this situation. The first one tries to avoid degenerate segments adding a white noise with small variance to the time series. The second one considers that the cost of any degenerate distribution is equal to a constant.

When the number of change points is unknown an usual way is to introduce a non decreasing penalty function  $g(k)$ . Then, the new optimization problem, called penalized case [3], is:

$$\min_{k, \tau_{1:k}} \sum_{i=1}^{k+1} C(\mathbf{y}_{\tau_{i-1}+1:\tau_i}) + g(k) \quad (2.2)$$

This problem can be solved in  $O(Kn^2f(n))$ . However, if the penalty function is linear in  $k$ , the problem can be formulated more efficiently and solved in  $O(n^2f(n))$ .

Also, there are several pruning algorithms to speedup the computation [3–5], in general trying to reduce the  $\tau$  search space but maintaining optimality.

## 2.5 HMM (Hidden Markov Model)

The idea that each segment is associated with a specific latent configuration of the procedure that generated the time series has a direct interpretation to a HMM model [6–8]. In this context, each window is related to a hidden state of a HMM, and the

observation distribution of this state represents the distribution of that segment. Therefore, the mechanism models the time series using a HMM, and through the hidden state path assesses the times when a transition between different hidden states occur.

There are several approaches in the detection and training phases. For example, given a trained HMM, is possible to analyze the most probable hidden state path that a time series can follow through the viterbi algorithm. Also, it is possible to evaluate the probability of a transition between different hidden states at time  $t$ , and then apply a threshold and peak detection methods, as well as in sliding windows techniques. For the training step, it is possible to use several time series to train a single HMM, and then use this model to detect change points in all time series. Another way is to, for each time series, train a single model using only the target time series.

It is important to note that the structure of the hidden state graph has a large impact on the performance. Using a fully connected graph, the number of states defines the maximum number of distribution configurations. Employing a left to right structure, the number of hidden states will induce the maximum number of segments.

In [8] is stated that when using a fully connected structure, the time interval that a time series stays in the same hidden state is low, which can not reflect real data. To overcome this problem, [8] suggests to increase the time that a time series stands in the same hidden state, using a dirichlet prior regularization.

## 2.6 Bayesian Inference

There are several Bayesian methods with the objective to assess the probability that a point is a change point. Following an offline fashion, the work of [9] recursively calculates, for each  $i$ , the probability of  $\mathbf{y}_{i:n}$  given a change point at  $i$ . With these probabilities is possible to simulate the time of the first change point, and then, compute the conditional distribution of the time of the second change given the first, and so on. To achieve this, the mechanism assumes that observations are independents, and that each segment is modeled by conjugate priors. Also, the procedure considers priors to model the number of changes and the time between two consecutive change points. The overall complexity of this method is  $O(n^2)$ , considering that the likelihood of a segment can be evaluated in  $O(1)$ .

In [10] it is also considered that parameters of different segments are independents, and that data within a window is iid. However, through an online mode, the procedure is concerned with the estimation of the distribution of the length of the current time since the last change point, called run length, given the data so

far observed. To achieve this, the method assumes the probability of current run length given the last run length as a prior. Assuming exponential-family likelihoods to model a segment, the time complexity to process a point is linear in the number of points already observed.



# Chapter 3

## Dataset

- 3.1 Description of End-to-End Packet Loss Measurements Time Series
- 3.2 Change Points Classification Survey
- 3.3 Description of Change Points Dataset
- 3.4 Performance Evaluation

# Chapter 4

## Conclusions

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