# 11-641: Homework 1

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### 1: Hyperplane properties.

(a) Suppose any two points  $x_1$ ,  $x_2$  on hyperplane h, we can represent vector on h as  $\vec{x_1} - \vec{x_2}$  According to hyperplane defination

$$\vec{w} \cdot (\vec{x_1} - \vec{x_2}) = w^T x_1 - w^T x_2 = b - b = 0$$

Since  $x_1$ ,  $x_2$  are any two points on h, we can prove vector  $\vec{w}$  is perpendicular to hyperplane h. Then the shortest distance from the origin to hyperplane h can be defined as the absolute value of projection of vector  $\vec{x}$  on vector  $\vec{w}$ , where vector  $\vec{x}$  is vector from origin to any point on h

$$d = \frac{\vec{x} \cdot \vec{w}}{||w||} = \frac{|b|}{||w||}$$

(b) Suppose any point  $x_0$  on hyperplane h, we can represent vector from  $x_0$  to point x as  $\vec{x} - \vec{x_0}$ Since vector  $\vec{w}$  is perpendicular to hyperplane h, the projection of  $x_0 \to x$  on vector  $\vec{w}$  is

$$P = \frac{(\vec{x} - \vec{x_0}) \cdot \vec{w}}{||w||} = \frac{\vec{x} \cdot \vec{w} - \vec{x_0} \cdot \vec{w}}{||w||} = \frac{w^T x - b}{||w||}$$

The perpendicular distance from point x to h can be defined as the absolute value of projection of  $x_0 \to x$  on vector  $\vec{w}$ , and projection value P will be smaller than 0 if point x is on the same side with original, so perpendicular distance d is

$$d = \frac{y(w^T x - b)}{||w||}, \ y \in \{-1, 1\}$$

## 2: Eigenvalues and eigenvectors.

(a) For matrix A

$$A = \begin{pmatrix} 3 & 1 \\ 1 & 3 \end{pmatrix}$$

The eigenvectors  $\mathbf{v}$  of transformation satisfy the equation  $A\mathbf{v} = \lambda \mathbf{v}$ , rearrange this equation to obtain  $(A - \lambda I)\mathbf{v} = 0$ . Set the determinant to zero to obtain the polynomial equation

$$p(\lambda) = |A - \lambda I| = 8 - 6\lambda + \lambda^2 = 0$$

So we can calculate the roots as  $\lambda_1 = 2$  and  $\lambda_2 = 4$ For  $\lambda_1 = 2$ , the equation becomes

$$(A - 2I)\mathbf{v_1} = \begin{pmatrix} 1 & 1 \\ 1 & 1 \end{pmatrix} \begin{pmatrix} v_1 \\ v_2 \end{pmatrix} = \begin{pmatrix} 0 \\ 0 \end{pmatrix}$$

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which has the solution,

$$\mathbf{v_1} = \begin{pmatrix} 1 \\ -1 \end{pmatrix}$$

For  $\lambda_2 = 4$ , the equation becomes

$$(A-4I)\mathbf{v_2} = \begin{pmatrix} -1 & 1\\ 1 & -1 \end{pmatrix} \begin{pmatrix} v_1\\ v_2 \end{pmatrix} = \begin{pmatrix} 0\\ 0 \end{pmatrix}$$

which has the solution,

$$\mathbf{v_2} = \begin{pmatrix} 1 \\ 1 \end{pmatrix}$$

Thus, the vectors  $\mathbf{v_1}$  and  $\mathbf{v_2}$  are eigenvectors of A associated with the eigenvalues  $\lambda_1 = 2$  and  $\lambda_2 = 4$ , respectively.

(b) Define the matrix P composed of eigenvectors

$$P = \begin{pmatrix} \mathbf{v_1} & \mathbf{v_2} \end{pmatrix}$$

Define diagonal matrix D composed of eigenvalues

$$D = \begin{pmatrix} \lambda_1 & 0 \\ 0 & \lambda_2 \end{pmatrix}$$

Then AP = PD, so  $A = PDP^{-1}$ So we can prove that

$$A^{2} = (PDP^{-}1)(PDP^{-}1)$$

$$= PD(P^{-}1P)DP^{-}1$$

$$= PD^{2}P^{-}1$$
(1)

By induction, we can further prove  $A^k = PD^kP^{-1}$  where

$$D^k = \begin{pmatrix} \lambda_1^k & 0\\ 0 & \lambda_2^k \end{pmatrix}$$

So the eigenvalues of  $A^k$  are  $\lambda_1^k$ ,  $\lambda_2^k$ , the kth powers of the eigenvalues of matrix A, and that each eigenvector of A is still an eigenvector of  $A^k$ .

#### 3: Maximum likelihood estimate

(a) For a binomial process of coin tossing

$$L(k \text{ headup out of } n \mid p) = f(k \text{ headup out of } n \mid p)$$

$$= {n \choose n} \cdot p^k \cdot (1-p)^{n-k}$$
(2)

$$F = \ln(L(k \text{ headup out of } n \mid p))$$

$$= const + k \ln p + (n - k) \ln(1 - p)$$
(3)

So for maximum likelihood,

$$\hat{p} = \arg\max_{p} F(p)$$

Set derivative  $\frac{\partial F}{\partial p} = 0$ 

$$\frac{k}{p} - \frac{n-k}{1-p} = 0$$

Thus we can get  $\hat{p} = \frac{k}{n}$ 

(b) For a multinomial process

$$L(n_1, \dots, n_m|p) = f(n_1, \dots, n_m|p) = \binom{n}{n_1 \dots n_m} \cdot \prod_j p_j^{n_j}$$

The log-likelihood is

$$F = \ln(L(n_1, \cdots, n_m|p)) = const + \sum_j n_j \ln p_j$$

Since  $n_1 + n_2 + \cdots + n_m = n$ ,  $p_1 + p_2 + \cdots + p_m = 1$ , we use Lagrange multiplier to maximize this function

$$F' = \sum_{j} n_{j} \ln p_{j} + \lambda (1 - \sum_{j} p_{j})$$
$$\frac{\partial F'}{\partial p_{j}} = \frac{n_{j}}{p_{j}} - \lambda = 0$$
$$\frac{\partial F'}{\partial \lambda} = (1 - \sum_{j} p_{j}) = 0$$

So

$$n_{j} = \lambda \cdot p_{j}$$

$$\sum_{j} n_{j} = \lambda \cdot \sum_{j} p_{j}$$

$$n = \lambda$$

$$\hat{p}_{j} = \frac{n_{j}}{n}$$

References: http://www.cs.ubc.ca/murphyk/Teaching/CS340-Fall06/reading/bernoulli.pdf

## 4: Calculus

(a) For  $u = \frac{1}{1 + e^{-x}}$ , make  $Y = e^{-x}$ 

$$\frac{du}{dx} = \frac{d\frac{1}{1+Y}}{dY} \cdot \frac{dY}{dx} 
= -\frac{1}{(1+e^{-x})^2} \cdot (-e^{-x}) 
= \frac{1}{1+e^{-x}} \cdot \frac{e^{-x}}{1+e^{-x}} 
= u(1-u)$$
(4)

(b) Gradient of l

$$\nabla l = \frac{dl}{du} \cdot \frac{du}{dz} \cdot \nabla z$$

$$= \left(\frac{y}{u} - \frac{1 - y}{1 - u}\right) \cdot u(1 - u) \cdot (1, x_1, \dots, x_m)^T$$

$$= (y - u) \cdot (1, x_1, \dots, x_m)^T$$

$$= (y - u, x_1(y - u), \dots, x_m(y - u))^T$$
(5)

(c) Pairwise 2nd order derivative  $H_{jj'}$ 

$$\nabla \nabla l = \begin{pmatrix} \frac{\partial^{2}l}{\partial w_{0}\partial w_{0}} & \frac{\partial^{2}l}{\partial w_{1}\partial w_{0}} & \cdots & \frac{\partial^{2}l}{\partial w_{m}\partial w_{0}} \\ \frac{\partial^{2}l}{\partial w_{0}\partial w_{1}} & \frac{\partial^{2}l}{\partial w_{1}\partial w_{1}} & \cdots & \frac{\partial^{2}l}{\partial w_{m}\partial w_{1}} \\ \cdots & \cdots & \ddots & \cdots \\ \frac{\partial^{2}l}{\partial w_{0}\partial w_{m}} & \frac{\partial^{2}l}{\partial w_{1}\partial w_{m}} & \cdots & \frac{\partial^{2}l}{\partial w_{m}\partial w_{m}} \end{pmatrix}$$

$$= \begin{pmatrix} u(u-1) & x_{1}u(u-1) & \cdots & x_{m}u(u-1) \\ x_{1}u(u-1) & x_{1}^{2}u(u-1) & \cdots & x_{1}x_{m}u(u-1) \\ \cdots & \cdots & \cdots & \cdots \\ x_{m}u(u-1) & x_{1}x_{m}u(u-1) & \cdots & x_{m}^{2}u(u-1) \end{pmatrix}$$

$$(6)$$

(d) The cost log-likelihood function in logistic regression is defined as l shown in previous question. By defination of concave, a differentiable function l is concave on an interval if its derivative function l' is monotonically decreasing on that interval: a concave function has a decreasing slope.

Thus for a twice-differentiable function l, if the second derivative l'' is negative, then the graph is concave.

According to previous prove,  $\frac{\partial^2 l}{\partial w_j \partial w_j} = x_j^2 u(u-1)$ , which is the diagonal values of the pairwise 2nd order derivative matrix  $H_{jj'}$ , since  $0 \le u \le 1$ , so  $\frac{\partial^2 l}{\partial w_j \partial w_j} \le 0$ . Then we can prove the log-likelihood cost function of logestic regression is concave.