Macroeconometrics

Lecture 15 Modeling effects of monetary policy

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A small-open economy model revisited

Modeling effects of monetary policy in the US

Modeling effects of monetary policy in the US

extended analysis

Materials

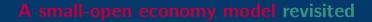
A zip file L15 mcxs-all.zip for the reproduction of the results

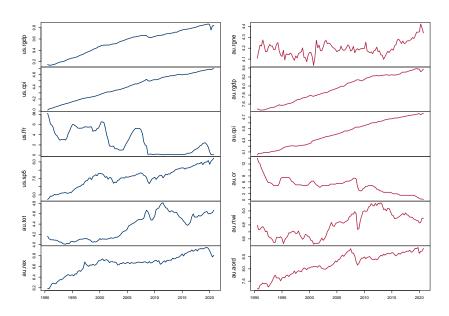
Objectives.

- ► To present the outcomes of Waggoner & Zha (2003) estimation algorithm
- ► To work with a variation on the small-open economy model
- ► To investigate the effects of the monetary policy shock in the U.S.

Learning outcomes.

- ▶ Implementing rotations in the estimation outcomes
- Specifying classical monetary policy models
- Emphasising prior specification sensitivity





$$\begin{bmatrix} B_{0.11} & \mathbf{0}_{6 \times 6} \\ B_{0.21} & B_{0.22} \end{bmatrix} \begin{bmatrix} y_t^f \\ y_d^d \end{bmatrix} = \begin{bmatrix} b_{0.1} \\ b_{0.2} \end{bmatrix} + \begin{bmatrix} B_{1.11} & B_{1.12} \\ B_{1.21} & B_{1.22} \end{bmatrix} \begin{bmatrix} y_{t-1}^f \\ y_{t-1}^d \end{bmatrix} + \dots + \begin{bmatrix} u_t^f \\ u_t^d \end{bmatrix}$$

$$y_t^{f\prime} = \begin{bmatrix} rgdp_t & cpi_t & FFR_t & sp500_t & tot_t & rex_t \end{bmatrix}$$

$$y_t^{d\prime} = \begin{bmatrix} rgne_t & rgdp_t & cpi_t & CR_t & rtwi_t & aord_t \end{bmatrix}$$

$$u_t^{f\prime} = \begin{bmatrix} u_{1.t} & u_{2.t} & u_{3.t}^{us.mps} & u_{4.t} & u_{5.t} & u_{6.t} \end{bmatrix}$$

$$u_t^{d\prime} = \begin{bmatrix} u_{7.t} & u_{8.t} & u_{9.t} & u_{10.t}^{au.mps} & u_{11.t} & u_{12.t} \end{bmatrix}$$

Foreign block.

 $rgdp_t$ – real GDP, cpi_t – CPI, FFR_t – federal funds rate, $sp500_t$ – S&P 500 index, tot_t – Australian terms of trade, rex_t – Australian real export

Australian block.

 $rgne_t$ - real gross national expenditure, $rgdp_t$ - real GDP, cpi_t - CPI, CR_t - cash rate, $rtwi_t$ - real trade weighted index, $aord_t$ - All Ordinaries Index

Shocks of interest.

 $u_{10.t}^{au.mps}$ – Australian monetary policy shock $u_{3.t}^{us.mps}$ – US monetary policy shock

$$\begin{bmatrix} B_{0.11} & \mathbf{0}_{6\times 6} \\ B_{0.21} & B_{0.22} \end{bmatrix} \begin{bmatrix} y_t^f \\ y_t^d \end{bmatrix} = \begin{bmatrix} b_{0.1} \\ b_{0.2} \end{bmatrix} + \begin{bmatrix} B_{1.11} & B_{1.12} \\ B_{1.21} & B_{1.22} \end{bmatrix} \begin{bmatrix} y_{t-1}^f \\ y_{t-1}^d \end{bmatrix} + \dots + \begin{bmatrix} u_t^f \\ u_t^d \end{bmatrix}$$

SVAR for a small-open economy

 $B_{0.11}$ – identification of foreign shocks: lower-triangular matrix

 $B_{0.22}$ – identification of domestic shocks: lower-triangular matrix

 $B_{0.12} = \mathbf{0}_{6 \times 6}$ – small-open economy assumption

 $B_{l.12} = \mathbf{0}_{6 \times 6}$ – small-open economy assumption (not imposed)

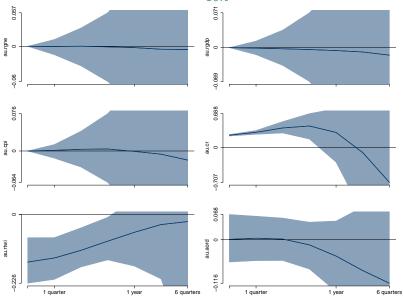
 $B_{0.21}$ – small-open economy assumption: foreign shocks affect domestic variables

$$\begin{bmatrix} B_{0.11} & \mathbf{0}_{6\times 6} \\ B_{0.21} & B_{0.22} \end{bmatrix} \begin{bmatrix} y_t^f \\ y_t^d \end{bmatrix} = \begin{bmatrix} b_{0.1} \\ b_{0.2} \end{bmatrix} + \begin{bmatrix} B_{1.11} & B_{1.12} \\ B_{1.21} & B_{1.22} \end{bmatrix} \begin{bmatrix} y_{t-1}^f \\ y_{t-1}^d \end{bmatrix} + \dots + \begin{bmatrix} u_t^f \\ u_t^d \end{bmatrix}$$

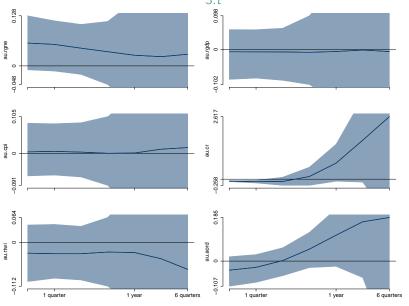
SVAR for a small-open economy

- ▶ The model was estimated using Waggoner & Zha (2003) algorithm
- ▶ \underline{B} and $\underline{\Omega}$ prior matrices are set the same way as the RF Minnesota prior with $\kappa_1 = \kappa_2 = 1$, $\kappa_3 = 0.95$
- ▶ $\underline{S} = \kappa_4 I_N$ and $\kappa_4 = 1$, $\underline{\nu} = N$ which make the generalized-normal prior for B_0 set to $\mathcal{N}(\mathbf{0}_N, I_N)$
- lacktriangle The results are sensitive to the specification of κ_1 and κ_4
- ▶ The Gibbs sampler for B_0 used $S_1 = 100$ draws in the burin-in and $S_2 = 5000$ in the final sampler
- ► The potential extensions include: estimation of prior hyper-parameters and heteroskedasticity of the structural shocks

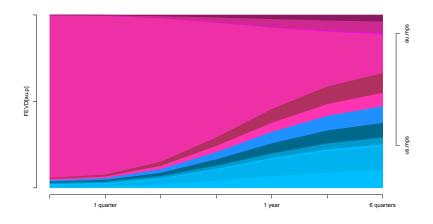
IRFs of domestic sector to $u_{10.t}^{au.mps}$



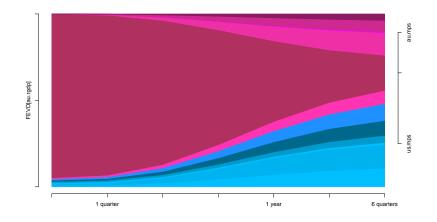
IRFs of domestic sector to $u_{3.t}^{us.mps}$



Forecast error variance decomposition of $au.cpi_{t+h|t}$



Forecast error variance decomposition of $au.rgdp_{t+h|t}$



$$\begin{bmatrix} B_{0.11} & \mathbf{0}_{6 \times 6} \\ B_{0.21} & B_{0.22} \end{bmatrix} \begin{bmatrix} y_t^f \\ y_t^d \end{bmatrix} = \begin{bmatrix} b_{0.1} \\ b_{0.2} \end{bmatrix} + \begin{bmatrix} B_{1.11} & B_{1.12} \\ B_{1.21} & B_{1.22} \end{bmatrix} \begin{bmatrix} y_{t-1}^f \\ y_{t-1}^d \end{bmatrix} + \dots + \begin{bmatrix} u_t^f \\ u_t^d \end{bmatrix}$$

$$y_t^{f'} = \begin{bmatrix} rgdp_t & cpi_t & FFR_t & sp500_t & tot_t & rex_t \end{bmatrix}$$

$$y_t^{d'} = \begin{bmatrix} rgne_t & rgdp_t & cpi_t & CR_t & rtwi_t & aord_t \end{bmatrix}$$

Shocks of interest.

 u_t^f – foreign shocks

 u_t^d – domestic shocks

Objective.

To identify to what extent the foreign shocks determine the business cycle in Australia jointly

$$\begin{bmatrix} B_{0.11} & \mathbf{0}_{6\times 6} \\ B_{0.21} & B_{0.22} \end{bmatrix} \begin{bmatrix} y_t^f \\ y_t^d \end{bmatrix} = \begin{bmatrix} b_{0.1} \\ b_{0.2} \end{bmatrix} + \begin{bmatrix} B_{1.11} & B_{1.12} \\ B_{1.21} & B_{1.22} \end{bmatrix} \begin{bmatrix} y_{t-1}^f \\ y_{t-1}^d \end{bmatrix} + \dots + \begin{bmatrix} u_t^f \\ u_t^d \end{bmatrix}$$

SVAR for a small-open economy.

 B_{11} , B_{22} – unrestricted

 $B_{12} = \mathbf{0}_{6 \times 6}$ – small-open economy assumption

 $B_{0.21}$ – small-open economy assumption: unrestricted

Identification.

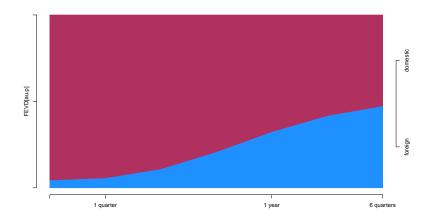
The model is identified up to a block diagonal rotation matrix

$$Q = \begin{bmatrix} Q_1 & \mathbf{0}_{6\times6} \\ \mathbf{0}_{6\times6} & Q_2 \end{bmatrix}$$

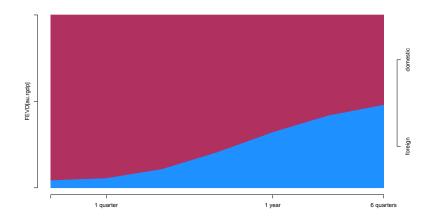
 Q_1 , $Q_2 - 6 \times 6$ rotation matrices (drawn from Haar distribution)

The model has been estimated by premultiplying every draw of matrices B_0 and B_+ by the corresponding draw of matrix Q

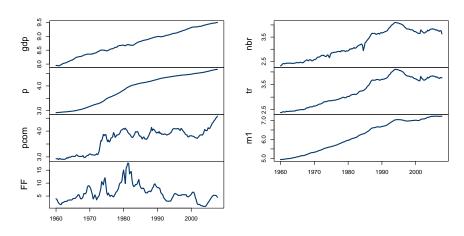
Forecast error variance decomposition of $au.cpi_{t+h|t}$



Forecast error variance decomposition of $au.rgdp_{t+h|t}$



Modeling effects of monetary policy in the US



Quarterly data from 1961 Q2 to 2007 Q4.

 $rgdp_t$ - real GDP, p_t - GDP price deflator, $pcom_t$ - commodity price index, FF_t - federal funds rate, nbr_t - non-borrowed reserves, tr_t - total reserves, m_t - monetary aggregate M1

$$B_0 y_t = b_0 + B_1 y_{t-1} + \dots + B_p y_{t-p} + u_t$$
$$u_t | s_t \sim \mathcal{N} (\mathbf{0}_N, I_N)$$

SVAR for a small-open economy

- ▶ The model was estimated using Waggoner & Zha (2003) algorithm
- ▶ \underline{B} and $\underline{\Omega}$ prior matrices are set the same way as the RF Minnesota prior with $\kappa_1 = 0.1$, $\kappa_2 = 10$, and $\kappa_3 = 1$
- ▶ $\underline{S} = \kappa_4 I_N$ and $\kappa_4 = 10$, $\underline{\nu} = N$ which make the generalized-normal prior for B_0 set to $\mathcal{N}(\mathbf{0}_N, 10I_N)$
- ▶ The results are sensitive to the specification of κ_1 and κ_4
- ► The Gibbs sampler for B_0 used $S_1 = 100$ draws in the burin-in and $S_2 = 5000$ in the final sampler
- ► The potential extensions include: estimation of prior hyper-parameters and heteroskedasticity of the structural shocks

FF policy shock by Bernanke & Blinder (1992, AER), Sims (1992, EER)

a_{11}	0	0	0	0	0	0]	$[rgdp_t]$
a ₂₁	a ₂₂	0	0	0	0	0	p_t
a ₃₁	a ₃₂	<i>a</i> 33	0	0	0	0	pcom _t
a ₄₁	<i>a</i> ₄₂	<i>a</i> 43	<i>a</i> 44	0	0	0	FFt
a ₅₁	a ₅₂	a ₅₃	a ₅₄	a ₅₅	0	0	nbr _t
a ₆₁	a ₆₂	a ₆₃	<i>a</i> ₆₄	a ₆₅	<i>a</i> ₆₆	0	tr _t
a ₇₁	a ₇₂	a ₇₃	a ₇₄	a ₇₅	a ₇₆	a ₇₇]	$\begin{bmatrix} m_t \end{bmatrix}$

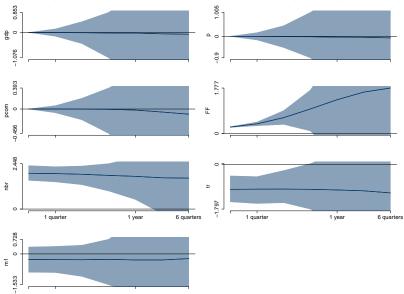
NBR policy shock by Christiano & Eichenbaum (1992)

a_{11}	0	0	0	0	0	0]	$\lceil rgdp_t \rceil$
a ₂₁	a ₂₂	0	0	0	0	0	p_t
a ₃₁	a ₃₂	<i>a</i> 33	0	0	0	0	pcomt
a ₄₁	<i>a</i> ₄₂	<i>a</i> 43	<i>a</i> 44	<i>a</i> 45	0	0	FF _t
a ₅₁	a ₅₂	a ₅₃	0	a ₅₅	0	0	nbr _t
a ₆₁	a ₆₂	a ₆₃	a ₆₄ a ₇₄	a ₆₅	a ₆₆	0	tr _t
a_{71}	a ₇₂	a ₇₃	a ₇₄	a ₇₅	a ₇₆	a ₇₇]	$[m_t]$

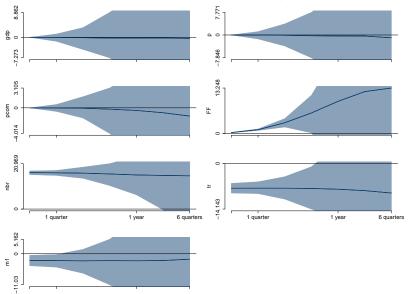
NBR/TR policy shock by Strongin (1995, JME)

a_{11}	0	0	0	0	0	0]	$[rgdp_t]$
a ₂₁	a ₂₂	0	0	0	0	0	pt
a ₃₁	a ₃₂	a ₃₃	0	0	0	0	pcom _t
a ₄₁	<i>a</i> ₄₂	<i>a</i> 43	<i>a</i> 44	<i>a</i> 45	<i>a</i> 46	0	FF _t
a ₅₁	a ₅₂	<i>a</i> 53	0	a ₅₅	a ₅₆	0	nbr _t
a ₆₁	32 342 352 362	a ₆₃	0	0	a ₆₆	0	tr _t
a ₇₁	a ₇₂	a ₇₃	<i>a</i> 74	a ₇₅	a ₇₆	a ₇₇]	$\begin{bmatrix} m_t \end{bmatrix}$

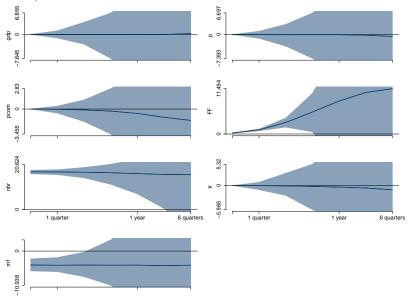
FF policy shock by Bernanke & Blinder (1992, AER), Sims (1992, EER)



NBR policy shock by Christiano & Eichenbaum (1992)



NBR/TR policy shock by Strongin (1995, JME)



Modeling effects of monetary policy in the US Extended analysis

Not Examinable

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In reference to:
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Christiano, Eichenbaum, & Evans (1999, HM) Ramey (2016, HM)

as well as:

Normandin, Phaneuf (2004, JME) Lanne, Lütkepohl (2008, JEDC) Woźniak & Droumaguet (2019)

Heteroskedastic Structural Vector Autoregressions

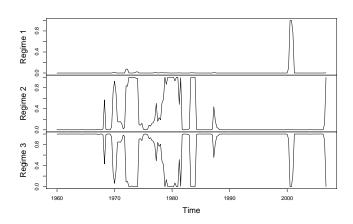
$$B_0 y_t = b_0 + B_1 y_{t-1} + \dots + B_p y_{t-p} + u_t$$
$$u_t | s_t \sim \mathcal{N} \left(\mathbf{0}_N, \operatorname{diag} \left(\lambda_{s_t} \right) \right)$$

$$B_{0.n} = b_n V_n$$
 — restrictions for rows $\sum_{s_t=1}^M \lambda_{n.s_t} = 1$ — standardization

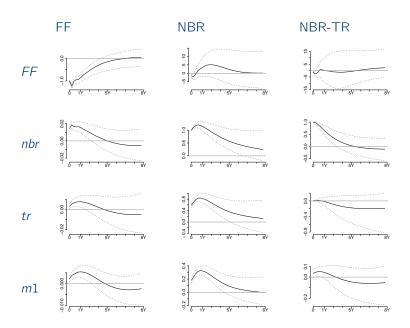
 $s_t = m \in \{1, ..., M\}$ – Markov process **P** – transition matrix $(\kappa_1, \kappa_2, \kappa_4)$ – estimated hyper-parameters

Volatility of Structural Shocks

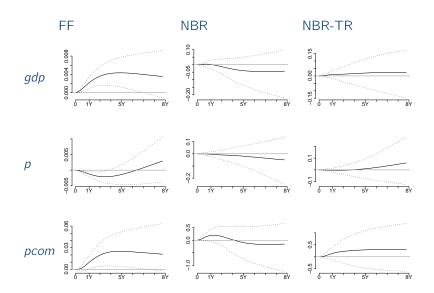
Marginal posterior state probabilities: $Pr[s_t|\mathbf{y}]$



Monetary Policy Models for U.S. Data



Monetary Policy Models for U.S. Data



Identification Through Heteroskedasticity

Heteroskedasticity.

Let there be 2 covariance matrices associated with data:

$$\Sigma_1$$
, and Σ_2

Contemporaneous effects matrix.

All N^2 elements of matrix B_0 are identified:

$$\begin{split} \Sigma_1 &= B_0^{-1} \text{diag}(\lambda_1) B_0^{-1'} \\ \Sigma_2 &= B_0^{-1} \text{diag}(\lambda_2) B_0^{-1'} \end{split}$$

Identification Through Heteroskedasticity

Heteroskedasticity.

Let there be M covariance matrices associated with data:

$$\Sigma_m = B_0^{-1} \operatorname{diag}(\lambda_m) B_0^{-1'}$$

Contemporaneous effects matrix.

All N^2 elements of matrix B_0 are identified.

Just-identifying restrictions in the homoskedastic case **over identify the system in the heteroskedastic** one.

These restrictions can be tested!

Monetary Policy Models for U.S. Data

Markov-switchi	ing heteroskedasticity
	rig ricicioshedusticity
2 2577.3 2655.2	2653.4 2660.1
3 2660.6 2710.0	2708.2 2731.4

Reported values: In $p(\mathbf{y}|\mathcal{M}_i)$

References

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Ramey (2016) Macroeconomic Shocks and Their Propagation, Handbook of Macroeconomics

Sims (1992) Interpreting the macroeonomic time series facts, The effects of monetary policy, European Economic Review

Strongin (1995) The identification of monetary policy disturbances explaining the liquidity puzzle, Journal of Monetary Economics

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