Probabilistic Inference by Hashing and Optimization

Focusing on Approximate Model Counting

Stefano Ermon slides by Dor Cohen

June 13, 2017

Introduction

- Many problems in ML and Statistics involve computation of high-dimensional integrals (e.g. evaluate posterior probabilities)
- Computing expectations w.r.t high dimensional probability distributions known to be intractable in worst-case (Roth, 2016)
- ▶ Difficulty arises as number of possible states scales exponentially *curse of dimensionality* (Belmman, 1996)

Standard approaches

Focusing on *discrete* probability distributions and *approximately* computing expectations, there are two main standard approaches:

- Monte Carlo sampling techniques (1950s): Approximate complex distributions with small number of representative samples
- ▶ Variational methods: Approximate complex models using families of tractable distributions (e.g., estimate $P(Z|X) \approx Q(Z)$ and compare distributions using KL-divergence)

These don't provide tight guarantees on accuracy

MCMC - Monte Carlo Markov Chain

Goal: Sample from a *discrete* distributions vector $\pi = (\pi_1, ... \pi_n)$ **Idea:**

- ▶ Build a Markov chain with *n* states such that it will be *ergodic*
- ▶ Choose weights for graph edges (probabilities) such that the chain will be *invertible* w.r.t to the desirable distribution π
- ightharpoonup ergodic and invertible chain imply that the "last" state of long random walk on the chain is distributed like π

MCMC - Monte Carlo Markov Chain

Key difficulty

is to to draw proper samples

- Needs to set up a Markov Chain over entire state space
- Has to reach equilibrium distribution - requires exponential time (Madras, 2002)
- In practice will only give approximate answer. Chain may "trap" in less relevant areas of the state space

Convergence of Metropolis — Hastings

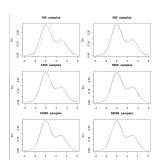


Figure: MCMC attempts to approximate blue distribution with orange one [Wiki]

A new approach

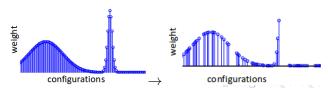
Approximate inference/counting algorithms based on randomized hashing and optimizations:

- provably accurate results assuming access to an optimization oracle
- Can compute parity functions, marginal probabilities providing an approximately correct answer
- ▶ Specifically, within a factor of $1+\epsilon$ of the true value for any desired ϵ , with high probability (at least $1-\lambda$ for any desired $\lambda>0$)

Hashing and optimization

Statistics are still computed using small subset of samples. However these samples aren't drawn randomly from the distribution using an MC, but:

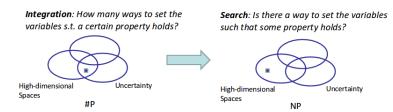
- Samples are obtained by random projection of original high-dimensional states to lower-dimensional ones using Universal hash functions
- Then look for "Most likely" configuration in the projected subspace (current optimization tools can handle millions of variables)



Complexity perspective

Inference or counting problems we consider are in #P, believed to be significantly harder than NP (Valiant, 1979)

▶ If we allow small failure probability, and small error we can reduce the #P problem to smaller number of instances of a NP-equivalent (combinatorial) optimization problem



Problem statement - setup

- Let Σ be a large but finite set (e.g. the set of all possible assignments in a model) and let $w: \Sigma \to \mathbb{R}^+$ be a non-negative function that assigns a weight for each configuration (an element of Σ)
- ▶ We are given 2ⁿ configurations, and non-negative weights w (Bayes net, Factor graph, weighted CNF)
- ► **Goal:** (approximately) compute the total weight of the set, defined as:

$$W = \sum_{\sigma \in \sum} w(\sigma)$$

Example: n = 100 variables, sum over 2^{100} configurations

Problem statement - setup

Assumption: We assume that we have access to an *optimization oracle* that can solve the following optimization problem:

$$\max_{\sigma \in \sum} w(\sigma) \mathbb{1}_{\{o\}}(\sigma)$$

Where $\mathbb{1}_{\{o\}}: \sum \to \{0,1\}$ is an indicator function for compactly represented subset $o \subseteq \sum$, i.e., $\mathbb{1}_{\{o\}} = 1$ if $\sigma \in o$ and 0 otherwise.

 o may be compactly represented using a smaller subset of constraints

SAT

- ▶ A boolean formula φ is said to be in *CNF* if its a logical conjunction of a set of clauses $C_1, ... C_n$, where each clause C is a logical disjunction of a set of literals. e.g., $(x_1 \lor \neg x_2)$
- ▶ SAT: deciding if there exists an assignment that satisfies φ

Example:
$$\varphi = (x_1 \vee \neg x_2 \vee x_3) \wedge (x_4 \vee \neg x_1) \wedge (x_2 \vee \neg x_3)$$

Satisfying assignment: $\{x_1 = 1, x_2 = 1, x_3 = 1, x_4 = 1\}$
 $\implies \varphi$ is SATISFIABLE

Model counting

- Let V be the set of boolean variables of φ , |V|=n, and let $\sigma=\{0,1\}^n$ be the set of all possible assignments to these variables
- \blacktriangleright An assignment $\sigma\in \sum$ is a mapping that assigns a value in $\{0,1\}$ to each variable in V
- ▶ Define the weight $w(\sigma)$ to be 1 if φ is satisfied by σ and 0 otherwise

In this context:

$$W = \sum_{\sigma \in \sum} w(\sigma) = \#(\varphi)$$

Approximate model counting

▶ Problem of approximately counting the number of solutions, assuming access to an NP oracle (e.g. SAT solver) was first considered by Stockmeyer (1985) - Important result he established: #P can be approximated in BPP^{NP} 1

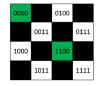
Intuition behind his algorithm:

- ▶ Let $S \subseteq \sum$ be the set of solutions to φ
- \blacktriangleright Reliably estimate |S| by **repeating** the following process
 - 1. Randomly partition \sum into 2^m cells
 - 2. Select one of the cells
 - Compute if S has at least one element in this cell (invoke a SAT solver)

 $^{^1}$ algorithms that have bounded-error probabilistic polynomial time and access to an NP oracle

Approximate model counting - Example

0000	0001	0100	0101
0010	0011	0110	0111
1000	1001	1100	1101
1010	1011	1110	1111





- ► **Example:** $\Sigma = \{0,1\}^4$, vars: x_1, x_2, x_3, x_4 . $Sols = \{(0,0,0,0), (0,0,1,0), (1,1,0,0), (1,0,1,0)\}$.
- ▶ **Middle:** matrix partitioned into 2 cells, based on the parity of $x_3 \oplus x_4$ two solutions are left, corresponding to the assignments satisfying $x_3 \oplus x_4 = 0$.
- ▶ **Right:** after partitioning again based on the parity of $x_2 \oplus x_3$, only solution (0,0,0,0) is left.

Approximate model counting

To summarize:

- ▶ Estimate |S|, by defining progressively smaller cells, until no element of S can be found inside a randomly chosen cell
- ▶ The larger |S| is, the smaller the cells have to be
- ► Correctness of the algorithm relies crucially on how the space is randomly partitioned into cells
- ➤ To achieve strong worst-case (probabilistic) guarantees, algorithm relies on universal hash functions (Vadhan, 2011; Goldreich, 2011)

Universal hash functions

Definition: A family of functions $\mathcal{H} = \{h : \{0,1\}^n \mapsto \{0,1\}^m\}$ is strongly universal (pairwise independent) if the following two conditions hold when h is chosen uniformly at random from \mathcal{H} :

- 1. $\forall x \in \{0,1\}^n$, the random variable H(x) is uniformly distributed in $\{0,1\}^m$
- 2. $\forall x_1, x_2 \in \{0,1\}^n | x_1 \neq x_2$, the random variables $H(x_1)$ and $H(x_2)$ are independent

Universal hash functions

- ▶ Considering the set of all possible functions from $\{0,1\}^n$ to $\{0,1\}^m$ we establish statistically optimal functions. Easy to verify this is a family of *fully* independent functions
- ► However, functions like these require $m2^n$ bits for specifying \rightarrow not useful for large n
- pairwise independent hash functions can be specified compactly, with number of bits linear in n

Pairwise independent hash functions

▶ Generally, these functions are based on modular arithmetic constraints of the form $Ax = b \mod 2$. Implies that Ax is congruent to $b \mod 2$

Proposition: Let $A \in \{0,1\}^{m \times n}$, $b \in \{0,1\}^m$. The family $\mathcal{H} = \{h_{A,b}(x) : \{0,1\}^n \mapsto \{0,1\}^m\}$ where $h_{A,b}(x) = Ax + b \mod 2$ is a family of pairwise independent hash functions.

Algorithm: ApproxModelCount

Algorithm 9.1 ApproxModelCount $(F, \mathcal{O}_S, \Delta)$

```
Let S denote the set of solutions to the input formula F
    T \leftarrow \left\lceil \frac{\log(1/\Delta)}{\alpha} \log n \right\rceil
    i = 0
     while i \le n do
          for t = 1, \dots, T do
5:
              Sample hash function h_{A,b}^i: \Sigma \to \{0,1\}^i, i.e.
6:
                   sample uniformly A \in \{0, 1\}^{i \times n}, b \in \{0, 1\}^i
              Let S(h_A^i) = |\{x \in S \mid Ax \equiv b \pmod{2}\}|
8:
              w_i^t \leftarrow \mathbb{I}[S(h_{A.b}^i) \ge 1], using \mathbb{O}_S to check if \{x \in S \mid Ax \equiv b \pmod{2}\} is empty
9:
           end for
10:
           if Median(w_i^1, \dots, w_i^T) < 1 then
11:
12:
                break
13:
           end if
           i = i + 1
14:
      end while
15:
      Return |2^{i-1}|
16:
```

Algorithm: ApproxModelCount

Based on pairwise independence, possible to show that ApproxModelCount provides with high probability an accurate estimate of the true value, summarized by the following Theorem:

Theorem: For any $\Delta>0$, positive constant $a\leq 0.0042$, ApproxModelCount makes $\Theta(nlnnln1/\delta)$ queries to the NP oracle O_S and, with probability of at least $(1-\Delta)$, outputs a 16-approximation of |S|, the number of solutions of a formula φ

By the uniformity property of *Universal hash functions*: Given any solution $x \in S$, we can see that in iteration i it holds that: $P[Ax = b(mod 2)] = (\frac{1}{2})^i$

```
Algorithm 9.1 ApproxModelCount (F, O_S, \Delta)
    Let S denote the set of solutions to the input formula F
     T \leftarrow \left[\frac{\log(1/\Delta)}{\alpha} \log n\right]
     while i \le n do
         for t = 1, \dots, T do
              Sample hash function h_{A,b}^i: \Sigma \to \{0,1\}^i, i.e.
                  sample uniformly A \in \{0, 1\}^{i \times n}, b \in \{0, 1\}^{i}
              Let S(h_A^i) = |\{x \in S \mid Ax \equiv b \pmod{2}\}|
9:
              w_i^t \leftarrow \mathbb{I}[S(h_{A,h}^i) > 1], using O_S to check if \{x \in S \mid Ax \equiv b\}
10
           if Median(w_i^1, \dots, w_i^T) < 1 then
               break
           end if
           i = i + 1
      end while
```

Return $|2^{i-1}|$

By linearity of expectation: $E[S(h_{A,b}^i)] = \frac{|S|}{2^i}$. It implies that in expectation the while loop should break for $i \approx log|S|$ i.e., when the corresponding expected number of "surviving" solutions is less than 1. When this happens, the algorithm provides an accurate estimate for |S|

```
Algorithm 9.1 ApproxModelCount (F, O_S, \Delta)

    Let S denote the set of solutions to the input formula F

     T \leftarrow \left[\frac{\log(1/\Delta)}{\alpha} \log n\right]
     while i \le n do
         for t = 1, \dots, T do
              Sample hash function h_{A,b}^{i}: \Sigma \rightarrow \{0,1\}^{i}, i.e.
                  sample uniformly A \in \{0, 1\}^{i \times n}, b \in \{0, 1\}^i
              Let S(h_{A,b}^i) = |\{x \in S \mid Ax \equiv b \pmod{2}\}|
              w_i^t \leftarrow \mathbb{I}[S(h_{A,b}^i) > 1], using O_S to check if \{x \in S \mid Ax \equiv b\}
           if Median(w_i^1, \dots, w_i^T) < 1 then
11:
12.
               break
13:
           end if
           i = i + 1
      end while
      Return |2^{i-1}|
```

Because the hash functions family is pairwise independent, $S(h_{A,b}^i)$ is the sum of pairwise independent random variables (one for each element, corresponding to whether that solution satisfies the random constraints)

```
Algorithm 9.1 ApproxModelCount (F, O_S, \Delta)
    Let S denote the set of solutions to the input formula F
     T \leftarrow \left[\frac{\log(1/\Delta)}{\alpha} \log n\right]
     while i \le n do
         for t = 1, \dots, T do
              Sample hash function h_{A,b}^i: \Sigma \to \{0,1\}^i, i.e.
                 sample uniformly A \in \{0, 1\}^{i \times n}, b \in \{0, 1\}^{i}
              Let S(h_{A,b}^i) = |\{x \in S \mid Ax \equiv b \pmod{2}\}|
              w_i^t \leftarrow \mathbb{I}[S(h_{A,h}^i) > 1], using O_S to check if \{x \in S \mid Ax \equiv b\}
10:
          if Median(w_i^1, \dots, w_i^T) < 1 then
12:
               break
           end if
          i = i + 1
      end while
      Return |2^{i-1}|
```

Therefore, the variance of $S(h_{A,b}^i)$ is the sum of **individual** variances. And can be shown to be "small" compared to the mean.

By standard concentration inequalities, $S(h_{A,b}^i)$ takes values close to it's mean reasonably often.

Computing median over T runs, guarantees an accurate estimate with high probability.

Algorithm 9.1 ApproxModelCount $(F, \mathcal{O}_S, \Delta)$

```
    Let S denote the set of solutions to the input formula F

     T \leftarrow \left[\frac{\log(1/\Delta)}{\alpha} \log n\right]
     while i \le n do
         for t = 1, \dots, T do
              Sample hash function h_{A,b}^i : \Sigma \to \{0,1\}^i, i.e.
                  sample uniformly A \in \{0, 1\}^{i \times n}, b \in \{0, 1\}^i
              Let S(h_{A,b}^i) = |\{x \in S \mid Ax \equiv b \pmod{2}\}|
              w_i^t \leftarrow \mathbb{I}[S(h_{A,b}^i) > 1], using O_S to check if \{x \in S \mid Ax \equiv b\}
10:
           if Median(w_i^1, \dots, w_i^T) < 1 then
12:
                break
           end if
           i = i + 1
      end while
16: Return | 2<sup>i-1</sup> |
```

Improving the approximation factor

Given the k-approximation algorithm and any $\epsilon>0$, it's possible to design a $(1+\epsilon)$ -approximation algorithm, with the following construction:

Let $I = log_{1+\epsilon} k$, define new set of configurations $\sum_{l=0}^{l=0} x_{l} \sum_{l=0}^{l=0} ... x_{l} \sum_{l=0}^{l=0} ... x_{l} \sum_{l=0}^{l=0} ... x_{l} \sum_{l=0}^{l=0} ... x_{l} \sum_{l=0}^{l=0} w(\sigma_{1}) w(\sigma_{2}) ... w(\sigma_{l})$

Idea is to estimate the size of $S(h_{A,b}^i)$ with multiple calls to an NP-oracle , e.g. check if $S(h_{A,b}^i)$ contains at least k elements \to Reduces the variance, can be used to improve accuracy, but requires more calls to the NP-oracle

Compactly representing constraints

- ▶ In model counting, the calls to an *NP*-oracle are implemented by using a SAT-solver. This is accomplished by adding to the formula *i* parity constraints, and checking for satisfiability.
- Naive encoding of a parity constraint over p variables requires 2^{p-1} clauses of length p ruling out 2^p possible assignments (ones with wrong parity)
- Constraints can be compactly represented (introducing O(p) extra variables) using standard Tseitin transformation (Tseitin, 1983).

Example: $x_1 \oplus x_2 \oplus x_3 \oplus x_4 = 0$ can be written as $\{x_1 \oplus x_2 = z_1, x_2 \oplus x_3 = z_2, z_1 \oplus z_2 = 0\}$

Practical implementations

- ► First practical implementation by Gomes et al. (2006) who used a SAT solver as an *NP*-oracle.
- Their algorithm leverages decades of research and engineering in SAT solving techniques for approximate model counting, resulted in huge improvements
- ► Recently, Chakraborty et al. (2013); Ivrii et al. (2015) provided several practical improvements
- ▶ Specifically, the former introduced the use of *pivots*, where an NP-oracle is used to check the existence of at least k>1 solutions (k=1 corresponds to earlier discussed algorithm), in order to improve the accuracy of the estimated count

Probabilistic Models and Approx Inference: WISH Algorithm

- ▶ Generally, when the weight function is "close" for being constant on a subset of states, and zero elsewhere, then the hashing-based algorithm of Chakraborty et al. (2013) can be used - as in the model counting problem.
- Typical models in ML are unlikely to satisfy this restriction (e.g., weight function that is log-linear can have large variability)
- An alternative algorithm (WISH) based on universal hashing and combinatorial optimization which can handle general weight function, was introduced by Ermon et al. (2013)

WISH Algorithm

Algorithm 9.2 WISH $(w : \Sigma \to \mathbb{R}^+, n = \log_2 |\Sigma|, \Delta, \alpha)$

```
1: T \leftarrow \left[\frac{\ln(n/\delta)}{\alpha}\right]
      for i = 0, \dots, n do
           for t = 1, \dots, T do
3:
                Sample hash function h_{A,b}^i: \Sigma \to \{0,1\}^i, i.e.
4:
                     sample uniformly A \in \{0,1\}^{i \times n}, b \in \{0,1\}^i
5:
6:
                w_i^t \leftarrow \max_{\sigma} w(\sigma) subject to A\sigma \equiv b \pmod{2}
7:
           end for
           M_i \leftarrow \text{Median}(w_i^1, \cdots, w_i^T)
      end for
       Return M_0 + \sum_{i=0}^{n-1} M_{i+1} 2^i
10:
```

ApproxModelCount Proof sketch Improvements WISH Algorithm

Thank you for listening!

Any questions?