

Assignment 1

$$1) X \sim P_x(m) = \frac{e^{-\lambda} \lambda^m}{m!}$$

$$Y \sim P_y(y) = \frac{e^{-\mu} \mu^y}{y!}$$

Distribution of X given that $X+Y=n$ can be shown by

$$P(X=m | X+Y=n) \\ \text{or } P(X=m | Y=n-X) = \frac{P(X,Y)}{P(X+Y=n)} \quad \text{--- (3)}$$

$$\begin{aligned} P(X+Y=n) &= \sum_{m=0}^n P_x(m) P_y(n-m) \\ &= \sum_{m=0}^n \left(\frac{e^{-\lambda} \lambda^m}{m!} \right) \left(\frac{e^{-\mu} \mu^{n-m}}{(n-m)!} \right) \\ &= e^{-(\lambda+\mu)} \sum_{m=0}^n \left(\frac{1}{n!} \right) \left(\frac{n!}{m! (n-m)!} \right) \lambda^m \mu^{n-m} \\ &= \frac{e^{-(\lambda+\mu)}}{n!} \sum_{m=0}^n n C_m \lambda^m \mu^{n-m} \\ &= \frac{e^{-(\lambda+\mu)}}{n!} (\lambda+\mu)^n \quad (\text{Binomial expansion}) \quad \text{--- (4)} \end{aligned}$$

$P(X,Y) = P_x(X=m) P_y(Y=y)$ [Both are independent random variables]

$$\begin{aligned} &= P_x(X=m) P_y(Y=n-m) \\ &= \left(\frac{e^{-\lambda} \lambda^m}{m!} \right) \left(\frac{e^{-\mu} \mu^{n-m}}{(n-m)!} \right) \\ &= \frac{e^{-(\lambda+\mu)}}{n!} \left(\frac{n!}{m! (n-m)!} \right) \lambda^m \mu^{n-m} = \frac{e^{-(\lambda+\mu)}}{n!} \binom{n}{m} \lambda^m \mu^{n-m} \quad \text{--- (5)} \end{aligned}$$

Using ①, ②, ③

$$\begin{aligned}
 p(X=m | Y=n-X) &= \frac{\frac{e^{-(\lambda+\mu)}}{m!} \binom{n}{m} \lambda^m \mu^{n-m}}{\frac{e^{-(\lambda+\mu)}}{n!} (\lambda+\mu)^n} \\
 &= \binom{n}{m} \frac{\lambda^m \mu^{n-m}}{(\lambda+\mu)^n} = \binom{n}{m} \frac{\lambda^m \mu^{n-m}}{(\lambda+\mu)^m (\lambda+\mu)^{n-m}} \\
 &= \binom{n}{m} \left(\frac{\lambda}{\lambda+\mu}\right)^m \left(\frac{\mu}{\lambda+\mu}\right)^{n-m} \\
 &= \binom{n}{m} \pi^m (1-\pi)^{n-m}, \quad \pi = \frac{\lambda}{\lambda+\mu}
 \end{aligned}$$

Q3) Given: $P(X > 0) = 1$

↳ This implies

$$\int_{-\infty}^0 f_X(x) dx = 0 - 0 = 0 \quad (m=0)$$

$$P(X \leq 0) = 0$$

$$E(X) = \int_{-\infty}^{\infty} x f_X(x) dx$$

$$= \int_0^{\infty} x f_X(x) dx$$

$$= \int_0^{\infty} x \left(\int_0^x f_X(t) dt \right) dx = \int_0^{\infty} \left(\int_0^x f_X(t) dx \right) dt$$

$$= \int_0^{\infty} F_X(x) dx$$

$$= \int_0^{\infty} (1 - P(X > x)) dx = \int_0^{\infty} (1 - P(X > x)) dx$$

$$= \int_0^{\infty} 1 dx - \int_0^{\infty} P(X > x) dx$$

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$$= \int_0^{\infty} P(X > x) dx - \int_0^{\infty} (1 - F_X(x)) dx$$

since $\lim_{x \rightarrow \infty} F_X(x) = 1$

$$\therefore \int_0^{\infty} (1 - F_X(x)) dx = 0$$

$$= \int_0^{\infty} P(X > x) dx, \text{ hence proved}$$

$$24) \quad p(x) = N(x|M, E) \quad \text{and} \quad q(x) = N(x|m, L)$$

$$q(x) = N(x|m, L)$$

$$D_{KL}(p(x)||q(x)) = \int_{-\infty}^{\infty} p(x) \ln\left(\frac{p(x)}{q(x)}\right) dx$$

$$= \int_{-\infty}^{\infty} p(x) \ln\left(\sqrt{\frac{L}{E}} e^{\left\{-\frac{(x-M)^2}{2E} + \frac{(x-m)^2}{2L}\right\}}\right) dx$$

$$= \int_{-\infty}^{\infty} p(x) \ln\left(\sqrt{\frac{L}{E}}\right) - \frac{(x-M)^2}{2E} p(x) + \frac{(x-m)^2}{2L} p(x) dx$$

$$= \ln\left(\sqrt{\frac{L}{E}}\right) \int_{-\infty}^{\infty} p(x) dx - \int_{-\infty}^{\infty} \frac{(x-M)^2}{2E} p(x) dx + \int_{-\infty}^{\infty} \frac{(x-m)^2}{2L} p(x) dx$$

$$= \frac{1}{2} \ln\left(\frac{L}{E}\right) - \frac{E_p((x-M)^2)}{2E} + \frac{E_p((x-m)^2)}{2L}$$

$$= \frac{1}{2} \ln\left(\frac{L}{E}\right) - \frac{1}{2} + \frac{E_p(m^2 - 2mx + x^2)}{2L}$$

$$(E = E_p((x-M)^2))$$

$$E_p(x^2) = E + M^2$$

$$E_p(x) = M$$

$$= \frac{1}{2} \ln\left(\frac{L}{E}\right) - \frac{1}{2} + \frac{1}{2L} (E + M^2 - 2mM + m^2)$$

$$= \frac{1}{2} \ln\left(\frac{L}{E}\right) - \frac{1}{2} + \frac{1}{2L} (E + (M-m)^2)$$

$$5) f(x, y) = \frac{1}{2\pi(1-\rho^2)^{1/2}} \exp \left(-\frac{1}{2(1-\rho^2)} (x^2 - 2\rho xy + y^2) \right)$$

$$\text{Corr}(X, Y) = \frac{\text{Cov}(X, Y)}{\sigma_X \sigma_Y}$$

$$\sigma_X, \sigma_Y = 1$$

$$\mu_X, \mu_Y = 0$$

$$\therefore \text{Corr}(X, Y) = \text{Cov}(X, Y) = E(XY)$$

$$E(XY) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} xy \frac{1}{2\pi(1-\rho^2)^{1/2}} \exp \left(-\frac{1}{2(1-\rho^2)} (x^2 - 2\rho xy + y^2) \right) dx dy$$

$$\begin{aligned} x^2 - 2\rho xy + y^2 &= x^2 - \rho^2 x^2 + (y^2 - 2\rho xy + \rho^2 x^2) \\ &= x^2(1-\rho^2) + (y - \rho x)^2 \end{aligned}$$

$$\begin{aligned} &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \frac{xy}{2\pi(1-\rho^2)^{1/2}} \exp \left\{ -\frac{x^2}{2} - \frac{(y - \rho x)^2}{2(1-\rho^2)} \right\} dx dy \\ &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \left(\frac{x e^{-x^2/2}}{\sqrt{2\pi}} \right) \underbrace{\left(\frac{y}{\sqrt{2\pi(1-\rho^2)}} \exp \left(-\frac{(y - \rho x)^2}{2(1-\rho^2)} \right) \right)}_{\text{normal distribution}} dy dx \end{aligned}$$

This integral will be ρx since it is of form of normal distribution.

$$= \rho \int_{-\infty}^{\infty} \frac{x^2 e^{-x^2/2}}{\sqrt{2\pi}} dx$$

$$= \rho E(X^2) \quad ; \quad \begin{aligned} E(X^2) &= \text{Var}(X) + E(X)^2 \\ &= 1 + 0 \\ &= 1 \end{aligned}$$

$$= \rho \quad \therefore \underline{\underline{\text{Corr}(X, Y) = \rho}}$$

$$\text{Corr}(X^2, Y^2) = \frac{\text{Cov}(X^2, Y^2)}{\sigma_{X^2} \sigma_{Y^2}}$$

$$= \frac{(E(X^2 Y^2) - 1)}{\sigma_{X^2} \sigma_{Y^2} - \left(\frac{(y - \mu)^2}{2(1-p^2)} \right)}$$

$$E(X^2 Y^2) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \left(\frac{y^2 e^{-y^2/2}}{\sqrt{2\pi}} \right) \left(\frac{y^2 e^{-\frac{(y-\mu)^2}{2(1-p^2)}}}{\sqrt{2\pi(1-p^2)}} \right) dy d\eta$$

$$E(Y^2) = \text{Var}(Y) + E(Y)^2$$

$$= (1-p^2) + p^2 \eta^2$$

$$= \int_{-\infty}^{\infty} \left(\frac{y^2 e^{-y^2/2}}{\sqrt{2\pi}} \right) (1-p^2 + p^2 \eta^2) d\eta$$

$$= E(X^2) - p^2 E(X^2) + p^2 E(X^4)$$

$$= 1 - p^2 + p^2 E(X^4)$$

$$\sigma_{X^2} = \sqrt{E(X^4) - E(X^2)^2}$$

$$\sigma_{Y^2} = \sqrt{E(Y^4) - E(Y^2)^2}$$

$$\text{also } E(X^2) = E(Y^2) = 1$$

$$E(X^4) = E(Y^4)$$

$$\Rightarrow \sigma_{X^2} \sigma_{Y^2} = E(X^4) - 1$$

$$\therefore \text{Corr}(X^2, Y^2) = \frac{E(X^2 Y^2) - 1}{\sigma_{X^2} \sigma_{Y^2}}$$

$$= \frac{p^2 (E(X^4) - 1)}{E(X^4) - 1}$$

$$= p^2$$

$$\therefore \underline{\underline{\text{Corr}(X^2, Y^2) = p^2}}$$