#### AMBARTI models for agricultural experiments

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#### Agenda

- ▶ Additive Main Effect interaction (AMMI) models
- ► AMMI + BART = AMBARTI
- Simulation (AMMI and AMBARTI)
- Next steps
- Appendix

# Additive Main effects and Multiplicative Interactions (AMMI)

Linear-bilinear models are frequently used to analyse two-way data such as genotype-by-environment data.

An example of this class of models is the AMMI model:

$$y_{ij} = \mu + g_i + e_j + \sum_{q=1}^{Q} \lambda_q \gamma_{iq} \delta_{jq} + \epsilon_{ij}, \quad \epsilon_{ij} \sim \mathcal{N}\left(0, \tau^{-1}\right),$$

where  $g_i$  is the effect of genotype and  $e_i$  the effect of environment.

# Additive Main effects and Multiplicative Interactions (AMMI)

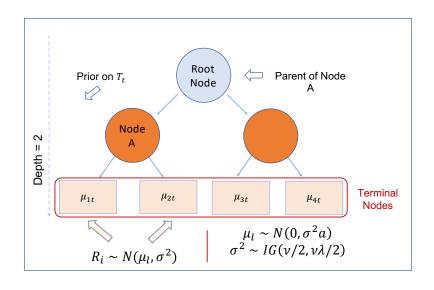
The following priors are assumed in its Bayesian version (Josse et al, JABES, 2014):

$$\begin{array}{rcl} \mu & \sim & \mathsf{N}\left(m,s_{\mu}^{2}\right), \\ g_{i} & \sim & \mathsf{N}\left(0,s_{g}^{2}\right), \\ e_{j} & \sim & \mathsf{N}\left(0,s_{e}^{2}\right), \\ (\lambda_{q})_{q=1,\ldots,Q} & \sim & \text{ordered sample of } Q \text{ independent } \mathsf{N}^{+}\left(0,s_{\lambda}^{2}\right), \\ \gamma_{1q} & \sim & \mathsf{N}^{+}(0,1) \quad \text{for } q=1,\ldots,Q, \\ \gamma_{iq} & \sim & \mathsf{N}(0,1) \quad \text{for } i>1 \text{ and } q=1,\ldots,Q, \\ \delta_{jq} & \sim & \mathsf{N}(0,1) \quad \text{for } j\geq 1 \text{ and } q=1,\ldots,Q, \\ \sigma_{E} & \sim & \mathsf{U}\left(0,S_{\mathrm{ME}}\right). \end{array}$$

Additive Main Effect Bayesian Additive

Regression Tree Interaction models (AMBARTI)

#### **BART**



#### AMMI + BART

BART is a flexible tree-based method that can be used for predicting when there are interactions and non-linear relationships.

$$y_{ij}|\mathbf{x}_{ij}, \mathcal{T}, \mathcal{M}, \Theta, \sigma^2 \sim N\left(g_i + e_j + \sum_{t=1}^T h(\mathbf{x}_{ij}, \mathcal{M}_t, \mathcal{T}_t), \sigma^2\right),$$

where  $y_{ij}$  is the yield for genotype i and environment j, and  $g_i$  and  $e_j$  are the genotype and environment effects, respectively.

$$\begin{split} &\mu_{t\ell}|\mathcal{T}_t \sim \mathsf{N}(\mu_{\mu} = 0, \sigma_{\mu}^2), \\ &g_i|\mathcal{T}_t \sim \mathsf{N}(\mu_g, \sigma_g^2), \\ &e_j|\mathcal{T}_t \sim \mathsf{N}(\mu_e, \sigma_e^2), \\ &\sigma_g^2 \sim \mathsf{IG}(a_g, b_g), \\ &\sigma_e^2 \sim \mathsf{IG}(a_e, b_e), \\ &\sigma^2 \sim \mathsf{IG}(a, b). \end{split}$$

Interesting fact...

#### Bayesian AMMI - Postprocessing (Josse et al, 2014)

Recall that

$$\mu_{ij} = \hat{y}_{ij} = \hat{\mu} + \hat{g}_i + \hat{e}_j + \sum_{q=1}^{Q} \hat{\lambda}_q \hat{\gamma}_{iq} \hat{\delta}_{jq}.$$

"This means that, concretely, S matrices of size  $I \times J$  are available as draws from the posterior distributions of the  $\mu_{ij}$ . Thus, it is possible to apply a postprocessing on each matrix (s=1,...,S) performing the classical procedure (in accordance with the chosen constraints): each matrix is centered by row and by column, and an SVD is applied on the resulting matrix. Consequently, for each s, new parameters  $(\mu, g, e, \gamma, \delta, \lambda_q)$  meeting the constraints are available. Consequently, draws in the posterior distribution of the parameters (taking the S new values) are available. Such a postprocessing makes it easier to interpret the results."

#### Simulation: AMMI scenarios (20 combinations)

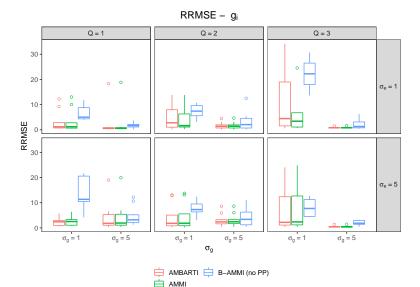
We consider the following setting to generate a set of simulated data sets:

$$y_{ij} = \mu + g_i + e_j + \sum_{q=1}^{Q} \lambda_q \gamma_{iq} \delta_{jq} + \epsilon_{ij}, \quad \epsilon_{ij} \sim \mathcal{N}\left(0, \tau^{-1}\right),$$

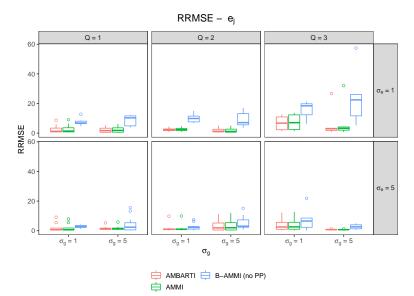
where

- $\mu = 100, \tau = 1.$
- I = J = 10 (without repetitions).
- $e_i \sim N(0, s_e)$ , with  $s_e = c(1, 5)$ .
- $g_i \sim N(0, s_g)$ , with  $s_g = c(1, 5)$ .
- ightharpoonup Q = c(1,2,3), with  $\lambda = c(8,12,[8,12],[10,12],[8,10,12])$ .
- We used RMSE (Root Mean Squared Error) and RRMSE (Relative Root Mean Squared Error).

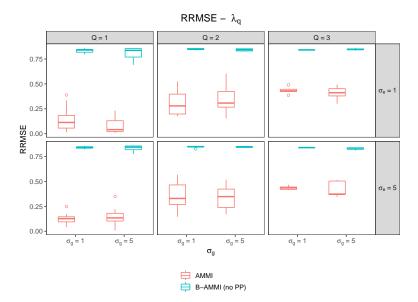
## Simulation AMMI $(y_{ij} = \mu + \mathbf{g}_i + \mathbf{e}_j + \sum_{q=1}^{Q} \lambda_q \gamma_{iq} \delta_{jq})$



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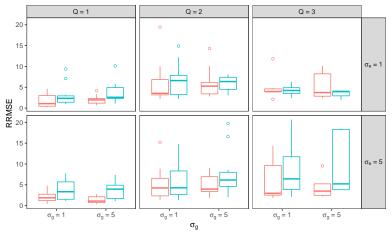


## Simulation AMMI $(y_{ij} = \mu + g_i + e_j + \sum_{q=1}^{Q} \frac{\lambda_q}{\lambda_q} \gamma_{iq} \delta_{jq})$



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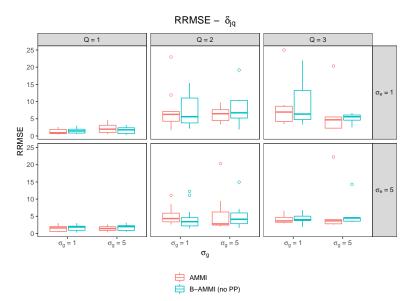




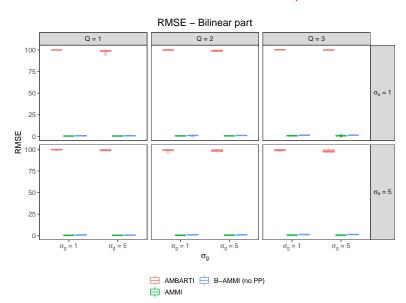




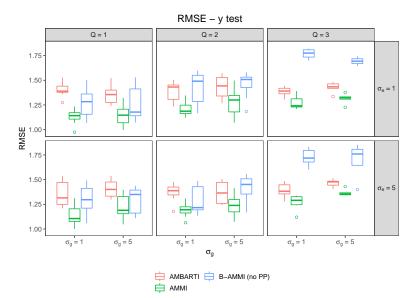
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#### Simulation: AMBARTI scenarios (4 combinations)

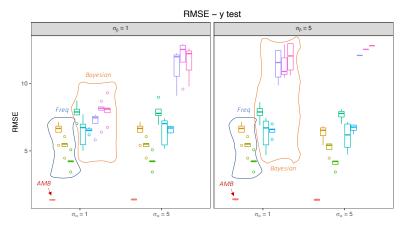
We consider the following setting to generate a set of simulated data sets:

$$y_{ij}|\mathbf{x}_{ij}, \mathcal{T}, \mathcal{M}, \Theta, \sigma^2 \sim N\left(g_i + e_j + \sum_{t=1}^T h(\mathbf{x}_{ij}, \mathcal{M}_t, \mathcal{T}_t), \sigma^2\right),$$

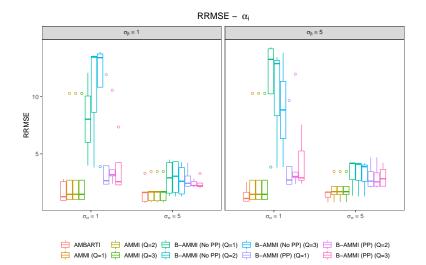
where  $y_{ij}$  is the yield for genotype i and environment j, and  $g_i$  and  $e_j$  are the genotype and environment effects, respectively.

- $\sigma^2 = 1$ , T = 200.
- ▶ I = J = 10 (without repetitions).
- $\qquad \qquad \mu_{t\ell}|\mathcal{T}_t \sim \mathsf{N}(\mu_{\mu}=0,\sigma_{\mu}^2=3),$
- $e_j \sim N(0, s_e)$ , with  $s_e = c(1, 5)$ .
- $g_i \sim N(0, s_g)$ , with  $s_g = c(1, 5)$ .

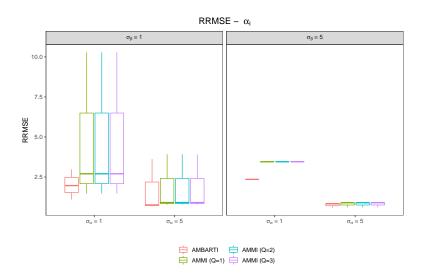
## Simulation AMBARTI $(y_{ij} = g_i + e_j + \sum_{t=1}^T h(\mathbf{x}_{ij}, \mathcal{M}_t, \mathcal{T}_t))$



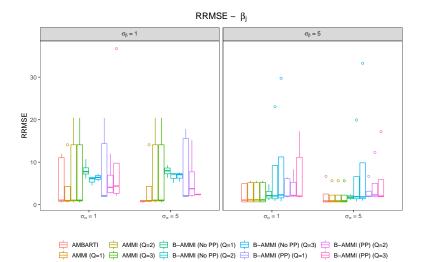
## Simulation AMBARTI $(y_{ij} = \mathbf{g_i} + e_j + \sum_{t=1}^{T} h(\mathbf{x}_{ij}, \mathcal{M}_t, \mathcal{T}_t))$



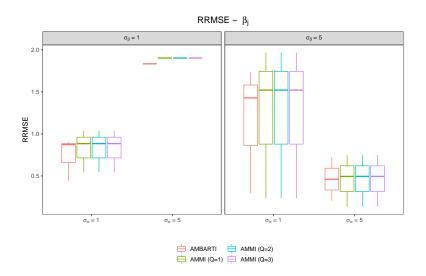
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#### Next steps

- 1. Adapt our implementation to simulate large (> 15) I and J.
- 2. Perform more simulations considering large I and J.
- 3. Analyse real data sets.

#### That's all, folks! Thank you!

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#### Simulation: full model (4 combinations)

We consider the following setting to generate a set of simulated data sets:

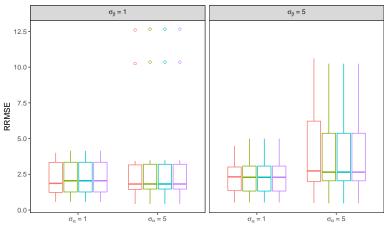
$$y_{ij} = \mu + g_i + e_j + g_i \times e_j + \epsilon_{ij}, \quad \epsilon_{ij} \sim \mathcal{N}\left(0, \tau^{-1}\right),$$

where

- $\mu = 100, \tau = 1.$
- I = J = 10 (without repetitions).
- $e_j \sim N(0, s_e)$ , with  $s_e = c(1, 5)$ .
- $g_i \sim N(0, s_g)$ , with  $s_g = c(1, 5)$ .

#### Simulation full model $(y_{ij} = g_i + e_j + g_i \times e_j)$

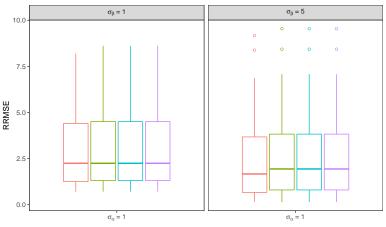






### Simulation full model $(y_{ij} = g_i + e_j + g_i \times e_j)$







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