

1A]

SHANNON DEFINED "PERFECT SECRECY", i.e., THE FACT THAT THE CYPHERTEXT CARRIES NO INFORMATION ABOUT THE PLAINTEXT.

From DEF:

LET M BE A RANDOM VARIABLE OVER \mathcal{M} , AND K BE A UNIFORM DISTRIBUTION OVER K . (ENC, DEC) HAS PERFECT SECRECY IF:
 $\forall M, M_m \in \mathcal{M}, c \in \mathcal{C}:$

$$\Pr[M = m] = \Pr[M = m | C = c]$$

I REFUTE THIS SENTENCE SHOWING AN ENCRYPTION SCHEME WITH NON-UNIFORM DISTRIBUTION C , BUT PERFECT SECRECY IS STILL VALID.

ASSUME Π IS A PERFECT SECRET ENCRYPTION SCHEME.

I BUILD $\Pi'(\text{Enc}', \text{Dec}')$ AS: $(\text{KEYSPACE } K, \text{MESSAGE SPACE } \mathcal{M}, \text{CYPHERTEXT } \mathcal{C}, M \text{ r.v. over } \mathcal{M}, \text{KEYSPACE } K, \text{MESSAGE SPACE } \mathcal{M}, \text{CYPHERTEXT } \mathcal{C}, M \text{ r.v. over } \mathcal{M})$

$\text{Enc}'(K, M) :$

$$C \leftarrow \Pi \text{Enc}(K, M)$$

$$b \leftarrow \$\{0, 1\}^{10}$$

IF $b \bmod 3 == 01$:

RETURN $C || 01$

ELSE RETURN $C || 00$

$\text{Dec}'(K, C) :$

$C' \leftarrow \text{FIRST } M-2 \text{ bit of } C$

RETURN $\Pi \text{Dec}(K, c)$

THIS ENCRYPTION SCHEME PRODUCES A NON-UNIFORM CYPHERTEXT DISTRIBUTION. THE PROBABILITY A CYPHERTEXT ENDS WITH 00 IS $\frac{3}{4}$ AND 01 IS $\frac{1}{4}$. STILL THIS ENCRYPTION SCHEME CARRIES NO INFO ABOUT THE MESSAGE, AND THE DECRYPTION DOES NOT DEPEND ON THE 2 LAST BITS OF CYPHERTEXT. SO, Π' IT'S PERFECTLY SECURE EVEN IF IT HAS A NON-UNIFORM CYPHERTEXT DISTRIBUTION.

SO, THE SENTENCE IS WRONG.

1b

(i)

$$M = \{0, 1, 2, 3, 4\} \quad \text{Enc} \Rightarrow C = k + m \bmod 5$$

$$K = \{0, 1, 2, 3, 4, 5\} \quad \text{Dec} \Rightarrow m = c - k \bmod 5$$

ENC

K	0	1	2	3	4	
M	0	1	2	3	4	
0	0	1	2	3	4	
1	1	2	3	4	0	
2	2	3	4	0	1	
3	3	4	0	1	2	
4	4	0	1	2	3	
5	0	1	2	3	4	

$$\Pr(C=0) = 6/30$$

$$\Pr(C=1) = 6/30$$

$$\Pr(C=2) = 6/30$$

$$\Pr(C=3) = 6/30$$

$$\Pr(C=4) = 6/30$$

$$\Pr(C=x) = 6/30 \quad \forall x \in \{0, 4\}$$

Hence THE CIPHERTEXT CARRIES NO INFORMATION ABOUT THE MESSAGE, SO:

$$\Pr[M=m] = \Pr[M=m | C=c]$$

AND THE PERFECT SECRECY IS VALID.

16

(ii)

$$\mathcal{M} = \{m \in \{0,1\}^l : \text{LAST BIT OF } m \text{ IS } 0\}$$

K UNIFORM OVER THE KEY SPACE $K = \{0,1\}^{l-1}$

$$\text{Enc}(K, m) \Rightarrow c = m \oplus (K \parallel 0)$$

$$\text{Dec}(k, c) \Rightarrow m = c \oplus (K \parallel 0)$$

CLEARLY THIS ENC. SCHEME IS PERFECTLY SECURE BECAUSE IS A OTP WITH AN ~~RELEVANT~~ IRRELEVANT MODIFICATION. BY THE WAY, LET'S FORMALIZE THIS:

$$\text{I HAVE TO SHOW } \Pr[M = m] = \Pr[M = m | C = c]$$

$$\bullet \Pr[M = m | C = c] = \frac{\Pr[C = c | M = m] \cdot \Pr[M = m]}{\Pr[C = c]} \quad (1)$$

$$\bullet \Pr[C = c | M = m] = \Pr[\text{Enc}(K, m) = c] =$$

$$\Pr[c = m \oplus (K \parallel 0)] = \frac{1}{2^{l-1}}$$

THIS BECAUSE THE PROBABILITY ONLY DEPENDS ON $l-1$ BITS M AND K. SINCE THEY ARE IN UNIFORM DISTRIBUTIONS $\frac{1}{2^{l-1}}$ IS THE PR.

$$\bullet \Pr[C = c] = \sum_{m \in \mathcal{M}} \Pr[C = c \wedge M = m] = \left\{ \text{by SATURATION} \right\}$$

$$= \sum_{m \in M} \Pr[C=c | M=m] \cdot \Pr[M=m] = \text{(by Bayes)}$$

$$= \sum_{m \in M} \Pr[\text{Enc}(k, M) = c | M=m] \cdot \Pr[M=m] =$$

$$= \sum_{m \in M} \Pr[\text{Enc}(k, m) = c] \cdot \Pr[M=m] =$$

$$= \Pr[\text{Enc}(k, m) = c] \cdot \sum_{m \in M} \Pr[M=m] =$$

$$\text{So, I HAVE } \Pr[C=c] = \frac{1}{2^{l-1}},$$

REPLACING THE FORMULA (1) WITH VALUES I GET:

$$\Pr[M=m | C=c] = \frac{\Pr[C=c | M=m] \cdot \Pr[M=m]}{\Pr[C=c]} =$$

$$= \frac{\left(\frac{1}{2^{l-1}}\right) \cdot \Pr[M=m]}{\left(\frac{1}{2^{l-1}}\right)} = \Pr[M=m] \quad \checkmark$$

2A

$$\mathcal{H} = \left\{ h_s : X \rightarrow Y \right\}_{s \in S} \quad t\text{-WISE INDEPENDENT} \vee \text{SEQUENCES}$$

OF DISTINCT INPUTS $x_1, \dots, x_t \in X$, \mathcal{H} OUTPUT SEQUENCE

$$y_1, \dots, y_t \in Y$$

$$\Pr[h_s(x_1) = y_1 \wedge \dots \wedge h_s(x_t) = y_t : s \in S] = \frac{1}{|Y|^t}$$

(i)

$\forall t \geq 2$:

\mathcal{H} t -WISE INDEPENDENT $\Rightarrow (t-1)$ -WISE INDEPENDENT

SINCE \mathcal{H} IS t -WISE INDEPENDENT, I CAN WRITE THIS AS:

$$\Pr[h_s(x_1) = y_1 \wedge \dots \wedge h_s(x_{t-1}) = y_{t-1}] \cdot \Pr[h_s(x_t) = y_t] =$$

$$= \underbrace{\Pr[h_s(x_1) = y_1 \wedge \dots \wedge h_s(x_{t-1}) = y_{t-1}]}_{(t-1)\text{-WISE INDEPENDENT}} \cdot \frac{1}{|Y|} =$$

$|Y| \leftarrow \text{probability of } t\text{-th element}$

(ii)

I SHOULD PROVE THAT (FOR A UNIFORMLY RANDOM S)

$$\Pr[h_s(x_1) = y_1 \wedge h_s(x_2) = y_2 \wedge h_s(x_3) = y_3] = \frac{1}{|Y|^3}$$

$$\Pr[h_s(x_1) = y_1 \wedge h_s(x_2) = y_2 \wedge h_s(x_3) = y_3] =$$

CONTINUE

$$= \Pr_r \left[\begin{cases} S_0 + S_1 \cdot x_1 + S_2 \cdot x_1^2 = q_1 \\ S_0 + S_1 \cdot x_2 + S_2 \cdot x_2^2 = q_2 \\ S_0 + S_1 \cdot x_3 + S_2 \cdot x_3^2 = q_3 \end{cases} \right] =$$

$$= \Pr_r \left[\begin{pmatrix} 1 & x_1 & x_1^2 \\ 1 & x_2 & x_2^2 \\ 1 & x_3 & x_3^2 \end{pmatrix} \cdot \begin{pmatrix} S_0 \\ S_1 \\ S_2 \end{pmatrix} = \begin{pmatrix} q_1 \\ q_2 \\ q_3 \end{pmatrix} \right] =$$

$$= \Pr_r \left[\begin{pmatrix} S_0 \\ S_1 \\ S_2 \end{pmatrix} = \begin{pmatrix} 1 & x_1 & x_1^2 \\ 1 & x_2 & x_2^2 \\ 1 & x_3 & x_3^2 \end{pmatrix}^{-1} \begin{pmatrix} q_1 \\ q_2 \\ q_3 \end{pmatrix} \right] = \left| \begin{array}{l} \text{THE RIGHT PART OF THE} \\ \text{EQUAL IS CONSTANT, SO} \\ \text{LET'S REPLACE IT FOR} \\ \text{SOME CONSTANTS } k_0, k_1, k_2 \end{array} \right.$$

$$= \Pr_r \left[\begin{pmatrix} S_0 \\ S_1 \\ S_2 \end{pmatrix} = \begin{pmatrix} k_0 \\ k_1 \\ k_2 \end{pmatrix} \right] = \frac{1}{|y|^3} \quad \text{5 UNIFORMLY RANDOM}$$

THIS IS VALID ONLY IF $\boxed{\Delta}$ CAN INVERT THE MATRIX, SO LET'S COMPUTE THE DETERMINANT.

$$\det \begin{pmatrix} 1 & x_1 & x_1^2 \\ 1 & x_2 & x_2^2 \\ 1 & x_3 & x_3^2 \end{pmatrix} = (x_2 x_3^2) + (x_1 x_2^2) + (x_1^2 x_3) - (x_2^2 \cdot x_3) - (x_1 \cdot x_3^2)$$

$$- (x_2^2 \cdot x_3) =$$

$$= x_2 x_3 (x_3 - x_2) + x_1 x_2 (x_2 - x_1) + x_1 x_3 (x_1 - x_3) \neq 0$$

INPUTS DISTINCT



2b]

SAY THAT X IS (k, n) -SOURCE IF $X \in \{0, 1\}^n$, $\text{MIN-ENTROPY}(X) \geq k$.

(i) $l = 128$

MINIMAL AMOUNT OF MIN ENTROPY NEEDED TO OBTAIN $\epsilon = 2^{-80}$

$$-\text{MIN ENTROPY} \Rightarrow 2 \log_2 \left(\frac{1}{\epsilon} \right) - 2 + l = 286 \text{ (k)}$$

$$-\text{ENTROPY LOSS} \Rightarrow k \geq l + \delta \Rightarrow 286 \geq 128 + 8 \Rightarrow \delta = 158$$

(ii)

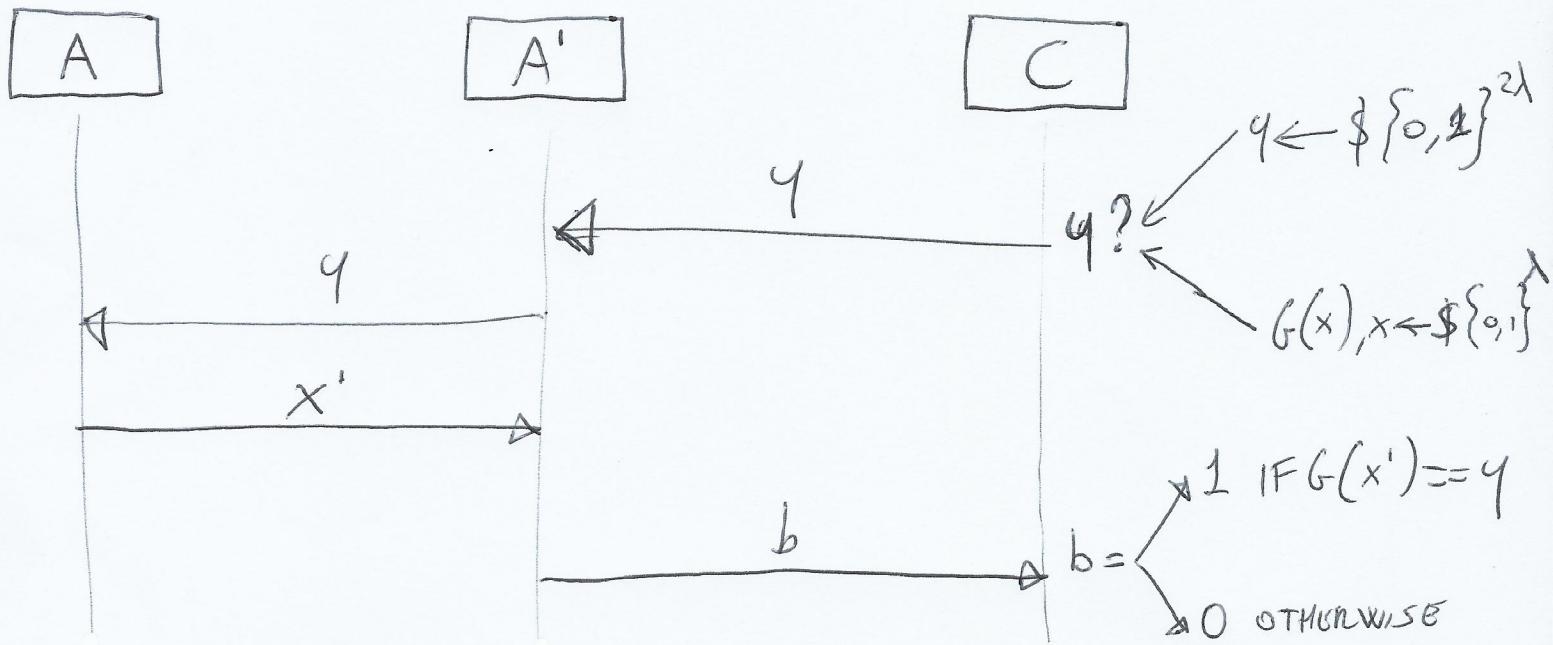
$$\delta = 158$$

$$238 \geq l + 158 \Rightarrow l \leq 238 - 158 = 80 = l$$

3A

$$G: \{0,1\}^\lambda \rightarrow \{0,1\}^{2\lambda} \text{ PRG with } \lambda\text{-bit stretch.}$$

G is OWF?



Let's assume G is not a OWF. So, there exists a PPT attacker A able to find pre-images of G with pr. negl .

$$\Pr[A(q) = 1, x \leftarrow \$\{0,1\}^\lambda, q = G(x)] \geq \frac{1}{\text{poly}(\lambda)} \quad (\text{A})$$

$$\Pr[A(q) = 1, q \leftarrow \$\{0,1\}^{2\lambda}] \leq \frac{2^\lambda}{2^{2\lambda}} = \frac{1}{2^\lambda} \left(\frac{\text{VALID INPUTS}}{\text{POSSIBLE INPUTS}} \right) \quad (\text{B})$$

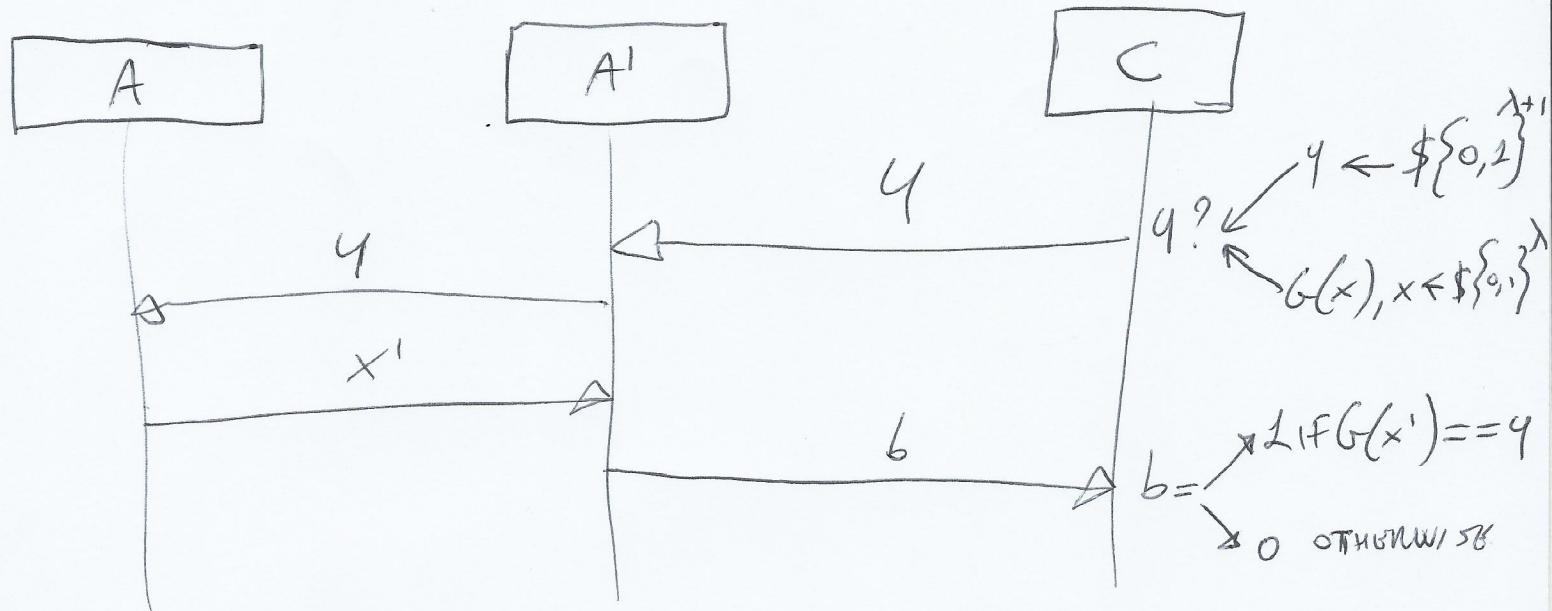
~~VALID INPUTS / POSSIBLE INPUTS~~

LET'S TAKE THE FIRST ELEMENT (A) AND THE SECOND (B).

$$\left| \Pr[(\text{A})] - \Pr[(\text{B})] \right| \geq \frac{1}{\text{poly}(\lambda)} - \frac{1}{2^\lambda} \notin \text{negl}(\lambda).$$

3b]

As I did in the previous solution, I assume G is not a OWE. Then exists a PPT attacker A able to find preimages of G with Pr. Non-NBC .



$$\Pr[A(y) = 1, x \in \{0,1\}^{\lambda}, y = G(x)] \geq \frac{1}{\text{poly}(\lambda)} \quad (A)$$

$$\Pr[A(y) = 1, y \in \{0,1\}^{2^\lambda}] = \frac{2^\lambda}{2^{\lambda+1}} = \frac{1}{2} \left(\frac{\text{VALID INPUTS}}{\text{POSSIBLE INPUTS}} \right) \quad (B)$$

$$\left| \Pr[(A)] - \Pr[(B)] \right| \geq \frac{1}{\text{poly}(\lambda)} - \frac{1}{2} \notin \text{NEGL}(\lambda).$$

4A

$$G_1, G_2 : \{0,1\}^\lambda \rightarrow \{0,1\}^{\lambda+l} \text{ with } l \geq 1.$$

AT LEAST ONE OF G_1, G_2 IS SECURE PRG.

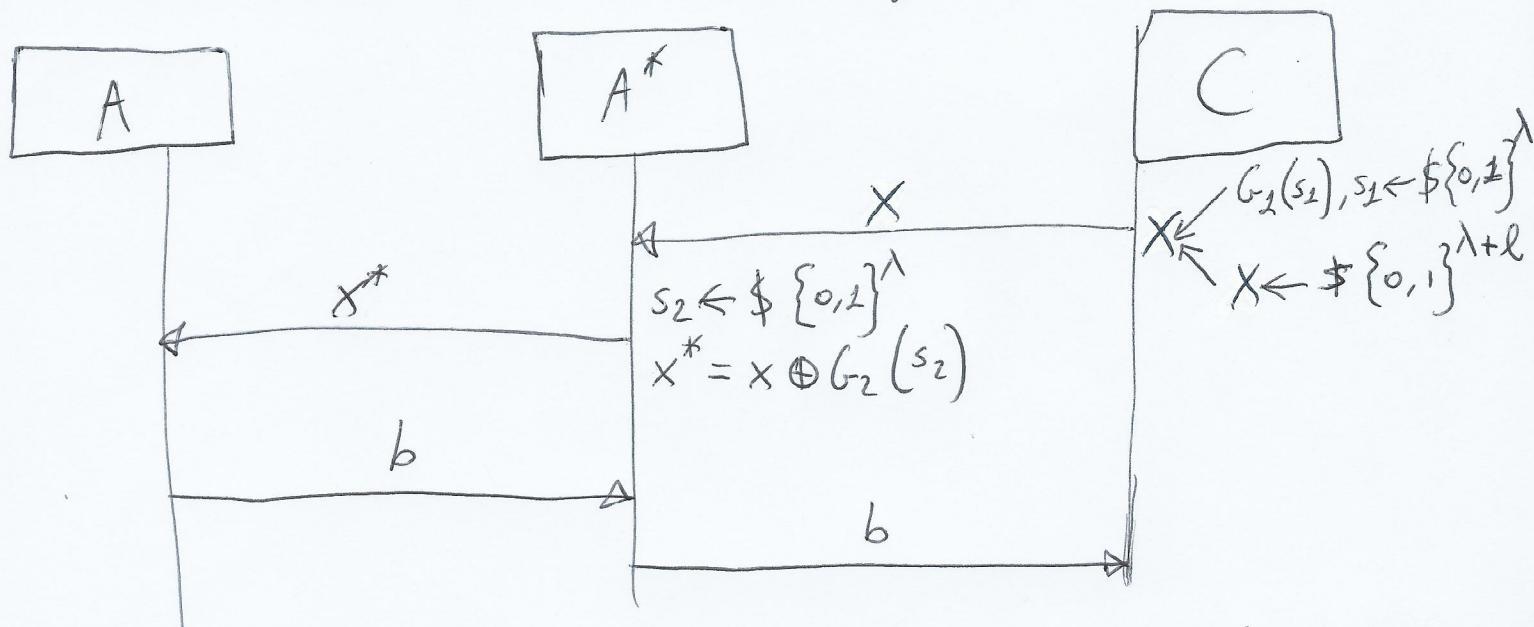
SHOW $G^* : \{0,1\}^{2\lambda} \rightarrow \{0,1\}^{\lambda+l}$ SECURE PRG COMBINING G_1 AND G_2 .

$$G^* = G_1(s_1) \oplus G_2(s_2), \text{ WITH } s_1, s_2 \leftarrow \mathbb{F}\{0,1\}^\lambda$$

I ASSUME G^* ISN'T A SECURE PRG.

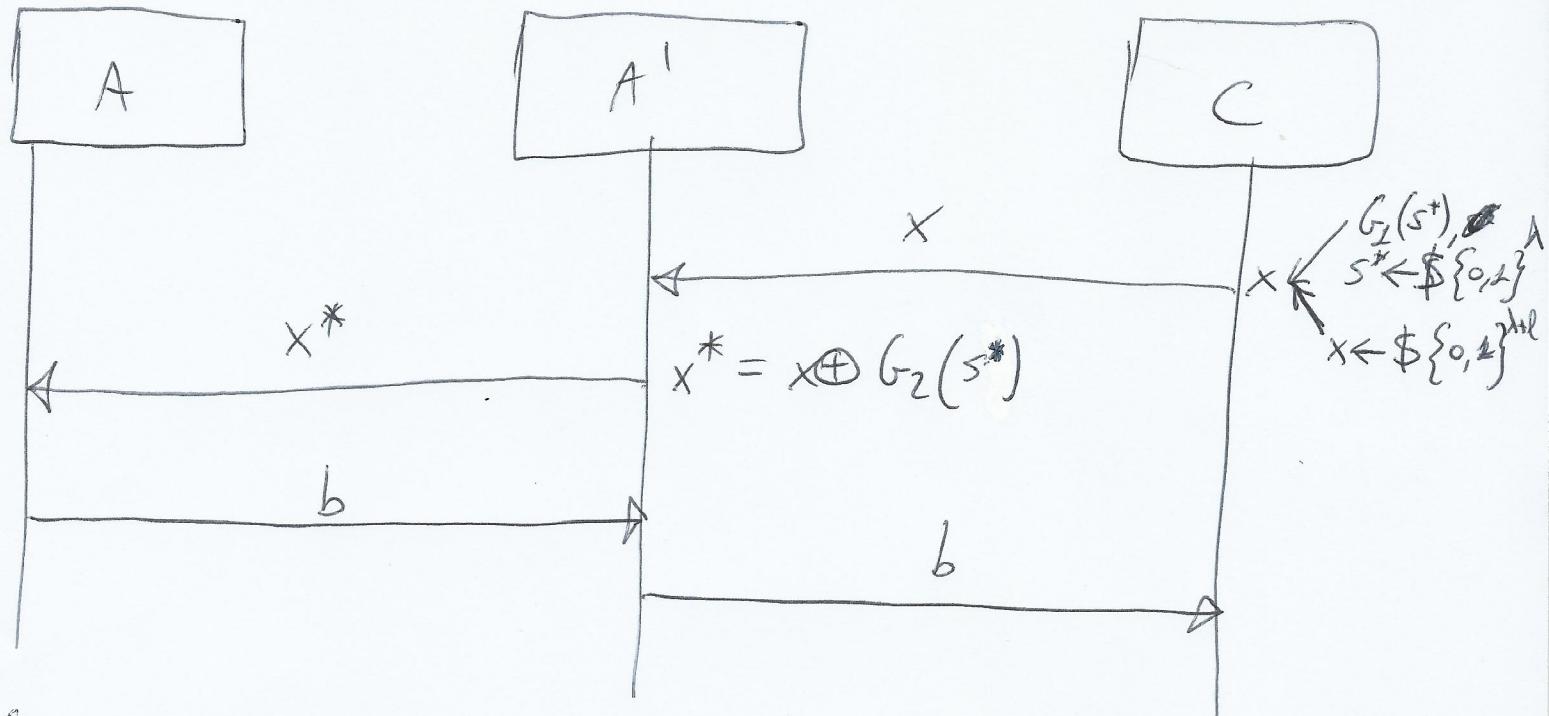
I ASSUME ALSO THAT G_1 IS A SECURE PRG AND EXISTS A PPT ATTACKER A ABLE TO DISTINGUISH BETWEEN G^* AND $\mathbb{U}_{\lambda+l}$.

I CAN BUILD AN ATTACKER A^* THAT BREAKS G_1 .



IF A CAN DISTINGUISH BETWEEN G^* AND $\mathbb{U}_{\lambda+l}$, THEN A CAN DISTINGUISH BETWEEN G_1 AND $\mathbb{U}_{\lambda+l}$ BUT THIS IS IMPOSSIBLE.

4b



GIVEN THE FACT THAT G_1 IS A SECURE PRG, IT PRODUCES A COMPUTATIONALLY INDISTINGUISHABLE OUTPUT (RANDOM).

THEN x^* IS THE RESULT OF x (RANDOM) AND $G_2(s^*)$.

IT WORKS, THIS IS THE ONLY COUNTEREXAMPLE I CAN GIVE IS WHEN $G_2 = G_1 \oplus a$, for $a \in \{0,1\}^{\lambda+1}$.

IN THAT CASE THE OUTPUT WILL BE a , ~~VS.~~

HOWEVER, IF THE SEEDS ARE CHOSEN BOTH RANDOMLY IT'S A NEGIGIBLE PROBABILITY.

5A

LET CONSIDER ANY PRF FAMILY $F_k : \{0,1\}^n \rightarrow \{0,1\}^l$, $k \in \{0,1\}^\lambda$
WITH $n, l \in \text{poly}(\lambda)$.

THE COMPUTATIONALLY UNBOUNDED DISTINGUISHER CAN BRUTE FORCE
 F_k . THIS MEANS THAT A (THE DISTINGUISHER) CAN QUERIES THE
~~CHALLENGER~~ CHALLENGER C WITH ALL INPUTS x_i .

A WILL RETURN 1 IF $\exists k : \forall i F_k(x_i) = q_i$.

IF THE CHALLENGER COMPUTES VALUES WITH PRF, A WILL
RETURN 1 (FOREVER, SO W.P. 1).

SO, $\Pr[A(q)=1, q=F_k(x), k \leftarrow \$\{0,1\}^l] = 1$ (A)

INSTEAD, THE PROBABILITY THAT THE CHALLENGER WILL TAKE
A FUNCTION INSTEAD F_k IS $\frac{2^\lambda}{(2^\ell)^{2^m}}$ $\left(\frac{\text{VARS INPUTS}}{\text{POSSIBLE INPUTS}} \right)$ (B)

SO, $\Pr[A(q)=1, q=G(x), G \leftarrow \$\text{Functions}(m, l)] \leq \frac{2^\lambda}{(2^\ell)^{2^m}}$

I HAVE THEN:

$$\left| P[A] - P[B] \right| \geq 1 - \frac{2^\lambda}{(2^\ell)^{2^m}} \notin \text{negl.}$$

5b]

(i) THIS DOES NOT WORK. LET'S TAKE TWO INPUTS x AND x' .

$$x \oplus x' = F_K(x) \oplus F_K(x')$$

IF I TAKE TWO INPUTS AND A DISTRIBUTION A , A QUERIES C WITH THOSE INPUTS. IF C COMPUTES F_K , I GET:

$$x \oplus x' = F_K(x) \oplus F_K(x') \Rightarrow x \oplus x' = G'(K) \oplus x \oplus G'(K) \oplus x' \Rightarrow x \oplus x' = x \oplus x'.$$

OBVIOUSLY A WILL DISTINGUISH WITH PROBABILITY 1.

IF, INSTEAD, C COMPUTES A RANDOM FUNCTION, A WILL RETURN 1 WITH PROBABILITY $\frac{1}{2^\lambda}$.

THUS SO,

$$\left| \Pr \left[A(y_1, y_2) = 1, y_1 = F_K(x_1), y_2 = F_K(x_2), K \leftarrow \mathbb{F}^{\{0,1\}^\lambda} \right] - \Pr \left[A(y_1, y_2) = 1, y_1 = G(x_1), y_2 = G(x_2), G \leftarrow \text{FUNCTIONS } (\lambda, \lambda + l) \right] \right|$$

$$\geq 1 - \frac{1}{2^\lambda} \neq \text{NGL.}$$

5b]

(ii)

$$F_K(x) := F_x(K), F: \{0,1\}^\lambda \times \{0,1\}^\lambda \rightarrow \{0,1\}^{\ell};$$

LET'S ASSUME A BAD CASE K^* :

$$F_{K^*}(x) = 0^\ell$$

THIS HAPPENS WITH NEGIGIBLE PROBABILITY IF THE KEY IS CHOSEN RANDOMLY.

SINCE THE ATTACKER A CAN QUERY THE INPUT (IT'S x), A CAN FIX THE INPUT.

WITH THIS:

- IF THE FUNCTION IS REAL $\rightarrow \Pr = 1$
- IF THE FUNCTION IS RANDOM $\rightarrow \Pr = 1/2^\lambda$

$$\left| \Pr[A(y) = 1, y = F_K(x), K^*] - \right.$$

$$\left. \Pr[A(y) = 1, y = G(x), G \in \text{FUNCTIONS}(\lambda, \lambda, \ell)] \right| \geq 1 - \frac{1}{2^\lambda}$$

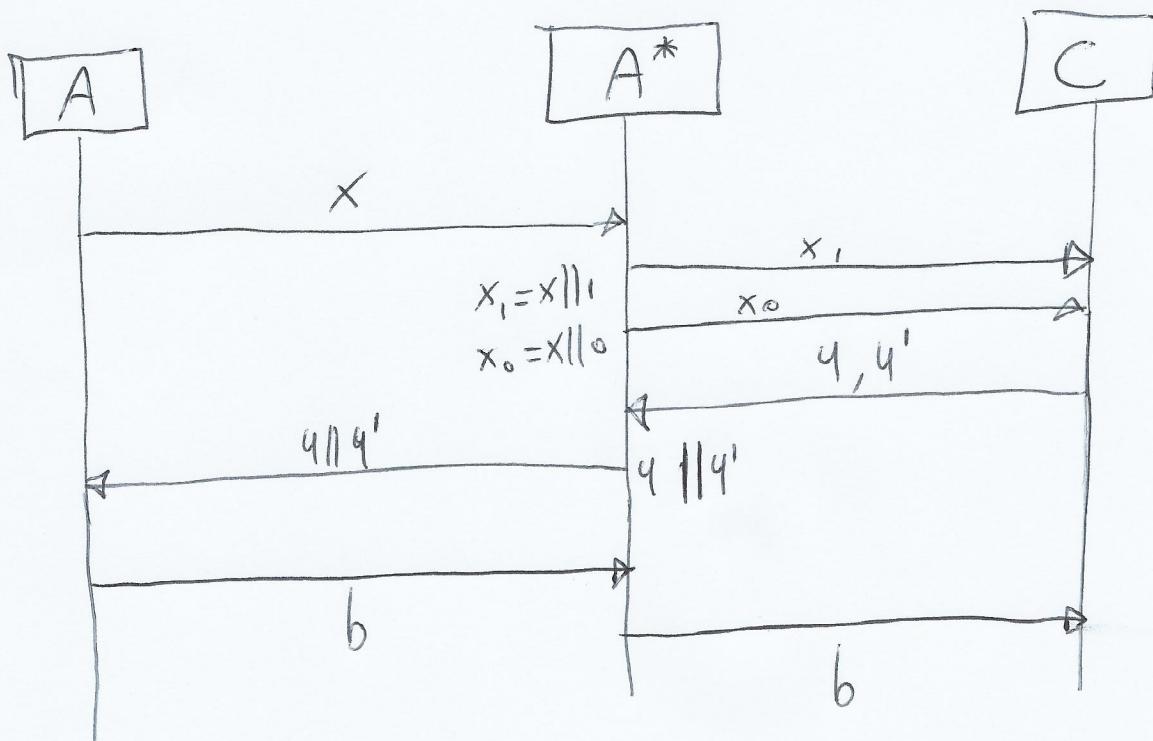
$\not\in$ NEGL.

56
(iii)

$$F'_k(x) = F_k(x||0) \parallel F_k(x||1), \quad x \in \{0,1\}^{m-1}$$

LET'S ASSUME THIS IS NOT A PRC.

SO IT HAS TO EXIST A PPT ATTACKER A ABLE TO BREAK THE SECURITY OF F'_k .



A CAN DISTINGUISH BETWEEN F'_k AND A TRUE RANDOM FUNCTION FUNCTIONS $(m-1, 2^m)$.

A QUERIES WITH x , THEN A^* FORWARDS $x_1 = x||1$ AND $x_0 = x||0$.

C RETURNS BACK q AND q' , AND THEN A^* RETURNS $q \parallel q'$.

A^* FORWARDS THIS VALUE TO A, THAT CAN DISTINGUISH BETWEEN F'_k AND RFUNCTIONS.

SO A^* JUST RETURNS BACK THE OUTPUT b.