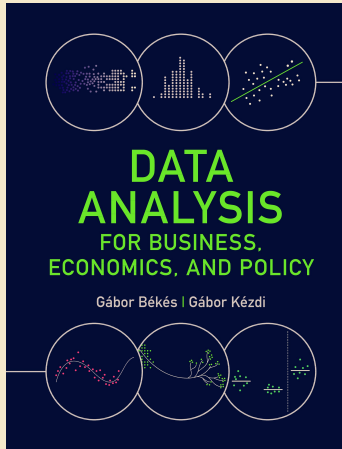


Békés-Kézdi: Data Analysis, Chapter 23: Methods for Panel Data



Data Analysis for Business, Economics, and Policy

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Plan for today

- ▶ Talk about a key method for observational data
- ▶ Panel data methods = generalization of difference in differences in several ways
- ▶ Widely used in academia and real life
- ▶ Very useful to get closer to causality

Multiple Time Periods Can Be Helpful

- ▶ Diff-in-diffs estimates the effect at a single point in time.
- ▶ Issue 1: Immediate effect in one period
- ▶ Most real-life situations: delayed effect, variation of impact over time
 - ▶ Having a single endline time period is not enough to tell the full story.
- ▶ To estimate how an effect plays out in time, need more time periods.
- ▶ Issue 2: subjects may be treated at various points in time
- ▶ Need method(s) that generalize diff-in-diffs for multiple periods.

Plan for today and next week

- ▶ Time series
- ▶ Pooled time series
- ▶ Panel data and first difference
- ▶ Panel data and fixed effects
- ▶ Panel data with first difference and fixed effects
- ▶ Tweak panel data to design the control group
- ▶ Event studies with placebo controls
- ▶ Talk other useful ideas in causal inference

Gábor Békés (Central European University)

Estimating Effects Using Observational Time Series

- ▶ Generalization: multiple periods
- ▶ Estimating an effect from a single time series: within subject comparisons only.
- ▶ An average effect across time for the same subject.
 - ▶ we care about a single country / shop; the intervention happens at one place.
- ▶ Time series regressions
- ▶ specified in levels as well as changes.
 - ▶ y_t variable is measured at which t time period. Could have lags.
 - ▶ Δ denotes change: $\Delta y_t = y_t - y_{t-1}$

Estimating Effects Using Observational Time Series

- ▶ Time series regression specified in levels:

$$y_t^E = \alpha + \beta x_t \quad (1)$$

- ▶ α is the average y when $x = 0$;
- ▶ β shows how much larger y is, on average, when x is larger by one unit.

Estimating Effects Using Observational Time Series

- ▶ Time series regression specified in terms of changes in y and changes in x :

$$\Delta y_t^E = \alpha + \beta \Delta x_t \quad (2)$$

- ▶ α : estimates the trend: the average change in y when x doesn't change.
- ▶ β : how much y changes, on average when x increases (or decreases), by one unit; *in addition* to the trend.
 - ▶ as y_t changes by the trend anyway, so "how much more" is the question.
- ▶ Difference: avoid estimating spurious effects due to trends and random walks
 - ▶ Applied when x is binary or quantitative.

Estimating Effects Using Observational Time Series

- ▶ Causal effect? Yes, if variation in Δx_t is exogenous.
- ▶ time periods with different changes in x would have experienced the same change in y , had x changed the same way for them.
 - ▶ Yes, units are the time periods, as we have a single subject
- ▶ Whatever makes x change at time t should be independent of all other things that would make y change at time t .
 - ▶ Within-subject criterion: changes in x and y are for the same subject.
 - ▶ A version of PTA. In time periods when the treatment status changed ($\Delta x_t \neq 0$), y would have changed the same way, had the treatment status remained the same, as it changed in time periods when the treatment status did remain the same.
- ▶ Please read and think about the gasoline example on p652, p654, p655

Estimating Effects Using Observational Time Series

- ▶ Time series problems are the same as before
 - ▶ Trend - first difference takes care of it
 - ▶ Seasonality - if suspected, add seasonal dummies
 - ▶ Standard error - Newey-West SE
- ▶ Review Chapter 12, especially 12.4, 12.8

Capture dynamics and reverse effects with
leads and lags

Lags to Estimate the Time Path of Effects

- ▶ Advantage of multiple time periods: estimate the time path of effects,
 - ▶ immediate effects,
 - ▶ effects in the near future,
 - ▶ long-run effects.
- ▶ Include appropriate lags of Δx_t .
 - ▶ Application of what we covered earlier

Lags to Estimate the Time Path of Effects

- ▶ With lags, we can estimate effects within the same time period (β_0 below), effects one time period later (β_1), etc.
- ▶ Time series regression that can estimate effects for up to K time periods has K lags of Δx :

$$\Delta y_t^E = \alpha + \beta_0 \Delta x_t + \beta_1 \Delta x_{t-1} + \dots + \beta_K \Delta x_{t-K} \quad (3)$$

Lags to Estimate the Time Path of Effects

- ▶ One-lagged effect: A change in January affects the outcome change in February (and stay that way)
- ▶ A change in how demand grows affects how sales grow next month and beyond.
 - ▶ new growth rate, no reversal back

Lags to Estimate the Time Path of Effects

- ▶ Long-run effect on y = adding up the coefficients on all lags
- ▶ Or apply trick to get cumulative effect:

$$\Delta y_t^E = \alpha + \beta_{cumul} \Delta x_{t-K} + \delta_0 \Delta(\Delta x_t) + \dots + \delta_{K-1} \Delta(\Delta x_{t-(K-1)}) \quad (4)$$

- ▶ $\beta_{cumul} = \beta_0 + \beta_1 + \dots + \beta_K$ above
- ▶ β_{cumul} shows the total change in y within K time periods after a unit change in x , on average.

Lags to Estimate the Time Path of Effects

- ▶ Use either method
- ▶ β_{cumul} shows the total effect of Δx_t on Δy over the long run.
- ▶ Causal effect condition the same: when variation in Δx is exogenous.

Leads to Examine Pre-trends and Reverse Effects

- ▶ We can also include **lead terms** of Δx in the regression.
- ▶ It's analogous to pre-trends in diff-in-diffs regressions
- ▶ It helps capture reverse causality

$$\Delta y_t^E = \alpha + \beta \Delta x_t + \gamma_1 \Delta x_{t+1} + \dots + \gamma_L \Delta x_{t+L} \quad (5)$$

Leads to Examine Pre-trends and Reverse Effects

- ▶ Include **lead terms** of Δx in the regression.
- ▶ Lead $x =$ lagged y
- ▶ Similar role as looking at pre-trends
- ▶ Examine how y did change in the previous time period(s)
- ▶ The parallel trends assumption we need here: analogous to pre-trends in diff-in-diffs regressions
 - ▶ Formally adding a few periods. Few defined by data limitations.

Leads to Examine Pre-trends and Reverse Effects

- ▶ Specific case of endogenous change in x – reverse causality effect: y affecting x .
- ▶ With observations from multiple time periods - capture this reverse effect.
- ▶ IF it takes time.
- ▶ Result of reverse effect: a change in x would tend to follow a change in y .
- ▶ One time period, Δy_t is associated with Δx_{t+1} ,
 - ▶ coefficient capture that reverse effect

Leads to Examine Pre-trends and Reverse Effects

- Include **lead terms** of Δx in the regression. With L leads:

$$\Delta y_t^E = \alpha + \beta \Delta x_t + \gamma_1 \Delta x_{t+1} + \dots + \gamma_L \Delta x_{t+L} \quad (6)$$

- The lead terms are Δx_{t+1} through Δx_{t+L} .
- γ_1 shows how y tends to change one time periods before x changes.
- γ_L shows how y tends to change L time periods before x changes.
- They show that because Δy_t is one time period **before** Δx_{t+1} , two time periods before Δx_{t+2} , etc.
- $\gamma_1 = \dots \gamma_L = 0$ would show that, regardless of how x changes, y tends to change the same way one through L time periods earlier.

Leads to Examine Pre-trends and Reverse Effects

- Causal model with a single series: combine leads and lags
- The lag terms help capture delayed effects.
- The lead terms help capture differences in pre-trends and reverse effects.
- A time series regression, in differences, with K lags and L leads, has the form

$$\Delta y_t^E = \alpha + \beta_0 \Delta x_t + \beta_1 \Delta x_{(t-1)} + \dots + \beta_K \Delta x_{(t-K)} + \gamma_1 \Delta x_{(t+1)} + \dots + \gamma_L \Delta x_{(t+L)} \quad (7)$$

Generalization 2: Multiple time periods, multiple subjects - pooled time series

Pooled Time Series to Estimate the Effect for One Unit

- ▶ Despite the advantages of estimating effects from time series, single time series are rarely used to estimate effects in practice.
- ▶ Time series are rarely long enough

Pooled Time Series to Estimate the Effect for One Unit

- ▶ Despite the advantages of estimating effects from time series, single time series are rarely used to estimate effects in practice.
- ▶ Time series are rarely long enough
- ▶ Even if long, are they relevant? Often, not.

- ▶ One solution: combine time series from several subjects i (cross-sectional units).
- ▶ Idea: time series of similar units are more representative than longer series of a single unit
- ▶ Use domain knowledge to select similar units

Pooled Time Series to Estimate the Effect for One Unit

- ▶ The simplest pooled time series regression estimates a single intercept and a single slope.
- ▶ Most often, though, we include separate intercepts for each i .
- ▶ Doing so allows for trends to be different across i .

$$\Delta y_{it}^E = \alpha_i + \beta \Delta x_{it} \quad (8)$$

- ▶ Here β shows the average change in y , across time and units i , when x increases by one unit.
- ▶ Conditional on i -specific trends: even if different subjects had different trends, this would not affect our estimate.

Pooled Time Series to Estimate the Effect for One Unit

- ▶ We can add leads and lags as before
- ▶ We had two ways to tackle serial correlation: Newey-West SE and adding lagged y_t . Here it's the lagged y_t
- ▶ Data table with pooled time series, N units, each with T_i observations.
 - ▶ There is no specific, ideal N , it's typically 5-20, depends on domain, could be more.
 - ▶ Ideally, each unit has same time series, but can work with them even if not —> end of lecture

Panel Regression: taking stock 1

- ▶ Pooled time series helps to estimate an effect for one unit
- ▶ Often in first difference - trend
- ▶ Using pooled time series, N is small, T is large
- ▶ Add lagged Δy_t to the right hand side of equation to tackle serial correlation

Case study – Import Demand and Industrial Production

- ▶ Interested in understanding how external demand affects production
- ▶ Thai industrial production and US total imports: individual time series
 - ▶ Industrial production in Thailand, in logs, monthly time series
 - ▶ US total imports, in logs, monthly time series
- ▶ Source: asia-industry dataset. N=243.
 - ▶ Monthly data, seasonally adjusted, February 1998–April 2018.

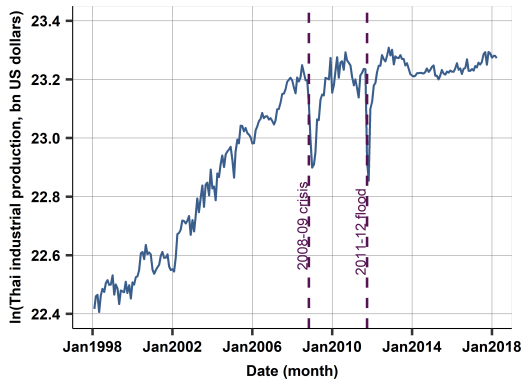
Case study – Thai industrial production and US total imports

- ▶ Question: how the import demand of the USA affects industrial production in Thailand.
- ▶ Causal question, but no explicit intervention.
- ▶ what happens in a mid-sized open economy when something changes externally - major trading partner.
- ▶ Mechanism: global supply chains, Thailand sells to USA directly, and indirectly (often through China).
- ▶ We care about coefficient not just if there is an effect – policy

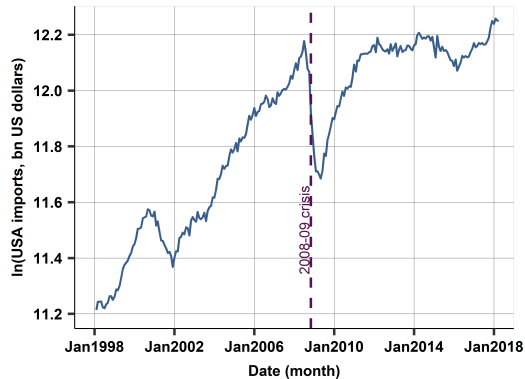
Case study – Thai industrial production and US total imports

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 - ▶ Monthly data, seasonally adjusted, February 1998–April 2018.

Case study – Thai industrial production and US total imports



Thailand IP, in logs, Feb 1998–April 2018, monthly



US total imports, in logs, monthly

Case study – Thai industrial production and US total imports [REV]

- ▶ There is a trend, an extreme event (2009 great crisis), care about relative change
- ▶ First difference. Log values.
- ▶ Lags=4 -a one-time change in U.S. imports can have an effect on how Thai industrial production changes through four months.
- ▶ No leads - expect no reverse causality
- ▶ TS regression estimate the effect of U.S. import demand on Thai industrial production (IP):

$$\begin{aligned} \Delta(\ln(ipTHA)_t) = & \alpha + \beta_0 \Delta(\ln(impUSA)_t) + \beta_1 \Delta(\ln(impUSA)_{t-1}) + \dots \\ & + \beta_4 \Delta(\ln(impUSA)_{t-4}) + \phi \Delta(\ln(ipTHA)_{t-1}) \end{aligned} \quad (9)$$

Case study – Import Demand and Industrial Production

- ▶ US imports and industrial production in Thailand and three other countries
- ▶ Dependent variable is change of log industrial production in each country;
- ▶ Explanatory variable cumulative effect of the change in log US imports, four lags.
- ▶ Add lagged dependent variable to capture serial correlation
- ▶ Monthly time series, seasonally adjusted, February 1998–April 2018. N=243 - for all units

Case study - US imports and IP in Thailand + 3 other countries

Variables	(1) Thailand	(2) Malaysia	(3) Philippines	(4) Singapore	(5) Pooled
USA imports log change, cumulative coeff.	0.400* (0.190)	0.358** (0.112)	0.556** (0.185)	0.367 (0.289)	0.437** (0.103)
Industrial production log change, lag	-0.119 (0.065)	-0.460** (0.059)	-0.242** (0.064)	-0.376** (0.061)	-0.315** (0.031)
Malaysia					0.000 (0.004)
Philippines					-0.001 (0.004)
Singapore					0.002 (0.004)
Constant	0.002 (0.003)	0.004* (0.002)	0.001 (0.003)	0.005 (0.004)	0.003 (0.003)
Observations	238	238	238	238	952
R-squared	0.070	0.231	0.140	0.183	0.123

TS regression; dep.var= change of log industrial production in country; log US imports change: 4 lags.
Monthly, SA, Feb 1998–April 2018. N=243. Standard error estimates in parentheses. ** $p < 0.01$, * $p < 0.05$.

Case study - US imports and IP in Thailand + 3 other countries

- ▶ Estimate is 0.44, 95% confidence interval is [0.24,0.64].
- ▶ Causality: we have good reasons to take estimate as causal effect
 - ▶ First difference takes care of level, trend.
 - ▶ Unlikely reverse causality (but may add leads **Try at home.**)
- ▶ What can go wrong?

Case study - US imports and IP in Thailand + 3 other countries

- ▶ Estimate is 0.44, 95% confidence interval is [0.24,0.64].
- ▶ Causality: we have good reasons to take estimate as causal effect
 - ▶ First difference takes care of level, trend.
 - ▶ Unlikely reverse causality (but may add leads **Try at home.**)
- ▶ What can go wrong?
- ▶ A confounder affecting the **change** in output and demand
- ▶ **Examples?**

Generalization 3: Multiple time periods and subjects - xt panel data, FE model

Panel Regression

- ▶ Pooled time series from a **few** subjects to estimate the expected effect of a causal variable x on outcome y . Policy question was for **one of the subjects**.
- ▶ Change of question: the average effect of x on y across many subjects.
- ▶ Same kind of question to diff-in-diffs, but multiple periods
- ▶ So we'll have: N units, over T periods
- ▶ Will look at different models, approaches

Panel Regression with Fixed Effects

- ▶ Typically N is large, T is relatively small
- ▶ \rightarrow time series aspects less important
- ▶ Start with levels (y_t and not Δy_t)

Panel Regression with Fixed Effects

- ▶ So, setup: multi-period panel data
- ▶ First model is the **fixed-effects regression** (FE regression).
- ▶ In FE regressions we have y and x (in levels)
- ▶ Fixed effects are separate intercepts for different cross-sectional units.
- ▶ The simplest linear panel regression with cross-section fixed effects:

$$y_{it}^E = \alpha_i + \beta x_{it} \quad (10)$$

- ▶ The fixed effects are denoted by α_i .
- ▶ Intercept varies for different cross-sectional units.

Panel Regression with Fixed Effects

- ▶ Like pooled time series in levels, but
 - ▶ ... many units,
 - ▶ ...a short time series
-
- ▶ We look for average relationship

Panel Regression with Fixed Effects

- ▶ Why do we include the fixed effects?
 - ▶ Separate intercepts for each xsec unit instead of a common intercept?
- ▶ IF subjects tend to have higher y on average due to some unobserved confounder that affects x or y in the same way at all times.
- ▶ THEN, fixed effects help avoid/mitigate bias. Including fixed effects = conditioning on all variables that don't change through time.
- ▶ Fixed effects condition on confounders that do not change in time (time-invariant)

Panel Regression with Fixed Effects - mean differencing

- Technical detour: fixed effects is like mean differencing.
- Inclusion of the cross-sectional fixed effects acts as a transformation of the y and x variables into differences from their cross-sectional means: $y_{it} - \bar{y}_i$ and $x_{it} - \bar{x}_i$,
- where \bar{y}_i and \bar{x}_i are average values of y and x across all time periods within cross-sectional unit i .
- β in the model $y_{it}^E = \alpha_i + \beta x_{it}$ is exactly the same as the β in the model $(y_{it} - \bar{y}_i)^E = \alpha + \beta(x_{it} - \bar{x}_i)$.

Panel Regression with Fixed Effects - coefficients

- In the FE regression, β shows how much larger y is, on average, compared to its mean within the cross-sectional unit, where and when x is higher by one unit compared to its mean within the cross-sectional unit.
- Compare two observations that are different in terms of the value of x compared to its i -specific mean. On average, y is larger, compared to its i -specific mean, by β , for the observation with the larger x value.

Panel Regression with Fixed Effects - coefficients

- ▶ In the FE regression, β shows how much larger y is, on average, compared to its mean within the cross-sectional unit, where and when x is higher by one unit compared to its mean within the cross-sectional unit.
- ▶ That's a within-subject comparison, and it's not affected by whether one subject has larger average y .
- ▶ That's why it's not affected by whether an unobserved confounder affects the average y values of the different subjects.

Panel Regression with Fixed Effects - coefficients: where / when

- ▶ In the FE regression, β shows how much larger y is, on average, compared to its mean within the cross-sectional unit, where and when x is higher by one unit compared to its mean within the cross-sectional unit.
- ▶ “where and when” - β approximates the average pattern of association across both time and space. Linear specification = an approximation to the average pattern of association :
 - ▶ different cross-sectional units (cross-sectional heterogeneity of the association),
 - ▶ and/or across different time periods (changing patterns of association).
 - ▶ + as always, the pattern itself may be different for different values of x (nonlinearity),

Panel Regression with Fixed Effects : Fruits and income

- ▶ how much more fruit and vegetables people eat (compared their 10-year average) when their income is higher than their average 10-year income.
- ▶ this effect is not confounded by anything that is stable over time (such as gender, genes)
- ▶ Fixed effects soak up all time-invariant variation
- ▶ Other potential confounders that affect both how income in year t is different from average and how fruit consumption is different to average.
- ▶ Example?

Panel Regression with Fixed Effects

- ▶ Technical note 1. R-squared has two versions
- ▶ within R-squared - based on the transformed model, ie comparing mean differenced y and x
- ▶ R-squared of full model - the FE model as is, with binary variables for all cross-sectional units
- ▶ Preference for within R-squared - more meaningful re model.
 - ▶ Key: make sure to report which one.
- ▶ Technical note 2. Don't publish constant - fixed effects include it.

Aggregate Trend in panel data

- ▶ Aggregate trend is a global trend that affects all unit the same way
- ▶ such as global business cycle
- ▶ varies across time periods but not units
- ▶ With xt panel data, we can condition on an aggregate trend, whatever form it has, including nonlinear trends or even ups and downs.

Aggregate Trend in panel data

- ▶ To condition on aggregate trends, we need to include **time dummies**: binary variables for each time period.
- ▶ Sometimes called **time fixed effects**

$$y_{it}^E = \alpha_i + \theta_t + \beta x_{it} \quad (11)$$

- ▶ β shows how much larger y is, on average, compared to its mean within the cross-sectional units and its mean within the time period, where and when x is higher by one unit compared to its mean within the cross-sectional unit and its mean within the time period.

Aggregate Trend in panel data

$$y_{it}^E = \alpha_i + \theta_t + \beta x_{it} \quad (12)$$

- β shows how much larger y is, on average, compared to its mean within the cross-sectional units and its mean within the time period, where and when x is higher by one unit compared to its mean within the cross-sectional unit and its mean within the time period.
- β shows how much larger y is, on average, compared to its *long run mean* and its *aggregate trend*, where and when x is higher by one unit compared to its *long run mean* and its *aggregate trend*.

Clustered Standard Errors

- ▶ Instead of heteroskedasticity robust SE (cross-section) or Newey West SE (time series), we'll use a new type called clustered standard error.
- ▶ Standard errors clustered at the level of cross-sectional units
- ▶ Idea: adjust standard errors to capture that observations in time series are not independent (serial correlation is likely)
- ▶ Clustered standard errors are robust in two aspects. They are fine in the presence of any kind of serial correlation, and they are also fine without any serial correlation.
- ▶ They are also fine in the presence of heteroskedasticity as well as homoskedasticity
- ▶ Thus, with panel models, we always use clustered SE.
- ▶ we need a not small (>30) number of units

Panel Regression: taking stock 2

- ▶ Fixed effect regression with dummies for aggregate trend - causality?
- ▶ The cross-sectional FE regression can get us **closer to causal effect** of x on y
- ▶ Conditioning on confounders that don't change;
- ▶ Condition on aggregate trends of any shape.
- ▶ Use clustered standard error

Case study B: Immunization against Measles and Saving Children

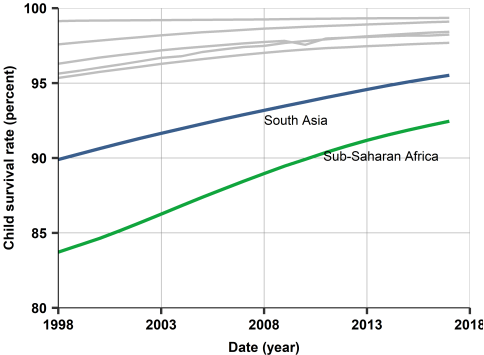
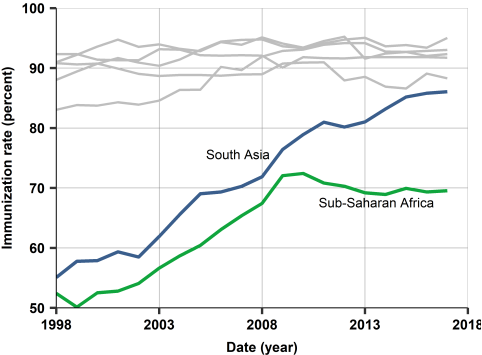
Case study – Immunization against Measles and Saving Children

- ▶ A case study about vaccines.
- ▶ We picked in 2018

Case study – Immunization against Measles and Saving Children

- ▶ Immunization against measles and child survival rate in seven regions of the world
 - ▶ Immunization rate
 - ▶ Child survival rate
 - ▶ Immunization rate: percentage of children of age 12 to 23 months who received vaccination against measles.
 - ▶ Child survival rate: 100% minus the percentage of children of age 0 to 5 years who died in the given year.
- ▶ Source: worldbank-immunization dataset.
- ▶ Annual data, 1998–2017, aggregated to seven geographical regions.
- ▶ Many, but not all countries, N=172

Case study - Immunization against measles and child survival rate in seven regions of the world



Immunization rate Child survival rate
Source: worldbank-immunization dataset. Annual data, 1998–2017, aggregated to seven geographical regions.
N=140

Case study - the effect of measles immunization on child survival. FE regressions

- ▶ The effect of measles immunization on child survival.
- ▶ FE regressions
- ▶ Within R-squared presented for FE regressions.
- ▶ Source: worldbank-immunization dataset;
- ▶ balanced yearly panel, years 1998–2017 in 172 countries.

Case study - the effect of measles immunization on child survival. FE regressions

Variables	(1) Survival rate	(2) Survival rate
Immunization rate	0.077** (0.010)	0.038** (0.011)
ln GDP per capita		1.593** (0.399)
ln population		12.049** (1.648)
Year dummies	Yes	Yes
Observations	3,440	3,440
R-squared	0.717	0.848
Number of countries	172	172

Within R-squared presented for FE regressions. Appropriate standard error estimates in parentheses. ** $p < 0.01$, * $p < 0.05$. Source: worldbank-immunization dataset; balanced yearly panel, years 1998–2017 in 172 countries.

Case study - the effect of measles immunization on child survival. FE regressions

- ▶ The slope parameter estimate on immunization is 0.077 without conditioning on any confounders
- ▶ drops to 0.038 when we condition on GDP per capita and population
- ▶ When we compare years with the same GDP and population, in years when the immunization rate is higher by 10 percentage points than its average rate within a country, child survival tends to be 0.38 percentage points higher than its average within the country, conditional on aggregate trends in the world.
- ▶ We can expect it to be 0.16 to 0.6 percentage points higher in the general pattern represented by our data.
 - ▶ 100 percent - 3.8 percent, but not a realistic improvement, 10% makes more sense

Case study - the effect of measles immunization on child survival.

- ▶ Clustered SE versus biased simple SE in a FE panel regression
- ▶ FE regressions with different SE estimates
- ▶ Clustered SE versus biased simple SE in a FE panel regression
- ▶ Measles immunization and child survival, FE panel regression estimates.
 - ▶ Within R-squared presented for FE regressions.
- ▶ Source: worldbank-immunization dataset;
- ▶ balanced yearly panel, years 1998–2017 in 172 countries.

Case study - the effect of measles immunization on child survival.

FE regressions with different Simple and Clustered SE estimates.

Variables	(1) Clustered SE	(2) Simple SE
Immunization rate	0.038** (0.011)	0.038** (0.002)
ln GDP per capita	1.593** (0.399)	1.593** (0.071)
ln population	12.049** (1.648)	12.049** (0.227)
Observations	3,440	3,440
R-squared	0.848	0.848
Number of countries	172	172

Within R-squared presented for FE regressions. Standard error estimates in parentheses. ** $p < 0.01$, * $p < 0.05$.

Source: worldbank-immunization dataset; balanced yearly panel, years 1998–2017 in 172 countries.

Generalization 4: Allowing for flexible dynamics in effect, Panel FD model

Panel Regression: taking stock 3

- ▶ So far 1: Pooled time series
 - ▶ $N = \text{small}$, $T = \text{long}$
 - ▶ First difference
 - ▶ Focus on time dimension, dynamics

- ▶ So far 1: FE Panel Regression
 - ▶ $N = \text{large}$, $T = \text{small}$ (may be large)
 - ▶ Level
 - ▶ Focus on within-unit comparison

- ▶ Now combine key ideas
 - ▶ Large N , focus on within unit variation
 - ▶ Allowing for dynamics

Panel Regression in First Differences

- ▶ Setup is the same: xt panel data, many cross-sectional units,
- ▶ panel regression in first differences or FD panel regression.
- ▶ FD = changes $\rightarrow \Delta y_{it} = y_{it} - y_{i(t-1)}$.
- ▶ FD panel regression with a common intercept across all i .

$$\Delta y_{it}^E = \alpha + \beta \Delta x_{it} \quad (13)$$

- ▶ Looks like a pooled a cross-section with first difference.
- ▶ But N is large
- ▶ We have a single intercept, α

Panel Regression in First Differences

$$\Delta y_{it}^E = \alpha + \beta \Delta x_{it}$$

- ▶ β shows the difference in the average change of y for units that experience a change in x during the same period.
- ▶ Comparing different cross-sectional units for the same time, or comparing different time periods for the same unit, β shows how much more y changes, on average, where and when x increases by one unit.

FD Panel Regression : what's new

- ▶ FD panel regression is alternative to FE panel regression
 - ▶ Similar setting
 - ▶ We can capture dynamics
- ▶ Related to pooled time series. But here N is large, T is small(er)

Lags and Leads in FD Panel Regressions

- ▶ Often, we want to estimate not only immediate effects but longer run effects, too.
- ▶ Multiple time periods allow us to capture the time path of the effects by including lags of Δx in the regression.
- ▶ Same idea as with pooled time series
- ▶ Regression in FD with K lags:

$$\Delta y_{it}^E = \alpha + \beta_0 \Delta x_{it} + \beta_1 \Delta x_{i(t-1)} + \dots + \beta_K \Delta x_{i(t-K)} \quad (14)$$

Lags and Leads in FD Panel Regressions

$$\Delta y_{it}^E = \alpha + \beta_0 \Delta x_{it} + \beta_1 \Delta x_{i(t-1)} + \dots + \beta_K \Delta x_{i(t-K)}$$

- ▶ α shows the trend in y : average change in y when x does/did not change
- ▶ β_0 is the contemporaneous slope and it shows how **much more** y changes, on average, for observations with a change in x in the same time period
 - ▶ ... but no change in all K preceding time periods.
- ▶ β_k is the slope on lag number k ($k = 1, 2, \dots, K$) – how much more y changes, on average, for observations with a one unit higher increase in x in the k -th preceding time period
 - ▶ ... but with the same change of x in the current time period as well as all K preceding time periods except for the k th.

Lags and Leads in FD Panel Regressions

- **cumulative effect** or long-run effect of the change of x = sum of the immediate effect and all lagged effects:

$$\beta_{cumul} = \beta_0 + \beta_1 + \dots + \beta_K \quad (15)$$

- Same as before, can add difference to get the cumulative effect directly out

Lags and Leads in FD Panel Regressions ***

- ▶ The trick in a regression with K lags is to include the K^{th} lag and the differences of the previous lags: the cumulative coefficient is then the one on the K^{th} lag.
- ▶ When variables are in first differences, the K^{th} lag is in difference, and the previous lags are differences of the difference.
- ▶ The formula for panel regression is

$$\begin{aligned}
 \Delta y_{it}^E &= \alpha + \beta_0 \Delta x_{it} + \beta_1 \Delta x_{i(t-1)} + \dots + \beta_K \Delta x_{i(t-K)} \\
 &= \alpha + \beta_{cumul} \Delta x_{i(t-K)} + \gamma_0 \Delta(\Delta x_{it}) + \gamma_1 \Delta(\Delta x_{i(t-1)}) + \\
 &\quad + \dots + \gamma_{K-1} \Delta(\Delta x_{i(t-K+1)}) \\
 &\text{with } \beta_{cumul} = \beta_0 + \beta_1 + \dots + \beta_K
 \end{aligned} \tag{16}$$

Lags and Leads in FD Panel Regressions

- ▶ We can also add lead terms to an FD regression to examine pre-trends and capture reverse effects, just like with single time series.
- ▶ An FD panel regression with K lags and L leads looks like this:

$$\Delta y_{it}^E = \alpha + \beta_0 \Delta x_{it} + \beta_1 \Delta x_{i(t-1)} + \dots + \beta_K \Delta x_{i(t-K)} + \gamma_1 \Delta x_{i(t+1)} + \dots + \gamma_L \Delta x_{i(t+L)} \quad (17)$$

- ▶ The γ coefficients on the lead terms are zero if, prior to time periods when x may change, y tends to change the same way regardless of whether and how much x actually changes.

Lags and Leads in FD Panel Regressions

- ▶ Causality if assumption holds re **pre-intervention trends** too: what happened to y before an intervention, or more generally, before a change in the causal variable x
- ▶ Adding leads directly to the model is a bit better than inspecting.
- ▶ But still about the past - it's still an assumption

Aggregate Trend in FD Models

- ▶ As for FE models, we can add time dummies to capture non-linear trend
- ▶ FD regression with K lags and time dummies (time FE) is the following:

$$\Delta y_{it}^E = \theta_t + \beta_0 \Delta x_{it} + \beta_1 \Delta x_{i(t-1)} + \dots + \beta_K \Delta x_{i(t-K)} \quad (18)$$

- ▶ θ_t = coefficients of the time dummies
 - ▶ = time-specific intercepts = time fixed effects.

Individual Trends in FD Models

- ▶ Time dummies capture an aggregate trend in a completely flexible way
- ▶ Cross-sectional units in the data may have their own trends, too.
 - ▶ Here we don't have the opportunity to estimate flexible trends, because we have only one observation for each time period for each unit.
- ▶ Can capture **individual linear trends**: allow the intercept to be different across cross-sectional units.
 - ▶ trend = average change per unit
 - ▶ as with pooled time series

Individual Trends in FD Models

- ▶ FD regression with K lags, time dummies, and individual-specific intercepts:

$$\Delta y_{it}^E = \alpha_i + \theta_t + \beta_0 \Delta x_{it} + \beta_1 \Delta x_{i(t-1)} + \dots + \beta_K \Delta x_{i(t-K)} \quad (19)$$

- ▶ α_i : the average change in y in cross-sectional unit i across all time periods
 - ▶ measured as a deviation from the flexibly estimated aggregate trend θ_t ,
 - ▶ and when x does not change (and didn't change for the past K time periods).

Panel Regression: taking stock 4

- ▶ Model in first difference takes care of level differences (as we look at differences by design)
- ▶ Add aggregate trend (as dummies) and individual linear trends (as unit specific intercept)
- ▶ So this model takes care of confounders that are
 - ▶ correlated with levels of x_t and y_t
 - ▶ correlated the global trends affecting x and y the same way
 - ▶ makes linear trends in unit specific x and y correlate

Case study – Immunization against Measles and Saving Children

- ▶ The immediate and lagged effect of measles immunization on child survival
- ▶ FD panel regression estimates
- ▶ Cumulative effect estimates calculated via transformation.
- ▶ Clustered standard error
- ▶ balanced yearly panel, years 1998–2017 in 172 countries.

Variables	(1) Δ_{surv}	(2) Δ_{surv}	(3) Δ_{surv}	(4) Δ_{surv}
Δimm	0.009** (0.002)	0.010** (0.002)		
Δimm lag 1		0.010** (0.002)		
Δimm lag 2		0.011** (0.002)		
Δimm lag 3		0.009** (0.002)		
Δimm lag 4		0.007** (0.002)		
Δimm lag 5		0.006** (0.002)		
Δimm lead 1				0.008** (0.002)
Δimm lead 2				0.007** (0.002)
Δimm lead 3				0.005 (0.003)
Δimm cumul			0.053** (0.010)	0.054** (0.008)
Constant	0.188** (0.024)	0.136** (0.018)	0.136** (0.018)	0.125** (0.018)
R-squared	0.013	0.078	0.078	0.093
Observations	3,268	2,408	2,408	1,892

Case study – Immunization against Measles and Saving Children

- ▶ The effect of measles immunization on child survival. FD panel regression estimates with year dummies, confounders, and country-specific trends
- ▶ FD panel regressions with 5 lags of all right-hand-side variables.
 - ▶ Cumulative coefficient on the change of immunization over the 5 lags.
 - ▶ Clustered standard error estimates in parentheses.
- ▶ Adding leads - 3 periods

Case study - The effect of measles immunization on child survival

The effect of measles immunization on child survival - FD model estimates

Variables	(1) Δ_{surv}	(2) Δ_{surv}	(3) Δ_{surv}
Δ_{imm} cumulative ,	0.052** (0.010)	0.030** (0.009)	0.011** (0.003)
Year dummies	Yes	Yes	Yes
Confounder variables	No	Yes	Yes
Country-specific trends	No	No	Yes
Observations	2,408	2,408	2,408
R-squared	0.088	0.212	0.331

FD panel regressions with 5 lags of all right-hand-side variables. Confounders: GDP per cap, population. Cumulative coefficient w 5 lags. Clustered SE estimates in parentheses. ** $p < 0.01$, * $p < 0.05$. Source: worldbank-immunization dataset; balanced yearly panel, years 1998–2017 in 172 countries.

Detour: good tables

- ▶ Focus on key causal variable
- ▶ Note but not publish values we don't care about (like global time trend dummies)
- ▶ Detailed footnote
- ▶ N of obs, key stats (here R-squared)
- ▶ Opted for long title - could be a shorter one

Case study - The effect of measles immunization on child survival

- ▶ Baseline result 0.05
- ▶ Year dummies + confounders: 0.030 - confounders clearly important
- ▶ Adding individual linear time trend: 0.011 - small but precisely measured
- ▶ A 10 percent increase in the immunization rate tends to be followed by a 0.1 percentage point increase in the child survival rate within five years in the data relative to its country-specific trend
- ▶ Corresponding expected increase in child survival is 0.05 to 0.17 percentage points in the general pattern represented by the data.

Case study - The effect of measles immunization on child survival

- ▶ Causal effect?
- ▶ We can't be certain. It's observational data.
 - ▶ Country-specific trends: can't be certain that this ensures that the parallel trends assumption,
- ▶ We did a great deal of efforts to condition on all kinds of confounders.
 - ▶ FD model with lags - takes out level differences and accounts for dynamics
 - ▶ Key confounders added: GDP per capita and population + individual linear trends
 - ▶ PTA - make a very good effort: Adding leads or confounders like population, gdp makes no difference.
- ▶ Good approximation to what the true effect: A 10 percent increase in the immunization rate **leads to** a 0.1 percentage point increase in the child survival rate within five years.

Case study - The effect of measles immunization on child survival

- ▶ Ok, so consider a nutrition supplement policy where a nurse visits people.
- ▶ this would help both survival and also makes families easier to get to vaccination.
- ▶ FE, or FD models take care of differences on average, ie like development
- ▶ FD + FE models take care the variation in speed of how such policies are implemented / carried out on site.

- ▶ Broader set of confounders: cross country differences in long-term efforts to make the health system better, with most of its elements getting improved in parallel.

- ▶ And so identification comes from different functional form of changes around individual unit trends

Case study - The effect of measles immunization on child survival

- ▶ Adding all these confounders - could it be too much?
- ▶ Is it possible we partial out some of exogenous variation in x ?
- ▶ Yes. Individual linear trends - if measles vaccination is linear
- ▶ Unlikely
- ▶ But 0.01 may be a lower bound, while 0.03 an upper bound.
- ▶ While of course, other confounders may lurk
- ▶ Analytical choice how to present.

Working with panel data, making decisions

Dealing with Unbalanced Panels

- ▶ Missing observations: missing at random or not
- ▶ If missing at random - okay to keep. Maybe FE models will be better.
- ▶ If not
 - ▶ Reduce T - focus only on more recent years when coverage is high
 - ▶ Reduce N - drop unit (countries) where coverage is low
- ▶ Sample design (filtering out observation) means we have a different sample, and may not be representative to what we started with.
- ▶ Many analytical choice, but must make notes

Panel Regressions and Causality

- ▶ FE regressions and FD regressions can estimate the effect of x on y without the bias due to confounders that don't change over time.
- ▶ Confounders that change through time need to be observed and included in the FE or FD regression.
- ▶ Conditioning on individual trends is feasible with FD regressions
 - ▶ Can do something similar in FE, but (even more) complicated
- ▶ Panel model allow us conditioning on a great deal of confounding factors
- ▶ But, as always, there can be omitted variables - so never certain.

First Differences or Fixed Effects?

- ▶ Have seen many models, which one to choose?
- ▶ FE and FD regressions are similar because both condition on confounders that affect the level of y and x and don't change through time.
 - ▶ FE regressions do that by comparing values of y and x to their cross-sectional means.
 - ▶ FD regressions do something similar by comparing values of y and x to their values in the previous time period.
- ▶ Confounders that affect the change in y or x still matter for both FE and FD regressions, whether the confounders themselves change through time or not

First Differences or Fixed Effects?

- ▶ FD main advantage 1: capture serial correlation by first differencing
 - ▶ important if time series properties key
- ▶ FD main advantage 2: capture transparent dynamics
- ▶ As long as we keep adding lags. But that means smaller and smaller panel for estimation.
 - ▶ FD takes care of linear trend automatically, but as we add anyway, no big deal
- ▶ FD main advantage 3: can easily capture individual linear trends

First Differences or Fixed Effects?

- ▶ FE main advantage 1: simple method of estimating longer run effects, easier to use
 - ▶ estimate of the average of short-term and longterm effects.
 - ▶ When the long-term effects kick in fast, that's a good approximation of the long-term effects themselves
- ▶ FE main advantage 2: Works when missing values in panel (see next bit)

First Differences or Fixed Effects?

- ▶ In many cases, both FD and FE can work.
- ▶ If similar result, use either, and show robustness of method
- ▶ If different - should investigate
 - ▶ Time path of effect
 - ▶ Missing values, too short series for lags
 - ▶ Very strong serial correlation

One more option: Long difference

- ▶ We have seen FE and FD.
- ▶ One other model is the long difference: considering the difference between the end and beginning of our panel.
- ▶ Intervention happens sometime during a long period
- ▶ Technically: a difference in differences regression
 - ▶ before and after are further apart
 - ▶ To capture long run difference
- ▶ α Expected long-term change in y when x does not change
- ▶ β : Compare two units with different changes in x across the long time horizon. y is expected to increase by β more more where or when x increases by one more unit.

One more option: Long difference

- ▶ LD is useful first step, simple to carry out, interpret.
- ▶ Especially when we do not really know when and how interventions happened.
- ▶ least likely to condition on confounders
- ▶ PTA is like: Compare countries that experienced different changes in immunization between 1998-2018.
- ▶ Had they experienced the same change in immunization, child survival would have changed the same way, on average.
- ▶ Least likely to be believed.
- ▶ Hence: in most cases, FD or FE is preferred.

Don't do it at home: POLS

- ▶ "Pooled OLS".

$$y_{it}^E = \theta_t + \beta x_{it}$$

- ▶ Has time dummies
- ▶ NO fixed effects for cross-sectional units.
- ▶ It is an average of cross-sectional OLS regressions.
 - ▶ Looks like a panel model, but in fact, it's not.
- ▶ It's better to pick a single year for cross section.
 - ▶ Maybe check for a few years.

Don't do it at home: RE

- ▶ Random Effects (RE) model
- ▶ A mixture of fixed effects and POLS
- ▶ In rare case, a model leads to it. In PhD only
- ▶ Otherwise, just don't do it.

Panel Regression: taking stock 4

- ▶ Panel data: first difference or fixed effects
 - ▶ Each with some technical issues to pay attention to.
- ▶ Often similar results, but not always
 - ▶ Worth investigating
- ▶ The strictest way is first difference model with cross-sectional unit FE. Protects against individual trends and time-invariant confounders.
- ▶ Some other models
 - ▶ LD is okay to do, less informative re causality
 - ▶ POLS, RE - to avoid.

Summary: Panel Regression

- ▶ Data with multiple time periods can help uncover short- and long-run effects and examine pretrends.
- ▶ When interested in the effects on a single cross-sectional unit, we may analyze a single time series or pool several time series of similar units.
- ▶ With panel data having multiple time periods, several modeling options
- ▶ use an FD regression to uncover the development of the effect over time, and an FD or an FE regression to uncover the long-run effect
- ▶ Watch out for interpretation - hard
- ▶ Overall big picture: using panel data methods can take us much closer to a causal interpretation.

Panel Regression: Some issues

- ▶ Panel data methods – some deeper understanding of possible issues in 2020s
- ▶ Ask your instructor
- ▶ Plus, online updates coming in 2025