Topics in statistics III/IV (MATH3361/4071)

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## Problem class handout 2: Likelihood methods

Lecturer: Georgios Karagiannis

georgios.karagiannis@durham.ac.uk

**Exercise 1.** Consider a  $2 \times 2$  contingency table where  $(n_{i,j})$  is the (i,j)th cell count, and  $\pi_{ij}$  is the (i,j)th cell probability.

1. Show that the marginal distribution of the MLE of the odd ratio  $\hat{\theta}$  is such that

$$\sqrt{n}(\log(\hat{\theta}) - \log(\theta)) \xrightarrow{D} N(0, \frac{1}{\pi_{11}} + \frac{1}{\pi_{12}} + \frac{1}{\pi_{21}} + \frac{1}{\pi_{22}}).$$

2. Show that

$$\frac{\log(\hat{\theta}) - \log(\theta)}{\sqrt{\frac{1}{n_{11}} + \frac{1}{n_{12}} + \frac{1}{n_{21}} + \frac{1}{n_{22}}}} \xrightarrow{D} N(0, 1).$$

**Hint-1:** It is  $\hat{\theta} = \frac{n_{11}n_{22}}{n_{21}n_{12}} = \frac{p_{11}p_{22}}{p_{21}p_{12}}$ , where  $p_{i,j} = n_{i,j}/n$ .

Hint-2: From Example 12 (Handout 2), it is:

$$\sqrt{n}(\boldsymbol{p} - \boldsymbol{\pi}) \xrightarrow{D} N(0, \operatorname{diag}(\boldsymbol{\pi}) - \boldsymbol{\pi} \boldsymbol{\pi}^T)$$

where

$$\operatorname{diag}(\boldsymbol{\pi}) - \boldsymbol{\pi} \boldsymbol{\pi}^T = \begin{bmatrix} (1 - \pi_{11})\pi_{11} & -\pi_{11}\pi_{12} & -\pi_{11}\pi_{21} & -\pi_{11}\pi_{22} \\ -\pi_{11}\pi_{12} & (1 - \pi_{12})\pi_{12} & -\pi_{12}\pi_{21} & -\pi_{12}\pi_{22} \\ -\pi_{11}\pi_{21} & -\pi_{12}\pi_{21} & (1 - \pi_{21})\pi_{21} & -\pi_{21}\pi_{22} \\ -\pi_{22}\pi_{11} & -\pi_{22}\pi_{12} & -\pi_{22}\pi_{21} & (1 - \pi_{22})\pi_{22} \end{bmatrix}$$

Solution.

1.

• In Example 12 (Handout 2), we showed that from the CLT, have

$$\sqrt{n}(\boldsymbol{p} - \boldsymbol{\pi}) \xrightarrow{D} N(0, \operatorname{diag}(\boldsymbol{\pi}) - \boldsymbol{\pi} \boldsymbol{\pi}^T)$$

where

$$\operatorname{diag}(\boldsymbol{\pi}) - \boldsymbol{\pi} \boldsymbol{\pi}^{T} = \begin{bmatrix} (1 - \pi_{11})\pi_{11} & -\pi_{11}\pi_{12} & -\pi_{11}\pi_{21} & -\pi_{11}\pi_{22} \\ -\pi_{11}\pi_{12} & (1 - \pi_{12})\pi_{12} & -\pi_{12}\pi_{21} & -\pi_{12}\pi_{22} \\ -\pi_{11}\pi_{21} & -\pi_{12}\pi_{21} & (1 - \pi_{21})\pi_{21} & -\pi_{21}\pi_{22} \\ -\pi_{22}\pi_{11} & -\pi_{22}\pi_{12} & -\pi_{22}\pi_{21} & (1 - \pi_{22})\pi_{22} \end{bmatrix}$$

for the whole vectorized quantities  $\pi = (\pi_{11}, \pi_{12}, \pi_{21}, \pi_{22})$ , and  $p = (p_{11}, ..., p_{22})$ .

- It is  $\hat{\theta} = \frac{p_{11}p_{22}}{p_{21}p_{12}} \implies \log(\hat{\theta}) = \log(p_{11}) + \log(p_{22}) \log(p_{12}) \log(p_{21})$  So I can specify  $g(x) = \log(x_{11}) + \log(x_{22}) \log(x_{12}) \log(x_{21})$
- It is

$$\dot{g}(x) = \frac{\mathrm{d}}{\mathrm{d}x}g(x) = (\frac{1}{x_{11}}, -\frac{1}{x_{12}}, -\frac{1}{x_{21}}, \frac{1}{x_{22}})$$

and hence  $\dot{g}(x)$  is continuous a.s.

• Because all the assumptions of Delta Method are satisfied, it is

$$\sqrt{n}(\log(\hat{\theta}) - \log(\theta)) \xrightarrow{D} N(0, \dot{g}(\pi)(\operatorname{diag}(\pi) - \pi\pi^{T})\dot{g}(\pi)^{T})$$

with

$$\dot{g}(\boldsymbol{\pi})(\operatorname{diag}(\boldsymbol{\pi}) - \boldsymbol{\pi}\boldsymbol{\pi}^T)\dot{g}(\boldsymbol{\pi})^T = \dot{g}(\boldsymbol{\pi})(1, -1, -1, 1)^T \\
= \frac{1}{\pi_{11}} + \frac{1}{\pi_{12}} + \frac{1}{\pi_{21}} + \frac{1}{\pi_{22}}$$

2. Using Slutsky theorem, and law of large numbers, similar to Example 19 (Handout 2), we find that

$$\frac{\sqrt{\frac{1}{p_{11}} + \frac{1}{p_{12}} + \frac{1}{p_{21}} + \frac{1}{p_{22}}}}{\sqrt{\frac{1}{\pi_{11}} + \frac{1}{\pi_{12}} + \frac{1}{\pi_{21}} + \frac{1}{\pi_{22}}}} \xrightarrow{P} 1$$

and by using Slutsky theorem as in Example 19 (Handout 2), we find

$$\frac{\log(\hat{\theta}) - \log(\theta)}{\sqrt{\frac{1}{n_{11}} + \frac{1}{n_{12}} + \frac{1}{n_{21}} + \frac{1}{n_{22}}}} \xrightarrow{D} N(0, 1).$$

**Exercise 2.** Consider random variables  $X, X_1, X_2, ...,$  where  $\mu_n = E(X - \mu)^n$ , and  $\mu = E(X)$ 

1. Show that,

$$\sqrt{n} \begin{pmatrix} \bar{X} \\ s_x^2 \end{pmatrix} - \begin{bmatrix} \mu \\ \sigma^2 \end{pmatrix} \end{pmatrix} \xrightarrow{D} \mathcal{N} \begin{pmatrix} 0 \\ 0 \end{pmatrix}, \begin{bmatrix} \sigma^2 & \mu_3 \\ \mu_3 & \mu_4 - \sigma^4 \end{bmatrix} \end{pmatrix}$$

2. Show that the asymptotic distribution of the coefficient of variation cv =  $\frac{s_x}{X}$ , is

$$\sqrt{n}(\frac{s_x}{\bar{X}} - \frac{\sigma}{\mu}) \xrightarrow{D} N(0, \frac{\mu_4 - \sigma^4}{4\mu^2\sigma^2} - \frac{\mu_3}{\mu^3} + \frac{\sigma^4}{\mu^4})$$

3. Show that the asymptotic distribution of the 3rd central moment  $m_3 = \frac{1}{n} \sum_{i=1}^n (X_j - \bar{X})^3$  is

$$\sqrt{n}(m_3 - \mu_3) \xrightarrow{D} N(0, \mu_6 - \mu_3^2 - 6\sigma^2\mu_4 + 9\sigma^6)$$

**Hint-1:** It is:

$$\begin{bmatrix} \operatorname{Var}(X_i - \mu) & \operatorname{Cov}((X_i - \mu), (X_i - \mu)^2) & \operatorname{Cov}((X_i - \mu), (X_i - \mu)^3) \\ \operatorname{Cov}((X_i - \mu), (X_i - \mu)^2) & \operatorname{Var}((X_i - \mu)^2) & \operatorname{Cov}((X_i - \mu)^2, (X_i - \mu)^3) \\ \operatorname{Cov}((X_i - \mu), (X_i - \mu)^3) & \operatorname{Cov}((X_i - \mu)^2, (X_i - \mu)^3) & \operatorname{Var}((X_i - \mu)^3) \end{bmatrix} \begin{bmatrix} \sigma^2 & \mu_3 \\ \mu_3 & \mu_4 - \sigma^2 \\ \mu_4 & \mu_5 - \sigma^2 \mu_3 \end{bmatrix}$$

Solution.

1.

• I observe that

$$\begin{bmatrix} \bar{X} \\ s_x^2 \end{bmatrix} - \begin{bmatrix} \mu \\ \sigma^2 \end{bmatrix} = \begin{bmatrix} \bar{X} - \mu \\ \frac{1}{n} \sum_{i=1}^n (x_i - \mu)^2 - (\bar{X} - \mu)^2 \end{bmatrix} - \begin{bmatrix} 0 \\ \sigma^2 \end{bmatrix}$$

$$= \begin{bmatrix} \frac{1}{n} \sum_{i=1}^n (X_i - \mu) \\ \frac{1}{n} \sum_{i=1}^n (X_i - \mu)^2 - (\frac{1}{n} \sum_{i=1}^n (X_i - \mu))^2 \end{bmatrix} - \begin{bmatrix} 0 \\ \sigma^2 \end{bmatrix}$$

$$= \begin{bmatrix} m_1' \\ m_2' - (m_1')^2 \end{bmatrix} - \begin{bmatrix} 0 \\ \sigma^2 \end{bmatrix}$$

where  $m'_1 = \frac{1}{n} \sum_{i=1}^n (X_i - \mu)$  and  $m'_2 = \frac{1}{n} \sum_{i=1}^n (X_i - \mu)^2$ .

- I will try to find the joint asymptotic distribution of  $(m'_1, m'_2)^T$  by CLT, and then the asymptotic distribution of  $(\bar{X}, s_x^2)^T$  by Delta method.
- Let

$$\xi_i = \begin{bmatrix} X_i - \mu \\ (X_i - \mu)^2 \end{bmatrix}$$

IID random vectors. It is

$$\mu_{\xi} = \mathrm{E}(\xi_i) = \begin{bmatrix} \mathrm{E}(X_i - \mu) \\ \mathrm{E}(X_i - \mu)^2 \end{bmatrix} = \begin{bmatrix} 0 \\ \sigma^2 \end{bmatrix}$$

and

$$\Sigma_{\xi} = \operatorname{Var}(\xi_{i}) = \begin{bmatrix} \operatorname{Var}(X_{i} - \mu) & \operatorname{Cov}((X_{i} - \mu), (X_{i} - \mu)^{2}) \\ \operatorname{Cov}((X_{i} - \mu), (X_{i} - \mu)^{2}) & \operatorname{Var}(X_{i} - \mu)^{2} \end{bmatrix}$$
$$= \begin{bmatrix} \sigma^{2} & \mu_{3} \\ \mu_{3} & \mu_{4} - \sigma^{4} \end{bmatrix}$$

since

$$Cov((X_{i} - \mu), (X_{i} - \mu)^{2}) = E(((X_{i} - \mu) - E(X_{i} - \mu))((X_{i} - \mu)^{2} - E(X_{i} - \mu)^{2}))$$

$$= E(((X_{i} - \mu) - \mu_{1})((X_{i} - \mu)^{2} - \mu_{2}))$$

$$= E((X_{i} - \mu)^{3} - (X_{i} - \mu)\mu_{2} - \mu_{1}(X_{i} - \mu)^{2} + \mu_{1}\mu_{2})$$

$$= E(X_{i} - \mu)^{3} - E(X_{i} - \mu)\mu_{2}^{2} - \mu_{1}E(X_{i} - \mu)^{2} + \mu_{1}\mu_{2}$$

$$= E(X_{i} - \mu)^{3} = \mu_{3}$$

It is

$$\bar{\xi} = \begin{bmatrix} m_1' \\ m_2' \end{bmatrix}$$

So by CLT, I have,

$$\sqrt{n}(\begin{bmatrix} m_1' \\ m_2' \end{bmatrix} - \begin{bmatrix} 0 \\ \sigma^2 \end{bmatrix}) \xrightarrow{D} \mathcal{N}(\begin{bmatrix} 0 \\ 0 \end{bmatrix}, \begin{bmatrix} \sigma^2 & \mu_3 \\ \mu_3 & \mu_4 - \sigma^4 \end{bmatrix})$$

• Now, I will calculate the asymptotic distribution of  $(\bar{X}, s_x^2)^T$  by Delta method. Let,

$$g(x,y) = \begin{bmatrix} x \\ y - x^2 \end{bmatrix}$$

with

$$\dot{g}(x,y) = \frac{\mathrm{d}g(x,y)}{\mathrm{d}(x,y)} = \begin{bmatrix} -1 & 0\\ -2x & 1 \end{bmatrix}$$

So

$$g(\underbrace{m_1', m_2'}) = \begin{bmatrix} m_1' \\ m_2' - (m_1')^2 \end{bmatrix} = \begin{bmatrix} \bar{X} - \mu \\ s_x^2 \end{bmatrix}; \qquad g(\underbrace{0, \sigma^2}) = \begin{bmatrix} 0 \\ \sigma^2 \end{bmatrix}$$

$$\dot{g}(\underbrace{0, \sigma^2}) = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}; \qquad \Sigma_g = \dot{g}(\underbrace{0, \sigma^2}) \Sigma_{\xi} \dot{g}(\underbrace{0, \sigma^2})^T = \begin{bmatrix} \sigma^2 & \mu_3 \\ \mu_3 & \mu_4 - \sigma^4 \end{bmatrix}$$

Then, according to Delta theorem

$$\sqrt{n}(g(\bar{\xi}) - g(\mu_{\xi})) \xrightarrow{D} \mathcal{N}(0, \dot{g}(\mu_{\xi}) \Sigma_{\xi} \dot{g}(\mu_{\xi})^{T})$$

$$\sqrt{n}(\begin{bmatrix} \bar{X} \\ s_{x}^{2} \end{bmatrix} - \begin{bmatrix} \mu \\ \sigma^{2} \end{bmatrix}) \xrightarrow{D} \mathcal{N}(\begin{bmatrix} 0 \\ 0 \end{bmatrix}, \begin{bmatrix} \sigma^{2} & \mu_{3} \\ \mu_{3} & \mu_{4} - \sigma^{4} \end{bmatrix})$$

- 2. Since I have the asymptotic distribution of  $(\bar{X}, s_x^2)^T$ , I can use the Delta method.
  - Let  $h(a,b) = \sqrt{b}/a$ , with  $\dot{h}(a,b) = (-\frac{\sqrt{b}}{a^2}, \frac{1}{2a\sqrt{b}})$
  - Then

$$\begin{split} h(\bar{X},s_x^2) &= \frac{s_x}{\bar{X}}; \\ \dot{h}(\mu,\sigma^2) &= \left[ -\frac{\sigma}{\mu^2}, \quad \frac{1}{2\mu\sigma} \right]; \end{split}$$

$$\Sigma_h = \dot{h}(\mu, \sigma^2) \Sigma_g \dot{h}(\mu, \sigma^2)^T$$
$$= \frac{\mu_4 - \sigma^4}{4\mu^2 \sigma^2} - \frac{\mu_3}{\mu^3} + \frac{\sigma^4}{\mu^4}$$

• Then, according to Delta theorem

$$\begin{split} \sqrt{n}(h(\bar{X}, s_x^2) - h(\mu, \sigma^2)) &\xrightarrow{D} \mathrm{N}(0, \dot{h}(\mu, \sigma^2) \Sigma_g \dot{h}(\mu, \sigma^2)^T) \\ \sqrt{n}(\frac{s_x}{\bar{X}} - \frac{\sigma}{\mu}) &\xrightarrow{D} \mathrm{N}(0, \frac{\mu_4 - \sigma^4}{4\mu^2 \sigma^2} - \frac{\mu_3}{\mu^3} + \frac{\sigma^4}{\mu^4}) \end{split}$$

3. I observe that

$$m_3 = \frac{1}{n} \sum_{i=1}^n ((\underbrace{X_i - \mu}_{=Z_i}) - (\underbrace{\bar{X} - \mu}_{=\bar{Z}}))^3 =$$

$$= \frac{1}{n} \sum_{i=1}^n Z_i^3 - 3\frac{1}{n} \sum_{i=1}^n Z_i^2 \bar{Z} + 2\bar{Z}$$

$$= m_3' - 3m_2' m_1' + 2(m_1')^2$$

where  $m'_j = \frac{1}{n} \sum_{i=1}^n Z_i^j = \frac{1}{n} \sum_{i=1}^n (X_i - \mu)^j$ , since  $Z_i = X_i - \mu$ .

• I will use the CLT to calculate the joint asymptotic distribution of  $(m'_1, m'_2, m'_3)^T$  and then I will use Delta method to calculate that of  $m_3$ .

I specify

$$\psi_{i} = \begin{bmatrix} Z_{i} \\ Z_{i}^{2} \\ Z_{i}^{3} \end{bmatrix} = \begin{bmatrix} X_{i} - \mu \\ (X_{i} - \mu)^{2} \\ (X_{i} - \mu)^{3} \end{bmatrix};$$

which are IID, with

$$\bar{\psi} = \frac{1}{n} \sum_{i=1}^{n} \psi_i = \begin{bmatrix} m_1' \\ m_2' \\ m_3' \end{bmatrix}$$

$$\mu_{\psi} = \mathbf{E}(\psi_i) = \begin{bmatrix} \mathbf{E}(X_i - \mu) \\ \mathbf{E}(X_i - \mu)^2 \\ \mathbf{E}(X_i - \mu)^3 \end{bmatrix} = \begin{bmatrix} \mu \\ \sigma^2 \\ \mu_3 \end{bmatrix};$$

$$\Sigma_{\psi} = \text{Var}(\psi_{i}) = \begin{bmatrix} \text{Var}(X_{i} - \mu) & \text{Cov}((X_{i} - \mu), (X_{i} - \mu)^{2}) & \text{Cov}((X_{i} - \mu), (X_{i} - \mu)^{3}) \\ \text{Cov}((X_{i} - \mu), (X_{i} - \mu)^{2}) & \text{Var}((X_{i} - \mu)^{2}) & \text{Cov}((X_{i} - \mu)^{2}, (X_{i} - \mu)^{3}) \\ \text{Cov}((X_{i} - \mu), (X_{i} - \mu)^{3}) & \text{Cov}((X_{i} - \mu)^{2}, (X_{i} - \mu)^{3}) & \text{Var}((X_{i} - \mu)^{3}) \end{bmatrix};$$

$$= \dots \text{calculations...} = \begin{bmatrix} \sigma^{2} & \mu_{3} & \mu_{4} \\ \mu_{3} & \mu_{4} - \sigma^{2} & \mu_{5} - \sigma^{2}\mu_{3} \\ \mu_{4} & \mu_{5} - \sigma^{2}\mu_{3} & \mu_{6} - \sigma^{2}\mu_{3} \end{bmatrix}$$

For instance, you can compute the covariance terms as

$$Cov((X_i - \mu)^2, (X_i - \mu)^3) = E(((X_i - \mu)^2 - E(X_i - \mu)^2) ((X_i - \mu)^3 - E(X_i - \mu)^3))$$

$$= E(((X_i - \mu)^2 - \mu_2) ((X_i - \mu)^3 - \mu_3))$$

$$= E((X_i - \mu)^5 - E(X_i - \mu)^2 \mu_3 - \mu_2 (X_i - \mu)^3 + \mu_2 \mu_3)$$

$$= \mu_5 - \mu_2 \mu_3$$

So by CLT

$$\sqrt{n}\left(\underbrace{\begin{bmatrix} m_1' \\ m_2' \\ m_3' \end{bmatrix}}_{=\bar{\psi}} - \underbrace{\begin{bmatrix} \mu \\ \sigma^2 \\ \mu_3 \end{bmatrix}}_{=\mu_{\psi}}\right) \xrightarrow{D} \text{N}\left(0 \underbrace{\begin{bmatrix} \sigma^2 & \mu_3 & \mu_4 \\ \mu_3 & \mu_4 - \sigma^2 & \mu_5 - \sigma^2 \mu_3 \\ \mu_4 & \mu_5 - \sigma^2 \mu_3 & \mu_6 - \sigma^2 \mu_3 \end{bmatrix}}_{=\Sigma_{\psi}}\right)$$

• Now, in order to find the asymptotic distribution of  $m_3 = m_3' - 3m_2'm_1' + 2(m_1')^2$ , I will use Delta method

Let

$$a(a, b, c) = c - 3ab + 2a^3$$

then

$$\dot{q}(a,b,c) = \frac{\mathrm{d}}{\mathrm{d}(a,b,c)} q(a,b,c) = \begin{bmatrix} -3b + 6a^2, & -3a, & 1 \end{bmatrix}$$

So

$$q(m'_1, m'_2, m'_3) = m'_3 - 3m'_2m'_1 + 2(m'_1)^2$$

$$q(\mu_1', \mu_2', \mu_3') = \mu_3' - 3\mu_2'\mu_1' + 2(\mu_1')^2 = \mu_3$$

and

$$\dot{q}(\mu, \sigma^2, \mu_3) = \begin{bmatrix} -3\sigma^2, & \mu, & 1 \end{bmatrix}$$

and

$$\dot{q}(\mu_{\psi}) \Sigma_{\psi} \dot{q}(\mu_{\psi})^{T} = \begin{bmatrix} -3\sigma^{2}, & \mu, & 1 \end{bmatrix} \begin{bmatrix} \sigma^{2} & \mu_{3} & \mu_{4} \\ \mu_{3} & \mu_{4} - \sigma^{2} & \mu_{5} - \sigma^{2}\mu_{3} \\ \mu_{4} & \mu_{5} - \sigma^{2}\mu_{3} & \mu_{6} - \sigma^{2}\mu_{3} \end{bmatrix} \begin{bmatrix} -3\sigma^{2}, & \mu, & 1 \end{bmatrix}^{T}$$
$$= \mu_{6} - \mu_{3}^{2} - 6\sigma^{2}\mu_{4} + 9\sigma^{6}$$

So the asymptotic distribution of  $m_3$  is such that

$$\sqrt{n}(q(\bar{\psi}) - q(\mu_{\psi})) \xrightarrow{D} N(0, \dot{q}(\mu_{\psi}) \Sigma_{\psi} \dot{q}(\mu_{\psi})^{T})$$

$$\sqrt{n}(m_{3} - \mu_{3}) \xrightarrow{D} N(0, \mu_{6} - \mu_{3}^{2} - 6\sigma^{2}\mu_{4} + 9\sigma^{6})$$