

# Chapter 7 - Estimating CDF and Estistical Functions

## Section 7.1 - The Empirical Distribution Function

$$X_1, \dots, X_n \stackrel{iid}{\sim} F$$

**Definition 7.1:** The empirical distribution function  $\hat{F}_n$  is the CDF that puts mass  $\frac{1}{n}$  at each data point  $x_i$ . Formally,

$$\hat{F}_n(x) = \frac{\sum_{i=1}^n I(X_i \leq x)}{n},$$

Where  $I(X_i \leq x) = \begin{cases} 1, & \text{if } x_i \geq x, \\ 0, & \text{if } x_i < x. \end{cases}$

**Theorem 7.3:** At any fixed value of  $x$ ,

- $\hat{F}_n(x) \xrightarrow{P} F(x)$
- $E(\hat{F}_n(x)) = F(x)$
- $\text{Var}(\hat{F}_n(x)) = \frac{F(x)(1-F(x))}{n}$

$$\text{MSE} = \frac{F(x)(1-F(x))}{n} \rightarrow 0$$

Implement it //

**Theorem 7.4:** (The Glivenko-Cantelli Theorem) Let  $X_1, \dots, X_n \stackrel{iid}{\sim} F$ . Then

$$\sup_x |\hat{F}_n(x) - F(x)| \xrightarrow{a.s.} 0.$$

**Theorem 7.5:** (The Dvoretzky-Kiefer-Wolfowitz (DKW) Inequality) Let  $X_1, \dots, X_n \stackrel{iid}{\sim} F$ , then for any  $\epsilon > 0$ , we have

$$P\left(\sup_x |\hat{F}_n(x) - F(x)| > \epsilon\right) \leq 2^{-2n\epsilon^2}$$

$$P\left(\sup_x |\hat{F}_n(x) - F(x)| > \epsilon\right) + P\left(\sup_x |\hat{F}_n(x) - F(x)| \leq \epsilon\right) = 1$$

$$P\left(\sup_x |\hat{F}_n(x) - F(x)| \leq \epsilon\right) = 1 - P\left(\sup_x |\hat{F}_n(x) - F(x)| > \epsilon\right)$$

$$\left[ \sup_x |\hat{F}_n(x) - F(x)| \leq \epsilon \right] \subset \left[ |\hat{F}_n(x) - F(x)| \leq \epsilon \right]$$

$$\underbrace{1 - 2^{-2n\epsilon^2}}_{1 - 2^{-2n\epsilon^2}} \leq P\left(\sup_x |\hat{F}_n(x) - F(x)| \leq \epsilon\right) \leq P\left(|\hat{F}_n(x) - F(x)| \leq \epsilon\right)$$

$$1 - \alpha = 1 - 2^{-2n\epsilon^2} \Rightarrow 2^{-2n\epsilon^2} = \alpha \Rightarrow -2n\epsilon^2 = \log_2(\alpha) \\ \Rightarrow \epsilon^2 = \frac{\log_2(\alpha)}{-2n} \Rightarrow \epsilon_n = \sqrt{\frac{\log_2(\alpha)}{-2n}}$$

For  $\alpha$  and fixed, we have:

$$- \epsilon_n + \hat{F}_n(x) \leq F(x) \leq \hat{F}_n(x) + \epsilon_n \text{ where } \epsilon_n = \sqrt{\frac{\log_2(\alpha)}{-2n}} \\ \Rightarrow IC(F(x); 1 - \alpha) = \left( \underbrace{\max \left( \sqrt{\frac{\log_2(\alpha)}{-2n}} + \hat{F}_n(x) \right)}_{L(x)}; \underbrace{\min \left( \sqrt{\frac{\log_2(\alpha)}{-2n}}; 1 \right)}_{U(x)} \right)$$

$$P(L(x) \leq F(x) \leq U(x)) \geq 1 - \alpha.$$

Implement:  $\begin{cases} * \text{ Sample (random)} \\ * \text{ real data} \end{cases}$

## Section 7.2 - Statistical Functionals

• Statistical functional (ST):  $T(F)$ ;  $\mu = \int x dF(x)$ ;  $\sigma^2 = \int (x - \mu)^2 dF(x)$ .

Definition 7.7: The plug-in estimator of  $\Theta = T(F)$  is defined by  $\hat{\Theta}_n = T(\hat{F}_n)$ .

In another words, just plug-in  $\hat{F}_n$  for the unknown  $F$ .

Definition 7.8: If  $T(F) = \int r(x) dF(x)$  for some  $r(x)$  then  $T$  is called a linear functional.

$$(T(aF + bG) = aT(F) + bT(G))$$

Theorem 7.9: The plug-in estimator for linear functional  $T(F) = \int r(x) dF(x)$  is:

$$T(\hat{F}_n) = \int r(x) d\hat{F}_n(x) = \frac{1}{n} \sum_{i=1}^n r(X_i)$$

Example 7.10: (The mean)  $\mu = T(F) \approx \int x d\hat{F}_n(x) = \frac{1}{n} \sum_{i=1}^n X_i = \bar{X}$

Example 7.11: (The Variance)  $\sigma^2 = T(F) = \int (x - \mu)^2 dF(x) \approx \int (x - \mu)^2 d\hat{F}_n(x) \\ = \frac{1}{n} \sum_{i=1}^n (X_i - \mu)^2$

Example 7.12: (The Skewness) -

$$K = \frac{\int (x-\mu)^3 dF(x)}{\left\{ \int (x-\mu)^2 dF(x) \right\}^{3/2}} \approx \frac{\int (x-\mu)^3 d\hat{F}_n(x)}{\left\{ \int (x-\mu)^2 d\hat{F}_n(x) \right\}^{3/2}} = \frac{\frac{1}{n} \sum_{i=1}^n (X_i - \hat{\mu})}{\left\{ \frac{1}{n} \sum_{i=1}^n (X_i - \hat{\mu})^2 \right\}^{3/2}}$$

Example 7.14: (Quantiles).  $T(F) = F^{-}(p) \Rightarrow T(F) \approx \hat{F}_n^{-}(p)$

$$\hat{F}_n^{-}(p) = \inf \{ x \mid \hat{F}_n(x) \geq p \}.$$