

Defn: Random Sample

If $X_1, \dots, X_n \stackrel{iid}{\sim} f$ $\leftarrow f$ is some distribution
 then we call the $\{X_i\}$ a random
 sample from f .
 of size n

Fact: If X_i are a random sample then

$$\begin{aligned}
 f(\underline{x}) &= f(x_1, x_2, \dots, x_n) \\
 \uparrow \\
 (x_1, \dots, x_n) &= f_{X_1}(x_1) f_{X_2}(x_2) \dots f_{X_n}(x_n) \quad (\text{b/c independence}) \\
 &= f(x_1) f(x_2) \dots f(x_n) \quad (\text{b/c identical distributions}) \\
 &= \prod_{i=1}^n f(x_i)
 \end{aligned}$$

All together: $f(\underline{x}) = \prod_{i=1}^n f(x_i)$

Defn: Statistic

If $\{X_i\}_{i=1}^n$ are a random sample then if T
 is function $T: \mathbb{R}^n \rightarrow \mathbb{R}$, we say

$T(\underline{X})$ is a statistic.

Ex. ① arithmetic mean

$$\bar{X} = \frac{1}{n} (X_1 + X_2 + X_3 + \dots + X_n)$$

Consider $T: \mathbb{R}^n \rightarrow \mathbb{R}$

$$\text{when } T(X_1, \dots, X_n) = \frac{1}{n} (X_1 + \dots + X_n)$$

then $T(\underline{X})$ is a statistic.

② Sample variance

$$S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2 = T(X_1, \dots, X_n)$$

③ Order Statistics

$$X_{(1)} = \text{minimum of } X_1, \dots, X_n = \min_{i=1, \dots, n} X_i$$

$$X_{(n)} = \text{maximum of } X_1, \dots, X_n = \max_{i=1, \dots, n} X_i$$

$$X_{(r)} = r^{\text{th}} \text{ smallest value among } X_1, \dots, X_n$$

④ range : $X_{(n)} - X_{(1)} = R$

median:

$$M = \begin{cases} X_{(\frac{n+1}{2})} & n \text{ odd} \\ \frac{X_{(n/2)} + X_{(n/2+1)}}{2} & n \text{ even} \end{cases}$$

Defn: Sampling Distribution of a Statistic.

The sampling dist. of a Stat $T(\underline{X})$ is simply the dist. of $T(\underline{X})$.
↑ univariate r.v.

Order Statistics Henceforth: $\{X_i\}_{i=1}^n$ is a random sample

Minimum $X_{(1)} = \min_{i=1, \dots, n} X_i$

Dist of $X_{(1)}$?

$$\boxed{P(X_{(1)} > t)} = P(X_1 > t, X_2 > t, X_3 > t, \dots, X_n > t)$$

$$= P(X_1 > t) P(X_2 > t) P(X_3 > t) \dots P(X_n > t) \quad \text{independence}$$

$$= \prod_{i=1}^n P(X_i > t)$$

$1 - F(t)$

same distribution
w/ a CDF
 F

$$= \prod_{i=1}^n (1 - F(t))$$

$$\boxed{= (1 - F(t))^n}$$

$$P(X_{(1)} > t) = (1 - F(t))^n$$

$$F_{X_{(1)}}(t) = 1 - P(X_{(1)} > t) = 1 - (1 - F(t))^n$$

For continuous r.v.s

$$f_{X_{(1)}}(t) = \frac{d}{dt} F_{X_{(1)}}(t) = \frac{d}{dt} (1 - (1 - F(t))^n) \\ = -n(1 - F(t))^{n-1} (-f(t))$$

$$f_{X_{(1)}}(t) = n(1 - F(t))^{n-1} f(t)$$

pdf for minimum

f PDF of X_i
 F CDF for X_i

Maximum :

$$F_{X_{(n)}}(t) = P(X_{(n)} \leq t) = P(X_1 \leq t, X_2 \leq t, \dots, X_n \leq t) \\ = \prod_{i=1}^n P(X_i \leq t) \\ = \prod_{i=1}^n F(t) \\ = F(t)^n$$

For continuous

$$f_{X_{(n)}}(t) = n F(t)^{n-1} f(t)$$

$$f_{X(n)}(t) = n F(t)^{n-1} f(t)$$

Ex. let $X_i \stackrel{iid}{\sim} \text{Exp}(\lambda)$

$$f(x) = \lambda e^{-\lambda x} \text{ for } x > 0$$

$$F(x) = 1 - e^{-\lambda x}$$

$$f_{X(1)}(t) = n (1 - F(t))^{n-1} f(t)$$

$$= n (e^{-\lambda t})^{n-1} \lambda e^{-\lambda t}$$

$$= (n\lambda) e^{-(n-1)\lambda t} e^{-\lambda t}$$

$$= (n\lambda) e^{-(n\lambda)t} \quad \text{see: this is the pdf of } \text{Exp}(n\lambda)$$

$$X_{(1)} \sim \text{Exp}(n\lambda)$$

$$f_{X(n)}(t) = n F(t)^{n-1} f(t)$$

$$= n (1 - e^{-\lambda t})^{n-1} \lambda e^{-\lambda t}$$

Theorem: $X_{(r)} = r^{\text{th}}$ smallest value among X_1, \dots, X_n

If X_i are continuous

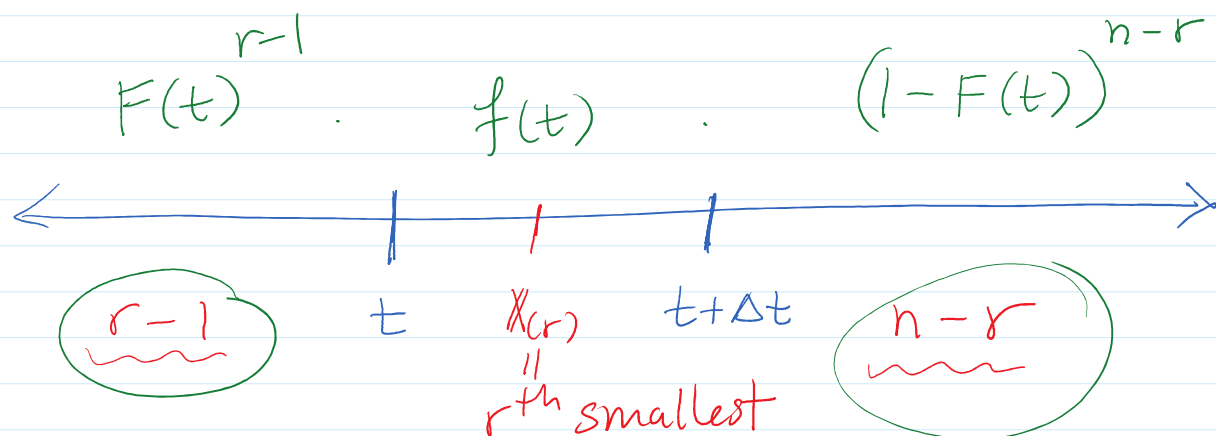
$$f_{X_{(r)}}(t) = \frac{n!}{(r-1)!(n-r)!} F(t)^{r-1} (1-F(t))^{n-r} f(t)$$

$$\frac{1}{n!} \frac{n!}{(r-1)!(n-r)!} \dots$$

F CDF of $\{X_i\}$
 f pdf of $\{X_i\}$

Notice: when $r=1$ or $r=n$
 we get the previous formula.

Pf.



$$f_{X(r)}(t) = \lim_{\Delta t \rightarrow 0} P(t \leq X(r) \leq t + \Delta t)$$

$$= \frac{n!}{(r-1)!(n-r)!} F(t)^{r-1} (1-F(t))^{n-r} f(t)$$

Ex. $X_i \stackrel{iid}{\sim} U(0,1)$

$$f(t) = 1 \quad \text{for } 0 \leq t \leq 1$$

$$F(t) = t \quad \text{for } 0 \leq t \leq 1$$

$$\text{So } f_{X_{(r)}}(t) = \frac{n!}{(r-1)!(n-r)!} F(t)^{r-1} (1-F(t))^{n-r} \underbrace{f(t)}_1$$

$$= \frac{n!}{(r-1)!(n-r)!} t^{r-1} (1-t)^{n-r}$$

for $0 \leq t \leq 1$ PDF of a Beta($r, n-r+1$)

So: $X_{(r)} \sim \text{Beta}(r, n-r+1)$

Theorem: Joint Distribution of Order Statistics

If $r < s$ then $X_{(r)} = r^{\text{th}}$ smallest
 $X_{(s)} = s^{\text{th}}$ smallest

$$X_{(r)} < X_{(s)}$$

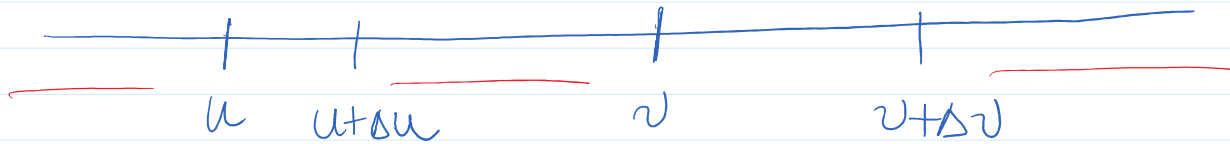
$$f_{X_{(r)}, X_{(s)}}(u, v) = \frac{n!}{(r-1)!(s-r-1)!(n-s)!}$$

$$\cdot F(u)^{r-1} (F(v)-F(u))^{s-r-1} (1-F(v))^{n-s} f(u)f(v)$$

PDF: $r-1$ $X_{(r)}$ $s-r-1$ $X_{(s)}$ $n-s$

1 1

$r-1$ 1 $a-r-1$ 1



$$\frac{n!}{(r-1)!(a-r-1)!(n-a)!} F(u)^{r-1} (1-F(u))^{n-a} (F(v)-F(u))^{a-r-1}$$

Ex. let $X_i \stackrel{iid}{\sim} U(0,1)$ then

$$f_{X_{(r)}, X_{(a)}}(u, v) = \frac{n!}{(r-1)!(n-a)!(a-r-1)!} u^{r-1} (v-u)^{a-r-1} (1-v)^{n-a}$$

recall: $f(t) = 1$ for $0 \leq t \leq 1$
 $F(t) = t$

Ex. $R = X_{(n)} - X_{(1)}$.

Q: what is the dist of R ?

If $X_i \stackrel{iid}{\sim} U(0,1)$

Bivariate Transformation.

$$\text{let } U = R = X_{(n)} - X_{(1)}$$

$$V = X_{(1)}$$

$$X_{(1)} = V \quad \text{so} \quad g_1^{-1}(u, v) = v$$

$$X_{(n)} = u + V \quad g_2^{-1}(u, v) = u + v$$

$$J = \begin{bmatrix} 0 & 1 \\ 1 & 1 \end{bmatrix} \quad \text{so} \quad \underline{|\det J| = 1}$$

So:

$$f_{u,v}(u, v) = f_{X_{(1)}, X_{(n)}}(v, u+v)$$

Aside: $r=1$ and $s=n$

$$f_{X_{(1)}, X_{(n)}}(a, b) = \frac{n!}{(1-1)!(n-n)!(n-1-1)!} \cdot F(a)^{1-1} (F(b)-F(a))^{n-1-1} (1-F(b))^{n-n}$$

Simplify:

$$= \boxed{n(n-1) (F(b)-F(a))^{n-2} f(a)f(b)}$$

plug in:

$n-2$

$$f_{u,v}(u,v) = n(n-1)(F(u+v) - F(v)) f(u+v) f(v)$$

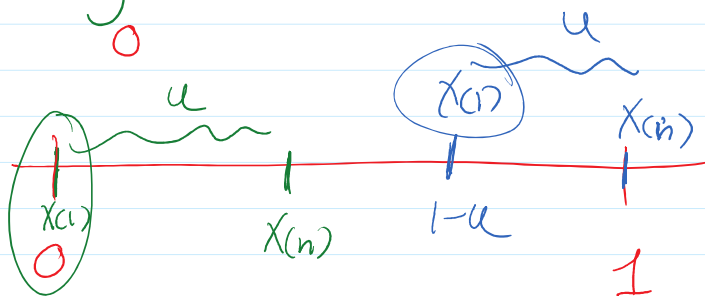
Since: $X_i \stackrel{iid}{\sim} U(0,1)$

$$f(t) = 1 \quad F(t) = t$$

$$= n(n-1) u^{n-2}$$

joint of
 $u = R = X_{(n)} - X_{(1)}$
 $V = X_{(1)}$

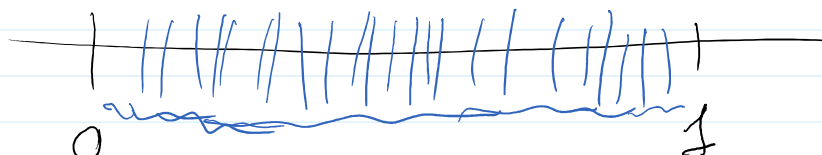
$$f_u(u) = \int_0^{1-u} n(n-1) u^{n-2} dv = n(n-1) u^{n-2} \int_0^{1-u} dv$$

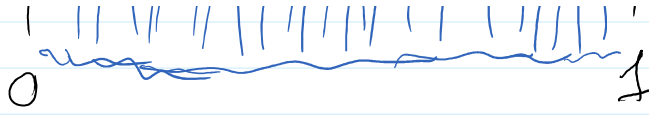


$$u = R = X_{(n)} - X_{(1)} \\ = n(n-1) u^{n-2} (1-u)$$

$$R \sim \text{Beta}(n, 2)$$

Fact: $E[R] = \frac{n-1}{n+1} \xrightarrow{n \rightarrow \infty} 1$





Theorem: Joint Dist. of All Order Stats.

$$X_{(1)}, X_{(2)}, X_{(3)}, \dots, X_{(n-1)}, X_{(n)}$$

$$f_{X_{(1)}, X_{(2)}, X_{(3)}, \dots, X_{(n)}}(u_1, u_2, u_3, \dots, u_n) = n! \prod_{i=1}^n f(u_i)$$

basically $n!$ • Joint ^{PDF} of my random sample.