

Defn: Variance

$$\begin{aligned}\text{Var}(X) &= \mathbb{E}[(X - \mathbb{E}[X])^2] \\ &= \mathbb{E}[X^2] - \mathbb{E}[X]^2\end{aligned}$$

Ex. $X \sim \text{Bin}(n, p)$

recall: $\mathbb{E}[X] = np$

$$\mathbb{E}[X^2] = \sum_{x=0}^n x^2 f(x)$$

$$= \sum_{x=0}^n x^2 \binom{n}{x} p^x (1-p)^{n-x}$$

$$x \binom{n}{x} = n \binom{n-1}{x-1}$$

$$= \sum_{x=1}^n x n \binom{n-1}{x-1} p^x (1-p)^{n-x}$$

$$\begin{aligned}y &= x-1 \\ x &= y+1\end{aligned}$$

$$= \sum_{y=0}^{n-1} (y+1) n \binom{n-1}{y} p^{y+1} (1-p)^{n-y-1}$$

$$= np \left(\sum_{y=0}^{n-1} y \binom{n-1}{y} p^y (1-p)^{(n-1)-y} + \sum_{y=0}^{n-1} \binom{n-1}{y} p^y (1-p)^{(n-1)-y} \right)$$

$$\begin{aligned}\mathbb{E}[\text{Bin}(n-1, p)] \\ &= (n-1)p\end{aligned}$$

$$\begin{aligned}\text{Sum of a Bin}(n-1, p) \\ \text{pmf} \\ &= 1\end{aligned}$$

$$= np((n-1)p + 1) = \mathbb{E}[X^2]$$

∴ $\text{Var}(X) = \mathbb{E}[X^2] - \mathbb{E}[X]^2$

$$\begin{aligned}
 \text{Var}(X) &= E[X^2] - E[X]^2 \\
 &= np(n-1)p + 1 - (np)^2 \\
 &= \cancel{n^2 p^2} - np^2 + np - \cancel{n^2 p^2} \\
 &= np - np^2 \\
 &= np(1-p)
 \end{aligned}$$

$$\begin{aligned}
 \text{Std}(X) &= \sqrt{\text{Var}(X)} \\
 &= \sqrt{np(1-p)}
 \end{aligned}$$

Defn:

If r is a pos. integer we define the r^{th} moment of a r.v. as

$$\mu_r \stackrel{\text{def}}{=} E[X^r]$$

Ex. $\mu_1 = E[X] = \mu$
 $\mu_2 = E[X^2] \quad \dots \quad \mu_3 = E[X^3], \dots$

We define the r^{th} central moment

$$\mu_r' \stackrel{\text{def}}{=} E[(X - \mu)^r]$$

Ex. $\mu_1' = E[X - \mu] = 0$

$$\mu_1 = E[X - \mu] = 0$$

$$\mu_2 = E[(X - \mu)^2] = \text{Var}(X) = \mu_2 - \mu_1^2$$

Defn: Moment Generating Function (MGF)

If X is a r.v. the MGF is a function

$$M: \mathbb{R} \rightarrow \mathbb{R}$$

defined for $t \in \mathbb{R}$ as

$$M(t) = E[e^{tX}]$$

Its defined in some neighborhood of zero.

(possible $M(t) = \infty$ for some t)

For continuous r.v.s,

$$M(t) = E[e^{tX}] = \int_{\mathbb{R}} e^{tx} f(x) dx$$

discrete

$$M(t) = \sum_x e^{tx} f(x)$$

Ex. $X \sim \text{Exp}(\lambda) \rightsquigarrow f(x) = \lambda e^{-\lambda x}$ for $x > 0$

$$E[e^{tX}] = \int_0^\infty e^{tx} \lambda e^{-\lambda x} dx$$

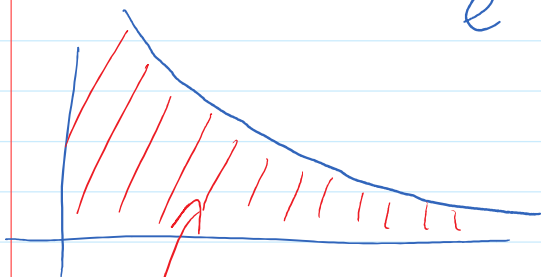
$$f(x) = \lambda e^{-\lambda x} \text{ for } x > 0$$

$$\begin{aligned} M(t) &= E[e^{tx}] = \int_{\mathbb{R}} e^{tx} f(x) dx \\ &= \int_0^{\infty} e^{tx} \lambda e^{-\lambda x} dx \\ &= \lambda \int_0^{\infty} e^{(t-\lambda)x} dx \end{aligned}$$

$$t < \lambda$$

$$t - \lambda < 0$$

$$e^{(t-\lambda)x}$$

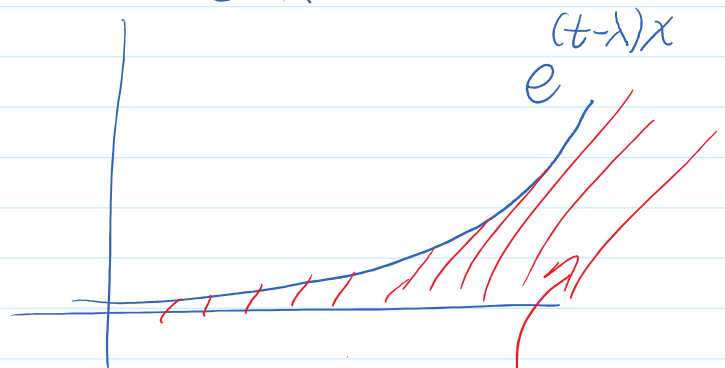


finite area

$$t \in (0, \lambda)$$

$$t \geq \lambda$$

$$t - \lambda \geq 0$$



infinite area

if $t < \lambda$ then

$$\lambda \int_0^{\infty} e^{(t-\lambda)x} dx = \frac{\lambda}{t-\lambda} \left[e^{(t-\lambda)x} \right]_0^{\infty}$$

$$= \frac{\lambda}{t-\lambda} [0 - 1]$$

$$\dots \lambda \quad \rho \quad \perp \quad \dots$$

$$M(t) = \frac{\lambda}{\lambda - t} \quad \text{for } t < \lambda$$

Consider:

$$\frac{dM}{dt} = \frac{\lambda}{(\lambda - t)^2}, \quad \left. \frac{dM}{dt} \right|_{t=0} = \frac{\lambda}{\lambda^2} = \frac{1}{\lambda} = E[X]$$

$$\frac{d^2 M}{dt^2} = \frac{d}{dt} \left(\frac{\lambda}{(\lambda - t)^2} \right) = \frac{2\lambda}{(\lambda - t)^3}$$

$$\left. \frac{d^2 M}{dt^2} \right|_{t=0} = \frac{2\lambda}{\lambda^3} = \frac{2}{\lambda^2} = E[X^2]$$

Theorem:

$$\left. \frac{d^r M}{dt^r} \right|_{t=0} = E[X^r] = \mu_r$$

pf.

$$\frac{d^r M}{dt^r} = \frac{d^r}{dt^r} E[e^{tx}] = \frac{d^r}{dt^r} \int_{\mathbb{R}} e^{tx} f(x) dx$$

$$= \int_{\mathbb{R}} \frac{d^r}{dt^r} e^{tx} f(x) dx$$

$$\frac{d}{dt} e^{tx} = x e^{tx}, \quad \frac{d^2}{dt^2} e^{tx} = \frac{d}{dt} (x e^{tx}) = x^2 e^{tx}$$

generally: $\frac{d^r}{dt^r} e^{tx} = x^r e^{tx}$

$$= \int_{\mathbb{R}} x^r e^{tx} f(x) dx = \frac{d^r M}{dt^r}$$

So $\left. \frac{d^r M}{dt^r} \right|_{t=0} = \int_{\mathbb{R}} x^r f(x) dx = \mathbb{E}[X^r] = \mu_r$

(2)

$$e^x = 1 + x + \frac{x^2}{2!} + \frac{x^3}{3!} + \frac{x^4}{4!} + \frac{x^5}{5!} + \dots$$

$$e^{tx} = 1 + tx + \frac{t^2 x^2}{2!} + \frac{t^3 x^3}{3!} + \dots$$

$$\frac{d}{dt} e^{tx} = x + \frac{2tx^2}{2!} + \frac{3t^2 x^3}{3!} + \dots$$

$$\mathbb{E}\left[\frac{d}{dt} e^{tx}\right] \text{ at } t=0 \rightarrow \mathbb{E}[X]$$

Ex. $X \sim \text{Bin}(n, p)$

$$\begin{aligned} M(t) &= \mathbb{E}[e^{tx}] = \sum_{x=0}^n e^{tx} f(x) \\ &= \sum_{x=0}^n e^{tx} \binom{n}{x} p^x (1-p)^{n-x} \\ &= \sum_{x=0}^n \binom{n}{x} (pe^t)^x (1-p)^{n-x} \end{aligned}$$

$$= \sum_{x=0}^n \binom{n}{x} (pe^t)^x (1-p)^{n-x}$$

Binomial Theorem

$$(a+b)^2 = a^2 + 2ab + b^2$$

$$= \binom{2}{0} a^2 b^0 + \binom{2}{1} a^1 b^1 + \binom{2}{2} a^0 b^2$$

$$\sum_{i=0}^2 \binom{2}{i} a^{2-i} b^i$$

$$(a+b)^n = \sum_{i=0}^n \binom{n}{i} a^i b^{n-i}$$

$$\sum_{x=0}^n \binom{n}{x} (pe^t)^x (1-p)^{n-x} = (pe^t + 1-p)^n = M(t).$$

$$\left. \frac{dM}{dt} \right|_{t=0} = \left. n(pe^t + 1-p)^{n-1} pe^t \right|_{t=0}$$

$$= n(p(1) + 1-p)^{n-1} p(1)$$

$$= np = E[X]$$

$$\begin{aligned} \frac{d^2 M}{dt^2} &= \frac{d}{dt} \left(n(pe^t + 1-p)^{n-1} pe^t \right) \\ &= n(n-1)(pe^t + 1-p)^{n-2} pe^t pe^t \\ &\quad + n(pe^t + 1-p)^{n-1} ne^t \end{aligned}$$

$$- n(n-1)p^2 + 1 - p + np^2 + n(n-1)p^2 + np^2$$

$$+ n(pe^t + 1 - p)^{n-1} pe^t$$

So

$$\left. \frac{d^2 M}{dt^2} \right|_{t=0} = n(n-1)p^2 + np$$

$$= n^2 p^2 - np^2 + np = E[X^2]$$

$$\text{Var}(X) = E[X^2] - (E[X])^2$$

$$= n^2 p^2 - np^2 + np - n^2 p^2$$

$$= np - np^2$$

$$= np(1-p)$$

Theorem: If $a, b \in \mathbb{R}$, and $Y = aX + b$

$$M_Y(t) = e^{bt} M_X(at)$$

pf.

$$M_Y(t) = E[e^{tY}] = E[e^{t(aX+b)}]$$

$$= E[e^{atX} e^{bt}]$$

$$= e^{bt} E[e^{atX}]$$

$$= e^{bt} M_X(at)$$

Theorem:

Theorem:

If X and Y are r.v.s. and

$$M_X(t) = M_Y(t)$$

for t in some neighborhood of zero,
then $X \stackrel{d}{=} Y$.

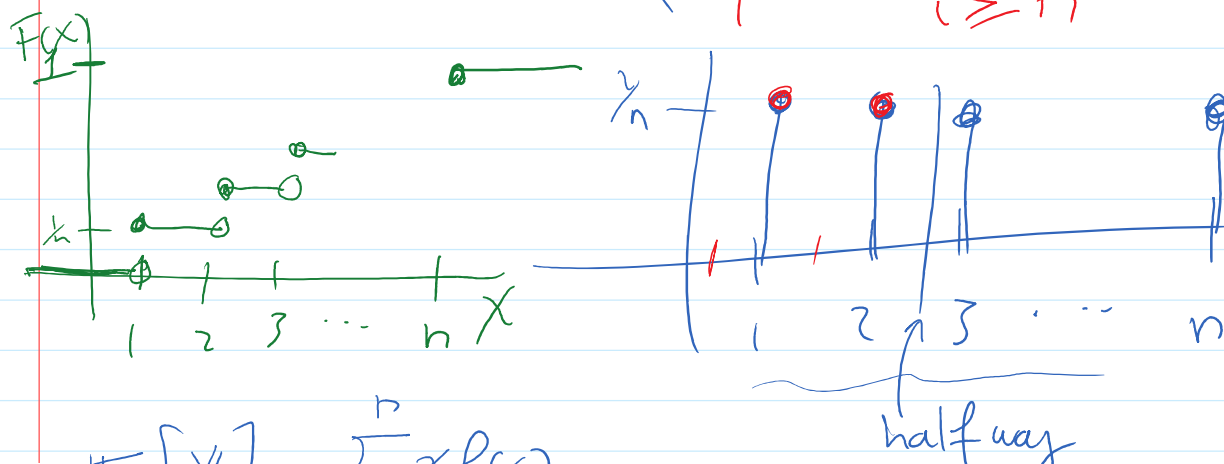
Discrete Uniform

$$X \sim U(\{1, \dots, n\})$$

means $f(x) = 1/n$ for $x=1, \dots, n$

CDF:

$$F(x) = \sum_{i \leq x} f(i) = \begin{cases} 0 & i < 1 \\ 1/n & 1 \leq i < 2 \\ 2/n & 2 \leq i < 3 \\ \vdots & \vdots \\ 1 & i \geq n \end{cases}$$



$$E[X] = \sum_{x=1}^n x f(x)$$

half way

$$= \sum_{x=1}^n x \frac{1}{n} = \frac{1}{n} \frac{n(n+1)}{2} = \frac{n+1}{2}$$

$$E[X^2] = \sum_{x=1}^n x^2 \frac{1}{n} = \frac{1}{n} \frac{(n+1)(2n+1)}{6}$$

$$\text{Var}(X) = E[X^2] - E[X]^2 = \dots = \boxed{\frac{(N+1)(N-1)}{12}}$$