

Defn: Bivariate Expectation

If (X, Y) and $g: \mathbb{R}^2 \rightarrow \mathbb{R}$

then

$$E[g(X, Y)] = \begin{cases} \sum_x \sum_y g(x, y) f(x, y) \\ \int_{\mathbb{R}^2} g(x, y) f(x, y) dx dy \end{cases}$$

recall: univariate case

Law of Unconscious Stat. $E[g(X)] = \begin{cases} \sum_x g(x) f(x) \\ \int g(x) f(x) dx \end{cases}$

Ex. Let (X, Y) where

support $\{(2, 1), (2, 2), \dots, (2, 6), (3, 1), \dots, (3, 6), \dots, (6, 1), \dots, (6, 6)\} \leftarrow X = \text{sum of two dice rolls}$

$\{(0, 1), (1, 0), \dots, (5, 0)\} \leftarrow Y = \text{abs. value of difference between the rolls}$

$$\begin{array}{lll} \text{roll 1: } 1 & \Rightarrow X = 2 & f(2, 0) = 1/36 \\ \text{roll 2: } 1 & \Rightarrow Y = |1 - 1| = 0 & \end{array}$$

$$\begin{array}{lll} \text{roll 1: } 5 & \Rightarrow X = 6 & f(6, 4) = 2/36 \\ \text{roll 2: } 1 & \Rightarrow Y = 4 & \end{array}$$

$(5,1)$, or $(1,5)$

$$f(3,0) = 0$$

		X										
		2	3	4	5	6	7	8	9	10	11	12
Y		0	$\frac{1}{36}$									
y		1	$\frac{2}{36}$									
2			$\frac{2}{36}$									
3				$\frac{2}{36}$								
4					$\frac{2}{36}$							
5						$\frac{2}{36}$						

Table $f(x,y)$

$$\mathbb{E}[XY] = \sum_{x=2}^{12} \sum_{y=0}^5 xy f(x,y) = \boxed{\frac{245}{18}}$$

Leave for you to fill in

$$= (2)(0) \underbrace{f(2,0)}_{\frac{1}{36}} + (2)(1) \underbrace{f(2,1)}_0 \dots$$

Theorem: \mathbb{E} linear

If $g_1: \mathbb{R}^2 \rightarrow \mathbb{R}$ and $g_2: \mathbb{R}^2 \rightarrow \mathbb{R}$

then for any $a, b \in \mathbb{R}$ we have

then for any $a, b \in \mathbb{R}$ we have

$$\rightarrow E[a g_1(X, Y) + b g_2(X, Y)] \\ = a E[g_1(X, Y)] + b E[g_2(X, Y)]$$

Pf. As in univariate case this comes from linearity of addition (discrete) and integration (cts)

$$\iint (a g_1(x, y) + b g_2(x, y)) f(x, y) dx dy \\ = a \iint g_1(x, y) f(x, y) dx dy \\ + b \iint g_2(x, y) f(x, y) dx dy$$

Defn: Covariance

We define the covariance between X, Y as

defn $\boxed{\text{Cov}(X, Y) = E[(X - E[X])(Y - E[Y])]}$

$$\begin{aligned} \mu_X &= E[X] \in \mathbb{R} \\ \mu_Y &= E[Y] \in \mathbb{R} \end{aligned}$$

$$\Rightarrow = E[(X - \mu_X)(Y - \mu_Y)]$$

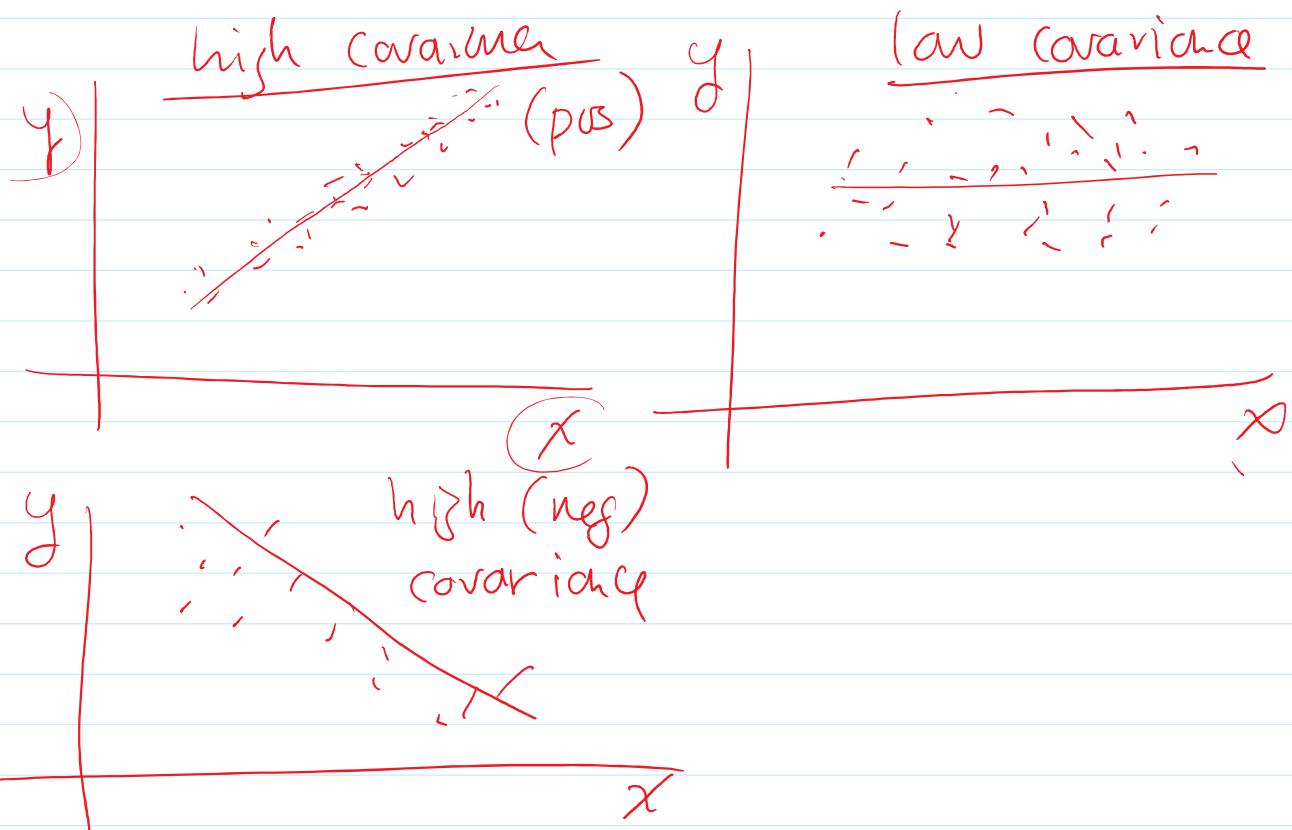
$$\gamma = \mathbb{E}[(X - \mu_X)(Y - \mu_Y)]$$

$\underbrace{\mathbb{E}[g(X, Y)]}$ where $g(X, Y) = (X - \mu_X)(Y - \mu_Y)$

Facts:

recall: $\text{Var}(X) = \mathbb{E}[(X - \mathbb{E}[X])^2]$

so $\text{Var}(X) = \text{Cov}(X, X)$



Defn: Correlation

We re-scale covariance to be in $[-1, 1]$
and define:

and define:

$$\text{Cor}(X, Y) = \frac{\text{Cov}(X, Y)}{\sqrt{\text{Var}(X) \text{Var}(Y)}}$$
$$= \frac{\text{Cov}(X, Y)}{\text{sd}(X) \text{sd}(Y)}$$
$$\text{sd}(X) = \sqrt{\text{Var}(X)}$$
$$\text{sd}(Y) = \sqrt{\text{Var}(Y)}$$

Theorem: If $a, b \in \mathbb{R}$ then

$$\text{Var}(aX + bY) = a^2 \text{Var}(X) + b^2 \text{Var}(Y)$$
$$+ 2ab \text{Cov}(X, Y).$$

different than \mathbb{E} :

$$\mathbb{E}[aX + bY] = a\mathbb{E}[X] + b\mathbb{E}[Y]$$

Pf. $Z = aX + bY$

$$\text{Var}(Z) = \mathbb{E}[(Z - \mathbb{E}[Z])^2]$$

$$\mathbb{E}[Z] = aX + bY - a\mathbb{E}[X] + b\mathbb{E}[Y]$$

$$= a(X - \mathbb{E}[X]) + b(Y - \mathbb{E}[Y]) \leftarrow$$

$$(Z - \mathbb{E}[Z])^2 = a^2(X - \mathbb{E}[X])^2 + b^2(Y - \mathbb{E}[Y])^2$$

$$(Z - E[Z])^2 = a^2(X - E[X])^2 + b^2(Y - E[Y])^2 + 2ab(X - E[X])(Y - E[Y])$$

$\text{Var}(aX + bY)$

$$\begin{aligned} \textcircled{3} \quad \text{Var}(Z) &= E[(Z - E[Z])^2] \\ &= a^2 E[(X - E[X])^2] + b^2 E[(Y - E[Y])^2] \\ &\quad + 2ab E[(X - E[X])(Y - E[Y])] \\ &= a^2 \text{Var}(X) + b^2 \text{Var}(Y) + 2ab \text{Cov}(X, Y). \end{aligned}$$

Theorem:

If $a, b, c, d \in \mathbb{R}$ then

$$\rightarrow \boxed{\text{Cov}(a+bX, c+dY) = bd \text{Cov}(X, Y)}$$

recall: $\text{Var}(a+bX) = b^2 \text{Var}(X)$ ↗
 $\text{Cov}(a+bX, a+bX) = \text{Var}(a+bX)$ ↗

pf. $\text{Cov}(a+bX, Z)$

$$= E[(a+bX - E(a+bX))(Z - E[Z])]$$

$$= E[(a+bX - (a+bE[X]))(Z - E[Z])]$$

$$= b \mathbb{E}[(X - \mathbb{E}[X])(Z - \mathbb{E}[Z])]$$

$= b \operatorname{Cov}(X, Z)$ (finish using symmetry)

Theorem: $a, b, c, d \in \mathbb{R}$ then

$$\text{Cor}(a+bX, c+dY)$$

$$= \text{Sign}(b) \text{Sign}(d) \operatorname{Cor}(\mathcal{X}, \mathcal{Y})$$

$$\text{Sign}(b) = \begin{cases} 1 & \text{if } b > 0 \\ -1 & \text{if } b < 0 \\ 0 & \text{if } b = 0 \end{cases}$$

$$\text{Alt. } |\operatorname{Cor}(a + bX, cfdY)| = |\operatorname{Cor}(X, Y)|$$

$$\underline{\text{Ej.}} \quad \text{Cor}(X, -Y) = -\text{Cor}(X, Y)$$

$$\begin{aligned}
 \text{Pf. } \text{Cov}(a+bX, c+dY) &= \frac{\text{Cov}(a+bX, c+dY)}{\sqrt{\text{Var}(a+bX) \text{Var}(c+dY)}} \\
 &= \frac{bd \text{Cov}(X, Y)}{\sqrt{b^2 \text{Var}(X) d^2 \text{Var}(Y)}} \\
 &= \frac{b}{|b|} \frac{d}{|d|} \frac{\text{Cov}(X, Y)}{\sqrt{\text{Var}(X) \text{Var}(Y)}}
 \end{aligned}$$

Fact:-

$$\underline{b} = \text{sign}(b)$$

$$\frac{b}{|b|} = \text{Sign}(b)$$

for $b \neq 0$

$$= \text{Sign}(b) \text{Sign}(d) \text{Cor}(\tilde{X}, Y_1)$$

Theorem: $|\text{Cor}(X, Y)| \leq 1$

pf. $\tilde{X} = \frac{X - \mu_X}{\text{sd}(X)}$ $\text{Var}(\tilde{X}) = 1$
 $E(\tilde{X}) = 0$

$$\tilde{Y} = \frac{Y - \mu_Y}{\text{sd}(Y)} \quad \text{linear transformations}$$

$$|\text{Cor}(X, Y)| = |\text{Cor}(\tilde{X}, \tilde{Y})| = \rho \geq 0$$

$$\text{Var}(\tilde{X} + \tilde{Y}) = \text{Var}(\tilde{X}) + \text{Var}(\tilde{Y}) + 2 \text{Cov}(\tilde{X}, \tilde{Y})$$

$$\text{Var}(\tilde{X} - \tilde{Y}) = \underbrace{\text{Var}(\tilde{X})}_{1} + \underbrace{\text{Var}(\tilde{Y})}_{1} - 2 \text{Cov}(\tilde{X}, \tilde{Y})$$

$$\text{Cov}(\tilde{X}, \tilde{Y}) = \text{Cov}(\tilde{X}, \tilde{Y}) = \rho$$

$$0 \leq \text{Var}(\tilde{X} \pm \tilde{Y}) = 1 + 1 \pm 2\rho = 2(1 \pm \rho)$$

$$2(1 \pm \rho) \geq 0 \quad \rho > 0$$

or $1 \pm \rho \geq 0$

$$\text{or } |1-p| \leq 1$$

Theorem: Short-cut formula

recall: $\text{Var}(X) = E[X^2] - [E[X]]^2$

Now: $\boxed{\text{Cov}(X, Y) = E[XY] - E[X]E[Y].}$

Pf. $\text{Cov}(X, Y) = E[(X - E[X])(Y - E[Y])]$

$$= E\left[\cancel{XY} - X\cancel{E[Y]} - \cancel{Y}\cancel{E[X]} + E[X]E[Y]\right]$$

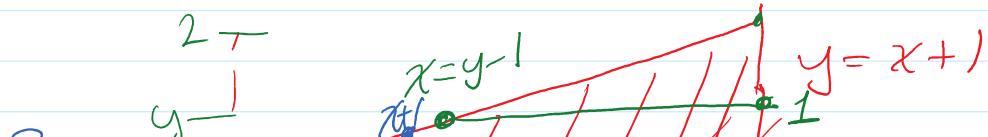
$$= \underline{E[XY]} - \underline{E[X]E[Y]} - \underline{E[Y]E[X]} + \cancel{E[X]E[Y]}$$

Ex. (X, Y) has pdf of

$$\boxed{f(x, y) = 1}$$

for $0 < x < 1$
and $x < y < x+1$

Calculate

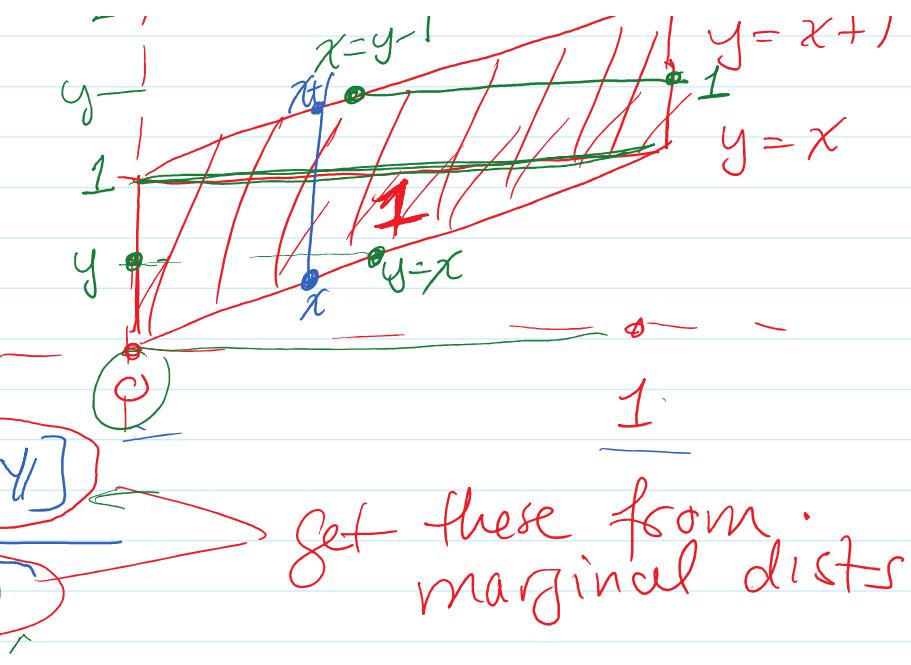


(calculate

$$\text{Cor}(X, Y) = ?$$

$$= \frac{\text{Cov}(X, Y)}{\sqrt{\text{Var}(X) \text{Var}(Y)}}$$

$$= \frac{E[XY] - E[X]E[Y]}{\sqrt{\text{Var}(X) \text{Var}(Y)}}$$



get these from
marginal dists

Marginal of X :

$$f(x) = \int f_{XY}(y) dy = \int_x^{x+1} 1 dy = y \Big|_x^{x+1} = (x+1) - x = 1$$

for $0 < x < 1$

i.e. $X \sim U(0, 1)$

$$\boxed{E[X] = \frac{1}{2} \quad \text{Var}(X) = \frac{1}{12}}$$

Marginal of Y

$$f(y) = \int f_{XY}(x) dx =$$

$$\begin{cases} \int_0^y 1 dx = y & 0 < y < 1 \\ \int_{y-1}^1 1 dx = 2 - y & 1 \leq y < 2 \end{cases}$$

- (y , $0 < y < 1$)

$$= \begin{cases} y & 0 < y < 1 \\ 2-y & 1 \leq y < 2 \end{cases}$$

Exercise: $E[Y] = 1$ $\text{Var}(X) = 1/6$

Last part: $E[XY]$

$$E[XY] = \iint_{\Omega} xy f(x,y) dy dx = \underset{\substack{(x+1) \\ \text{exercise}}}{\text{Calc III}} = 7/12$$

All together:

$$\text{Cor}(X,Y) = \frac{7/12 - (1/2)(1)}{\sqrt{1/2 \cdot 1/6}} = \text{some number } r > 0$$

Conditional Distributions

Recall: $P(A|B) = \frac{P(A \cap B)}{P(B)}$

If X, Y discrete let $A = \{X=x\}, B = \{Y=y\}$

$$\rightarrow P(X=x|Y=y) = \frac{P(X=x, Y=y)}{P(Y=y)}$$

$$\rightarrow P(X=x | Y=y) = \frac{\text{Probability of } X=x \text{ and } Y=y}{P(Y=y)}$$

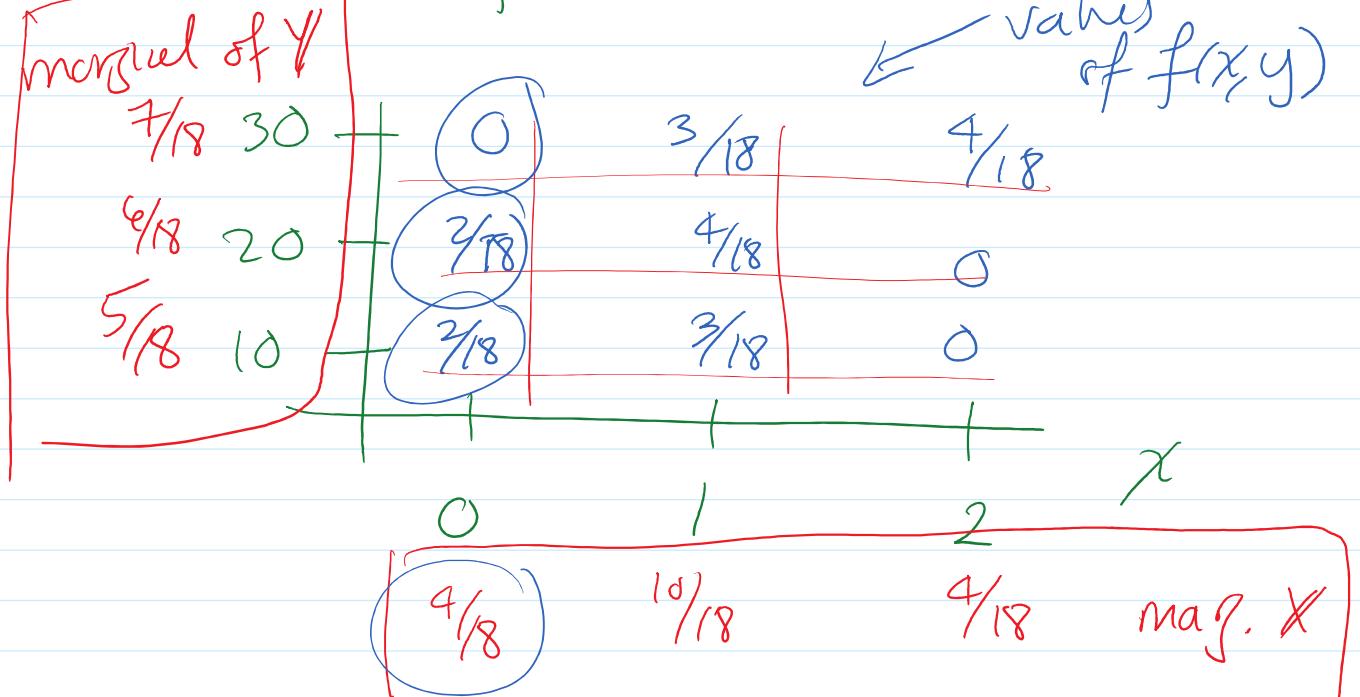
Defn: Conditional PMF

The conditional pmf of X given $Y=y$ is defined as

$$f_{X|Y=y}(x) = f(x|y) = \frac{f(x,y)}{f_y(y)}$$

Consider $Z = "X|Y=y"$ as a univariate random variable w/ pmf $f(x|y)$.

Ex. Joint PMF of X and Y



(7/8)

1/8

7/18

may. *

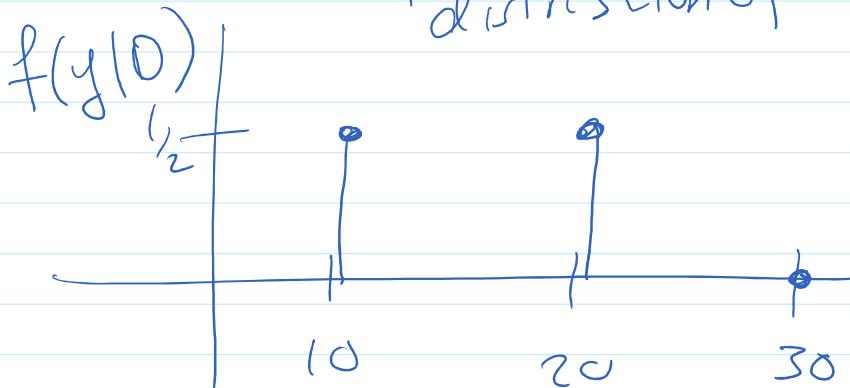
Distr of $Y/X = \chi$

$$f(y|X=0) = \frac{f(0,y)}{f_X(0)}$$

$$\begin{array}{c} f(10|X=0) \quad f(20|X=0) \quad f(30|X=0) \\ \hline \frac{f(0,10)}{f_X(0)} = \frac{3/8}{4/18} = \frac{3/8}{1/2} = 1 \quad \frac{f(0,30)}{f_X(0)} = 0 \end{array}$$

$$= 1/2$$

↑
distribution of $Y/X = 0$



Cts Case?

$$P(X=x|Y=y) = \frac{P(X=x, Y=y)}{P(Y=y)}$$

problem! $P(Y=y) = 0$
in cts case.

Defn: Conditional Distr (cts case)

Weku: Conditionale dist (cts case)

If X, Y continuous then the conditional pdf of X given $Y = y$ is

$$f(x|y) = \frac{f(x,y)}{f_y(y)}$$

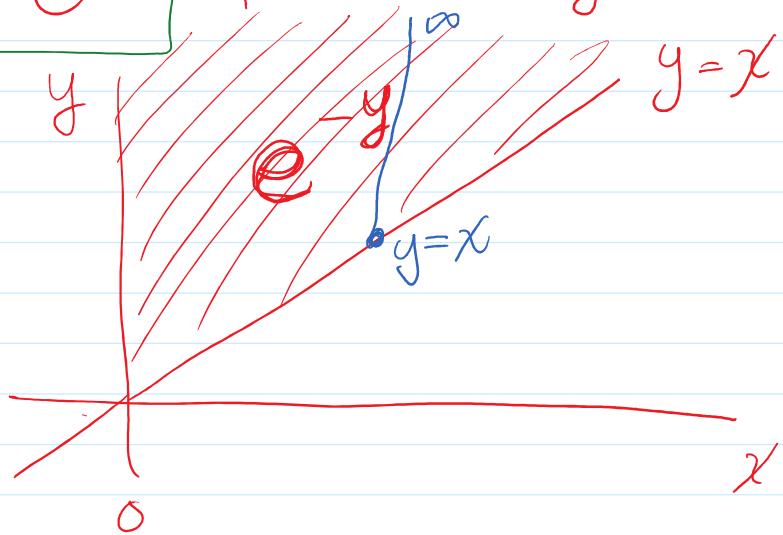
joint pdf ← marginal pdf ←

Ex. X, Y has joint pdf of

$$f(x,y) = e^{-y} \quad \text{for } 0 < x < y$$

Want: $f(y|x)$

i.e. pdf of $Y|X=x$



Formula:

$$f(y|x) = \frac{f(x,y)}{f_x(x)}$$

Marginal of X : $f_X(x) = \int f(x,y) dy = \int_x^{\infty} e^{-y} dy = e^{-x}$

$$\text{So } f(y|x) = \frac{e^{-y}}{e^{-x}} = e^{x-y} \quad \text{for } 0 < x < y$$

looks like an exponential dist.
Called: shifted exponential:

