Appendix A

Summary of Basic Formulae

Summary

Two sets of tables are provided for reference. The first records the definition, and the first two moments of the most common probability distributions used in this volume. The second records the basic elements of standard Bayesian inference processes for a number of special cases. In particular, it records the appropriate likelihood function, the sufficient statistics, the conjugate prior and corresponding posterior and predictive distributions, the reference prior and corresponding reference posterior and predictive distributions.

A.1 PROBABILITY DISTRIBUTIONS

The first section of this Appendix consists of a set of tables which record the notation, parameter range, variable range, definition, and first two moments of the probability distributions (discrete and continuous, univariate and multivariate) used in this volume.

Univariate Discrete Distributions

$Br(x \mid \theta)$ Bernoulli $(p. 115)$	
$0 < \theta < 1$	x = 0, 1
$p(x) = \theta^x (1 - \theta)^{1 - x}$	
$E[x] = \theta$	$V[x] = \theta(1 - \theta)$
$Bi(x \mid \theta, n)$ Binomial $(p. 115)$	
$0 < \theta < 1, n = 1, 2, \dots$	$x = 0, 1, \dots, n$
$p(x) = \binom{n}{x} \theta^x (1 - \theta)^{n-x}$	
$E[x] = n\theta$	$V[x] = n\theta(1 - \theta)$
$Bb(x \mid \alpha, \beta, n)$ Binomial-Beta (p. 117)	
$\alpha > 0, \beta > 0, n = 1, 2, \dots$	$x = 0, 1, \dots, n$
$p(x) = c \binom{n}{x} \Gamma(\alpha + x) \Gamma(\beta + n - x)$	$c = \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)\Gamma(\alpha + \beta + n)}$
$E[x] = n \frac{\alpha}{\alpha + \beta}$	$V[x] = \frac{n\alpha\beta}{(\alpha+\beta)^2} \frac{(\alpha+\beta+n)}{(\alpha+\beta+1)}$
$ Hy(x \mid N, M, n) \textit{Hypergeometric} \ \ (p. 115) $	
N = 1, 2, M = 1, 2, n = 1,, N + M	$x = a, a + 1, \dots, b$ $a = \max(0, n - M)$ $b = \min(n, N)$
$p(x) = c \binom{N}{x} \binom{M}{n-x}$	$c = \binom{N+M}{n}^{-1}$
$E[x] = n \frac{N}{N+M}$	$V[x] = \frac{nNM}{(N+M)^2} \frac{N+M-n}{N+M-1}$

Univariate Discrete Distributions (continued)

$Nb(x | \theta, r)$ Negative-Binomial (p. 116)

$$0 < \theta < 1, r = 1, 2, \dots$$

$$x = 0, 1, 2, \dots$$

$$p(x) = c \binom{r+x-1}{r-1} (1-\theta)^x$$

$$c=\theta^r$$

$$E[x] = r\theta$$

$$V[x] = r \frac{1 - \theta}{\theta^2}$$

$Nbb(x \mid \alpha, \beta, r)$ Negative-Binomial-Beta (p. 118)

$$\alpha > 0, \ \beta > 0, \ r = 1, 2 \dots$$

$$x = 0, 1, 2, \dots$$

$$p(x) = c \begin{pmatrix} r+x-1 \\ r-1 \end{pmatrix} \frac{\Gamma(\beta+x)}{\Gamma(\alpha+\beta+r+x)} \quad c = \frac{\Gamma(\alpha+\beta)\Gamma(\alpha+r)}{\Gamma(\alpha)\Gamma(\beta)}$$

$$c = \frac{\Gamma(\alpha + \beta)\Gamma(\alpha + r)}{\Gamma(\alpha)\Gamma(\beta)}$$

$$E[x] = \frac{r\beta}{\alpha - 1}$$

$$E[x] = \frac{r\beta}{\alpha - 1} \qquad V[x] = \frac{r\beta}{(\alpha - 1)} \left[\frac{\alpha + \beta + r - 1}{(\alpha - 2)} + \frac{r\beta}{(\alpha - 1)(\alpha - 2)} \right]$$

$Pn(x \mid \lambda)$ Poisson (p. 116)

$$\lambda > 0$$

$$x = 0, 1, 2, \dots$$

$$p(x) = c \frac{\lambda^x}{x!}$$

$$c=e^{-\lambda}$$

$$E[x] = \lambda$$

$$V[x] = \lambda$$

$Pg(x \mid \alpha, \beta, n)$ Poisson-Gamma (p. 119)

$$\alpha > 0, \ \beta > 0, \ \nu > 0$$

$$x = 0, 1, 2, \dots$$

$$p(x) = c \frac{\Gamma(\alpha + x)}{x!} \frac{\nu^x}{(\beta + \nu)^{\alpha + x}}$$

$$c = \frac{\beta^{\alpha}}{\Gamma(\alpha)}$$

$$E[x] = \nu \frac{\alpha}{\beta}$$

$$V[x] = \frac{\nu\alpha}{\beta} \left[1 + \frac{\nu}{\beta} \right]$$

Univariate Continuous Distributions

$Be(x \mid \alpha, \beta) \textit{Beta} \ (p. 116)$	
$\alpha > 0, \beta > 0$	0 < x < 1
$p(x) = c x^{\alpha - 1} (1 - x)^{\beta - 1}$	$c = \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)}$
$E[x] = \frac{\alpha}{\alpha + \beta}$	$V[x] = \frac{\alpha\beta}{(\alpha+\beta)^2(\alpha+\beta+1)}$
$Un(x \mid a, b) Uniform \ (p. 117)$	
b > a	a < x < b
p(x) = c	$c = (b - a)^{-1}$
$E[x] = \frac{1}{2}(a+b)$	$V[x] = \frac{1}{12}(b - a)^2$
$Ga(x \mid \alpha, \beta)$ Gamma (p. 118)	
$\alpha > 0, \beta > 0$	x > 0
$p(x) = c \ x^{\alpha - 1} e^{-\beta x}$	$c = \frac{\beta^{\alpha}}{\Gamma(\alpha)}$
$E[x] = \alpha \beta^{-1}$	$V[x] = \alpha \beta^{-2}$
$Ex(x \mid \theta)$ Exponential (p. 118)	
$\theta > 0$	x > 0
$p(x) = c e^{-\theta x}$	$c = \theta$
$E[x] = 1/\theta$	$V[x] = 1/\theta^2$
$Gg(x \mid \alpha, \beta, n)$ Gamma-Gamma (p. 120)	
$\alpha > 0, \beta > 0, n > 0$	x > 0
$p(x) = c \frac{x^{n-1}}{(\beta + x)^{\alpha + n}}$	$c = \frac{\beta^{\alpha}}{\Gamma(\alpha)} \frac{\Gamma(\alpha + n)}{\Gamma(n)}$
$E[x] = n \frac{\beta}{\alpha - 1}$	$V[x] = \frac{\beta^{2}(n^{2} + n(\alpha - 1))}{(\alpha - 1)^{2}(\alpha - 2)}$

A.1 Probability Distributions

431

Univariate Continuous Distributions (continued)

$\chi^2(x \mid \nu) = \chi^2_{\nu}$ Chi-squared (p.	120)
$\nu > 0$	x > 0
$p(x) = c x^{(\nu/2)-1} e^{-x/2}$	$c = \frac{(1/2)^{\nu/2}}{\Gamma(\nu/2)}$
$E[x] = \nu$	$V[x] = 2\nu$
$\frac{1}{\chi^2(x \mid \nu, \lambda)}$ Non-central Chi-squ	

$$\begin{split} \nu > 0, \lambda > 0 & x > 0 \\ p(x) = \sum_{i=0}^{\infty} \Pr\left(i \mid \frac{\lambda}{2}\right) \chi^2(x \mid \nu + 2i) & \\ E[x] = \nu + \lambda & V[x] = 2(\nu + 2\lambda) \end{split}$$

$Ig(x \mid \alpha, \beta)$ Inverted-Gamma (p. 119)

$\alpha > 0, \beta > 0$	x > 0
$p(x) = c x^{-(\alpha+1)} e^{-\beta/x}$	$c = \frac{\beta^{\alpha}}{\Gamma(\alpha)}$
$E[x] = \frac{\beta}{\alpha - 1}$	$V[x] = \frac{\beta^2}{(\alpha - 1)^2(\alpha - 2)}$

$$\begin{array}{ll} \chi^{-1}(x\,|\,\nu) & \textit{Inverted-Chi-squared} \;\; (p.\,119) \\ \hline \nu > 0 & x > 0 \\ p(x) = c \;\; x^{-(\nu/2+1)} e^{-1/2x^2} & c = \frac{(1/2)^{\nu/2}}{\Gamma(\nu/2)} \\ E[x] = \frac{1}{\nu - 2} & V[x] = \frac{2}{(\nu - 2)^2(\nu - 4)} \end{array}$$

$Ga^{-1/2}(x \mid \alpha, \beta)$ Square-root Inverted-Gamma (p. 119)

$$\begin{split} \alpha > 0, \, \beta > 0 & x > 0 \\ p(x) = c \; x^{-(2\alpha+1)} e^{-\beta/x^2} & c = \frac{2\beta^{\alpha}}{\Gamma(\alpha)} \\ E[x] = \frac{\sqrt{\beta}\Gamma(\alpha - 1/2)}{\Gamma(\alpha)} & V[x] = \frac{\beta}{\alpha - 1} - E[x]^2 \end{split}$$

Univariate Continuous Distributions (continued)

$Pa(x \mid \alpha, \beta)$ Pareto $(p. 120)$	
$\alpha > 0, \beta > 0$ $p(x) = c \ x^{-(\alpha+1)}$	$\beta \le x < +\infty$ $c = \alpha \beta^{\alpha}$
$E[x] = \frac{\beta\alpha}{\alpha - 1} \; , \alpha > 1$	$V[x] = \frac{\beta^2 \alpha}{(\alpha - 1)^2 (\alpha - 2)} , \alpha > 2$
$Ip(x \mid \alpha, \beta)$ Inverted-Pareto (p. 120)	
$\alpha > 0, \beta > 0$ $p(x) = c x^{\alpha - 1}$ $E[x] = \beta^{-1} \alpha (\alpha + 1)^{-1}$	$0 < x < \beta^{-1}$ $c = \alpha \beta^{\alpha}$ $V[x] = \beta^{-2} \alpha (\alpha + 1)^{-2} (\alpha + 2)^{-1}$
$N(x \mid \mu, \lambda)$ Normal (p. 121)	
$-\infty < \mu < +\infty, \lambda > 0$ $p(x) = c \exp\left\{-\frac{1}{2}\lambda(x - \mu)^2\right\}$ $E[x] = \mu$	$-\infty < x < +\infty$ $c = \lambda^{1/2} (2\pi)^{-1/2}$ $V[x] = \lambda^{-1}$
$St(x \mid \mu, \lambda, \alpha) Student \ t \ (p. 122)$	
$-\infty < \mu < +\infty, \lambda > 0, \alpha > 0$ $p(x) = c \left[1 + \alpha^{-1} \lambda (x - \mu)^2 \right]^{-(\alpha + 1)/2}$ $E[x] = \mu$	$-\infty < x < +\infty$ $c = \frac{\Gamma\left(\frac{1}{2}(\alpha+1)\right)}{\Gamma(\frac{1}{2}\alpha)} \left(\frac{\lambda}{\alpha\pi}\right)^{1/2}$ $V[x] = \lambda^{-1}\alpha(\alpha-2)^{-1}$
$F(x \mid \alpha, \beta) = F_{\alpha,\beta}$ Snedecor F (p. 123)	
$\alpha > 0, \beta > 0$ $p(x) = c \frac{x^{\alpha/2 - 1}}{(\beta + \alpha x)^{(\alpha + \beta)/2}}$ $E[x] = \frac{\beta}{\beta - 2}, \beta > 2$	$x > 0$ $c = \frac{\Gamma(\frac{1}{2}(\alpha + \beta)) \alpha^{\alpha/2} \beta^{\beta/2}}{\Gamma(\frac{1}{2}\alpha)\Gamma(\frac{1}{2}\beta)}$ $V[x] = \frac{2\beta^2(\alpha + \beta - 2)}{\alpha(\beta - 2)^2(\beta - 4)}, \beta > 4$

Univariate Continuous Distributions (continued)

$$Lo(x \mid \alpha, \beta) \quad Logistic \quad (p. 122)$$

$$-\infty < \alpha < +\infty, \beta > 0 \qquad -\infty < x < +\infty$$

$$p(x) = \beta^{-1} \exp\left\{-\beta^{-1}(x - \alpha)\right\} \left[1 + \exp\left\{-\beta^{-1}(x - \alpha)\right\}\right]^{-2}$$

$$E[x] = \alpha \qquad V[x] = \beta^2 \pi^2 / 3$$

Multivariate Discrete Distributions

$\begin{aligned} & \underbrace{\mathsf{Mu}_{k}(\boldsymbol{x} \,|\, \boldsymbol{\theta}, n) \quad \mathsf{Multinomial} \ (p.\, 133)}_{\boldsymbol{\theta} = (\theta_{1}, \dots, \theta_{k}) } & \boldsymbol{x} = (x_{1}, \dots, x_{k}) \\ & 0 < \theta_{i} < 1, \quad \sum_{\ell=1}^{k} \theta_{\ell} \le 1 & \sum_{\ell=1}^{k} x_{\ell} \le n \\ & n = 1, 2, \dots & x_{i} = 0, 1, 2, \dots \\ & p(\boldsymbol{x}) = \frac{n!}{\prod_{\ell=1}^{k+1} x_{\ell}!} \prod_{\ell=1}^{k+1} \theta^{x_{\ell}}, \quad \theta_{k+1} = 1 - \sum_{\ell=1}^{k} \theta_{\ell}, \quad x_{k+1} = n - \sum_{\ell=1}^{k} x_{\ell} \\ & E[x_{i}] = n\theta_{i} & V[x_{i}] = n\theta_{i} (1 - \theta_{i}) & C[x_{i}, x_{j}] = -n\theta_{i}\theta_{j} \end{aligned}$

$Md_k(\boldsymbol{x} \mid \boldsymbol{\theta}, n)$ Multinomial-Dirichlet (p. 135)

$$\begin{array}{ll} \pmb{\alpha} = (\alpha_{1}, \dots, \alpha_{k+1}) & \pmb{x} = (x_{1}, \dots, x_{k}) \\ \alpha_{i} > 0 & x_{i} = 0, 1, 2, \dots \\ n = 1, 2, \dots & \sum_{\ell=1}^{n} x_{\ell} \leq n \\ p(\pmb{x}) = c \prod_{\ell=1}^{k+1} \frac{\alpha_{\ell}^{[x_{\ell}]}}{x_{\ell}!} & c = \frac{n!}{\left(\sum_{\ell=1}^{k+1} \alpha_{\ell}\right)^{[n]}} \\ \alpha^{[s]} = \prod_{\ell=1}^{s} (\alpha + \ell - 1) & x_{k+1} = n - \sum_{\ell=1}^{k} x_{\ell} \\ E[x_{i}] = np_{i} & V[x_{i}] = \frac{n + \sum_{\ell=1}^{k+1} \alpha_{\ell}}{1 + \sum_{\ell=1}^{k+1} \alpha_{\ell}} np_{i}(1 - p_{i}) \\ p_{i} = \frac{\alpha_{i}}{\sum_{\ell=1}^{k+1} \alpha_{\ell}} & C[x_{i}, x_{j}] = -\frac{n + \sum_{\ell=1}^{k+1} \alpha_{\ell}}{1 + \sum_{\ell=1}^{k+1} \alpha_{\ell}} np_{i}p_{j} \end{array}$$

Multivariate Continuous Distributions

$Di_k(\boldsymbol{x} \mid \boldsymbol{\alpha})$ Dirichlet (p. 134)

$$\alpha = (\alpha_1, \dots, \alpha_{k+1}) \qquad x = (x_1, \dots, x_k)$$

$$\alpha_i > 0 \qquad 0 < x_i < 1, \quad \sum_{\ell=1}^k x_\ell \le 1$$

$$p(x) = c \left(1 - \sum_{\ell=1}^k x_\ell\right)^{\alpha_{k+1} - 1} \prod_{\ell=1}^k x_\ell^{\alpha_{\ell} - 1} \qquad c = \frac{\Gamma(\sum_{\ell=1}^{k+1} \alpha_{\ell})}{\prod_{\ell=1}^{k+1} \Gamma(\alpha_{\ell})}$$

$$E[x_i] = \frac{\alpha_i}{\sum_{\ell=1}^{k+1} \alpha_{\ell}} \quad V[x_i] = \frac{E[x_i](1 - E[x_i])}{1 + \sum_{\ell=1}^{k+1} \alpha_{\ell}} \quad C[x_i, x_j] = \frac{-E[x_i]E[x_j]}{1 + \sum_{\ell=1}^{k+1} \alpha_{\ell}}$$

$Ng(x, y | \mu, \lambda, \alpha, \beta)$ Normal-Gamma (p. 136)

$$\begin{split} &\mu\in\Re,\;\lambda>0,\;\alpha>0,\;\beta>0, &x\in\Re,\;y>0\\ &p(x,y)=\mathrm{N}(x\,|\,\mu,\lambda y)\;\mathrm{Ga}(y\,|\,\alpha,\beta)\\ &E[x]=\mu \qquad E[y]=\alpha\beta^{-1} \qquad V[x]=\beta\lambda^{-1}(\alpha-1)^{-1} \qquad V[y]=\alpha\beta^{-2}\\ &p(x)=\mathrm{St}(x\,|\,\mu,\alpha\beta^{-1}\lambda,2\alpha) \end{split}$$

$N_k(\boldsymbol{x} \mid \boldsymbol{\mu}, \boldsymbol{\lambda})$ Multivariate Normal (p. 136)

$$oldsymbol{\mu} = (\mu_1, \dots, \mu_k) \in \Re^k$$
 $oldsymbol{x} = (x_1, \dots, x_k) \in \Re^k$ $oldsymbol{\lambda}$ symmetric positive-definite $p(oldsymbol{x}) = c \, \exp\left\{-\frac{1}{2}(oldsymbol{x} - oldsymbol{\mu})^t oldsymbol{\lambda}(oldsymbol{x} - oldsymbol{\mu})\right\}$ $c = |oldsymbol{\lambda}|^{1/2} (2\pi)^{-k/2}$ $E[oldsymbol{x}] = oldsymbol{\mu}$ $V[oldsymbol{x}] = oldsymbol{\lambda}^{-1}$

$Pa_2(x, y \mid \alpha, \beta_0, \beta_1)$ Bilateral Pareto (p. 141)

$$(\beta_{0}, \beta_{1}) \in \Re^{2}, \ \beta_{0} < \beta_{1}, \ \alpha > 0 \qquad (x, y) \in \Re^{2}, \ x < \beta_{0}, \ y > \beta_{1}$$

$$p(x, y) = c \ (y - x)^{-(\alpha + 2)} \qquad c = \alpha(\alpha + 1)(\beta_{1} - \beta_{0})^{\alpha}$$

$$E[x] = \frac{\alpha\beta_{0} - \beta_{1}}{\alpha - 1} \quad E[y] = \frac{\alpha\beta_{1} - \beta_{0}}{\alpha - 1} \qquad V[x] = V[y] = \frac{\alpha(\beta_{1} - \beta_{0})^{2}}{(\alpha - 1)^{2}(\alpha - 2)}$$

Multivariate Continuous Distributions (continued)

$Ng_k(x, y \mu, \lambda, \alpha, \beta)$ Multivariate Normal-Gamma (p. 140)		
$-\infty < \mu_i < +\infty, \ \alpha > 0, \ \beta > 0$ $\lambda \text{ symmetric positive-definite}$	$(\boldsymbol{x}, y) = (x_1, \dots, x_k, y)$ $-\infty < x_i < \infty, y > 0$	
$p(\boldsymbol{x}, y) = N_k(\boldsymbol{x} \mid \boldsymbol{\mu}, \boldsymbol{\lambda} y) \text{ Ga}(y \mid \alpha, \beta)$	1)-10) -1	
$E[\mathbf{x}, y] = (\boldsymbol{\mu}, \ \alpha \beta^{-1}), \qquad V[\mathbf{x}] = (p(\mathbf{x}) = \operatorname{St}_k(\mathbf{x} \mid \boldsymbol{\mu}, \boldsymbol{\lambda} \alpha \beta^{-1}, 2\alpha)$	$p(y) = Ga(y \mid \alpha, \beta)$	
$\overline{\mathrm{Nw}_k(oldsymbol{x},oldsymbol{y} oldsymbol{\mu},\lambda,lpha,oldsymbol{eta})}$ Multivariate	Normal-Wishart (p. 140)	
$-\infty < \mu_i < +\infty, \ \lambda > 0$	$\boldsymbol{x}=(x_1,\ldots,x_k)$	
$2\alpha > k-1$	$-\infty < x_i < +\infty$	

$$2\alpha > k-1$$
 $-\infty < x_i < +\infty$ $m{y}$ symmetric non-singular $m{y}$ symmetric positive-definite $p(m{x}, m{y}) = \mathrm{N}_k(m{x} \mid m{\mu}, \lambda m{y}) \ \mathrm{Wi}_k(m{y} \mid \alpha, m{\beta})$ $E[m{x}, m{y}] = \{m{\mu}, \alpha m{\beta}^{-1}\}$ $V[m{x}] = (\alpha - 1)^{-1} m{\beta} \lambda^{-1}$

 $p(\boldsymbol{y}) = \operatorname{Wi}_k(\boldsymbol{y} \mid \alpha, \boldsymbol{\beta})$

 $\operatorname{St}_k(\boldsymbol{x} \mid \boldsymbol{\mu}, \boldsymbol{\lambda}, \alpha)$ Multivariate Student (p. 139)

 $p(\boldsymbol{x}) = \operatorname{St}_k(\boldsymbol{x} \mid \boldsymbol{\mu}, \lambda \alpha \boldsymbol{\beta}^{-1}, 2\alpha)$

$$\begin{aligned} &-\infty < \mu_i < +\infty, \ \alpha > 0 & \boldsymbol{x} = (x_1, \dots, x_k) \\ &\boldsymbol{\lambda} \text{ symmetric positive-definite} & -\infty < x_i < +\infty \end{aligned}$$

$$p(\boldsymbol{x}) = c \left[1 + \frac{1}{\alpha} (\boldsymbol{x} - \boldsymbol{\mu})^t \boldsymbol{\lambda} (\boldsymbol{x} - \boldsymbol{\mu}) \right]^{-(\alpha + k)/2} c = \frac{\Gamma\left(\frac{1}{2}(\alpha + k)\right)}{\Gamma(\frac{1}{2}\alpha)(\alpha\pi)^{k/2}} |\boldsymbol{\lambda}|^{1/2}$$

$$E[\boldsymbol{x}] = \boldsymbol{\mu}, \quad V[\boldsymbol{x}] = \boldsymbol{\lambda}^{-1} (\alpha - 2)^{-1} \alpha$$

 $\operatorname{Wi}_k(\boldsymbol{x} \mid \alpha, \boldsymbol{\beta})$ Wishart (p. 138)

$$\begin{aligned} &2\alpha>k-1 & &\pmb{x} \text{ symmetric positive-definite} \\ &\pmb{\beta} \text{ symmetric non-singular} \\ &p(\pmb{x})=c\,|\pmb{x}|^{\alpha-(k+1)/2}\exp\{-\operatorname{tr}(\pmb{\beta}\pmb{x})\} & &c=\frac{\pi^{-k(k-1)/4}|\pmb{\beta}|^\alpha}{\prod_{\ell=1}^k\Gamma(\frac{1}{2}(2\alpha+1-\ell))} \\ &E[\pmb{x}]=\alpha\pmb{\beta}^{-1}, \quad E[\pmb{x}^{-1}]=(\alpha-\frac{k+1}{2})^{-1}\pmb{\beta} \end{aligned}$$

A.2 INFERENTIAL PROCESSES

The second section of this Appendix records the basic elements of the Bayesian learning processes for many commonly used statistical models.

For each of these models, we provide, in separate sections of the table, the following: the sufficient statistic and its sampling distribution; the conjugate family, the conjugate prior predictives for a single observable and for the sufficient statistic, the conjugate posterior and the conjugate posterior predictive for a single observable.

When clearly defined, we also provide, in a final section, the reference prior and the corresponding reference posterior and posterior predictive for a single observable. In the case of uniparameter models this can always be done. We recall, however, from Section 5.2.4 that, in multiparameter problems, the reference prior is only defined *relative to an ordered parametrisation*. In the univariate normal model (Example 5.17), the reference prior for (μ, λ) happens to be the same as that for (λ, μ) , namely $\pi(\mu, \lambda) = \pi(\lambda, \mu) \propto \lambda^{-1}$, and we provide the corresponding reference posteriors for μ and λ , together with the reference predictive distribution for a future observation.

In the multinomial, multivariate normal and linear regression models, however, there are very many different reference priors, corresponding to different inference problems, and specified by different ordered parametrisations. These are not reproduced in this Appendix.

Bernoulli model

```
\begin{aligned} & \boldsymbol{z} = \{x_1, \dots, x_n\}, & x_i \in \{0, 1\} \\ & p(x_i \mid \boldsymbol{\theta}) = \operatorname{Br}(x_i \mid \boldsymbol{\theta}), & 0 < \boldsymbol{\theta} < 1 \end{aligned} \begin{aligned} & t(\boldsymbol{z}) = r = \sum_{i=1}^n x_i \\ & p(r \mid \boldsymbol{\theta}) = \operatorname{Bi}(r \mid \boldsymbol{\theta}, n) \end{aligned} \begin{aligned} & p(\boldsymbol{\theta}) = \operatorname{Be}(\boldsymbol{\theta} \mid \boldsymbol{\alpha}, \boldsymbol{\beta}) \\ & p(\boldsymbol{x}) = \operatorname{Bb}(r \mid \boldsymbol{\alpha}, \boldsymbol{\beta}, 1) \\ & p(r) = \operatorname{Bb}(r \mid \boldsymbol{\alpha}, \boldsymbol{\beta}, n) \\ & p(\boldsymbol{\theta} \mid \boldsymbol{z}) = \operatorname{Be}(\boldsymbol{\theta} \mid \boldsymbol{\alpha} + r, \boldsymbol{\beta} + n - r) \\ & p(\boldsymbol{x} \mid \boldsymbol{z}) = \operatorname{Bb}(\boldsymbol{x} \mid \boldsymbol{\alpha} + r, \boldsymbol{\beta} + n - r, 1) \end{aligned} \begin{aligned} & \pi(\boldsymbol{\theta}) = \operatorname{Be}(\boldsymbol{\theta} \mid \frac{1}{2}, \frac{1}{2}) \\ & \pi(\boldsymbol{\theta} \mid \boldsymbol{z}) = \operatorname{Be}(\boldsymbol{\theta} \mid \frac{1}{2} + r, \frac{1}{2} + n - r) \\ & \pi(\boldsymbol{x} \mid \boldsymbol{z}) = \operatorname{Bb}(\boldsymbol{x} \mid \frac{1}{2} + r, \frac{1}{2} + n - r, 1) \end{aligned}
```

Poisson Model

$$\begin{aligned} & \boldsymbol{z} = \{x_1, \dots, x_n\}, & x_i = 0, 1, 2, \dots \\ & p(x_i \mid \lambda) = \operatorname{Pn}(x_i \mid \lambda), & \lambda \geq 0 \end{aligned}$$

$$\begin{aligned} & \boldsymbol{t}(\boldsymbol{z}) = r = \sum_{i=1}^n x_i \\ & p(r \mid \lambda) = \operatorname{Pn}(r \mid n\lambda) \end{aligned}$$

$$\begin{aligned} & p(\lambda) &= \operatorname{Ga}(\lambda \mid \alpha, \beta) \\ & p(x) &= \operatorname{Pg}(x \mid \alpha, \beta, 1) \\ & p(r) &= \operatorname{Pg}(x \mid \alpha, \beta, n) \\ & p(\lambda \mid \boldsymbol{z}) &= \operatorname{Ga}(\lambda \mid \alpha + r, \beta + n) \\ & p(x \mid \boldsymbol{z}) &= \operatorname{Pg}(x \mid \alpha + r, \beta + n, 1) \end{aligned}$$

$$\pi(\lambda) &\propto \lambda^{-1/2}$$

$$\pi(\lambda \mid \boldsymbol{z}) &= \operatorname{Ga}(\lambda \mid r + \frac{1}{2}, n) \\ & \pi(x \mid \boldsymbol{z}) &= \operatorname{Pg}(x \mid r + \frac{1}{2}, n, 1) \end{aligned}$$

Negative-Binomial model

$$\begin{aligned} & \boldsymbol{z} = (x_1, \dots, x_n), & x_i = 0, 1, 2, \dots \\ & p(x_i \mid \boldsymbol{\theta}) = \operatorname{Nb}(x_i \mid \boldsymbol{\theta}, r), & 0 < \boldsymbol{\theta} < 1 \end{aligned}$$

$$\begin{aligned} & \boldsymbol{t}(\boldsymbol{z}) = s = \sum_{i=1}^n x_i \\ & p(s \mid \boldsymbol{\theta}) = \operatorname{Nb}(s \mid \boldsymbol{\theta}, nr) \end{aligned}$$

$$\begin{aligned} & p(\boldsymbol{\theta}) = \operatorname{Be}(\boldsymbol{\theta} \mid \boldsymbol{\alpha}, \boldsymbol{\beta}) \\ & p(x) = \operatorname{Nbb}(x \mid \boldsymbol{\alpha}, \boldsymbol{\beta}, r) \\ & p(s) = \operatorname{Nbb}(s \mid \boldsymbol{\alpha}, \boldsymbol{\beta}, nr) \\ & p(\boldsymbol{\theta} \mid \boldsymbol{z}) = \operatorname{Be}(\boldsymbol{\theta} \mid \boldsymbol{\alpha} + nr, \boldsymbol{\beta} + s) \\ & p(x \mid \boldsymbol{z}) = \operatorname{Nbb}(x \mid \boldsymbol{\alpha} + nr, \boldsymbol{\beta} + s, r) \end{aligned}$$

$$\begin{aligned} & \pi(\boldsymbol{\theta}) \propto \boldsymbol{\theta}^{-1}(1 - \boldsymbol{\theta})^{-1/2} \\ & \pi(\boldsymbol{\theta} \mid \boldsymbol{z}) = \operatorname{Be}(\boldsymbol{\theta} \mid nr, s + \frac{1}{2}) \\ & \pi(x \mid \boldsymbol{z}) = \operatorname{Nbb}(x \mid nr, s + \frac{1}{2}, r) \end{aligned}$$

Exponential Model

$$z = \{x_1, \dots, x_n\}, \quad 0 < x_i < \infty$$

$$p(x_i \mid \theta) = \operatorname{Ex}(x_i \mid \theta), \quad \theta > 0$$

$$t(z) = t = \sum_{i=1}^{n} x_i$$

$$p(t \mid \theta) = \operatorname{Ga}(t \mid n, \theta)$$

$$p(\theta) = \operatorname{Ga}(\theta \mid \alpha, \beta)$$

$$p(x) = \operatorname{Gg}(x \mid \alpha, \beta, 1)$$

$$p(t) = \operatorname{Gg}(t \mid \alpha, \beta, n)$$

$$p(\theta \mid z) = \operatorname{Ga}(\theta \mid \alpha + n, \beta + t)$$

$$p(x \mid z) = \operatorname{Gg}(x \mid \alpha + n, \beta + t, 1)$$

$$\pi(\theta) \propto \theta^{-1}$$

$$\pi(\theta \mid z) = \operatorname{Ga}(\theta \mid n, t)$$

$$\pi(x \mid z) = \operatorname{Gg}(x \mid n, t, 1)$$

Uniform Model

$$\begin{aligned} & \boldsymbol{z} = \{x_1, \dots, x_n\}, & 0 < x_i < \theta \\ & p(x_i \mid \theta) = \operatorname{Un}(x_i \mid 0, \theta), & \theta > 0 \end{aligned}$$

$$\begin{aligned} & \boldsymbol{t}(\boldsymbol{z}) = \boldsymbol{t} = \max\{x_1, \dots, x_n\} \\ & p(t \mid \theta) = \operatorname{Ip}(t \mid n, \theta^{-1}) \end{aligned}$$

$$\begin{aligned} & p(\theta) &= \operatorname{Pa}(\theta \mid \alpha, \beta) \\ & p(x) &= \frac{\alpha}{\alpha + 1} \operatorname{Un}(x \mid 0, \beta), & \text{if } x \leq \beta, & \frac{1}{\alpha + 1} \operatorname{Pa}(x \mid \alpha, \beta), & \text{if } x > \beta \end{aligned}$$

$$p(t) &= \frac{\alpha}{\alpha + n} \operatorname{Ip}(t \mid n, \beta^{-1}), & \text{if } t \leq \beta, & \frac{n}{\alpha + n} \operatorname{Pa}(t \mid \alpha, \beta), & \text{if } t > \beta \end{aligned}$$

$$p(\theta \mid \boldsymbol{z}) &= \operatorname{Pa}(\theta \mid \alpha + n, \beta_n), & \beta_n = \max\{\beta, t\} \\ p(x \mid \boldsymbol{z}) &= \frac{\alpha + n}{\alpha + n + 1} \operatorname{Un}(x \mid 0, \beta_n), & \text{if } x \leq \beta_n, & \frac{1}{\alpha + n + 1} \operatorname{Pa}(x \mid \alpha, \beta_n), & \text{if } x > \beta_n \end{aligned}$$

$$\pi(\theta) \propto \theta^{-1}$$

$$\pi(\theta \mid \boldsymbol{z}) &= \operatorname{Pa}(\theta \mid n, t)$$

$$\pi(x \mid \boldsymbol{z}) &= \frac{n}{n + 1} \operatorname{Un}(x \mid 0, t), & \text{if } x \leq t, & \frac{1}{n + 1} \operatorname{Pa}(x \mid n, t), & \text{if } x > t \end{aligned}$$

Normal Model (known precision \lambda)

$$t(\boldsymbol{z}) = \bar{x} = n^{-1} \sum_{i=1}^{n} x_i$$

$$p(\bar{x} \mid \mu, \lambda) = \mathbf{N}(x \mid \mu, n\lambda)$$

$$p(\mu) = \mathbf{N}(\mu \mid \mu_0, \lambda_0)$$

$$p(x) = \mathbf{N} \left(x \mid \mu_0, \lambda \lambda_0 (\lambda_0 + \lambda)^{-1} \right)$$

$$p(\overline{x}) = N(\overline{x} | \mu_0, n\lambda \lambda_0 \lambda_n^{-1}), \quad \lambda_n = \lambda_0 + n\lambda,$$

$$p(\mu \mid \mathbf{z}) = \mathbf{N}(\mu \mid \mu_n, \lambda_n), \quad \mu_n = \lambda_n^{-1}(\lambda_0 \mu_0 + n\lambda \overline{x})$$
$$p(x \mid \mathbf{z}) = \mathbf{N}\left(x \mid \mu_n, \lambda \lambda_n (\lambda_n + \lambda)^{-1}\right)$$

$$p(x \mid \mathbf{z}) = \mathbf{N} \left(x \mid \mu_n, \lambda \lambda_n (\lambda_n + \lambda)^{-1} \right)$$

$$\pi(\mu) = \text{constant}$$

$$\pi(\mu \,|\, \boldsymbol{z}) = \mathbf{N}(\mu \,|\, \bar{x}, n\lambda)$$

$$\pi(x \mid z) = \mathbf{N}(x \mid \bar{x}, \lambda n(n+1)^{-1})$$

Normal Model (known mean μ)

$$z = \{x_1, \dots, x_n\}, \quad -\infty < x_i < \infty$$

 $p(x_i | \mu, \lambda) = \mathbf{N}(x_i | \mu, \lambda), \quad \lambda > 0$

$$t(z) = t = \sum_{i=1}^{n} (x_i - \mu)^2$$

$$p(t \mid \mu, \lambda) = Ga(t \mid \frac{1}{2}n, \frac{1}{2}\lambda), \qquad p(\lambda t) = \chi^2(\lambda t \mid n)$$

$$p(\lambda) = Ga(\lambda \mid \alpha, \beta)$$

$$p(x) = \operatorname{St}(x \mid \mu, \alpha \beta^{-1}, 2\alpha)$$

$$p(t) = \operatorname{Gg}(x \mid \alpha, 2\beta, \frac{1}{2}n)$$

$$p(\lambda \mid \boldsymbol{z}) = \operatorname{Ga}(\lambda \mid \alpha + \frac{1}{2}n, \beta + \frac{1}{2}t)$$

$$p(x \,|\, \pmb{z}) = \mathrm{St}(x \,|\, \mu, (\alpha + \tfrac{1}{2}n)(\beta + \tfrac{1}{2}t)^{-1}, 2\alpha + n)$$

$$\pi(\lambda) \propto \lambda^{-1}$$

$$\pi(\lambda \mid \boldsymbol{z}) = \operatorname{Ga}(\lambda \mid \frac{1}{2}n, \frac{1}{2}t)$$

$$\pi(x \,|\, \boldsymbol{z}) = \operatorname{St}(x \,|\, \mu, nt^{-1}, n)$$

Normal Model (both parameters unknown)

$$\begin{split} & \mathbf{z} = \{x_1, \dots, x_n\}, \quad -\infty < x_i < \infty \\ & p(x_i \mid \mu, \lambda) = \mathbf{N}(x_i \mid \mu, \lambda), \quad -\infty < \mu < \infty, \quad \lambda > 0 \\ & t(\mathbf{z}) = (\bar{x}, s), \quad n\bar{x} = \sum_{i=1}^n x_i, \quad ns^2 = \sum_{i=1}^n (x_i - \bar{x})^2 \\ & p(\bar{x} \mid \mu, \lambda) = \mathbf{N}(\bar{x} \mid \mu, n\lambda) \\ & p(ns^2 \mid \mu, \lambda) = \mathbf{Ga}(ns^2 \mid \frac{1}{2}(n-1), \frac{1}{2}\lambda), \quad p(\lambda ns^2) = \chi^2(\lambda ns^2 \mid n-1) \\ & p(\mu, \lambda) = \mathbf{Ng}(\mu, \lambda \mid \mu_0, n_0, \alpha, \beta) = \mathbf{N}(\mu \mid \mu_0, n_0\lambda) \, \mathbf{Ga}(\lambda \mid \alpha, \beta) \\ & p(\mu) = \mathbf{St}(\mu \mid \mu_0, n_0\alpha\beta^{-1}, 2\alpha) \\ & p(\lambda) = \mathbf{Ga}(\lambda \mid \alpha, \beta) \\ & p(x) = \mathbf{St}(x \mid \mu_0, n_0(n_0 + 1)^{-1}\alpha\beta^{-1}, 2\alpha) \\ & p(\bar{x}) = \mathbf{St}(\bar{x} \mid \mu_0, n_0n(n_0 + n)^{-1}\alpha\beta^{-1}, 2\alpha) \\ & p(ns^2) = \mathbf{Gg}(ns^2 \mid \alpha, 2\beta, \frac{1}{2}(n-1)) \\ & p(\mu \mid \mathbf{z}) = \mathbf{St}(\mu \mid \mu_n, (n+n_0)(\alpha+\frac{1}{2}n)\beta_n^{-1}, 2\alpha+n), \\ & \mu_n = (n_0+n)^{-1}(n_0\mu_0+n\bar{x}), \\ & \beta_n = \beta + \frac{1}{2}ns^2 + \frac{1}{2}(n_0+n)^{-1}n_0n(\mu_0-\bar{x})^2 \\ & p(\lambda \mid \mathbf{z}) = \mathbf{Ga}(\lambda \mid \alpha+\frac{1}{2}n, \beta_n) \\ & p(x \mid \mathbf{z}) = \mathbf{St}(x \mid \mu_n, (n+n_0)(n+n_0+1)^{-1}(\alpha+\frac{1}{2}n)\beta_n^{-1}, 2\alpha+n) \\ & \pi(\mu, \lambda) = \pi(\lambda, \mu) \propto \lambda^{-1}, \quad n > 1 \\ & \pi(\mu \mid \mathbf{z}) = \mathbf{St}(\mu \mid \bar{x}, (n-1)s^{-2}, n-1) \\ & \pi(\lambda \mid \mathbf{z}) = \mathbf{Ga}(\lambda \mid \frac{1}{2}(n-1), \frac{1}{2}ns^2) \\ & \pi(x \mid \mathbf{z}) = \mathbf{St}(x \mid \bar{x}, (n-1)(n+1)^{-1}s^{-2}, n-1) \end{split}$$

Multinomial Model

$$\begin{aligned} & \boldsymbol{z} = \{r_1, \dots, r_k, n\}, & r_i = 0, 1, 2, \dots, & \sum_{\ell=1}^k r_i \leq n \\ & p(\boldsymbol{z} \mid \boldsymbol{\theta}) = \operatorname{Mu}_k(\boldsymbol{z} \mid \boldsymbol{\theta}, n), & 0 < \theta_i < 1, & \sum_{i=1}^k \theta_i \leq 1 \end{aligned}$$

$$\begin{aligned} & \boldsymbol{t}(\boldsymbol{z}) = (\boldsymbol{r}, n), & \boldsymbol{r} = (r_1, \dots, r_k) \\ & p(\boldsymbol{r} \mid \boldsymbol{\theta}) = \operatorname{Mu}_k(\boldsymbol{r} \mid \boldsymbol{\theta}, n) \end{aligned}$$

$$p(\boldsymbol{\theta}) = \operatorname{Di}_k(\boldsymbol{\theta} \mid \boldsymbol{\alpha}), & \boldsymbol{\alpha} = \{\alpha_1, \dots, \alpha_{k+1}\} \\ & p(\boldsymbol{r}) = \operatorname{Md}_k(\boldsymbol{r} \mid \boldsymbol{\alpha}, n) \\ & p(\boldsymbol{\theta} \mid \boldsymbol{z}) = \operatorname{Di}_k\left(\boldsymbol{\theta} \mid \alpha_1 + r_1, \dots, \alpha_k + r_k, \alpha_{k+1} + n - \sum_{\ell=1}^k r_\ell\right) \\ & p(\boldsymbol{x} \mid \boldsymbol{z}) = \operatorname{Md}_k\left(\boldsymbol{x} \mid \alpha_1 + r_1, \dots, \alpha_k + r_k, \alpha_{k+1} + n - \sum_{\ell=1}^k r_\ell, n\right) \end{aligned}$$

Multivariate Normal Model

$$\begin{aligned} & \boldsymbol{z} = \{\boldsymbol{x}_1, \dots, \boldsymbol{x}_n\}, & \boldsymbol{x}_i \in \Re^k \\ & p(\boldsymbol{x}_i \mid \boldsymbol{\mu}, \boldsymbol{\lambda}) = \mathrm{N}_k(\boldsymbol{x}_i \mid \boldsymbol{\mu}, \boldsymbol{\lambda}), & \boldsymbol{\mu} \in \Re^k, & \boldsymbol{\lambda} \quad k \times k \text{ positive-definite} \end{aligned}$$

$$\begin{aligned} & \boldsymbol{t}(\boldsymbol{z}) = (\bar{\boldsymbol{x}}, \boldsymbol{S}), & \overline{\boldsymbol{x}} = n^{-1} \sum_{i=1}^n \boldsymbol{x}_i, & \boldsymbol{S} = \sum_{i=1}^n (\boldsymbol{x}_i - \bar{\boldsymbol{x}})(\boldsymbol{x}_i - \bar{\boldsymbol{x}})^t \\ & p(\bar{\boldsymbol{x}} \mid \boldsymbol{\mu}, \boldsymbol{\lambda}) = \mathrm{N}_k(\bar{\boldsymbol{x}} \mid \boldsymbol{\mu}, n\boldsymbol{\lambda}) \\ & p(\boldsymbol{S} \mid \boldsymbol{\lambda}) = \mathrm{Wi}_k \left(\boldsymbol{S} \mid \frac{1}{2}(n-1), \frac{1}{2}\boldsymbol{\lambda}\right) \end{aligned}$$

$$p(\boldsymbol{\mu}, \boldsymbol{\lambda}) = \mathrm{Nw}_k(\boldsymbol{\mu}, \boldsymbol{\lambda} \mid \boldsymbol{\mu}_0, n_0, \alpha, \boldsymbol{\beta}) = \mathrm{N}_k(\boldsymbol{\mu} \mid \boldsymbol{\mu}_0, n_0\boldsymbol{\lambda}) \, \mathrm{Wi}_k(\boldsymbol{\lambda} \mid \alpha, \boldsymbol{\beta}) \\ p(\boldsymbol{x}) = \mathrm{St}_k \left(\boldsymbol{x} \mid \boldsymbol{\mu}_0, (n_0+1)^{-1} n_0(\alpha - \frac{1}{2}(k-1))\boldsymbol{\beta}^{-1}, 2\alpha - k + 1\right) \end{aligned}$$

$$p(\boldsymbol{\mu} \mid \boldsymbol{z}) = \mathrm{St}_k \left(\boldsymbol{\mu} \mid \boldsymbol{\mu}_n, (n+n_0)\alpha_n\boldsymbol{\beta}_n^{-1}, 2\alpha_n\right), \\ \boldsymbol{\mu}_n = (n_0+n)^{-1}(n_0\boldsymbol{\mu}_0 + n\bar{\boldsymbol{x}}), \\ \boldsymbol{\beta}_n = \boldsymbol{\beta} + \frac{1}{2}\boldsymbol{S} + \frac{1}{2}(n+n_0)^{-1}nn_0(\boldsymbol{\mu}_0 - \overline{\boldsymbol{x}})(\boldsymbol{\mu}_0 - \overline{\boldsymbol{x}})^t \end{aligned}$$

$$p(\boldsymbol{\lambda} \mid \boldsymbol{z}) = \mathrm{Wi}_k(\boldsymbol{\lambda} \mid \alpha + \frac{1}{2}n, \boldsymbol{\beta}_n)$$

$$p(\boldsymbol{x} \mid \boldsymbol{z}) = \mathrm{St}_k(\boldsymbol{x} \mid \boldsymbol{\mu}_n, (n_0+n+1)^{-1}(n_0+n)\alpha_n\boldsymbol{\beta}_n^{-1}, 2\alpha_n), \\ \boldsymbol{\alpha}_n = \boldsymbol{\alpha} + \frac{1}{2}n - \frac{1}{2}(k-1) \end{aligned}$$

Linear Regression

$$\begin{aligned} & \boldsymbol{z} = (\boldsymbol{y}, \boldsymbol{X}), \quad \boldsymbol{y} = (y_1, \dots, y_n)^t \in \Re^n, \quad \boldsymbol{x}_i = (x_{i1}, \dots, x_{ik}) \in \Re^k, \quad \boldsymbol{X} = (x_{ij}) \\ & p(\boldsymbol{y} \mid \boldsymbol{X}, \boldsymbol{\theta}, \lambda) = \mathrm{N}_n(\boldsymbol{y} \mid \boldsymbol{X} \boldsymbol{\theta}, \lambda \boldsymbol{I}_n), \quad \boldsymbol{\theta} \in \Re^k, \quad \lambda > 0 \end{aligned}$$

$$\begin{aligned} & \boldsymbol{t}(\boldsymbol{z}) = (\boldsymbol{X}^t \boldsymbol{X}, \boldsymbol{X}^t \boldsymbol{y}) \\ & p(\boldsymbol{\theta}, \lambda) = \mathrm{Ng}(\boldsymbol{\theta}, \lambda \mid \boldsymbol{\theta}_0, \boldsymbol{n}_0, \alpha, \beta) = \mathrm{N}_k(\boldsymbol{\theta} \mid \boldsymbol{\theta}_0, \boldsymbol{n}_0 \lambda) \operatorname{Ga}(\lambda \mid \alpha, \beta) \\ & p(\boldsymbol{\theta}) = \mathrm{St}_k(\boldsymbol{\theta} \mid \boldsymbol{\theta}_0, \quad \boldsymbol{n}_0 \alpha \beta^{-1}, \quad 2\alpha) \\ & p(\lambda) = \mathrm{Ga}(\lambda \mid \alpha, \beta) \\ & p(\boldsymbol{y} \mid \boldsymbol{x}) = \mathrm{St} \left(\boldsymbol{y} \mid \boldsymbol{x} \boldsymbol{\theta}_0, \quad f(\boldsymbol{x}) \alpha \beta^{-1}, \quad 2\alpha\right) \\ & f(\boldsymbol{x}) = 1 - \boldsymbol{x}(\boldsymbol{x}^t \boldsymbol{x} + \boldsymbol{n}_0)^{-1} \boldsymbol{x}^t, \\ & p(\boldsymbol{\theta} \mid \boldsymbol{z}) = \mathrm{St}_k \left(\boldsymbol{\theta} \mid \boldsymbol{\theta}_n, \quad (\boldsymbol{n}_0 + \boldsymbol{X}^t \boldsymbol{X})(\alpha + \frac{1}{2}n)\beta_n^{-1}, \quad 2\alpha + n\right), \\ & \boldsymbol{\theta}_n = (\boldsymbol{n}_0 + \boldsymbol{X}^t \boldsymbol{X})^{-1} (\boldsymbol{n}_0 \boldsymbol{\theta}_0 + \boldsymbol{X}^t \boldsymbol{y}), \\ & \boldsymbol{\beta}_n = \beta + \frac{1}{2} (\boldsymbol{y} - \boldsymbol{X} \boldsymbol{\theta}_n)^t \boldsymbol{y} + \frac{1}{2} (\boldsymbol{\theta}_0 - \boldsymbol{\theta}_n)^t \boldsymbol{n}_0 \boldsymbol{\theta}_0 \end{aligned}$$

$$p(\lambda \mid \boldsymbol{z}) = \mathrm{Ga}(\lambda \mid \alpha + \frac{1}{2}n, \beta_n) \\ p(\boldsymbol{y} \mid \boldsymbol{x}, \boldsymbol{z}) = \mathrm{St}(\boldsymbol{y} \mid \boldsymbol{x} \boldsymbol{\theta}_n, \quad f_n(\boldsymbol{x})(\alpha + \frac{1}{2}n)\beta_n^{-1}, \quad 2\alpha + n), \\ & f_n(\boldsymbol{x}) = 1 - \boldsymbol{x}(\boldsymbol{x}^t \boldsymbol{x} + \boldsymbol{n}_0 + \boldsymbol{X}^t \boldsymbol{X})^{-1} \boldsymbol{x}^t \end{aligned}$$

$$\pi(\boldsymbol{\theta}, \lambda) = \pi(\lambda, \boldsymbol{\theta}) \propto \lambda^{-(k+1)/2} \quad \text{(for all reorderings of the } \boldsymbol{\theta}_j) \\ \pi(\boldsymbol{\theta} \mid \boldsymbol{z}) = \mathrm{St}_k \left(\boldsymbol{\theta} \mid \hat{\boldsymbol{\theta}}_n, \quad \frac{1}{2} \boldsymbol{X}^t \boldsymbol{X}(n - k) \hat{\boldsymbol{\beta}}_n^{-1}, \quad n - k\right), \\ & \hat{\boldsymbol{\theta}}_n = (\boldsymbol{X}^t \boldsymbol{X})^{-1} \boldsymbol{X}^t \boldsymbol{y}, \\ & \hat{\boldsymbol{\theta}}_n = \frac{1}{2} (\boldsymbol{y} - \boldsymbol{X} \hat{\boldsymbol{\theta}}_n)^t \boldsymbol{y} \\ \pi(\lambda \mid \boldsymbol{z}) = \mathrm{Ga}(\lambda \mid \frac{1}{2} (n - k), \hat{\boldsymbol{\beta}}_n) \\ \pi(\boldsymbol{y} \mid \boldsymbol{x}, \boldsymbol{z}) = \mathrm{St}(\boldsymbol{y} \mid \boldsymbol{x} \hat{\boldsymbol{\theta}}_n, \quad \frac{1}{2} f_n(\boldsymbol{x})(n - k) \hat{\boldsymbol{\beta}}_n^{-1}, \quad n - k), \end{aligned}$$

 $f_n(\boldsymbol{x}) = 1 - \boldsymbol{x}(\boldsymbol{x}^t \boldsymbol{x} + \boldsymbol{X}^t \boldsymbol{X})^{-1} \boldsymbol{x}^t$