

# **GENERAL LINEAR MODELS**

**2 NOVEMBER 2018**

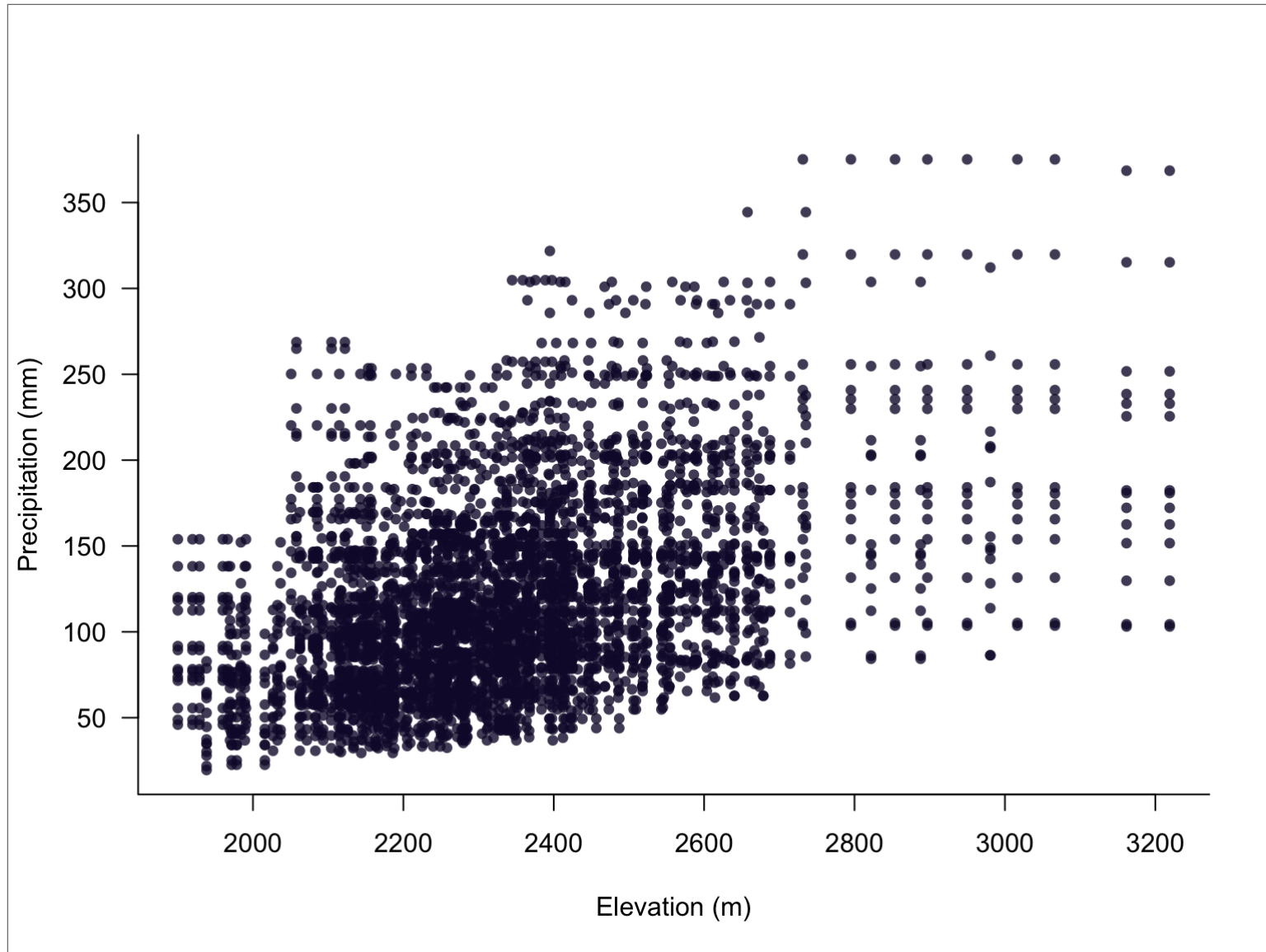
# WORKSHOP OVERVIEW

- **[https://github.com/goldingn/linear\\_models\\_workshop](https://github.com/goldingn/linear_models_workshop)**
- general linear models
- mixed effects models
- generalised linear models
- Bayesian inference

## **EXPECTED OUTCOMES**

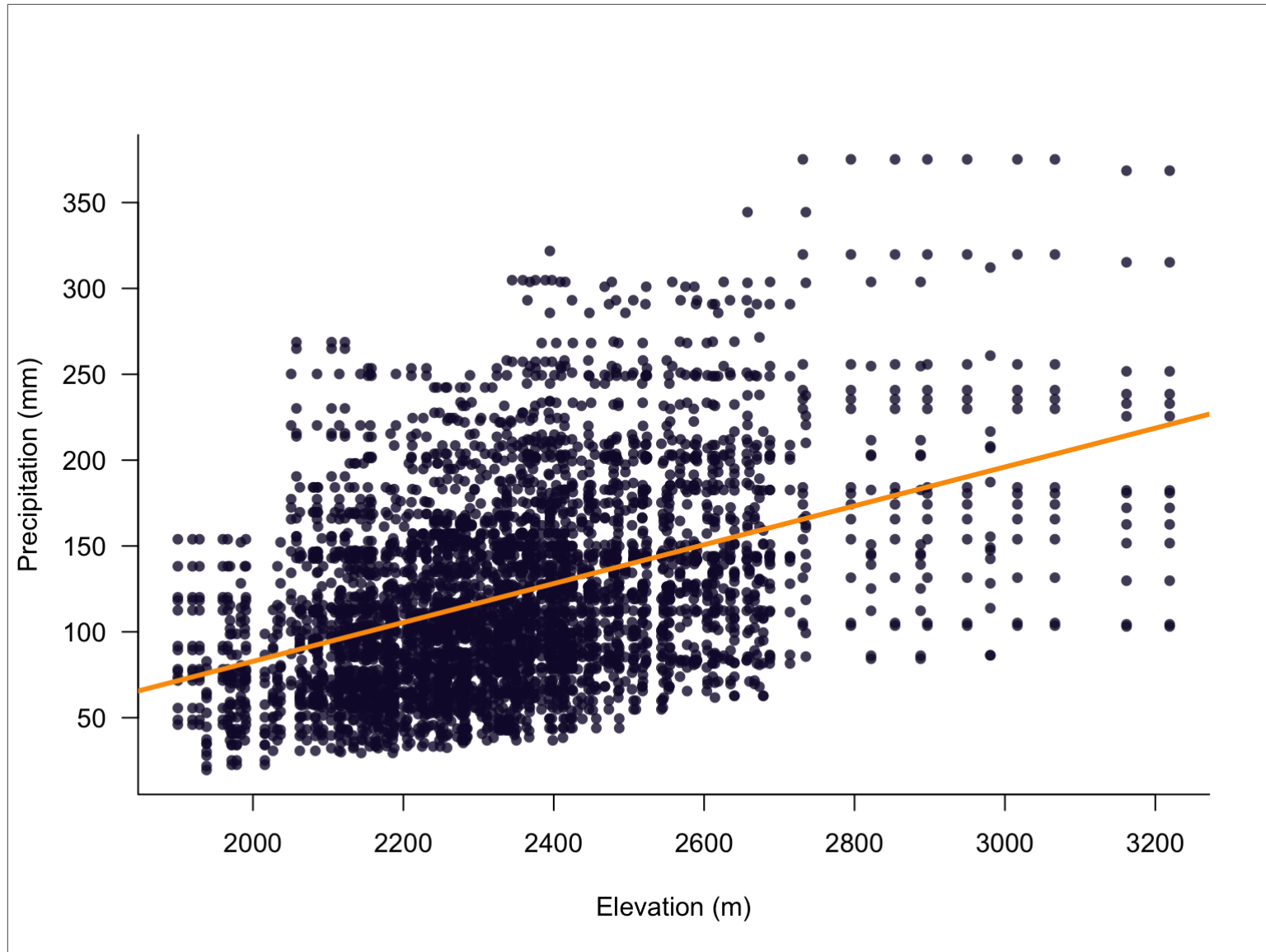
- learn some new terms
- identify appropriate models for your data
- understand assumptions of common models
- know where to look for help

# AN EXAMPLE





# AN EXAMPLE





# LINEAR REGRESSION

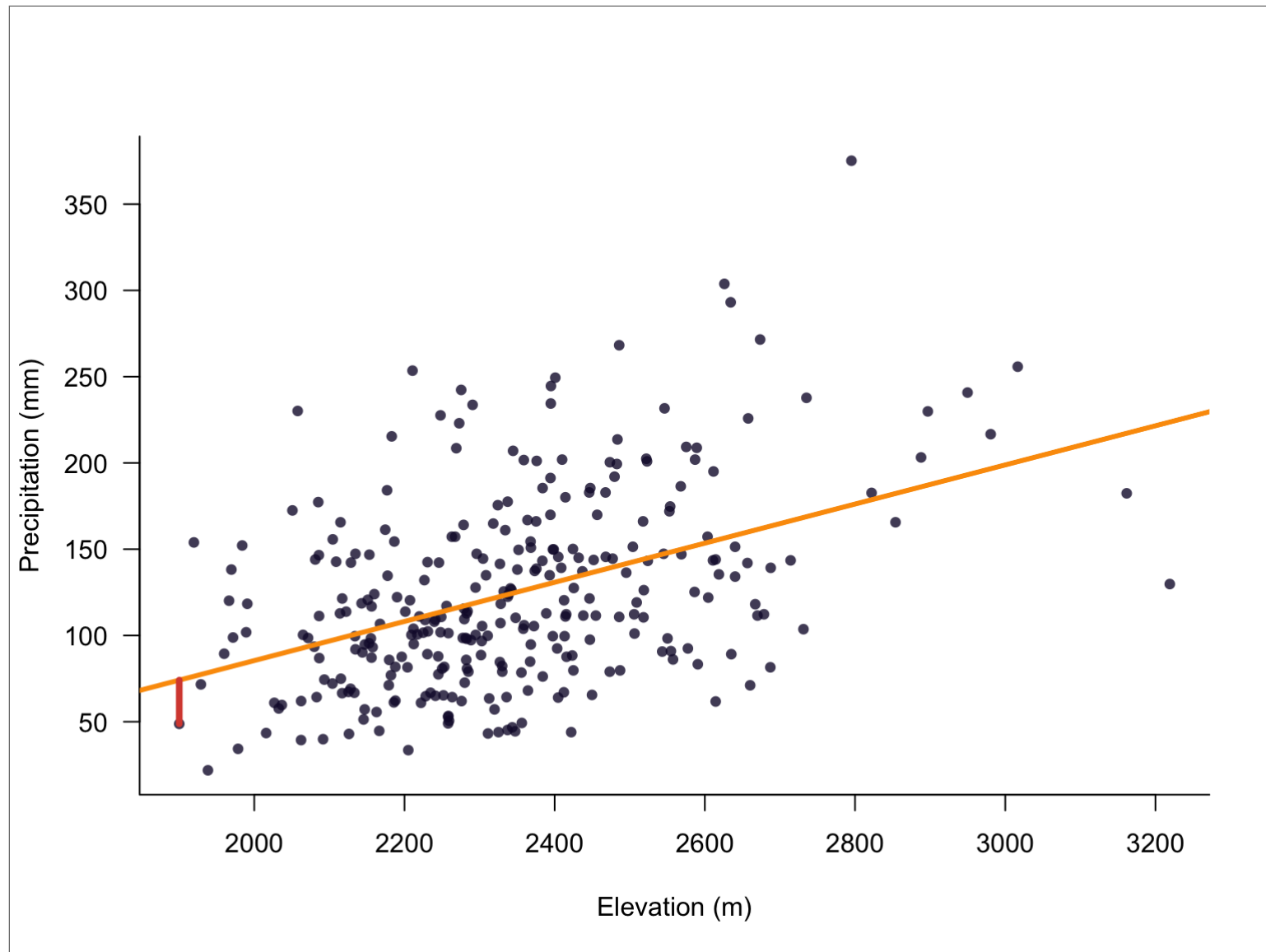
- what characterises this example?
  - continuous response
  - continuous predictor



# LINEAR REGRESSION

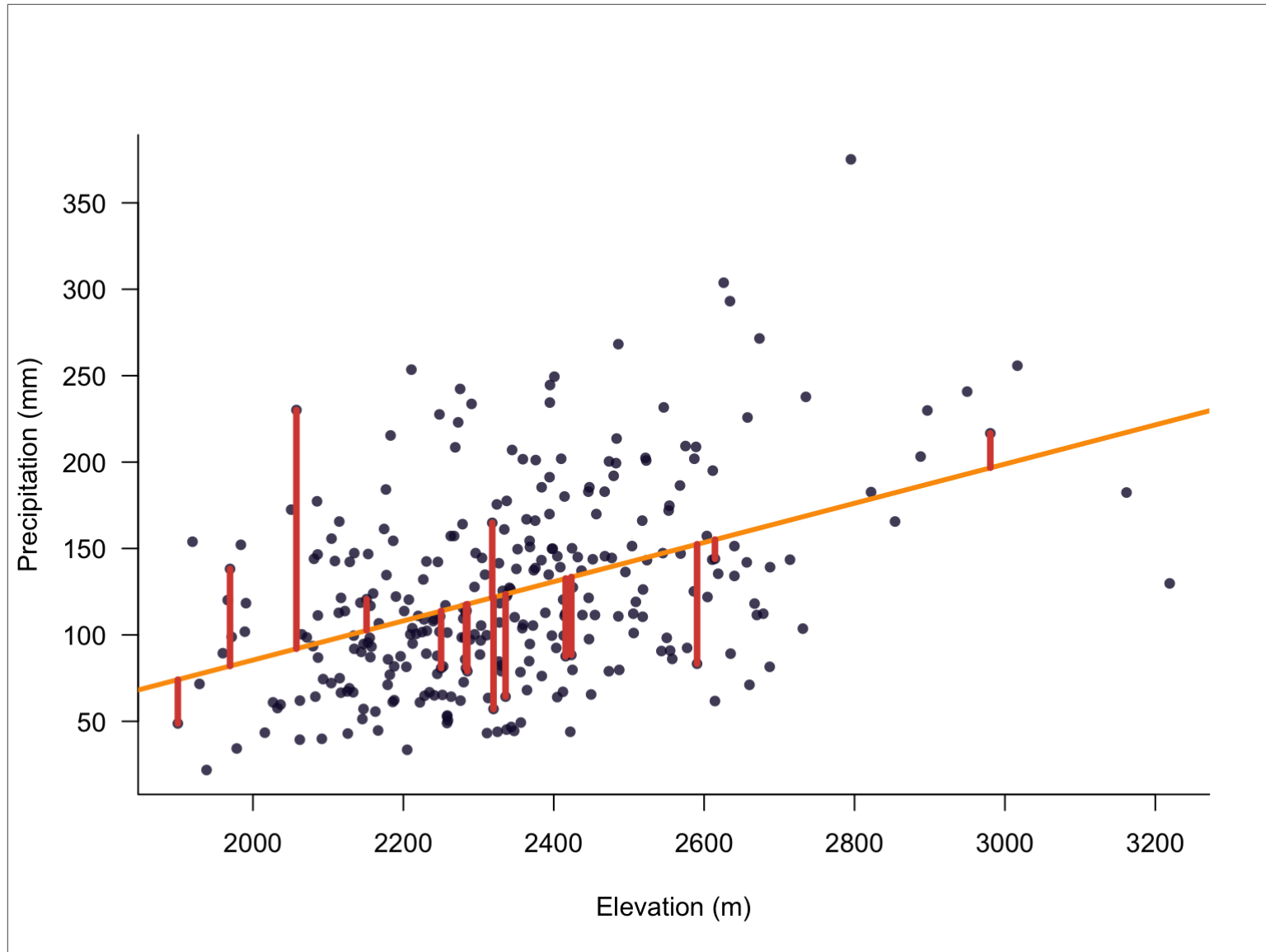
- need an equation for a line:
  - $response = intercept + slope \times predictor + residual$
  - $y_i = \alpha + \beta x_i + \epsilon_i$
- our goal is to estimate the best values of  $\alpha$  and  $\beta$

# LINEAR REGRESSION





# LINEAR REGRESSION





# **LINEAR REGRESSION: ASSUMPTIONS**

- observations are independent
- residuals are normally distributed
- residual variance is constant

# **LINEAR REGRESSION: INDEPENDENT OBSERVATIONS**

- each independent observations gives a certain amount of information
- non-independent observations give less information
- how to avoid issues: good study design
- how to address issues: mixed models (second session)

# LINEAR REGRESSION: RESIDUAL DISTRIBUTION

- $y_i = \alpha + \beta x_i + \epsilon_i$
- we assume  $\epsilon_i$  is normally distributed
  - needed to define *likelihood*
- how to identify issues: diagnostic checks
- how to address issues: transformations, generalised linear models



# LINEAR REGRESSION: CONSTANT RESIDUAL VARIANCE

- $y_i = \alpha + \beta x_i + \epsilon_i$
- we assume  $\epsilon_i$  all come from one distribution
- how to identify issues: diagnostic checks
- how to address issues: transformations, GLMs, hierarchical models

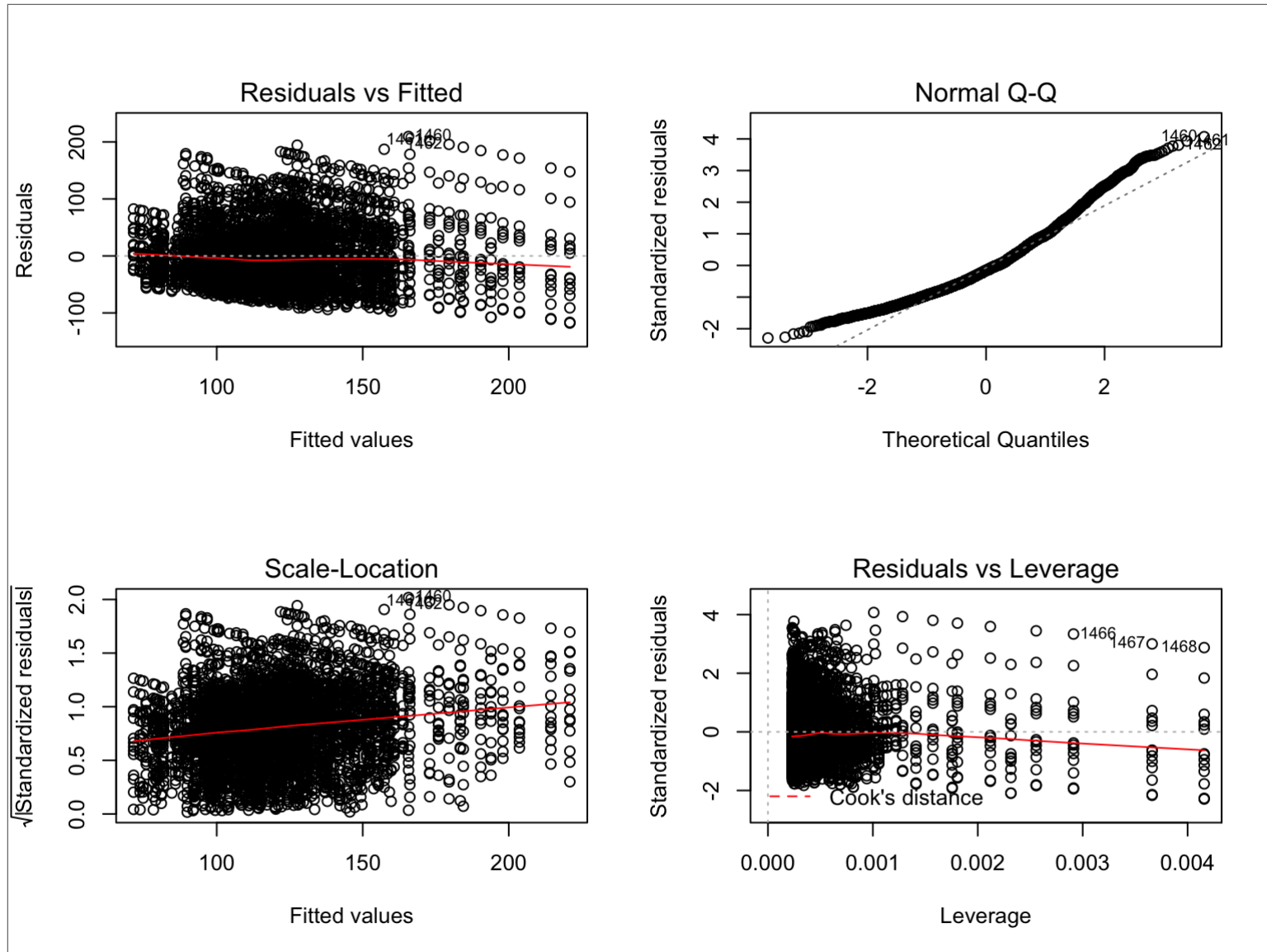
# LINEAR REGRESSION: FITTING A MODEL IN R

```
# fit model
mod <- lm(precipitation ~ elevation)

# check model assumptions
plot(mod)

# summarise model
summary(mod)
```

# LINEAR REGRESSION: ASSESSING A FITTED MODEL

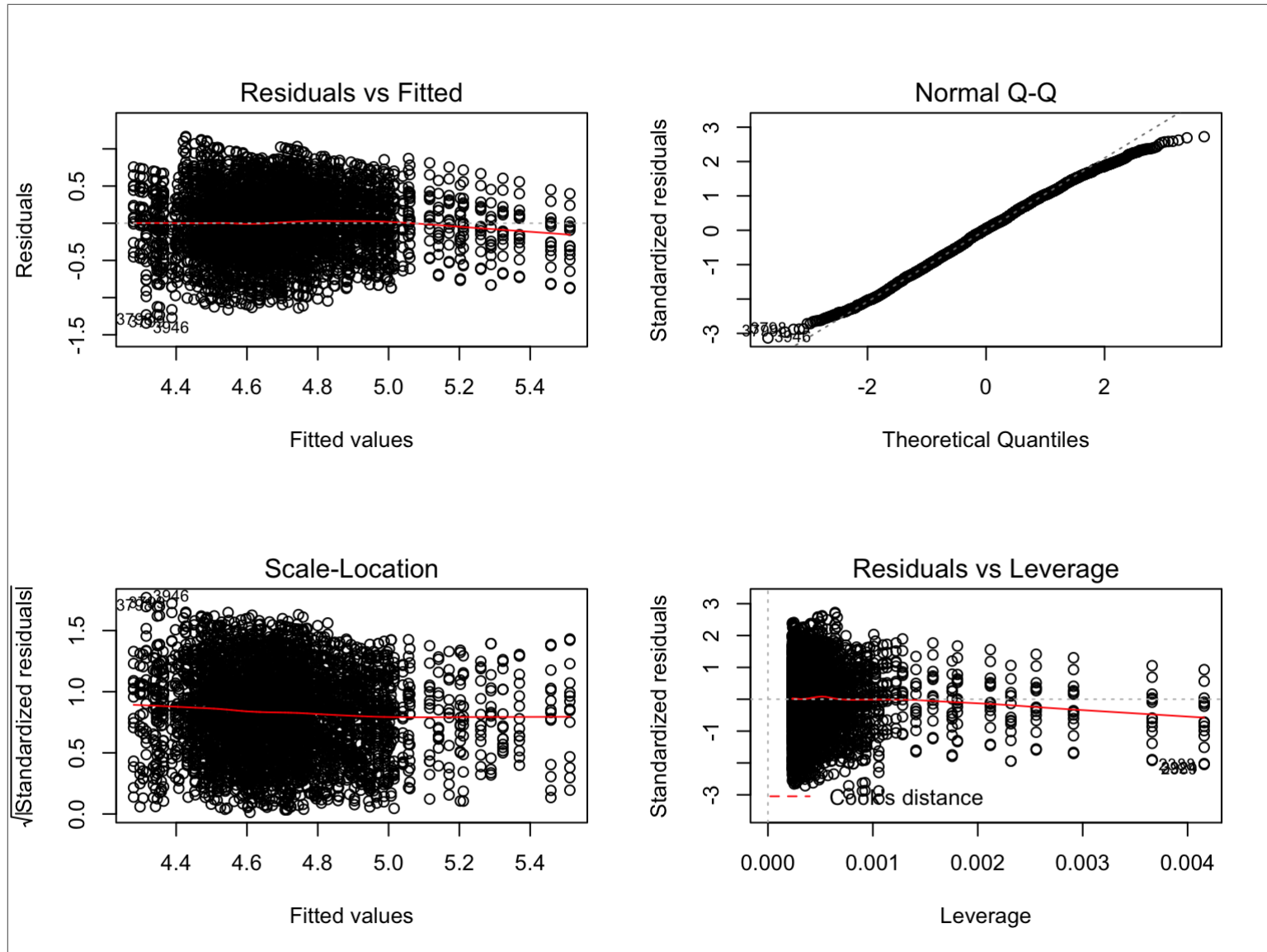




# LINEAR REGRESSION: ASSESSING A FITTED MODEL

```
# fit model with log-transformed response  
mod <- lm(log(precipitation) ~ elevation)  
  
# is it any better?  
plot(mod)
```

## LINEAR REGRESSION: ASSESSING A FITTED MODEL





# LINEAR REGRESSION: INTERPRETING A FITTED MODEL

```
summary(mod)
```

```
##
## Call:
## lm(formula = precipitation ~ elevation)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -117.788  -37.852   -8.118   30.345  209.488
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept) -1.437e+02  8.632e+00  -16.64  <2e-16 ***
## elevation    1.133e-01  3.672e-03   30.84  <2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 51.48 on 4296 degrees of freedom
## Multiple R-squared:  0.1813, Adjusted R-squared:  0.1811
## F-statistic: 951.1 on 1 and 4296 DF,  p-value: < 2.2e-16
```



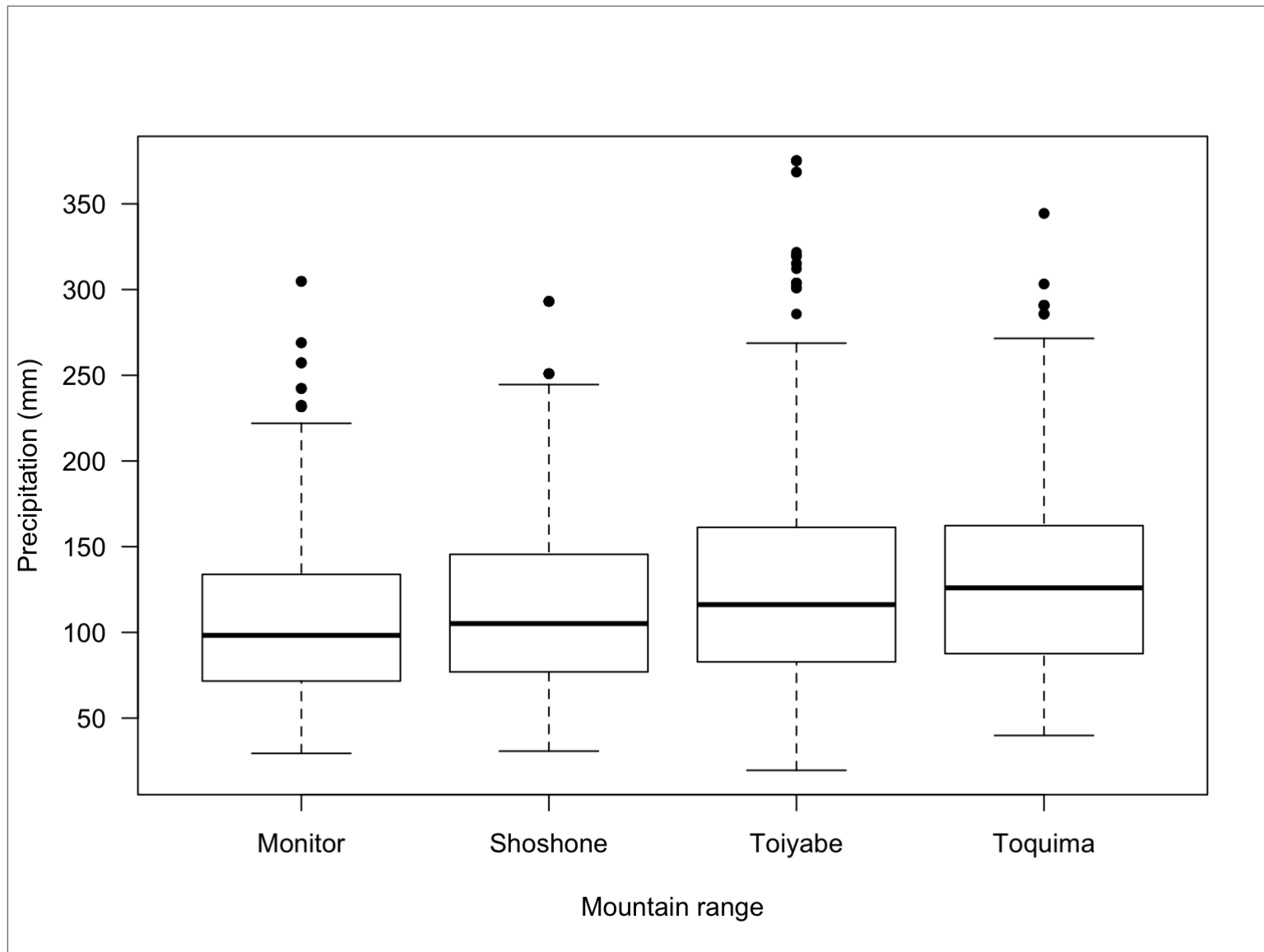
# LINEAR REGRESSION: INTERPRETING A FITTED MODEL

- how well does the model fit?
  - typically use  $r^2$
- is there statistical support for an association?
  - often use *p-values*
- is a statistically supported association meaningful?
  - look at the coefficients

## **LINEAR REGRESSION: PRESENTING A FITTED MODEL**

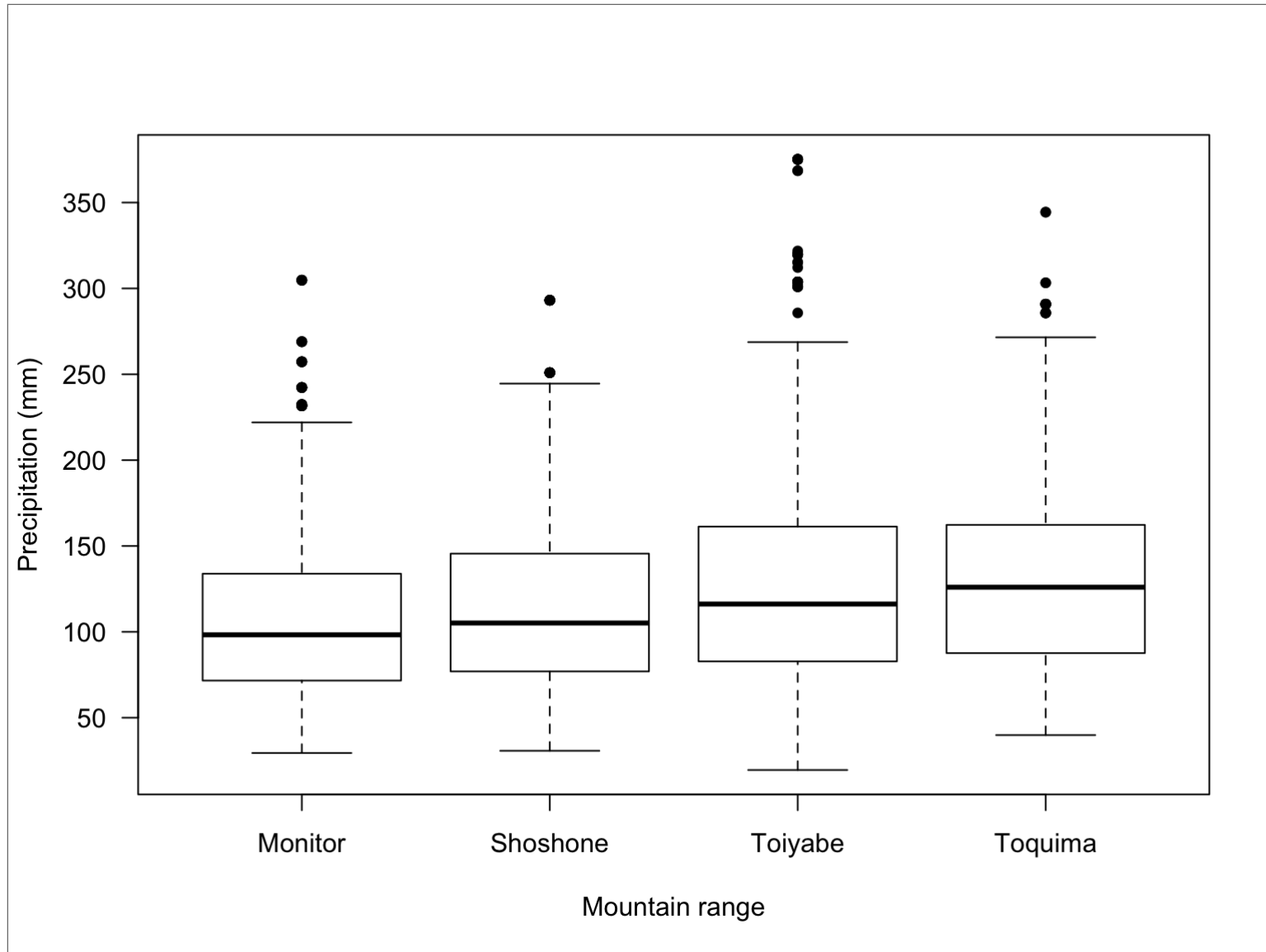
- is the model adequate? (assumptions, diagnostics)
- does the model fit the data? (diagnostics,  $r^2$ )
- is the model statistically meaningful? (p-values, test statistics)
- is the model actually meaningful? (parameter estimates)
- can I see it? (scatterplots)

## ANOTHER EXAMPLE



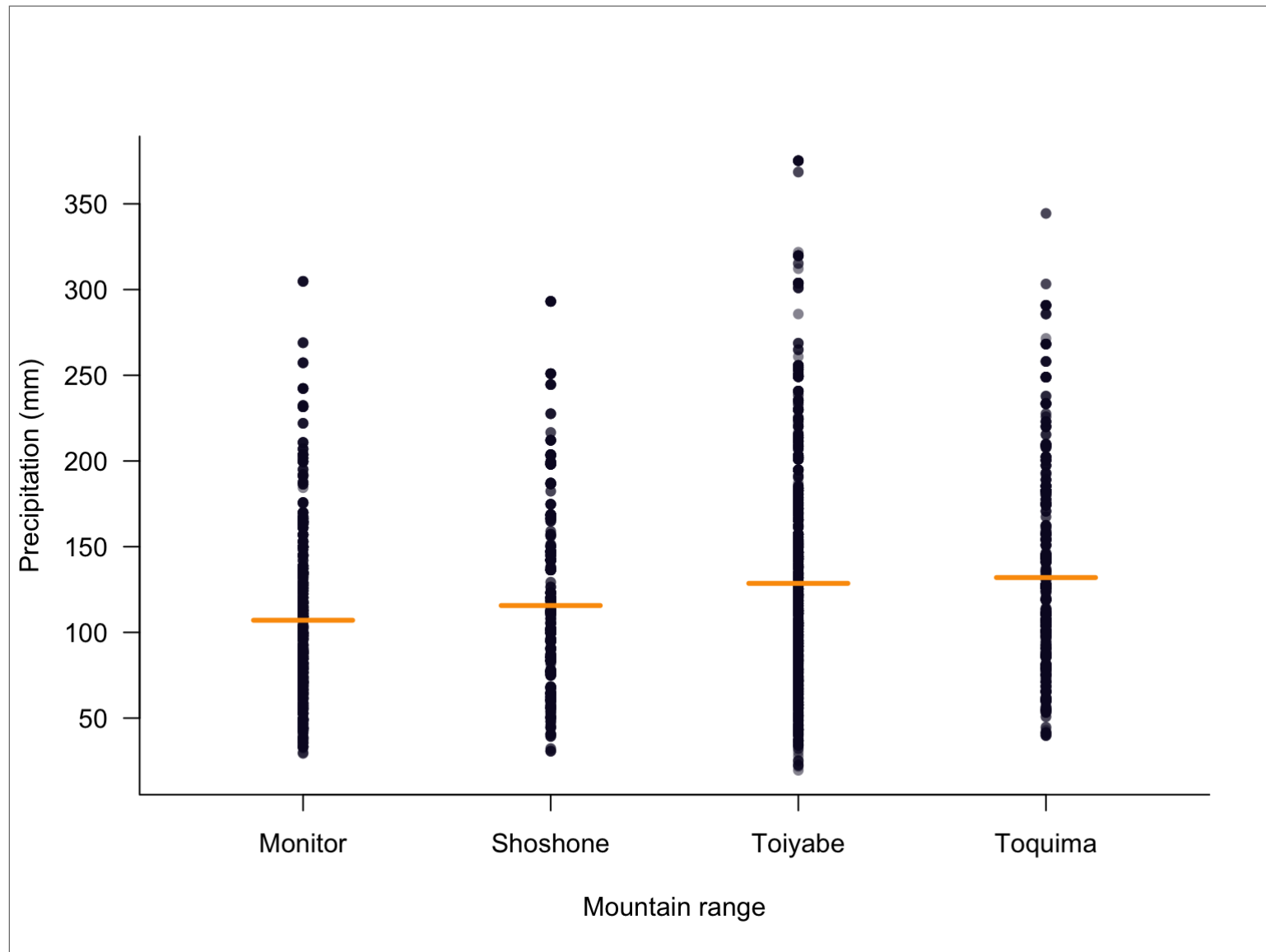


# ANOVA





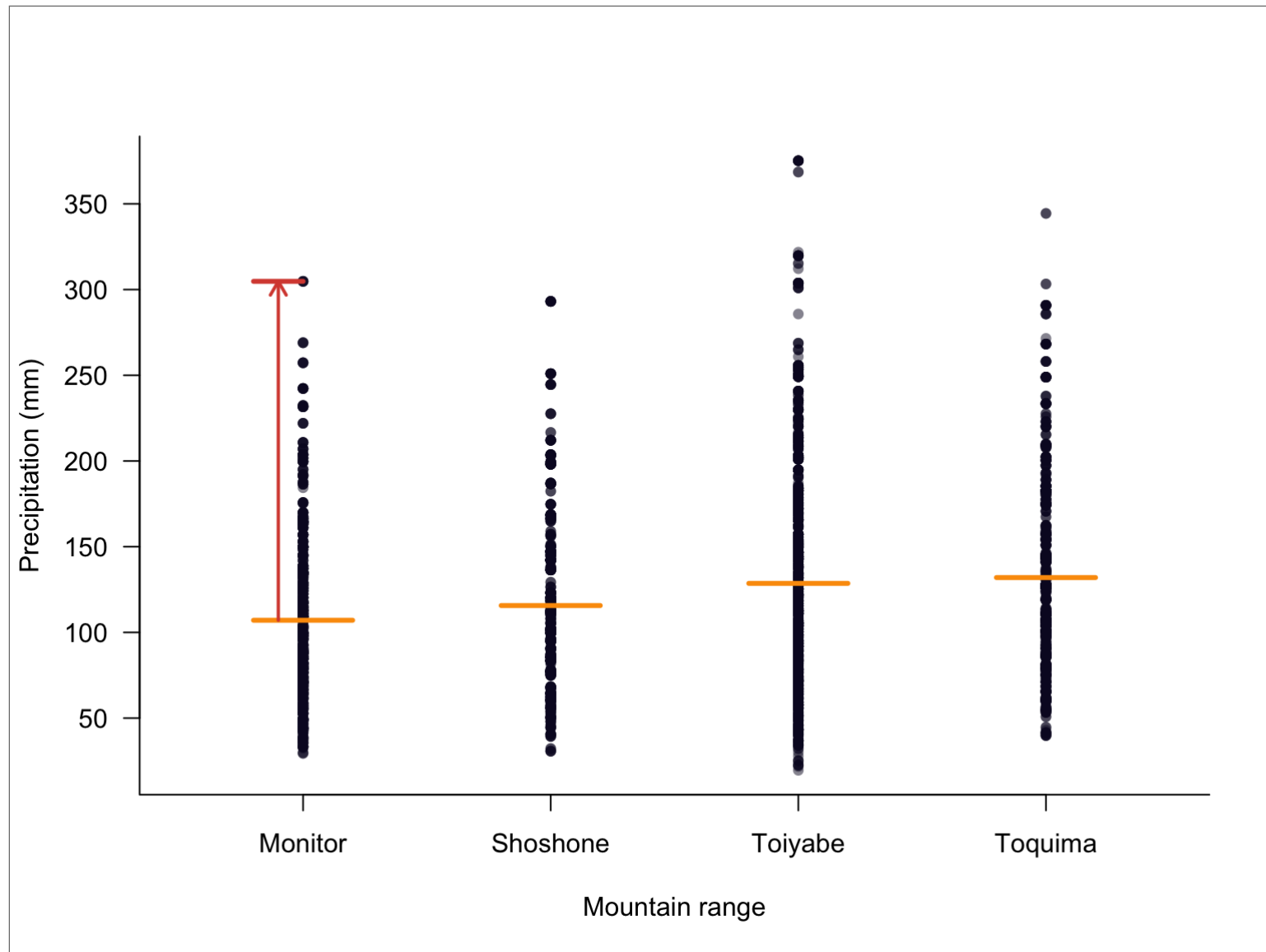
# ANOVA







# ANOVA





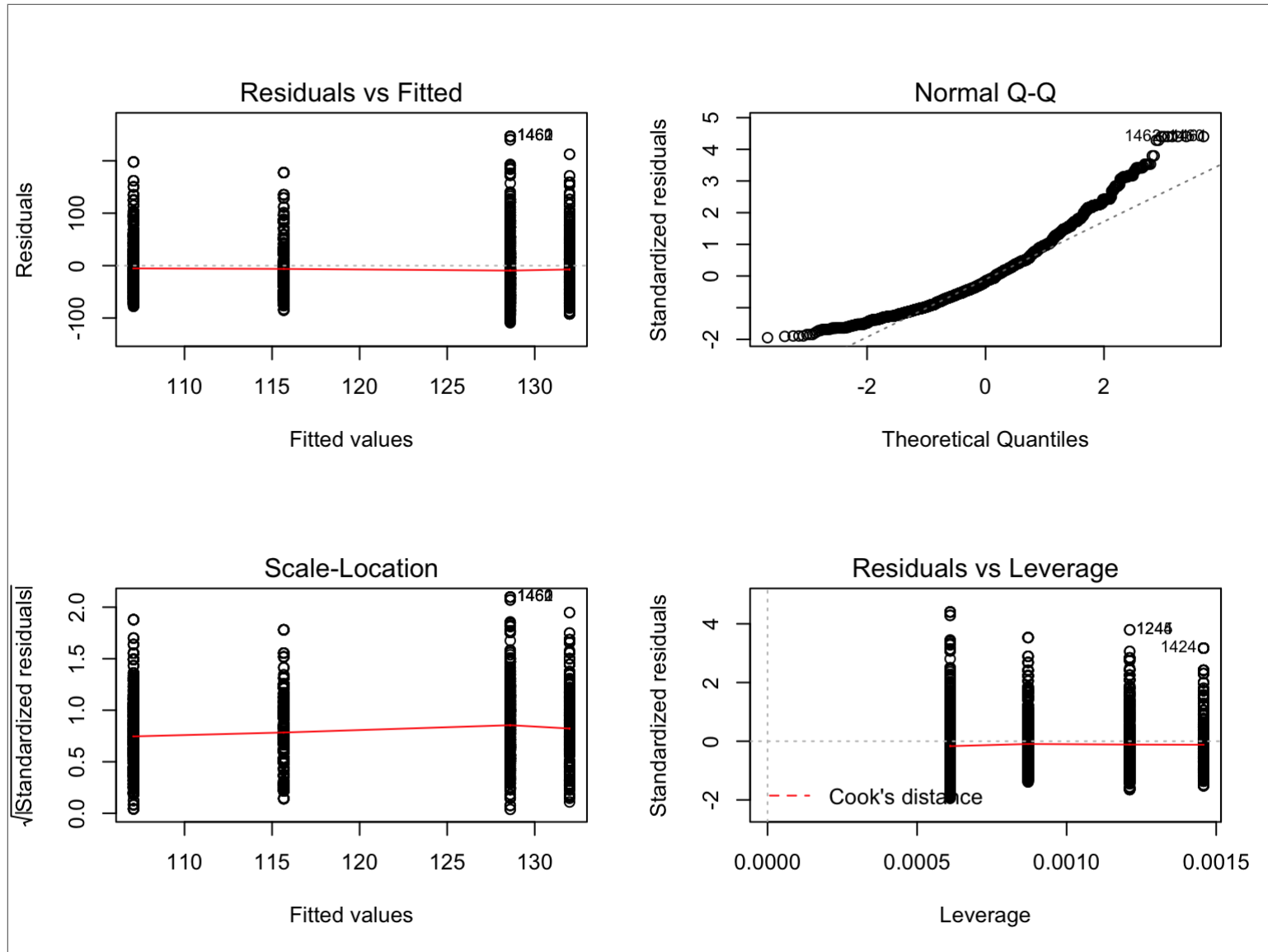
# ANOVA

- what characterises this example?
  - continuous response
  - *discrete* predictor
- assumptions: identical to linear regression
- *response = overall intercept + group intercept + residual*
- $y_i = \alpha + \beta_{g(i)} + \epsilon_i$

# ANOVA: FITTING A MODEL IN R

```
# fit a model  
mod <- lm(response ~ predictor, data = data_set)  
  
# does the model meet assumptions?  
plot(mod)  
  
# summarise the model  
summary(mod)
```

# ANOVA: ASSESSING A FITTED MODEL





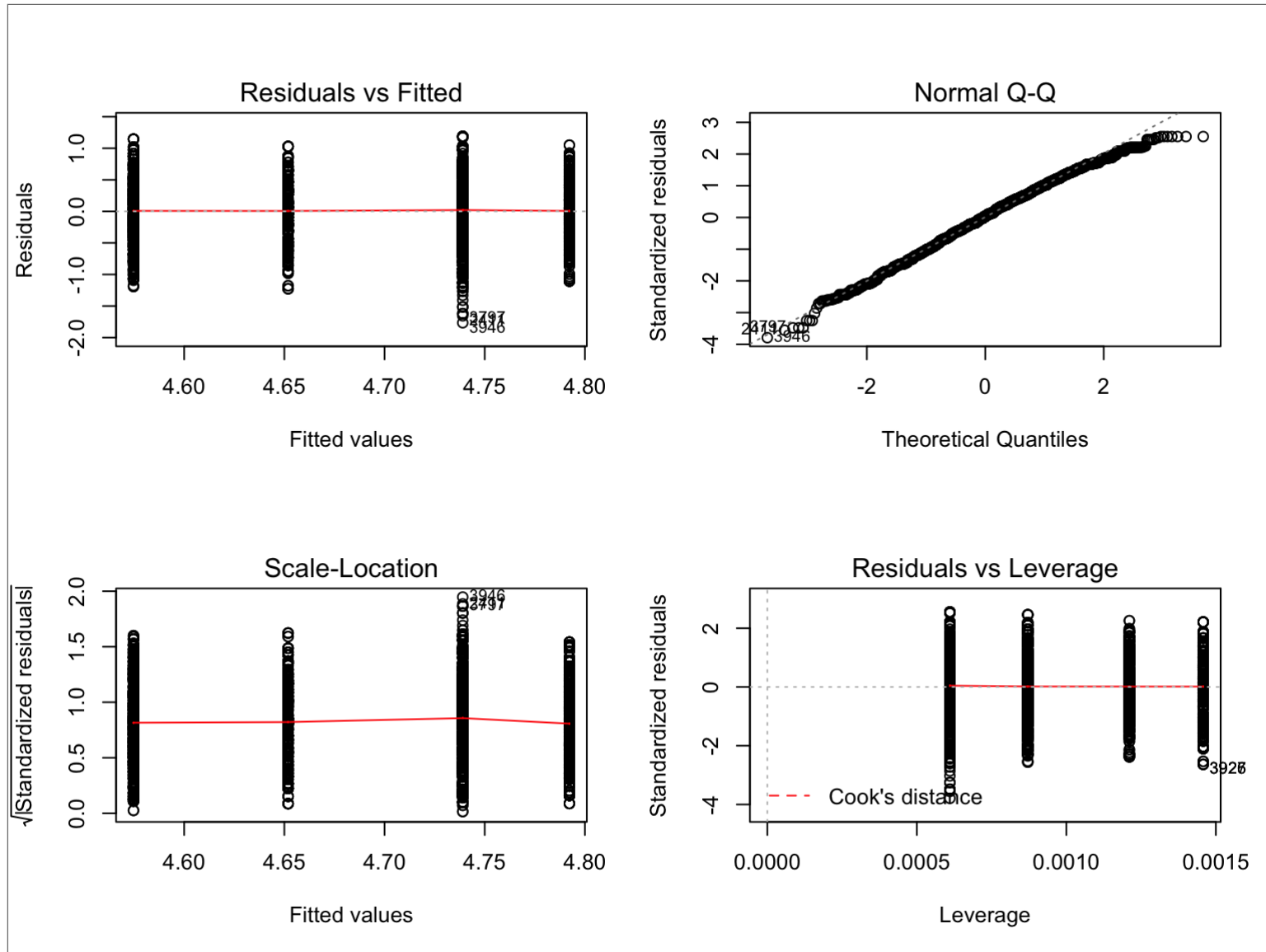
# ANOVA: FITTING A MODEL IN R

```
# fit a model to log-transformed data
mod <- lm(log(response) ~ predictor, data = data_set)

# does the model meet assumptions?
plot(mod)

# summarise the model
summary(mod)
```

## ANOVA: ASSESSING A FITTED MODEL







# ANOVA: INTERPRETING A FITTED MODEL

```
##
## Call:
## lm(formula = log(precipitation) ~ mountain_range)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -1.76554 -0.31034  0.01769  0.32208  1.18833
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)      4.57471    0.01375 332.777 < 2e-16 ***
## mountain_rangeShoshone 0.07717    0.02248   3.433 0.000602 ***
## mountain_rangeToiyabe  0.16431    0.01793   9.165 < 2e-16 ***
## mountain_rangeToquima  0.21755    0.02125  10.237 < 2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.4658 on 4294 degrees of freedom
## Multiple R-squared:  0.03001,    Adjusted R-squared:  0.02933
## F-statistic: 44.28 on 3 and 4294 DF,  p-value: < 2.2e-16
```

## **ANOVA: INTERPRETING A FITTED MODEL**

- now we have lots of p-values. . .
- can use *post hoc* tests but not universally accepted
- can pre-specify *contrasts* for specific hypotheses

## **ANOVA: PRESENTING A FITTED MODEL**

- is the model adequate? (assumptions, diagnostics)
- does the model fit the data? (diagnostics,  $r^2$ )
- is the model statistically meaningful? (p-values, test statistics)
- is the model actually meaningful? (parameter estimates)
- can I see it? (boxplots)

## ASIDE: DISCRETE PREDICTOR WITH TWO LEVELS

- special case: t-test (it's still an ANOVA)

```
mod <- t.test(response ~ predictor, data = data_set)
summary(mod)
```

```
##
##  Welch Two Sample t-test
##
## data:  precipitation by region
## t = -4.2244, df = 4284.3, p-value = 2.446e-05
## alternative hypothesis: true difference in means is not equal to 0
## 95 percent confidence interval:
##  -10.652670  -3.899172
## sample estimates:
## mean in group East mean in group West
##           117.5083           124.7842
```

# **GENERAL LINEAR MODELS**

- linear regression, ANOVA, t-test: they're all the same
- just needs a special setup for discrete predictors

# MATRIX NOTATION

- $y_i = \alpha + \beta^T \mathbf{x}_i + \epsilon_i$

$$\beta = \begin{pmatrix} \beta_1 \\ \vdots \\ \beta_k \end{pmatrix}; \mathbf{x}_i = \begin{pmatrix} x_{i,1} \\ \vdots \\ x_{i,k} \end{pmatrix}$$

$$\beta^T = ( \beta_1 \quad \dots \quad \beta_k )$$

## ANOVA: HOW DOES THIS WORK?

- code the  $x_{i,k}$  values as 1 or 0

$$\beta^T = ( \beta_1 \quad \beta_2 \quad \beta_3 )$$

$$\mathbf{x}_i = \begin{pmatrix} 0 \\ 1 \\ 0 \end{pmatrix}$$

$$\beta^T \mathbf{x}_i = ( 0 \quad \beta_2 \quad 0 )$$



## ALL TOO MUCH?

- R will do this for you! (this is one reason R has factors)

```
model.matrix( ~ discrete_predictor)
```

```
##      (Intercept) mountain_rangeShoshone mountain_rangeToiyabe mountain_rangeToquima
## [1,]           1              0              0              0
## [2,]           1              1              0              0
## [3,]           1              0              0              0
## [4,]           1              0              0              1
## [5,]           1              0              1              0
```

## MORE THAN ONE PREDICTOR

- same setup, but now the  $\mathbf{x}_i$  values don't have to be 0 or 1

$$\beta^T = ( \beta_1 \quad \beta_2 \quad \beta_3 )$$

$$\mathbf{x}_i = \begin{pmatrix} x_{i,1} \\ x_{i,2} \\ x_{i,3} \end{pmatrix}$$

$$\beta^T \mathbf{x}_i = ( \beta_1 x_{i,1} \quad \beta_2 x_{i,2} \quad \beta_3 x_{i,3} )$$

# MORE THAN ONE PREDICTOR

```
model.matrix( ~ predictor1 + predictor2 + predictor3)
```

```
##      (Intercept) predictor1 predictor2 predictor3
## [1,]           1    2285.1      78.99      20.8
## [2,]           1    2304.6      67.31      17.1
## [3,]           1    2330.1      64.27      16.7
## [4,]           1    2589.2     144.02      15.8
## [5,]           1    3016.6     235.46       6.2
```

# MORE THAN ONE PREDICTOR

- the scale of the variables matters
- good to standardise continuous predictors

```
# standardise continuous predictors  
predictors_std <- scale(predictors)
```

```
##      predictor1 predictor2 predictor3  
## [1,] -0.7065051 -0.5328227  1.00678496  
## [2,] -0.6438887 -0.6923145  0.32702139  
## [3,] -0.5620058 -0.7338261  0.25353344  
## [4,]  0.2699889  0.3551696  0.08818554  
## [5,]  1.6424108  1.6037937 -1.67552534
```

# CONTINUOUS AND DISCRETE PREDICTORS

- can include continuous and discrete predictors in one model

```
mod <- lm(response ~ continuous1 + continuous2 + discrete)
```

```
##      (Intercept) continuous1 continuous2 discretel1 discrete2 discrete3
## [1,]           1      2285.1       78.99           0           0           0
## [2,]           1      2304.6       67.31           1           0           0
## [3,]           1      2330.1       64.27           0           0           0
## [4,]           1      2589.2      144.02           0           0           1
## [5,]           1      3016.6      235.46           0           1           0
```

## MULTIPLE PREDICTORS: NEW ASSUMPTIONS

- all the same assumptions as before
- plus: predictors are assumed to be independent(ish)
  - technical term: *multicollinearity*
- if two predictors are highly correlated the model can't tell them apart
- how to address issues: careful predictor choice, remove correlated predictors

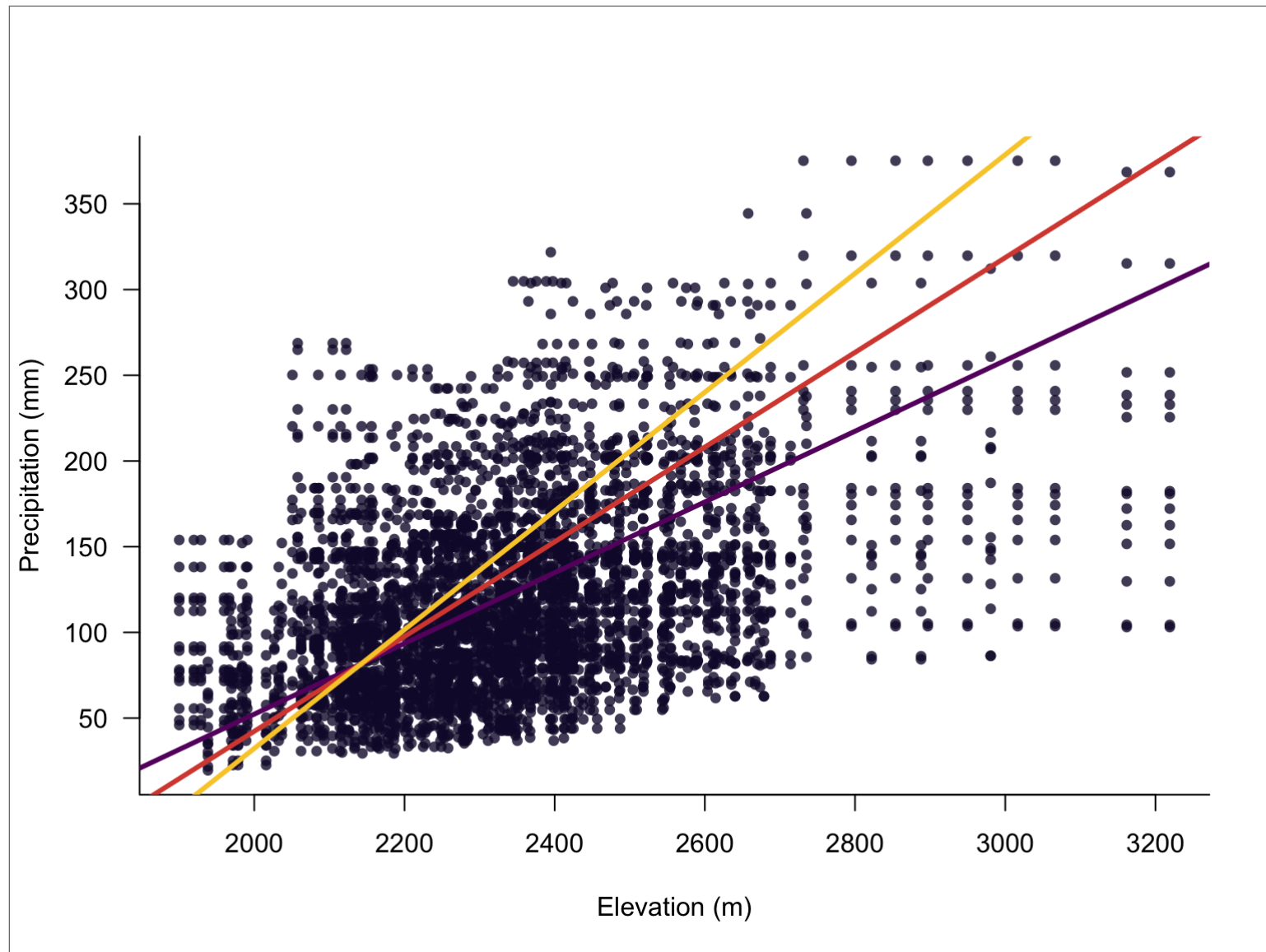
# MULTICOLLINEARITY IN R

```
round(cor(predictor_variables), 2)
```

```
##           predictor1 predictor2 predictor3 predictor4
## predictor1         1.00         0.34         0.43        -0.53
## predictor2         0.34         1.00         0.20        -0.22
## predictor3         0.43         0.20         1.00        -0.52
## predictor4        -0.53        -0.22        -0.52         1.00
```

- solution: remove variables until none are highly correlated
  - removing predictor4 is a good option here

# MULTIPLE PREDICTORS: INTERACTIONS







# MULTIPLE PREDICTORS: INTERACTIONS

```
mod <- lm(precipitation ~ elevation * mountain_range)
summary(mod)
```

```
##
## Call:
## lm(formula = precipitation ~ elevation * mountain_range)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -120.75  -35.97   -9.48   30.03  202.73
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)   -2.104e+02  3.185e+01  -6.607  4.4e-11 ***
## elevation      1.364e-01  1.366e-02   9.981  < 2e-16 ***
## mountain_rangeShoshone  1.842e+01  4.067e+01   0.453  0.65053
## mountain_rangeToiyabe   9.473e+01  3.346e+01   2.832  0.00465 **
## mountain_rangeToquima   1.798e+01  4.192e+01   0.429  0.66795
## elevation:mountain_rangeShoshone -3.521e-03  1.747e-02  -0.202  0.84028
## elevation:mountain_rangeToiyabe  -3.089e-02  1.435e-02  -2.152  0.03145 *
## elevation:mountain_rangeToquima  -2.771e-03  1.767e-02  -0.157  0.87538
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 50.62 on 4290 degrees of freedom
## Multiple R-squared:  0.2096, Adjusted R-squared:  0.2083
## F-statistic: 162.5 on 7 and 4290 DF, p-value: < 2.2e-16
```



## **MULTIPLE PREDICTORS: INTERACTIONS**

- difficult to interpret coefficients
  - effect of one depends on value of the other
  - particularly hard if both are continuous
- it is possible to include higher-order interactions
  - even more difficult to interpret

