Graduate Statistics, 541A, Spring 2019, USC		Instructor:	Steven Heilman		
Name:	USC ID:	Date:			
Signature: By signing here, I certify that I have taken this test while refraining from cheating.)					

Mid-Term 2

This exam contains 7 pages (including this cover page) and 4 problems. Check to see if any pages are missing. Enter all requested information on the top of this page.

You may *not* use your book or any calculator on this exam. You *cannot* use your homeworks. You are required to show your work on each problem on the exam. The following rules apply:

- You have 50 minutes to complete the exam, starting at the beginning of class.
- If you use a theorem or proposition from class or the notes or the book you must indicate this and explain why the theorem may be applied. It is okay to just say, "by some theorem/proposition from class."
- Organize your work, in a reasonably neat and coherent way, in the space provided. Work scattered all over the page without a clear ordering will receive very little credit.
- Mysterious or unsupported answers will not receive full credit. A correct answer, unsupported by calculations, explanation, or algebraic work will receive no credit; an incorrect answer supported by substantially correct calculations and explanations might still receive partial credit.
- If you need more space, use the back of the pages; clearly indicate when you have done this. Scratch paper appears at the end of the document.

Do not write in the table to the right. Good luck!^a

Problem	Points	Score
1	10	
2	10	
3	10	
4	10	
Total:	40	

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Reference sheet

Below are some definitions that may be relevant.

We say that a sequence of random variables $X_1, X_2, \ldots : \Omega \to \mathbf{R}$ converges in probability to a random variable $X : \Omega \to \mathbf{R}$ if: for all $\varepsilon > 0$

$$\lim_{n \to \infty} \mathbf{P}(|X_n - X| > \varepsilon) = 0.$$

We say that a sequence of real-valued random variables X_1, X_2, \ldots converges in distribution to a real-valued random variable X if, for any $t \in \mathbf{R}$ such that $\mathbf{P}(X \leq t)$ is continuous at t,

$$\lim_{n \to \infty} \mathbf{P}(X_n \le t) = \mathbf{P}(X \le t).$$

We say that a sequence of random variables $X_1, X_2, \ldots : \Omega \to \mathbf{R}$ converges in L_2 to a random variable $X : \Omega \to \mathbf{R}$ if

$$\lim_{n \to \infty} \mathbf{E} \left| X_n - X \right|^2 = 0.$$

We say that a sequence of random variables $X_1, X_2, \ldots : \Omega \to \mathbf{R}$ converges almost surely to a random variable $X : \Omega \to \mathbf{R}$ if

$$\mathbf{P}(\lim_{n\to\infty} X_n = X) = 1.$$

Suppose $X = (X_1, ..., X_n)$ is a random sample of size n from a distribution f where $f \in \{f_\theta : \theta \in \Theta\}$ is a family of densities (such as an exponential family). Let $t : \mathbf{R}^n \to \mathbf{R}^k$, so that $Y := t(X_1, ..., X_n)$ is a statistic.

We say that Y is a **sufficient statistic** for θ if, for every $y \in \mathbf{R}^k$ and for every $\theta \in \Theta$, the conditional distribution of (X_1, \ldots, X_n) given Y = y (with respect to probabilities given by f_{θ}) does not depend on θ .

We say Y is **minimal sufficient** for θ if Y is sufficient for θ and, for every statistic $Z \colon \Omega \to \mathbf{R}^m$ that is sufficient for θ , there exists a function $r \colon \mathbf{R}^m \to \mathbf{R}^k$ such that Y = r(Z).

We say Y is **complete** for $\{f_{\theta} \colon \theta \in \Theta\}$ if the following holds:

For any $f: \mathbf{R}^m \to \mathbf{R}$ such that $\mathbf{E}_{\theta} f(Y) = 0 \quad \forall \theta \in \Theta$, it holds that f(Y) = 0.

We say Y is **ancillary** for θ if the distribution of Y does not depend on θ .

Let $X, Y, Z: \Omega \to \mathbf{R}$ be discrete or continuous random variables. Let A be the range of Y. Define $g: A \to \mathbf{R}$ by $g(y) := \mathbf{E}(X|Y=y)$, for any $y \in A$. We then define the **conditional expectation** of X given Y, denoted $\mathbf{E}(X|Y)$, to be the random variable g(Y).

- 1. Label the following statements as TRUE or FALSE. If the statement is true, explain your reasoning. If the statement is false, provide a counterexample and explain your reasoning. In any case, you must **justify your answer** as stated on the front cover page of this exam.
 - (a) (2 points) Let $n \geq 2$ be an integer. Let X_1, \ldots, X_n be a random sample of size n from the Gaussian distribution with mean $\mu \in \mathbf{R}$ and variance $\sigma^2 > 0$. Let $\overline{X} := \frac{1}{n} \sum_{i=1}^n X_i$. and let $S := \sqrt{\frac{1}{n-1} \sum_{i=1}^n (X_i \overline{X})^2}$. Then \overline{X} and S are independent random variables.

TRUE FALSE (circle one)

(b) (2 points) A complete sufficient statistic always exists. (In your answer here, you are allowed to cite a result from the homework, though try to be specific in your response.)

TRUE FALSE (circle one)

(c) (2 points) A constant function is both ancillary and complete.

TRUE FALSE (circle one)

(d) (2 points) A complete sufficient statistic is minimal sufficient.

TRUE FALSE (circle one)

(e) (2 points) Let X_1, \ldots, X_n be a random sample of size n from the Gaussian distribution with mean $\mu \in \mathbf{R}$ and variance $\sigma = 1$, so that μ is unknown. Let Y be a complete sufficient statistic for μ . Let Z be an ancillary statistic for μ . Then Y and Z are independent.

TRUE FALSE (circle one)

2.	(10 points) that Y is no	Give an example of a statistic ot sufficient.	Y that is complete and nonconstant, but such

- 3. (10 points) Suppose X_1, \ldots, X_n is a random sample of size n from the Gaussian distribution with unknown mean $\mu \in \mathbf{R}$ and unknown variance $\sigma^2 > 0$. You may freely use that the sample mean \overline{X} is UMVU for μ and (\overline{X}, S^2) is complete sufficient for (μ, σ^2) , where $S^2 := \frac{1}{n-1} \sum_{i=1}^n (X_i \overline{X})^2$.
 - Fix $\sigma^2 > 0$. Give an explicit expression for a statistic Y that is UMVU for μ^2 , and prove that Y is UMVU for μ^2 .

4. (10 points) Prove the Rao-Blackwell Theorem:

Let Z be a sufficient statistic for $\{f_{\theta} \colon \theta \in \Theta\}$ and let Y be an estimator for $g(\theta)$. Define $W := \mathbf{E}_{\theta}(Y|Z)$. Let $\theta \in \Theta$ with $r(\theta, Y) < \infty$ and such that $\ell(\theta, y)$ is convex in $y \in \mathbf{R}$. Then

$$r(\theta, W) \le r(\theta, Y).$$

(Recall that $\ell \colon \Theta \times \mathbf{R}^k \to \mathbf{R}$, and $r(\theta, Y) := \mathbf{E}_{\theta} \ell(\theta, Y)$.)

(In your response, you can cite results from a homework, if you want.)

(Scratch paper)