

Regression Discontinuity Designs¹

Prepared for the PPRN Methods Masterclass

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March 12, 2021
<https://ucl.zoom.us>

¹NSF support through grants [SES-1357561](#), [SES-1459931](#), [SES-1947805](#), [SES-2019432](#).

Overview

Part 1: Introduction to Regression Discontinuity Designs

Part 2: Graphical illustration of RD models

Part 3: RD Estimation and Inference

3(a): Continuity-based Estimation and Inference

3(b): Local Randomization Estimation and Inference

Part 4: RD Falsification Analysis

Part 1:

Introduction to Regression Discontinuity Designs

Causal Inference and Program Evaluation

- Main goal: learn about treatment effect of policy or intervention
- If treatment randomization available → easy to estimate effects
- If treatment randomization not available → observational studies
 - ▶ Selection on observables
 - ▶ Instrumental variables, etc.
- **Regression discontinuity (RD) design**
 - ▶ Simple assignment, based on known external factors
 - ▶ Objective basis to evaluate assumptions
 - ▶ *Careful*: very local!

Regression Discontinuity Design

Defined by the triplet: score, cutoff, treatment.

- Units receive a **score**.
- A treatment is assigned based on the score and a *known* **cutoff**.
- The **treatment** is:
 - ▶ given to units whose score is greater than the cutoff.
 - ▶ withheld from units whose score is less than the cutoff.
- Under assumptions, the abrupt change in the probability of treatment assignment allows us to learn about the effect of the treatment.

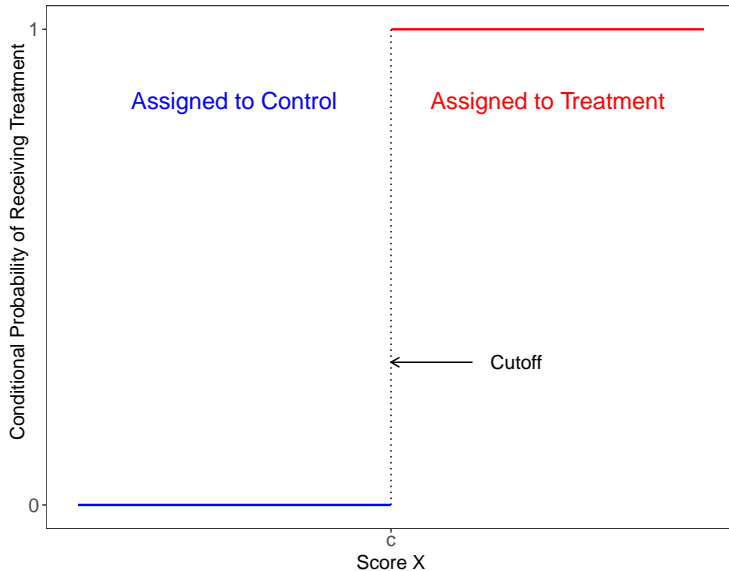
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- **Some examples:**

	X_i	Y_i
Education:	entry test score	test score, enrollment, performance, etc
Development:	pov index	educ, labor, health, etc
Health:	age / birthdate	insurance coverage, mortality, etc.

Treatment Assignment in (Sharp) RD Design



Sharp Regression Discontinuity Design

- n units, indexed by $i = 1, 2, \dots, n$
- Unit's score is X_i , treatment is $T_i = \mathbf{1}(X_i \geq c)$
- Each unit has two potential outcomes:

$Y_i(1)$: outcome that would be observed if i received treatment

$Y_i(0)$: outcome that would be observed if i received control

- The *observed* outcome is

$$Y_i = \begin{cases} Y_i(0) & \text{if } X_i < c, \\ Y_i(1) & \text{if } X_i \geq c. \end{cases}$$

- **Fundamental problem of causal inference**: only observe $Y_i(0)$ for units below cutoff and only observe $Y_i(1)$ for units above cutoff

RD Designs: Taxonomy

- **Frameworks.**

- ▶ Identification: Continuity/Extrapolation, Local Randomization.
- ▶ Score: Continuous, Many Repeated, Few Repeated.

- **Settings.**

- ▶ Sharp, Fuzzy, Kink, Kink Fuzzy.
- ▶ Multiple Cutoff, Multiple Scores, Geographic RD.
- ▶ Dynamic, Continuous Treatments, Time, etc.

- **Parameters of Interests.**

- ▶ Average Effects, Quantile/Distributional Effects, Partial Effects.
- ▶ Heterogeneity, Covariate-Adjustment, Differences, Time.
- ▶ Extrapolation.

RCTs

- **Notation:** $(Y_i(0), Y_i(1), X_i), i = 1, 2, \dots, n$.
- **Treatment:** $T_i \in \{0, 1\}$, T_i independent of $(Y_i(0), Y_i(1), X_i)$.
- **Data:** $(Y_i, T_i, X_i), i = 1, 2, \dots, n$, with

$$Y_i = \begin{cases} Y_i(0) & \text{if } T_i = 0 \\ Y_i(1) & \text{if } T_i = 1 \end{cases}$$

- **Average Treatment Effect:**

$$\tau_{\text{ATE}} = \mathbb{E}[Y_i(1) - Y_i(0)] = \mathbb{E}[Y_i|T = 1] - \mathbb{E}[Y_i|T = 0]$$

(Sharp) RD Designs

- **Notation:** $(Y_i(0), Y_i(1), X_i)$, $i = 1, 2, \dots, n$, X_i score.
- **Treatment:** $T_i \in \{0, 1\}$, $T_i = \mathbb{1}(X_i \geq c)$, c cutoff.
- **Data:** (Y_i, T_i, X_i) , $i = 1, 2, \dots, n$, with

$$Y_i = \begin{cases} Y_i(0) & \text{if } T_i = 0 \\ Y_i(1) & \text{if } T_i = 1 \end{cases}$$

- **Average Treatment Effect at the cutoff** (Continuity-based):

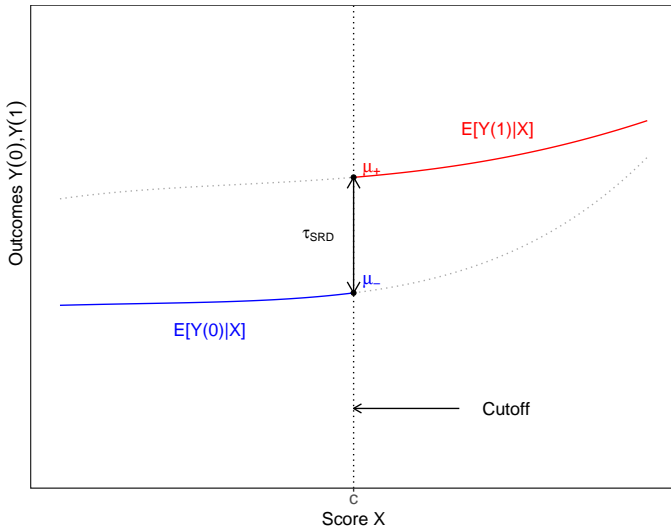
$$\tau_{\text{SRD}} = \mathbb{E}[Y_i(1) - Y_i(0) | X_i = c] = \lim_{x \downarrow c} \mathbb{E}[Y_i | X_i = x] - \lim_{x \uparrow c} \mathbb{E}[Y_i | X_i = x]$$

- **Average Treatment Effect in a neighborhood** (LR-based):

$$\tau_{\text{LR}} = \mathbb{E}[Y_i(1) - Y_i(0) | X_i \in \mathcal{W}] = \frac{1}{N_1} \sum_{X_i \in \mathcal{W}, T_i=1} Y_i - \frac{1}{N_0} \sum_{X_i \in \mathcal{W}, T_i=0} Y_i$$

Continuity-based RD

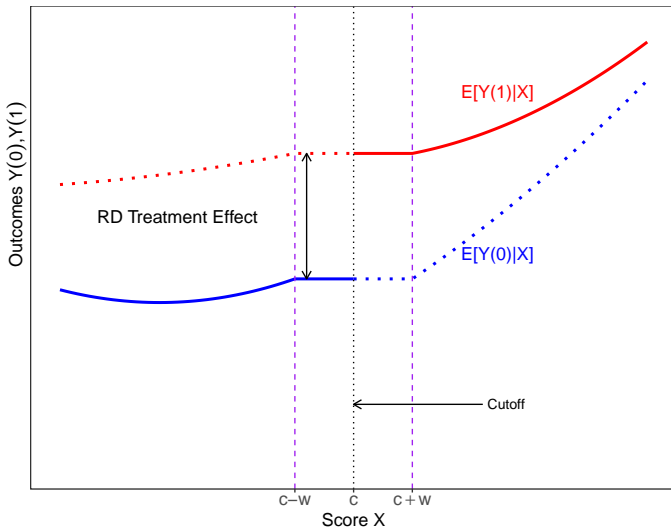
$$\tau_{\text{SRD}} = \underbrace{\mathbb{E}[Y_i(1) - Y_i(0)|X_i = c]}_{\text{Unobservable}} = \underbrace{\lim_{x \downarrow c} \mathbb{E}[Y_i|X_i = x]}_{\text{Estimable}} - \underbrace{\lim_{x \uparrow c} \mathbb{E}[Y_i|X_i = x]}_{\text{Estimable}}$$



Local Randomization RD

T_i independent of $(Y_i(0), Y_i(1))$ for all $X_i \in \mathcal{W} = [c - w, c + w]$

+ exclusion restriction



Part 3:

Graphical illustration of RD models

RD plots

- Main ingredients:
 - ▶ Global fit: smooth approximation to the unknown regression functions
4th or 5th order polynomials, separately above and below the cutoff.
 - ▶ Local sample means:
disjoint intervals (bins) of score, calculating mean of outcome within each bin.
- Main goals:
 - ▶ Graphical (heuristic) representation.
 - ▶ Detection of discontinuities.
 - ▶ Representation of variability.
- Tuning parameters:
 - ▶ Global polynomial degree.
 - ▶ Location (ES or QS) and number of bins.
- **Great to convey ideas but horrible to draw conclusions.**

Example: Incumbency Advantage in U.S. Senate

- **Problem:** incumbency advantage (U.S. senate).

- **Data:**

Y_i = Democratic election outcome at $t + 1$.

T_i = whether Democratic party wins election at t .

X_i = margin of victory at t ($c = 0$).

Z_i = covariates (*demvoteshlag1*, *demvoteshlag2*, *dopen*, etc.).

- **Potential outcomes:**

$Y_i(0)$ = election outcome at $t + 1$ if **had not been** incumbent.

$Y_i(1)$ = election outcome at $t + 1$ if **had been** incumbent.

- **Causal Inference:**

$$Y_i(0) \neq Y_i|T_i = 0 \quad \text{and} \quad Y_i(1) \neq Y_i|T_i = 1$$

- Cattaneo, Frandsen & Titiunik (2015, JCI).

RD Packages

<https://rdpackages.github.io/>

- `rdrobust`: estimation, inference and graphical presentation using local polynomials, partitioning, and spacings estimators; bandwidth selection.
- `rdlocrand` package: covariate balance, binomial tests, randomization inference methods (window selection & inference).
- `rddensity`: discontinuity in density test at cutoff (a.k.a. manipulation testing) using novel local polynomial density estimator.
- `rdmulti`: RD plots, estimation, inference, and extrapolation with multiple cutoffs and multiple scores.
- `rdpower` : power calculation and sample selection for local polynomial methods.

Part 3:

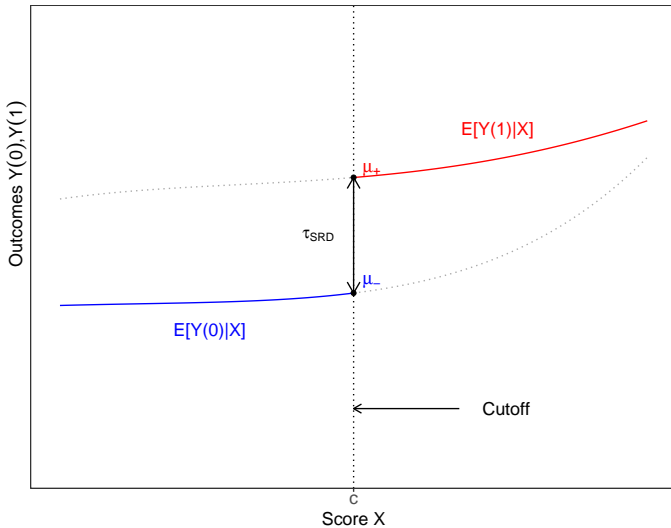
RD Estimation and Inference

Estimation and Inference Methods

- **Continuity/Extrapolation:** Local polynomial approach.
 - ▶ Localization: bandwidth selection (trade-off bias and variance).
 - ▶ Point estimation: “flexible” (nonparametric).
 - ▶ Inference: robust bias-corrected methods.
- **Local Randomization:** finite-sample and large-sample inference.
 - ▶ Localization: window selection (via local independence implications).
 - ▶ Point estimation: parametric, finite-sample (Fisher) or large-sample (Neyman/SP).
 - ▶ Inference: randomization inference (Fisher) or large-sample (Neyman/SP).
- Many refinements and other methods exist (EL, Bayesian, Uniformity, etc.).
 - ▶ Do not offer much improvements in applications.
 - ▶ Can be overly complicated (lack of transparency).
 - ▶ Can depend on user-chosen tuning parameters (lack of replicability).

Continuity-based RD Estimation and Inference

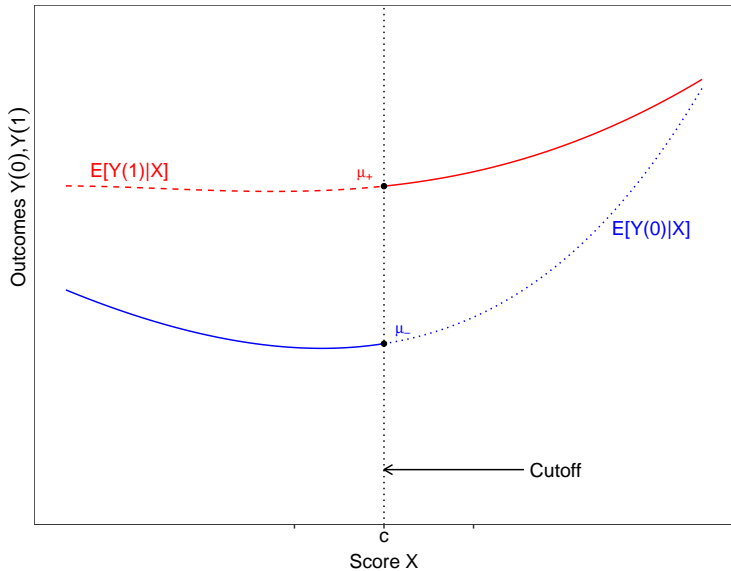
$$\tau = \underbrace{\mathbb{E}[Y_{1i} - Y_{0i} | X_i = x_0]}_{\text{Unobservable}} = \lim_{x \downarrow x_0} \underbrace{\mathbb{E}[Y_i | X_i = x]}_{\text{Observable}} - \lim_{x \uparrow x_0} \underbrace{\mathbb{E}[Y_i | X_i = x]}_{\text{Observable}}$$



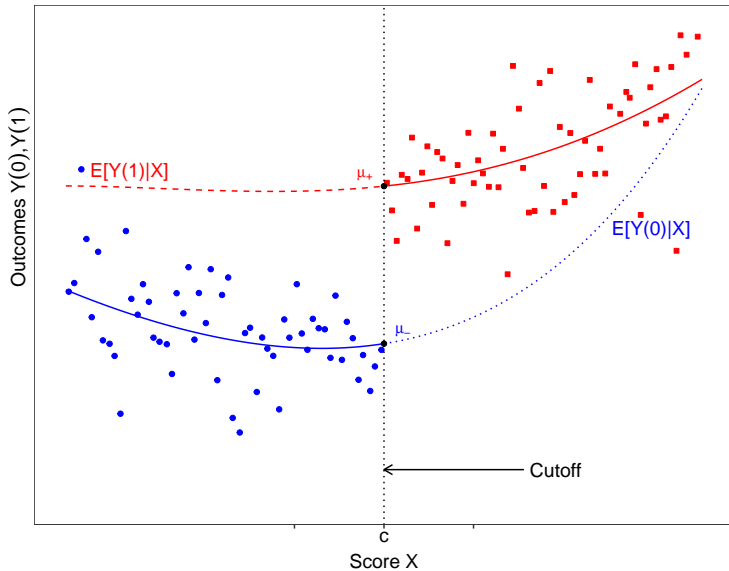
Local Polynomial Treatment Effect Estimation

- $\mathbb{E}[Y_i|X_i = x]$ approximated in neighborhood of x_0 by polynomial function
- Local polynomial estimation:
 - ▶ Choose order of polynomial p
 - ▶ Choose bandwidth h to keep observations in $[x_0 - h, x_0 + h]$
 - ▶ Choose kernel function to weigh observations, $w_i = K(\frac{x_i - x_0}{h})$

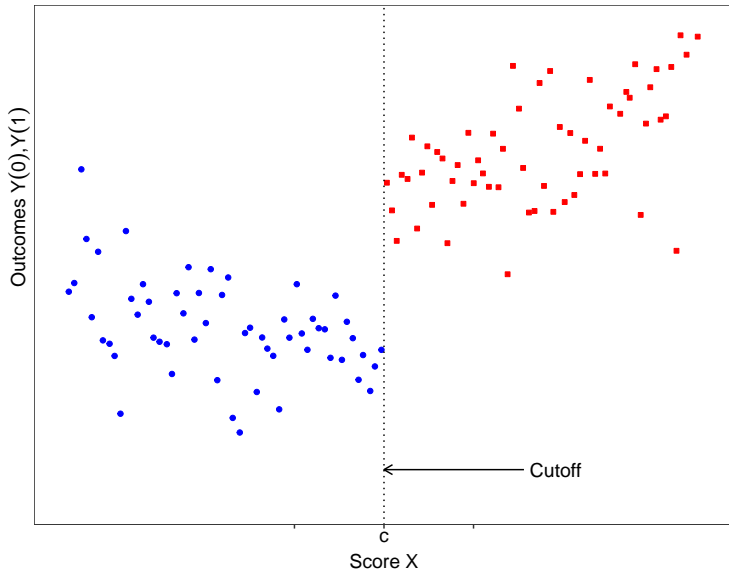
Local Polynomial Estimation



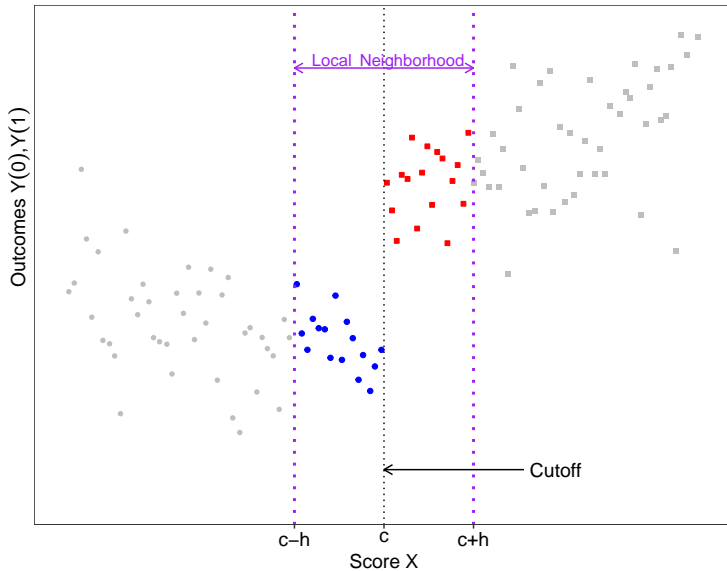
Local Polynomial Estimation



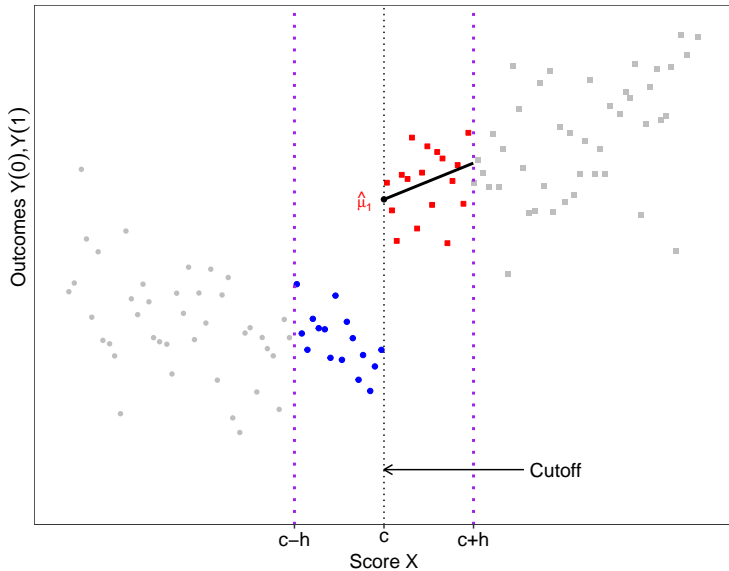
Local Polynomial Estimation



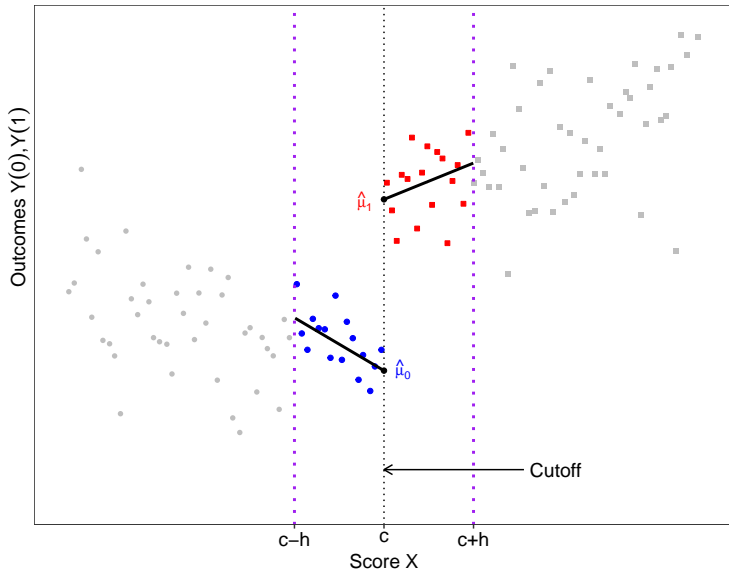
Local Polynomial Estimation



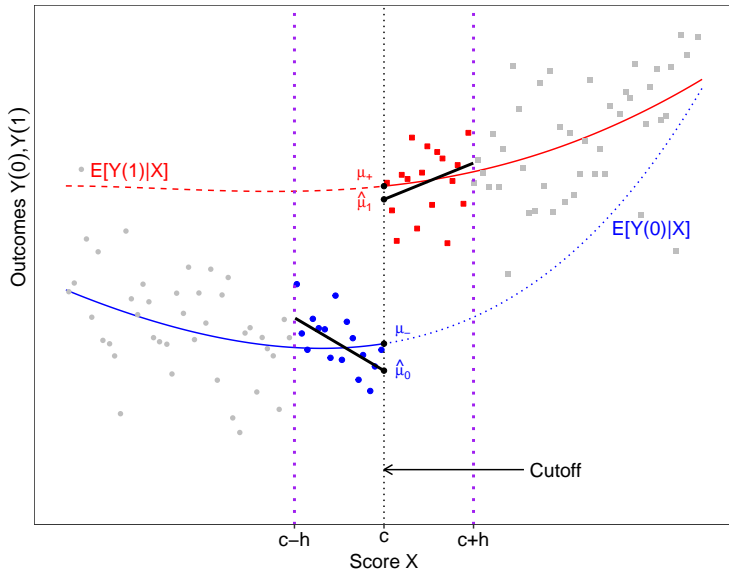
Local Polynomial Estimation



Local Polynomial Estimation



Local Polynomial Estimation



RD Local Polynomial Estimation and Inference

Choose low p and a kernel function $K(\cdot)$



Choose bandwidth h : MSE-optimal or CER-optimal

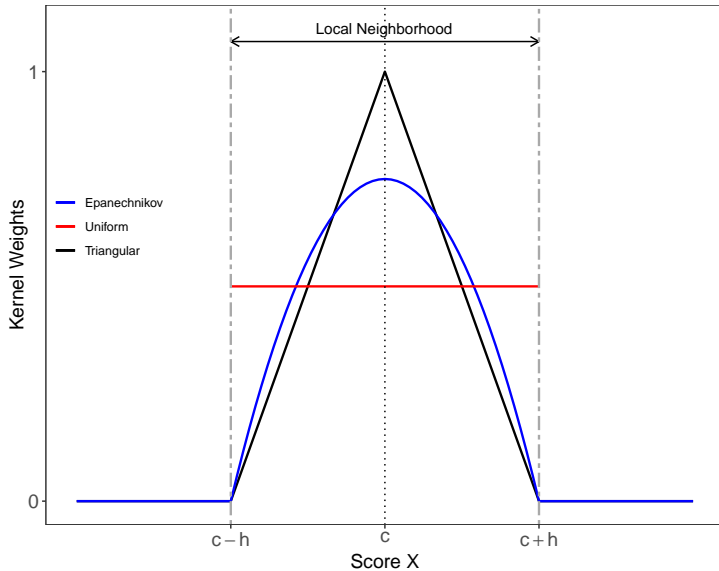


Construct point estimator $\hat{\tau}_n$ (optimal)



Given above steps, how do we make inferences about τ ?

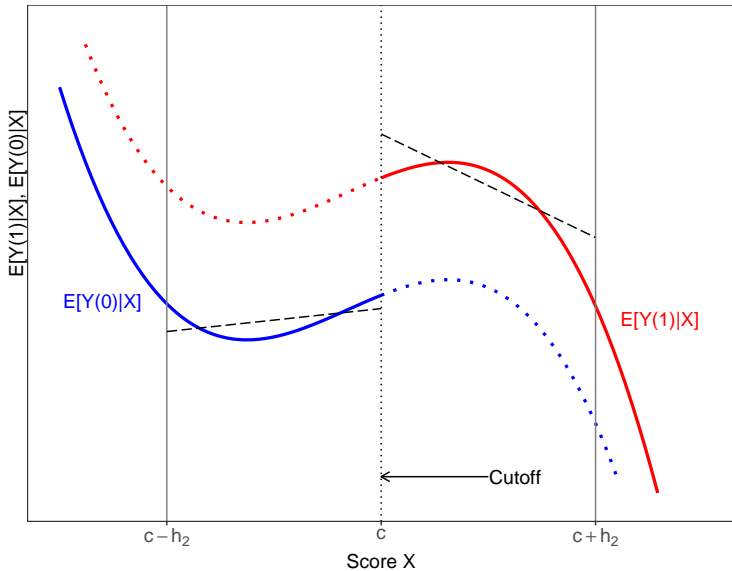
Choice of Kernel Weights



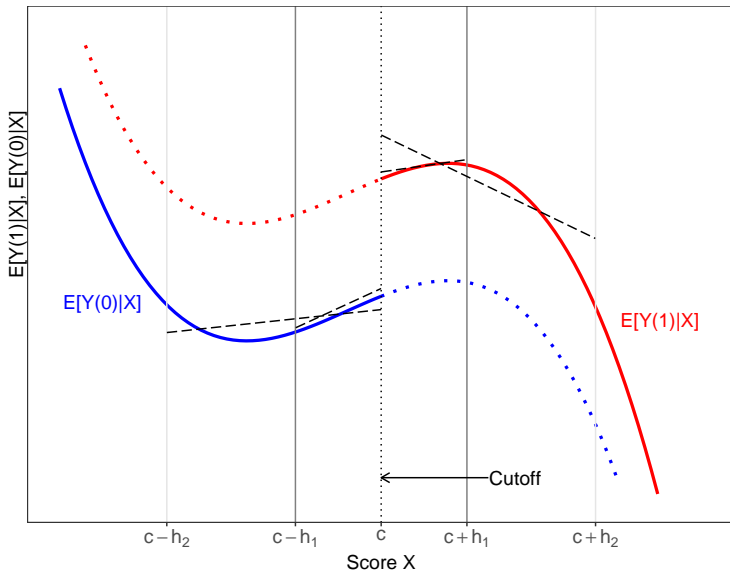
Choice of Polynomial Order (p)

- The higher p , the more flexible the approximation
- However, since approximation is local, p should be low to avoid overfitting
- Given p , approximation can be improved by focusing on a smaller neighborhood around the cutoff
- Standard practice is to choose $p = 1$ (“local linear”)

Approximation for fixed $p = 1$



Approximation for fixed $p = 1$



Choice of Bandwidth

- Given p , find h to ensure optimal properties of the point estimator $\tau_{RD}^{\hat{}}$
- MSE-optimal

$$MSE(\tau_{RD}^{\hat{}}) = Bias^2 + Variance \approx h^{2(p+1)} \mathcal{B}^2 + \frac{1}{nh} \mathcal{V}$$

$$h_{MSE} = C_{MSE}^{1/(2p+3)} \cdot n^{-1/(2p+3)} \quad C_{MSE} = C(K) \cdot \frac{\text{Var}(\hat{\tau}_{SRD})}{\text{Bias}(\hat{\tau}_{SRD})^2}$$

- **Key idea:** trade-off bias and variance of point estimator $\hat{\tau}$

$$\uparrow \text{Bias}(\hat{\tau}) \implies \downarrow \hat{h} \quad \text{and} \quad \uparrow \text{Var}(\hat{\tau}) \implies \uparrow \hat{h}$$

- Coverage Error Rate (CER) optimal

$$h_{CER} = n^{-\frac{p}{(3+p)(3+2p)}} \times h_{MSE}$$

- **Key idea:** choose optimal bandwidth rate to minimize coverage error of the RBC confidence intervals.

Conventional Local Polynomial Point Estimation

- “Local-linear” estimator (w/ weights $K(\cdot)$):

$$\begin{array}{c|c} -h_n \leq X_i < c : & c \leq X_i \leq h_n : \\ Y_i = \alpha_0 + (X_i - c) \cdot \beta_0 + \varepsilon_{0,i} & Y_i = \alpha_1 + (X_i - c) \cdot \beta_1 + \varepsilon_{1,i} \end{array}$$

- RD effect: $\hat{\tau}_n = \hat{\alpha}_1 - \hat{\alpha}_0$
- When choosing MSE-optimal h , this point estimator $\hat{\tau}_n$ is optimal (also consistent)

Conventional Local Polynomial RD Inference

- RD effect: $\hat{\tau}_n = \hat{\alpha}_1 - \hat{\alpha}_0$
- Once $\hat{\tau}_n$ is estimated with optimal h , we might be tempted to use conventional (OLS) inference
- Construct usual t-statistic. For $H_0 : \tau = 0$,

$$T = \frac{\hat{\tau}_n}{\sqrt{V_n}} = \frac{\hat{\alpha}_1 - \hat{\alpha}_0}{\sqrt{V_{1,n} + V_{0,n}}} \rightarrow_d \mathcal{N}(0, 1)$$

- 95% Confidence interval:

$$CI = \left[\hat{\tau}_n \pm 1.96 \cdot \sqrt{V_n} \right]$$

Conventional Local Polynomial RD Inference

- However, with conditions on $h_n \rightarrow 0$, the distributional approximation

$$T = \frac{\hat{\tau}_n}{\sqrt{V_n}} \rightarrow_d \mathcal{N}(B_n, 1) \neq \mathcal{N}(0, 1)$$

- ▶ Bias B_n in RD point estimator captures “curvature” of regression functions
- ▶ In particular, the bias B_n occurs when the MSE-optimal bandwidth is used
- Conventional approach \rightarrow assume bias negligible or undersmoothing

$$T = \frac{\hat{\tau}_n}{\sqrt{V_n}} \rightarrow_d \mathcal{N}(0, 1) \quad \Bigg| \quad \text{CI} = [\hat{\tau}_n \pm 1.96 \cdot \sqrt{V_n}]$$

\implies Not clear guidance & power loss!

- Bias-correction approach

$$T^{\text{bc}} = \frac{\hat{\tau}_n - B_n}{\sqrt{V_n}} \rightarrow_d \mathcal{N}(0, 1) \quad \Bigg| \quad \text{CI}^{\text{bc}} = \left[\left(\hat{\tau}_n - \hat{B}_n \right) \pm 1.96 \cdot \sqrt{V_n} \right]$$

\implies Poor finite sample properties!

Robust Local Polynomial Inference

- Key observation: $\hat{\mathbf{B}}_n$ is constructed to estimate leading bias

$$\mathbf{T}^{\text{bc}} = \frac{\hat{\tau}_n - \hat{\mathbf{B}}_n}{\sqrt{\mathbf{V}_n}} = \underbrace{\frac{\hat{\tau}_n - \mathbf{B}_n}{\sqrt{\mathbf{V}_n}}}_{\rightarrow_d \mathcal{N}(0,1)} + \underbrace{\frac{\mathbf{B}_n - \hat{\mathbf{B}}_n}{\sqrt{\mathbf{V}_n}}}_{\rightarrow_p 0}$$

- Our robust approach** \rightarrow *Non-standard Asymptotics*

$$\mathbf{T}^{\text{bc}} = \frac{\hat{\tau}_n - \hat{\mathbf{B}}_n}{\sqrt{\mathbf{V}_n}} = \underbrace{\frac{\hat{\tau}_n - \mathbf{B}_n}{\sqrt{\mathbf{V}_n}}}_{\rightarrow_d \mathcal{N}(0,1)} + \underbrace{\frac{\mathbf{B}_n - \hat{\mathbf{B}}_n}{\sqrt{\mathbf{V}_n}}}_{\rightarrow_d \mathcal{N}(0,\gamma)}$$

- Robust Bias-Correction Approach:**

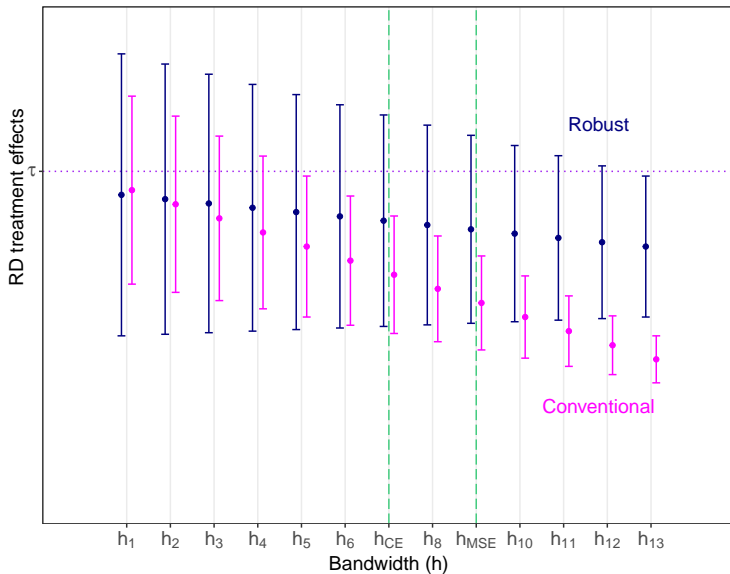
$$\mathbf{T}^{\text{rbc}} = \frac{\hat{\tau}_n - \hat{\mathbf{B}}_n}{\sqrt{\mathbf{V}_n + \mathbf{W}_n}} \rightarrow_d \mathcal{N}(0,1)$$

$$\text{CI}^{\text{rbc}} = \left[\left(\hat{\tau}_n - \hat{\mathbf{B}}_n \right) \pm 1.96 \cdot \sqrt{\mathbf{V}_n + \mathbf{W}_n} \right]$$

Table: Local Polynomial Confidence Intervals

	Centered at	Standard Error
Conventional: CI_{us}	$\hat{\tau}_{\text{SRD}}$	$\sqrt{\hat{\mathcal{V}}}$
Bias Corrected: CI_{bc}	$\hat{\tau}_{\text{SRD}} - \hat{\mathcal{B}}$	$\sqrt{\hat{\mathcal{V}}}$
Robust Bias Corrected: CI_{rbc}	$\hat{\tau}_{\text{SRD}} - \hat{\mathcal{B}}$	$\sqrt{\hat{\mathcal{V}}_{\text{bc}}}$

Confidence Intervals for Different Bandwidths



Local Randomization RD Estimation and Inference

Local Randomization Approach to RD Design

- Gives an alternative that can be used as a robustness check.
- **Key assumption:** exists window $W = [-w, w]$ around cutoff ($-w < c < w$) where (assuming random potential outcomes)

$$T_i \text{ independent of } (Y_i(0), Y_i(1)) \quad (\text{for all } X_i \in W)$$

- Thus, inside W_0 subjects are as-if randomly assigned to either side of cutoff
 - ▶ The distribution of running variable same for all units inside W_0
 - ▶ Potential outcomes in W_0 depend on running variable only through threshold indicators within W_0
- Stronger than Continuity-Based Approach \Rightarrow **Relevant population functions are not only continuous at x_0 , but also completely unaffected by the running variable in W_0**

Local Randomization Framework

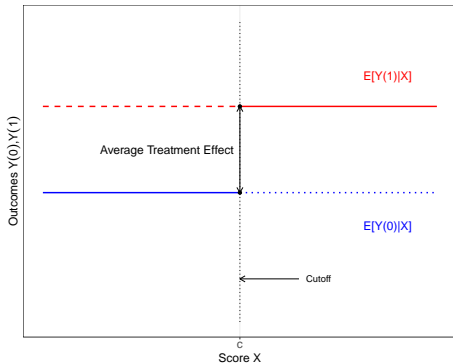
- **Key idea:** treatment assignment as-if randomly assigned “near” cutoff. There exists window $\mathcal{W} = [-w, w]$, with $-w < c < w$, such that

$$\text{for all } X_i \in \mathcal{W} \implies T_i \text{ independent of } (Y_i(0), Y_i(1))$$

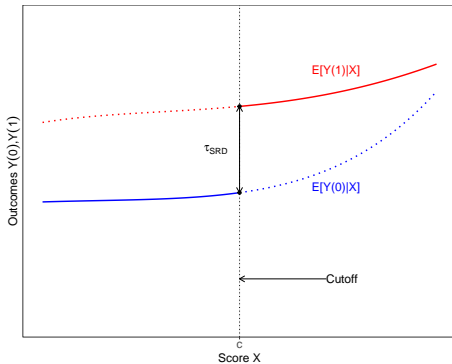
and possibly other conditions hold (e.g., knowledge of assignment mechanism).

- ▶ Conceptually different from continuity/extrapolation based methods.
 - ▶ Challenge: window (neighborhood) selection.
 - ▶ Challenge: small sample (estimation and) inference.
- **Two Steps** (analogous to local polynomial methods):
 1. Select window \mathcal{W} based on “covariate balance” idea.
 2. Given window \mathcal{W} , (estimation and) inference is “standard”: superpopulation, large-samples designed-based methods, randomization inference methods.
 - **Catch:** as-if random assumption good approximation *only very near cutoff*!

Experiment versus RD Design

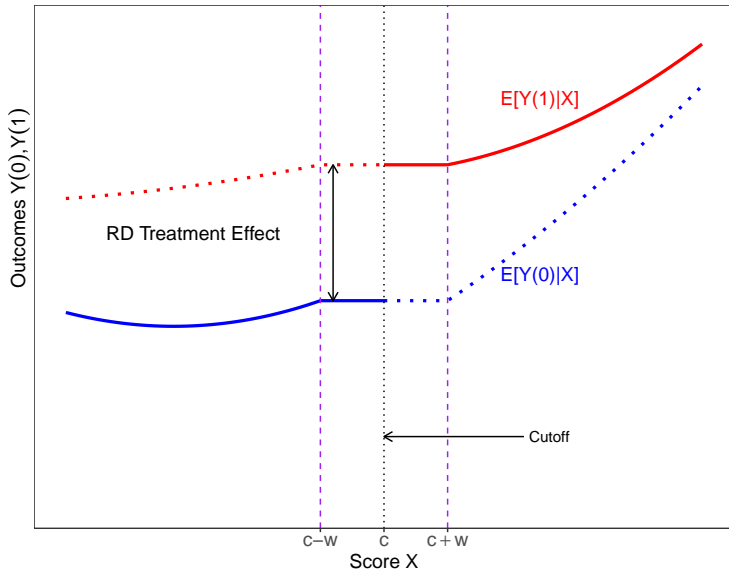


(a) Randomized Experiment



(b) RD Design

If as-if random interpretation is true:
Local Randomization RD

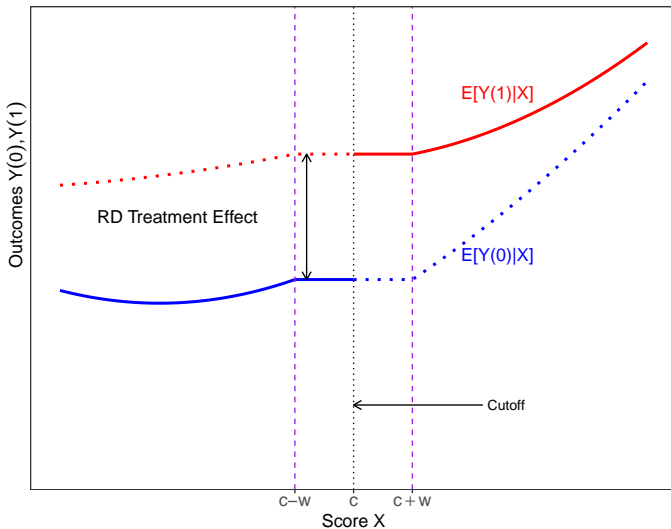


Step 1: Choose the window W_0

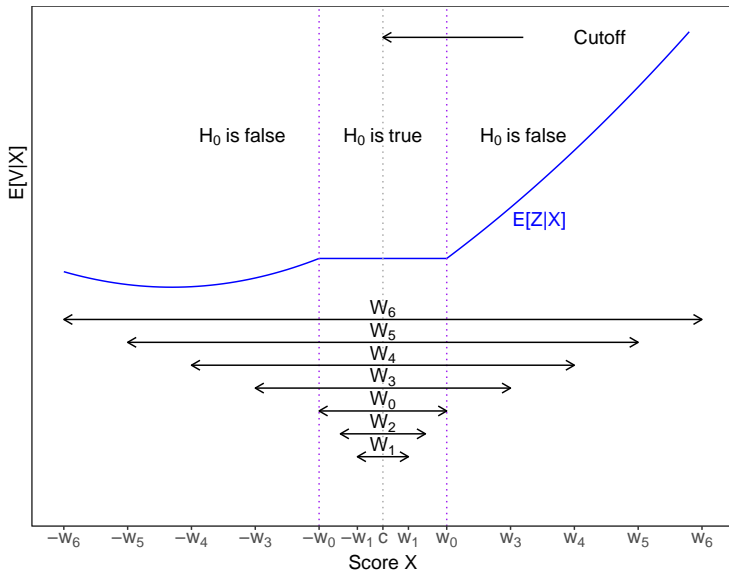
- How to choose \mathcal{W} ?
 - ▶ Find neighborhood where (pre-intervention) covariate-balance holds.
 - ▶ Find neighborhood where outcome and score independent.
 - ▶ Domain-specific or application-specific choice.

T_i independent of $(Y_i(0), Y_i(1))$ for all $X_i \in \mathcal{W} = [c - w, c + w]$

+ exclusion restriction



Window Selector Based on Covariate Balance in Locally Random RD



Step 2: Finite-sample and Large-sample Methods in \mathcal{W}

- Given \mathcal{W} where local randomization holds:
 - ▶ Randomization inference (Fisher): sharp null, finite-sample exact.
 - ▶ Design-based (Neyman): large-sample valid, conservative.
 - ▶ Large-sample standard: random potential outcomes, large-sample valid.
- All methods require window (\mathcal{W}) selection, and choice of statistic.
First two also require choice/assumptions assignment mechanism.
Covariate-adjustments (score or otherwise) possible.

Example: Incumbency Advantage in U.S. Senate

- **Problem:** incumbency advantage (U.S. senate).

- **Data:**

Y_i = election outcome at $t + 1$.

T_i = whether party wins election at t .

X_i = margin of victory at t ($c = 0$).

Z_i = covariates (*demvoteshlag1*, *demvoteshlag2*, *dopen*, etc.).

- **Potential outcomes:**

$Y_i(0)$ = election outcome at $t + 1$ if **had not been** incumbent.

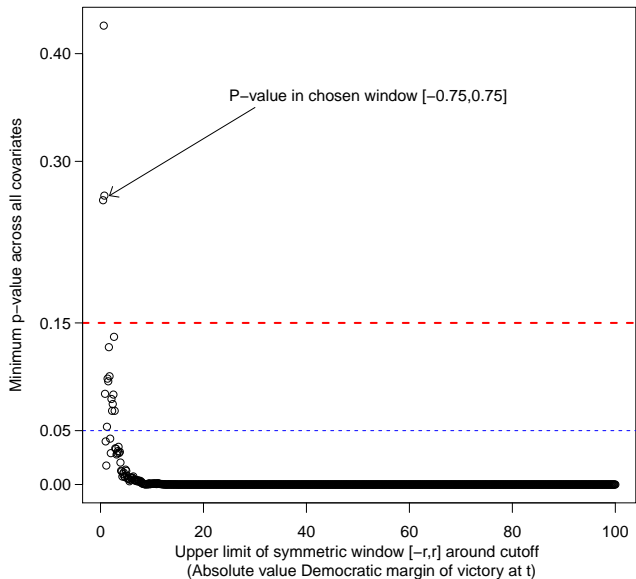
$Y_i(1)$ = election outcome at $t + 1$ if **had been** incumbent.

- **Causal Inference:**

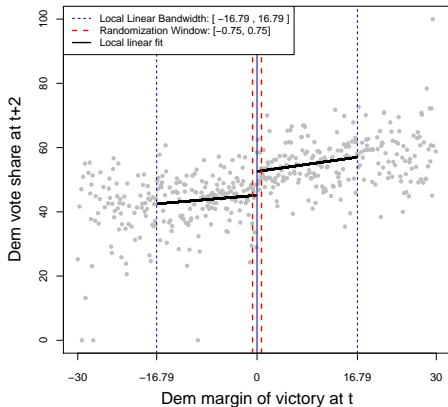
$$Y_i(0) \neq Y_i|T_i = 0 \quad \text{and} \quad Y_i(1) \neq Y_i|T_i = 1$$

- Cattaneo, Frandsen & Titiunik (CFT, 2015, Journal of Causal Inference).

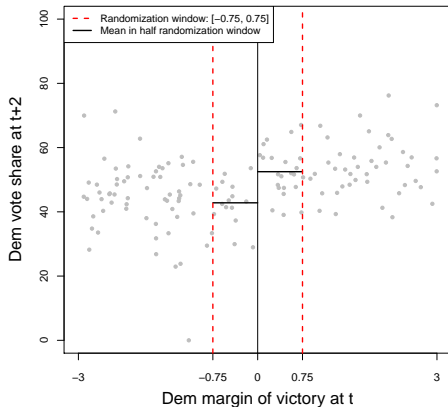
Window Selection Based on Covariates, CFT



Continuity-Based vs Local Randomization Analysis, CFT



(a) Continuity-Based Analysis



(b) Local Randomization Analysis

Part 4:

Falsification Analysis for RD designs

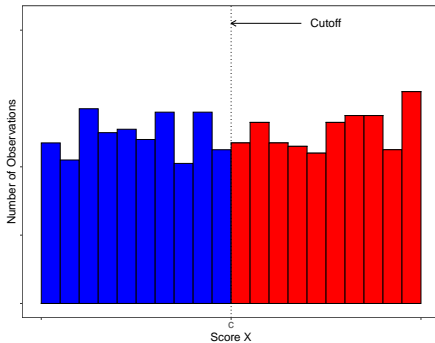
Falsification Methods

- RD rule of treatment assignment is not by itself enough to guarantee that continuity or local randomizations are met
- Qualitative information and quantitative falsification tests play crucial role
 - ▶ Qualitative information: were there mechanisms to appeal score? did people change their score?
 - ▶ Falsification: various statistical tests

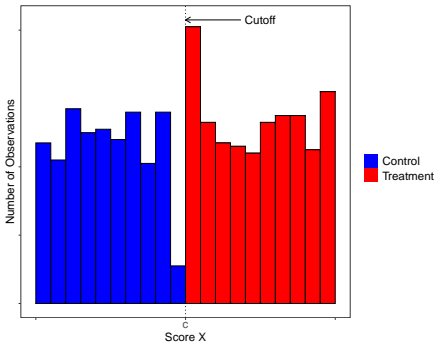
Falsification Methods

- Density test of “sorting”: is number of observations below the cutoff surprisingly different from number of observations above it?
- Treatment effect on
 - ▶ Predetermined covariates
 - ▶ Placebo outcomes
- Also: effect at different cutoffs, effect at different bandwidths, doughnut hole

Falsification Methods: Density Test



(a) No sorting



(b) Sorting

Falsification Methods: Tests on Predetermined Covariates and Placebo Outcomes

- Continuity-based falsification:
 - ▶ Test of continuity of density of the running variable
 - ▶ Local polynomial effects with optimal bandwidth
 - ▶ Robust Inference
 - ▶ CRUCIAL: each covariate/placebo outcome must have its own optimal bandwidth
- Local randomization falsification:
 - ▶ Within chosen window, density test
 - ▶ Test that covariate and placebo outcome distributions are indistinguishable for treated and control
 - ▶ CRUCIAL: all tests are conducted within the same window for each covariate/placebo outcome

Thank you!

<https://rdpackages.github.io/>