Targeted minimum loss-based estimation and double/debiased machine learning

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References

- ► "Targeted Maximum Likelihood Learning" by van der Laan & Rubin (2006)
- "CV-TMLE and double machine learning" in van der Laan's website
- ► "Machine learning in the estimation of causal effects: targeted minimum loss-based estimation and double/debiased machine learning" by Díaz (2020)
- "Double/debiased machine learning for treatment and structural parameters"
 Chernozhukov, Chetverikov, Demirer, Duflo, Hansen, Newey, & Robins (2018)
- "Locally Robust Semiparametric Inference with Debiased GMM" by Chernozhukov, Escanciano, Ichimura, Newey, & Robins (2022)

Regularization Bias

- Variables
 - Y: outcome
 - D: treatment indicator
 - X: covariates
- ▶ Observe $(y_i, d_i, x_i) \sim P_0$ i.i.d. for i = 1, ..., N
- ▶ Easy to get \sqrt{N} -consistent θ_0 if

$$Y = D\theta_0 + X^{\top}\beta_0 + U, E[U|X, D] = 0, \beta_0 \in \mathbb{R}^p, \text{ where } p \text{ is small enough.}$$

What happens if we apply lasso to the following model?

$$Y = D\theta_0 + X^{\mathsf{T}}\beta_0 + U, E[U|X,D] = 0, \beta_0 \in \mathbb{R}^p$$
, where p is very large.

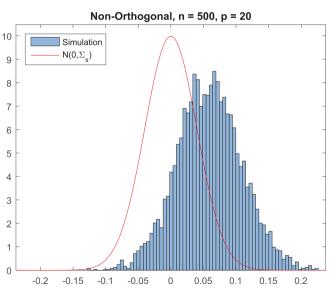
Regularization Bias

What happens if we apply ML prediction approach to the following model?

$$Y = D\theta_0 + g_0(X) + U, E[U|X, D] = 0.$$

- 1. Start from a guess of $\theta_0 \Rightarrow \hat{\theta}^0$
- 2. Apply ML to predict $Y-D\hat{\theta}^0$ using $X\Rightarrow \hat{g}^1(\cdot)$
- 3. Regress $Y \hat{g}^1(X)$ on $D \Rightarrow \hat{\theta}^1$
- 4. Iterate until convergence $\Rightarrow \hat{\theta}_0$

Regularization Bias



Frish-Waugh-Lowell Theorem

Consider

$$Y = D\theta_0 + X^{\mathsf{T}}\beta_0 + U, E[U|X,D] = 0, \beta_0 \in \mathbb{R}^p$$
, where p is small enough.

 θ_0 can be consistently estimated by regressing

ightharpoonup residual of regression Y on X

on

ightharpoonup residual of regression D on X.

Double/Debiased Machine Learning Estimator

What happens if we apply FWL-style estimation to the following?

$$Y = D\theta_0 + X^{\mathsf{T}}\beta_0 + U, E[U|X,D] = 0, \beta_0 \in \mathbb{R}^p$$
, where p is very large.

- 1. Apply lasso to predict D by X, and collect the residual $\Rightarrow \hat{V}$
- 2. Apply lasso to predict Y by X, and collect the residual $\Rightarrow \hat{W}$
- 3. Regress \hat{W} on $\hat{V} \Rightarrow {\sf DML}$ estimator $\hat{\theta}_0$

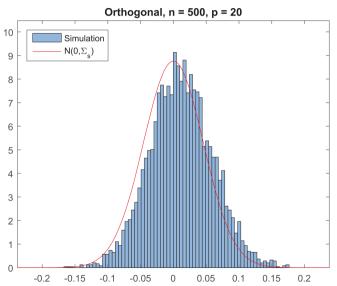
Double/Debiased Machine Learning Estimator

Consider a more general situation:

$$Y = D\theta_0 + g_0(X) + U, E[U|X, D] = 0.$$

- 1. Apply ML to predict D by X, and collect the residual $\Rightarrow \hat{V}$
- 2. Apply ML to predict Y by X, and collect the residual $\Rightarrow \hat{W}$
- 3. Regress \hat{W} on $\hat{V} \Rightarrow {\sf DML}$ estimator $\hat{\theta}_0$

Double/Debiased Machine Learning Estimator

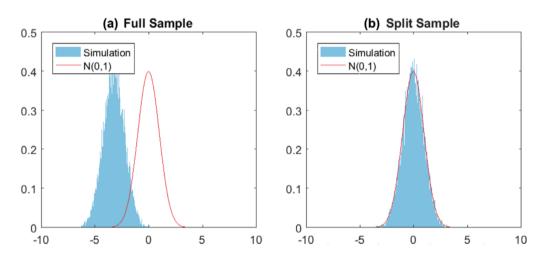


Split Sample

We need to use independent sample sets for implementing

- 1. Estimation for residuals \hat{V} and \hat{W}
- 2. Regression \hat{W} on \hat{V} to get consistency.

Split Sample



Orthogonality: Non-linear Moment Condition

► General non-linear moment condition:

$$E[\psi(W;\theta_0,\eta_0)] = 0.$$

- ▶ In previous examples, W = (Y, D, X) and $\eta_0 = (\beta_0, \gamma_0)$ or (g_0, m_0) .
- Orthogonality condition:

$$\partial_{\eta} E[\psi(W;\theta_0,\eta_0)] = 0.$$

Example: ATE

Consider

$$\begin{cases} Y = g_0(D, X) + U, & E[U|X, D] = 0 \\ D = m_0(X) + V, & E[V|X] = 0 \end{cases}.$$

► Want to estimate ATE:

$$\theta_0 = E[g_0(1, X) - g_0(0, X)].$$

▶ Score can be $\psi(W; \theta, \eta) =$

$$(g(1,X) - g(0,X)) + \frac{D(Y - g(1,X))}{m(X)} - \frac{(1-D)(Y - g(0,X))}{1 - m(X)} - \theta,$$

where $\eta = (g, m)$.

13 / 17

Example: ATE

- 1. Apply ML to predict Y by (D,X), and get $\hat{g}(D,X)$
 - Caution: naive estimator $\hat{\theta} = \sum_{i=1}^{N} \{\hat{g}(1, x_i) \hat{g}(0, x_i)\}$ will be biased!
- 2. Apply ML to predict D by X, and get $\hat{m}(X)$
- 3. DML estimator $\hat{\theta} =$

$$\sum_{i=1}^{N} \left\{ (\hat{g}(1, x_i) - \hat{g}(0, x_i)) + \frac{d_i(y_i - \hat{g}(1, x_i))}{\hat{m}(x_i)} - \frac{(1 - d_i)(y_i - \hat{g}(0, x_i))}{1 - \hat{m}(x_i)} \right\}$$

is unbiased!!

DML estimator coincides with the "doubly robust" estimator for ATE.

Targeted minimum loss-based estimation (TMLE)

- Variables
 - Y: outcome
 - D: treatment indicator
 - X: covariates
- ▶ Observe $(y_i, d_i, x_i) \sim P_0$ i.i.d. for i = 1, ..., N
- ightharpoonup Statistical model \mathcal{M} contains P_0
- ▶ Want to estimate $\Psi(P)$: e.g, ATE

$$\Psi(P) = E_P[E_P(Y|D=1,X) - E_P(Y|D=0,X)]$$

- ► Idea of TMLE:
 - 1. Estimate P_0 by MLE with a correction term $\epsilon \Rightarrow P_{\hat{\epsilon}}$
 - 2. Plug $P_{\hat{\epsilon}}$ into estimator of $\Psi \Rightarrow \mathsf{TMLE}$ estimator $\hat{\Psi}(P_{\hat{\epsilon}})$
- Also use sample splitting.

Example: ATE

- 1. Set statistical model: $P(Y|D,X) \sim \mathcal{N}(m(D,X),\sigma^2(D,X))$
- 2. Obtain correction formula: $P_{\epsilon}(Y|D,X) \sim \mathcal{N}(m(D,X) + \epsilon h(P)(D,X), \sigma^2(D,X))$

where
$$h(P)(D,X) = \left(\frac{\mathbb{1}(D=1)}{E_P[D=1|X]} - \frac{\mathbb{1}(D=0)}{E_P[D=0|X]}\right)\sigma^2(D,X)$$

- Correction is made to satisfy the "efficient influence curve" condition
- 3. Estimate $m(D,X), \sigma^2(D,X), \epsilon$ with MLE by fitting $P_{\epsilon}(Y|D,X)$ to data $\Rightarrow \hat{m}(D,X), \hat{h}(D,X), \hat{\epsilon}$
- 4. $\hat{m}(D,X) + \hat{\epsilon}\hat{h}(D,X)$ is the corrected estimator for $E_P(Y|D,X)$
- 5. TMLE estimator

$$\hat{\Psi}(P_{\hat{\epsilon}}) = \sum_{i=1}^{N} \left[\left\{ \hat{m}(1, X) + \hat{\epsilon} \hat{h}(1, X) \right\} - \left\{ \hat{m}(0, X) + \hat{\epsilon} \hat{h}(0, X) \right\} \right]$$

Difference between TMLE and DML?

- ▶ van der Laan and Díaz argue TMLE is more general framework ⇒ No, they ignore Chernozhukov et al. (2022), which provides a general framework of Chernozhukov et al. (2018)
- ▶ Both are coming from the semiparametric idea of first stage estimator can be high-dimensional when the second stage parameter of interest is low dimension.
- ► They are just following different contexts:
 - Biostatistics, statistical model
 - Econometrics, GMM