Stochastic Local Volatility

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1 The Model

1.1 General description

Let T be a finite time horizon (the final maturity of the deal we want to price, for example) Let S be an asset of drift μ_t under the domestic risk-neutral probability measure. We assume the following dynamics under \mathbb{Q}_d

$$\frac{dS_t}{S_t} = \mu_t dt + \sigma(S_t, t)e^{X_t} dW_t \tag{1}$$

where $(S,t) \mapsto \sigma(S,t)$ is the local volatility function (deterministic function of spot and time, calibrated to the vanilla option market at time 0. We will develop later a methodology to calibrate that function σ . X is an Ornstein-Ulhenbeck process so that:

$$\forall t \in [0, T], \mathbb{E}\left[e^{2X_t}\right] = 1 \tag{2}$$

The goal of the stochastic volatility would be to calibrate to the forward smile (to be consistent with the

1.2 Stochastic volatility part

Let us assume that X follows the following dynamics under the domestic risk-neutral probability measure

$$dX_t = (\theta(t) - \lambda X_t)dt + \sigma_X dW_t^X$$

$$X_0 = 1$$
(3)

The coefficients are kept constant for now for simplicity. We have

$$X_t = e^{-\lambda t} + \int_0^t \theta(s)e^{-\lambda(t-s)}ds + \int_0^t \sigma_X e^{-\lambda(t-s)}dW_s^X$$
(4)

Then,

$$\begin{split} \exp\left(2e^{-\lambda t} + 2\int_0^t \theta(s)e^{-\lambda(t-s)}ds + 2\int_0^t \sigma_X^2 e^{-2\lambda(t-s)}ds\right) &= 1 \\ e^{-\lambda t} + \int_0^t \theta(s)e^{-\lambda(t-s)}ds + \int_0^t \sigma_X^2 e^{-2\lambda(t-s)}ds &= 0 \\ e^{-\lambda t} + \int_0^t \theta(s)e^{\lambda s}dse^{-\lambda t} + \int_0^t \sigma_X^2 e^{-2\lambda(t-s)}ds &= 0 \\ -\lambda e^{-\lambda t} - \int_0^t \theta(s)e^{\lambda s}ds\lambda e^{-\lambda t} + 2\sigma_X^2 e^{-2\lambda t} + \theta(t) &= 0 \end{split}$$

Setting $f(t) = \theta(t)e^{\lambda t}$ it yields

$$\lambda \int_0^t f(s)ds + f(t) = c(t)$$

$$c(t) = \lambda - 2\sigma_X^2 e^{-\lambda t}$$
(5)

Then f satisfies the following ODE

$$f'(t) + \lambda f(t) = -2\sigma_X^2 \lambda e^{-\lambda t}$$

$$f(0) = \lambda - 2\sigma_X^2$$
(6)

Then

$$f(t) = Ce^{-\lambda t} - 2\sigma_X^2 \lambda t e^{-\lambda t}$$

$$= (\lambda - 2\sigma_X^2) e^{-\lambda t} - 2\sigma_X^2 \lambda t e^{-\lambda t}$$

$$f(t) = \lambda e^{-\lambda t} - 2\sigma_X^2 (1 + \lambda t) e^{-\lambda t}$$
(7)

Then

$$\theta(t) = \lambda e^{-2\lambda t} - 2\sigma_X^2 (1 + \lambda t) e^{-2\lambda t} \tag{8}$$

The stochastic volatility X can be computed quite easily and be simulated exactly

$$X_t \sim \mathcal{N}\left(e^{-\lambda t} + \int_0^t \theta(s)e^{-\lambda(t-s)}ds, \sigma^2 \frac{1 - e^{-2\lambda t}}{2\lambda}\right)$$
(9)

1.3 Calibration of the local volatility

Theorem 1.1 (Markovian Projection). Put here the theorem of the markovian projection of general volatility models into a local volatility

In our case, we have

$$\sigma_D^2(x,t) = \mathbb{E}\left[e^{2X_t}\sigma^2(S_t,t)\big|S_t = S\right]$$

$$= \mathbb{E}\left[e^{2X_t}\big|S_t = S\right]\sigma^2(S,t)$$
(10)

1.3.1 Calibration via a PDE

Let us introduce the Green function $G(S_0, X_0, t_0; S, X, t)$. Let us set $b(X) = e^X$. G satisfies

$$\frac{\partial G}{\partial T} - \frac{1}{2} \frac{\partial^2}{\partial S^2} \left(\sigma^2 b^2 S^2 G \right) - \frac{\partial^2}{\partial X^2} \left(\sigma_X^2 G \right) - \frac{\partial^2}{\partial S \partial X} \left(\rho b \sigma S \sigma_X G \right) + \frac{\partial}{\partial X} \left((\theta - \lambda X) G \right) + \frac{\partial}{\partial S} (\mu S G) = 0 \tag{11}$$

Let us assume (in relation of the previous section), that

- 1. σ_X , θ and λ are not functions of X
- 2. μ is not a function of S
- 3. ρ is neither a function of S nor X

The calibration PDE degenerates into

$$\frac{\partial G}{\partial T} - \frac{1}{2}e^{2X}\frac{\partial^{2}}{\partial S^{2}}\left(\sigma^{2}S^{2}G\right) - \sigma_{X}^{2}\frac{\partial^{2}}{\partial X^{2}}\left(G\right) - \rho\frac{\partial^{2}}{\partial S\partial X}\left(b\sigma S\sigma_{X}G\right) + \left(\theta - \lambda X\right)\frac{\partial}{\partial X}\left(G\right) + \mu\frac{\partial}{\partial S}\left(SG\right) = 0 \tag{12}$$

Let us denote by Q the density of S

$$Q(S,t) = \int_{\mathbb{D}} G(S,X,t)dX \tag{13}$$

Then

$$\frac{\partial Q}{\partial t} - \frac{1}{2} \frac{\partial^2}{\partial S^2} \left(\sigma^2 S^2 \int_{\mathbb{R}} e^{2X} G(S, X, t) dX \right) + \mu \frac{\partial}{\partial S} (SG) = 0$$
 (14)

Thus,

$$\frac{\partial Q}{\partial t} - \frac{1}{2} \frac{\partial^2}{\partial S^2} \left(\sigma^2 \frac{\int_{\mathbb{R}} e^{2X} G(S, X, t) dX}{\int_{\mathbb{R}} G(X, S, t) dX} S^2 Q \right) + \mu \frac{\partial}{\partial S} (SG) = 0$$
 (15)

We recognize the Dupire equation which defines a local volatility function $\sigma_D(S,t)$

$$\sigma_D^2(S,t) = \sigma^2(S,t) \frac{\int_{\mathbb{R}} e^{2X} G(S,X,t) dX}{\int_{\mathbb{R}} G(X,S,t) dX}$$
(16)

which permits to re-write the calibration PDE

$$\frac{\partial G}{\partial T} - \frac{1}{2} \frac{\partial^{2}}{\partial S^{2}} \left(S^{2} e^{2X} \sigma_{D}^{2} \frac{\int_{\mathbb{R}} G(S, X, t) dX}{\int_{\mathbb{R}} e^{2X} G(S, X, t) dX} \right) + \mu \frac{\partial}{\partial S} \left(SG \right) - \frac{\sigma_{X}^{2}}{2} \frac{\partial^{2} G}{\partial X^{2}} + \frac{\partial}{\partial X} \left((\theta - \lambda X)G \right) - \rho \sigma_{X} \frac{\partial^{2}}{\partial S \partial X} \left(e^{X} S \sigma_{D} \sqrt{\frac{\int_{\mathbb{R}} G(S, X, t) dX}{\int_{\mathbb{R}} e^{2X} G(S, X, t) dX}} G \right) = 0$$
(17)

The only SLV models tractables with a PDE calibration are then mono-asset SLV models. We will develop in the following another method which enable us to calibrate via Monte-Carlo SLV Models. However, a Monte-Carlo will have more noise than a PDE calibration.

1.3.2 Calibration by Monte-Carlo

In this section, we will develop a method to calibrate the SLV models in a higher dimension problem using a Monte-Carlo simulation.

Let us re-write the Dupire equation associated with the SLV model we just described

$$\frac{\partial C}{\partial T} + P_d(0, T) \mathbb{E}^{\mathbb{Q}_T} \left[\left(K r_d(T) - S_T r_f(T) \right) \mathbf{1}_{\{S_T > K\}} \right] - \frac{1}{2} \sigma_D^2(K, T) K^2 \frac{\partial^2 C}{\partial K^2} = 0$$

$$\frac{\partial C}{\partial T} + P_d(0, T) \mathbb{E}^{\mathbb{Q}_T} \left[\left(K r_d(T) - S_T r_f(T) \right) \mathbf{1}_{\{S_T > K\}} \right] - \frac{1}{2} \sigma^2(K, T) \mathbb{E} \left[e^{2X_T} \middle| S_T = K \right] K^2 \frac{\partial^2 C}{\partial K^2} = 0$$
(18)

To write the calibration equation, we are left with, the calculation of the following integral

$$\mathbb{E}\left[e^{2X_T}\middle|S_T = K\right] = \frac{\mathbb{E}\left[e^{2X_T}\mathbf{1}_{S_T = K}\right]}{\mathbb{P}(S_T = K)} \tag{19}$$

Re-writing the above equation as a function of log-spot $y_t = \log(S_t)$ yields

$$\mathbb{E}\left[e^{2X_T}\middle|y_T=k\right] = \frac{\mathbb{E}\left[e^{2X_T}\mathbf{1}_{y_T=k}\right]}{\mathbb{P}(y_T=k)}$$
(20)

where $k = \log(K)$.

Let us assume that we have an expression for the forward transition probabilities $(y, X, T) \mapsto p(y, X, T) = G(e^y, X, T)$, the two expectations that come into play in equation (20) can be refactored such as

$$\sigma(k,T) = \sigma_D(e^k, T) \sqrt{\frac{I_0(k,T)}{I_1(k,T)}}$$
(21)

where

$$I_n(k,T) = \int e^{nx} p(y,x,T) dx$$
, for $n \in \{0,1\}$ (22)

Two Monte-Carlo method can then be used for the computation of this conditional expectation

- 1. Estimation of the integral via a full Monte-Carlo simulation of (S_T, X_T)
- 2. Approximation of the conditional expectation via a Least-Square approach using well-choosen polynomial functions

Estimation via Least-Squares As we want to compute a S_T -measurable integral, we can approximate

$$\mathbb{E}\left[e^{2X_T}\middle|S_T\right] \approx \sum_{m=1}^M \beta_m \zeta_m\left(S_T\right) \tag{23}$$

where the coefficients $(\beta_1, \ldots, \beta_M)$ are found by a least-square approach. Let us assume we have N realisations of the random variables (S_T, X_T) denoted by $((S_T^1, X_T^1), \ldots, (S_T^N, X_T^N))$

Let us write the minimization problem associated with Least-Square in our case. Let us set $\beta = (\beta_1, \dots, \beta_M)$

$$\hat{\beta} = \underset{\beta \in \mathbb{R}^M}{\operatorname{argmin}} \left\{ \mathbb{E} \left[\left(\mathbb{E} \left[e^{2X_T} \middle| S_T \right] - \sum_{m=1}^M \beta_m \zeta_m \left(S_T \right) \right)^2 \right] \right\}$$
(24)

Denoting by $M_{\zeta\zeta}$ and $M_{\zeta X}$ the following quantities, according to [2]

$$(M_{\zeta\zeta})_{r,s} = \mathbb{E}\left[\zeta_r(S_T)\zeta_s(S_T)\right]$$

$$(M_{\zeta X})_r = \mathbb{E}\left[\zeta_r(S_T)e^{2X_T}\right]$$
(25)

we have the following estimation for the "regression coefficients" β

$$\beta = (M_{\zeta\zeta})^{-1} M_{\zeta X} \tag{26}$$

Those quantities can be estimated using the Monte-Carlo paths $((S_T^1, X_T^1), \dots, (S_T^N, X_T^N))$

$$(\hat{M}_{\zeta\zeta})_{r,s} = \frac{1}{N} \sum_{n=1}^{N} \zeta_r(S_T^n) \zeta_s(S_T^n)$$

$$(\hat{M}_{\zeta X})_r = \frac{1}{N} \sum_{n=1}^{N} \zeta_r(S_T^n) e^{2X_T^n}$$
(27)

However, those estimations can be high-biased because the same paths are used to compute the regression coefficients are used to the conditional expectation. [2] treats the analytical removal of the bias and the use of Bessel inequality (to prove the under-estimation of variance of e^{2X_t}

Estimation via Monte-Carlo

2 Overview of the methodology proposed in [4]

2.1 Main equations

We have the following dynamics for the FX Spot S

$$\frac{dS_t}{S_t} = \mu_t^d dt + \sqrt{v_t} A(S_t, t) dW_t^{(1)}$$

$$v_t \text{ is a stochastic process} : \text{ Heston or exponential-OU,...}$$

$$dv_t = \kappa(m(t) - v_t) dt + \alpha \sqrt{v_t} dW_t^{(2)} \text{ for Heston}$$

$$d\langle W^{(1)}, W^{(2)} \rangle_t = \rho dt$$
(28)

2.1.1 PDE for volatility as Heston process

Let p(y, v, t) be the joint probability distribution for log-spot y and stochastic variance v at time t.

$$\frac{\partial p}{\partial t} = \frac{1}{2} \frac{\partial^2 \left[vA^2(y,t)p \right]}{\partial y^2} + \rho \alpha \frac{\partial^2 \left[vA(y,t)p \right]}{\partial y \partial v} + \frac{1}{2} \alpha^2 \frac{\partial^2 \left[v^2p \right]}{\partial v^2} + \frac{\partial}{\partial y} \left[\left(\frac{1}{2} vA^2(y,t) - \mu_t^d \right) p \right] + \kappa \frac{\partial \left[(v-m_t) p \right]}{\partial v}$$
(29)

It then reduces to

$$\frac{\partial p}{\partial t} = \frac{A^2 v}{2} \frac{\partial^2 p}{\partial y^2} + \rho \alpha A v \frac{\partial^2 p}{\partial y \partial v} + \frac{\alpha^2 v}{2} \frac{\partial^2 p}{\partial v^2} + \left(\frac{1}{2} A^2 v + 2AA' v + \rho \alpha A - \mu_t^d\right) \frac{\partial p}{\partial y} + \left(\kappa \left[v - m_t\right] + \alpha^2 + \rho \alpha A' v\right) \frac{\partial p}{\partial v} + \left[AA' v + \rho \alpha A' + AA'' v + \left[A'\right]^2 + \kappa\right] p$$
(30)

2.1.2 PDE for volatility as exponential of Ornstein-Ulhenbeck process

3 Solving the SLV calibration PDE

In this section we will develop numerical methods to solve equation (17)

References

- [1] Frédéric Abergel, Rémi Tachet, A Non-linear integro-differential equation from Mathematical Finance, AIMS Journal, Version 1, 8th Sept. 2011
- [2] Alexandre Humeau, Credit Valuation Adjustment, Master's thesis Ecole Centrale Paris Version 2, 1st November 2013
- [3] Anthonie W. Van Der Stoep, Lech A. Grzelak, Cornelis W. Oosterlee, *The Heston Stochastic-Local volatility Model: Efficient Monte-Carlo simulation*, Version 1, 11th June 2013
- [4] Iain J. Clark, Foreign Exchange option pricing: A Practitioner's Guide, Wiley Finance, ,2010