

EAI Math Reading Group II

Wishart Matrices

&

The Marcenko - Pastur  
distribution

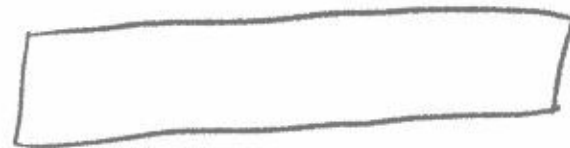
# Setup: Sample Covariance Matrices

data:  $T$  observations of  $N$  variables  $\{x_{i,t}^t\}_{i,t}$   $i \in \{1, \dots, N\}$   
 $t \in \{1, \dots, T\}$

↳ e.g. time series from stocks, neurons, ... 

• Main parameter:  $q = \frac{N}{T}$

$q < 1$ :



more samples than  
variables

$q > 1$ :



more variables than  
samples

(e.g. genetic studies)

Problem: Estimate the true covariance of the data

Assuming zero-mean  $x_i^t$  (otherwise remove the mean)

Sample  
Covariance  
matrix

$$E = \frac{1}{T} H H^T$$

$\leadsto$  symmetric, positive definite

$$E_{ij} = \frac{1}{T} \sum_{t=1}^T x_i^t x_j^t$$

(Covariance of the  
variables)

$$F = \frac{1}{N} H^T H$$

(Covariance of the  
observations)

•  $TE$  and  $NF$  have the same eigenvalues

•  $q < 1$ :  $F$  has  $N$  non-zero eigenvalues  $\frac{1}{q} \lambda_k^E$ ,  $T-N$  zero eigenvalues

$$\Rightarrow g_T^F(z) = \frac{1}{T} \sum_{k=1}^T \frac{1}{1 - \lambda_k^F} = q^2 g_N^E(z) + \frac{1-q}{z}$$

## First & Second Moments of a Wishart Matrix

Suppose the columns of  $H$  are drawn iid  $\sim N(0, C)$

$$\Rightarrow \mathbb{E}[H_{it} H_{js}] = C_{ij} \delta_{ts} \quad \begin{array}{l} \text{true} \\ \text{covariance} \end{array}$$

•  $E = \frac{1}{T} H H^T$  (Wishart Matrix)

$$\mathbb{E}[E_{ij}] = \frac{1}{T} \mathbb{E}\left[\sum_{t=1}^T H_{it} H_{jt}\right] = \frac{1}{T} \sum_{t=1}^T C_{ij} = C_{ij}$$

$$\Rightarrow \mathbb{E}[E] = C \quad \leadsto \text{the sample covariance matrix is an unbiased estimator of the true covariance}$$

•  $\mathcal{Z}(E^2) = \frac{1}{N T^2} \mathbb{E}[\text{tr}(H H^T H H^T)] = \frac{1}{N T^2} \sum_{i,j,t,s} \mathbb{E}[H_{it} H_{jt} H_{js} H_{is}]$

$$= \dots = \mathcal{Z}(C^2) + \frac{N}{T} \mathcal{Z}(C)^2 + \frac{1}{T} \mathcal{Z}(C^2) \rightarrow \mathcal{Z}(C^2) + \underbrace{q \mathcal{Z}(C)^2}_{\text{Wick's theorem}}$$

$\uparrow$  Wick's theorem

# Law of Wishart Matrices

$$P(\{H_{ij}\}) = \frac{1}{\sqrt{(2\pi)^N \det C}} \exp \left[ -\frac{1}{2} H_{it} (C^{-1})_{ij} H_{jt} \right]$$

$$P[H] = \frac{1}{\sqrt{(2\pi)^N \det C^{T/2}}} \exp \left[ -\frac{1}{2} \underbrace{\text{tr}(H^T C^{-1} H)}_{= \text{tr}(E C^{-1})} \right]$$

$$H \rightarrow E$$

$$P[E] = \frac{(T/2)^{NT/2}}{\Gamma_N(T/2)} \frac{(\det E)^{\frac{T-N-1}{2}}}{(\det C)^{\frac{T}{2}}} \exp \left[ -\frac{T}{2} \text{tr}(E C^{-1}) \right]$$

↳ multivariate  
gamma distribution

$$\leadsto \left[ \begin{array}{l} N=1: P(x) = \frac{b^a}{\Gamma(a)} x^{a-1} e^{-bx} \\ b = \frac{T}{2C} \quad a = \frac{T}{2} \end{array} \right.$$

→ The Wishart Ensemble

generalizes the Gamma distribution to Gaussian matrices

## White Wishart Matrices

$$\begin{aligned} \Sigma &= O \Lambda O^T \\ \log \Sigma &= O \underbrace{\log \Lambda}_{\text{diagonal}} O^T \end{aligned}$$

By the identity  $\det \Sigma = \exp(\text{tr} \log \Sigma)$ , we get

$$P[\Sigma] = \frac{(T/2)^{\frac{NT}{2}}}{\Gamma_N(T/2)} \frac{1}{(\det C)^{\frac{T}{2}}} \exp \left[ -\frac{T}{2} \text{tr}(\Sigma C^{-1}) + \frac{T-N-1}{2} \text{tr} \log \Sigma \right]$$

• Let  $C = I$  (white case), then

$$P[W] \approx \exp \left[ -\frac{N}{2} \text{tr} V(W) \right] \quad \text{where } V(W) = \left(1 - \frac{1}{q}\right) \log W + \frac{1}{q} W$$

↳ rotationally  
invariant

The Marcenko-Pastur distribution (White case  $C = I$ )

Let  $H \in \mathbb{R}^{N \times T} \sim \text{iid } N(0, 1)$ ,  $W = \frac{1}{T} H H^T$

Let  $z \in \mathbb{C} \setminus \sigma(W)$ ,  $M = zI - W$ , then by Schur's complement

$$\frac{1}{(G(z))_{11}} = M_{11} - M_{12} (M_{22})^{-1} M_{21} \quad \text{where} \quad M = \begin{bmatrix} M_{11} & M_{12} \\ M_{21} & M_{22} \end{bmatrix} \begin{matrix} 1 \\ N-1 \end{matrix}$$

$$G(z) = (zI - W)^{-1}$$

$$= z - W_{11} - \frac{1}{T^2} \sum_{t,s=1}^T \sum_{j,k=2}^N H_{1t} H_{jt} \underbrace{(M_{22}^{-1})_{jk}}_{\substack{1 \\ N-1}} H_{ks} H_{1s}$$

$$= \underbrace{z - W_{11}}_{= 1 + O(T^{-\frac{1}{2}})} - \frac{1}{T} \sum_{j,k=2}^N \underbrace{\frac{\sum_t H_{kt} H_{jt}}{T}}_{W_{kj}} (M_{22}^{-1})_{jk} + \underbrace{O(T^{-\frac{1}{2}})}$$

$$\begin{aligned}
\frac{1}{(G(z))_{11}} &= z - 1 - \frac{1}{T} \operatorname{tr} W_2 G_2(z) + O(T^{-\frac{1}{2}}) \\
&= \operatorname{tr}(W_2(zI - W_2)^{-1}) \\
&= -\operatorname{tr} I + z \operatorname{tr}((zI - W_2)^{-1}) \\
&= -\operatorname{tr} I + z \operatorname{tr} G_2(z)
\end{aligned}$$

$$\Rightarrow \frac{1}{(G(z))_{11}} = \underbrace{z - 1 + q - qz g(z)} + \underbrace{O(N^{-\frac{1}{2}})}$$

$$\downarrow N, T \rightarrow \infty \quad \frac{N}{T} = q$$

$$\boxed{\frac{1}{g(z)} = z - 1 + q - qz g(z)}$$



## Solution of the Helmholtz transform

$$g(z) = \frac{z + q - 1 \pm \sqrt{(z + q - 1)^2 - 4qz}}{2qz}$$

$$\lambda_{\pm} = (1 \pm \sqrt{q})^2$$

Correct branch:

$$g(z) = \frac{z - (1 - q) - \sqrt{z - \lambda_+} \sqrt{z - \lambda_-}}{2qz}$$

Eigenvalue density

$$\rho(x) = \frac{1}{\pi} \lim_{\eta \rightarrow 0} \text{Im } g(x - i\eta) = \frac{\sqrt{(\lambda_+ - x)(x - \lambda_-)}}{2\pi qx} + \underbrace{\frac{q-1}{q} f(x)}_{\text{zero eigenvals when } N > T}$$

# General Wishart Matrices

$$E_C = \frac{1}{T} H_C H_C^T$$

$$H_C \sim N(0, C)$$

$$\text{Let } C = O \Lambda O^T \text{ (diagonalized form)}$$

$$\Lambda = \begin{pmatrix} \sigma_1^2 & & \\ & \ddots & \\ & & \sigma_N^2 \end{pmatrix}$$

$$C^{\frac{1}{2}} = O \Lambda^{\frac{1}{2}} O^T$$

$$\Lambda^{\frac{1}{2}} = \begin{pmatrix} \sigma_1 & & \\ & \ddots & \\ & & \sigma_N \end{pmatrix}$$

To generate  $y \sim N(0, C)$ , generate  $x \sim N(0, I) \leadsto y = C^{\frac{1}{2}} x$

$$\Rightarrow E[yy^T] = E[C^{\frac{1}{2}} x x^T C^{\frac{1}{2}}] = C$$

$$E_C = \frac{1}{T} H_C H_C^T = \frac{1}{T} \underbrace{C^{\frac{1}{2}} H H^T C^{\frac{1}{2}}}_{\text{"Free Product"}} = C^{\frac{1}{2}} \underbrace{W_q}_{\text{White Wishart w/ } q = \frac{N}{T}} C^{\frac{1}{2}}$$

"Free Product"

$\hookrightarrow$  White Wishart w/  $q = \frac{N}{T}$

$\hookrightarrow$  "Free Probability"