

form/shape, direction, strength, points follow recognizable form), unusual features. **PC**
Coeff: strength of linear association; -1 to 1; 0 means no linear; r does not change if variables are rescaled. **2: Residual:** observed - predicted; $e_i = y_i - \hat{y}_i$.
Least Squares: minimize SSE: $\sum_{i=1}^n e_i^2$. **Interp Coeff:** slope coeff as the pred change in Y associated with a one-unit change in X_i , holding all other predictors constant; intercept as the predicted Y value when all preds are 0 (could be nonsensical). **R^2 :** coeff of determination; $= \frac{SSE(\bar{y}) - SSE(\hat{y})}{SSE(\bar{y})}$; unexplained variation in y / total var in y (look at points distribution). **R^2 Interp:** percent reduction in SSE by taking into account the predictors; percentage of variation in Y explained by the regression function with predictors; range is 0 to 1 and 1 if perfect predictions; $R^2 = r^2$ for simple linear regression. **3: Least Squares Estimate of β :** $SSE = \sum_{i=1}^n e_i^2 = e^T e$; $\hat{\beta} = (X^T X)^{-1} X^T y$; $\hat{y} = X \hat{\beta} = H y$. **Hat Matrix:** $H = X(X^T X)^{-1} X^T$; symmetric and h_{ij} describes the weight each of the values in the ith row of H have on the predicted value of y_i . **Contributions:** $h_{ij} y_j$ is the actual contribution of j makes to the value \hat{y}_i . **4: MSE:** $s^2 = \frac{SSE}{n-p}$; s or RMSE is the typical prediction error; expect 95% of observed y values to lie roughly within 2s of predicted values; called **residual standard error in R. s vs R^2 :** both small s and large R^2 is the goal. **5: Perm. Test:** is there a relation between y and x find test stat that measures association for all possible perms of the resp var and compute the prop. of times an observed test stat as extreme as the one from original sample; p-value from the graph by counting. **p-value:** probability of obtaining a result (test stat) at least as extreme as the one observed, if the null were true; < 0.1 is some, < 0.05 is fairly strong, < 0.01 is very strong, < 0.001 is extremely strong; small p-value means result is unlikely to have occurred by chance alone, if the null were true making it statistically significant. **Inference on β_j :** $t = \frac{\hat{\beta}_j - \beta_j}{SE(\hat{\beta}_j)}$; $1 - \alpha$ confidence limits: $\hat{\beta}_j \pm t_{n-p, \alpha/2} SE(\hat{\beta}_j)$. **Standard Error of β :** $SE(\hat{\beta}_j) = \sqrt{MSE(C_{j+1} C_{j+1}^{-1})}$ where C is the jth column of $(X^T X)^{-1}$; **summary(fit)** gives null = 0 and alternative $\neq 0$; **model.matrix(fit)** gives design matrix X. **One Sided H:** **pt()** with lower tail false gives area to the right ($t > 0$) and true gives left tail area. **68/95/99.7 Rule:** values like within 1/2/3 std dev of mean. **t-value:** $t > 2$ or < -2 means results are significant. **Critical Value:** **qt(p, df)** gives t value for area p to the left of it; 90/95/99 percent CI is $p = 0.95/0.975/0.995$. **CI Slope Interp:** We are 95% confident that the expected *change* in the *response* for each one *unit* increase in the X_i falls between l and u, (after adjusting for other predictors in the model) **use only when interpreting about the effect of one predictor with multiple predictors, say for conclusions too in inference tests.** **Bonferroni:** $1 - \frac{1-\alpha}{g}$ where C is the confidence level and g is the number of coefficients being tested; want each individual confidence to be above C in order for the joint level to be above C; **only for intervals.** **Joint CI Interp:** We are (at least) 95% confident that all intervals correctly capture the population parameters. **6: FINE Assumptions:** Form, Independence, Normality, Equal Variance; Form: expect linear form, Independence: errors are independent (in data description), Normality: errors follow a normal dist, Equal Variance: variance of errors is the same. **Plots:** Form: residuals vs fitted (no trend/curve) or residuals vs each X_i **for multiple predictors**; Equal Variance: residuals vs fitted (no fan shape); Normality: qq plot (no big departure from straight line) or histogram (no big skew); Independence: look at residuals vs observations number (want to be random). **Formal tests:** wilks \Rightarrow normality, pagan \Rightarrow equal variance, low p means violated. **7: Interp of Slope:** only log transforms can be restated in terms of the original vars, preds can always be restated in terms of original vars. **Transforming Y:** non-linearity, non-constant variance, and non-normality; X: non-linearity, high leverage, influence; **Ladder of Powers:** $p = 2, 1, 0.5, 0, -0.5, -1, -2$; $y^* = y^2, y, \sqrt{y}, 1/\sqrt{y}, 1/y, 1/y^2$; right to become better; log and sqrt not defined for zero or negative values, so transform $y/x + c$, where c makes all values ≥ 1 . **Strategies:** skewed residuals: right skew is y down, left skew is y up; residual var inc. as x incr: y down, decr as x incr: y up; non-linear: correct non-normal and unequal var then y, only non-linear then x. **Non-linearity Bulges Point:** up and left is y up or x down; up and right is y up or x up; down and left is y down or x down; down and right is y down or x up. **Interp of X Transform:** If we multiply x_i by b (chosen log base) we predict a *change* of $\hat{\beta}_i$ in the mean value of y after adjusting for the other vars in model. **Interp of Y Transform:** Each one unit *change* in x_i *changes* the predicted median value of y by a factor of $b^{\hat{\beta}_i}$ after ... **Interp of Both:** A c-fold *change* in x_i *changes* the predicted median value of y by a factor of $c^{\hat{\beta}_i}$ after ... **Non-log Transform:** Each increase of one in $1/x_i$ is associated with an increase of $\hat{\beta}_i$ in predicted \sqrt{y} . **Median:** median instead of mean because $E[\log(Y)] \neq \log(E[Y])$ but for median it is true. **Box-Cox:** round lambda to nearest 0.5. **Matrix Scatterplot:** don't reflect preds act jointly.

8A: R-sq equation: $SSTO = SSR + SSE$; $R^2 = \frac{SSR}{SSTO}$; variation explained by model / total variation in response; Adding vars moves SSE to SSR. **Type 1 SS:** **anova(fit)**; order matters, after adjusting for vars earlier in model; additional var in y explained by adding x to model already containing other x's. **Type 2 SS:** **anova(fit, type = 'II')**; after adjusting for all other vars in model; output row is var being entered last. **8B: DF Type 1:** total = n-1, error = n-p, reg = p-1; MS = SS / df, MS = SSR / k, MSE = SSE / (n-k-1). **R-sq:** Adding additional vars can never decrease R^2 ; $R_{adj}^2 = 1 - \frac{SSE/(n-k-1)}{SSTO/(n-1)} = 1 - \frac{n-1}{n-k-1} (1 - R^2)$; can decrease when new unnecessary var is added to model. **General F Test Props:** p-values found in right tail (upper tail), df = F(r, n-k-1), where r is number of vars in full model; FINE assumptions met for partial F test. **Partial F-Test:** **anova(fit.reduced, fit.full)**; reduced is model without variables; H_0 : r of β_j is 0, H_a : at least one $\beta_j \neq 0$; $F = \frac{(SSE_{reduced} - SSE_{full})/r}{MSE(full)}$; can replace SSE with SSR but full is first; numerator is reduction in SSE (boost in SSR) for full compared to reduced, dividing by r accounts for change in model complexity, standardize by MSE of full model. **Model Utility Test:** **summary(fit)**; H_0 : all β_j are 0, H_a : at least one $\beta_j \neq 0$; $F = \frac{(SSTO - SSE(full))/k}{MSE(full)} = \frac{MSR}{MSE}$; numerator is the SSR(full); reduced model is mean model (no predictors); **Single Coef Test:** **summary(fit)** or **Anova(fit, type='II')**; H_0 : $\beta_j = 0$, H_a : $\beta_j \neq 0$; $F = \frac{(SSR(full) - SSR(all-but-x_j))/1}{MSE(full)}$; reduced model is full model without x_j ; **Equivalent to t-test with $t^2 = F$ and p-values are the same.** **SS Total: Type 1 ANOVA table contains SST, not Type 2 b/c not sequential.** **MUT Decision:** reject H_0 means we have sufficient evidence to conclude that at least one of vars is not equal to 0. Can conclude that at least one of vars is significantly useful in predicting y. **SCT Decision:** large p-val means we do not have enough evidence to believe that var is any different from zero after adjusting for other vars, that is var does not significantly improve the model containing other vars. **PFT Decision:** Fail to reject the null, adding r vars to model that already contains other vars does not significantly improve model. Not enough evidence that full containing all vars is better than reduced model, so keep reduced b/c of parsimony; Reject null, means larger has significant reduction in SSE, bost in SSR and R^2 , and improves pred of Y. **Dropping Vars:** cannot just drop all variables if not significant, need to come out one at a time or perform partial F test. **10: Point Estimate:** $\hat{y}_0 = x_0^T \hat{\beta}$; in R, **predict()**. **Prediction Interval:** $t_{n-k-1, \alpha/2} \sqrt{1 + x_0^T (X^T X)^{-1} x_0}$; predicting the future response y_0 of an individual for a particular value of x_0 , use "individual" in interp; wider than CI b/c of greater uncertainty around predicting individual observation; wider as we move further away from original combination of data; refer to as a chance since it is a future observation. **Confidence Interval:** same but without the 1 + term; **predicting mean response for a particular value of x_0** ; use "all" keyword. **11: Signs of MultiCol:** corr matrix values close to 1 or -1; coeffs do not expected signs; overall F test significant but no t-tests significant; large VIFs. **VIF:** $VIF_j = \frac{1}{1 - R_j^2}$, where R_j^2 is the proportion of variation in x_i that is explained by its linear relationship to other vars; measures increase in variability of ith sample reg coeff due to linear assoc of x_i with other preds in model; $VIF_j = 1$ if not linearly related, SD is unchanged when other preds enter model; 5 or 10 means problems in estimation, 4 means SD doubles, 10 means SD triples; Ex. regressing x on other x's; **Added Variable Plot:** *residual($y \sim x_1 + x_2$)* vs *residual($x_3 \sim x_1 + x_2$)* where first is y with the effect of x1 and x2 removed and second is x3 with the effect of x1 and x2 removed; **whether x_i has anything new to tell us (unique contribution) after removing effect of other vars**; intercept should always be 0 since SSE = 0; reg line goes through average which is 0; linear relationship means useful; ex. association between left over (unexplained) variation in y and x1 after adding x2 to the model (looking at x1); **Slope of LS Line:** what it means to adjust for other vars; slope of LS line in plot is the reg coeff of xi in multiple reg model; entire row is same in output; large p-val means xi not useful to add to model that already has other vars. **12: Leverage:** outlying in x-space; ith diagonal of hat matrix is leverage of ith obs; h_{ii} is leverage of ith obs; 0 to 1, $h_i > 2\bar{h} = 2(k+1)/n$ are leverage points. **Influential:** removing obs changes fit of model; outlying in x and y not consistent; large residuals mean outlying in y. **Internally Studentized Residual:** $r_i = e_i / \sqrt{1 - h_i}$; $r_i > 3$ is large residual; **rstandard(fit).** **Externally Studentized**

Residual: $t_i = e_i/s_{(i)}\sqrt{1-h_i}$; $s_{(i)}$ is s when ith obs removed; $t_i > Bonferroni$ is extreme; `rstudent(fit)`. **DFITS:** $\frac{\hat{y}_i - \hat{y}_{i(i)}}{s\sqrt{h_i}}$; comparing fitted values w/wo
 obs in data set and standardize by denom. **Cook's Distance:** $D_i = \frac{1}{p+1} \frac{e_i^2}{s^2} \frac{h_{ii}}{1-h_{ii}}$ which is constant*internally studentized residual* leverage; it is an aggregated
 measure of the effect of removing ith obs on all predicted values. **Guidelines Both:** > 1 as being influential for $n < 30$, greater than $2\sqrt{(k+1)/n}$ for large, D_i
 greater than 0.5 or 1 as influential.