

# Lecture 10

## Quantitative Political Science

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# Agenda

1. Finishing up last lecture's example
2. Hypothesis Testing
3. Relation to CIs
4. Two- versus one-tailed tests

# Example Time!

- Poll of 1,203 adults between Sep. 15 and 20, 2023 asking about a hypothetical vote choice if the election were held tomorrow, found that 52.5% of respondents indicated they would support Trump, and 47.5% indicated they would support Biden. This marks a reduction in Trump support from a previous tracking poll fielded a week earlier of 1,203 adults who indicated 55.6% support for Trump and 44.4% support for Biden.
- How confident are we that the change in Trump's support over this period is not due to sampling error?
- Parameter we seek is  $p_1 - p_2$  where  $p_1$  is Trump's **true** support in the first poll and  $p_2$  is his **true** support in the second poll. Consider the polls as binomial experiments in which  $Y_1$  is the number of "successes" (here, the # favoring Trump) in the first poll and  $Y_2$  is the number of "successes" in the second poll.
- Intuitive estimator:  $\hat{p}_1 - \hat{p}_2$ . Is this unbiased?
- Calculate estimator's standard errors:  $\sqrt{VAR(\hat{p}_1 - \hat{p}_2)} = \sqrt{VAR(\hat{p}_1) + VAR(\hat{p}_2)} = \sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}$

# Example Time!

- Continuing from the previous example, what is the 95% confidence interval for this estimator?
- Does this interval include zero? How can we interpret that?
- What about the 90% confidence interval? Does it still include zero?
- At what level of confidence would we conclude Trump's support changed between the two surveys?
- **Think:** want to find  $\alpha$  (call it  $\alpha^*$ ) s.t. the *lower bound of the CI is greater than zero*

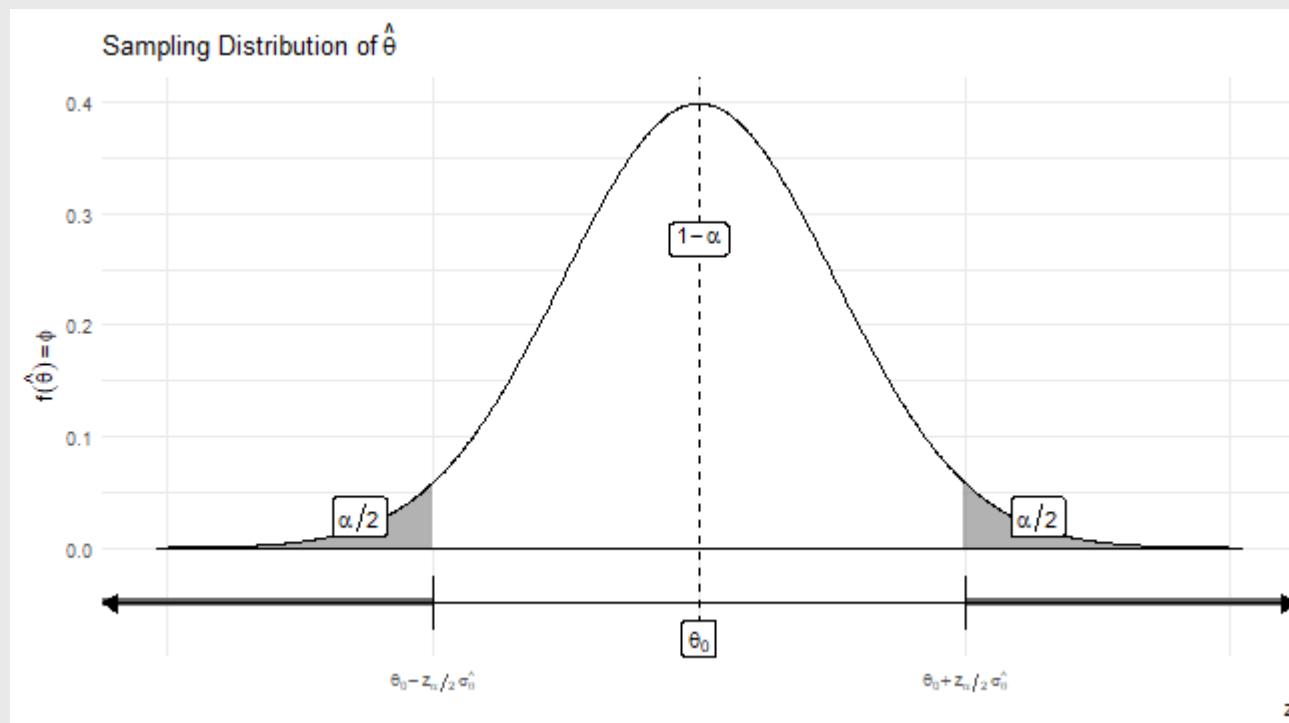
$$\begin{aligned}\hat{p}_1 - \hat{p}_2 - z_{\alpha^*/2} \left( \sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}} \right) &> 0 \\ -z_{\alpha^*/2} \left( \sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}} \right) &> -(\hat{p}_1 - \hat{p}_2) \\ z_{\alpha^*/2} &< \frac{\hat{p}_1 - \hat{p}_2}{\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}}\end{aligned}$$

# Hypothesis Testing

- **Hypothesis Test** consists of four elements:
  1. **Null** hypothesis about a parameter:  $H_0$
  2. **Alternative** hypothesis about the parameter:  $H_A$
  3. **Test statistic** derived from estimator of the parameter
  4. **Rejection region**: range of values of test statistic for which  $H_0$  should be *rejected* in favor of  $H_A$
- Choosing the RR trades off two kinds of errors:
  - **Type I error**: reject  $H_0$  when it is actually true
  - **Type II error**: accept  $H_0$  when  $H_A$  is actually true

# Type I Error

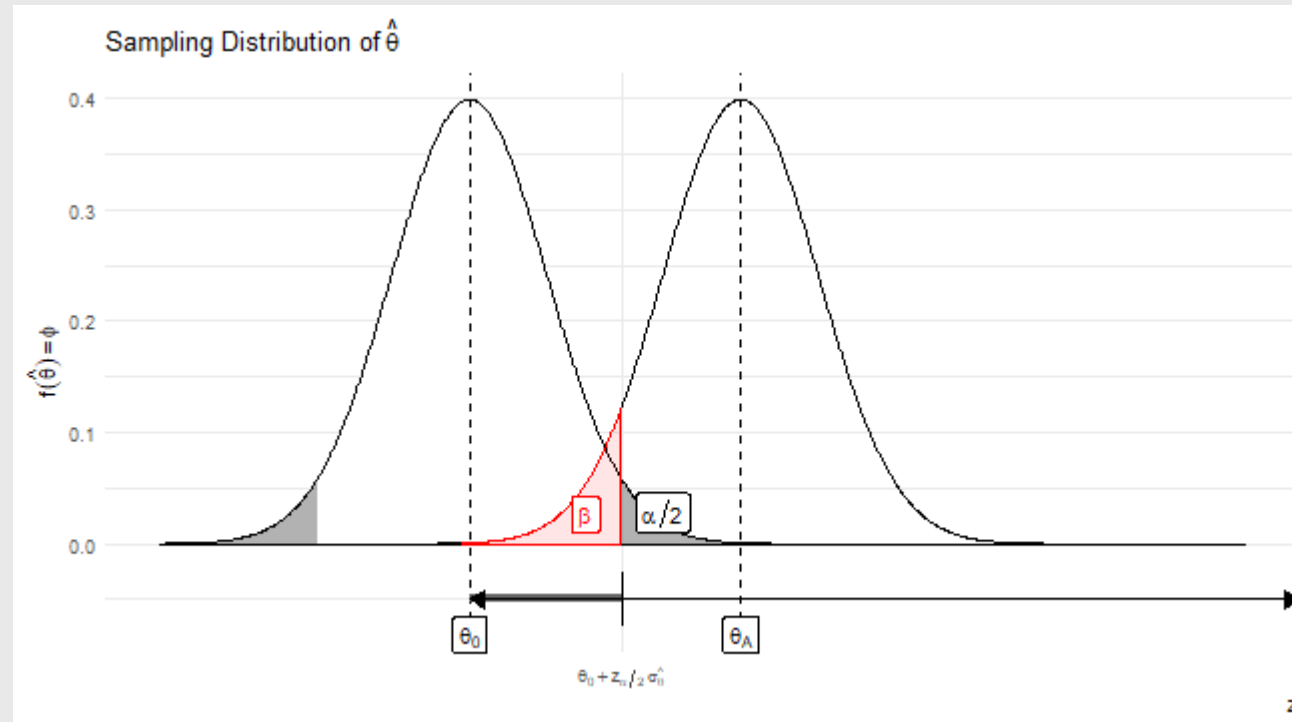
- **Type I error:** reject  $H_0$  when it is actually true
  - What does this look like?



# Type I error

- We will (purely by chance):
  - Observe an estimated  $\hat{\theta}$  in the  $RR$   $100 * \alpha\%$  of the time
  - Thus falsely reject the null even though it's true
- This is Type I error!

# Type II error





# Type II error

- Suppose that the alternative hypothesis is true
- But we always conduct our hypothesis test **under the assumption that the null is true**
- If the sampling distribution of our estimator  $\hat{\theta} \sim \mathcal{N}(\theta_A, \sigma_{\hat{\theta}})$ , we will mistakenly accept the null  $100 * \beta$  % of the time
- Define **power** as  $1 - \beta$

$$\begin{aligned}\text{Power} &= 1 - \beta \\ &= 1 - \Pr(\text{reject } H_0 | H_A \text{ true}) \\ &= 1 - \Pr(\hat{\theta} < \theta_0 + z_{\alpha/2} \sigma_{\hat{\theta}} | \theta = \theta_A)\end{aligned}$$

# Type I and II error

- Thus  $P(\text{Type I}) = \alpha$  and  $P(\text{Type II}) = \beta$
- Ideally, we want the hypothesis test's level of significance to be **low** and its power to be **high**
- Why is this a trade-off?
- Re-evaluate the example:
  - $H_0 : p_1 - p_2 = 0$
  - $H_A : p_1 - p_2 \neq 0$
- Test statistic is  $\hat{p}_1 - \hat{p}_2$
- Rejection region is all values of statistic for which we reject  $H_0$  for chosen  $\alpha$ 
  - I.e., values of  $\hat{p}_1 - \hat{p}_2$  where the constructed CI **does not include zero**

# Type I and II error

- Recall our CI:  $(\hat{p}_1 - \hat{p}_2) \pm z_{\alpha/2} \sqrt{\frac{\hat{p}_1(1-\hat{p}_1)}{n_1} + \frac{\hat{p}_2(1-\hat{p}_2)}{n_2}}$
- We could have rejected  $H_0$  if  $(\hat{p}_1 - \hat{p}_2) - z_{\alpha/2} \sqrt{\frac{\hat{p}_1(1-\hat{p}_1)}{n_1} + \frac{\hat{p}_2(1-\hat{p}_2)}{n_2}} > 0$  and concluded Trump's popularity did fall
- Rewriting:

$$\begin{aligned} (\hat{p}_1 - \hat{p}_2) &> z_{\alpha/2} \sqrt{\frac{\hat{p}_1(1-\hat{p}_1)}{n_1} + \frac{\hat{p}_2(1-\hat{p}_2)}{n_2}} \\ \frac{(\hat{p}_1 - \hat{p}_2)}{z_{\alpha/2} \sqrt{\frac{\hat{p}_1(1-\hat{p}_1)}{n_1} + \frac{\hat{p}_2(1-\hat{p}_2)}{n_2}}} &> z_{\alpha/2} \end{aligned}$$

# Sample Size

- What sample size would have been needed for our difference in sample populations to be statistically significant from zero with 95% confidence? (Assume  $n_1 = n_2 = n$ )
- Plug in the numbers!
- Power calculation is a crucial tool for determining how big your sample must be to avoid committing Type II error
- (We will come back to this soon)

# Relation to CIs

- Walk through a question
  - Conventional wisdom says that  $\theta = \theta_0$ , but I theorize that  $\theta \neq \theta_0$
  - I obtain a point estimate  $\hat{\theta} \neq \theta_0$
  - How sure am I that  $\theta \neq \theta_0$ ?
- This is the core language of **hypothesis tests**
- Often  $\theta_0 = 0$ , but it could be any value
- Regardless, we can fully define the distribution of our estimator if  $\theta_0$  is true
- We know that CLT tells us that the standardized version of **any** estimator is  $\frac{\hat{\theta} - \theta_0}{\sigma_{\hat{\theta}}} \sim \mathcal{N}(0, 1)$

# Hypothesis Testing

- Remember...4 components!
  1. Null hypothesis  $H_0$
  2. Alternative hypothesis  $H_A$
  3. Test statistic  $\hat{\theta}$
  4. Rejection region
- Start by choosing  $\alpha$  which is now defined as **the probability of Type I error**
- Then identify the range of values of  $\hat{\theta}$  we will observe  $\alpha$  percent of the time in repeated sampling, which is our **rejection region**
- If we observe  $\hat{\theta}$  in this region, we reject  $H_0 : \theta = \theta_0$  in favor of  $H_A : \theta \neq \theta_0$
- In practice, we reject  $H_0$  if  $\hat{\theta} < \theta_0 - z_{\alpha/2}\sigma_{\hat{\theta}}$  or if  $\hat{\theta} > \theta_0 + z_{\alpha/2}\sigma_{\hat{\theta}}$

# One-Tailed Hypothesis Test

- What if we have a stronger alternative hypothesis?
  - Our alternative is **signed**
  - Instead of  $H_A : \theta \neq \theta_0$ , I have theoretical reason to believe  $H_A : \theta > \theta_0$
- Again, pick  $\alpha$
- Then look at the standard Normal and identify range of values for  $\hat{\theta}$  greater than  $\theta_0$  that we will observe  $\alpha\%$  of the time in repeated sampling
- Beware of cooking the books! Say you have some  $\hat{\theta} > \theta_0$ , and you make an *ex post* hypothesis that  $H_A : \hat{\theta} \geq \theta_0$ . This is not based on theory, and looks very suspicious!