Lecture 16 Quantitative Political Science

Prof. Bisbee

Vanderbilt University

Lecture Date: 2023/11/02

Slides Updated: 2023-12-23

Agenda

- 1. Variance of OLS estimators
- 2. Heteroskedasticity

Recap

- We went over 4 assumptions to characterize the bias of our OLS estimators
- 1. Relationship between \boldsymbol{x} and \boldsymbol{y} is **linear in its parameters**
- 2. x and y are drawn from a random sample, making them **i.i.d.**
- 3. $VAR(X) \neq 0$
- 4. E(u|x) = 0
- With these, we demonstrated that \hat{eta}_0 and \hat{eta}_1 are **unbiased** for eta_0 and eta_1

Sampling Distributions

- ullet \hat{eta}_0 and \hat{eta}_1 are statistics, just like $ar{Y}$
- When we evaluate statistics, we care about both their bias (last class) and their variance
 - How far can we expect them to be from their true value (i.e., the population parameter) on average?
- ullet In the univariate case, we were interested in the **sampling distribution** of $ar{Y}$
- ullet Here, we are also interested in the sampling distributions of \hat{eta}_0 and \hat{eta}_1

- We already know the means of \hat{eta}_0 and \hat{eta}_1 : they are eta_0 and eta_1 (from last class)
- To compute the variances of $\hat{\beta}_0$ and $\hat{\beta}_1$, we need a fifth assumption

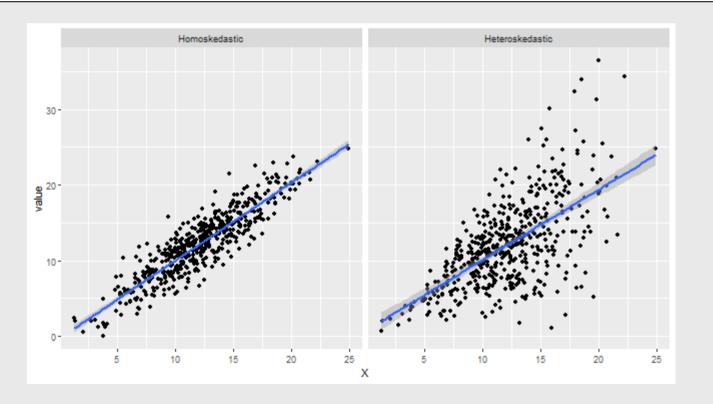
Assumption 5: $VAR(u|x) = \sigma^2$

- ullet The error has the **same variance** regardless of the value of x
- This is known as homoskedasticity
- If this fails, we have heteroskedasticity

```
require(tidyverse)
set.seed(123)
X <- rnorm(500,mean = 12,sd = 4)
Y <- rnorm(500,mean = X,sd = X/3)
Y2 <- rnorm(500,mean = X,sd = 2)

p <- data.frame(X = X,Heteroskedastic = Y,Homoskedastic = Y2) %>%
    gather(outcome,value,-X) %>%
    mutate(outcome = factor(outcome,levels = c('Homoskedastic','Heteroskedastic'))) %>%
    ggplot(aes(x = X,y = value)) +
    geom_point() +
    geom_smooth(method = 'lm') +
    facet_wrap(~outcome)
```

p



- Note the difference between Assumptions 4 and 5!
 - $\circ E(u|x) = 0$
 - $\circ VAR(u|x) = \sigma^2$
- We don't need assumption 5 for unbiasedness, but we do for variance!
- What is VAR(y|x)?

$$egin{aligned} VAR(y|x) &= VAR(eta_0 + eta_1 x + u|x) \ &= VAR(eta_0|x) + VAR(eta_1 x|x) + VAR(u|x) \ &= 0 + 0 + \sigma^2 \ &= \sigma^2 \end{aligned}$$

- What is σ^2 ?
- It is a measure of the extent to which unexplained factors are affecting \boldsymbol{y}
 - \circ These factors are not related to x (from assumption 4)
 - \circ These factors are constant regardless of x (from assumption 5)
 - \circ When σ^2 is big, it means that other factors explain a lot of variation in y beyond just x
 - \circ When σ^2 is small, it means that x explains a lot of variation in y
- Note that σ^2 is a **parameter**, something that exists in the population

Variance of estimators

•
$$VAR(\hat{eta}_0)=rac{\sigma^2rac{\sum x_i^2}{n}}{SST_x}$$
 and $VAR(\hat{eta}_1)=rac{\sigma^2}{SST_x}$

- I'll leave it to you to prove $VAR(\hat{\beta}_0)$, but let's dig into $VAR(\hat{\beta}_1)$

$$eta_1 = eta_1 + rac{\sum (x_i - ar{x})u_i}{SST_x}$$
 $VAR(\hat{eta}_1|x) = VARigg[eta_1 + rac{\sum (x_i - ar{x})u_i}{SST_x} \ | xigg]$
 $= VAR(eta_1|x) + rac{1}{SST_x^2} \sum (x_i - ar{x})^2 VAR(u_i \mid x)$
 $= 0 + rac{SST_x}{SST_x^2} VAR(u_i|x)$
 $= rac{\sigma^2}{SST_x}$

Sampling Variance of \hat{eta}_1

- So $VAR(\hat{eta}_1) = rac{\sigma^2}{SST_x}$
- Want this to be as small as possible (bias-variance tradeoff)
- As σ^2 gets smaller, so does $VAR(\hat{eta}_1)$
- As SST_x gets bigger, $VAR(\hat{eta}_1)$ gets smaller
- Unpack SST_x for more insights!

Sampling Variance of \hat{eta}_1

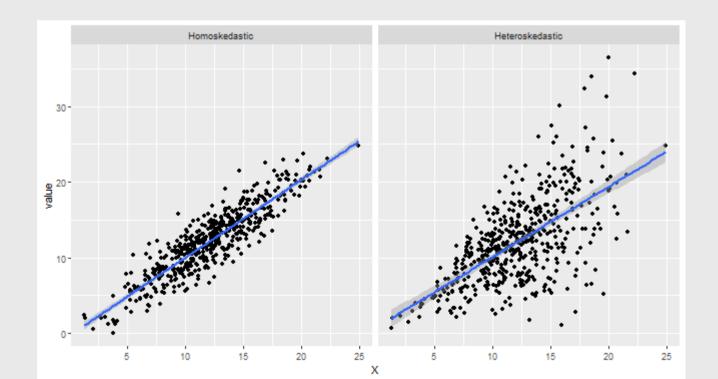
$$egin{aligned} SST_x &= \sum (x_i - ar{x})^2 \ rac{SST_x}{n} &= rac{\sum (x_i - ar{x})^2}{n} \ &= var(x) ext{ sample variance of } x \ SST_x &= n * var \ VAR(\hat{eta}_1) &= rac{\sigma^2}{n * var(x)} \end{aligned}$$

- What can we actually manipulate here as a researcher?
 - $\circ \sigma^2$ is a parameter: it declines when x explains y well. But we don't have a lot of control over this.
 - $\circ var(x)$ is the empirical variance of x in our sample. It approximates the population variance, but we don't have a ton of control over this either.
 - \circ n: we choose this!

Heteroskedasticity

- ullet The preceding results rely on the assumption of $VAR(u|x)=\sigma^2$
- What if this doesn't hold?

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Heteroskedasticity

ullet We would then say that the variance of the errors conditional on x is specific to that unit

$$\circ VAR(u_i|x_i) = \sigma_i^2$$

ullet Recall that $\hat{eta}_1=rac{\sum (x_i-ar{x})u_i}{\sum (x_i-ar{x})^2}$ and that

$$egin{align} VAR(\hat{eta}_1) &= VARigg[rac{\sum (x_i - ar{x})u_i}{\sum (x_i - ar{x})^2}igg] \ &= rac{1}{SST_x^2} \sum (x_i - ar{x})^2 VAR(u_i|x_i) \ &= rac{\sum (x_i - ar{x})^2 \sigma_i^2}{SST_x^2} \end{split}$$

Heteroskedasticity

- What to do?
- In 1980, one of the most cited economics papers was written by Halbert White
- ullet In it, he proposed calculating heteroskedastic robust standard errors as $\widehat{VAR}(\hat{eta}_1)=rac{\sum (x_i-ar{x})^2\hat{u}_i^2}{SST_x^2}$
 - \hat{u}_i^2 is just the squared residual associated with each observation i
- These standard errors have LOTS of different names:
 - "White standard errors"
 - "Huber-White standard errors"
 - "Robust standard errors"
 - "Heteroskedasticity-robust standard errors"

A Preview of What's to Come

- The easiest thing to do is just calculate robust standard errors with $\widehat{VAR}(\hat{eta}_1) = rac{\sum (x_i ar{x})^2 \hat{u}_i^2}{SST_x^2}$
- ullet However, if our fifth assumption fails (remember: $VAR(u|x)=\sigma^2$), it means that the OLS estimator is no longer the best" estimator
- What do we mean by best"? The lowest-variance!
 - As an aside, Best Linear Unbiased Estimator or BLUE is a commonly used acronym for the OLS estimators.
 We will come back to this next week in more detail, but for now, note that we can prove Unbiasedness with the first four assumptions, and assumption 5 gives us Best