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MATH2021-1 HIGH DIMENSIONAL ANALYSIS DATA Gentiane Haesbroeck

Project 3 Supervised classification

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Log repression: 3 /8 [this part is really not lear! It is lear is

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Preliminaries for the supervised classification 1

A very visual way to find out if other variables can explain our categorisation is to plot them individually and colour the graph using the binary variable.

So if we look at the variables "radius mean" [10], "concave.points mean" [10] and "concavity" [10]

we can see that the variables tends to separate into two groups.

By continuing the analysis on the variables "radius_se"[11], "concave.points_se"[11], "concavity_se"[11], "radius_worst"[20], "concave.points_worst"[20] and "concavity_worst"[20]. We come to more

or less the same conclusion. Let us note all the same for, "concave.points_se"[11] and "concavity_se"[11].

brusica variables

The two groups are very confused. clearer

Also, as we discussed in project 2, this dataset seems to be very suitable for classification and the second secon Indeed it has been showed outliers strongly depended on the class of individuals: outstanding values of explanatory variables were most of the time related to nuclei cell with malign cancer.

non ousliem.

Duld

2 Classification using the logistic regression model

2.1 Logistic model

In order to find the best model of logistic regressions explaining our binary variable. It was necessary to proceed with several iterations of the basic model. The basic model was the one that included all the variables. As can be seen in Figure 1, the model does not converge. To overcome this problem, the same regressions were applied to the 3 main dataset groups (" mean ", " se " and " worst "). This made it possible to identify the variables (perimeter mean, perimeter_se, perimeter_worst, compactness_mean, compactness_se, fractal_dimension_worst). The identification of this one was done through the study of their p-values. This made it possible to develop model 3. Although there are still many variables with large p-values, the graphical representation shows that the model is very good. In order to improve the model, the same process was carried out to manufacture model 4 but it did not show any significant improvement. An evolution of each model can be found in the appendix.

2.2 Residual deviance

I do not understand how An analysis of the residual deviance of each model allows us to see that it is becoming more defined and correct. and more balanced and correct.

Considering the values obtained, it is not surprising to see that the model2 has the best

model1	model2	model3	model4
32006.76421	24.43389	41.68973	42.59127

Tab. 1: Residual deviance

nt you do not take into account number of

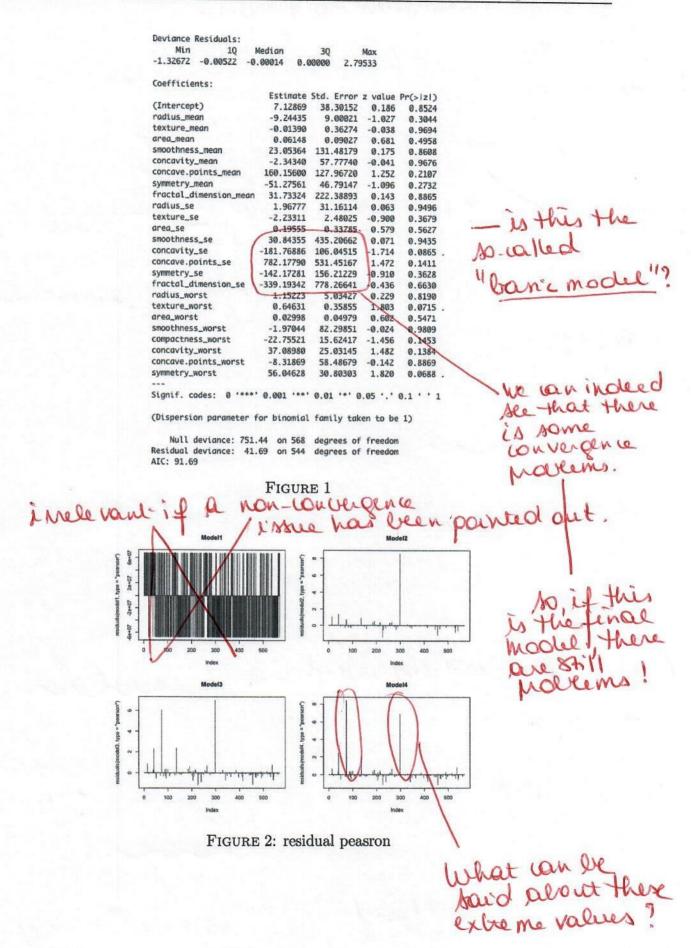
behaviour in terms of pearson residues. Even if the analysis of the figures of the logistic regressions would have led us to think that the model 3 would have the best behaviour.

in order to long pare models with \$ numbers of parameters, an approach based on the

High dimensional data analysis

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2.3 Model

Even unna R-sluggt, it possi ele to undera and.

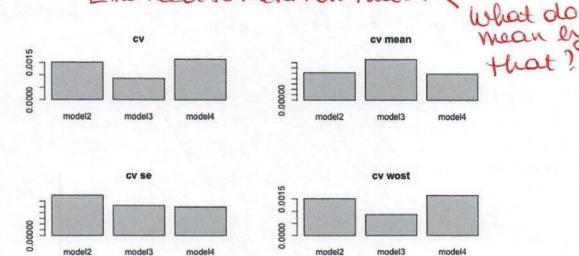
By using the coefficients table our model can be written as follows.

Model = 7.12869 * Intercept - 9.24435 * radius mean-0.01390 * texture mean + 0.06148 * area mean+23.05364 * smoothness mean - 2.34340 * concavity mean+160.15600 * concave.points mean - 51.27561 * symmetry mean $+31.73324 * fractal_dimension_mean + 1.96777 * radius_se$ -2.23311 * texture se + 0.19555 * area se+30.84355 * smoothness se - 181.76886 * concavity se $+782.17790 * concave.points_se - 142.17281 * symmetry_se$ -339.19342 * fractal dimension sev + 1.15223 * radius worst+0.64631 * texture worst + 0.02998 * area worst $-1.97044 * smoothness_worst - 22.75521 * compactness_worst$ +37.08980 * concavity worst - 8.31869 * concave.points worst+56.04628 * symmetry worstWhat can be said Are these renets conerent/logical

(1)

Cross-Validation 2.4

To ensure the accuracy of the model. The function cv.glm()\$delta allows to analyze values showing the behavior of the model. By making the difference for each delta we obtain the figure below. me need to ment on that



-it is now model 3 which is the best??

The value for model 3 (in cv) being almost equal to zero shows us once again that our model t was polled experiencely to use is almost perfect

2.5

the leave one out technique while suipt is laxed on so fold cross The study of the confusion matrix of model 3 gives us very small value for our false positive and negative and very large value for the true positive and negative. This shows once again

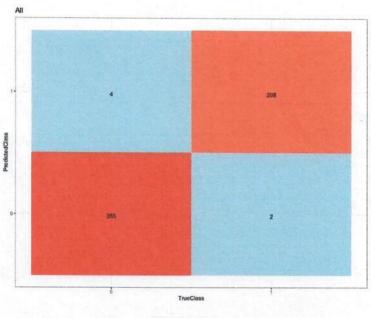


FIGURE 4

how well the classification of model 3 works. We can therefore conclude the sensitivity and specific y which is 0.9119497 and 0.9626866 respectively. Which once again shows us the good behaviour of our model

3 Classification based on LDA scores

Linear discriminant analysis idea is to find a vector (if two classes) u which will discriminate the best a class from the other on this vector. This means getting values on this vector as different as possible from a class to the other.

In order to find the vector u, the algorithm maximise variability between projected classes and minimise variance within each projected class. The da function on R software allows to get coefficients of explanatory variables to construct vector u.

Once the ${\bf u}$ vector is determined, it is then possible to obtain the projected value (score) of each individual on this new canonical variable. Distributions of classes scores define a measure of the discriminant potential of ${\bf u}$ (power of the statistical test).

3.1 Model 0: full.

In this part all columns of the dataset, except for the qualitative one, are used as explanatory variables. We will use this model as a baseline. Coefficients of \mathbf{u}_0 are displayed in 2. This canonical vector leads to a discriminant power of 0.774 which is very satisfactory to establish a decision rule based on scores.

An approach we could have intent is to let down explanatory variables having very low coefficients. Those are poorly correlated to the new canonical variable and thus probably do not influence much power discriminant. However this approach won't be discussed as two models are already developed below.

On this plot, one can observe a difference in distributions dispersion's of classes. Let's compare their overlap evolving as we tweak the model in the section below.

and? interpretation?

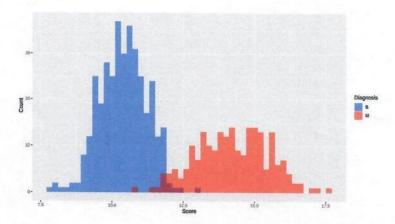


FIGURE 5: Distribution of each class on u₀.

3.2 Model 1: reducing redundancy.

Based on discussion we had in project 2, explanatory variables related to *area* and *perimeter* were removed from the dataset as a strong correlation existed between those and the *radius* measures. The model is thus reduced to 24 explanatory variables.

The discriminant power of this model is 0.750. Compared to the baseline, the loss in discriminant power is quite small (0.024). However the model size hasn't been reduced much.

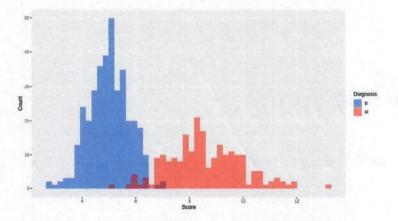


FIGURE 6: Distribution of each class on \mathbf{u}_1 .

The loss of discriminant power can be explained here as we get a larger overlap of distributions.

3.3 Model 2: visual way.

As discussed in section 1, it is possible to visually identify which explanatory variables should provide the most discriminant power. One should look for variables having well separated centroïds and clearly defined fringe between classes on the 1D projection of individuals.

Based on the discussion above, the following explanatory variables were sliced from the dataset:

• "smoothness_mean",

• "symmetry_mean",

• "texture_mean",

• "fractal_dimension_mean",

- "texture se",
- "smoothness_se",
- "compactness_se",
- "symmetry_se",

- "fractal_dimension_se",
- "symmetry_worst",
- "fractal_dimension_worst"

That reduces the model to 19 explanatory variables. The discriminant power decreases by only 0.01 compared to the baseline giving 0.763. This simplification seems to be better than the previous one.

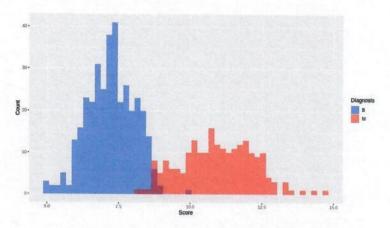


FIGURE 7: Distribution of each class on u₂.

This plot is much more similar to the baseline we defined, with small overlap but higher count where the overlap occurs.

3.4 Decision rule.

As we found a subset of explanatory variables which has a decent discriminant power in regard to the full model, a decision rule can be derived from it. To find the optimal trade-off between false positive and true positive one can inspect the ROC plot. This leads to a

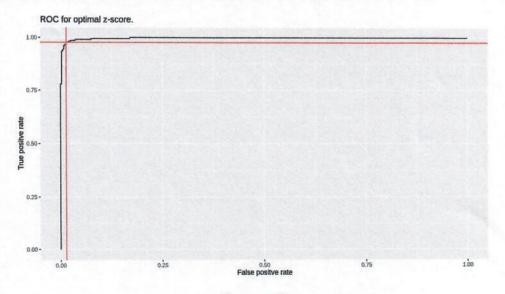


FIGURE 8

normalised z score of 0.128, a false positive rate of 0.014 and a true positive rate of 0.976. The confusion plot below shows good performance with the classification rule.

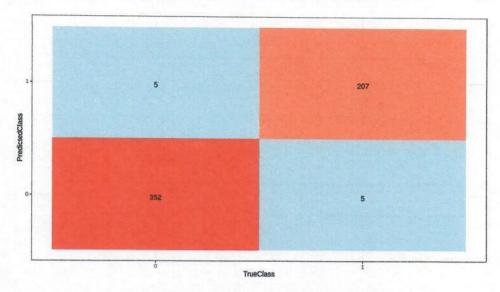


FIGURE 9: Confusion matrix with optimal cutoff of model 2.

3.5 Discuss the homoscedasticity assumption.

First, by looking at plots displaying distribution on **u** of each model it seems classes do not share the same variance/covariance matrix and thus do not verify homoscedasticity assumption.

Some boxplots were displayed in appendices. It is clearly observable that for those explanatory variables the behaviour of the variance is not the same from a class to the other. It is the same for most explanatory variables in the dataset.

Also we tried to perform a statistical test to verify the intuition above. Sadly results we got from bartlett.test function in R that were incoherent. The p-value was way to low compared to what we could expect. We probably misused this tool.(see LDAclass.r script)

If the p-value was not suspicious we could have compared it to a significance level of 0.05 and determine whether or not to accept the null hypothesis: distribution respect homoscedasticity.

- you were expetiting toro praenes as you were concinced that the variances are not simplar!

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