Confidence interval on the cross entropy

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Goal

- True distribution F (density f)
- We observed an iid sample (y_1, \ldots, y_n) from F
- We considered the parametric model $Q_{ heta}$ (density q(y| heta))
- The cross-entropy is

$$\beta = \max_{\theta} \mathbb{E}\left[\log(q(y|\theta))\right] = \max_{\theta} \log(q(y|\theta))f(y)dy$$

Construct a confidence interval on β

Estimate of β

- Clearly, $\beta = \gamma(F)$
- \hat{F}_n the empirical distribution associated to (y_1, \ldots, y_n)
- $\hat{\theta}_n$ the MLE of θ
- An estimate of β is

$$\hat{\beta}_n = \gamma(\hat{F}_n) = \frac{1}{n} \sum_{i=1}^n \log(q(y_i|\hat{\theta}_n))$$

We propose to bootstrap the statistics $S = \hat{\beta}_n - \beta$

A gaussian toy example $F=\mathcal{N}(0,1)$; $Q_{ heta}=\mathcal{N}(heta,1)$

In such a case

•
$$\beta = -1/2 \log(2\pi) - 1/2$$

•
$$\hat{\theta}_n = \bar{y}_n$$
; $\hat{\beta}_n = -1/2\log(2\pi) - 1/(2n)\sum_{i=1}^n (y_i - \bar{y}_n)^2$

■ The bootstrap version of $S = \hat{\beta}_n - \beta = \gamma(\hat{F}_n) - \gamma(F)$ is

$$S^* = 1/(2n) \sum_{i=1}^{n} (y_i - \bar{y}_n)^2 - 1/(2n) \sum_{i=1}^{n} (y_i^* - \bar{y}_n^*)^2 = 1/2s_n^2 - 1/2s_n^*$$

where (y_1^*, \ldots, y_n^*) is a bootstrap sample from (y_1, \ldots, y_n)

Numerical experiments for the gaussian toy example $F = \mathcal{N}(0,1)$; $Q_{\theta} = \mathcal{N}(\theta,1)$

```
beta <-\frac{-1}{2}*\log(2*pi)-\frac{1}{2}
N \leftarrow 1000; n \leftarrow 500; B \leftarrow 500; alpha \leftarrow 0.1
statboot <- rep(0,B); cover <- 0
for (i in 1:N) {
  x \leftarrow rnorm(n) : sn2 \leftarrow var(x)*(n-1)/n
  for (i in 1:B) {
    xstar \leftarrow sample(x,n,rep=TRUE); sn2star \leftarrow var(xstar)*(n-1)/n
     statboot[i] <- sn2/2-sn2star/2 }
  interva <- -1/2*log(2*pi)-sn2/2-quantile(statboot,c(1-alpha/2,alpha/2),
  names=FALSE)
  cover <- cover + (beta >= interva[1] & beta <= interva[2]) }</pre>
cover/N
```

we get 0.903

A gaussian example with unknown variance ; $F=\mathcal{N}(0,1)$; $Q_{(\mu,\sigma^2)}=\mathcal{N}(\mu,\sigma^2)$

In such a case

- $\beta = -1/2\log(2\pi) 1/2$
- $\hat{\theta}_n = (\bar{y}_n, s_n^2)$; $\hat{\beta}_n = -1/2(\log(2\pi) + 1) 1/2\log(s_n^2)$
- The bootstrap version of $S = \hat{\beta}_n \beta = \gamma(\hat{F}_n) \gamma(F)$ is

$$S^* = 1/2\log(s_n) - 1/2\log(s_n^*)$$

Numerical experiments for the gaussian toy example $F=\mathcal{N}(0,1)$; $Q_{(\mu,\sigma^2)}=\mathcal{N}(\mu,\sigma^2)$

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  for (i in 1:B) {
    xstar \leftarrow sample(x,n,rep=TRUE); sn2star \leftarrow var(xstar)*(n-1)/n
     statboot[i] \leftarrow log(sn2)/2-log(sn2star)/2 }
  interva <- -1/2*log(2*pi)-sn2/2-quantile(statboot,c(1-alpha/2,alpha/2),
  names=FALSE)
  cover <- cover + (beta >= interva[1] & beta <= interva[2]) }</pre>
cover/N
```

we get 0.896