Problem Set 2

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Problem 1

```
suppressMessages(library(lmtest))
suppressMessages(library(car))
ts <- read.csv("D:/Dropbox/16summer/Macroeconometrics/Problem Set2/1.csv")</pre>
```

i

Estimate the dynamic version of Okun's law

```
GDP <- ts[,3]
L.GDP <- c(NA,GDP[-239])
L2.GDP <- c(NA,L.GDP[-239])
L3.GDP <- c(NA,L2.GDP[-239])
arima(ts[,4],order = c(2,0,0),xreg =cbind(GDP,L.GDP,L2.GDP))</pre>
```

```
##
## Call:
## arima(x = ts[, 4], order = c(2, 0, 0), xreg = cbind(GDP, L.GDP, L2.GDP))
##
## Coefficients:
##
            ar1
                     ar2 intercept
                                         GDP
                                                L.GDP
                                                        L2.GDP
##
         0.2787 -0.1252
                             1.4694
                                    -0.2036
                                             -0.1393
                                                       -0.0686
                  0.0668
                             0.1083
                                      0.0158
## s.e. 0.0655
                                               0.0154
                                                        0.0155
##
## sigma^2 estimated as 0.7897: log likelihood = -308.36, aic = 630.72
```

Although the magnitude of the coefficients are not the same as the paper, the sign of the coefficients are consistent with the paper.

ii

```
arima(ts[-(2:3),4],order = c(2,0,0),xreg =cbind(GDP,L.GDP,L2.GDP)[-c(2:3),])$aic

## [1] 628.5374

arima(ts[-(2:3),4],order = c(3,0,0),xreg =cbind(GDP,L.GDP,L2.GDP,L3.GDP)[-c(2:3),])$aic

## [1] 629.604
```

```
arima(ts[-(2:3),4],order = c(3,0,0),xreg =cbind(GDP,L.GDP,L2.GDP)[-c(2:3),])$aic
## [1] 630.0607
arima(ts[-(2:3),4],order = c(2,0,0),xreg =cbind(GDP,L.GDP,L2.GDP,L3.GDP)[-c(2:3),])$aic
## [1] 628.6941
arima(ts[-(2:3),4], order = c(1,0,0), xreg = cbind(GDP, L.GDP)[-c(2:3),])$aic
## [1] 645.908
arima(ts[-(2:3),4], order = c(2,0,0), xreg = cbind(GDP, L.GDP)[-c(2:3),])$aic
## [1] 644.9237
arima(ts[-(2:3),4], order = c(1,0,0), xreg = cbind(GDP, L.GDP, L2.GDP)[-c(2:3),])$aic
## [1] 629.4827
The aic of the original model is the smallest. So the author chose the right model.
iii
SSRp \leftarrow sum(arima(ts[,4],order = c(2,0,0),xreg = cbind(GDP,L.GDP,L2.GDP)) \\ \\ sresiduals[-(1:2)]^2)
SSR1 <- sum(arima(ts[1:143,4],order = c(2,0,0),
                   xreg =cbind(GDP,L.GDP,L2.GDP)[1:143,])$residuals[-(1:2)]^2)
SSR2 <- sum(arima(ts[144:239,4],order = c(2,0,0),xreg =cbind(GDP,L.GDP,L2.GDP)[144:239,])$residuals^2)
chow <- ((SSRp-SSR1-SSR2)/(SSR1+SSR2))*((239-2*6)/6)
print(chow)
## [1] 2.961727
qf(0.95,6,(239-6*2))
## [1] 2.138668
The chow test shows that the F-statistic is significant, so the structure is unstable over the pre-1984 and
post-1984.
```

iv

Estimate the difference version of Okun's law

```
lm(ts[,4]~GDP)
##
## Call:
## lm(formula = ts[, 4] ~ GDP)
## Coefficients:
## (Intercept)
                      GDP
##
       0.9968
                  -0.2757
Test stability
D \leftarrow c(rep(0,143), rep(1,(239-143)))
reg <- lm(ts[,4]~D+GDP*D)
coeftest(reg)
##
## t test of coefficients:
##
              Estimate Std. Error t value Pr(>|t|)
##
## (Intercept) 1.200087 0.116497 10.3015 < 2e-16 ***
## D
             ## GDP
## D:GDP
              0.055751 0.057582 0.9682 0.33394
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
linearHypothesis(reg,c("D=0","D:GDP=0"))
## Linear hypothesis test
##
## Hypothesis:
## D = 0
## D:GDP = 0
##
## Model 1: restricted model
## Model 2: ts[, 4] \sim D + GDP * D
##
##
    Res.Df
             RSS Df Sum of Sq
                                  F Pr(>F)
       237 299.90
## 1
       235 288.44 2
## 2
                      11.458 4.6675 0.01028 *
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
Reject the null hypothesis. It is unstable over pre-1984 and post-1984, so we have to calculate the GDP
growth rate separately.
reg1 <- lm(ts[1:143,4]~GDP[1:143])
```

pre-1984 GDP growth rate

 $reg2 \leftarrow lm(ts[144:239,4] \sim GDP[144:239])$

-reg1\$coefficients[1]/reg1\$coefficients[2] ## (Intercept) ## 4.214554 post-1984 GDP growth rate

-reg2\$coefficients[1]/reg2\$coefficients[2]

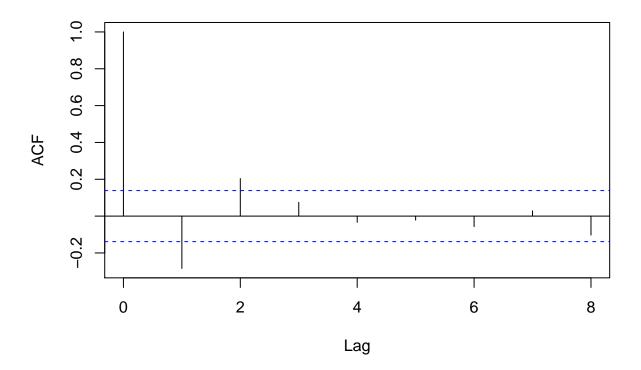
```
## (Intercept)
## 2.585415
```

Problem 2

ii

```
suppressMessages(library(forecast))
masim <- arima.sim(list(ma=c(-0.3,0.17)),n=200)
acf(masim,8)$acf</pre>
```

Series masim



```
##
   , , 1
##
##
                 [,1]
    [1,]
         1.00000000
##
##
    [2,] -0.28383716
##
    [3,] 0.20384244
    [4,] 0.07459754
##
##
    [5,] -0.03351591
##
    [6,] -0.02172824
##
    [7,] -0.05685270
##
    [8,] 0.02833621
    [9,] -0.10289715
```

Yes, it matches the theoretical prediction because the lag 1 and 2 are significant and similar to what we calculated, and all following lags are non-significant.

auto.arima(masim)

##

data: auto.arima(masim)\$residuals[1:8]

X-squared = 0.45316, df = 1, p-value = 0.5008

```
## Series: masim
## ARIMA(2,0,2) with zero mean
##
## Coefficients:
##
            ar1
                     ar2
                              ma1
                                      ma2
##
         0.1180
                 -0.2585
                          -0.3910 0.5665
## s.e. 0.2004
                  0.2178
                           0.1819 0.1620
##
## sigma^2 estimated as 1.012: log likelihood=-283.13
## AIC=576.27
                AICc=576.58
                              BIC=592.76
Box.test(auto.arima(masim)$residuals[1:8],type="Ljung-Box")
##
##
   Box-Ljung test
```

Based on AIC, the best model is not MA(2), and it has a non-significant Q-statistic, so it is a valid model. It is because the sampling error that we cannot fit a simulated model into its exactly original Data Generating Process.

Problem 4

```
library(forecast)
library(tseries)
\mathbf{a}
ts1<- read.csv("D:/Dropbox/16summer/Macroeconometrics/Problem Set2/41.csv")
adf.test(ts1[,2])
##
##
   Augmented Dickey-Fuller Test
##
## data: ts1[, 2]
## Dickey-Fuller = -3.4015, Lag order = 4, p-value = 0.05802
## alternative hypothesis: stationary
Non-stationary. We should allow auto.arima to do difference.
auto.arima(ts1[,2],stationary=FALSE)
## Series: ts1[, 2]
## ARIMA(2,1,1)
##
## Coefficients:
##
            ar1
                    ar2
                             ma1
##
         0.3372 0.2269 -0.9594
## s.e. 0.1099 0.1069 0.0491
## sigma^2 estimated as 5.178: log likelihood=-229.98
## AIC=467.96
              AICc=468.37 BIC=478.5
Box.test(auto.arima(ts1[,2])$residuals,type="Ljung-Box")
##
## Box-Ljung test
## data: auto.arima(ts1[, 2])$residuals
## X-squared = 0.0054382, df = 1, p-value = 0.9412
The final model is ARIMA(2,1,1).
b
ts2<- read.csv("D:/Dropbox/16summer/Macroeconometrics/Problem Set2/42.csv")
adf.test(ts2[,2])
```

```
##
  Augmented Dickey-Fuller Test
##
##
## data: ts2[, 2]
## Dickey-Fuller = -3.1678, Lag order = 6, p-value = 0.09395
## alternative hypothesis: stationary
Non-stationary. We should allow auto.arima to do difference.
auto.arima(ts2[,2],stationary=FALSE)
## Series: ts2[, 2]
## ARIMA(3,1,0)
##
## Coefficients:
##
            ar1
                     ar2
                             ar3
         0.4099 -0.0310 0.2529
##
## s.e. 0.0624 0.0678 0.0623
## sigma^2 estimated as 0.03326: log likelihood=68.87
## AIC=-129.73 AICc=-129.56 BIC=-115.83
Box.test(auto.arima(ts2[,2])$residuals,type="Ljung-Box")
##
## Box-Ljung test
## data: auto.arima(ts2[, 2])$residuals
## X-squared = 0.024162, df = 1, p-value = 0.8765
The final model is ARIMA(3,1,0)
\mathbf{c}
ts3<- read.csv("D:/Dropbox/16summer/Macroeconometrics/Problem Set2/43.csv")
adf.test(ts3[,2])
##
  Augmented Dickey-Fuller Test
##
## data: ts3[, 2]
## Dickey-Fuller = -1.2572, Lag order = 4, p-value = 0.8846
## alternative hypothesis: stationary
Non-stationary. We should allow auto.arima to do difference.
auto.arima(ts3[,2],stationary=FALSE)
```

```
## Series: ts3[, 2]
## ARIMA(0,2,1)
##
## Coefficients:
##
             ma1
         -0.9323
##
        0.0490
## s.e.
##
## sigma^2 estimated as 0.2821: log likelihood=-79.92
## AIC=163.83
               AICc=163.95
                              BIC=169.06
Box.test(auto.arima(ts3[,2])$residuals,type="Ljung-Box")
##
##
   Box-Ljung test
##
## data: auto.arima(ts3[, 2])$residuals
## X-squared = 0.13323, df = 1, p-value = 0.7151
The final model is ARIMA(0,2,1)
Problem 5
library(forecast)
ts<- read.csv("D:/Dropbox/16summer/Macroeconometrics/Problem Set2/ps2prob5_data.csv")
First let's find out what lag length is the best based on AIC.
aic <- rep(NA,10)
for(i in 1:10){
  aic[i] \leftarrow arima(ts[,2], order = c(i,0,0), include.mean = 0)aic
}
aic
    [1] 60.99179 62.57473 54.54597 56.44198 56.29581 58.13532 59.07763
   [8] 60.41373 62.41164 64.39334
Clearly AR(3) is the best based on AIC.
arima(ts[,2],c(3,0,0),include.mean = 0)
##
## Call:
## arima(x = ts[, 2], order = c(3, 0, 0), include.mean = 0)
##
## Coefficients:
##
            ar1
                    ar2
                              ar3
##
         0.2308 0.0072 -0.2607
## s.e. 0.0802 0.0832
                          0.0808
```

$sigma^2$ estimated as 0.08056: log likelihood = -23.27, aic = 54.55

Box.test(arima(ts[,2],c(3,0,0),include.mean = 0)\$residuals,type="Ljung-Box")

```
##
## Box-Ljung test
##
## data: arima(ts[, 2], c(3, 0, 0), include.mean = 0)$residuals
## X-squared = 0.019525, df = 1, p-value = 0.8889
```

AR(3) is valid, but it is actually not the "true" model. In order to get the "true" model AR(8), we have to build our own information criterion.

One way is to correct the parameter of penalty term of lag length in the AIC. We know that when there is no penalty(parameter=0), a higher lag length will always get a higher maximized log-likelihood. When in the AIC the parameter of penalty term is 2, we get the best lag length of 3. So there will definitely be some numbers between 0 and 2 that will make the 8th lag length the best model. So we use numerical computation method to compute the interval of all feasible parameters.

```
loglik <- rep(NA,10)
for(i in 1:10){
    loglik[i] <- arima(ts[,2],order = c(i,0,0),include.mean = 0)$loglik
}
MyIC <- rep(NA,10)
step <- seq(0,2,1e-5)
parameter <- rep(NA,length(step))
l <- c(1:10)
for (j in 1:length(step)) {
    MyIC<- -2*loglik/145*step[j]*l/145
    if(MyIC[8]==min(MyIC)){
        parameter [j] <- step[j]
    }
}
min(subset(parameter,!is.na(parameter)))</pre>
```

[1] 0.0102

```
max(subset(parameter,!is.na(parameter)))
```

```
## [1] 0.62735
```

So here we get our own information criterion as following: IC=-2 \times loglikelihood/T+parameter \times laglength/T where parameter \in [0.0102,0.62735]

```
arima(ts[,2],c(8,0,0),include.mean = 0)
```

```
##
## Call:
## arima(x = ts[, 2], order = c(8, 0, 0), include.mean = 0)
##
## Coefficients:
##
                    ar2
                             ar3
                                                                          ar8
            ar1
                                       ar4
                                               ar5
                                                       ar6
                                                                ar7
##
         0.2196 0.0554
                         -0.2504
                                  -0.0851
                                            0.1077
                                                    0.0583
                                                            -0.0751
                                                                     -0.0702
## s.e. 0.0832 0.0856
                          0.0855
                                   0.0885 0.0879 0.0869
                                                             0.0868
                                                                      0.0860
## sigma^2 estimated as 0.07818: log likelihood = -21.21, aic = 60.41
```

Box.test(arima(ts[,2],c(8,0,0),include.mean = 0)\$residuals,type="Ljung-Box")

```
##
## Box-Ljung test
##
## data: arima(ts[, 2], c(8, 0, 0), include.mean = 0)$residuals
## X-squared = 0.0031579, df = 1, p-value = 0.9552
```

The "true" model AR(8) is valid.