

109550220 Prob HW2 胡景斌

Problem 1

$$(a) P_X(k) = \frac{e^{-\lambda T} (\lambda T)^k}{k!} \quad P_X(\lambda T) = \frac{e^{-\lambda T} (\lambda T)^{\lambda T}}{(\lambda T)!}$$

$$\text{let } \Delta k \geq 0, \in \mathbb{N}, \\ P_X(\lambda T + \Delta k) = \frac{e^{-\lambda T} (\lambda T)^{\lambda T + \Delta k}}{(\lambda T + \Delta k)!} = \frac{e^{-\lambda T} (\lambda T)^{\lambda T} (\lambda T)^{\Delta k}}{(\lambda T)! \prod_{n=1}^{\Delta k} (\lambda T + n)}$$

$$= P_X(\lambda T) \frac{(\lambda T)^{\Delta k}}{\prod_{n=1}^{\Delta k} (\lambda T + n)}, \text{ monotonically decrease as } \Delta k \text{ increase}$$

Similarly

$$P_X(\lambda T - \Delta k) = P_X(\lambda T) \frac{(\lambda T)^{\Delta k}}{\prod_{n=1}^{\Delta k} (\lambda T - n)}, \text{ monotonically non-decrease as } \Delta k \text{ decrease}$$

$$\Rightarrow k^* = \lfloor \lambda T \rfloor = \operatorname{argmax} (k \in \mathbb{N} \cup \{0\} | P_X(k))$$

$$(b) X = \min(X_1, X_2, \dots, X_n)$$

the CDF of X is $F_X(k) = P(X \leq k) = P(\text{at least one of } X_i \leq k)$

$$= 1 - P(\text{all } X_i > k)$$

$$= 1 - P(X_1 > k) P(X_2 > k) \dots P(X_n > k) \text{ (independent)}$$

$$= 1 - P(X_1 > k)^n \text{ (share the same } p)$$

$$= 1 - [1 - \underbrace{F_{X_1}(k)}_{\text{CDF of } X_1}]^n$$

$$p_X(k) = F_X(k) - F_X(k-1) \quad F_{X_1}(1) = 1 - (1-p)^k$$

$$= 1 - [1 - F_{X_1}(k)]^n - \{1 - [1 - F_{X_1}(k-1)]^n\}$$

$$= [1 - F_{X_1}(k-1)]^n - [1 - F_{X_1}(k)]^n$$

$$= \{1 - [1 - (1-p)^{k-1}]\}^n - \{1 - [1 - (1-p)^k]\}^n$$

$$= (1-p)^{n(k-1)} - (1-p)^{nk}$$

$$= (1-p)^{n(k-1)} (1 - (1-p)) = p(1-p)^{n(k-1)}$$

X is a Geometric Random Variable

Problem 2

let $X_i = \#$ of balls in box i , $i \in \mathbb{N}$

$X_1 + X_2 + \dots + X_n = r$, there are $H_r^n = \binom{n+r-1}{r}$ cases

(a) if $X_1 = k \in \mathbb{N} \cup \{0\}$

$X_2 + X_3 + \dots + X_n = r-k$, there are $H_{r-k}^{n-1} = \binom{(n-1)+r-k-1}{r-k}$ cases

thus $q_k = \frac{\binom{n+r-k-2}{r-k}}{\binom{n+r-1}{r}}$ (since all cases have equal probabilities)

$$\begin{aligned} (b) \quad q_k &= \frac{(n+r-k-2)!}{(r-k)!} \cdot \frac{r!}{(n+r-1)!} = \frac{(n+r-1-k-1)!}{(n+r-1)!} \times \frac{r!}{(r-k)!} \\ &= \frac{r(r-1) \dots (r-k+1)}{(n+r-1)(n+r-2) \dots (n+r-1-k)} = \prod_{i=0}^{k-1} \left(\frac{r-i}{n+r-1-i} \right) \times \left(\frac{1}{r-k} \right) \end{aligned}$$

$$\lim_{\substack{n \rightarrow \infty \\ r \rightarrow \infty \\ n \rightarrow \infty}} q_k = \lim_{\substack{n \rightarrow \infty \\ r \rightarrow \infty \\ n \rightarrow \infty}} \prod_{i=0}^{k-1} \left(\frac{\frac{r}{n} - \frac{1}{n}}{\frac{n}{n} + \frac{r}{n} - \frac{1}{n} - \frac{i}{n}} \right) \times \left(\frac{\frac{1}{r}}{\frac{r}{n} - \frac{k}{n}} \right), \quad \frac{r}{n} = \lambda \text{ as } r \rightarrow \infty, n \rightarrow \infty$$

$$= \prod_{i=0}^{k-1} \left(\frac{\lambda}{1-\lambda} \right) \times \frac{1}{\lambda} = \left(\frac{\lambda}{1-\lambda} \right)^{k+1} \times \left(\frac{1}{\lambda} \right) = \frac{\lambda^k}{(1-\lambda)^{k+1}}$$

X in the limit is a Geometric Random Variable

Problem 3

(a) let V = total transmitted bits

$$p_v(k) = \frac{e^{-\lambda T} (\lambda T)^k}{k!}$$

X = # of "1" transmitted

$$p_x(k) = P(X=k)$$

$$= \sum_{n=0}^{\infty} P(X=k | V=k+n) \cdot P(V=k+n)$$

$$= \sum_{n=0}^{\infty} C_k^{k+n} p^k (1-p)^n \times p_v(k+n)$$

$$= \sum_{n=0}^{\infty} \frac{(k+n)!}{k! n!} p^k (1-p)^n \times \frac{e^{-\lambda T} (\lambda T)^{k+n}}{(k+n)!}$$

$$= e^{-\lambda T} \frac{p^k (\lambda T)^k}{k!} \sum_{n=0}^{\infty} \frac{(1-p)^n (\lambda T)^n}{n!}$$

$$= e^{-\lambda T} \frac{p^k (\lambda T)^k}{k!} e^{(1-p)\lambda T} = \frac{e^{-\lambda p T} (\lambda p T)^k}{k!}$$

\Rightarrow average rate is λp

(b) Similarly, let X' = # of 0's transmitted

$$p_x(n) = P(X'=n)$$

$$= \sum_{k=0}^{\infty} P(X'=n | V=k+n) \cdot P(V=k+n)$$

$$= \frac{e^{-\lambda T (1-p)} (\lambda T (1-p))^n}{n!}$$

$$p_x(Y) = P(\text{receive 1} | \text{sent 1}) + P(\text{receive 1} | \text{sent 0})$$

$$= (1-\alpha_0) \frac{e^{-\lambda p T} (\lambda p T)^k}{k!} + \alpha_1 \frac{e^{-\lambda T (1-p)} (\lambda T (1-p))^n}{n!}$$

Problem 4

$$\begin{aligned} (a) E[X] &= \sum_{k=1}^{\infty} k \cdot (1-p)^{k-1} p \\ &= p \left(\sum_{k=1}^{\infty} (1-p)^{k-1} + \sum_{k=2}^{\infty} (1-p)^{k-1} + \sum_{k=3}^{\infty} (1-p)^{k-1} + \dots \right) \\ &= p \left[\frac{1}{p} + \frac{(1-p)}{p} + \frac{(1-p)^2}{p} + \dots \right] = 1 + (1-p) + (1-p)^2 + \dots = \lim_{n \rightarrow \infty} \frac{1 - (1-p)^{n+1}}{1 - (1-p)} = \frac{1}{p} \end{aligned}$$

$$\text{Var}[X] = E[X^2] - E[X]^2 = \frac{2-p}{p^2} - \frac{1}{p^2} = \frac{1-p}{p^2}$$

$$E[X^2] = \sum_{k=1}^{\infty} k^2 q^{k-1} p \quad (q = 1-p)$$

$$= p \sum_{k=1}^{\infty} \frac{d}{dq} (kq)^k = p \frac{d}{dq} \sum_{k=1}^{\infty} (kq)^k = p \frac{d}{dq} \left\{ \frac{q}{1-q} E[X] \right\}$$

$$= p \frac{d}{dq} \left\{ q(1-q)^{-2} \right\} = p \left\{ \frac{1}{(1-q)^2} - \frac{2q}{(1-q)^3} \right\}$$

$$= \frac{1}{p} + \frac{2(1-p)}{p^2} = \frac{2-p}{p^2}$$

(b) Since $E[X^m] = E[Y^m]$ for all $m \in \mathbb{N}$

X, Y share the same moment generating function

let $S = \{a_1, a_2, \dots, a_n\}$

$$\Rightarrow \sum_{x \in S} e^{tx} p_X(x) = \sum_{y \in S} e^{ty} p_Y(y)$$

let $A = e^t$

$$\sum_{x \in S} A^x p_X(x) - \sum_{y \in S} A^y p_Y(y) = 0$$

$$\sum_{x \in S} A^x p_X(x) - \sum_{x \in S} A^x p_Y(x) = 0$$

$$A^x \sum_{x \in S} p_X(x) - p_Y(x) = 0$$

$\Rightarrow p_X(x) = p_Y(x) \forall x \in S$, the PMF of X, Y are the same

(c)

$$E[z^2] = \sum_{n=1}^{\infty} z_n^2 \frac{6}{\pi^2 n^2} = \frac{6}{\pi^2} \sum_{n=1}^{\infty} \frac{n}{n^2} = \frac{6}{\pi^2} \sum_{n=1}^{\infty} \frac{1}{n} \Rightarrow \text{diverges}$$

$\text{Var}[z]$ does not exist

$$E[z^3] = \sum_{n=1}^{\infty} z_n^3 = \frac{6}{\pi^2} \sum_{n=1}^{\infty} (-1)^n \frac{1}{\sqrt{n}}$$

$$\text{Since } 0 < \frac{1}{\sqrt{n+1}} < \frac{1}{\sqrt{n}}, \quad \lim_{n \rightarrow \infty} \frac{1}{\sqrt{n}} = 0$$

the series converge by alternating series test, $E[z^3]$ exist

$$E[z^{10}] = \sum_{n=1}^{\infty} z_n^{10} = \frac{6}{\pi^2} \sum_{n=1}^{\infty} n^3, \text{ the series diverge } E[z^{10}] \text{ does not exist}$$