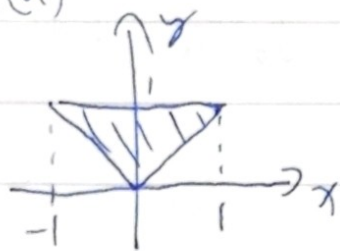


109550020 Probability HW3 胡景弘

1. (a)



$$f_X(x) = \begin{cases} \int_x^1 1 dy = 1-x, & \text{if } 0 < x < 1 \\ \int_{-x}^1 1 dy = 1+x, & \text{if } -1 < x < 0 \\ 0, & \text{otherwise} \end{cases}$$

$$f_Y(y) = \begin{cases} \int_{-y}^y 1 dx = 2y, & \text{if } 0 < y < 1 \\ 0, & \text{otherwise} \end{cases}$$

$$E[X] = \int_{-1}^0 x(1+x) dx + \int_0^1 x(1-x) dx = \left[ \frac{x^2}{2} + \frac{x^3}{3} \right]_{-1}^0 + \left[ \frac{x^2}{2} - \frac{x^3}{3} \right]_0^1$$

$$= -\frac{1}{2} + \frac{1}{3} + \frac{1}{2} - \frac{1}{3} = 0$$

$$E[Y] = \int_0^1 y \cdot 2y dy = \frac{2}{3} y^3 \Big|_0^1 = \frac{2}{3}$$

$$E[XY] = \int_0^1 y \int_{-y}^y x \cdot 1 dx dy = \int_0^1 y \left[ \frac{x^2}{2} \right]_{-y}^y dy = \int_0^1$$

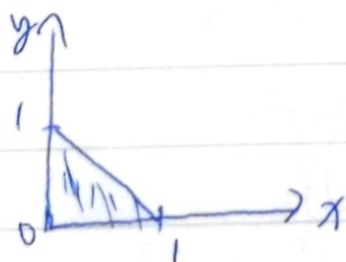
$$= \int_0^1 \frac{y^3}{2} - \frac{y^3}{2} dy = 0$$

$$E[XY] = E[X]E[Y] = 0$$

$$(b) f_{XY}\left(\frac{1}{4}, \frac{1}{2}\right) = 1 \quad f_X\left(\frac{1}{4}\right) = 1 - \frac{1}{4} = \frac{3}{4} \quad f_Y\left(\frac{1}{2}\right) = 2 \times \frac{1}{2} = 1$$

$$f_{XY}\left(\frac{1}{4}, \frac{1}{2}\right) = 1 \neq \frac{3}{4} = f_X\left(\frac{1}{4}\right) f_Y\left(\frac{1}{2}\right)$$

2. (a)



$$\frac{1 \times 1 \times 1}{2} = 1 \quad f_{XY}(x,y) = \begin{cases} 2, & \text{if } x > 0, y > 0, x+y < 1 \\ 0, & \text{otherwise} \end{cases}$$

$$D=2$$

$$f_X(x) = \begin{cases} \int_0^{1-x} 2 dy = 2(1-x), & \text{if } x > 0, y > 0, x+y < 1 \\ 0, & \text{otherwise} \end{cases}$$

$$f_Y(y) = \begin{cases} \int_0^{1-y} 2 dx = 2(1-y), & \text{if } x > 0, y > 0, x+y < 1 \\ 0, & \text{otherwise} \end{cases}$$

$$f_{X|Y}(x|y) = \frac{2}{2(1-y)} = \frac{1}{1-y}$$

$$(b) E[X|Y=y] = \int_0^1 x \frac{1}{1-y} dx = \left[ \frac{1}{2} \frac{x^2}{1-y} \right]_0^1 = \frac{1}{2(1-y)}$$

$$E[X] = E[E[X|Y=y]] = \int_0^1 E[X|Y=y] \cdot f_Y(y) dy$$

$$= \int_0^1 \frac{2(1-y)}{2(1-y)} dy = y \Big|_0^1 = 1$$

3. (c)  $X \sim U(-1, 3)$

$$M_X(t) = E[e^{tx}] = \int_{-1}^3 \frac{1}{4} e^{tx} dx = \frac{1}{4t} e^{tx} \Big|_{-1}^3 = \frac{e^{3t} - e^{-t}}{4t}$$

$$M_X'(t) = \frac{1}{4t^2} \left[ (3e^{3t} + e^{-t})4t - (e^{3t} - e^{-t})4 \right]$$

$$= \frac{1}{4t^2} \left[ (3t-1)e^{3t} + (t+1)e^{-t} \right]$$

$$M_X'(0) = \lim_{t \rightarrow 0} M_X'(t) = \lim_{t \rightarrow 0} \frac{(3t-1)e^{3t} + (t+1)e^{-t}}{4t^2} \left( \frac{0}{0} \right)$$

$$\stackrel{1^{\text{st}}}{=} \frac{3e^{3t} + 3(3t-1)e^{3t} + e^{-t} + -(t+1)e^{-t}}{8t} \left( \frac{0}{0} \right)$$

$$\stackrel{1^{\text{st}}}{=} \frac{9e^{3t} + 9e^{3t} + 9(3t-1)e^{3t} - e^{-t} + (t+1)e^{-t}}{8} = \frac{9+9-9-1-1+1}{8} = 1$$

$$\begin{aligned}
 M_X''(t) &= \frac{1}{4t^4} \left[ (3e^{3t} + 3(3t-1)e^{3t} + e^{-t} - (t+1)e^{-t})t^2 \right. \\
 &\quad \left. - 2t[(3t-1)e^{3t} + (t+1)e^{-t}] \right] \\
 &= \frac{1}{4t^3} \left[ 3te^{3t} + 9t^2e^{3t} - 3te^{3t} + te^{-t} - t^2e^{-t} - te^{-t} \right. \\
 &\quad \left. - 6te^{3t} + 2e^{3t} - 2te^{-t} - 2e^{-t} \right] \\
 &= \frac{1}{4t^3} \left[ (9t^2 - 6t + 2)e^{3t} - (t^2 + 2t + 2)e^{-t} \right]
 \end{aligned}$$

$$M_X''(0) = \lim_{t \rightarrow 0} M_X'(t) = \lim_{t \rightarrow 0} \frac{1}{4t^3} \left[ (9t^2 - 6t + 2)e^{3t} - (t^2 + 2t + 2)e^{-t} \right] \left( \frac{0}{0} \right)$$

$$\stackrel{1^H}{=} \lim_{t \rightarrow 0} \frac{1}{12t^2} \left[ (18t - 6)e^{3t} + 3(9t^2 - 6t + 2)e^{3t} \right. \\ \left. - (2t + 2)e^{-t} + (t^2 + 2t + 2)e^{-t} \right] \left( \frac{0}{0} \right)$$

$$\stackrel{1^H}{=} \lim_{t \rightarrow 0} \frac{1}{24t} \left[ 18e^{3t} + 3(18t - 6)e^{3t} + 3(18t - 6)e^{3t} + 9(9t^2 - 6t + 2)e^{3t} \right. \\ \left. - 2e^{-t} + (2t + 2)e^{-t} + (2t + 2)e^{-t} - (t^2 + 2t + 2)e^{-t} \right] \left( \frac{0}{0} \right)$$

$$\stackrel{1^H}{=} \lim_{t \rightarrow 0} \frac{1}{24} \left[ 54e^{3t} + 2(54e^{3t} + 9(18t - 6)e^{3t}) + 9(18t - 6)e^{3t} + 2(9t^2 - 6t + 2)e^{3t} \right. \\ \left. + 2e^{-t} + 2(2e^{-t} - (2t + 2)e^{-t}) - (2t + 2)e^{-t} + (t^2 + 2t + 2)e^{-t} \right]$$

$$= \frac{1}{24} (54 + (54 - 54) \times 2 + 54 - 54 + 2 + 2(2 - 2) - 2 + 2) = \frac{56}{24} = \frac{7}{3}$$

$$E[X] = 1, \quad E[X^2] = \frac{7}{3}, \quad \text{Var}[X] = \frac{7}{3} - 1^2 = \frac{4}{3}$$

$$(b) M_Y(t) = E[e^{tY}] = \sum_{k=1}^{\infty} \frac{6}{\pi^2 k^2} e^{tk}$$

$$\text{for } t \neq 0, \quad \lim_{k \rightarrow \infty} \left| \frac{\text{Pr}(k+1)e^{t(k+1)}}{\text{Pr}(k)e^{tk}} \right| = \frac{(k+1)^2}{k^2} e^t > 1, \quad M_Y(t) \text{ does not exist}$$

(series converge by ratio test)



$$4. (a) M_X(t) = \left(\frac{1}{3}e^t + \frac{2}{3}\right)^5$$

$$= \binom{5}{0} \left(\frac{2}{3}\right)^5 \left(\frac{1}{3}e^t\right)^0 + \binom{5}{1} \left(\frac{2}{3}\right)^4 \left(\frac{1}{3}e^t\right)^1 + \dots + \binom{5}{5} \left(\frac{2}{3}\right)^0 \left(\frac{1}{3}e^t\right)^5$$

$$p_X(k) = \begin{cases} \binom{5}{k} \frac{2^{5-k}}{3^5}, & \text{if } k = 0, 1, 2, 3, 4, 5 \\ 0, & \text{otherwise} \end{cases}$$

$$(b) M_X(t) = e^{5(e^t - 1)} \quad X \sim \text{poisson}(5)$$

$$p_X(k) = \frac{e^{-5} 5^k}{k!} \quad \text{if } k = 0, 1, 2, \dots$$

$$5. X_1 = \sigma_1 Z + \mu_1 = g_1(z, w)$$

$$X_2 = \sigma_2(\rho Z + \sqrt{1-\rho^2}W) + \mu_2 = g_2(z, w)$$

$$Z = g_1^{-1}(x, w) = \frac{x_1 - \mu_1}{\sigma_1}$$

$$W = g_2^{-1}(z, w) = \frac{\frac{x_2 - \mu_2}{\sigma_2} - \rho \frac{x_1 - \mu_1}{\sigma_1}}{\sqrt{1-\rho^2}}$$

$$J = \begin{vmatrix} \frac{1}{\sigma_1} & 0 \\ -\rho & \frac{1}{\sqrt{1-\rho^2}\sigma_2} \end{vmatrix} = \frac{1}{\sqrt{1-\rho^2}\sigma_1\sigma_2}$$

$$f_{ZW}(z, w) = f_z(z) \cdot f_w(w) = \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}} \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{w^2}{2}} = \frac{1}{2\pi} e^{-\frac{1}{2}(z^2 + w^2)}$$

$$= \frac{1}{2\pi} \exp \left\{ -\frac{1}{2} \left( \frac{(x_1 - \mu_1)^2}{\sigma_1^2} + \frac{(x_2 - \mu_2)^2}{\sigma_2^2} + \frac{1}{1 - \rho^2} \left( \frac{(x_1 - \mu_1)^2}{\sigma_1^2} - \frac{2\rho(x_1 - \mu_1)(x_2 - \mu_2)}{\sigma_1 \sigma_2} + \frac{\rho^2(x_1 - \mu_1)^2}{\sigma_1^2} \right) \right) \right\}$$

$$= \frac{1}{2\pi} \exp \left[ -\frac{1}{2(1-\rho^2)} \left( \frac{(x_1 - \mu_1)^2}{\sigma_1^2} - \frac{2\rho(x_1 - \mu_1)(x_2 - \mu_2)}{\sigma_1 \sigma_2} + \frac{(x_2 - \mu_2)^2}{\sigma_2^2} \right) \right]$$

$$f_{X_1, X_2}(x_1, x_2) = J \cdot f_{ZW}(z, w) = \frac{1}{2\pi \sigma_1 \sigma_2 \sqrt{1-\rho^2}} \exp \left[ -\frac{\left( \frac{(x_1 - \mu_1)^2}{\sigma_1^2} - 2\rho \frac{(x_1 - \mu_1)(x_2 - \mu_2)}{\sigma_1 \sigma_2} + \frac{(x_2 - \mu_2)^2}{\sigma_2^2} \right)}{2(1-\rho^2)} \right]$$

Q.E.D

$$6. E[|XY|] \leq E[|X|^p]^{\frac{1}{p}} E[|Y|^q]^{\frac{1}{q}} \quad (1)$$

$$\Rightarrow \sum_{x \in X, y \in Y} |xy| \leq \left( \sum_{x \in X} |x|^p \right)^{\frac{1}{p}} \left( \sum_{y \in Y} |y|^q \right)^{\frac{1}{q}} \quad (2) \quad \text{holds homogeneously i.e.}$$

$$\begin{aligned} \Rightarrow \sum_{x \in X, y \in Y} |\alpha x (\beta y)| &= |\alpha \beta| \sum_{x \in X, y \in Y} |xy| \\ &\leq |\alpha \beta| \left( \sum_{x \in X} |x|^p \right)^{\frac{1}{p}} \left( \sum_{y \in Y} |y|^q \right)^{\frac{1}{q}} = \left( \sum_{x \in X} |\alpha x|^p \right)^{\frac{1}{p}} \left( \sum_{y \in Y} |\beta y|^q \right)^{\frac{1}{q}} \quad (3) \end{aligned}$$

$$(3) \text{ holds when } X \text{ and } Y \text{ satisfying } \sum_{x \in X} |x|^p = 1 = \sum_{y \in Y} |y|^q \quad (4)$$

$$\text{then we can reduce (2) to } \sum_{x \in X, y \in Y} |xy| \leq 1$$

$$\begin{aligned} \text{by Young's inequality, } \sum_{x \in X, y \in Y} |xy| &\leq \sum_{x \in X, y \in Y} \left( \frac{|x|^p}{p} + \frac{|y|^q}{q} \right) \\ &= \frac{1}{p} \sum_{x \in X} |x|^p + \frac{1}{q} \sum_{y \in Y} |y|^q \end{aligned}$$

$$\stackrel{\text{by (4)}}{=} \frac{1}{p} + \frac{1}{q} = 1$$

Q. E. D.