# ILAA Tutorial

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## 1 Introduction

Iterative Linear Association Analysis (ILAA) is a computational method that estimates the Exploratory Residualization Transform (ERT) as seen in Figure 1, and Appendix A describes in detail the ILAA algorithm. ERT is estimated from a sample of multidimensional data. and mitigates muticollinearity issues via variable residualization.

The dataframe with reduced multicollinarity issues, Q, is estimated by:

$$Q = WX$$

where X is the observed data frame and W is the Exploratory Residualization Transform (ERT) matrix.

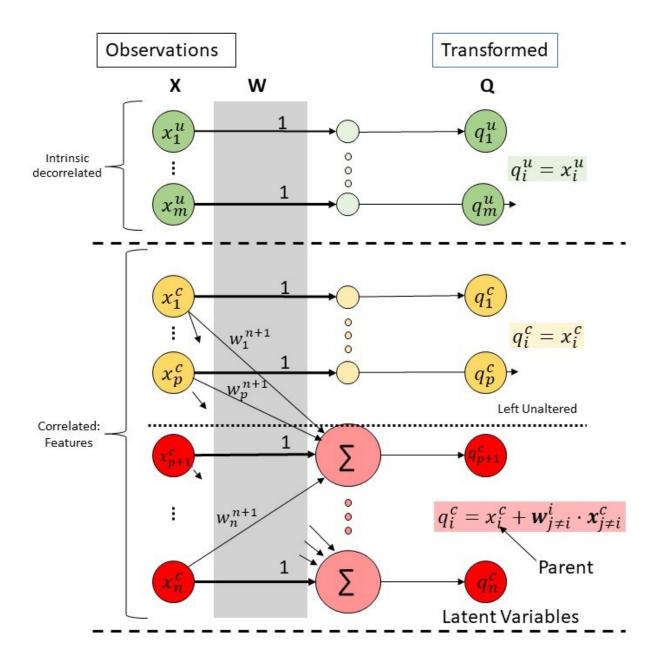


Figure 1: The UPLTM: any linear transformation matrix that does not reduce data dimensionality and has a one-to-one association between units of measurement, i.e., it is preserving the space metric. The ERT is special form of UPLTM that is preserving the metric, and aims to create latent variables with controlled association between them.

The returned transformation matrix can be used to:

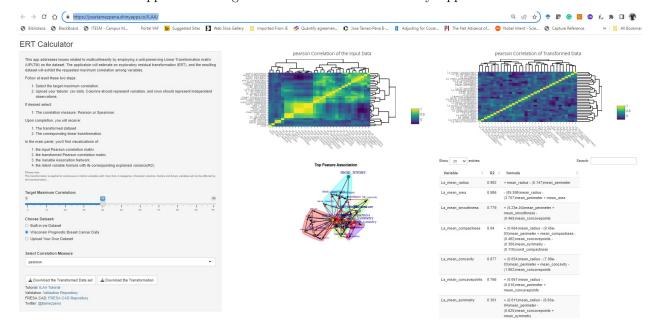
- 1. Do an exploratory analysis of latent variables and their association to the observed variables
- 2. Do exploratory discovery of latent variables associated with a specific outcome-target
- 3. Addressing multicollinearity issues in regression models
  - 1. Better estimation and interpretation of model variables
  - 2. Improve linear model performance
- 4. Simplify the multidimensional search space for many ML algorithms

The objective of this tutorial is to guide users in using the ILAA to effectively accomplish the aforementioned tasks. The tutorial will showcase:

- Transform a data frame affected by data multicollinearity into a new a data frame with a maximum degree of data correlation among variables
- Visualize the transformation matrix
- Explore the returned formulas for each one of the returned latent variables
- Understand and interpret the returned latent variables
- Use ILAA as a pre-processing step to model a specific target outcome using linear models
  - Explore the model in the transformed space
  - Get the observed variables coefficients.

## 1.1 Shiny App

Users can test ILLA application using the ERT calculator: ILLA Shiny App:



#### 1.2 The Libraries

ILAA is a wrapper of the more general method of data decorrelation algorithm (IDeA) implemented in R, and both are part of the FRESA.CAD 3.4.6 package.

```
## From git hub
#First install package devtools
#library(devtools)
#install_github("joseTamezPena/FRESA.CAD")

## For ILAA
library("FRESA.CAD")

## For network analysis
library(igraph)

## For multicollinearity
library(multiColl)
library(car)
library("colorRamps")
```

## 2 Test Data

For this tutorial I'll use the body-fat prediction data set. The data was downloaded from Kaggle:

https://www.kaggle.com/datasets/fedesoriano/body-fat-prediction-dataset

The Kaggle data disclaimer:

"Source The data were generously supplied by Dr. A. Garth Fisher who gave permission to freely distribute the data and use for non-commercial purposes.

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email address: rwjohnso@silver.sdsmt.edu web address: http://silver.sdsmt.edu/~rwjohnso"

#### 2.1 Loading the Data

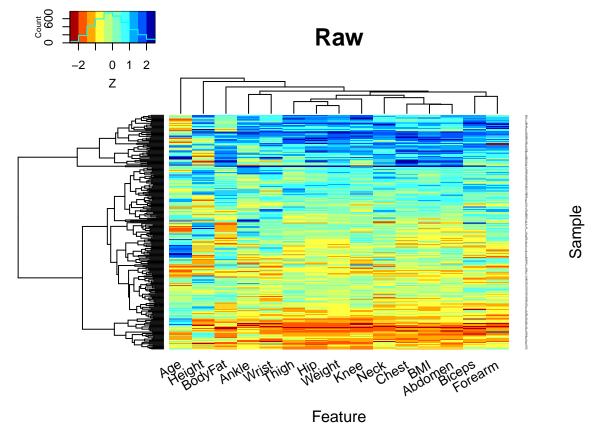
The BodyFat dataset contains the density information, a not direct measurement. In this tutorial, we will remove the density and we will try to model the body fat based on antrophometry.

The following code snippet loads the data and removes the density information from the data. It also computes the Body Mass Index (BMI)

#### 2.1.1 The Heatmap of the Raw Data

Now, here we show the heatmap of the dataframe:

```
bkcolors <-seq(-2.5, 2.5, by = 0.5)
smap <- FRESAScale(body_fat,method="OrderLogit")$scaledData</pre>
```



par(op)

# 3 ILAA Unsupervised Processing

The ILLA function is defined as follows:

```
bootstrap=0
)
```

#### where:

- data: The source data-frame
- thr: The target correlation goal.
- method: Defines the correlation measure
- Outcome: The name of the target variable, and it is required for supervised learning
- drivingFeatures: Defines a set of variables that are aimed to be basis unaltered vectors
- maxLoops: The maximum number of iterations cycles
- verbose: Display the evolution of the algorithm.
- bootstrap: The number of bootstrap estimations. (True bootstrap when n>500, 5% constrained resampling at n<=500)

At return of the ILLA function is a decorrelated dataframe that shares the same dimensions as the input dataframe. The dataframe has the following attributes:

```
RTM <- attr(decorrelatedData,"UPLTM")
fscore <- attr(decorrelatedData,"fscore");
drivingFeatures <- attr(decorrelatedData,"drivingFeatures");
adjustedpvalue <- attr(decorrelatedData,"unipvalue")
RCritical <- attr(decorrelatedData,"R.critical")
EvolutionData <- attr(decorrelatedData,"IDeAEvolution")
VarRatio <- attr(decorrelatedData,"VarRatio")</pre>
```

#### Attributes details:

- UPLTM: The UPLTM matrix that can be used to decorrelated or analyze variables associations
- fscore: A numeric vector with the final feature score of each analyzed variable. The fscore contains the number of times a variable was used as an independent variable minus the times it was a dependent variable.
- drivingFeatures: The ordered character vector indicating the hierarchy of the variables for tiebreak
- unipvalue: The adjusted p-values used to define a true variable-to-variable association inside the linear modeling
- R.critical: The pearson R critical value used to filter-out false association between variables.
- IDeAEvolution: A list with two elements:
  - Corr: The evolution of the maximum observed correlation.
  - Spar: The evolution of the matrix sparcity.
- VarRatio: A vector indicating the ratio of the observed variance explained by the latent variable model.

### 3.1 ILLA Auxiliary Functions

FRESA.CAD provide the following auxiliary functions:

```
newTransformedData <- predictDecorrelate(decorrelatedData,NewData)
theBetaCoefficientts <- getLatentCoefficients(decorrelatedData)
fromLatenttoObserved <- getObservedCoef(decorrelatedData,latentModel)</pre>
```

- predictDecorrelate() Rotates any new data set based on the output of an ILAA transformed data set.
- getLatentCoefficients() Returns a list of all the beta coefficients for each one of the discovered latent variables. The attribute: "LatentCharFormulas" returns a list of the character string of the corresponding latent variable formula.
- getObservedCoef() returns the beta coefficients on the observed space of any linear model that was trained on the UPLTM space.

## 3.2 Sample Usage

By default, the ILAA function will target a correlation lower than 0.8 using the Pearson correlation measure. But user has the freedom to chose between robust fitting with Spearman correlation measure, and/or set the level of feature association by lowering the threshold. The following snippet shows the different options.

```
# Default call
body_fat_Decorrelated <- ILAA(body_fat)
varRatio_D <- attr(body_fat_Decorrelated, "VarRatio")
yCor_D <- attr(body_fat_Decorrelated, "IDeAEvolution") $Corr
ySpar_D <- attr(body_fat_Decorrelated, "IDeAEvolution") $Spar

# Explore the convergence metrics in verbose mode
body_fat_Decorrelated <- ILAA(body_fat, verbose=TRUE)</pre>
```

fast | LM | Weight Body Fat Age Weight Height Neck Chest 0.40000000 0.06666667 1.00000000 0.13333333 0.53333333 0.73333333

```
Included: 15, Uni p: 0.01, Base Size: 1, Rcrit: 0.1467743
```

```
 \begin{array}{l} 1<R=0.890, thr=0.900>, \ Top: \ 2<1>Fa=2, <|><>Tot \ Used: 5 \ , \ Added: 3 \ , \ Zero \ Std: 0 \ , \ Max \ Cor: 0.888 \ 2 \ <R=0.861, thr=0.800>, \ Top: \ 1<5>Fa=2, <|><>Tot \ Used: 9 \ , \ Added: 5 \ , \ Zero \ Std: 0 \ , \ Max \ Cor: 0.860 \ 3 \ <R=0.860, thr=0.800>, \ Top: \ 1<1>Fa=2, <|><>Tot \ Used: 10 \ , \ Added: 1 \ , \ Zero \ Std: 0 \ , \ Max \ Cor: 0.959 \ 4 \ <R=0.959, thr=0.950>, \ Top: \ 1<1>Fa=2, <|><>Tot \ Used: 10 \ , \ Added: 1 \ , \ Zero \ Std: 0 \ , \ Max \ Cor: 0.735 \ 5 \ <R=0.735, thr=0.800> \ [5], \ 0.4782625 \ Decor \ Dimension: \ 10 \ Nused: \ 10 \ . \ Cor \ to \ Base: \ 7 \ , \ ABase: \ 15 \ , \ Outcome \ Base: 0 \end{array}
```

```
# Robust Linear Fitting with the Spearman correlation measure
body_fat_Decorrelated <- ILAA(body_fat,method="spearman")</pre>
varRatio_S <- attr(body_fat_Decorrelated, "VarRatio")</pre>
yCor S <- attr(body fat Decorrelated, "IDeAEvolution") $Corr
ySpar_S <- attr(body_fat_Decorrelated,"IDeAEvolution")$Spar</pre>
# Lowering the threshold
body_fat_Decorrelated <- ILAA(body_fat,thr=0.4)</pre>
varRatio_P_40 <- attr(body_fat_Decorrelated, "VarRatio")</pre>
yCor_P_40 <- attr(body_fat_Decorrelated, "IDeAEvolution") $Corr
ySpar_P_40 <- attr(body_fat_Decorrelated,"IDeAEvolution")$Spar</pre>
# Trying to achieve the maximum independence
# between variables, i.e., thr=0.0
body_fat_Decorrelated <- ILAA(body_fat,thr=0.0)</pre>
varRatior P 00 <- attr(body fat Decorrelated, "VarRatio")</pre>
yCor_P_00 <- attr(body_fat_Decorrelated, "IDeAEvolution") $Corr
ySpar_P_00 <- attr(body_fat_Decorrelated,"IDeAEvolution")$Spar</pre>
```

#### 3.2.1 Latent Models Variance Ratios

Every change in parameters will create different solutions of the ERT transform.

Here we will check the variance ratio of each latent model. Where the variance ratio can be interpreted as the percentage of the observed variance still present in the latent variable.

Note: Every time the variance ratio is 1, is an indication that the observed variable was not modeled by any other variable in the dataframe.

Table 1: Unexplained variance ratio of latent models (continued below)

	BodyFat	Age	Ankle	Forearm	Wrist	Weight
Default	1.000	1.000	1.000	1.000	1.000	1.0000
Spearman	1.000	1.000	1.000	1.000	1.000	1.0000
$\mathbf{At}\mathbf{\underline{40}}$	0.309	1.000	0.624	0.602	0.459	1.0000
$\mathbf{At}\mathbf{\underline{0}}$	0.290	0.483	0.540	0.518	0.403	1.0000
${f BMI\_BF}$	0.475	0.491	0.583	0.517	0.386	0.2108
${\bf BodyFat\_Driven}$	1.000	0.494	0.565	0.513	0.383	0.0687

Table 2: Table continues below

	Biceps	Neck	Knee	Thigh	BMI	Chest	Abdomen
Default	0.359	0.303	0.272	0.242	0.211	0.1710	0.1500
Spearman	1.000	0.303	0.273	0.242	0.212	0.1737	0.1500
$At\_40$	0.359	0.303	0.272	0.194	0.211	0.1085	0.1500
$\mathbf{At}\mathbf{\_0}$	0.320	0.269	0.228	0.183	0.211	0.1007	0.1277
$\mathbf{BMI}\mathbf{\_BF}$	0.308	0.294	0.243	0.234	1.000	0.0980	0.0734
BodyFat_Driven	0.315	0.367	0.256	0.190	0.124	0.0984	1.0000

	Hip	Height
Default	0.1096	0.0209
Spearman	0.1099	0.0221
$At\_40$	0.1096	0.0209
$\mathbf{At}\mathbf{\underline{0}}$	0.0993	0.0209
$\mathbf{BMI}\mathbf{\_BF}$	0.0779	0.0209
${f BodyFat\_Driven}$	0.2344	0.0210

#### 3.2.2 Plotting the Evolution

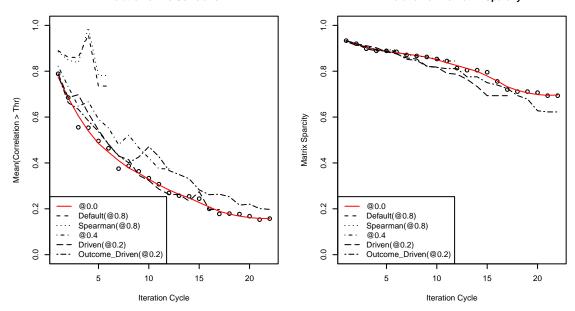
Here we will use the attr(dataTransformed,"IDeAEvolution") to plot the evolution of the correlation measure and the evolution of the matrix sparsity.

```
par(mfrow=c(1,2),cex=0.5)
# Correlation
yval <- yCor_P_00</pre>
xidx <- c(1:length(yval))</pre>
plot(xidx,yval,
     xlab="Iteration Cycle",
     ylab="Mean(Correlation > Thr)",
     ylim=c(0,1.0),
     main="Evolution of the Correlation")
  lfit <-try(loess(yval~xidx,span=0.5));</pre>
  if (!inherits(lfit,"try-error"))
    plx <- try(predict(lfit,se=TRUE))</pre>
    if (!inherits(plx,"try-error"))
      lines(xidx,plx$fit,lty=1,col="red")
    }
 }
lines(xidx,yCor_D[xidx],lty=2)
lines(xidx,yCor_S[xidx],lty=3)
lines(xidx,yCor_P_40[xidx],lty=4)
lines(xidx,yCor_P_20_D[xidx],lty=5)
lines(xidx,yCor_P_20_OD[xidx],lty=6)
legend("bottomleft",
       legend=c("@0.0","Default(@0.8)","Spearman(@0.8)",
                "@0.4", "Driven(@0.2)", "Outcome_Driven(@0.2)"),
       lty=c(1:6),
       col=c("red","black","black","black","black","black"))
```

```
# Sparsity
yval <- ySpar_P_00</pre>
plot(xidx,yval,
     xlab="Iteration Cycle",
     ylab="Matrix Sparcity",
     ylim=c(0,1.0),
     main="Evolution of the Matrix Sparcity")
  lfit <-try(loess(yval~xidx,span=0.5));</pre>
  if (!inherits(lfit,"try-error"))
  {
    plx <- try(predict(lfit,se=TRUE))</pre>
    if (!inherits(plx,"try-error"))
      lines(xidx,plx$fit,lty=1,col="red")
    }
 }
lines(xidx,ySpar_D[xidx],lty=2)
lines(xidx,ySpar_S[xidx],lty=3)
lines(xidx,ySpar_P_40[xidx],lty=4)
lines(xidx,ySpar_P_20_D[xidx],lty=5)
lines(xidx,ySpar_P_20_0D[xidx],lty=6)
legend("bottomleft",
       legend=c("@0.0","Default(@0.8)","Spearman(@0.8)",
                 "@0.4", "Driven(@0.2)", "Outcome_Driven(@0.2)"),
       lty=c(1:6),
       col=c("red","black","black","black","black","black"))
```

#### **Evolution of the Correlation**

#### **Evolution of the Matrix Sparcity**



For the next part of the tutorial I'll set the correlation goal to 0.2

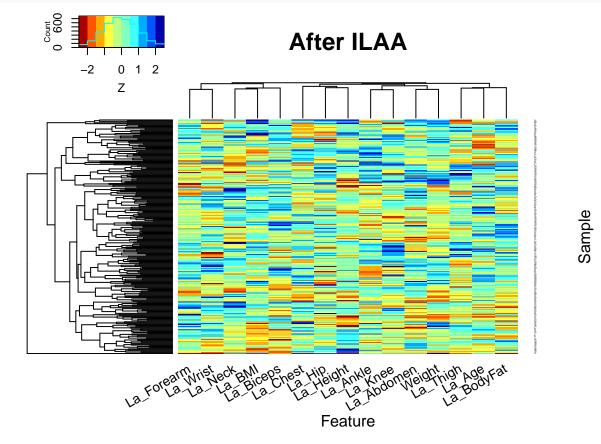
```
# Calling ILAA to achieve a final correlation of 0.2
body_fat_Decorrelated <- ILAA(body_fat,thr=0.2)
pander::pander(attr(body_fat_Decorrelated,"VarRatio"))</pre>
```

Table 4: Table continues below

Weight	${\rm La\_Ankle}$	La_Fo	rearm La	_Age La_	$_{ m Wrist}$	La_	Biceps	La_BodyFat
1	0.555	0.5	18 0	.483	0.403	0	.32	0.29
La_Neck	La_Knee	La_BMI	La_Thigh	La_Abdome	en La_	Chest	La_Hip	La_Height
0.279	0.228	0.211	0.183	0.132	0.	104	0.0993	0.0209

#### 3.3 The Heatmap of the Transformed Data

Here we review the transformed data using a heatmap of the data



```
par(op)
```

#### 3.4 Data Frame Attributes

The returned data matrix contains the following attributes

```
attr(body_fat_Decorrelated,"UPLTM") #The transformation matrix
attr(body_fat_Decorrelated,"fscore") #The score of each feature
attr(body_fat_Decorrelated,"drivingFeatures") #The list of driving features
attr(body_fat_Decorrelated,"R.critical") #The estimated minimum achieviable correlation
attr(body_fat_Decorrelated,"IDeAEvolution") #Evolution of the algorithm
attr(body_fat_Decorrelated,"VarRatio") #Variance Ratios: var(Latent)/Var(obs)
```

The main attributes is "UPLTM". That stores the specific linear transformation matrix from observed variables to the latent variable.

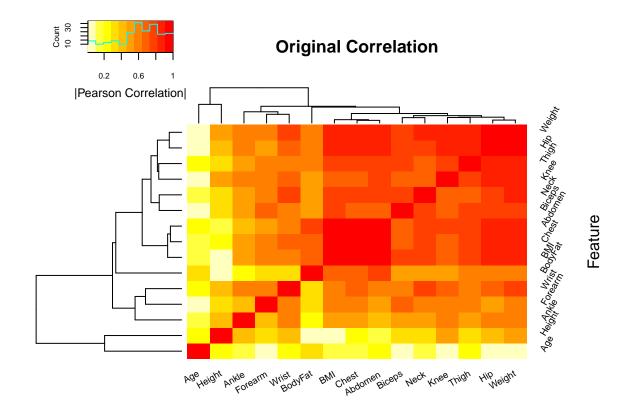
The next relevant attribute is the "VarRatio", this attributive stores the fraction of the original feature variance that is still present in the latent variable. All non-altered variables return a "VarRatio" of 1.

The "IDeAEvolution" attribute can be used to verify if the algorithm achieved the target correlation goal, and the sparsity of the returned matrix.

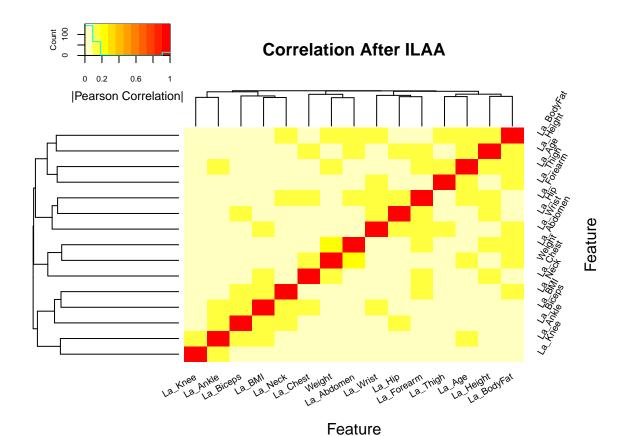
#### 3.5 The ILAA Transformed Data

Before exploring into more detail, the properties of the ILAA results. Let us first verify that the returned matrix does not contain features with very high correlation among them.

Here I'll plot the original correlation and the correlation of the returned data set.



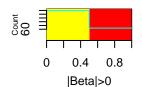
## **Feature**



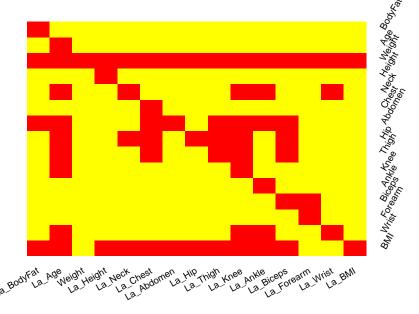
## 3.6 Exploring the Transformation

The attr(body\_fat\_Decorrelated, "UPLTM") returns the transformation matrix. The UPLTM is sparse, here I show a heat map of the transformation matrix that shows which elements are different from zero.

```
UPLTM <- attr(body_fat_Decorrelated,"UPLTM")</pre>
gplots::heatmap.2(1.0*(abs(UPLTM)>0),
                  trace = "none",
                  mar = c(5,5),
                  col=rev(heat.colors(2)),
                  Rowv=NULL,
                  Colv="Rowv",
                  dendrogram="none",
                  main = "Transformation matrix",
                  cexRow = 0.75,
                  cexCol = 0.75,
                 srtCol=30,
                 srtRow=60,
                  key.title=NA,
                  key.xlab="|Beta|>0",
                  xlab="Output Feature", ylab="Input Feature")
```



# **Transformation matrix**



**Output Feature** 

#### 3.7 The Latent Formulas

The sparsity of the UPLTM matrix can be analyzed to get the formula for each one of the latent formulas. The getLatentCoefficients() and its attribute: attr(LatentFormulas, "LatentCharFormulas") can be used to display the formula of the latent variables.

```
# Get a list with the latent formulas' coefficients
LatentFormulas <- getLatentCoefficients(body_fat_Decorrelated)

# A string character with the formulas can be obtained by:
charFormulas <- attr(LatentFormulas, "LatentCharFormulas")
pander::pander(as.matrix(charFormulas))</pre>
```

La_BodyFat	+ BodyFat $+$ (0.120)Weight $-$ (0.800)Abdomen $-$
	(0.480) BMI
${f La\_Age}$	+ Age + (0.363)Weight - $(0.636)$ Neck -
	(1.117)Abdomen - $(8.09e-04)$ Hip + $(2.273)$ Thigh
	-(1.732)Knee $-(5.032)$ Wrist $-(0.864)$ BMI
La_Height	- (0.191)Weight + Height + $(1.339)$ BMI
${ m La\_Neck}$	-(0.100)Weight + Neck + $(0.172)$ Hip -
	(0.074)BMI
$La\_Chest$	- $(0.140)$ Weight + Chest - $(0.363)$ Abdomen +
	(0.419)Hip + $(0.265)$ Thigh - $(1.082)$ BMI
$La\_Abdomen$	- $(0.094)$ Weight + Abdomen - $(1.865)$ BMI
${ m La\_Hip}$	-(0.181)Weight + Hip - $(0.430)$ BMI
${ m La\_Thigh}$	-(0.056)Weight $+(0.137)$ Abdomen $-(0.489)$ Hip
	+ Thigh - $(0.256)$ BMI

La_Knee	- $(0.056)$ Weight + $(0.067)$ Neck -
	(0.017)Abdomen - $(0.046)$ Hip - $(0.121)$ Thigh +
	Knee - $(0.406)$ Wrist + $(0.229)$ BMI
${f La\_Ankle}$	-(0.035)Weight $+(0.098)$ Neck $+$
	(0.069)Abdomen + Ankle - $(0.594)$ Wrist -
	(0.128)BMI
$La\_Biceps$	-(0.081)Weight $+(0.075)$ Abdomen $+$
	(0.098)Hip - $(0.200)$ Thigh + Biceps - $(0.140)$ BMI
$La\_Forearm$	- $(0.017)$ Weight - $(0.323)$ Biceps + Forearm
${f La\_Wrist}$	- $(0.012)$ Weight - $(0.165)$ Neck + Wrist
$La\_BMI$	- $(0.111)$ Weight + BMI

## 3.8 Latent Variable Interpretation

The ILAA returns the Unit Preserving Linear Transformation Matrix (UPLTM). This specific transformation is the combination of statistically significant linear association analysis between feature pairs. Each significant association is modeled by a linear equation; henceforth, the interpretation of each feature is as follows:

• Each discovered latent variable is the residual of the observed parent variable vs. the suitable model of the variables associated with the parent variable. For example:

```
LaWrist = Wrist - 0.012Weight - 0.165Neck.
```

Describes that the Wrist is associated with the Weight and Neck. The latent variable LaWrist is the amount of information in the Wrist not found by Weight nor the Neck.

• Therefore, the model of the Wrist is:

```
Wrist = +0.012Weight + 0.165Neck + b_o
```

where  $b_o$  is the bias term. It can be estimated using the difference between the mean of the raw observations and the mean of the model.

## 3.9 The Formula Network

The graph\_from\_adjacency\_matrix() function from igraph can be used to visualize the association between variables.

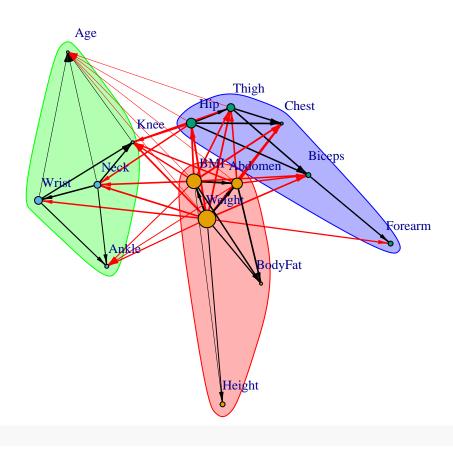
```
transform <- attr(body_fat_Decorrelated, "UPLTM") != 0
colnames(transform) <- str_remove_all(colnames(transform), "La_")
# The weights are proportional to the observed correlation
transform <- abs(transform*cor(body_fat[,rownames(transform)]))

# The size depends on the variable independence relevance (fscore)

VertexSize <- attr(body_fat_Decorrelated, "fscore")

names(VertexSize) <- str_remove_all(names(VertexSize), "La_")
# Normalization
VertexSize <- 10*(VertexSize-min(VertexSize))/(max(VertexSize)-min(VertexSize))</pre>
```

## **Feature Association**



## par(op)

#### 3.9.1 ILAA Solution and Perturbations

ILLA solutions depends on the observed data. The provided function can add data perturbations aiming to improve the sensitivity to find multicollinearity issues.

**3.9.1.1 Bootstrapping ILLA** To handle the data sensitivity to the input data, ILAA allows for bootstrapping estimation of the transformation matrix.

# ## Here we petrubate only 5% of the data body\_fat\_Decorrelated <- ILAA(body\_fat,thr=0.2,bootstrap=100) pander::pander(attr(body\_fat\_Decorrelated,"VarRatio"))</pre>

Table 7: Table continues below

Weight	La_Ankle	La_Forearm	La_Age	La_Wrist	La_Biceps	La_BodyFat
1	0.551	0.517	0.484	0.402	0.32	0.291

La_Neck	La_Knee	La_BMI	La_Thigh	La_Abdomen	La_Chest	La_Hip	La_Height
0.276	0.241	0.211	0.182	0.13	0.103	0.0992	0.0208

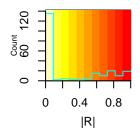
#### ## Getting the formulas

LatentFormulas <- getLatentCoefficients(body\_fat\_Decorrelated)
charFormulas <- attr(LatentFormulas, "LatentCharFormulas")

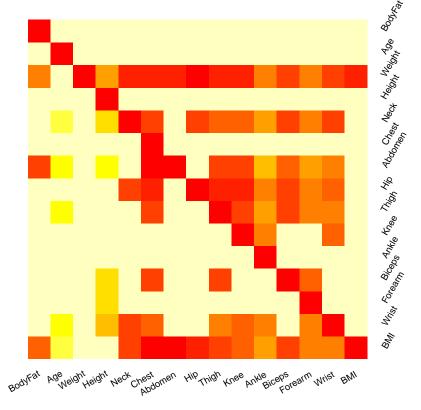
pander::pander(as.matrix(charFormulas))

${\bf La\_BodyFat}$	+ Body Fat $+$ (0.119) Weight - (0.795) Abdomen -
	(0.492)BMI
${f La\_Age}$	+ Age + (0.365)Weight - (0.678)Neck -
	(1.136)Abdomen - $(0.115)$ Hip + $(2.159)$ Thigh -
	(1.331)Knee - $(5.499)$ Wrist - $(0.571)$ BMI
${f La\_Height}$	- (0.192)Weight + Height - $(1.24e-05)$ Neck -
	(1.22e-04)Abdomen + $(6.60e-04)$ Biceps -
	(1.91e-03)Forearm + $(7.88e-05)$ Wrist +
	(1.342)BMI
${f La\_Neck}$	-(0.098)Weight + Neck + $(0.170)$ Hip -
	(3.87e-03)Wrist - $(0.093)$ BMI
${f La\_Chest}$	-(0.142)Weight $-(1.30e-04)$ Neck + Chest -
	(0.371)Abdomen + $(0.457)$ Hip + $(0.195)$ Thigh -
	(9.67e-04)Biceps + $(7.56e-04)$ Wrist - $(1.023)$ BMI
${f La\_Abdomen}$	- $(0.102)$ Weight + Abdomen - $(1.850)$ BMI
${f La\_Hip}$	-(0.182)Weight + $(1.37e-03)$ Neck + Hip -
	(0.419)BMI
${f La\_Thigh}$	- $(0.054)$ Weight - $(3.69e-04)$ Neck +
	(0.136)Abdomen - $(0.499)$ Hip + Thigh -
	(3.37e-03)Biceps + $(2.15e-03)$ Wrist - $(0.253)$ BMI
$La\_Knee$	- $(0.060)$ Weight + $(0.035)$ Neck -
	(8.58e-03)Abdomen - $(0.027)$ Hip - $(0.062)$ Thigh
	+  Knee - (0.210) Wrist + (0.098) BMI
${f La\_Ankle}$	- (0.032)Weight + $(0.094)$ Neck +
	(0.051)Abdomen + $(1.08e-04)$ Hip +
	(2.26e-04)Thigh - $(0.032)$ Knee + Ankle -
	(0.575)Wrist - $(0.094)$ BMI
${f La\_Biceps}$	- $(0.081)$ Weight - $(3.95e-03)$ Neck +
	(0.067)Abdomen + $(0.096)$ Hip - $(0.194)$ Thigh +
	Biceps - $(0.124)$ BMI

```
## The transformation
par(op)
# The non-zero coefficients
transform <- attr(body_fat_Decorrelated, "UPLTM") != 0</pre>
# For network analysis
colnames(transform) <- str_remove_all(colnames(transform), "La_")</pre>
# The weights are proportional to the observed correlation
transform <- abs(transform*cor(body_fat[,rownames(transform)]))</pre>
gplots::heatmap.2(transform,
                     trace = "none",
                     mar = c(5,5),
                     Rowv=NULL,
                     Colv="Rowv",
                     dendrogram="none",
                     col=rev(heat.colors(11)),
                     main = "(Transform <> 0)*Correlation",
                     cexRow = 0.75,
                     cexCol = 0.75,
                    srtCol=30,
                    srtRow=60,
                     key.title=NA,
                     key.xlab="|R|",
                     xlab="Output Feature", ylab="Input Feature")
  par(op)
```



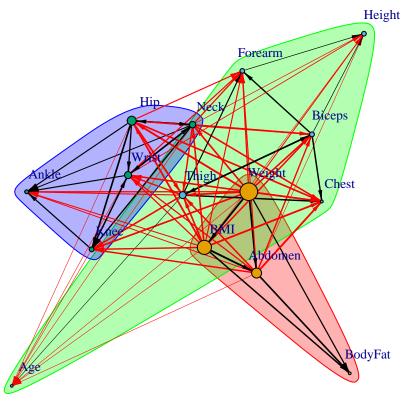
# (Transform <> 0)\*Correlation

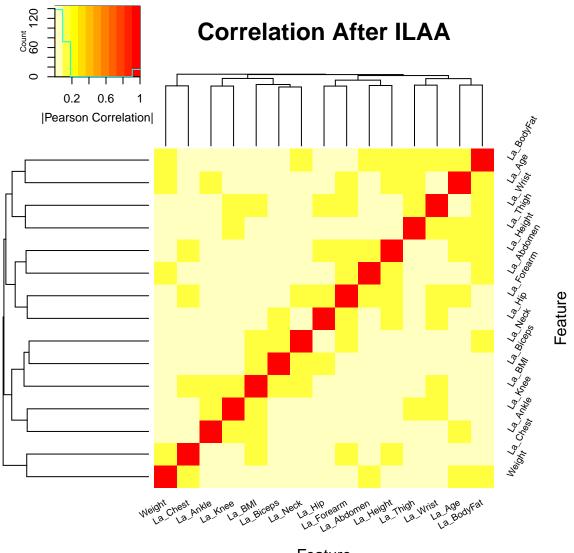


## **Output Feature**

```
edge.width=2*E(gr)$weight,
edge.arrow.size=0.5,
edge.arrow.width=0.5,
vertex.size=VertexSize,
vertex.label.cex=0.85,
vertex.label.dist=2,
main="Bootstrap: Feature Association")
```

# **Bootstrap: Feature Association**





## Feature

```
par(op)
diag(cormat) <- 0
pander::pander(max(abs(cormat)))</pre>
```

0.166

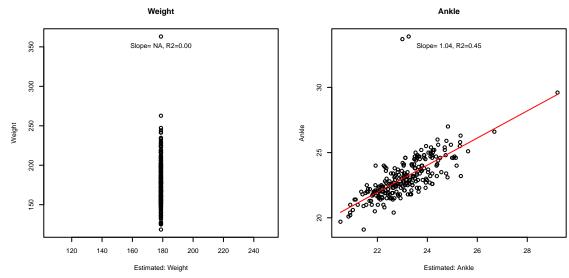
#### 3.9.2 Association Plots

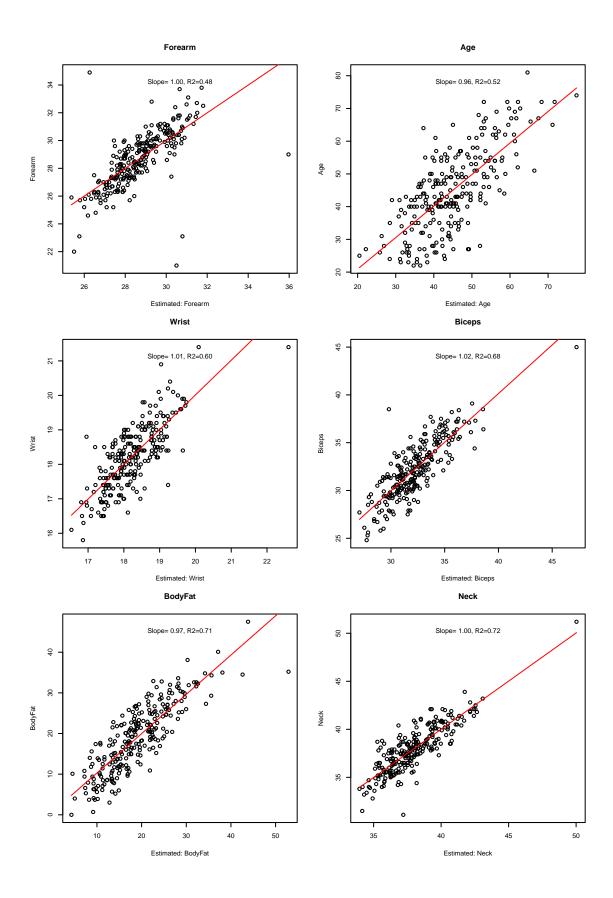
**3.9.2.1 Direct Transform Estimation** The transformation matrix can be used to get estimation of each variable from the latent models.

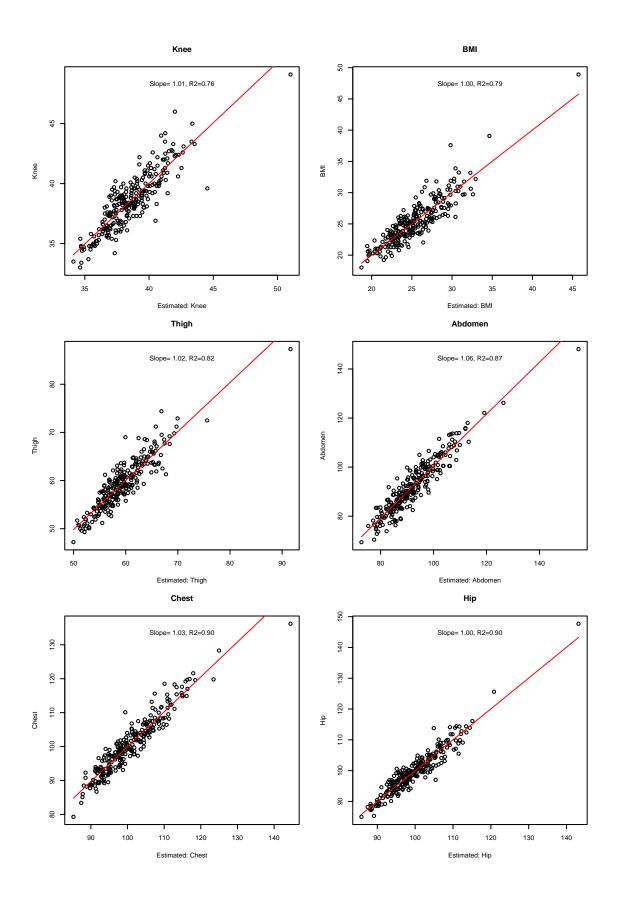
To to this just set the diagonal of the transformation to zero, then rotate the input matrix, multiply by -1, and the the output is the estimated observation from the independent variables. The bias term is estimated by computing the observed mean minus the transformed mean.

```
transform <- attr(body_fat_Decorrelated,"UPLTM")
varratio <- attr(body_fat_Decorrelated,"VarRatio")
# Set the diagonal to zero
diag(transform) <- 0</pre>
```

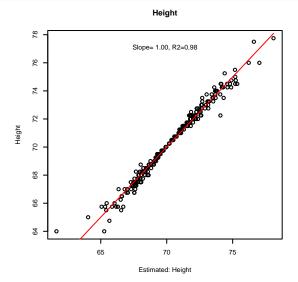
```
#Estimating the observation
obsestim <- -1*as.data.frame(</pre>
  as.matrix(body_fat[,rownames(transform)]) %*% transform)
#Bias estimation
bias <- apply(body_fat[,rownames(transform)],2,mean) -</pre>
  apply(obsestim[,colnames(transform)],2,mean)
#Plotting
par(mfrow=c(1,2),cex=0.45)
for (vn in names(varratio))
  oname <- str_remove_all(vn,"La_")</pre>
  plot(obsestim[,vn] + bias[oname],
       body_fat[,oname],xlab=paste("Estimated:",oname),
       ylab=oname,main=oname)
  indx <- obsestim[,vn]+bias[oname]</pre>
  lmtvals <- lm(body_fat[,oname] ~ indx )</pre>
  xvals <- c(min(obsestim[,vn]+ bias[oname]), max(obsestim[,vn]+ bias[oname]))</pre>
  pred <- lmtvals$coefficients[1] + lmtvals$coefficients[2] * xvals</pre>
  lines(x=xvals,y=pred,col="red")
 ylim <- c(min(body_fat[,oname]),max(body_fat[,oname]))</pre>
 text(xvals[1]+(xvals[2]-xvals[1])/2,0.95*(ylim[2]-ylim[1])+ylim[1],
       sprintf("Slope= %.2f, R2=%3.2f",
                lmtvals$coefficients[2],1.0-varratio[vn])
```











The visual inspection of the above-displayed figures shows that some latent variables are not associated with the original parent variable, but their model is fully correlated to the observed parent variable. A clear example is the last plot.

## 4 ILAA for Supervised Learning

The rerecorded use of ILAA transformation in supervised learning is to split the data into training and validation sets. Henceforth, the next lines of code will split the data into training (75%) and testing (25%)

#### 4.1 Split into Training Testing Sets

```
# 75% for training 25% for testing
set.seed(2)
trainsamples <- sample(nrow(body_fat),3*nrow(body_fat)/4)

trainingset <- body_fat[trainsamples,]
testingset <- body_fat[-trainsamples,]</pre>
```

#### 4.2 Data Train Analysis and Prediction of the Test Set

By default, ILAA() transforms are blind to outcome associations. but in supervised learning the user is free to specify a target outcome to drive the shape of the transformation matrix. Outcome-driven transformations try to keep unaltered features strongly associated with the target.

The predictDecorrelate() function can be used to predict any new dataset from an ILAA transformed object.

The next code snippet shows the process of transforming the training set and then using the returned object to transform the testing set using both outcome-blind and outcome-driven transformations.

## 4.3 Train a Regression Model for Body Fat Prediction

Once we have a transformed training and testing set, we can proceed to train a linear model of the body fat content. For this example we will use the LASSO\_MIN() function of the FRESA.CAD package to model the BodyFat using all the variables in the transformed training set.

```
## Outcome-Blind
modelBodyFatRaw <- LASSO_MIN(BodyFat~.,trainingset)
pander::pander(as.matrix(modelBodyFatRaw$coef),caption="Raw Coefficients")</pre>
```

Table 10: Raw Coefficients

(Intercept)	3.04784
$\mathbf{Age}$	0.04975
${f Height}$	-0.39330
$\mathbf{Neck}$	-0.14960
Chest	-0.15167
Abdomen	0.80580
$\mathbf{Thigh}$	0.10851
$\mathbf{Ankle}$	0.10517
Biceps	0.13697
Forearm	0.00417
Wrist	-1.39800

```
## Outcome-Blind
modelBodyFat <- LASSO_MIN(BodyFat~.,body_fat_Decorrelated_train)
pander::pander(as.matrix(modelBodyFat$coef),caption="Outcome-Blind Coefficients")</pre>
```

Table 11: Outcome-Blind Coefficients

(Intercept)	-56.0734
${f La\_Age}$	0.0372
${f Weight}$	0.1769
${ m La\_Height}$	0.5626
$La\_Neck$	-0.2083
${f La\_Chest}$	0.1769
$La\_Abdomen$	0.9393
La_Hip	0.2382
${ m La\_Thigh}$	0.1103
La_Knee	0.0819
$\mathbf{La}_{-}\mathbf{A}\mathbf{n}\mathbf{k}\mathbf{l}\mathbf{e}$	0.0997
$La\_Biceps$	0.1537
${f La\_Wrist}$	-0.9001
$La\_BMI$	1.9868

```
## Outcome-Driven
modelBodyFatD <- LASSO_MIN(BodyFat~.,body_fat_Decorrelated_trainD)
pander::pander(as.matrix(modelBodyFatD$coef),caption="Outcome-Driven Coefficients")</pre>
```

Table 12: Outcome-Driven Coefficients

(Intercept)	-29.7164
${f La\_Age}$	0.0172
${f La\_Weight}$	-0.0989
$La\_Neck$	-0.5088
${f La\_Chest}$	-0.1315
${f Abdomen}$	0.6441
La_Hip	-0.1942
La_Thigh	0.1177
$La\_Ankle$	0.0920
${f La\_Biceps}$	0.1394
${ m La\_Wrist}$	-0.5890

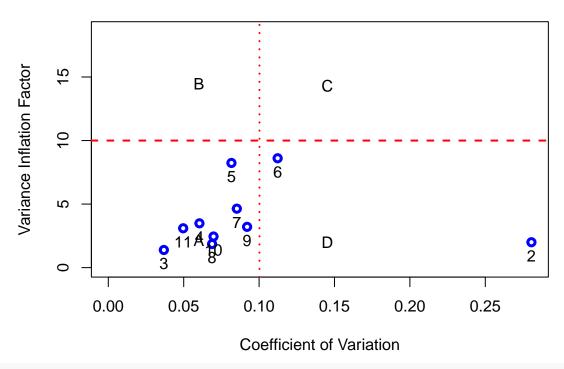
The printed beta coefficients of the models show that the LASSO models are different between the Outcomedriven and outcome-blind ILAA methods.

 $\textbf{4.3.0.1} \quad \textbf{Multicollinear Analysis} \quad \text{Here we check the Variance inflation factor (VIF) on the train and testing sets}$ 

```
frm <- paste("BodyFat~",str_flatten(modelBodyFatRaw$selectedfeatures," + "))

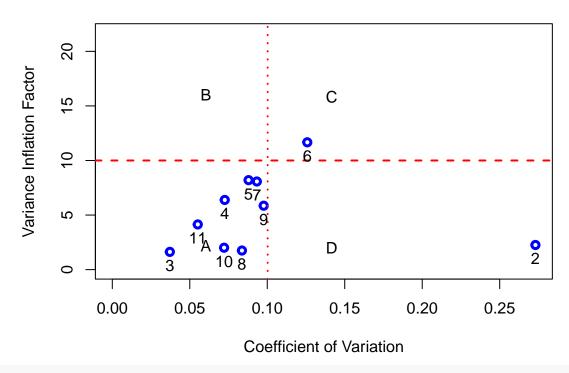
X <- model.matrix(formula(frm),trainingset);
mc <- multiCol(X)
vifd <- VIF(X)
vifx <-vif(lm(formula(frm),trainingset))
title("Raw Train VIF")</pre>
```

# **Raw Train VIF**



```
X <- model.matrix(formula(frm),testingset);
mc <- multiCol(X)
vifd <- VIF(X)
vifx <-vif(lm(formula(frm),testingset))
title("Raw Test VIF")</pre>
```

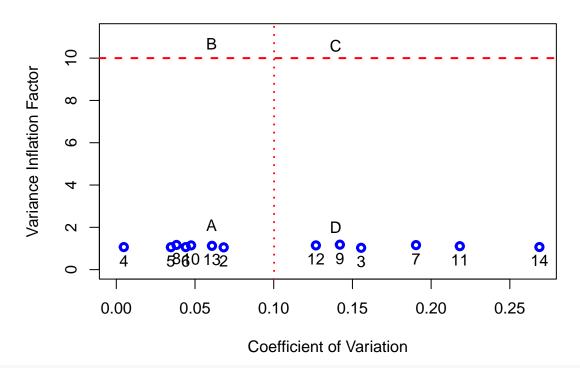
# **Raw Test VIF**



```
frm <- paste("BodyFat~",str_flatten(modelBodyFat$selectedfeatures," + "))

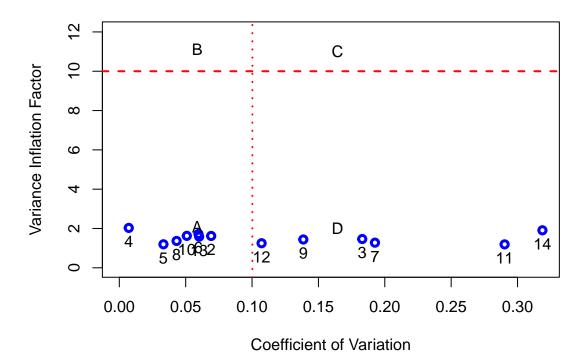
X <- model.matrix(formula(frm),body_fat_Decorrelated_train);
mc <- multiCol(X)
vifd <- VIF(X)
vifx <-vif(lm(formula(frm),body_fat_Decorrelated_train))
title("Blind: Train VIF")</pre>
```

**Blind: Train VIF** 



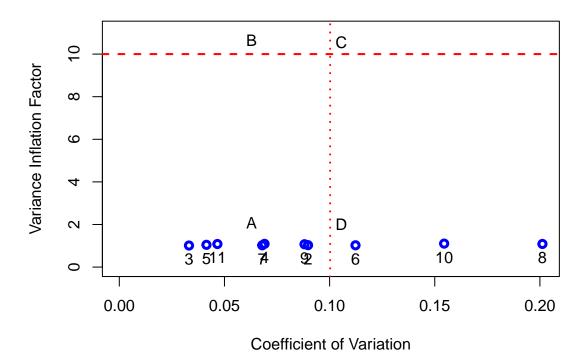
X <- model.matrix(formula(frm),body\_fat\_Decorrelated\_test);
mc <- multiCol(X)
title("Blind: Test VIF")</pre>

**Blind: Test VIF** 



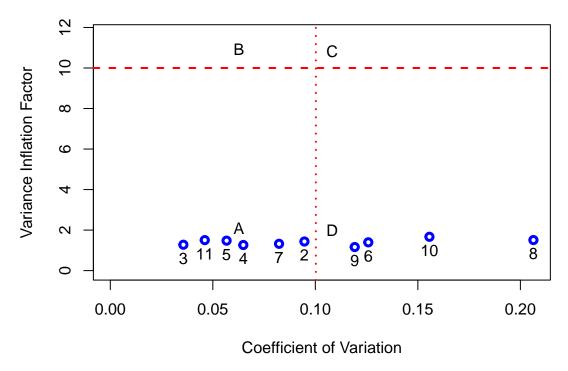
```
frm <- paste("BodyFat~",str_flatten(modelBodyFatD$selectedfeatures," + "))
X <- model.matrix(formula(frm),body_fat_Decorrelated_trainD);
mc <- multiCol(X)
title("Driven: Train VIF")</pre>
```

# **Driven: Train VIF**



```
X <- model.matrix(formula(frm),body_fat_Decorrelated_testD);
mc <- multiCol(X)
title("Driven: Test VIF")</pre>
```

# **Driven: Test VIF**



The plots clearly indicate that both models do not have colinearity issues

#### 4.3.1 The Model Coefficients in the Observed Space

The FRESA.CAD package provides a handy function, getObservedCoef()m to get the linear beta coefficients from the transformed object. The next code shows the procedure.

```
# Get the coefficients in the observed space for the outcome-blind
observedCoef <- getObservedCoef(body_fat_Decorrelated_train,modelBodyFat)
pander::pander(as.matrix(observedCoef$coefficients),caption="Blind Coefficients")</pre>
```

Table 13: Blind Coefficients

(Intercept)	-56.07340
$\mathbf{Age}$	0.03722
$\mathbf{Weight}$	-0.18055
${f Height}$	0.56259
$\mathbf{Neck}$	-0.07292
${f Chest}$	-0.28360
Abdomen	0.89781
${f Hip}$	-0.14482
${f Thigh}$	0.15616
$\mathbf{K}\mathbf{nee}$	-0.00665
${f Ankle}$	0.09967
Biceps	0.15371
$\mathbf{Wrist}$	-1.19584
BMI	1.36021

```
# The outcome-driven coefficients
observedCoefD <- getObservedCoef(body_fat_Decorrelated_trainD,modelBodyFatD)
pander::pander(as.matrix(observedCoefD$coefficients),caption="Driven Coefficients")</pre>
```

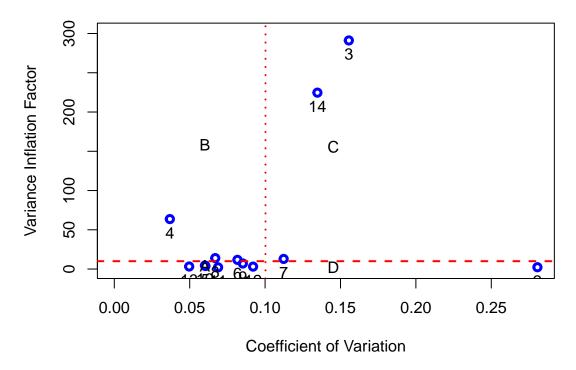
Table 14: Driven Coefficients

(Intercept)	-29.7164
$\mathbf{Age}$	0.0172
${f Weight}$	-0.0790
Neck	-0.1669
Chest	-0.1315
Abdomen	0.8662
${ m Hip}$	-0.0162
$\mathbf{Thigh}$	0.1124
Knee	-0.0417
$\mathbf{Ankle}$	0.0920
Biceps	0.1394
$\mathbf{Wrist}$	-0.7604
BMI	0.2150

**4.3.1.1** Muticollinear Analysis on the observed space Here we check the Variance inflation factor (VIF) on the train and testing sets using the observed variables

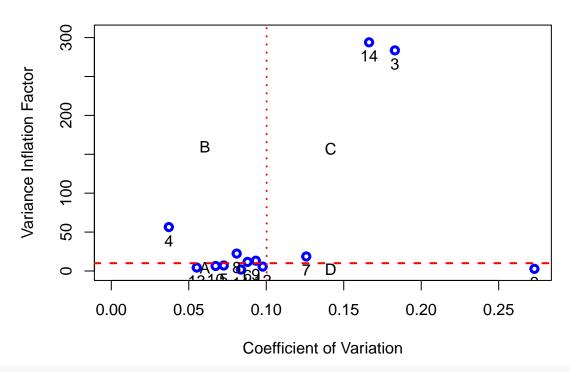
```
X <- model.matrix(formula(observedCoef$formula),trainingset);
mc <- multiCol(X)
title("Observed Training VIF")</pre>
```

# **Observed Training VIF**



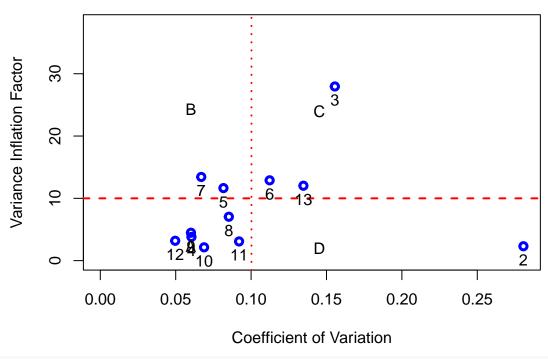
```
X <- model.matrix(formula(observedCoef$formula),testingset);
mc <- multiCol(X)
title("Observed Testing VIF")</pre>
```

# **Observed Testing VIF**



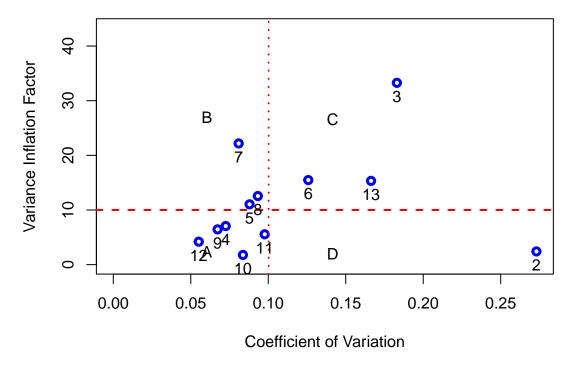
```
X <- model.matrix(formula(observedCoefD$formula),trainingset);
mc <- multiCol(X)
title("Driven: Observed Training VIF")</pre>
```

# **Driven: Observed Training VIF**



```
X <- model.matrix(formula(observedCoefD$formula),testingset);
mc <- multiCol(X)
title("Driven: Observed Testing VIF")</pre>
```

# **Driven: Observed Testing VIF**



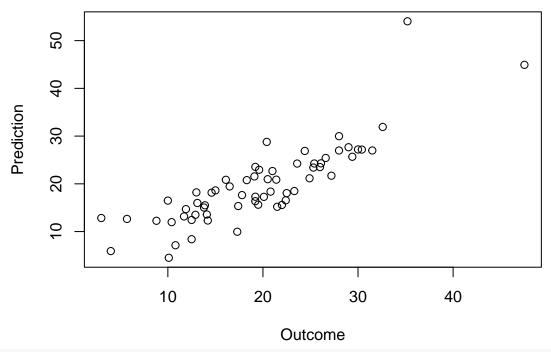
The results indicate that the models created using the observed variables have strong collinearity issues.

### 4.3.2 Predict Using the Transformed Data-Set

The user can predict the BodyFat content using the handy predict() function. After that we can measure the testing performance using the predictionStats\_regression() function.

Body Fat: Blind

**Body Fat: Blind** 



pander::pander(rmetrics)

• corci:

cor		
0.848	0.76	0.905

• biasci: 0.0434, -1.0731 and 1.1599

• **RMSEci**: 4.43, 3.78 and 5.37

• spearmanci:

50%	2.5%	97.5%
0.86	0.759	0.922

• MAEci:

50%	2.5%	97.5%
3.37	2.77	4.14

### • pearson:

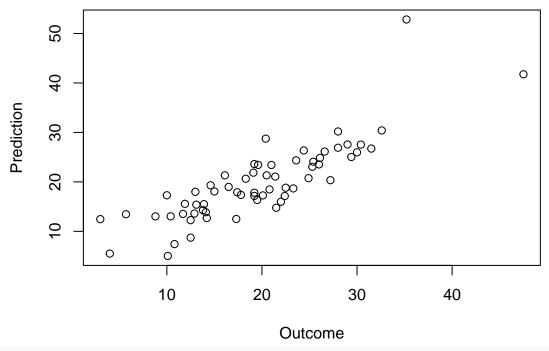
Table 18: Pearson's product-moment correlation: predictions[, 1] and predictions[, 2]

Test statistic	df	P value	Alternative hypothesis	cor
12.5	61	1.84e-18 * * *	two.sided	0.848

# 

Body Fat: Driven

# **Body Fat: Driven**



pander::pander(rmetrics)

• corci:

cor		
0.842	0.751	0.902

• biasci: 0.139, -0.972 and 1.249

- RMSEci: 4.41, 3.76 and 5.34
- spearmanci:

50%	2.5%	97.5%
0.849	0.745	0.912

• MAEci:

50%	2.5%	97.5%
3.39	2.76	4.16

#### • pearson:

Table 22: Pearson's product-moment correlation: predictions[, 1] and predictions[, 2]

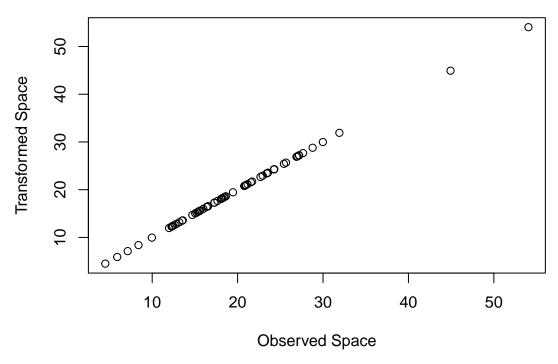
Test statistic	df	P value	Alternative hypothesis	cor
12.2	61	5.56e-18 * * *	two.sided	0.842

The reported metrics indicated that the model predictions are highly correlated to the real BodyFat

### 4.3.3 Prediction Using the Observed Features

An ILAA user has the option to predict the BodyFat content from the observed testing set using the computed beta coefficients. The next lines of code show how to do the prediction using model.matrix() R function and the dot product %\*%:

# **Test Predictions: Observed vs. Transformed**



The last plot shows the expected result: that both predictions are identical.

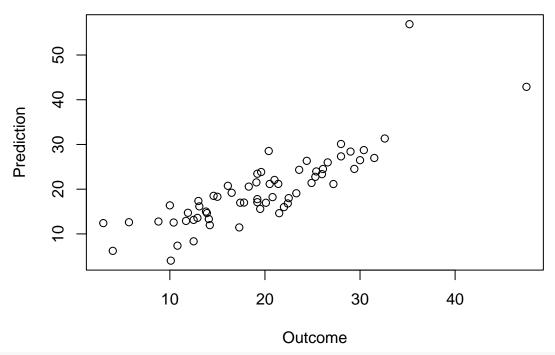
### 4.3.4 Comparison to Raw Model

A last experiment is to compare the differences between a LASSO model created from the observed features to the model created from the transformed observations.

The next lines of code compute the linear model using LASSO from the original observed data. Then, it computes the predicted performance.

Raw: Body Fat

Raw: Body Fat



pander::pander(rmetrics)

• corci:

cor		
0.837	0.743	0.898

• biasci: 0.0882, -1.0731 and 1.2494

• **RMSEci**: 4.61, 3.93 and 5.59

• spearmanci:

50%	2.5%	97.5%
0.858	0.755	0.919

• MAEci:

50%	2.5%	97.5%
3.38	2.7	4.23

• pearson:

Table 26: Pearson's product-moment correlation: predictions[, 1] and predictions[, 2]

Test statistic	df	P value	Alternative hypothesis	cor
11.9	61	1.27e-17***	two.sided	0.837

The evaluation of the testing results indicates that the observed model predictions have a correlation of 0.875. Slightly superior, but not statistically significant, to the one observed from the model estimated from the transformed space: (  $\rho_t = 0.863$  vs.  $\rho_o = 0.875$  )

#### 4.3.5 Comparing the Feature Significance on the Models

The main advantage of the ILAA transformation is that the returned latent variables are not colinear hence the statistical significance of the beta coefficients are not affected by multicolinearity. The next code snippet shows how to get the beta coefficients using the lm(), and summary.lm() functions.

The inspection of the summary results clearly shows that most of the beta coefficients on the transformed data set are significant.

Table 27: Fitting linear model: BodyFat  $\sim$  .

	Estimate	Std. Error	t value	$\Pr(> t )$	
(Intercept)	7.7499	10.4950	0.738	4.61e-01	
$\mathbf{Age}$	0.0619	0.0357	1.733	8.48e-02	
Height	-0.4207	0.1476	-2.851	4.87e-03	* *
Neck	-0.2136	0.2639	-0.810	4.19e-01	
$\mathbf{Chest}$	-0.2765	0.1133	-2.442	1.56e-02	*
Abdomen	0.8843	0.0919	9.627	6.62e-18	* * *
$\mathbf{Thigh}$	0.1278	0.1383	0.924	3.57e-01	
Ankle	0.2219	0.2788	0.796	4.27e-01	
Biceps	0.2260	0.1956	1.155	2.49e-01	
Forearm	0.1075	0.2538	0.423	6.72 e-01	
$\mathbf{Wrist}$	-1.6853	0.6294	-2.678	8.11e-03	* *

```
plot(rawlm)
```

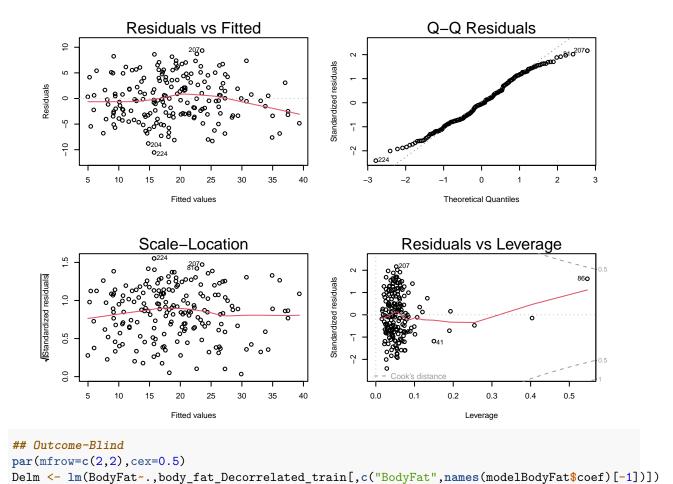


Table 28: Fitting linear model: BodyFat  $\sim$  .

pander::pander(Delm,add.significance.stars=TRUE)

	Estimate	Std. Error	t value	$\Pr(> t )$	
(Intercept)	-90.7305	70.5488	-1.286	2.00e-01	
$\mathbf{La}\mathbf{Age}$	0.0497	0.0384	1.294	1.97e-01	
${f Weight}$	0.1812	0.0119	15.167	1.17e-33	* * *
${ m La\_Height}$	0.9638	0.9999	0.964	3.36e-01	
$La\_Neck$	-0.3390	0.2571	-1.318	1.89e-01	
$La\_Chest$	0.2466	0.1225	2.014	4.56e-02	*
$La\_Abdomen$	0.9979	0.0995	10.032	5.63e-19	* * *
${ m La\_Hip}$	0.3015	0.1619	1.862	6.42e-02	
La_Thigh	0.1665	0.1582	1.052	2.94e-01	
La_Knee	0.2044	0.2819	0.725	4.69e-01	
La_Ankle	0.2925	0.2991	0.978	3.29e-01	
$La\_Biceps$	0.2946	0.1921	1.534	1.27e-01	
${ m La\_Wrist}$	-1.1125	0.5793	-1.921	5.64 e-02	
La_BMI	2.0726	0.2075	9.990	7.34e-19	* * *

plot(Delm)

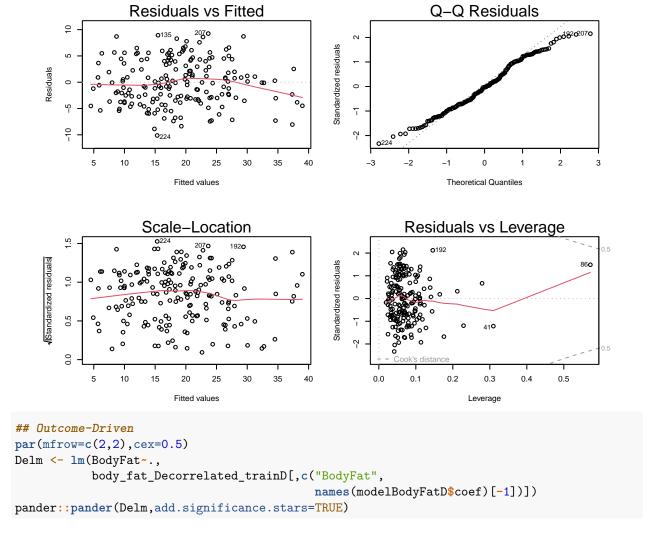
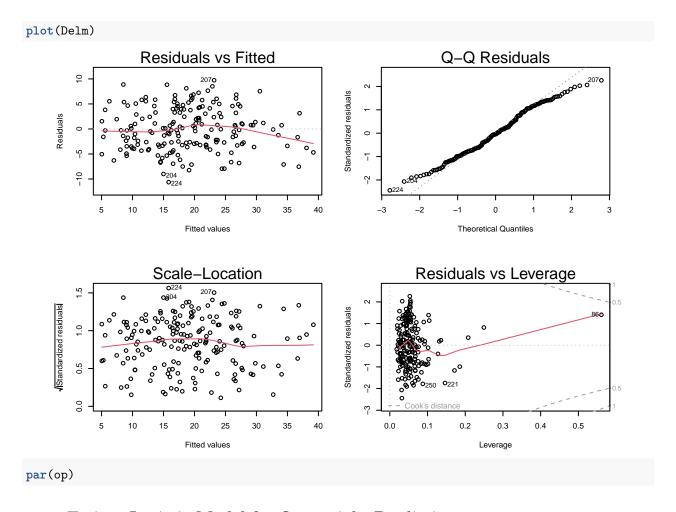


Table 29: Fitting linear model: BodyFat  $\sim$  .

	Estimate	Std. Error	t value	$\Pr(> t )$	
(Intercept)	-24.3110	18.3582	-1.32	1.87e-01	
$La\_Age$	0.0493	0.0381	1.29	1.97e-01	
$La\_Weight$	-0.1317	0.0432	-3.05	2.65e-03	* *
$La\_Neck$	-0.6998	0.2252	-3.11	2.20e-03	* *
$La\_Chest$	-0.2334	0.1341	-1.74	8.36e-02	
Abdomen	0.6707	0.0316	21.25	1.39e-50	* * *
La_Hip	-0.2604	0.1007	-2.59	1.05e-02	*
La_Thigh	0.1999	0.1503	1.33	1.85e-01	
$La\_Ankle$	0.3420	0.2892	1.18	2.39e-01	
$La\_Biceps$	0.2839	0.1909	1.49	1.39e-01	
$La\_Wrist$	-0.9602	0.5653	-1.70	9.12e-02	



## 4.4 Train a Logistic Model for Overweight Prediction

This last experiment showcases the effect of data transformation on logistic modeling. This experiment starts by creating a data-frame that does not includes the BMI, Height, and Weight variables. The target outcome is to identify if the person is Overweight or normal. (BMI>=25). The next lines of code compute the new data frames and remove the above mentioned variables.

### 4.4.1 Data Conditioning

First Remove Height and Weight from Training and Testing Sets

```
trainingsetBMI <- trainingset[,!(colnames(trainingset) %in% c("Weight","Height"))]
testingsetBMI <- testingset[,!(colnames(trainingset) %in% c("Weight","Height"))]
trainingsetBMI$Overweight <- 1*(trainingsetBMI$BMI>=25)
testingsetBMI$Overweight <- 1*(testingsetBMI$BMI>=25)
trainingsetBMI$BMI <- NULL
testingsetBMI$BMI <- NULL
# The number of subjects
pander::pander(table(trainingsetBMI$Overweight),caption="Training Distribution")</pre>
```

Table 30: Training Distribution

0	1
96	92

pander::pander(table(testingsetBMI\$Overweight), caption="Testing Distribution")

Table 31: Testing Distribution

0	1
29	34

The last code snippet transforms the observed features using ILLA and setting a target variable and setting the convergence not to be affected by the target outcome.

### 4.4.2 The Logistic Model

LASSO MIN with a binomial family is used to compute the logistic model of overweight.

Table 32: Training: Blind

(Intercept)	-60.8766
${f La\_BodyFat}$	0.0276
${f Chest}$	0.6239
${f La\_Abdomen}$	0.1753
${f La\_Hip}$	0.1748
${f La\_Thigh}$	0.0502
$La\_Knee$	-0.0198
$La\_Ankle$	0.0707

La_Biceps	0.0363
$La\_Wrist$	0.2129

Table 33: Training: Driven

(Intercept)	-49.36391
${f La\_BodyFat}$	0.00335
$\mathbf{Chest}$	0.51300
$La\_Abdomen$	0.12957
La_Hip	0.10793
${f La\_Ankle}$	0.01892
$La\_Biceps$	0.01039
${f La\_Wrist}$	0.01784

#### 4.4.3 The Model Coefficients in the Observed Space

Once the logistic model is created in the transformed space, we can compute the beta coefficients for each one of the observed variables.

Table 34: Observed Coefficients

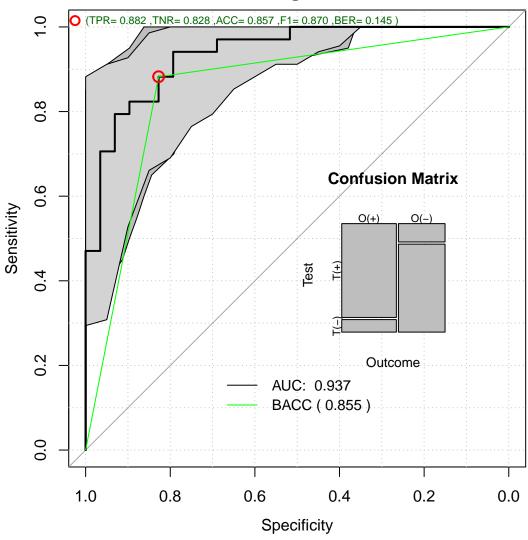
(Intercept)	-60.8766
${f BodyFat}$	0.0276
${f Neck}$	-0.0539
${f Chest}$	0.3680
${f Abdomen}$	0.1495
${f Hip}$	0.0518
${f Thigh}$	0.0405
$\mathbf{K}\mathbf{nee}$	-0.0376
${f Ankle}$	0.0707
Biceps	0.0363
$\mathbf{Wrist}$	0.1712

### 4.4.4 Predict Using the Transformed Data Set

The predictions of the testing set can be done using the handy predict() function. The evaluation of the testing results can be evaluated using the predictionStats\_binary() function.

```
## Outcome-blind
predicOverweight <- predict(modelOverweight,OW_Decorrelated_test)</pre>
```

# **Overweight: Blind**



## pander::pander(pr\$ClassMetrics)

• accci:

50%	2.5%	97.5%
0.857	0.762	0.937

• senci:

50%	2.5%	97.5%
0.886	0.763	0.973

• speci:

	50%	2.5%	97.5%
	0.828	0.667	0.96
aucci:			
	50%	2.5%	97.5%
	0.856	0.762	0.936
berci:			
	50%	2.5%	97.5%
	0.144	0.0639	0.238
preci:			
	50%	2.5%	97.5%

0.861

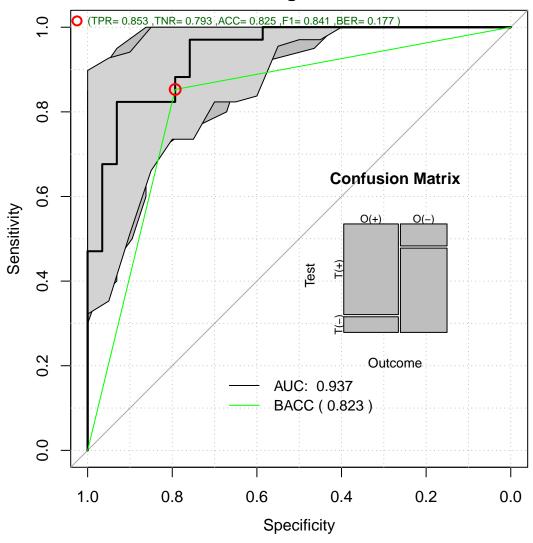
• F1ci:

50%	2.5%	97.5%
0.871	0.774	0.944

0.73

0.969

# **Overweight: Driven**



pander::pander(pr\$ClassMetrics)

• accci:

50%	2.5%	97.5%
0.825	0.73	0.921

• senci:

50%	2.5%	97.5%
0.853	0.719	0.969

• speci:

50%	2.5%	97.5%
0.8	0.64	0.929

• aucci:

50%	2.5%	97.5%
0.824	0.722	0.917

• berci:

50%	2.5%	97.5%
0.176	0.0833	0.278

• preci:

50%	2.5%	97.5%
0.829	0.694	0.941

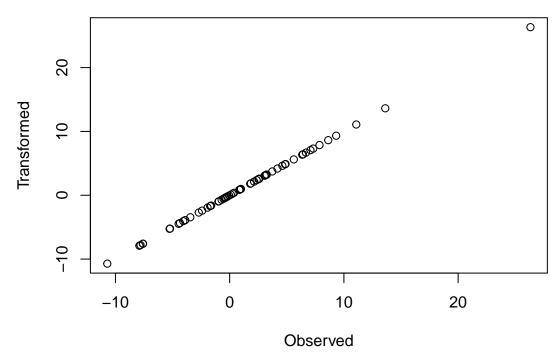
• F1ci:

50%	2.5%	97.5%
0.842	0.73	0.923

### 4.4.5 Prediction Using the Observed Features

The predict of the testing set can be done using the model.matrix() and the dot product %\*%.

# **Test predictions: Observed vs. Transformed**



The last plot shows the expected result: both predictions are identical.

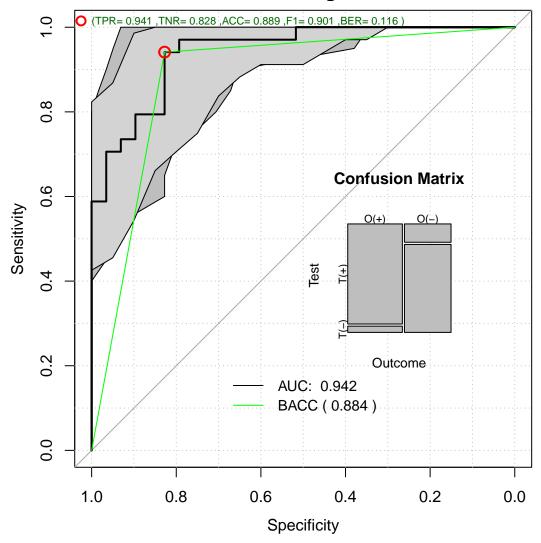
### 4.4.6 Comparison to Raw Model

To showcase the advantage of transformed modeling vs. raw modeling, here I'll estimate the logistic model from the observed variables and contrast to the model generated from the transformed space.

The next lines of code compute the logistic model and display its testing performance:

(Intercept)	BodyFat	Chest	Abdomen	Hip	Thigh	Ankle	Biceps	Wrist
-66.5	0.044	0.331	0.188	0.0226	0.0946	0.132	0.0818	0.106

# Overweight



pander::pander(pr\$ClassMetrics)

• accci:

50%	2.5%	97.5%
0.889	0.81	0.968

• senci:

50%	2.5%	97.5%
0.943	0.853	1

• speci:

	50%	2.5%	97.5%
	0.833	0.686	0.963
aucci:			
	50%	2.5%	97.5%
	0.889	0.803	0.963
berci:			
	50%	2.5%	97.5%
	0.111	0.0366	0.197

50%

0.871

• F1ci:

50%	2.5%	97.5%
0.904	0.821	0.97

2.5%

0.75

97.5%

0.971

The model created from the observed data has an ROC AUC that is not statistically significant to the transformed model

### 4.4.7 Comparing the Feature Significance on the Models

This last lines of code will compute the significance of the beta coefficients for both the observed model and the latent-based model. The user can clearly see that all the betas of the latent-based model are statically significant. An effect that is not seen in the logistic observed model.

```
par(mfrow=c(2,2),cex=0.5)

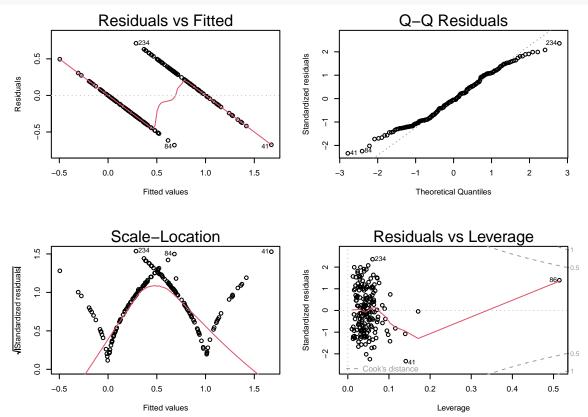
## Raw model
rawlm <- lm(Overweight~.,trainingsetBMI[,c("Overweight",names(rawmodelOverweight$coef)[-1])])
pander::pander(rawlm,add.significance.stars=TRUE)</pre>
```

Table 57: Fitting linear model: Overweight  $\sim$  .

	Estimate	Std. Error	t value	$\Pr(> t )$	
(Intercept)	-3.9632	0.55256	-7.172	1.88e-11	* * *
BodyFat	0.0060	0.00511	1.173	2.42e-01	
$\mathbf{Chest}$	0.0156	0.00783	1.993	4.78e-02	*
Abdomen	0.0176	0.00830	2.116	3.57e-02	*
${ m Hip}$	-0.0142	0.01013	-1.402	1.63e-01	

	Estimate	Std. Error	t value	$\Pr(> t )$	
Thigh	0.0087	0.01024	0.849	3.97e-01	
Ankle	0.0200	0.01927	1.039	3.00e-01	
${f Biceps}$	0.0268	0.01271	2.112	3.60e-02	*
$\mathbf{Wrist}$	0.0401	0.03859	1.039	3.00e-01	

# plot(rawlm)

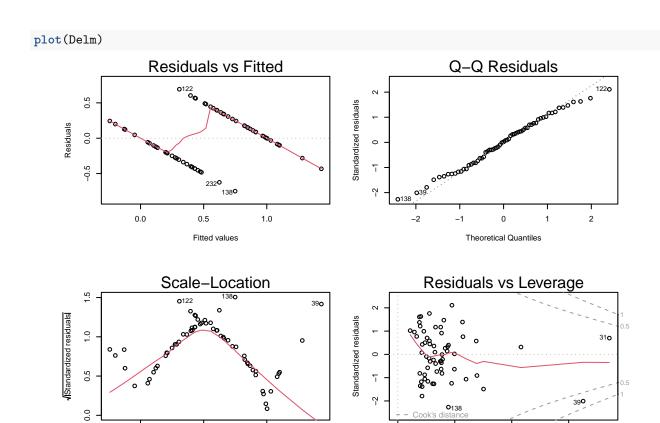


## Outcome-blind
par(mfrow=c(2,2),cex=0.5)

Delm <- lm(Overweight~.,OW\_Decorrelated\_test[,c("Overweight",names(modelOverweight\$coef)[-1])])
pander::pander(Delm,add.significance.stars=TRUE)

Table 58: Fitting linear model: Overweight  $\sim$  .

	Estimate	Std. Error	t value	$\Pr(> t )$	
(Intercept)	-1.923103	0.97461	-1.97320	5.37e-02	
$La\_BodyFat$	-0.002238	0.01292	-0.17315	8.63e-01	
$\mathbf{Chest}$	0.038521	0.00541	7.12645	2.82e-09	* * *
${f La\_Abdomen}$	0.015506	0.01297	1.19583	2.37e-01	
La_Hip	-0.006350	0.01331	-0.47691	6.35 e-01	
$La\_Thigh$	0.069865	0.02348	2.97580	4.39e-03	* *
La_Knee	-0.000289	0.03721	-0.00777	9.94e-01	
$La\_Ankle$	-0.042165	0.03006	-1.40281	1.67e-01	
$La\_Biceps$	0.012316	0.03435	0.35855	7.21e-01	
La_Wrist	-0.004661	0.08887	-0.05245	9.58 e-01	



## Outcome-Driven
par(mfrow=c(2,2),cex=0.5)

0.0

0.5

Fitted values

1.0

Delm <- lm(Overweight~.,OW\_Decorrelated\_testD[,c("Overweight",names(modelOverweightD\$coef)[-1])])
pander::pander(Delm,add.significance.stars=TRUE)</pre>

0.0

0.2

0.4

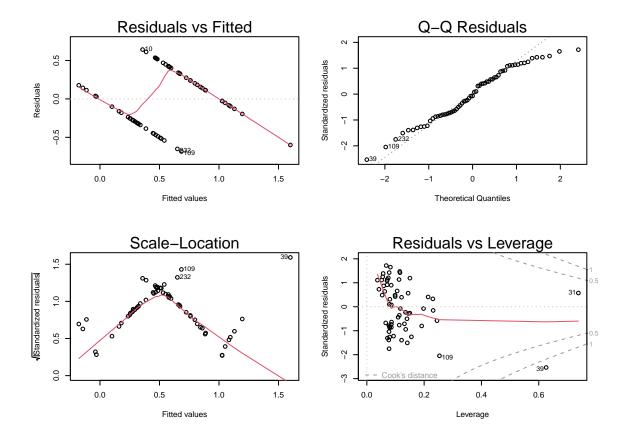
Leverage

0.6

Table 59: Fitting linear model: Overweight  $\sim$  .

	Estimate	Std. Error	t value	$\Pr(> t )$	
(Intercept)	-1.80519	1.00514	-1.796	7.80e-02	
La_BodyFat	-0.00573	0.01364	-0.420	6.76e-01	
$\mathbf{Chest}$	0.03517	0.00561	6.269	5.91e-08	* * *
$La\_Abdomen$	0.01632	0.01374	1.187	2.40e-01	
La_Hip	-0.00797	0.01388	-0.574	5.68e-01	
La_Ankle	-0.05489	0.03159	-1.738	8.79e-02	
$La\_Biceps$	0.04162	0.03492	1.192	2.38e-01	
${f La\_Wrist}$	-0.02560	0.09185	-0.279	7.81e-01	

plot(Delm)



## 5 Conclusion

In conclusion, ILAA (Iterative Linear Association Analysis), stands as an unsupervised computer-based methodology adept at estimating linear transformation matrices. These matrices enable the conversion of datasets into a fresh latent-based space, offering a user-controlled degree of correlation. This report has effectively demonstrated the practical application of ILAA, providing comprehensive insights into its functions for estimating, predicting, and scrutinizing transformations. Such capabilities hold significant promise in supervised learning scenarios, encompassing regression and logistic models.

# 6 Appendix A

## 6.1 The ILAA Algorithm

There are many transformation matrices W that moves a the observed feature space X affected by collinearity into new space Q, with controlled collinearity i.e.,:

$$Q = WX$$

and they can be found in many different ways. The ILAA algorithm aims to find a single realization based on variable residualization, and unit preserving transformations. i.e., it aims to find the Exploratory Residualization Transform (ERT).

The algorithm iterates by estimating the residuals of the linear fitting between observed feature pairs  $(x_i, x_j)$ :

$$x_j = \beta_{i,j} x_i + b_{i,j} + \epsilon_{i,j}.$$

Hence the residual variable:

$$q_j = x_j - \beta_{i,j} x_i, \quad \forall (i,j) \in \mathbb{B},$$

is a latent variable with zero correlation between the "independent" feature i and the "dependent" feature j.

The residualization is done on all the features pairs with significant association. Then, to address multiple collinearity the algorithm must be iterated.

The iteration:

$$W^n = W^{n-1}(I - \beta^n),$$

where  $\beta^n$  is the matrix of the beta coefficients of all independent variables to the dependent variables, will estimate the final shape of the ERT transformation until the desired maximum absolute correlation on Q is achieved.

The main issue with iterated realization is the heuristic decision of which feature is "independent" and who is the "dependent" feature. The ILAA function decides the hierarchy based on the degree of correlation and/or association to variables.

If we define  $\mathbb{B}$  as the basis vector of the observed matrix X, and we aim to return a Q matrix whose maximum correlation is lower than T, then the following steps estimate a single realization of W:

- 1. Set  $Q^1 = X$ , and  $W^1 = I$  as the initial solutions.
- 2. Compute the absolute of the correlation matrix of  $Q^1$ :  $R^1 = |corr(Q)|$  then set  $R^1_{i,i\in\mathbb{B}} = 0$ .
- 3. Extract the vector of the maximum correlation of each feature:

$$\rho = \max(R_{k}^1): \forall k \in \mathbb{B}$$

4. Extract the set of unaltered bases: U:

Set  $\mathbb{K} = \mathbb{B}$ , then sort  $\mathbb{K}$  from the most relevant feature to the least relevant then iterate trough all sorted features:

$$\mathbb{U}^{m+1} = \mathbb{U}^m \cup \{ \mathbb{K}_m \mid R^1_{\mathbb{K}_m \times \mathbb{U}^m} < T_0 \},$$

where  $T_0$  is the critical value of significant association between variables, and  $\mathbb{U}^1 = \emptyset$ .

5. At iteration n, define the  $\mathbb{I}$  (independent) and the  $\mathbb{J}$  (dependent) paired set:  $\mathbb{I} \times \mathbb{J}$ . First, get all the correlation pairs above the correlation goal:

$$\mathbb{H}^n = \{(i,j) \mid R_{i,i}^n \ge T : \forall (i,j) \in \mathbb{B}\}.$$

Second, sort the relevance of the pairs.

Third, interactively join all the pairs:

$$\mathbb{I}^m \times \mathbb{J}^m = \mathbb{I}^{m-1} \times \mathbb{J}^{m-1} \cup \{(i,j) \mid \rho_i \geq \rho_j : \forall (i,j) \in \mathbb{H}^n \land i \notin \mathbb{J}^{m-1} \land j \in \mathbb{T}\};$$

where

$$\mathbb{T} = \{ k \in \mathbb{B} \setminus (\mathbb{U} \cup \mathbb{I}^{m-1}) \},\$$

and

$$\mathbb{I}^1 = \mathbb{J}^1 = \emptyset.$$

6. Compute the  $\beta^n$  matrix loading based on the  $\mathbb{I} \times \mathbb{J}$  pairs. Estimate the slope coefficients  $\beta_{i,j}$  via linear fitting

$$x_i = \beta_{i,j} x_i + a_j, \quad \forall (i,j) \in \mathbb{I} \times \mathbb{J}$$

Then, set all non-statistically significant  $\beta_{i,j}$  values to zero. Statically significant values are defined by  $p(\beta_{i,j}) < p_{th}$ ; where  $p_{th}$  is the maximum probability of acceptance, and  $p(\beta_{i,j})$  is the probability that the slope coefficient is zero.

7. Update

$$W^n = W^{n-1}(I - \boldsymbol{\beta}^n),$$

then compute  $Q^n = W^n X$ .

- 8. Update the correlation matrix:  $R^n = |corr(Q^n)|$ , then set  $R^n_{j,j} = 0$ :  $\forall j \in \mathbb{B}$ , and the maximum correlation vector:  $\boldsymbol{\rho} = \max(R^n_{k,:})$ .
- 9. Repeat steps 5 to 8 until the maximum value of  $R^n$  is less than T, i.e.,  $\max(R^n) < T$ .

#### 6.1.1 Implementation Notes:

- 1. The correlation metric could be Person, or Spearman.
- 2. The critical value  $T_0$  at step 4 is the critical Pearson's correlation found using the t-distribution approximation. It is found by setting a significant q-value adjusted for multiple comparisons.  $q = 0.15/|\mathbb{B}|$ ;

```
q= 0.15/ncol(data)
ndf <- nrow(data)-2
tvalue <- qt(1.0 - q,ndf)
T_0 <- tvalue/sqrt(ndf + tvalue^2)</pre>
```

- 3. The FRESA.CAD::ILAA() linear fitting of step 6 is standard linear fitting for Pearson's Correlation, or Robust fitting for Spearman's rank correlation.
- 4. The  $p_{th}$  for statistical significance uses a q-value of 0.15 and it is adjusted for multiple comparisons using the Bonferroni correction:  $p_{th} = 0.15/|B|$ .
- 5. The ILAA implementation in FRESA.CAD uses relaxed iterations. i.e., the threshold values changes from an initial value of 0.95 until it reaches the user supplied threshold: T
- 6. In FRESA.CAD::ILAA the sorting of step 4 can be set to association to a user defined target (Driven) or the user can provide their own ranked variables.
- 7. In FRESA.CAD::ILAA the sorting of step 5 can be selected from maximum correlation or weighted correlation.
- 8. The correlation matrix R depends on the data. FRESA.CAD::ILAA() can be set to aggregate estimations of W with bootstrapped perturbations of the data.