A REPRINT OF "A METHOD OF ESTIMATING PLANE VULNERABILITY BASED ON DAMAGE OF SURVIVORS" BY ABRAHAM WALD

Abraham Wald

CENTER FOR NAVAL ANALYSES

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2.   This Research Contribution contains a series of
memoranda written by Abraham Wald of the Statistical
Research Group at Columbia University during World War II.
Unfortunately, this work was never published externally,
although some copies of his original memoranda have been
available and his methodology has been employed in the
analysis of data from both the Korean and Vietnam Wars.
It is published by CNA not only as a matter of historical
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CXA *Operations Evaluation Group*

CENTER FOR NAVAL ANALYSES

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FOREWORD

This CNA research contribution is composed of a series of
memoranda prepared in 1943 by Abraham Wald of the Statistical
Research Group (SRG), Columbia University, for the Applied
Mathematics Panel (AMP), National Defense Research Committee.
The memoranda present methods of estimating the vulnerability of
various parts of an aircraft on the basis of damage observed on
returning planes. Unfortunately, this work was never published
externally, although some copies of the original memoranda have
been available and the methodology has been employed in the
analysis of data in both the Korean and Vietnam wars. It is
published by CNA not only as a matter of historical interest but
also because the methodology is still relevant.

The eight memoranda in the series were published separately, but
actually represent parts I through VIII of a larger attempt to
address plane vulnerability. The parts are kept separate here,
and their original AMP and SRG memorandum numbers are given.

Only very minor format changes have been made to accommodate CNA
style and to smooth the transition from one part (memorandum) to
another. The substance and original wording, however, have been
retained.

Copies of the memoranda were acquired through the National
Archives in Washington, D.C.

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PART I

AN EQUATION SATISFIED BY THE PROBABILITIES THAT A PLANE WILL BE DOWNED BY i HITS[1]

INTRODUCTION

Denote by P_i (i = 1,2,..., ad inf.) the probability that a plane will be downed by i hits. Denote by p_i the conditional probability that a plane will be downed by the i-th hit knowing that the first i - 1 hits did not down the plane. Let $Q_i = 1 - P_i$ and $q_i = 1 - p_i$ (i = 1,2,..., ad inf.). It is clear that

$$Q_i = q_1 q_2 \cdots q_i \tag{1}$$

and

$$P_i = 1 - q_1 q_2 \cdots q_i . \tag{2}$$

Suppose that p_i and P_i (i = 1,2,...) are unknown and our information consists only of the following data concerning planes participating in combat:

- The total number N of planes participating in combat.

- For any integer i (i = 0,1,2,...) the number A_i of planes that received exactly i hits but have not been downed, i.e., have returned from combat.

Denote the ratio $\frac{A_i}{N}$ by a_i (i = 0,1,2,...) and let L be the proportion of planes lost. Then we have

$$\sum_{i=0}^{\infty} a_i = 1 - L. \tag{3}$$

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[1]This part of "A Method of Estimating Plane Vulnerability Based on Damage of Survivors" was published as SRG memo 85 and AMP memo 76.1.

The purpose of this memorandum is to draw inferences concerning the unknown probabilities p_i and P_i on the basis of the known quantities a_o, a_1, a_2,..., etc.

To simplify the discussion, we shall neglect sampling errors, i.e., we shall assume that N is infinity. Furthermore, we shall assume that

$$0 < p_i < 1 \qquad (i = 1,2,..., \text{ad inf.}). \qquad (4)$$

From equation 4 it follows that

$$0 < P_i < 1 \qquad (i = 1,2,..., \text{ad inf.}). \qquad (5)$$

We shall assume that there exists a non-negative integer n such that $a_n > 0$ but $a_i = 0$ for $i > n$.

We shall also assume that there exists a positive integer m such that the probability is zero that the number of hits received by a plane is greater than or equal to m. Let m' be the smallest integer with the property that the probability is zero that the number of hits received by a plane is greater than or equal to m'. Then the probability that the plane receives exactly m' - 1 hits is positive. We shall prove that m' = n + 1. Since $a_n > 0$, it is clear that m' must be greater than n. To show that m' cannot be greater than n + 1, let y be the proportion of planes that received exactly m' - 1 hits. Then y > 0 and $y(1 - p_{m'-1}) = a_{m'-1}$. Since y > 0 and $1 - p_{m'-1} > 0$, we have $a_{m'-1} > 0$. Since $a_i = 0$ for $i > n$, we see that $m' - 1 \leq n$, i.e., $m' \leq n + 1$. Hence, m' = n + 1 must hold.

Denote by x_i (i = 1,2,...) the ratio of the number of planes downed by the i-th hit to the total number of planes participating in combat. Since m' = n + 1, we obviously have $x_i = 0$ for $i > n$. It is clear that

$$\sum_{i=1}^{n} x_i = L = 1 - a_o - a_1 - \cdots - a_n . \qquad (6)$$

CALCULATION OF x_i IN TERMS OF a_o, a_1, ..., a_n, p_1, ..., p_n

Since the proportion of planes that received at least one hit is equal to $1 - a_o$, we have

$$x_1 = p_1(1 - a_o) . \tag{7}$$

The proportion of planes that received at least two hits and the first hit did not down the plane is obviously equal to $1 - a_o - a_1 - x_1$. Hence,

$$x_2 = p_2(1 - a_o - a_1 - x_1). \tag{8}$$

In general, we obtain

$$x_i = p_i(1 - a_o - a_1 - \ldots - a_{i-1} - x_1 - x_2 - \ldots - x_{i-1})$$
$$(i = 2, 3, \ldots, n) \tag{9}$$

Putting

$$c_i = 1 - a_o - a_1 - \ldots - a_{i-1} , \tag{10}$$

equation 9 can be written

$$x_i + p_i(x_1 + \ldots + x_{i-1}) = p_i c_i \quad (i = 2, 3, \ldots, n). \tag{11}$$

Substituting $i - 1$ for i, we obtain from equation 11

$$x_{i-1} + p_{i-1}(x_1 + \ldots + x_{i-2}) - p_{i-1}c_{i-1}$$
$$(i = 3, 4, \ldots, n). \tag{12}$$

Dividing by p_{i-1}, we obtain

$$\frac{x_{i-1}}{p_{i-1}} + (x_1 + \ldots + x_{i-2}) = c_{i-1} \quad (i = 3, 4, \ldots, n). \tag{13}$$

Adding $x_{i-1}\left(1 - \dfrac{1}{p_{i-1}}\right) = \dfrac{-q_{i-1}}{p_{i-1}}\, x_{i-1}$ to both sides of

equation 13, we obtain

$$x_1 + \ldots + x_{i-1} = c_{i-1} - \frac{q_{i-1}}{p_{i-1}}\, x_{i-1} \qquad (14)$$
$$(i = 3,4,\ldots,n+1).$$

From equations 11 and 14, we obtain

$$x_i + p_i\left(c_{i-1} - \frac{q_{i-1}}{p_{i-1}}\, x_{i-1}\right) = p_i c_i \; . \qquad (15)$$

Hence,

$$x_i = p_i(c_i - c_{i-1}) + \frac{p_i q_{i-1}}{p_{i-1}}\, x_{i-1} \quad (i = 3,4,\ldots,n). \qquad (16)$$

Let

$$d_i = p_i(c_i - c_{i-1}) = -p_i a_{i-1} \qquad (i = 3,4,\ldots,n) \qquad (17)$$

and

$$t_i = \frac{p_i q_{i-1}}{p_{i-1}} \qquad\qquad (i = 3,4,\ldots,n). \qquad (18)$$

Then equation 16 can be written as

$$x_i = d_i + t_i x_{i-1} \qquad\qquad (i = 3,4,\ldots,n). \qquad (19)$$

Denote $p_1(1 - a_o)$ by d_1, $-p_2 a_1$ by d_2, and $\dfrac{p_2 q_1}{p_1}$ by t_2; then we have

$$x_1 = d_1 \text{ and } x_2 = t_2 x_1 + d_2 \; . \qquad (20)$$

From equations 19 and 20, we obtain

$$\begin{cases} x_1 = d_1 \\ \\ x_i = \displaystyle\sum_{j=1}^{i-1} d_j t_{j+1} t_{j+2} \cdots t_i + d_i \qquad (i = 2, 3, \ldots, n). \quad (21) \end{cases}$$

EQUATION SATISFIED BY $q_1, \ldots, q_n$

To derive an equation satisfied by $q_1, \ldots, q_n$, we shall express $\displaystyle\sum_{i=1}^{n} x_i$ in terms of the quantities t_i and d_i $(i = 1, \ldots, n)$.

Substituting i for i - 1 in equation 14, we obtain

$$\tag{22}$$

$$x_i = \sum_{j=1}^{i} x_j = c_i - \frac{q_i}{p_i} x_i = c_i - \frac{q_i}{p_i} \left[\sum_{j=1}^{i-1} (d_j t_{j+1} \cdots t_i) + d_i \right] .$$

Hence, in particular

$$X_n = \sum_{j=1}^{n} x_j = c_n - \frac{q_n}{p_n} \left[\sum_{j=1}^{n-1} (d_j t_{j+1} \cdots t_n) + d_n \right] = L . \tag{23}$$

Since $c_n - L = a_n$, and since $t_{j+1} \cdots t_n = \dfrac{p_n}{p_j} q_j \cdots q_{n-1}$, we

obtain from equation 23

$$a_n - \left(\sum_{j=1}^{n-1} \frac{d_j}{p_j} q_j \cdots q_n \right) + q_n a_{n-1} = 0 . \tag{24}$$

Dividing by $q_1 \cdots q_n$ and substituting $-p_j a_{j-1}$ for d_j, we obtain

$$\frac{a_n}{q_1 \cdots q_n} + \frac{a_{n-1}}{q_1 \cdots q_{n-1}} - \sum_{j=1}^{n-1} \frac{d_j}{p_j q_1 \cdots q_{j-1}}$$

$$= \frac{a_n}{q_1 \cdots q_n} + \frac{a_{n-1}}{q_1 \cdots q_{n-1}}$$

$$+ \sum_{j=2}^{n-1} \frac{a_{j-1}}{q_1 \cdots q_{j-1}} - \frac{d_1}{p_1} \tag{25}$$

$$= \sum_{j=1}^{n} \frac{a_j}{q_1 \cdots q_j} - (1 - a_o) = 0$$

or

$$\sum_{j=1}^{n} \frac{a_j}{q_1 \cdots q_j} = 1 - a_o . \tag{26}$$

If it is known a priori that $q_1 = \ldots = q_n$, then our problem is completely solved. The common value of $q_1, \ldots, q_n$ is the root (between 0 and 1) of the equation

$$\sum_{j=1}^{n} \frac{a_j}{q^j} = 1 - a_o .$$

It is easy to see that there exists exactly one root between zero and one. We can certainly assume that $q_1 \geq q_2 \geq \ldots \geq q_n$. We shall investigate the implications of these inequalities and equation 26 later.

ALTERNATIVE DERIVATION OF EQUATION 26

Let b_i be the hypothetical proportion of planes that would have been hit exactly i times if dummy bullets would have been used. Clearly $b_i \geq a_i$. Denote $b_i - a_i$ by y_i ($i = 0, 1, 2, \ldots, n$). Of course, $b_o = a_o$, i.e., $y_o = 0$. We have $\sum_{j=0}^{n} b_i = 1$. Clearly

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$$y_i = P_i b_i = P_i(a_i + y_1) \qquad (i = 1, 2, \ldots, n). \qquad (27)$$

Hence,

$$y_i = \frac{P_i}{Q_i} a_i = \frac{1 - q_1 \cdots q_i}{q_1 \cdots q_i} a_i = \frac{a_i}{q_1 \cdots q_i} - a_i . \qquad (28)$$

Since $\sum\limits_{i=1}^{n} y_i = L$, we obtain from equation 28

$$\sum_{i=1}^{n} \frac{a_i}{q_1 \cdots q_i} = L + \sum_{i=1}^{n} a_i = 1 - a_o . \qquad (29)$$

This equation is the same as equation 26. This is a simpler derivation than the derivation of equation 26 given before. However, equations 21 and 22 (on which the derivation of equation 26 was based) will be needed later for other purposes.

As mentioned before, equation 29 leads to a solution of our problem if it is known that $q_1 = \ldots = q_n$. In the next memorandum (part II) we shall investigate the implications of equation 29 under the condition that $q_1 \geq q_2 \geq \ldots \geq q_n$.

NUMERICAL EXAMPLES

N is the number of planes participating in combat. A_o, A_1, A_2, $\ldots, A_n$ are the number returning with no hits, one hit, two hits, $\ldots, n$ hits, respectively. Then

$$a_i = \frac{A_i}{N} \qquad (i = 0, 1, 2, \ldots, n)$$

i.e., a_i is the proportion of planes returning with i hits. The computations below were performed under the following two assumptions:

- The bombing mission is representative so that there is no sampling error.

- The probability that a plane will be shot down does not depend on the number of previous non-destructive hits.

Example 1: Let $N = 400$
and $A_o = 320$ then $a_o = .80$

$A_1 = 32$ $a_1 = .08$

$A_2 = 20$ $a_2 = .05$

$A_3 = 4$ $a_3 = .01$

$A_4 = 2$ $a_4 = .005$

$A_5 = 2$ $a_5 = .005$

We assume $q_1 = q_2 = \ldots = q_5 = q_i$, where q_i is the probability of a plane surviving the i-th hit, knowing that the first $i - 1$ hits did not down the plane.

Then equation 26,

$$\sum_{j=1}^{n} \frac{a_j}{q_1 \cdots q_j} = 1 - a_o ,$$

reduces to

$$\sum_{j=1}^{n} \frac{a_j}{q^j} = 1 - a_o .$$

Substituting values of a_i

$$\frac{.08}{q} + \frac{.05}{q^2} + \frac{.01}{q^3} + \frac{.005}{q^4} + \frac{.005}{q^5} = .20$$

or

$$.200q^5 - .080q^4 - .050q^3 - .010q^2 - .005q - .005 = 0.$$

The Birge-Vieta method of finding roots described in Marchant Method No. 225 is used to solve this equation (table 1). We find $q = q_i = .851$, $p_i = .149$ where p_i is the probability of a plane being downed by the i-th hit, knowing that the first $i - 1$ hits did not down the plane.

x_i equals the ratio of the number of planes downed by the i-th hit to the total number of planes participating in combat. Using equation 9

$$x_i = p_i(1 - a_o - a_1 - \ldots - a_{i-1} - x_1 - x_2 - \ldots - x_{i-1})$$

$$(i = 2,3,\ldots,n)$$

for $n = 5$, we obtain

$$x_1 = p_1(1 - a_o) = .030$$
$$x_2 = p_2(1 - a_o - a_1 - x_1) = .013$$
$$x_3 = p_3(1 - a_o - a_1 - a_2 - x_1 - x_2) = .004$$
$$x_4 = p_4(1 - a_o - a_1 - a_2 - a_3 - x_1 - x_2 - x_3) = .002$$
$$x_5 = p_5(1 - a_o - a_1 - a_2 - a_3 - a_4 - x_1 - x_2 - x_3 - x_4) = .001$$

Example 2: Let $a_o = .3$, $a_1 = .2$, $a_2 = .1$, $a_3 = .1$, $a_4 = .05$, and $a_5 = .05$. Then the following results are obtained: $q = .87$, $p = 1 - q = .13$, $x_1 = .09$, $x_2 = .05$, $x_3 = .03$, $x_4 = .02$, and $x_5 = .01$.

The value of q in the second example is nearly equal to the value in the first example in spite of the fact that the values a_i ($i = 0,1,\ldots,5$) differ considerably. The difference in the values a_i in these two examples is mainly due to the fact that the probability that a plane will receive a hit is much smaller in the first example than in the second example. The probability that a plane will receive a hit has, of course, no relation to the probability that a plane will be downed if it receives a hit.

TABLE 1

1. Assume $q = 1 = y_1$

.200	-.080	-.050	-.010	-.005	-.005
	+.200	+.120	+.070	+.060	+.055
.200	+.120	+.070	+.060	+.055	+.050 = A_0
	+.200	+.320	+.390	+.450	
.200	+.320	+.390	+.450	+.505 = A_1	

$$y_2 = y_1 - \frac{A_0}{A_1} = 1 - .0990 = .9010$$

2. Assume $q = .9010 = y_2$

.2000	-.0800	-.0500	-.0100	-.0050	-.0050
	+.1802	+.0903	+.0363	+.0237	+.0168
.2000	+.1002	+.0403	+.0263	+.0187	+.0118 = B_0
	+.1802	+.2526	+.2639	+.2615	
.2000	+.2804	+.2929	+.2902	+.2802 = B_1	

$$y_3 = y_2 - \frac{B_0}{B_1} = .9010 - .042113 = .858887$$

3. Assume $q = .858887 = y_3$

.200000	-.060000	-.050000	-.010000	-.005000	-.005000
	+.171777	+.078826	+.024758	+.012675	+.006592
.200000	+.091777	+.028826	+.014758	+.007675	+.001592 = C_0
	+.171777	+.226363	+.219179	+.200925	
.200000	+.263554	+.255189	+.233937	+.208600 = C_1	

$$y_4 = y_3 - \frac{C_0}{C_1} = .858887 - .007632 = .851255$$

4. Assume $q = .851255 = y_4$

.2000000	+.080000	-.050000	-.010000	-.005000	-.005000
	+.170251	+.076827	+.022837	+.010928	+.005046
.2000000	+.090251	+.026827	+.012837	+.005928	+.000046 = D_0
	+.170251	+.221754	+.211606	+.191058	
.2000000	+.260502	+.248531	+.224443	+.196986 = D_1	

$$y_5 = y_4 - \frac{D_0}{D_1} = .851255 - .000234 = .851021$$

PART II

MAXIMUM VALUE OF THE PROBABILITY THAT A PLANE WILL BE DOWNED BY A GIVEN NUMBER OF HITS[1]

The symbols defined and the results obtained in part I will be used here without further explanation. The purpose of this memorandum is to derive the least upper bound of $X_i = \sum_{j=1}^{i} x_j$ and that of P_i ($i = 1, \ldots, n$) under the restriction that

$$q_1 \geq q_2 \geq \ldots, \geq q_n.$$

First, we shall show that X_i is a strictly increasing function of p_j for $j \leq i$. Let us replace p_j by $p_j + \Delta$ ($\Delta > 0$) and let us study the effect of this change on $x_1, \ldots, x_i$. Denote the changes in $x_1, \ldots, x_i$ by $\Delta_1, \ldots, \Delta_i$, respectively. Clearly, $\Delta_1 = \ldots = \Delta_{j-1} = 0$. It follows easily from equation 9 that

$\Delta_j > 0$ and

$$\Delta_{j+1} = -p_{j+1} \, \Delta_j \ .$$

Hence,

$$\Delta_j + \Delta_{j+1} = (1 - p_{j+1}) \, \Delta_j > 0.$$

Similarly, we obtain from equation 9

$$\Delta_{j+2} = -p_{j+2}(\Delta_j + \Delta_{j+1}) = -p_{j+2}(1 - p_{j+1}) \, \Delta_j \ .$$

Hence,

$$\Delta_j + \Delta_{j+1} + \Delta_{j+2} = (1 - p_{j+2})(1 - p_{j+1}) \, \Delta_j > 0.$$

In general

$$\Delta_j + \Delta_{j+1} + \ldots + \Delta_{j+k} = (1 - p_{j+1}) \ldots (1 - p_{j+k}) \, \Delta_j > 0$$

$$(k = 1, \ldots, i-j)$$

Hence, we have proved that X_i is a strictly increasing function of P_j ($j = 1, \ldots, i$).

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[1]This part of "A Method of Estimating Plane Vulnerability Based on Damage of Survivors" was published as SRG memo 87 and AMP memo 76.2.

On the basis of the inequalities $p_i \geq p_{i-1}$, we shall derive the least upper bound of X_i. For the purpose of this derivation we shall admit 0 and 1 as possible values of p_i ($i = 1,\ldots,n$), thus making the domain of all possible points $(p_1,\ldots,p_n)$ to be a closed and bounded subset of the n-dimensional Cartesian space. Since X_i is a continuous function of the probabilities p_1, p_2, etc. (X is a polynomial in $p_1,\ldots,p_i$), the maximum of X_i exists and coincides, of course, with the least upper bound. Hence, our problem is to determine the maximum of X_i.

First, we show that the value of X_i is below the maximum if $p_n > p_i$. Assume that $p_n > p_i$ and let k be the smallest positive integer for which $p_k > p_i$. Obviously $k > i$. Let $p_j' = p_j(1 + \varepsilon)$ for $j = 1,\ldots,k-1$, and $p_j' = p_j(1 - \eta)$ for $j = k,k+1,\ldots,n$, where $\varepsilon > 0$ and η is a function $\eta(\varepsilon)$ of ε determined so that $\sum_{j=1}^{n} x_j' = L$ (x_j' is the proportion of planes that would have been brought down with the j-th hit if $p_1',\ldots,p_n'$ were the true probabilities). Since X_r ($r = 1,\ldots,n$) is a strictly monotonic function of $p_1,\ldots,p_r$, it is clear that for sufficiently small such a function $\eta(\varepsilon)$ exists. It is also clear that for sufficiently small ε the condition $p'_1 \leq p'_2 \leq \ldots \leq p'_n$ is fulfilled. Since $p_j' > p_j$ ($j = 1,\ldots,i$), we see that $X_i' > X_i$ (X_i does not depend on p_r' for $r > i$). Hence, we have proved that if $p_1,\ldots,p_n$ is a point at which X_i becomes a maximum, we must have $p_i = p_{i+1} = \ldots = p_n$.

Now we shall show that if X_i is a maximum then $p_1 = p_2 = \ldots = p_i$. For this purpose assume that $p_i > p_1$ and we shall derive a contradiction. Let j be the greatest integer for which $p_j = p_1$. Since $p_i > p_1$, we must have $j < i$. Let $p_r' = p_r(1 + \varepsilon)$ for $r = 1,\ldots,j$ and $p_r' = p_r(1 - \eta)$ for $r = j+1,\ldots,i$, where $\varepsilon > 0$ and η is determined so that $\sum_{k=1}^{i} x_k' = \sum_{k=1}^{i} x_k$. Then for the probabilities $p_1',\ldots,p_i'$, $p_{i+1},\ldots,p_n$ the proportion of lost

planes is not changed, i.e., it is equal to L. Now let $p'_r = p'_i$ for $r > i$. Then the proportion L' of lost planes corresponding to $p'_1, \ldots, p'_n$ is less than L. Hence, there exists a positive Δ so that the proportion L" of lost planes corresponding to the probabilities $p''_r = p'_r (1 + \Delta)$ is equal to L. But, since $p''_r > p'_r$ $(r = 1, \ldots, i)$ we must have $\sum_{j=1}^{i} x''_j > \sum_{j=1}^{i} x'_j = \sum_{j=1}^{i} x_j$. Hence, we arrived at a contradiction and our statement that $p_1 = p_2 = \ldots = p_i$ is proved. Thus, we see that the maximum of X_i is reached when $p_1 = p_2 = \ldots = p_n$.

LEAST UPPER BOUND OF P_i

Now we shall calculate the least upper bound of P_i. Admitting the values 0 and 1 for p_j, the maximum of P_i exists and is equal to the least upper bound of P_i. Since $P_i = 1 - q_1 \ldots q_i$, maximizing P_i is the same as minimizing $q_1 \ldots q_i$. We know that $q_1, \ldots, q_n$ are subject to the restriction

$$\sum_{j=1}^{n} \frac{a_j}{q_1 \ldots q_j} = 1 - a_o \quad . \tag{30}$$

Let $q_1^o, \ldots, q_n^o$ be a set of values of $q_1, \ldots, q_n$ (satisfying equation 30) for which $q_1 \ldots q_j$ becomes a minimum. First, we show that $q_i^o = q_{i+1}^o = \ldots = q_n^o$. Suppose that $q_n^o < q_i^o$. Consider the set of probabilities $q'_r = q_r^o$ for $r \leq 1$ and $q'_r = q_i^o$ for $r > i$. Then

$$\sum_{j=1}^{n} \frac{a_j}{q'_1 \ldots q'_j} < 1 - a_o \quad .$$

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Hence, there exists a positive factor $\lambda < 1$ so that

$$\sum_{j=1}^{n} \frac{a_j}{q_1'' \cdots q_j''} = 1 - a_0 \quad ,$$

where $q_i'' = \lambda q_i'$ $(i = 1, \ldots, n)$. Then

$$q_1'' q_2'' \cdots q_1'' < q_1^0 q_2^0 \cdots q_i^0$$

in contradiction to our assumption that $q_1^0 \cdots q_i^0$ is a minimum. Hence, we have proved that $q_i^0 = \ldots = q_n^0$.

Now we show that there exists at most one value j such that $1 > q_j^0 > q_i^0$. Suppose there are two integers j and k such that $1 > q_j^0 \geq q_k^0 > q_i^0$. Let j' be the smallest integer for which $q_{j'}^0 = q_j^0$ and let k' be the largest integer for which $q_{k'}^0 = q_k^0$. Let $\bar{q}_{j'} = (1 + \varepsilon) q_{j'}^0$, $\bar{q}_{k'} = \frac{1}{1 + \varepsilon} q_{k'}^0$ $(\varepsilon > 0)$, and $\bar{q}_r = q_r^0$ for $r \neq j'$, $\neq k'$. Then

$$\bar{q}_1 \cdots \bar{q}_i = q_1^0 \cdots q_i^0 \quad \text{and} \quad \sum_{r=1}^{n} \frac{a_r}{\bar{q}_1 \cdots \bar{q}_r} < 1 - a_0 \quad .$$

Hence, there exists a positive factor $\lambda < 1$ such that

$$\sum_{r=1}^{n} \frac{a_r}{q_1^* \cdots q_r^*} = 1 - a_0 \quad ,$$

where $q_r^* = \lambda \bar{q}_r$. But $q_1^* \cdots q_i^* < \bar{q}_1 \cdots \bar{q}_i = q_1^0 \cdots q_i^0$, which contradicts the assumption that $q_1^0 \cdots q_i^0$ is a minimum. This proves our statement.

It follows from our results that the minimum of q_1 is the root of the equation

$$\sum_{r=1}^{n} \frac{a_r}{q_1^r} = 1 - a_0 \quad . \tag{32}$$

Now we shall calculate the minimum of $q_1 q_2$. First, we know that $q_i = q_2$ ($i \geq 2$) if $q_1 q_2$ be a minimum. Hence, we have to minimize $q_1 q_2$ under the restriction

$$\frac{1}{q_1}\left(a_1 + \frac{a_2}{q_2} + \frac{a_3}{q_2^2} + \dots + \frac{a_n}{q_2^{n-1}}\right) = 1 - a_0 \quad . \tag{33}$$

Using the Lagrange multiplier method we obtain the equations

$$q_2 - \frac{\lambda}{q_1^2}\left(a_1 + \frac{a_2}{q_2} + \frac{a_3}{q_2^2} + \dots + \frac{a_n}{q_2^{n-1}}\right) = 0 \tag{34}$$

$$\text{(Lagrange multiplier} = \lambda)$$

$$q_1 - \frac{\lambda}{q_1}\left(\frac{a_2}{q_2^2} + \frac{2a_3}{q_2^3} + \dots + \frac{(n-1)a_n}{q_2^n}\right) = 0 \quad . \tag{35}$$

Because of equation 33, we can write equation 34 as follows:

$$q_2 - \frac{\lambda}{q_1}(1 - a_0) = 0; \quad \lambda = \frac{q_1 q_2}{1 - a_0} \quad .$$

Substituting for λ in equation 35, we obtain

$$q_1 - \frac{1}{1 - a_0}\left(\frac{a_2}{q_2} + \frac{2a_3}{q_2^2} + \frac{3a_4}{q_2^3} + \dots + \frac{(n-1)a_n}{q_2^{n-1}}\right) = 0 \tag{36}$$

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or

$$q_1 = \frac{1}{1 - a_o} \left(\frac{a_2}{q_2} + \frac{2a_3}{q_2^2} + \dots + \frac{(n-1)a_n}{q_2^{n-1}} \right) \quad . \qquad (37)$$

On the other hand, from equation 33 we obtain

$$q_1 = \frac{1}{1 - a_o} \left(a_1 + \frac{a_2}{q_2} + \frac{a_3}{q_2^2} + \dots + \frac{a_n}{q_2^{n-1}} \right) \quad . \qquad (38)$$

Equating the right-hand sides of equations 37 and 38, we obtain

$$\frac{a_3}{q_2^2} + \frac{2a_4}{q_2^3} + \frac{3a_5}{q_2^4} + \dots + \frac{(n-2)a_n}{q_2^{n-1}} - a_1 = 0. \qquad (39)$$

It is clear that equation 39 has exactly one positive root. The root is less than or equal to 1 if and only if

$$a_3 + 2a_4 + 3a_5 + \dots + (n-2)a_n \le a_1 \quad . \qquad (40)$$

Equations 38 and 39 have exactly one positive root in q_1 and q_2. We shall show that if the roots satisfy the inequalities $1 \ge q_1 \ge q_2$, then for these roots $q_1 q_2$ becomes a minimum. We can assume that $2 < n$, since the derivation of the minimum value of $q_1 \dots q_n$ will be given later in this memorandum. It is clear that for any

value $q_1 > \dfrac{a_1}{1 - a_o}$ equation 38 has exactly one positive root in q_2. Denote this root by $\phi(q_1)$. Hence, $\phi(q_1)$ is defined for

all values $q_1 > \dfrac{a_1}{1 - a_o}$. It is easy to see that

$$\lim_{q_1 \to \frac{a_1}{1-a_0}} \phi(q_1) = +\infty$$

Hence (assuming $a_1 > 0$)

$$\lim_{q_1 \to \frac{a_1}{1-a_0}} \psi(q_1) = +\infty$$

where $\psi(q_1) = q_1 \phi(q_1)$.

It is clear that $\lim_{q_1 \to \infty} \phi(q_1) = 0$. Since $a_n > 0$, it follows from equation 38 that $q_1 [\phi(q_1)]^{n-1}$ has a positive lower bound when $q_1 \to \infty$. But then, since $n > 2$, $\lim_{q_1 \to \infty} q_1 \phi(q_1) = +\infty$. From the relations $\lim_{q_1 \to \frac{a_1}{1-a_0}} \psi(q_1) = \lim_{q_1 \to \infty} \psi(q_1) = +\infty$ it follows that the absolute minimum value of $\psi(q_1)$ is reached for some positive value q_1 . Since equations 38 and 39 have exactly one positive root in q_1 and q_2, the absolute minimum value of $\psi(q_1)$ must be reached for this root. This proves our statement that if the roots of equations 38 and 39 satisfy the inequalities $1 \geq q_1 \geq q_2$, then for these roots $q_1 q_2$ becomes a minimum consistent with our restrictions on q_1 and q_2. If $1 \geq q_1 \geq q_2$ is not satisfied by the roots of equations 38 and 39, then q_1 is equal either to 1 or to q_2 and the minimum value of $q_1 q_2$ is either $\phi(1)$ or q^2 , where q is the root of the equation

$$\sum_{r=1}^{n} \frac{a_r}{q_1^r} = 1 - a_o \quad .$$

Now we shall determine the minimum of $q_1 \ldots q_i$ $(2 < i < n)$. First, we determine the minimum M_{i1} of $q_1 \ldots q_i$ under the restriction that $q_2 = q_i$. Thus, we have to minimize $q_1 q_2^{i-1}$ under the restriction that

$$\frac{a_1}{q_1} + \frac{a_2}{q_1 q_2} + \frac{a_3}{q_1 q_2^2} + \ldots + \frac{a_n}{q_1 q_2^{n-1}} = 1 - a_o. \qquad (40a)$$

Using the Lagrange multiplier method, we obtain

$$q_2^{i-1} - \frac{\lambda}{q_1} \left(\frac{a_1}{q_1} + \ldots + \frac{a_n}{q_1 q_2^{n-1}} \right) = q_2^{i-1} - \frac{\lambda}{q_1} (1 - a_o) = 0; \qquad (41)$$

and

$$(i - 1) q_1 q_2^{i-2} - \frac{\lambda}{q_1} \left(\frac{a_2}{q_2^2} + \frac{2a_3}{q_2^3} + \ldots + \frac{(n-1)a_n}{q_2^n} \right) = 0 \quad . \qquad (41a)$$

Substituting $\dfrac{q_1 q_2^{i-1}}{1 - a_o}$ for λ (the value of λ obtained from equation 41), we obtain

$$(i - 1) q_1 - \frac{1}{1 - a_o} \left(\frac{a_2}{q_2} + \frac{2a_3}{q_2^2} + \ldots + \frac{(n-1)a_n}{q_2^{n-1}} \right) = 0 \quad . \qquad (42)$$

From equation 40a

$$(i - 1) q_1 - \frac{i - 1}{1 - a_o} \left(a_1 + \frac{a_2}{q_2} + \ldots + \frac{a_n}{q_2^{n-1}} \right) = 0 \quad . \qquad (43)$$

From equations 42 and 43, we obtain

$$(i - 1)a_1 + \frac{(i - 2)a_2}{q_2} + \frac{(i - 3)a_3}{q_2^2} + \ldots + \frac{(i - n)a_n}{q_2^{n-1}} = 0. \tag{44}$$

From Descartes' sign rule it follows that equation 44 has exactly one positive root.

Let $q_1 = q_1^o$ and $q_2 = q_2^o$ be the roots of the equations 43 and 44. If $1 \geq q_1^o \geq q_2^o$, then $M_{i1} = q_1^o(q_2^o)^{i-1}$. If $1 \geq q_1^o \geq q_2^o$ does not hold, then M_{i1} is either $(q')^i$ or $(q'')^{i-1}$, where q' is the root of the equation

$$\sum_{j=1}^{n} \frac{a_j}{(q')^j} = 1 - a_o \tag{45}$$

and q'' is the root of the equation

$$a_1 + \frac{a_2}{q''} + \frac{a_3}{(q'')^2} + \ldots + \frac{a_n}{(q'')^{n-1}} = 1 - a_o. \tag{46}$$

Let M_{ir} ($r = 2,\ldots,i-1$) be the minimum of $q_1 \ldots q_i$ under the restriction that $q_1 = \ldots = q_{r-1} = 1$ and $q_{r+1} = q_i$. Then M_{ir} can be calculated in the same way as M_{i1}; we have merely to make the substitutions

$$n^* = n - r + 1$$
$$a_o^* = a_o + a_1 + \ldots + a_{r-1}$$
$$a_j^* = a_{j+r-1} \qquad (j = 1,\ldots,n^*)$$
$$q_j^* = q_{j+r-1} \qquad (j = 1,\ldots,n^*)$$
$$i^* = i - r + 1,$$

and we have to calculate the minimum of $q_1^* \ldots q_{i^*}^*$. Thus, we have to solve the equations corresponding to equations 43 and 44, i.e., the equations

$$(i^* - 1)q_1^* - \frac{i^* - 1}{1 - a_o^*}\left(a_1^* + \frac{a_2^*}{q_2^*} + \frac{a_3^*}{(q_2^*)^2} + \ldots + \frac{a_{n^*}^*}{(q_2^*)^{n^*-1}}\right) = 0$$

$$(43^*)$$

and

$$(i^* - 1)a_1^* + \frac{(i^* - 2)a_2^*}{q_2^*} + \frac{(i^* - 3)a_3^*}{(q_2^*)^2} + \ldots + \frac{(i^* - n^*)a_{n^*}^*}{(q_2^*)^{n^*-1}} = 0.$$

$$(44^*)$$

Let $q_1^* = v_1$ and $q_2^* = v_2$ be the positive roots of the equations 43* and 44*. If $1 \geq v_1 \geq v_2$, then $M_{ir} = v_1 v_2^{i^*-1}$. If $1 \geq v_1 \geq v_2$ does not hold, then M_{ir} is equal to either $(v')^{i^*}$ or $(v'')^{i^*-1}$, where v' is the positive root of the equation

$$\sum_{j=1}^{n^*} \frac{a_j^*}{(v')^j} = 1 - a_o^*$$

$$(45^*)$$

and v'' is the positive root of the equation

$$a_1^* + \frac{a_2^*}{v''} + \frac{a_3^*}{(v'')^2} + \ldots + \frac{a_{n^*}^*}{(v'')^{n^*-1}} = 1 - a_o^*.$$

$$(46^*)$$

The minimum M_i of $q_1 \ldots q_i$ ($i = 2,3,\ldots,n-1$) is equal to the smallest of the $i - 1$ values $M_{i1},\ldots,M_{i,i-1}$.

Now we shall determine the minimum of $q_1 \ldots q_n$. We show that the minimum is reached when $q_1 = \ldots = q_{n-1} = 1$. Suppose that this is not true and we shall derive a contradiction. Let j be the smallest integer for which $q_j < 1$ ($j < n$). Let $\bar{q}_j = (1 + \epsilon)q_j$ ($\epsilon > 0$), $\bar{q}_n = \frac{q_n}{1 + \epsilon}$, and $\bar{q}_r = q_r$ for all $r \neq j, \neq n$. Then $\bar{q}_1 \ldots \bar{q}_n \ldots = q_1 \ldots q_n$ and

$$\sum_{r=1}^{n} \frac{a_r}{\overline{q}_1 \cdots \overline{q}_r} < 1 - a_o \quad .$$

Hence, there exists a positive $\lambda < 1$ such that

$$\sum_{r=1}^{n} \frac{a_r}{q_1^* \cdots q_r^*} = 1 - a_o ,$$

where

$$q_r^* = \lambda \overline{q}_r \quad .$$

But then $q_1^* \cdots q_n^* < \overline{q}_1 \cdots \overline{q}_n = q_1 \cdots q_n$ in contradiction to the assumption that $q_1 \cdots q_n$ is a minimum. Hence, we must have $q_1 = \cdots = q_{n-1} = 1$. Then, from equation 26 it follows that the minimum value of $q_1 \cdots q_n$ is given by

$$\frac{a_n}{1 - a_o - a_1 - \cdots - a_{n-1}} \quad .$$

If $i > 1$ but $< n$, the computation of the minimum value of $q_1 \cdots q_i$ is involved, since a large number of algebraic equations have to be solved. In the next part we shall discuss some approximation methods by means of which the amount of computational work can be considerably reduced.

APPROXIMATE DETERMINATION OF THE MAXIMUM VALUE OF THE PROBABILITY
THAT A PLANE WILL BE DOWNED BY A GIVEN NUMBER OF HITS[1]

The symbols defined in parts I and II will be used here without
further explanations. We have seen in part II that the exact
determination of the maximum value of P_i (i < n) involves a con-
siderable amount of computational work, since a large number of
algebraic equations have to be solved. The purpose of this
memorandum is to derive some approximations to the maximum of P_i
which can be computed much more easily than the exact values.

Let us denote the maximum of P_i by P_i^O and let $Q_i^O = 1 - P_i^O$.

Thus, Q_i^O is the minimum value of Q_i. Before we derive approxi-

mate values of P_i^O (or Q_i^O) we shall discuss some simplifications

that can be made in calculating the exact value P_i^O (or Q_i^O)

assuming 1 < i < n. We have seen in part II that Q_i^O is equal to

the smallest of the i - 1 values $M_{11}, \ldots, M_{i,i-1}$. We shall

make some simplifications in calculating M_{ir} (r = 1,...,i-1).

For this purpose consider the equation

$$\frac{a_r}{u} + \frac{a_{r+1}}{uv} + \ldots + \frac{a_n}{uv^{n-r}} = 1 - a_0 - a_1 - \ldots - a_{r-1}. \quad (47)$$

It is clear that for any value $u > \dfrac{a_r}{1 - a_0 - \ldots - a_{r-1}}$, equation

47 has exactly one positive root in v. Denote this root by $\phi_r(u)$.

Thus, $\phi_r(u)$ is defined for all values $u > \dfrac{a_r}{1 - a_0 - \ldots - a_{r-1}}$.

In all that follows we shall assume that $a_i > 0$ (i = 1,...,n).
We shall prove that

$$\lim_{u \to \frac{a_r}{1-a_0 - \ldots - a_{r-1}}} \left(u \left[\phi_r(u) \right]^{i-r} \right) = + \infty \quad (48)$$

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[1]This part of "A Method of Estimating Plane Vulnerability
Based on Damage of Survivors" was published as SRG memo 88 and
AMP memo 76.3.

and

$$\lim_{u \to \infty} \left(u \left[\phi_r(u) \right]^{i-r} \right) = +\infty \quad . \tag{49}$$

It follows easily from equation 47 that if $u \to \dfrac{u_r}{1 - a_o - \ldots - a_{r-1}}$, then $\phi_r(u) \to +\infty$. Since $i > r$, we see that equation 48 must hold. It follows easily from equation 47 that $\lim\limits_{u = +\infty} \phi_r(u) = 0$. We also see from equation 47 that if $u \to \infty$, the product $u \left[\phi_r(u) \right]^{n-r}$ must have a positive lower bound. Equation 49 follows from this and the fact that $\lim\limits_{u \to \infty} \phi_r(u) = 0$.

We have seen in part II that equations 43* and 44* have exactly one positive root in the unknowns, q_1^* and q_2^*. Let the root in q_1^* be u_{ir}^o. Then the root in q_2^* is equal to $\phi_r(u_{ir}^o)$. From equations 48 and 49 it follows that $u \left[\phi_r(u) \right]^{i-r}$ is strictly decreasing in the interval $\dfrac{a_r}{1 - a_o - \ldots - a_{r-1}} < u < u_{ir}^o$, and is strictly increasing in the interval $u_{ir}^o < u < +\infty$.

Denote by u_r' the positive root of the equation

$$\frac{a_r}{u} + \frac{a_{r+1}}{u^2} + \ldots + \frac{a_n}{u^{n-r+1}} = 1 - a_o - \ldots - a_{r-1}. \tag{50}$$

It is clear that $u_r' < 1$ and $\phi_r(u_r') = u_r'$. The value M_{ir} is equal to the smallest of the three values

$$u_r' \left[\phi_r(u_r') \right]^{i-r}, \quad \left[\phi_r(1) \right]^{i-r}, \text{ and } u_{ir}^o \left[\phi_r(u_{ir}^o) \right]^{i-r}.$$

A simplification in the calculation of M_{ir} can be achieved by the fact that in some areas M_{ir} can be determined without calculating the value u_{ir}^o . We consider three cases.

Case A: $\quad u_r' \left[\phi_r(u_r') \right]^{i-r} < \left[\phi_r(1) \right]^{i-r} \quad .$

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In this case,

$$M_{ir} = u'_r \left[\phi_r(u'_r) \right]^{i-r} \text{ if } \frac{d}{du} u \left[\phi_r(u) \right]^{i-r} \geq 0 \text{ for } u = u'_r$$

and

$$M_{ir} = u^o_{ir} \left[\phi_r(u^o_{ir}) \right]^{i-r} \text{ if } \frac{d}{du} u \left[\phi_r(u) \right]^{i-r} < 0 \text{ for } u = u'_r.$$

Case B: $u'_r \left[\phi_r(u'_r) \right]^{i-r} > \left[\phi_r(1) \right]^{i-r}$.

In this case,

$$\bar{M}_{ir} = \left[\phi_r(1) \right]^{i-r} \text{ if } \frac{d}{du} u \left[\phi_r(u) \right]^{i-r} \leq 0 \text{ for } u = 1$$

and

$$M_{ir} = u^o_{ir} \left[\phi_r(u^o_{ir}) \right]^{i-r} \text{ if } \frac{d}{du} u \left[\phi_r(u) \right]^{i-r} > 0 \text{ for } u = 1.$$

Case C: $u'_r \left[\phi(u'_r) \right]^{i-r} = \left[\phi(1) \right]^{i-r}$.

In this case,

$$M_{ir} = u^o_{ir} \left[\phi(u^o_{ir}) \right]^{i-r} .$$

We can easily calculate the value of $\frac{d}{du} u \left[\phi_r(u) \right]^{i-r}$ for $u = u'_r$ and $u = 1$. In fact, we have

$$\frac{d}{du} \left[\phi_r(u) \right]^{i-r} = \left[\phi_r(u) \right]^{i-r} + (i - r)u \left[\phi_r(u) \right]^{i-r-1} \frac{d\phi_r(u)}{du} \tag{51}$$

and $\frac{d\phi_r(u)}{du} = \frac{dv}{du}$ can be obtained from equation 47 as follows.

Denote $\dfrac{a_r}{u} + \dfrac{a_{r+1}}{uv} + \ldots + \dfrac{a_n}{uv^{n-r}}$ by $G(u,v)$. Then

$$\frac{d\phi_r(u)}{du} = \frac{dv}{du} = - \frac{\dfrac{\partial}{\partial u} G(u,v)}{\dfrac{\partial}{\partial v} G(u,v)}$$

<div align="right">(52)</div>

$$= \frac{-\dfrac{1}{u}\left(\dfrac{a_r}{u} + \dfrac{a_{r+1}}{uv} + \ldots + \dfrac{a_n}{uv^{n-r}}\right)}{\dfrac{1}{u}\left(\dfrac{a_{r+1}}{v^2} + \dfrac{2a_{r+2}}{v^3} + \ldots + \dfrac{(n-r)a_n}{v^{n-r+1}}\right)}$$

$$= \frac{-(1 - a_0 - a_1 - \ldots - a_{r-1})}{\left(\dfrac{a_{r+1}}{v^2} + \dfrac{2a_{r+2}}{v^3} + \ldots + \dfrac{(n-r)a_n}{v^{n-r+1}}\right)} \quad \left(v = \phi_r(u)\right).$$

On the basis of equations 51 and 52, we can easily obtain the value of $\dfrac{d}{du}u\left[\phi_r(u)\right]^{i-r}$ for $u = u'_r$ and $u = 1$ if u'_r and $\phi_r(1)$ have been calculated. If $u = u'_r$, then $\phi_r(u) = v = u'_r$; if $u = 1$, then $v = \phi_r(1)$.

Since $\phi_r(1)$ is equal to the root of the equation in v

$$a_r + \frac{a_{r+1}}{v} + \ldots + \frac{a_n}{v^{n-r}} = 1 - a_0 - a_1 - \ldots - a_{r-1},$$

it follows from equation 50 that

$$\phi_r(1) = u'_{r+1}.$$

<div align="right">(53)</div>

Thus, for carrying out the investigations of cases A, B, and C for $r = 1,\ldots,i-1$, we merely have to calculate $u'_1,\ldots,u'_i$.

If we want to calculate Q^O_i for all values $i < n$, then it seems best to compute first the n quantities $u'_1,\ldots,u'_n$.

Since $u'_r = \phi_r(u'_r)$ and $\phi_r(1) = u'_{r+1}$, we can say that M_{ir} is the smallest of the three values

$$\left(u'_r\right)^{i-r+1}, \quad \left(u'_{r+1}\right)^{i-r}, \quad \text{and } u^O_{ir}\left[\phi_r(u^O_{ir})\right]^{i-r} .$$

Since Q^O_i is equal to the minimum of the $i - 1$ values, $M_{i1},\ldots,M_{i,i-1}$, we see that

$$Q^O \leq t_i, \tag{54}$$

where

$$t_i = \text{Min}\left[(u'_1)^i, \ (u'_2)^{i-1},\ldots, \ (u'_{i-1})^2, \ u'_i\right] . \tag{55}$$

If n is large, it can be expected that Q^O_i will be nearly equal to t_i. Thus, t_i can be used as an approximation to Q^O_i. In order to see how good this approximation is, we shall derive a lower bound z_i for Q^O_i. If the difference $t_i - z_i$ is small, we are certain to have a satisfactory approximation to Q^O_i. If $t_i - z_i$ is large, then t_i still may be a good approximation to Q^O_i, since it may be that z_i is considerably below Q^O_i.

To obtain a lower bound z_i of Q^O_i, denote by y_j $(j = 0,1,\ldots,i-1)$ the proportion of planes (number of planes divided by the total number of planes participating in combat) that would be downed out of the returning planes with j hits if they were subject to $i - j$ additional hits. Then

$$P_i = y_0 + y_1 + \ldots + y_{i-1} + x_1 + x_2 + \ldots + x_i . \tag{56}$$

It is clear that $a_j P_i > y_j$ $(j = 0, 1, \ldots, i-1)$ and consequently

$$(a_0 + a_1 + \ldots + a_{i-1}) P_i > y_0 + y_1 + \ldots + y_{i-1} .$$

Hence,

$$\frac{y_0 + y_1 + \ldots + y_{i-1}}{a_0 + a_1 + \ldots + a_{i-1}} < P_i . \tag{57}$$

Equation 56 can be written

$$P_i = (a_0 + \ldots + a_{i-1}) \frac{y_0 + y_1 + \ldots + y_{i-1}}{a_0 + \ldots + a_{i-1}} \tag{58}$$

$$+ (1 - a_0 - \ldots - a_{i-1}) \frac{x_1 + \ldots + x_i}{1 - a_0 - \ldots - a_{i-1}} .$$

Hence, P_i is a weighted average of $\dfrac{y_0 + \ldots + y_{i-1}}{a_0 + \ldots + a_{i-1}}$ and

$\dfrac{x_i + \ldots + x_i}{1 - a_0 - \ldots - a_{i-1}}$. Then, from equation 57 it follows that

$$P_i < \frac{x_1 + \ldots + x_i}{1 - a_0 - a_1 - \ldots - a_{i-1}} . \tag{59}$$

Since $y_j > 0$, we obtain from equations 56 and 59

$$x_1 + \ldots + x_i < P_i < \frac{x_1 + \ldots + x_i}{1 - a_0 - \ldots - a_{i-1}} . \tag{60}$$

Hence,

$$1 - \frac{x_1 + \ldots + x_i}{1 - a_0 - \ldots - a_{i-1}} < Q_i < 1 - (x_1 + \ldots + x_i) . \tag{61}$$

In part II we have calculated the maximum value of $x_1 + \ldots + x_i$. Denote this maximum value by A_i. Then a lower bound of Q_i^o is given by

$$z_i = 1 - \frac{A_i}{1 - a_o - \ldots - a_{i-1}} < Q_i^o \; . \tag{62}$$

NUMERICAL EXAMPLE

The same notation will be used as in the numerical examples for part I. q_i is the probability of a plane surviving the i-th hit, knowing that the first $i - 1$ hits did not down the plane. Then the probability that a plane will survive i hits is given by

$$Q_i = q_1 q_2 \cdots q_i \; .$$

In part I it was assumed that

$$q_1 = q_2 = \ldots = q_i = q_o \quad \text{(say)},$$

which is equivalent to the assumption that the probability that a plane will be shot down does not depend on the number of previous non-destructive hits. Under this assumption

$$Q_i = q_o^i \; .$$

The example below is based on the assumption that

$$q_1 \geq q_2 \geq \ldots \geq q_n,$$

i.e., the probability of surviving the $i + 1$ hit is less than or equal to the probability of surviving the i-th hit. In this case, it is not possible to find an explicit formula for Q_i, but a lower bound can be obtained. That is, a value of Q_i can be found such that the actual value of Q_i must lie above it. The greatest lower bound is denoted by Q_i^o. Hence, we have

$$Q_i^o \leq Q_i \; .$$

If

$$P_i^o = 1 - Q_i^o,$$

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P_i^O is the least upper bound of P_i; that is, the probability of being downed by i bullets cannot be greater than P_i^O.

Since the computation of the exact value of Q_i^O is relatively complex, an approximate formula has been developed. This approximation is called t_i and $t_i \geq Q_i^O$. Another approximation (z_i) is available such that $z_i \leq Q_i^O$. However, z_i is not as accurate as t_i. Whenever the full computation is to be omitted, it is recommended that t_i be used.

The observed data of example 1, part I, will be used. Thus,

$$a_o = .80, \ a_1 = .08, \ a_2 = .05, \ a_3 = .01, \ a_4 = .005, \ a_5 = .005$$

The calculations are in three sections:

- The calculation of $t_i \geq Q_i^O$.
- The calculation of $z_i \leq Q_i^O$.
- The exact value of Q_i^O.

1. Calculation of t_i $(t_i \geq Q_i^O)$

(1) Calculate u_r', the positive root of equation 50:

$$\frac{a_r}{u} + \frac{a_{r+1}}{u^2} + \ldots + \frac{a_n}{u^{n-r+1}} = 1 - a_o - \ldots - a_{r-1} \ .$$

For r = 1, we obtain

$$\frac{a_1}{u} + \frac{a_2}{u^2} + \frac{a_3}{u^3} + \frac{a_4}{u^4} + \frac{a_5}{u^5} = 1 - a_o \ ,$$

which reduces to

$$.20u^5 - .08u^4 - .05u^3 - .01u^2 - .005u - .005 = 0$$

$$u_1' = .851 .$$

For r = 2,

$$\frac{a_2}{u} + \frac{a_3}{u^2} + \frac{a_4}{u^3} + \frac{a_5}{u^4} = 1 - a_o - a_1,$$

which reduces to

$$.12u^4 - .05u^3 - .01u^2 - .005u = 0$$

$$u_2' = .722 .$$

For r = 3,

$$\frac{a_3}{u} + \frac{a_4}{u^2} + \frac{a_5}{u^3} = 1 - a_o - a_1 - a_2 ,$$

which reduces to

$$.07u^3 - .01u^2 - .005u - .005 = 0$$

$$u_3' = .531 .$$

For r = 4,

$$\frac{a_4}{u} + \frac{a_5}{u^2} = 1 - a_o - a_1 - a_2 - a_3,$$

which reduces to

$$.06u^2 - .005u - .005 = 0$$

$$u_4' = .333 .$$

(2) $t_1, \ldots, t_5$ are given by equation 54:

$$t_i = \text{Min} \left[(u_1')^1, (u_2')^{i-1}, \ldots, (u_{i-1}')^2, (u_i') \right] .$$

We have

$$u_1' = .851, \quad u_2' = .722, \quad u_3' = .531, \quad u_4' = .333 .$$

Hence,

$$
\begin{aligned}
t_1 &= \text{Min} \left[(u_1') \right] = u_1' \\
&= .851
\end{aligned}
$$

$$
\begin{aligned}
t_2 &= \text{Min} \left[(u_1')^2, (u_2') \right] \\
&= \text{Min} \left[.724, .722 \right] \\
&= .722
\end{aligned}
$$

$$
\begin{aligned}
t_3 &= \text{Min} \left[(u_1')^3, (u_2')^2, (u_3') \right] \\
&= \text{Min} \left[.616, .521, .531 \right] \\
&= .521
\end{aligned}
$$

$$
\begin{aligned}
t_4 &= \text{Min} \left[(u_1')^4, (u_2')^3, (u_3')^2, (u_4') \right] \\
&= \text{Min} \left[.524, .376, .282, .333 \right] \cdot \\
&= .282
\end{aligned}
$$

t_5 is not calculated since the exact value of Q_5^o can be easily obtained.

2. **Calculation of z_i ($z_i \leq Q_i^o$)**

The following values must be obtained:

q_o, the root of equation 26A

$$\frac{a_1}{q} + \frac{a_2}{q^2} + \frac{a_3}{q^3} + \frac{a_4}{q^4} + \frac{a_5}{q^5} = 1 - a_o .$$

This has already been obtained as u_1' . Thus $q_o = .851$. The values of $x_1, \ldots, x_5$ have been calculated in part I:

$$x_1 = .030, \; x_2 = .013, \; x_3 = .004, \; x_4 = .002, \; x_5 = .001.$$

$$A_i = x_1 + x_2 + \ldots + x_i.$$

$$A_1 = x_1 = .030$$
$$A_2 = x_1 + x_2 = .043$$
$$A_3 = x_1 + x_2 + x_3 = .047$$
$$A_4 = x_1 + x_2 + x_3 + x_4 = .049$$
$$A_5 = x_1 + x_2 + x_3 + x_4 + x_5 = .050.$$

From equation 62 the lower bounds z_i are calculated:

$$z_i = 1 - \frac{A_i}{1 - a_o - \ldots - a_{i-1}} < Q_i^o.$$

Then

$$z_1 = 1 - \frac{A_1}{1 - a_o} = 1 - \frac{.030}{.20} = .850$$

$$z_2 = 1 - \frac{A_2}{1 - a_o - a_1} = 1 - \frac{.043}{.12} = .642$$

$$z_3 = 1 - \frac{A_3}{1 - a_o - a_1 - a_2} = 1 - \frac{.047}{.07} = .329$$

$$z_4 = 1 - \frac{A_4}{1 - a_o - a_1 - a_2 - a_3} = 1 - \frac{.049}{.06} = .183$$

z_5 is not calculated since Q_5^o can be obtained directly.

3. The Exact Value of Q_i^o

We have calculated t_i and z_i such that

$$z_i \leq Q_i^O \leq t_i \qquad (i = 1,2,\ldots,5) \ .$$

The exact value of Q_i^O is obtained as follows:

$$M_{ir} = \text{Min} \left\{ (u_r')^{i-r+1}, \ (u_{r+1}')^{i-r}, \ u_{ir}^O \left[\phi_r(u_{ir}^O) \right]^{i-r} \right\} ,$$

where u_{ir}^O and $\phi_r(u_{ir}^O)$ will be defined below.

$$Q_i^O = \text{Min} \ [M_{i1}, \ldots, M_{i,i-1}]$$

or combining these equations with the definition of t_i we obtain

$$Q_1^O = \text{Min} \ \{t_1\} = .851$$
$$Q_2^O = \text{Min} \ \{t_2, \ u_{21}^O [\ \phi_1(u_{21}^O)]\}$$
$$Q_3^O = \text{Min} \ \{t_3, \ u_{31}^O [\ \phi_1(u_{31}^O)]^2, \ u_{32}^O [\ \phi_2(u_{32}^O)]\}$$
$$Q_4^O = \text{Min} \ \{t_4, \ u_{41}^O [\ \phi_1(u_{41}^O)]^3, \ u_{42}^O [\ \phi_2(u_{42}^O)]^2, \ u_{43}^O [\ \phi_3(u_{43}^O)]\}$$

If $u_{ir}^O > 1$, $\left[\phi_r(u_{ir}^O) \right] > 1$, or $u_{ir}^O < \phi_r(u_{ir}^O)$, then

$u_{ir}^O \left[\phi_r(u_{ir}^O) \right]^{i-r}$ is neglected in the equations above.

$$Q_5^O = \frac{a_5}{1 - a_0 - a_1 - a_2 - a_3 - a_4} = \frac{.005}{.055} = .091 .$$

In the equation of Q_i^O the additional quantities we have to compute are

$$
\begin{array}{ll}
u_{21}^O & \phi_1(u_{21}^O) \\
u_{31}^O & \phi_1(u_{31}^O) \\
u_{32}^O & \phi_2(u_{32}^O) \\
u_{41}^O & \phi_1(u_{41}^O) \\
u_{42}^O & \phi_2(u_{42}^O) \\
u_{43}^O & \phi_3(u_{43}^O)
\end{array}
$$

The following equations have exactly one positive root in q_1^*, q_2^*. The root in q_1^* is u_{ir}^O; the root in q_2^* is $\phi_r(u_{ir}^O)$.

$$a_1^* + \frac{a_2^*}{q_2^*} + \frac{a_3^*}{(q_2^*)^2} + \ldots + \frac{a_{n^*}^*}{(q_2^*)^{n^*-1}} = (1 - a_o^*)q_1^* \text{ ,}$$

where q_2^* satisfies

$$(i^* - 1)a_1^* + \frac{(i^* - 2)a_2^*}{q_2^*} + \frac{(i^* - 3)a_3^*}{(q_2^*)^2} + \ldots + \frac{(i^* - n^*)a_{n^*}^*}{(q_2^*)^{n^*-1}} = 0,$$

where

$$n^* = n - r + 1$$

$$a_o^* = a_o + a_1 + \ldots + a_{r-1}$$

$$a_j^* = a_{j+r-1} \qquad (j = 1, 2, \ldots, n^*)$$

$$i^* = i - r + 1 \text{ .}$$

The details of the computation are given in tables 2 and 3.

TABLE 2

u_{ir}^O	i	r	n^*	i^*	a_o^*	a_1^*	a_2^*	a_3^*	a_4^*	a_5^*
u_{21}^O	2	1	5	2	.80	.08	.05	.01	.005	.005
u_{31}^O	3	1	5	3	.80	.08	.05	.01	.005	.005
u_{32}^O	3	2	4	2	.88	.05	.01	.005	.005	
u_{41}^O	4	1	5	4	.80	.08	.05	.01	.005	.005
u_{42}^O	4	2	4	3	.88	.05	.01	.005	.005	
u_{43}^O	4	3	3	2	.93	.01	.005	.005		

where

$$a_o = .80, \ a_1 = .08, \ a_2 = .05, \ a_3 = .01, \ a_4 = .005, \ a_5 = .005$$

TABLE 3

Computation of	Equation	Numerical Equation	Result Obtained
$\phi_1(u_{21}^o)$	$(i^*-1)a_1^* + \dfrac{(i^*-2)a_2^*}{q_2^*} + \dfrac{(i^*-3)a_3^*}{(q_2^*)^2} + \dfrac{(i^*-4)a_4^*}{(q_2^*)^3} + \dfrac{(i^*-5)a_5^*}{(q_2^*)^4} = 0$	$.08(q_2^*)^4 - .01(q_2^*)^2 - .01(q_2^*) - .015 = 0$	.774
u_{21}^o	$a_1^* + \dfrac{a_2^*}{q_2^*} + \dfrac{a_3^*}{(q_2^*)^2} + \dfrac{a_4^*}{(q_2^*)^3} + \dfrac{a_5^*}{(q_2^*)^4} = (1-a_0^*)q_1^*$	$.08 + \dfrac{.05}{.774} + \dfrac{.01}{(.774)^2} + \dfrac{.005}{(.774)^3} + \dfrac{.005}{(.774)^4} = .20q_1^*$	.932
$\phi_1(u_{31}^o)$	$(i^*-1)a_1^* + \dfrac{(i^*-2)a_2^*}{q_2^*} + \dfrac{(i^*-3)a_3^*}{(q_2^*)^2} + \dfrac{(i^*-3)a_4^*}{(q_2^*)^3} + \dfrac{(i^*-5)a_5^*}{(q^*_2)^4} = 0$	$.16(q_2^*)^4 + .05(q_2^*)^3 - .005(q_2^*) - .01 = 0$	.463
u_{31}^o	$a_1^* + \dfrac{a_2^*}{q_2^*} + \dfrac{a_3^*}{(q_2^*)^2} + \dfrac{a_4^*}{(q_2^*)^3} + \dfrac{a_5^*}{(q_2^*)^4} = (1-a_0^*)q_1^*$	$.08 + \dfrac{.05}{.463} + \dfrac{.01}{(.463)^2} + \dfrac{.005}{(.463)^3} + \dfrac{.005}{(.463)^4} = .20q_1^*$	1.968[a]
$\phi_2(u_{32}^o)$	$(i^*-1)a_1^* + \dfrac{(i^*-2)a^*_2}{q_2^*} + \dfrac{(i^*-3)a_3^*}{(q_2^*)^2} + \dfrac{(i^*-4)a_4^*}{(q_2^*)^3} = 0$	$.05(q^*_2)^3 - .005(q_2^*) - .01 = 0$	.642
u_{32}^o	$a_1^* + \dfrac{a_2^*}{q_2^*} + \dfrac{a_3^*}{(q_2^*)^2} + \dfrac{a_4^*}{(q_2^*)^3} = (1-a_0^*)q_1^*$	$.05 + \dfrac{.01}{.642} + \dfrac{.005}{(.642)^2} + \dfrac{.005}{(.642)^3} = .12q_1^*$	.805
$\phi_1(u_{41}^o)$	$(i^*-1)a_1^* + \dfrac{(i^*-2)a_2^*}{q_2^*} + \dfrac{(i^*-3)a_3^*}{(q_2^*)^2} + \dfrac{(i^*-4)a_4^*}{(q_2^*)^3} + \dfrac{(i^*-5)a_5^*}{(q_2^*)^4} = 0$	$.24(q_2^*)^4 + .10(q_2^*)^3 + .01(q_2^*)^2 - .005 = 0$	.290

TABLE 3 (Continued)

Computation of	Equation	Numerical Equation	Result Obtained
u^o_{41}	$a^*_1 + \dfrac{a^*_2}{q^*_2} + \dfrac{a^*_3}{(q^*_2)^2} + \dfrac{a^*_4}{(q^*_2)^3} + \dfrac{a^*_5}{(q^*_2)^4} = (1-a^*_o)q^*_1$	$.08 + \dfrac{.05}{.290} + \dfrac{.01}{(.290)^2} + \dfrac{.005}{(.290)^3} + \dfrac{.005}{(.290)^4} = .20q^*_1$	6.402[b]
$\phi_2(u^o_{42})$	$(i^*-1)a^*_1 + \dfrac{(1^*-2)a^*_2}{q^*_2} + \dfrac{(i^*-3)a^*_3}{(q^*_2)^2} + \dfrac{(i^*-4)a^*_4}{(q^*_2)^3} = 0$	$.10(q^*_2)^3 + .01(q^*_2)^2 - .005 = 0$	$.338$
u^o_{42}	$a^*_1 + \dfrac{a^*_2}{q^*_2} + \dfrac{a^*_3}{(q^*_2)^2} + \dfrac{a^*_4}{(q^*_2)^3} = (1-a^*_o)q^*_1$	$.05 + \dfrac{.01}{.338} + \dfrac{.005}{(.338)^2} + \dfrac{.005}{(.338)^3} = .12q^*_1$	2.108[c]
$\phi_3(u^o_{43})$	$(i^*-1)a^*_1 + \dfrac{(i^*-2)a^*_2}{q^*_2} + \dfrac{(i^*-3)a^*_3}{(q^*_2)^2} = 0$	$.01(q^*_2)^2 - .005 = 0$	$.707$
u^o_{43}	$a^*_1 + \dfrac{a^*_2}{q^*_2} + \dfrac{a^*_3}{(q^*_2)^2} = (1-a^*_o)q^*_1$	$.01 + \dfrac{.005}{.707} + \dfrac{.005}{(.707)^2} = .07q^*_1$	$.387$[d]

[a] $1.968 > 1$ ∴ $[u^o_{31} \, \phi_1(u^o_{31})]^2$ is not used.

[b] $6.402 > 1$ ∴ $[u^o_{41} \, \phi_1(u^o_{41})]^3$ is not used.

[c] $2.108 > 1$ ∴ $u^o_{42}[\phi_2(u^o_{42})]^2$ is not used.

[d] $.387 < \phi_3(u^o_{43})$ ∴ $u^o_{43}[\phi_3(u^o_{43})]$ is not used.

Substituting the values from table 3 in equation A and neglecting several terms as explained in table 3, we have

$$Q_1^o = .851$$

$$Q_2^o = Min \ \{.722, \ .721\} = .721$$

$$Q_3^o = Min \ \{.521, \ .517\} = .517$$

$$Q_4^o = .282$$

$$Q_5^o = .091$$

The results obtained are shown in table 4.

TABLE 4

i	z_i	Q_i^o	t_i	q_o^i
1	.851	.851	.851	.851
2	.642	.721	.722	.724
3	.329	.517	.521	.616
4	.183	.282	.282	.524
5	——	.091	——	.446

Thus, with the observed data, this example, if all the information available about the q_i's is that

$$q_1 \geq q_2 \geq \cdots \geq q_5 \ ,$$

all we can say about the Q_i is that

$$Q_1 \geq .85, \ Q_2 \geq .72, \ Q_3 \geq .52, \ Q_4 \geq .28, \ Q_5 = .09 \ .$$

Note that

$$z_1 = Q_1^o = t_1 = q_o \ .$$

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This is always true.

It is interesting to compare Q_i^O with the values of Q_i obtained under the assumption that all the q_i's are equal and have the value q_O. Under this assumption,

$$Q_i = q_O^i \qquad (i = 1, 2, \ldots, 5).$$

In table 4, $Q_1^O = q_O$ and Q_2^O is very close to q_O^2. Q_3^O and q_O^3 differ by approximately .1 and the agreement between Q_i^O and q_O^i gets progressively worse. It will usually be true that q_O^i and Q_i^O are approximately equal for small values of i; but will differ widely as i increases.

MINIMUM AND MAXIMUM VALUE OF THE PROBABILITY THAT A PLANE WILL BE DOWNED BY A GIVEN NUMBER OF HITS CALCULATED UNDER SOME FURTHER RESTRICTIONS ON THE PROBABILITIES $q_1,\ldots,q_n$ [1]

In parts I, II, and III we merely assumed that $q_1 \geq q_2 \geq \cdots \geq q_n$
In many cases we may have some further a priori knowledge concerning the values $q_1,\ldots,q_n$. We shall consider here the case when it is known a priori that $\lambda_1 q_j \leq q_{j+1} \leq \lambda_2 q_j$ ($j = 1,\ldots,n-1$), where λ_1 and λ_2 ($\lambda_1 < \lambda_2 < 1$) are known positive constants.

We shall also assume that

$$\sum_{j=1}^{n} \frac{a_j}{\lambda_1^{\frac{j(j-1)}{2}}} < 1 - a_o . \tag{63}$$

Since $a_1 + a_2 + \cdots + a_n < 1 - a_o$, the inequality in equation 63 is certainly fulfilled if λ_1 is sufficiently near 1. It follows immediately from equations 63 and 26 that $q_1 < 1$.

CALCULATION OF THE MINIMUM VALUE OF $Q_i = 1 - P_i$ ($i < n$)

Let $q_1^o,\ldots,q_n^o$ be the values of $q_1,\ldots,q_n$ for which Q_i becomes a minimum. We shall prove the following.

Lemma 1: The relations

$$q_{j+1}^o = \lambda_2 q_j^o \qquad\qquad (j = i,\ldots,n-1) \tag{64}$$

must hold.

Proof: Suppose that the relation in equation 64 does not hold for at least one value $j \geq i$ and we shall derive a contradiction.

Let $q'_r = q^o_r$ for $r = 1,\ldots,i$ and $q'_{j+1} = \lambda_2 q'_j$ for $j = i,\ldots,n-1$. Then we have

$$q'_1 \cdots q'_i = q^o_1 \cdots q^o_i \text{ and } \sum_{j=1}^{n} \frac{a_j}{q'_1 \cdots q'_j} < 1 - a_o. \quad (65)$$

Hence, there exists a positive value $\Delta < 1$ such that

$$\sum_{j=1}^{n} \frac{a_j}{q''_1 \cdots q''_j} = 1 - a_o,$$

where $q''_j = \Delta q'_j$ $(j = 1,\ldots,n)$. But then

$$q''_1 \cdots q''_i < q'_1 \cdots q'_i \, q^o_1 \cdots q^o_i$$

in contradiction to our assumption that $q^o_1 \cdots q^o_i$ is a minimum. Hence, Lemma 1 is proved.

<u>Lemma 2:</u> If j is the smallest integer such that $q^o_{k+1} = \lambda_2 q^o_k$ for all $k \geq j$, then $q^o_r = \lambda_1 q^o_{r-1}$ for $r = 2,3,\ldots,j-1$.

Proof: Assume that Lemma 2 does not hold and we shall derive a contradiction. Let u be the smallest integer greater than one such that $q^o_u > \lambda_1 q^o_{u-1}$. It follows from the definition of the integer u that if $u > 2$, then $q^o_{u-1} = \lambda_1 q^o_{u-2}$. From assumption 63 it follows that $q^o_1 < 1$. Hence, if we replace q^o_{u-1} by $q'_{u-1} = (1 + \varepsilon)q^o_{u-1}$ $(\varepsilon > 0)$, then for sufficiently small ε the inequalities $\lambda_1 q_r \leq q_{r+1} \leq \lambda_2 q_r$ $(r = 1,\ldots,n-1)$ will not be disturbed. Let v be the smallest integer greater than or equal to u such that $q^o_{v+1} < \lambda_2 q^o_v$. Since by assumption j is the smallest integer such that $q^o_{k+1} = \lambda_2 q^o_k$ for all $k \geq j$, we must have $q_j < \lambda_2 q_{j-1}$. Hence, $v \leq j-1$. It is clear that replacing q^o_v by $q'_v = \dfrac{q^o_v}{1 + \varepsilon}$ we shall not disturb the inequalities $\lambda_1 q_r \leq q_{r+1} \leq \lambda_2 q_r$ $(r = 1,\ldots,n-1)$. Hence, if

$$q'_{u-1} = (1 + \varepsilon) q^O_{u-1}, \quad q'_v = \frac{q^O_v}{1 + \varepsilon}, \quad \text{and} \quad q'_r = q^O_r$$

for $r \neq u, \neq v$, then $\lambda_1 q'_k \leq q'_{k+1} \leq \lambda_2 q'_k$ ($k = 1, \ldots, n-1$) is fulfilled. Furthermore, we have

$$q'_1 \ldots q'_i = q^O_1 \ldots q^O_i \quad \text{and} \quad \sum_{j=1}^{n} \frac{a_j}{q'_1 \ldots q'_j} < 1 - a_o.$$

Hence, there exists a positive $\Delta < 1$ such that

$$\sum_{j=1}^{n} \frac{a_j}{q''_1 \ldots q''_j} = 1 - a_o$$

and $q''_j = \Delta q'_j$ ($j = 1, \ldots, n$). But then

$$q''_1 \ldots q''_i < q'_1 \ldots q'_i = q^O_1 \ldots q^O_i$$

in contradiction to the assumption that $q^O_1 \ldots q^O_i$ is a minimum. Hence, Lemma 2 is proved.

Let E_{ir} ($r = 1, \ldots, i-1$) be the minimum value of Q_i under the restriction that $q_{j+1} = \lambda_2 q_j$ for $j = r+1, \ldots, n-1$ and $q_{j+1} = \lambda_1 q_j$ for $j = 1, \ldots, r-1$. From Lemma 1 and 2 it follows that the minimum of Q_i is equal to the smallest of the $i - 1$ values $E_{i1}, \ldots, E_{i,i-1}$. The computation of the exact value of E_{ir} can be carried out in a way similar to the computation of M_{ir} described in part II. Since these computations are involved if n is large, we shall discuss here an approximation method.

Let E^*_{ir} ($r = 1, \ldots, i-1$) be the value of Q_i if $q_{j+1} = \lambda_2 q_j$ for $j = r+1, \ldots, n-1$ and $q_{j+1} = \lambda_1 q_j$ for $j = 1, \ldots, r$. Furthermore, let E^*_{io} be the value of Q_i if $q_{j+1} = \lambda_2 q_j$ ($j = 1, \ldots, n-1$). Then, if n is large, the minimum of $E^*_{i,r-1}$ and E^*_{ir} will be nearly equal to E_{ir}. Hence, we obtain an approximation to the minimum of Q_i by taking the minimum of the i numbers $E^*_{io}, E^*_{i1}, \ldots, E^*_{i,i-1}$. The quantity E_{ir} can be computed as follows. Let g_r be the positive root in q of the equation

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$$\sum_{j=1}^{r+1} \frac{a_j}{\lambda_1^{\frac{j(j-1)}{2}} q^j} + \sum_{j=1}^{n-r-1} \frac{a_{r+1+j}}{\lambda_1^{\frac{r(r+1)}{2}+rj} \lambda_2^{\frac{j(j+1)}{2}} q^{r+1+j}} = 1 - a_0 \tag{66}$$

$$(r = 0,1,\ldots,i-1).$$

Then

$$E_{ir}^* = \lambda_1^{\frac{r(r+1)}{2}+r(i-r-1)} \lambda_2^{\frac{(i-r)(i-r-1)}{2}} g_r^i . \tag{67}$$

MINIMUM OF Q_n

Let $q_1^o,\ldots,q_n^o$ be values of $q_1,\ldots,q_n$ for which Q_n becomes a minimum. We shall prove that $q_{j+1}^o = \lambda_1 q_j^o$ $(j = 1,\ldots,n-1)$. Assume that there exists a value $j < n$ such that $q_{j+1}^o > \lambda_1 q_j^o$ and we shall derive a contradiction. Let u be the smallest integer such that $q_{u+1}^o > \lambda_1 q_u^o$ and let v be the largest integer such that $q_{v+1}^o > \lambda_1 q_v^o$. Let $q_u' = (1 + \epsilon)q_u^o$ $(\epsilon > 0)$, $q_{v+1}' = \frac{q_{v+1}^o}{1 + \epsilon}$, and $q_j' = q_j^o$ for $j \neq u, \neq v+1$. Then for sufficiently small ϵ we shall have $\lambda_1 q_r' \leq q_{r+1}' \leq \lambda_2 q_r'$ $(r = 1,\ldots,n-1)$. Furthermore, we have

$$q_1' \cdots q_n' = q_1^o \cdots q_n^o \quad \text{and} \quad \sum_{j=1}^{n} \frac{a_j}{q_1' \cdots q_j'} < 1 - a_0 .$$

Hence, there exists a positive $\Delta < 1$ such that $q_j'' = \frac{q_j'}{\Delta}$ $(j = 1,\ldots,n)$ and

$$\sum_{j=1}^{n} \frac{a_j}{q_1'' \cdots q_j''} = 1 - a_0.$$

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But then $q_1'' \ldots q_n'' < q_1^0 \ldots q_n^0$ in contradiction to the assumption that $q_1^0 \ldots q_n^0$ is a minimum. Hence, our statement is proved.

If q is the root of the equation

$$\sum_{j=1}^{n} \frac{a_j}{\lambda_1^{\frac{j(j-1)}{2}} q^j} = 1 - a_0 \; ,$$

then the minimum of Q_n is equal to $\lambda_1^{\frac{n(n-1)}{2}} q^n$.

MAXIMUM OF Q_i $(i < n)$

Let $q_1^*, \ldots, q_n^*$ be values of $q_1, \ldots, q_n$ for which Q_i becomes a maximum. We shall prove the following:

Lemma 3: The relations

$$q_{j+1}^* = \lambda_1 q_j^* \qquad (j = i, \ldots, n-1) \qquad\qquad (68)$$

must hold.

Proof: Assume that there exists an integer $j \geq i$ such that $q_{j+1}^* > \lambda_1 q_j^*$ and we shall derive a contradiction. Let $q_r' = q_r^*$ for $r = 1, \ldots, i$ and let $q_{j+1}' = \lambda_1 q_j'$ $(j = i, \ldots, n-1)$. Then

$$q_1' \ldots q_i' = q_1^* \ldots q_i^* \quad \text{and} \quad \sum_{j=1}^{n} \frac{a_j}{q_1' \ldots q_j'} > 1 - a_0 \; .$$

Hence, there exists a value $\Delta > 1$ such that

$$\sum_{j=1}^{n} \frac{a_j}{q_1'' \ldots q_j''} = 1 - a_0 \, ,$$

where $q_j'' = \Delta q_j'$ $(j = 1, \ldots, n)$. But then $q_1'' \ldots q_i'' > q_1^* \ldots q_i^*$ in contradiction to the assumption that $q_1^* \ldots q_i^*$ is a maximum. Hence, Lemma 3 is proved.

Lemma 4: If for some $j < i$ we have $q^*_{j+1} > \lambda_1 q^*_j$, then $q^*_{k+1} = \lambda_2 q^*_k$ for $k = 1, \ldots, j-1$.

Proof: Assume that $q^*_{j+1} > \lambda_1 q^*_j$ for some $j < i$ and that there exists an integer $k \leq j-1$ such that $q^*_{k+1} < \lambda_2 q^*_k$. We shall derive a contradiction from this assumption. Let u be the smallest integer such that $q^*_{u+1} < \lambda_2 q^*_u$. Furthermore, let v be the smallest integer greater than or equal to $u + 1$ such that $q^*_{v+1} > \lambda_1 q^*_v$. It is clear that $v \leq j$. Let $q'_u = \dfrac{q^*_u}{1 + \varepsilon}$ $(\varepsilon > 0)$, $q'_v = (1 + \varepsilon) q^*_v$, and $q'_r = q^*_r$ for $r \neq u, \neq v$. Then for sufficiently small ε we have

$$\lambda_1 q'_j \leq q'_{j+1} \leq \lambda_2 q'_j \qquad (j = 1, \ldots, n-1) .$$

Furthermore, we have

$$q'_1 \cdots q'_i = q^*_1 \cdots q^*_i \quad \text{and} \quad \sum_{j=1}^{n} \frac{a_j}{q'_1 \cdots q'_j} > 1 - a_o .$$

Hence, there exists a value $\Delta > 1$ such that

$$\sum_{j=1}^{n} \frac{a_j}{q''_1 \cdots q''_j} = 1 - a_o ,$$

where $q''_j = \Delta q'_j$ $(j = i, \ldots, n)$. But then $q''_1 \cdots q''_i > q^*_1 \cdots q^*_i$ in contradiction to the assumption that $q^*_1 \cdots q^*_i$ is a maximum.

Let D_{ir} $(r = 1, \ldots, i-1)$ be the maximum of Q_i under the restriction that $q_{j+1} = \lambda_1 q_j$ for $j = r+1, \ldots, n-1$ and $q_{j+1} = \lambda_2 q_j$ for $j = 1, \ldots, r-1$. From Lemma 3 and 4 it follows that the maximum of Q_i is equal to the maximum of the $i - 1$ values $D_{i1}, \ldots, D_{i,i-1}$. The computation of the exact value of D_{ir} can be carried out in a way similar to the computation of M_{ir} in part II. Since these computations are involved if n is large, we shall discuss here only an approximation method.

Let D^*_{ir} ($r = 1,\ldots,i-1$) be the value of Q_i if $q_{j+1} = \lambda_1 q_j$ for $j = r+1,\ldots,n-1$ and $q_{j+1} = \lambda_2 q_j$ for $j = 1,\ldots,r$. Furthermore, let D^*_{io} be the value of Q_i if $q_{j+1} = \lambda_1 q_j$ ($j = 1,\ldots,n-1$). Then, if λ_1 is not much below one, the maximum of D^*_{ir} and $D^*_{i,r-1}$ ($r = 1,\ldots,i-1$) will be nearly equal to D_{ir}. Hence, we obtain an approximation to the maximum value of Q_i by taking the largest of the i values $D^*_{io},\ldots,D^*_{i,i-1}$.

The value of D^*_{ir} can be determined as follows. Let g_r be the root in q of the equation

$$\sum_{j=1}^{r+1} \frac{a_j}{\lambda_2^{\frac{j(j-1)}{2}} q^j} + \sum_{j=1}^{n-r-1} \frac{a_{r+1+j}}{\lambda_2^{\frac{r(r+1)}{2}+jr} \lambda_1^{\frac{j(j+1)}{2}} q^{r+1+j}} = 1 - a_o.$$

Then

$$D^*_{ir} = \lambda_2^{\frac{r(r+1)}{2}+(i-r-1)} \lambda_1^{\frac{(i-r-1)(i-r)}{2}} g_r^i.$$

MAXIMUM OF Q_n

We shall prove that the maximum of Q_n is reached when $q_{j+1} = \lambda_2 q_j$ ($j = 1,\ldots,n-1$). Denote by $q^*_1 \ldots q^*_n$ the values of $q_1 \ldots q_n$ for which Q_n becomes a maximum. We shall assume that there exists a value $j < n$ such that $q^*_{j+1} < \lambda_2 q^*_j$ and we shall derive a contradiction from this assumption. Let u be the smallest and v be the largest integer such that $q^*_{u+1} < \lambda_2 q^*_u$ and $q^*_{v+1} < \lambda_2 q^*_v$.

Let $q'_u = \frac{q^*_u}{1+\epsilon}$ ($\epsilon > 0$), $q'_{v+1} = (1+\epsilon) q^*_{v+1}$, and $q'_r = q^*_r$ for $r \neq u, \neq v+1$. Then for sufficiently small ϵ we shall have $\lambda_1 q'_r \leq q'_{r+1} \leq \lambda_2 q'_r$ ($r = 1,\ldots,n-1$).

Furthermore, we have

$$q_1' \ldots q_n' = q_1^* \ldots q_n^* \quad \text{and} \quad \sum_{j=1}^{n} \frac{q_j}{q_1' \ldots q_j'} > 1 - a_o \ .$$

Hence, there exists a value $\Delta > 1$ such that $q_j'' = \Delta q_j'$
$(j = 1, \ldots, n)$ and

$$\sum_{j=1}^{n} \frac{a_j}{q_1'' \ldots q_j''} = 1 - a_o \ .$$

But then $q_1'' \ldots q_n'' > q_1^* \ldots q_n^*$ in contradiction to the assumption
that $q_1^* \ldots q_n^*$ is a maximum. Hence, our statement is proved.

The maximum of Q_n is equal to

$$\lambda_2^{\frac{n(n-1)}{2}} q^n \ ,$$

where q is the root of the equation

$$\sum_{j=1}^{n} \frac{a_j}{\lambda_2^{\frac{j(j-1)}{2}} q^j} = 1 - a_o \ .$$

NUMERICAL EXAMPLE

The same notation will be used as in the previous numerical
examples. The assumption of no sampling error, which is common
to all the previous examples, is retained. In part I it was
assumed that the q_i, the probability of a plane surviving the
i-th hit, knowing that the first $i - 1$ hits did not down the
plane, were equal for all i $(q_1 = q_2 = \ldots = q_n = q_o$ (say)).
Under this assumption, the exact value of the probability of a
plane surviving i hits is given by

$$Q_i = q_o^i \ .$$

In part III it was assumed that $q_1 \geq q_2 \geq \ldots \geq q_n$. Since no
lower limit is assumed in the decrease from q_i to q_{i+1}, only a

lower bound to the Q_i could be obtained. The assumption here is that the decrease from q_i to q_{i+1} lies between definite limits. Therefore, both an upper and lower bound for the Q_i can be obtained.

We assume that

$$\lambda_1 q_i \leq q_{i+1} \leq \lambda_2 q_i \ ,$$

where $\lambda_1 < \lambda_2 < 1$ and such that the expression

$$\sum_{j=1}^{n} \frac{a_j}{\lambda_1^{j(j-1)} } < 1 - a_o \qquad \text{(A)}$$

is satisfied.

The exact solution is tedious but close approximations to the upper and lower bounds to the Q_i for $i < n$ can be obtained by the following procedure. The set of hypothetical data used is

$$a_o = .780 \qquad a_3 = .010$$
$$a_1 = .070 \qquad a_4 = .005$$
$$a_2 = .040 \qquad a_5 = .005$$
$$\lambda_1 = .80 \qquad \lambda_2 = .90$$

Condition A is satisfied, since by substitution

$$.07 + \frac{.04}{.8} + \frac{.01}{(.8)^3} + \frac{.005}{(.8)^6} + \frac{.005}{(.8)^{10}} = .20529 \ ,$$

which is less than

$$1 - a_o = .22 \ .$$

THE LOWER LIMIT OF Q_i

The first step is to solve equation 66. This involves the solution of the following four equations for positive roots g_o, g_1, g_2, g_3.

$$\frac{a_1}{q} + \frac{a_2}{\lambda_2 q^2} + \frac{a_3}{\lambda_2^3 q^3} + \frac{a_4}{\lambda_2^6 q^4} + \frac{a_5}{\lambda_2^{10} q^5} = 1 - a_0 = .22 \qquad (B)$$

$$\frac{.07}{q} + \frac{.04}{.9q^2} + \frac{.01}{.729q^3} + \frac{.005}{.531441q^4} + \frac{.005}{.348678q^5} = .22$$

$$.22q^5 - .07q^4 - .044444q^3 - .013717q^2 - .009408q - .014340 = 0$$

$$g_0 = .844.$$

$$\frac{a_1}{q} + \frac{a_2}{\lambda_1 q^2} + \frac{a_3}{\lambda_1^2 \lambda_2 q^3} + \frac{a_4}{\lambda_1^3 \lambda_2^3 q^4} + \frac{a_5}{\lambda_1^4 \lambda_2^6 q^5} = 1 - a_0 \qquad (C)$$

$$\frac{.07}{q} + \frac{.04}{.8q^2} + \frac{.01}{(.64)(.9)q^3} + \frac{.005}{(.512)(.729)q^4} + \frac{.005}{(.4096)(.531441)q^5} = .22$$

$$.22q^5 - .07q^4 - .05q^3 - .017361q^2 - .013396q - .022970 = 0$$

$$g_1 = .904.$$

$$\frac{a_1}{q} + \frac{a_2}{\lambda_1 q^2} + \frac{a_3}{\lambda_1^3 q^3} + \frac{a_4}{\lambda_1^5 \lambda_2 q^4} + \frac{a_5}{\lambda_1^7 \lambda_2^3 q^5} = 1 - a_0 \qquad (D)$$

$$\frac{.07}{q} + \frac{.04}{.8q^2} + \frac{.01}{.512q^3} + \frac{.005}{(.32768)(.9)q^4} + \frac{.005}{(.209715)(.729)q^5} = .22$$

$$.22q^5 - .07q^4 - .05q^3 - .019531q^2 - .016954q - .032705 = 0$$

$$g_2 = .941.$$

$$\frac{a_1}{q} + \frac{a_2}{\lambda_1 q^2} + \frac{a_3}{\lambda_1^3 q^3} + \frac{a_4}{\lambda_1^6 q^4} + \frac{a_5}{\lambda_1^9 \lambda_2 q^5} = 1 - a_0 \qquad (E)$$

$$\frac{.07}{q} + \frac{.04}{.8q^2} + \frac{.01}{.512q^3} + \frac{.005}{.262144q^4} + \frac{.005}{(.134218)(.9)q^5} = .22$$

$$.22q^5 - .07q^4 - .05q^3 - .019531q^2 - .019073q - .041392 = 0$$

$$g_3 = .964 \; .$$

Next, calculate the i numbers defined by

$$E_{ir}^* = \lambda_1^{a(i,r)} \, \lambda_2^{b(i,r)} \, g_r^i \qquad (r = 0,1,\ldots,i-1),$$

where

$$a(i,r) = \frac{r(r+1)}{2} + r(i - r - 1)$$

$$b(i,r) = \frac{(i-r)(i-r-1)}{2}$$

$$g_0 = .844$$
$$g_1 = .904$$
$$g_2 = .941$$
$$g_3 = .964$$

The minimum of the E_{ir}^* ($r = 0,\ldots,i-1$) will be the lower limit of Q_i. The computations are given in table 5.

TABLE 5

COMPUTATION OF LOWER LIMIT OF Q_i

Q_i	i	r	a(i,r)	b(i,r)	g_r	g_r^i	E_{ir}^*
Q_1	1	0	0	0	.844	.844	.844

Min $[E_{10}^*]$ = .844

Q_i	i	r	a(i,r)	b(i,r)	g_r	g_r^i	E_{ir}^*
Q_2	2	0	0	1	.844	.712	.641
	2	1	1	0	.904	.817	.654

Min $[E_{20}^*,\ E_{21}^*]$ = .641

Q_i	i	r	a(i,r)	b(i,r)	g_r	g_r^i	E_{ir}^*
Q_3	3	0	0	3	.844	.601	.438
	3	1	2	1	.904	.739	.426
	3	2	3	0	.941	.833	.427

Min $[E_{30}^*,\ E_{31}^*,\ E_{32}^*]$ = .426

Q_i	i	r	a(i,r)	b(i,r)	g_r	g_r^i	E_{ir}^*
Q_4	4	0	0	6	.844	.507	.270
	4	1	3	3	.904	.668	.249
	4	2	5	1	.941	.784	.231
	4	3	6	0	.964	.864	.226

Min $[E_{40}^*,\ E_{41}^*,\ E_{42}^*,\ E_{43}^*]$ = .226

The lower limit of Q_5 can be obtained directly. The lower limit of

$$Q_5 = \lambda_1^{10} q^5 \; ,$$

where q is the positive root of

$$\frac{a_1}{q} + \frac{a_2}{\lambda_1 q^2} + \frac{a_3}{\lambda_1^3 q^3} + \frac{a_4}{\lambda_1^6 q^4} + \frac{a_5}{\lambda_1^{10} q^5} = 1 - a_0$$

$$\frac{.07}{q} + \frac{.04}{.8q^2} + \frac{.01}{.512q^3} + \frac{.005}{.262144q^4} + \frac{.005}{.107374q^5} = .22$$

$$q = .974 \; .$$

The lower limit of

$$Q_5 = (.8)^{10}(.974)^5 = .094 \; .$$

THE UPPER LIMIT OF Q_i

The computations for the upper limit of Q_1 are entirely analogous to the computations of the lower limit. First, we solve the equations of part IV, which for this example are the following:

$$\frac{a_1}{q} + \frac{a_2}{\lambda_1 q^2} + \frac{a_3}{\lambda_1^3 q^3} + \frac{a_4}{\lambda_1^6 q^4} + \frac{a_5}{\lambda_1^{10} q^5} = 1 - a_0$$

$$\frac{.07}{q} + \frac{.04}{.8q^2} + \frac{.01}{.512q^3} + \frac{.005}{.262144q^4} + \frac{.005}{.107374q^5} = .22$$

$$.22q^5 - .07q^4 - .05q^3 - .019531q^2 - .019073q - .046566 = 0$$

$$q_0^* = .974$$

$$\frac{a_1}{q} + \frac{a_2}{\lambda_2 q^2} + \frac{a_3}{\lambda_2^2 \lambda_1 q^3} + \frac{a_4}{\lambda_2^3 \lambda_1^3 q^4} + \frac{a_5}{\lambda_2^4 \lambda_1^6 q^5} = 1 - a_o$$

$$\frac{.07}{q} + \frac{.04}{.9q^2} + \frac{.01}{(.81)(.8)q^3} + \frac{.005}{(.729)(.512)q^4} + \frac{.005}{(.6561)(.262144)q^5} = .22$$

$$.22q^5 - .07q^4 - .044444q^3 - .015432q^2 - .013396q - .029071 = 0$$

$$g_1^* = .905$$

$$\frac{a_1}{q} + \frac{a_2}{\lambda_2 q^2} + \frac{a_3}{\lambda_2^3 q^3} + \frac{a_4}{\lambda_2^5 \lambda_1 q^4} + \frac{a_5}{\lambda_2^3 \lambda_1^7 q^5} = 1 - a_o$$

$$\frac{.07}{q} + \frac{.04}{.9q^2} + \frac{.01}{.729q^3} + \frac{.005}{(.59049)(.8)q^4} + \frac{.005}{(.512)(.478297)q^5} = .22$$

$$.22q^5 - .07q^4 - .044444q^3 - .013717q^2 - .010584q - .020417 = 0$$

$$g_2^* = .869$$

$$\frac{a_1}{q} + \frac{a_2}{\lambda_2 q^2} + \frac{a_3}{\lambda_2^3 q^3} + \frac{a_4}{\lambda_2^6 q^4} + \frac{a_5}{\lambda_2^9 \lambda_1 q^5} = 1 - a_o$$

$$\frac{.07}{q} + \frac{.04}{.9q^2} + \frac{.01}{.729q^3} + \frac{.005}{.531441q^4} + \frac{.005}{(.387420)(.8)q^5} = .22$$

$$.22q^5 - .07q^4 - .044444q^3 - .013717q^2 - .009408q - .016132 = 0$$

$$g_3^* = .851$$

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Next, calculate the i numbers defined by

$$D^*_{ir} = \lambda_2^{a(i,r)} \lambda_1^{b(i,r)} g_r^{*i} \qquad (r = 0,1,\ldots,i-1),$$

where

$$a(i,r) = \frac{r(r+1)}{2} + r(i - r - 1)$$

$$b(i,r) = \frac{(i-r)(i-r-1)}{2}$$

$$g_0^* = .974$$
$$g_1^* = .905$$
$$g_2^* = .869$$
$$g_3^* = .851$$

The maximum of the D^*_{ir} $(r = 0,\ldots,i-1)$ will be the upper limit of Q_i. The computations are given in table 6.

The upper limit of Q_5 can be obtained directly. The limit of

$$Q_5 = \lambda_2^{10} q^{*5},$$

where q^* is the positive root of

$$\frac{a_1}{q} + \frac{a_2}{\lambda_2 q^2} + \frac{a_3}{\lambda_2^3 q^3} + \frac{a_4}{\lambda_2^6 q^4} + \frac{a_5}{\lambda_2^{10} q^5} = 1 - a_0$$

$$\frac{.07}{q} + \frac{.04}{.9q^2} + \frac{.01}{.729q^3} + \frac{.005}{.531441q^4} + \frac{.005}{.348678q^5} = .22$$

$$q^* = .844.$$

TABLE 6

COMPUTATION OF UPPER LIMIT OF Q_i

Q_i	i	r	a(i,r)	b(i,r)	g_r^*	g_r^{*i}	D_{ir}^*
Q_1	1	0	0	0	.974	.974	.974

Max $[D_{10}^*]$ = .974

	i	r	a(i,r)	b(i,r)	g_r^*	g_r^{*i}	D_{ir}^*
Q_2	2	0	0	1	.974	.949	.759
	2	1	1	0	.905	.819	.737

Max $[D_{20}^*, D_{21}^*]$ = .759

	i	r	a(i,r)	b(i,r)	g_r^*	g_r^{*i}	D_{ir}^*
Q_3	3	0	0	3	.974	.924	.473
	3	1	2	1	.905	.741	.480
	3	2	3	0	.869	.656	.478

Max $[D_{30}^*, D_{31}^*, D_{32}^*]$ = .480

	i	r	a(i,r)	b(i,r)	g_r^*	g_r^{*i}	D_{ir}^*
Q_4	4	0	0	6	.974	.890	.236
	4	1	3	3	.905	.671	.250
	4	2	5	1	.869	.570	.269
	4	3	6	0	.851	.524	.279

Max $[D_{40}^*, D_{41}^*, D_{42}^*, D_{43}^*]$ = .279

The upper limit of

$$Q_5 = (.9)^{10}(.844)^5 = .149$$

Summarizing the results, the upper and lower limits of the probability of a plane surviving i hits are given by

$$.844 < Q_1 < .974$$
$$.641 < Q_2 < .759$$
$$.426 < Q_3 < .480$$
$$.226 < Q_4 < .279$$
$$.094 < Q_5 < .149$$

SUBDIVISION OF THE PLANE INTO SEVERAL EQUI-VULNERABILITY AREAS[1]

In parts I through IV we have considered the probability that a plane will be downed by a hit without any reference to the part of the plane that receives the hit. Undoubtedly, the probability of downing a plane by a hit will depend considerably upon the part that receives the hit. The purpose of this memorandum is to extend the previous results to the more general case where the probability of downing a plane by a hit depends on the part of the plane sustaining the hit. To carry out this generalization of the theory, we shall subdivide the plane into k equi-vulnerability areas $A_1, \ldots, A_k$. For any set of non-negative integers $i_1, \ldots, i_k$ let $P(i_1, \ldots, i_k)$ be the probability that a plane will be downed if the area A_1 receives i_1 hits, the area A_2 receives i_2 hits,..., and the area A_k receives i_k hits. Let $Q(i_1, \ldots, i_k) = 1 - P(i_1, \ldots, i_k)$. Then $Q(i_1, \ldots, i_k)$ is the probability that the plane will not be downed if the areas $A_1, \ldots, A_k$ receive $i_1, \ldots, i_k$ hits, respectively. We shall assume that $Q(i_1, \ldots, i_k)$ is a symmetric function of the arguments $i_1, \ldots, i_k$.

To estimate the value of $Q(i_1, \ldots, i_k)$ from the damage to returning planes, we need to know the probability distribution of hits over the k areas $A_1, \ldots, A_k$ knowing merely the total number of hits received. In other words, for any positive integer i we need to know the conditional probability $\gamma_i(i_1, \ldots, i_k)$ that the areas $A_1, \ldots, A_k$ will receive $i_1, \ldots, i_k$ hits, respectively, knowing that the total number of hits is i. Of course, $\gamma_i(i_1, \ldots, i_k)$ is defined only for values $i_1, \ldots, i_k$ for which $i_1 + \ldots + i_k = i$. To avoid confusion, it should be emphasized that the probability $\gamma_i(i_1, \ldots, i_k)$ is determined under the

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[1]This part of "A Method of Estimating Plane Vulnerability Based on Damage of Survivors" was published as SRG memo 96 and AMP memo 76.5.

assumption that dummy bullets are used. It can easily be shown that it is impossible to estimate both $\gamma_i(i_1,\ldots,i_k)$ and $Q(i_1,\ldots,i_k)$ from the damage to returning planes only. To see this, assume that k is equal to 2 and all hits on the returning planes were located in the area A_1. This fact could be explained in two different ways. One explanation could be that $\gamma_i(i_1,i_2) = 0$ for $i_2 > 0$. The other possible explanation would be that $Q(i_1,i_2) = 0$ for $i_2 > 0$. Hence, it is impossible to estimate both $\gamma_i(i_1,i_2)$ and $Q(i_1,i_2)$. Fortunately, $\gamma_i(i_1,\ldots,i_k)$ can be assumed to be known a priori (on the basis of the dispersion of the guns), or can be established experimentally by firing with dummy bullets and recording the hits scored. Thus, in what follows we shall assume that $\gamma_i(i_1,\ldots,i_k)$ is known for any set of integers $i_1,\ldots,i_k$.

Clearly, the probability that i hits will not down the plane is given by

$$Q_i = \sum_{i_k} \cdots \sum_{i_1} \gamma_i(i_1,\ldots,i_k)Q(i_1,\ldots,i_k), \qquad (69)$$

where the summation is to be taken over all non-negative integers $i_1,\ldots,i_k$ for which $i_1 + \ldots + i_k = i$.

Let $\delta_i(i_1,\ldots,i_k)$ be the conditional probability that the areas $A_1,\ldots,A_k$ received $i_1,\ldots,i_k$ hits, respectively, knowing that the plane received i hits and that the plane was not downed. Then we have

$$\delta_i(i_1,\ldots,i_k) = \frac{\gamma_i(i_1,\ldots,i_k)Q(i_1,\ldots,i_k)}{Q_i} . \qquad (70)$$

Of course, $\delta_i(i_1,\ldots,i_k)$ is defined only for non-negative integers $i_1,\ldots,i_k$ for which $i_1 + \ldots + i_k = i$.

The probability $\delta_i(i_1, \ldots, i_k)$ can be determined from the distribution of hits on returning planes. In fact, let $a(i_1, \ldots, i_k)$ be the proportion of planes (out of the total number of planes participating in combat) that returned with i_1 hits on area A_1, i_2 hits on area $A_2, \ldots,$ and i_k hits on area A_k. Then we obviously have

$$\delta_i(i_1, \ldots, i_k) = \frac{a(i_1, \ldots, i_k)}{a_i} . \tag{71}$$

From equations 70 and 71, we obtain

$$Q(i_1, \ldots, i_k) = \frac{Q_i a(i_1, \ldots, i_k)}{a_i \gamma_i(i_1, \ldots, i_k)} \qquad (i = i_1 + \ldots + i_k) \tag{72}$$

Since Q_i can be estimated by methods described in parts I through IV, estimates of $Q(i_1, \ldots, i_k)$ can be obtained from equation 72.

According to equation 29, the probabilities $Q_1, \ldots, Q_n$ satisfy the equation

$$\sum_{j=1}^{n} \frac{a_j}{Q_j} = 1 - a_o . \tag{73}$$

We have assumed that $q_1 \geq q_2 \geq \cdots \geq q_n$. This is equivalent to stating that

$$\frac{Q_{i+1}}{Q_i} \leq \frac{Q_{j+1}}{Q_j} \qquad \text{for } j \leq i . \tag{74}$$

A similar assumption can be made with respect to the probabilities $Q(i_1, \ldots, i_k)$. In fact, the conditional probability that an additional hit on the area A_r will not down the plane knowing that the areas $A_1, \ldots, A_k$ have already sustained $i_1, \ldots, i_k$ hits, respectively, is given by

$$\frac{Q(i_1,\ldots,i_{r-1},\ i_r+1,\ i_{r+1},\ldots,i_k)}{Q(i_1,\ldots,i_{r-1},\ i_r,\ i_{r+1},\ldots,i_k)} \quad . \tag{75}$$

Obviously, we can assume that if

$$j_1 \le i_1,\ j_2 \le i_2,\ldots,j_k \le i_k$$

then

$$\frac{Q(i_1,\ldots,i_{r-1},i_r+1,i_{r+1},\ldots,i_k)}{Q(i_1,\ldots,i_{r-1},i_r,i_{r+1},\ldots,i_k)} \le \frac{Q(j_1,\ldots,j_{r-1},j_r+1,j_{r+1},\ldots,j_k)}{Q(j_1,\ldots,j_{r-1},j_r,j_{r+1},\ldots,j_k)}$$

$$\tag{76}$$

for $r = 1,2,\ldots,k$.

Hence, the possible values of $Q_1,\ldots,Q_n$ are restricted to those for which equation 73 is fulfilled and for which the quantities $Q(i_1,\ldots,i_k)$ computed from equation 72 are less than or equal to one and satisfy the inequalities of equation 76. It should be remarked that the inequalities of equation 76 do not follow from the inequalities of equation 74. From equation 72 and the inequality $Q(i_1,\ldots,i_k) \le 1$, it follows that

$$Q_i \le \frac{a_i \gamma_i(i_1,\ldots,i_k)}{a(i_1,\ldots,i_k)} \quad . \tag{77}$$

If the right-hand side expression in equation 77 happens to be less than one, then equation 77 imposes a restriction on Q_i. Since

$$\sum_{i_k} \cdots \sum_{i_1} \frac{a(i_1,\ldots,i_k)}{a_i} = \sum_{i_k} \cdots \sum_{i_1} \gamma_i(i_1,\ldots,i_k) = 1$$

(the summation is taken over all values $i_1,\ldots,i_k$ for which $i_1 + \ldots + i_k = i$), we must have either

$$\frac{a_i \, \gamma_i(i_1,\ldots,i_k)}{a(i_1,\ldots,i_k)} = 1$$

for all values $i_1,\ldots,i_k$ for which $i_1 + \ldots + i_k = i$, or

$$\frac{a_i \, \gamma_i(i_1,\ldots,i_k)}{a(i_1,\ldots,i_k)} < 1$$

at least for one set of values $i_1,\ldots,i_k$ satisfying the condition $i_1 + \ldots + i_k = i$. Hence, equation 77 gives an upper bound for Q_i whenever there exists a set of integers $i_1,\ldots,i_k$ such that $i_1 + \ldots + i_k = i$ and

$$\frac{a(i_1,\ldots,i_k)}{a_i} \neq \gamma_i \; (i_1,\ldots,i_k) \; .$$

It is of interest to investigate the case of independence, i.e., the case when the probability that an additional hit will not down the plane does not depend on the number and distribution of hits already received. Denote by $q(i)$ the probability that a single hit on the area A_i will not down the plane. Then under the assumption of independence we have

$$Q(i_1,\ldots,i_k) = [q(1)]^{i_1} \, [q(2)]^{i_2} \, \ldots \, [q(k)]^{i_k} \; . \qquad (78)$$

Hence, the only unknown probabilities are $q(1),\ldots,q(k)$.

Let $\gamma(i)$ be the conditional probability that the area A_i is hit knowing that the plane received exactly one hit. Obviously

$$\gamma_i(i_1,\ldots,i_k) = \frac{i!}{i_1! \, \ldots \, i_k!} \, [\gamma(1)]^{i_1} \, \ldots \, [\gamma(k)]^{i_k} \; . \qquad (79)$$

Similarly, let $\delta(i)$ be the conditional probability that the area A_i is hit knowing that the plane received exactly one hit and this hit did not down the plane. Because of the assumption of independence, we have

$$\delta_i(i_1,\ldots,i_k) = \frac{i!}{i_1! \ldots i_k!} [\delta(1)]^{i_1} \ldots [\delta(k)]^{i_k} . \qquad (80)$$

Furthermore, we have

$$\delta(i) = \frac{\gamma(i)q(i)}{\sum\limits_{i=1}^{k} \gamma(i)q(i)} . \qquad (81)$$

Since the probability q that a single hit does not down the plane is equal to $\sum\limits_{i=1}^{k} \gamma(i)q(i)$, we obtain from equation 81

$$q(i) = \frac{\delta(i)}{\gamma(i)} q . \qquad (82)$$

Because of the assumption of independence, we see that $\delta(i)$ is equal to the ratio of the total number of hits in the area A_i of the returning planes to the total number of hits received by the returning planes. That is

$$\delta(i) = \frac{\sum\limits_{j_k} \cdots \sum\limits_{j_1} j_i a(j_1,\ldots,j_k)}{\sum\limits_{j_k} \cdots \sum\limits_{j_1} (j_1 + \ldots + j_k)a(j_1,\ldots,j_k)} . \qquad (83)$$

Since $\gamma(i)$ is assumed to be known and since $\delta(i)$ can be computed from equation 83, we see from equation 82 that $q(i)$ can be determined as soon as the value of q is known. The value of q can be obtained by solving the equation

$$\sum_{j=1}^{n} \frac{a_j}{q^j} = 1 - a_o . \qquad (84)$$

NUMERICAL EXAMPLE

In the examples for parts I, III, and IV we have estimated the
probability that a plane will be downed without reference to the
part of the plane that receives the hit. However, the vulner-
ability of a particular part (say the motors) may be of interest
and this example illustrates the methods of estimating part
vulnerabilities under the following assumptions:

- The number of planes participating in combat is large so
that sampling errors can be neglected.

- The probability that a hit will down the plane does not
depend on the number of previous non-destructive hits. That
is, $q_1 = q_2 = \ldots = q_n = q_0$.

- Given that a shot has hit the plane, the probability that
it hit a particular part is assumed to be known. In this
example it is put equal to the ratio of the area of this
part to the total surface area of the plane.[1]

- The division of the plane into several parts is repre-
sentative of all the planes of the mission. If the types of
planes are radically different so that no representative
division is possible, we may consider the different classes
of planes separately.

Consider the following example. Of 400 planes on a bombing
mission, 359 return. Of these, 240 were not hit, 68 had one hit,
29 had two hits, 12 had three hits, and 10 had four hits.
Following the example in part I we have

$$N = 400,$$

whence

$$A_0 = 240 \qquad a_0 = .600$$
$$A_1 = 68 \qquad a_1 = .170$$
$$A_2 = 29 \qquad a_2 = .072$$
$$A_3 = 12 \qquad a_3 = .030$$
$$A_4 = 10 \qquad a_4 = .025$$

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[1]By area is meant here the component of the area perpendic-
ular to the direction of the enemy attack. If this direction
varies during the combat, some proper average direction may be
taken.

As before, the probability that a single hit will not down the plane is given by the root of

$$\frac{a_1}{q_o} + \frac{a_2}{q_o^2} + \frac{a_3}{q_o^3} + \frac{a_4}{q_o^4} = 1 - a_o ,$$

which reduces to

$$.4q_o^4 - .170q_o^3 - .072q_o^2 - .030q_o - .025 = 0$$

and

$$q_o = .850 .$$

Suppose that we are interested in estimating the vulnerability of the engines, the fuselage, and the fuel system. Assume that the following data is representative of all the planes of the mission:

Part number	Description	Area of part	Ratio of area of part to total area ($\gamma(i)$)
1	2 engines	35 sq. ft.	$\frac{35}{130}$ = .269
2	Fuselage	45 sq. ft.	$\frac{45}{130}$ = .346
3	Fuel system	20 sq. ft.	$\frac{20}{130}$ = .154
4	All other parts	30 sq. ft.	$\frac{30}{130}$ = .231
	Total area	130 sq. ft.	

The ratio of the area of the i-th part to the total area is designated $\gamma(i)$. Given that the plane is hit, by the third assumption, $\gamma(i)$ is the probability that this hit occurred on part i. Thus

$$\gamma(1) = .269$$
$$\gamma(2) = .346$$
$$\gamma(3) = .154$$
$$\gamma(4) = .231$$

The only additional information we require is the number of hits on each part. Let the observed number of hits be 202. In general, the total number of hits (on returning planes) must be equal to

$$A_1 + 2A_2 + 3A_3 + \ldots + nA_n$$

and in this example

$$A_1 + 2A_2 + 3A_3 + 4A_4 = 68 + 2(29) + 3(12) + 4(10) = 202$$

The hits on the returning planes were distributed as follows:

Part number	Number of hits observed on part	Ratio of number of hits observed on part to total number of observed hits ($\delta(i)$)
1	39	.193
2	78	.386
3	31	.154
4	54	.267

Total number of hits   202

The ratio of the number of hits on part i to the total number of hits on surviving planes is designated $\delta(i)$. Then $q(i)$, the probability that a hit on the i-th part does not down the plane, is given by

$$q(i) = \frac{\delta(i)}{\gamma(i)} q_0 \quad ,$$

whence

$$q(1) = \frac{\delta(1)}{\gamma(1)} \, q_O = \frac{.193}{.269} \, (.850) = .61$$

$$q(2) = \frac{\delta(2)}{\gamma(2)} \, q_O = \frac{.386}{.346} \, (.850) = .95$$

$$q(3) = \frac{\delta(3)}{\gamma(3)} \, q_O = \frac{.154}{.154} \, (.850) = .85$$

$$q(4) = \frac{\delta(4)}{\gamma(4)} \, q_O = \frac{.267}{.231} \, (.850) = .98$$

The results may be summarized as follows:

Part	Probability of surviving a single hit ($q(i)$)	Probability of being downed by a single hit ($1 - q(i)$)
Entire plane	.85	.15
Engines	.61	.39
Fuselage	.95	.05
Fuel system	.85	.15
Other parts	.98	.02

Thus, for the observed data of this hypothetical example, the engine area is the most vulnerable in the sense that a hit there is most likely to down the plane. The fuselage has a relatively low vulnerability.

PART VI

SAMPLING ERRORS[1]

In parts I through V we have assumed that the total number of planes participating in combat is so large that sampling errors can be neglected altogether. However, in practice N is not excessively large and therefore it is desirable to take sampling errors into account. We shall deal here with the case when $q_1 = q_2 \ldots = q_n = q$ (say) and we shall derive confidence limits for the unknown probability q.

If there were no sampling errors, then we would have

$$(85)$$

$$x_i = p(1 - a_o - a_1 - \ldots - a_{i-1} - x_1 - x_2 - \ldots - x_{i-1})$$

$$(i = 2,3,\ldots),$$

where $p = 1 - q$. However, because of sampling errors we shall have the equation

$$x_i = \overline{p}_i(1 - a_o - \ldots - a_{i-1} - x_1 - \ldots - x_{i-1}), \qquad (86)$$

where $\overline{p}_i$ is distributed like the success ratio in a sequence of $N_i = N(1 - a_o - a_1 - \ldots - a_{i-1} - x_1 - \ldots - x_{i-1})$ independent trials, the probability of success in a single trial being equal to p.

Let $\overline{q}_i = 1 - \overline{p}_i$. Then, according to equation 26 we have

$$\sum_{j=1}^{n} \frac{a_j}{\overline{q}_1 \ldots \overline{q}_j} = 1 - a_o , \qquad (87)$$

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[1]This part of "A Method of Estimating Plane Vulnerability Based on Damage of Survivors" was published as SRG memo 103 and AMP memo 76.6.

provided that $x_i = 0$ for $i > n$. In part I we have shown that $x_i = 0$ for $i > n$ if there are no sampling errors. This is not necessarily true if sampling errors are taken into account. However, in the case of independence, i.e., when $q_i = q$ ($i = 1,2,\ldots$), x_i is very small for $i > n$ so that $\sum\limits_{i=n+1}^{\infty} x_i$ can be neglected.

In fact, if the number of planes that received more than n hits were not negligibly small, it follows from the assumption of independence that the probability is very high that at least some of these planes would return. Since no plane returned with more than n hits, for practical purposes we may assume that $\sum\limits_{i=n+1}^{\infty} x_i = 0$. In what follows we shall make this assumption.

Each of the quantities $\bar{q}_1,\ldots,\bar{q}_n$ can be considered as a sample estimate of the unknown probability q. However, the quantities $\bar{q}_1,\ldots,\bar{q}_n$ are unknown. It is merely known that they satisfy the relation in equation 87. Confidence limits for q may be derived on the basis of equation 87. However, we shall use another more direct approach.

To derive confidence limits for the unknown probability q we shall consider the hypothetical proportion b_i of planes that would have been hit exactly i times if dummy bullets would have been used. We shall treat the quantities $b_1,\ldots,b_k$ as fixed (but unknown) constants. This assumption does not involve any loss of generality, since the confidence limits for q obtained on the basis of this assumption remain valid also when $b_1,\ldots,b_k$ are random variables. Clearly, the probability distribution of Na_i ($i = 1,\ldots,n$) is the same as the distribution of the number of successes in a sequence of Nb_i independent trials, the probability of success in a single trial being q^i. Hence

$$E(Na_i) = q^i Nb_i \tag{88}$$

$$\sigma^2(Na_i) = Nb_i q^i(1 - q^i) . \tag{89}$$

From equations 88 and 89 we obtain

$$E\left(\frac{a_i}{q^i}\right) = b_i \tag{90}$$

$$\sigma^2\left(\frac{a_i}{q^i}\right) = \frac{b_i(1 - q^i)}{Nq^i} \quad . \tag{91}$$

Since the variates $\dfrac{a_1}{q}, \dfrac{a_2}{q^2}, \ldots, \dfrac{a_n}{q^n}$ are independently distributed, and since a_i is nearly normally distributed if N is not small, we can assume with very good approximation that the sum

$$\sum_{i=1}^{n} \frac{a_i}{q^i} \tag{92}$$

is normally distributed. We obtain from equations 90 and 91

$$E\left(\sum_{i=1}^{n} \frac{a_i}{q^i}\right) = \sum_{i=1}^{n} b_i = 1 - a_o \tag{93}$$

$$\sigma^2\left(\sum_{i=1}^{n} \frac{a_i}{q^i}\right) = \sum_{i=1}^{n} \frac{b_i(1 - q^i)}{Nq^i} \quad . \tag{94}$$

For any positive $\alpha < 1$ let λ_α be the value for which

$$\int_{-\lambda_\alpha}^{\lambda_\alpha} \frac{1}{\sqrt{2\pi}}\, e^{-\frac{t^2}{2}}\, dt = \alpha$$

The set of all values q for which the inequality

$$1 - a_o - \lambda_\alpha \sqrt{\sum_{i=1}^{n} \frac{b_i(1 - q^i)}{Nq^i}} \le \sum_{i=1}^{n} \frac{a_i}{q^i} \le 1 - a_o + \lambda_\alpha \sqrt{\sum_{i=1}^{n} \frac{b_i(1 - q^i)}{Nq^i}} \tag{95}$$

is fulfilled forms a confidence set for the unknown probability q with confidence coefficient α. However, formula 95 cannot be used, since it involves the unknown quantities $b, \ldots, b_n$. Since $\frac{a_i}{q^i}$ converges stochastically to b_i as $N \to \infty$, we change the standard deviation of $\sum \frac{a_i}{q^i}$ only by a quantity of order less than $\frac{1}{\sqrt{N}}$ if we replace b_i by $\frac{a_i}{q^i}$. Thus, the set of values q that satisfy the inequalities

$$1 - a_o - \lambda_\alpha \sqrt{\sum_{i=1}^{n} \frac{a_i(1 - q^i)}{Nq^{2i}}} \leq \sum_{i=1}^{n} \frac{a_i}{q^i} \leq 1 - a_o + \lambda_\alpha \sqrt{\sum_{i=1}^{n} \frac{a_i(1 - q^i)}{Nq^{2i}}} \qquad (96)$$

is an approximation to a confidence set with confidence coefficient α.

Denote by q_o the root of the equation in q

$$\sum_{1}^{n} \frac{a_j}{q^j} = 1 - a_o .$$

Then q_o converges stochastically to q as $N \to \infty$. A considerable simplification can be achieved in the computation of the confidence set by substituting q_o for q in the expression of the standard deviation of $\sum \frac{a_i}{q^i}$. The error introduced by this substitution is small if N is large. Making this substitution, the inequalities defining the confidence set are given by

$$1 - a_o - \lambda_\alpha \sqrt{\sum_{i=1}^{n} \frac{a_i(1 - q_o^i)}{Nq_o^{2i}}} \leq \sum_{i=1}^{n} \frac{a_i}{q^i} \leq 1 - a_o + \lambda_\alpha \sqrt{\sum_{i=1}^{n} \frac{a_i(1 - q_o^i)}{Nq_o^{2i}}} \qquad (97)$$

Hence, the confidence set is an interval. The upper end point of the confidence interval is the root of the equation

$$\sum_{i=1}^{n} \frac{a_i}{q^i} = 1 - a_o - \lambda_\alpha \sqrt{\sum_{i=1}^{n} \frac{a_i(1 - q_o^i)}{Nq_o^{2i}}} \tag{98}$$

and the lower end point of the confidence interval is the root of the equation

$$\sum_{i=1}^{n} \frac{a_i}{q^i} = 1 - a_o + \lambda_\alpha \sqrt{\sum_{i=1}^{n} \frac{a_i(1 - q_o^i)}{Nq_o^{2i}}} \tag{99}$$

NUMERICAL EXAMPLE

In all previous examples it was assumed that A_i (the number of planes returning with i hits) was compiled from such a large number of observations that they were not subject to sampling errors. If it is further assumed that the probability q that a hit will down a plane does not depend on the number of previous non-destructive hits, it is possible to obtain an exact solution for the probability that a hit will down a plane. Here we introduce the possibility that the $A_o,\ldots,A_n$ are subject to sampling errors but retain the assumption of independence. Under these less restrictive assumptions we cannot obtain the exact solution for q, but for any positive number $\alpha < 1$ we can construct two functions of the data, called confidence limits, such that the statement that q lies between the confidence limits will be true 100α percent of the time in the long run. The confidence limits are calculated for $\alpha = .95$ and $.99$.

Under the assumptions of part I, it was proved that no planes received more hits than the greatest number of hits observed on a returning plane. This is not necessarily true when the possibility of sampling error is introduced, but it is retained as an assumption, since the error involved is small.

If the a_i are subject to sampling error, and q is the true parameter,

$$\sum_{1}^{n} \frac{a_i}{q^i} \qquad (A)$$

will be approximately normally distributed with mean value $1 - a_o$.

In outlining the steps necessary to calculate the confidence limits, the following hypothetical set of data will be used. Given

$$N = 500 \qquad a_i = \frac{A_i}{N}$$

$$A_o = 400 \qquad a_o = .80$$
$$A_1 = 40 \qquad a_1 = .08$$
$$A_2 = 25 \qquad a_2 = .05$$
$$A_3 = 5 \qquad a_3 = .01$$
$$A_4 = 3 \qquad a_4 = .006$$
$$A_5 = \underline{2} \qquad a_5 = .004$$

$$475$$

The first step is to find the value q_o, for which expression A is equal to its mean value, by finding the positive root of

$$\frac{a_1}{q} + \frac{a_2}{q^2} + \frac{a_3}{q^3} + \frac{a_4}{q^4} + \frac{a_5}{q^5} = 1 - a_o .$$

We obtain

$$.20q^5 - .08q^4 - .05q^3 - .01q^2 - .006q - .004 = 0$$

$$q_o = .850 .$$

The next step is to calculate the standard deviation of expression A. This can be shown to be approximately equal to

$$\sigma = \sqrt{\sum_{i=1}^{n} \frac{a_i(1 - q_o^i)}{N q_o^{2i}}}$$

$$= \sqrt{\frac{a_1(1-q_o^1)}{N q_o^2} + \frac{a_2(1-q_o^2)}{N q_o^4} + \frac{a_3(1-q_o^3)}{N q_o^6} + \frac{a_4(1-q_o^4)}{N q_o^8} + \frac{a_5(1-q_o^5)}{N q_o^{10}}}$$

$$= .01226 .$$

Knowing that $\sum_{i=1}^{n} \frac{a_i}{q^i}$ is approximately normally distributed with mean value $1 - a_o$ and the standard deviation σ, we can determine the range in which $\sum \frac{a_i}{q^i}$ can be expected to be 100α percent of the time (say 95 and 99 percent) by determining $\lambda_{.95}$ and $\lambda_{.99}$ such that

$$\frac{1}{\sqrt{2\pi}} \int_{-\lambda_{.95}}^{\lambda_{.95}} \exp\left(-\frac{t^2}{2}\right) dt = .95$$

$$\frac{1}{\sqrt{2\pi}} \int_{-\lambda_{.99}}^{\lambda_{.99}} \exp\left(-\frac{t^2}{2}\right) dt = .99 .$$

From the table or the areas of a normal curve, it is found that

$$\lambda_{.95} = 1.959964$$
$$\lambda_{.99} = 2.575829 .$$

We can now calculate the confidence limits for each value of α by finding the two values of q for which the equality sign of the following expression holds:

$$\left| \sum_{i=1}^{n} \frac{a_i}{q^i} - (1 - a_o) \right| \leq \lambda_\alpha \sigma \quad .$$

It follows that for each α, the confidence limits are the positive roots of the equation

$$\sum_{i=1}^{n} \frac{a_i}{q^i} = 1 - a_o \pm \lambda_\alpha \sigma$$

α	λ_α	$.0122678\lambda_\alpha$	$1 - a_o - \lambda_\alpha \sigma$	$1 - a_o + \lambda_\alpha \sigma$
.95	1.959964	.024044	.175956	.224044
.99	2.575829	.031600	.168400	.231600

For $\alpha = .95$ the confidence limits of q_o are the positive roots of equation

$$\frac{a_1}{q} + \frac{a_2}{q^2} + \frac{a_3}{q^3} + \frac{a_4}{q^4} + \frac{a_5}{q^5} = .175956,$$

which reduces to

$$.175956q^5 - .08q^4 - .05q^3 - .01q^2 - .006q - .004 = 0$$

$$q = .912,$$

and equation

$$\frac{a_1}{q} + \frac{a_2}{q^2} + \frac{a_3}{q^3} + \frac{a_4}{q^4} + \frac{a_5}{q^5} = .224044,$$

which reduces to

$$.224044q^5 - .08q^4 - .05q^3 - .01q^2 - .006q - .004 = 0$$

$$q = .801.$$

Similarly, for $\alpha = .99$ we have

$$.168400q^5 - .08q^4 - .05q^3 - .01q^2 - .006q - .004 = 0$$

$$q = .935$$

$$.231600q^5 - .08q^4 - .05q^3 - .01q^2 - .006q - .004 = 0$$

$$q = .787 \ .$$

Summarizing the results we find that the 95-percent confidence limits of q are .801 and .912, and that the 99-percent confidence limits are .787 and .935.

PART VII

MISCELLANEOUS REMARKS[1]

1. <u>Factors that may vary from combat to combat but influence the</u>
<u>probability of surviving a hit.</u> The factors that influence the
probability of surviving a hit may be classified into two groups.
The first group contains those factors that do not vary from
combat to combat. This does not necessarily mean that the factor
in question has a fixed value of all combats; the factor may be a
random variable whose probability distribution does not vary from
combat to combat. The second group comprises those factors whose
probability distribution cannot be assumed to be the same for all
combats. To make predictions as to the proportions of planes
that will be downed in future combats, it is necessary to study
the dependence of the probability q of surviving a hit on the
factors in the second group. In part V we have already taken
into account such a factor. In part V we have considered a
subdivision of the plane into several equi-vulnerability areas
$A_1, \ldots, A_k$ and we expressed the probability of survival as a func-
tion of the part of the plane that received the hit. Since the
probability of hitting a certain part of the plane depends on the
angle of attack, this probability may vary from combat to combat.
Thus, it is desirable to study the dependence of the probability
of survival on the part of the plane that received the hit. In
addition to the factors represented by the different parts of the
plane, there may also be other factors, such as the type of gun
used by the enemy, etc., which belong to the second group. There
are no theoretical difficulties whatsoever in extending the
theory in part V to any number and type of factors. To
illustrate this, let us assume that the factors to be taken into
account are the different parts $A_1, \ldots, A_k$ of the plane and the
different guns $g_1, \ldots, g_m$ used by the enemy. Let $q(i,j)$ be the
probability of surviving a hit on part A_i knowing that the bullet
has been fired by gun g_j. We may order the km pairs (i,j) in a
sequence. We shall denote $q(i,j)$ by $q(u)$ if the pair (i,j) is
the u-th element in the ordered sequence of pairs. The problem
of determining the unknown probabilities $q(u)$ ($u = 1, \ldots, km$) can
be treated in exactly the same way as the problem discussed in

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[1]This part of "A Method of Estimating Plane Vulnerability
Based on Damage of Survivors" was published as SRG memo 109 and
AMP memo 76.7.

part V assuming that the plane consists of km parts. Any hit on part A_i by a bullet from gun g_j can be considered as a hit on part A_u in the problem discussed in part V where (i,j) is the u-th element in the ordered sequence of pairs.

2. Non-probabilistic interpretation of the results. It is interesting to note that a purely arithmetic interpretation of the results of parts I through V can be given. Instead of defining q_i as the probability of surviving the i-th hit knowing that the previous $i - 1$ hits did not down the plane, we define q_i as follows: Let M_i be the number of planes that received at least i hits and the i-th hit did not down the plane, and let N_i be the total number of planes that received at least i hits. Then $q_i = \dfrac{M_i}{N_i}$. Thus, q_i is defined in terms of what actually happened in the particular combat under consideration. To distinguish this definition of q_i from the probabilistic definition, we shall denote the ratio $\dfrac{M_i}{N_i}$ by $\overline{q}_i$. The quantity $\overline{q}$ is unknown, since we do not know the distribution of hits on the planes that did not return. However, it follows from the results of part I that these quantities must satisfy equation 26. If we can assume that in the particular combat under consideration we have $\overline{q}_i = \ldots = \overline{q}_n$ then the common value $\overline{q}$ of these quantities is the root of the equation

$$\sum \frac{a_j}{\overline{q}^j} = 1 - a_o \quad .$$

Assuming that $\overline{q}_1 \geq \overline{q}_2 \geq \ldots \geq \overline{q}_n$, the minimum value Q_i^o of Q_i derived in parts III and IV can be interpreted as the minimum value of $\overline{Q}_i = \overline{q}_1 \ldots \overline{q}_i$.

The minimum and maximum values of Q_i derived in part IV can also be interpreted as minimum and maximum values of $\overline{Q}_i = \overline{q}_1 \ldots \overline{q}_i$ if we assume that the inequalities $\lambda_1 \overline{q}_j \leq \overline{q}_{j+1} \leq \lambda_2 \overline{q}_j$ $(j = 1,\ldots,n-1)$ are fulfilled. Similarly, a pure arithmetic interpretation of the results of part V can be given.

3. <u>The case when $\gamma(i)$ is unknown.</u> In part V we have assumed that the probabilities $\gamma(1),\ldots,\gamma(k)$ are known. Since the exposed areas of the different parts $A_1,\ldots,A_k$ depend on the angle of attack, and since this angle may vary during the combat, it may sometimes be difficult to estimate the probabilities $\gamma(1),\ldots,\gamma(k)$. Thus, it may be of interest to investigate the question whether any inference as to the probabilities $q(1),\ldots,q(k)$ can be drawn when $\gamma(1),\ldots,\gamma(k)$ are entirely unknown. We shall see that frequently a useful lower bound for $q(i)$ can still be obtained. In fact, the value $q^*(i)$ of $q(i)$, calculated under the assumption that the parts $A_j (j \neq i)$ are not vulnerable $(q(j) = 1)$, is certainly a lower bound of the true value $q(i)$. Considering only the hits on part A_i, a lower bound of $q^*(i)$, and therefore also of $q(i)$, is given by the root of the equation

$$\sum_{r=1}^{n} \frac{a_r^*}{q^r} = 1 - a_o^* \quad , \tag{100}$$

where a_r^* $(r = 0,1,\ldots,n)$ is the ratio of the number of planes returned with exactly r hits on part A_i to the total number of planes participating in combat.

The lower limit obtained from equation 100 will be a useful one if it is not near zero. The root of equation 100 will be considerably above zero if $\sum_{r=1}^{n} a_r^*$ is not very small as compared with $1 - a_o^*$. This can be expected to happen whenever both $\gamma(i)$ and $q(i)$ are considerably above zero.

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PART VIII

VULNERABILITY OF A PLANE TO DIFFERENT TYPES OF GUNS[1]

In part V we discussed the case where the plane is subdivided into several equi-vulnerability areas (parts) and we dealt with the problem of determining the vulnerability of each of these parts. It was pointed out in part VII that the method described in part V can be applied to the more general problem of estimating the probability $q(i,j)$ that a plane will survive a hit on part i caused by a bullet fired from gun j. However, this method is based on the assumption that the value of $\gamma(i,j)$ is known where $\gamma(i,j)$ is the conditional probability that part i is hit by gun j knowing that a hit has been scored. In practice it may be difficult to determine the value of $\gamma(i,j)$ since the proportions in which the different guns are used by the enemy may be unknown. On the other hand, it seems likely that frequently we shall be able to estimate the conditional probability $\gamma(i|j)$ that part i is hit knowing that a hit has been scored by gun j. The purpose of this memorandum is to investigate the question whether $q(i,j)$ can be estimated from the data assuming that merely the quantities $\gamma(i|j)$ are known a priori. In what follows we shall restrict ourselves to the case of independence, i.e., it will be assumed that the probability of surviving a hit does not depend on the non-destructive hits already received.

Let $\delta(i,j)$ be the conditional probability that part i is hit by gun j knowing that a hit has been scored and the plane survived the hit. Furthermore, let q be the probability that the plane survives a hit (not knowing which part was hit and which gun scored the hit). Then, similar to equation 82, we shall have

$$q(i,j) = \frac{\delta(i,j)}{\gamma(i,j)} q \ . \tag{101}$$

Let $q(j)$ be the probability that the plane will survive a hit by gun j (not knowing the part hit). Then obviously

$$q(j) = \sum_1 \gamma(i|j)q(i,j) \quad . \tag{102}$$

Let $\delta(i|j)$ be the conditional probability that part i is hit by gun j knowing that a hit has been scored by gun j and the plane survived the hit. Clearly

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[1]This part of "A Method of Estimating Plane Vulnerability Based on Damage of Survivors" was published as SRG memo 126 and AMP memo 76.8.

$$\delta(i|j) = \frac{\gamma(i|j)q(i,j)}{\sum_i \gamma(i|j)q(i,j)} = \frac{\gamma(i|j)q(i,j)}{q(j)} \quad . \tag{103}$$

From equation 103, we obtain

$$q(i,j) = \frac{\delta(i|j)}{\gamma(i|j)} q(j) \quad . \tag{104}$$

The quantity $\delta(i|j)$ can be estimated on the basis of the observed hits on the returning planes. The best sample estimate of $\delta(i|j)$ is the ratio of the number of hits scored by gun j on part i of the returning planes to the total number of hits scored by gun j on the returning planes. Thus, on the basis of equation 104, the probability $q(i,j)$ can be determined if $q(j)$ is known.

Now we shall investigate the question whether $q(j)$ can be estimated. First, we shall consider the case when it is known a priori that a certain part of the plane, say part 1, is not vulnerable. Then $q(i,j) = 1$ and we obtain from equation 104

$$1 = \frac{\delta(1|j)}{\gamma(1|j)} q(j) \quad . \tag{105}$$

Hence,

$$q(j) = \frac{\gamma(1|j)}{\delta(1|j)} \quad . \tag{106}$$

Thus, in this case our problem is solved. If no part of the plane can be assumed to be invulnerable, then we can still obtain upper limits for $q(j)$. In fact, since $q(i,j) \leq 1$, we obtain from equation 104

$$q(j) \leq \frac{\gamma(i|j)}{\delta(i|j)} \quad . \tag{107}$$

Denote by $\rho(j)$ the minimum of $\frac{\gamma(i|j)}{\delta(i|j)}$ with respect to 1. Then we have

$$q(j) \leq \rho(j) \quad . \tag{108}$$

If there is a part of the airplane that is only slightly vulnerable (this is usually the case), then $q(j)$ will not be much below $\rho(j)$. Let the part i_j be the part of the plane least

vulnerable to gun j. If $q(i_j,j)$ has the same value for any gun j, then $q(j)$ is proportional to $\rho(j)$. Thus, the error is perhaps not serious if we assume that $q(j)$ is proportional to $\rho(j)$, i.e.,

$$q(j) = \lambda\rho(j). \tag{109}$$

The proportionality factor λ can be determined as follows. From equations 101 and 104 we obtain

$$\frac{\delta(i,j)}{\gamma(i,j)} q = \lambda\rho(j) \frac{\delta(i|j)}{\gamma(i|j)} \quad . \tag{110}$$

Hence,

$$\lambda\gamma(i,j) = q \frac{\delta(i,j)\gamma(i|j)}{\delta(i|j)\rho(j)} \quad . \tag{111}$$

Denote $\sum_{i} \delta(i,j)$ by $\delta(j)$. Then,

$$\delta(i|j) = \frac{\delta(i,j)}{\delta(j)} \quad . \tag{112}$$

From equations 111 and 112 we obtain

$$\lambda\gamma(i,j) = q \frac{\delta(j)\gamma(i|j)}{\rho(j)} \quad . \tag{113}$$

Since

$$\sum_{i} \gamma(i|j) = 1,$$

we obtain from equation 113

$$\lambda \sum_{j} \sum_{i} \gamma(i,j) = q \sum_{j} \frac{\delta(j)}{\rho(j)} \quad . \tag{114}$$

But

$$\sum_{j} \sum_{i} \gamma(i,j) = 1.$$

Hence,

$$\lambda = q \sum_j \frac{\delta(j)}{\rho(j)} \ .$$

(115)

Since $\delta(j)$ and $\rho(j)$ are known quantities, the proportionality factor λ can be obtained from equation 115. The probability q is the root of the equation

$$\sum_{j=1}^{n} \frac{a_j}{q^j} = 1 - a_o,$$

where a_j denotes the ratio of the number of planes returned with exactly j hits to the total number of planes participating in combat.

NUMERICAL EXAMPLE

In part V, the case of a plane subdivided into several equi-vulnerability areas was discussed, and the vulnerability of each part was estimated. The same method can be extended to solve the more general problem of estimating the probability that a plane will survive a hit on part i caused by a bullet fired from gun j, if assumptions corresponding to those of part V are made. The first three of the four assumptions that must be made to apply the method of part V directly are identical with those made in part V. They are:

- The number of planes participating in combat is large so that sampling errors can be neglected.

- The probability that a hit will not down the plane does not depend on the number of previous non-destructive hits. That is, $q_1 = q_2 = \ldots = q_o$ (say), where q_i is the conditional probability that the i-th hit will not down the plane, knowing that the plane is hit.

- The division of the plane into several parts is representative of all planes of the mission.

The fourth assumption necessary to apply the method of part V directly usually cannot be fulfilled in practice. It is:

- Given that a shot has hit the plane, the probability that it hit a particular part, and was fired from a particular type of gun, is known.

These probabilities depend upon the proportions in which different guns are used by the enemy. To overcome this difficulty a method that does not depend on these proportions is developed in part VIII. The assumptions necessary for the method of part VIII differ from those of part V only in that the fourth assumption is replaced by:

- Given that a shot has hit the plane, and given that it was fired by a particular type of gun, the probability that it hit a particular part is known.

The information necessary to satisfy this assumption is more readily available, and in the numerical example that follows a simplified method is suggested for estimating these probabilities.

The Data

The numerical example will be an analysis of a set of hypothetical data, which is based on an assumed record of damage of surviving planes of a mission of 1,000 planes dispatched to attack an enemy objective. Of the 1,000 planes dispatched, 634 (N) actually attacked the objective. Thirty-two planes were lost (L=32) in combat and the number of hits on returning planes was:

$$A_i = \text{number of planes returning with i hits}$$

$$
\begin{aligned}
A_0 &= 386 \\
A_1 &= 120 \\
A_2 &= 47 \\
A_3 &= 22 \\
A_4 &= 16 \\
A_5 &= 11
\end{aligned}
\tag{A}
$$

The total number of hits on all returning planes is

$$A_1 + 2A_2 + 3A_3 + 4A_4 + 5A_5 = \tag{B}$$

$$120 + 2 \times 47 + 3 \times 22 + 4 \times 16 + 5 \times 11 = 399 \quad .$$

These 399 hits were made by three types of enemy ammunition:

	B$_1$	Flak
	B$_2$	20-mm aircraft cannon
	B$_3$	7.9-mm aircraft machine gun

and the hits by these different types of ammunition were also recorded by part of airplane hit:

	C$_1$	Forward fuselage
	C$_2$	Engine
	C$_3$	Full system
	C$_4$	Remainder

The necessary information from the record of damage is given in table 7.

TABLE 7

NUMBER OF HITS OF VARIOUS TYPES BY PARTS

	Forward fuselage, C$_1$	Engine, C$_2$	Fuel system, C$_3$	Remainder, C$_4$	Total all parts
Flak, B$_1$	17	25	50	202	294
20-mm cannon, B$_2$	8	7	17	18	50
7.9-mm machine gun, B$_3$	7	13	17	18	55
Total all types	32	45	84	238	399

A Method of Estimating the Probability of Hitting a Particular Part Given That a Shot of a Particular Ammunition Has Hit the Plane[1]

The conditional probability that a plane will be hit on the i-th area, knowing that the hit is of the j-th type, must be determined from other sources of information than the record of

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[1]Necessary for fourth assumption.

damage. Although a simplified method is used in this example, more accurate estimates can be made if more technical data is at hand. The first step is to make definite boundaries for the areas C_1, C_2, C_3, C_4. Next, assume that each type of enemy fire B_1, B_2, B_3 has an average angle of fire θ_1, θ_2, θ_3. Finally, assume that the probability of hitting a part of the plane from a given angle is equal to the ratio of the exposed area of that part from the given angle to the total area exposed from that angle.

In this example it is assumed that flak (B_1) has the average angle of attack of 45 degrees in front of and below the plane, whereas 20-mm cannon and 7.9-mm machine gun fire both hit the plane head-on on the average. The area C_1 is so bounded that it includes areas which, if hit, will endanger the pilot and co-pilot. Area C_2 includes the engine area and area C_3 consists essentially of the area covering the fuel tanks. The results of computations, based on the above assumptions, are assumed to be as follows, where $\gamma(C_i|B_j)$[1] represents the probability that a hit is on part C_i knowing it is of type B_j (as estimated by determining the ratio of the area of C_i to the total area as viewed from the angle θ_j associated with ammunition B_j).

(C)

$\gamma(C_1|B_1) = .058$	$\gamma(C_1|B_2) = .143$	$\gamma(C_1|B_3) = .143$
$\gamma(C_2|B_1) = .092$	$\gamma(C_2|B_2) = .248$	$\gamma(C_2|B_3) = .248$
$\gamma(C_3|B_1) = .174$	$\gamma(C_3|B_2) = .303$	$\gamma(C_3|B_3) = .303$
$\gamma(C_4|B_1) = .676$	$\gamma(C_4|B_2) = .306$	$\gamma(C_4|B_3) = .306$

---

[1]This notation differs from the previous notation of part VIII. In the first part of part VIII, $\gamma(i|j)$ is used with the understanding that the first subscript refers to the part hit and the second subscript refers to the type of bullet. In the numerical example, the relationship is made explicit by letting C_i stand for the i-th part (or component) and B_j for the j-th type of bullet. The same device is used throughout this example.

Computations for Method of Part VIII

Let $q(C_i, B_j)$ be the probability of surviving a hit on part C_i by gun B_j. By equation 104, we have

$$q(C_i, B_j) = \frac{\delta(C_i | B_j)}{\gamma(C_i | B_j)} q(B_j) \quad , \tag{D}$$

where $\delta(C_i | B_j)$ is the probability of being hit on part C_i, knowing that the hit was scored by a bullet from gun B_j and that the plane survived; $\gamma(C_i | B_j)$ is the probability of being hit on part C_i, knowing that the hit was scored by a bullet of type B_j; and $q(B_j)$ is the probability that a plane will survive a hit of type B_j, knowing that the plane is hit. This can be estimated by taking the ratio of the number of hits of type B_j on part C_i to the total number of hits of type B_j on returning planes. Applying this method to the table we obtain

$$\tag{E}$$

$$\delta(C_1 | B_1) = .058 \qquad \delta(C_1 | B_2) = .160 \qquad \delta(C_1 | B_3) = .127$$
$$\delta(C_2 | B_1) = .085 \qquad \delta(C_2 | B_2) = .140 \qquad \delta(C_2 | B_3) = .236$$
$$\delta(C_3 | B_1) = .170 \qquad \delta(C_3 | B_2) = .340 \qquad \delta(C_3 | B_3) = .309$$
$$\delta(C_4 | B_1) = .687 \qquad \delta(C_4 | B_2) = .360 \qquad \delta(C_4 | B_3) = .327$$

The final quantity required to calculate $q(C_i, B_j)$ by equation D is $q(B_j)$. By equation 109, we have

$$q(B_j) = \lambda \rho(B_j) \quad , \tag{F}$$

where $\rho(B_j)$ is the minimum of $\dfrac{\gamma(C_i | B_j)}{\delta(C_i | B_j)}$ with respect to i.

$$\rho(B_j) = \min \left\{ \frac{\gamma(C_1 | B_j)}{\delta(C_1 | B_j)} , \frac{\gamma(C_2 | B_j)}{\delta(C_2 | B_j)} , \frac{\gamma(C_3 | B_j)}{\delta(C_3 | B_j)} , \frac{\gamma(C_4 | B_j)}{\delta(C_4 | B_j)} \right\}$$

$$\rho(B_1) = \min \left\{ \frac{.058}{.058} , \frac{.092}{.085} , \frac{.174}{.170} , \frac{.676}{.687} \right\}$$

$$= \min \left\{ 1 \quad , >1 \quad , >1 \quad , .984 \right\}$$

$$= .984$$

(G)

$$\rho(B_2) = \min \left\{ \frac{.143}{.160} , \frac{.248}{.140} , \frac{.303}{.340} , \frac{.306}{.360} \right\}$$

$$= \min \left\{ .894 , >1 \quad , .891 , .850 \right\}$$

$$= .850$$

$$\rho(B_3) = \min \left\{ \frac{.143}{.127} , \frac{.248}{.236} , \frac{.303}{.309} , \frac{.306}{.327} \right\}$$

$$= \min \left\{ >1 \quad , >1 \quad , .981 , .936 \right\}$$

$$= .936$$

The constant multiplier λ is defined by equation 115

$$\lambda = q \sum \frac{\delta(B_j)}{\rho(B_j)} ,$$

(H)

where $\delta(B_j)$ is the conditional probability that a hit is of type B_j.

The determination of q is identical with the procedure of part VII. From equation 26

$$\sum \frac{A_j}{q^j} = N - A_o$$

we substitute the values of equation A:

$$248q^5 - 120q^4 - 47q^3 - 22q^2 - 16q - 11 = 0$$

(I)

The root is .930 (= q_0, say).

The values $\delta(B_j)$ are obtained directly from table 7 by taking the ratio of hits of type B_j on returning planes to the total number of hits on returning planes.

$$\delta(B_1) = \frac{294}{399} = .737$$

$$\delta(B_2) = \frac{50}{399} = .125 \tag{J}$$

$$\delta(B_3) = \frac{55}{399} = .138$$

Substituting the results of equations G, I, and J in equation H, we obtain:

$$\lambda = q_0 \sum \frac{\delta(B_j)}{\rho(B_j)}$$

$$= .930 \left\{ \frac{.737}{.984} + \frac{.125}{.850} + \frac{.138}{.936} \right\}$$

$$= .930 \ (1.0433)$$

$$= .9703$$

Substituting in equation F

$$q(B_1) = (.9703) \ (.984) = .955$$

$$q(B_2) = (.9703) \ (.850) = .825 \tag{K}$$

$$q(B_3) = (.9703) \ (.936) = .908$$

The probabilities $q(C_i, B_j)$ can now be determined from equation D by using the values given in equations C, E, and K.

$$q(C_i, B_j) = \frac{\delta(C_i | B_j)}{\gamma(C_i | B_j)} \ q(B_j)$$

$$q(C_1, B_1) = (.058)(.955)/.058 = .955$$
$$q(C_2, B_1) = (.085)(.955)/.092 = .882$$
$$q(C_3, B_1) = (.170)(.955)/.174 = .933$$
$$q(C_4, B_1) = (.687)(.955)/.676 = .971$$

$$q(C_1, B_2) = (.160)(.825)/.143 = .923$$
$$q(C_2, B_2) = (.140)(.825)/.248 = .466$$
$$q(C_3, B_2) = (.340)(.825)/.303 = .926 \qquad (L)$$
$$q(C_4, B_2) = (.360)(.825)/.306 = .971$$

$$q(C_1, B_3) = (.127)(.908)/.143 = .806$$
$$q(C_2, B_3) = (.236)(.908)/.248 = .864$$
$$q(C_3, B_3) = (.309)(.908)/.303 = .926$$
$$q(C_4, B_3) = (.327)(.908)/.306 = .970$$

Comments on Results

The vulnerability of a plane to a hit of type B_j on part C_i is the probability that a plane will be destroyed if it receives a hit of type B_j on part C_i. Let $P(C_i, B_j)$ represent this vulnerability. The numerical value of $P(C_i, B_j)$ is obtained from the set L and the relationship

$$P(C_i, B_j) = 1 - q(C_i, B_j) \qquad (M)$$

The vulnerability of a plane to a hit to type B_j on part C_i is given in table 8.

This analysis of the hypothetical data would lead to the conclusion that the plane is most vulnerable to a hit on the engine area if the type of bullet is not specified, and is most vulnerable to a hit by a 20-mm cannon shell if the part hit is not specified. The greatest probability of being destroyed is .534, and occurs when a plane is hit by a 20-mm cannon shell

on the engine area. The next most vulnerable event is a hit by a
7.9-mm machine gun bullet on the cockpit. These, and other
conclusions that can be made from the table of vulnerabilities
derived by the method of analysis of part VIII, can be used as
guides for locating protective armor and can be used to make a
prediction of the estimated loss of a future mission.

TABLE 8

VULNERABILITY OF A PLANE TO A HIT OF A SPECIFIED TYPE
ON A SPECIFIED PART

	Forward fuselage	Engine	Fuel system	Remainder	Vulner-ability to specified type of hit when area is un-specified
Flak, B_1	.045	.118	.067	.029	.045
20-mm cannon, B_2	.077	.534	.074	.029	.175
7.9-mm machine gun, B_3	.194	.136	.074	.030	.092
Vulnerability to hit on specified area when type of hit is un-specified[a]	.114	.179	.074	.038	.070[b]

---

[a]These vulnerabilities are calculated using the method of
part V, and assuming that the $\gamma(C_i)$, the probability that part
C_i is hit, knowing that the plane is hit, are as follows:

$$\gamma(C_i) = .084 \quad \gamma(C_2) = .128 \quad \gamma(C_3) = .212 \quad \gamma(C_4) = .576 \ .$$

[b]This is the probability that a plane will be destroyed by a
hit, when neither the part hit nor the type of bullet is
specified.

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This research contribution consists of a series of eight memoranda originally published by the Statistical Research Group at Columbia University for the National Defense Research Committee in 1943 on methods of estimating the vulnerability of various parts of an aircraft based on damage to surviving planes. The methodology presented continues to be valuable in defense analysis and, therefore, has been reprinted by the Center for Naval Analyses in order to achieve wider dissemination.

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