

Lect 9 Prob 201 11/23/17

Recap:  $g(x)$  special case...

$$M_X(t) := E[e^{tX}] \text{ def of mgf}$$

Properties

I  $M_X(t) = M_Y(t) \Rightarrow X \stackrel{d}{=} Y$

Note:  $E[X^0] = 1 = M_X(0)$  (at  $x=0$  case)

II  $E[X^k] = m_X^{(k)}(0)$

III if  $Y = aX + c \Rightarrow M_Y(t) = e^{tc} M_X(at)$

IV if  $X, Y$  indep if  $X, Y$  iid  
 $M_{X+Y}(t) = M_X(t) M_Y(t) = (M_X(t))^2$

V Levy Continuity thm

if  $\lim_{n \rightarrow \infty} M_{X_n}(t) = M_Y(t) \Leftrightarrow X \rightarrow Y$  r.v.  $X$  converges to r.v.  $Y$   
 $\Rightarrow$  if  $n$  large  $X_n \stackrel{d}{\approx} Y$   $X, Y$  are approx. equally distr.  
 $\Rightarrow f_{X_n} \approx f_Y$  PDFs approx

if  $X \sim \text{Bern}(p) \Rightarrow M_X(t) = 1 - p + pe^t$  (from def)

$X \sim \text{Binom}(n, p) \Rightarrow M_X(t) = (1 - p + pe^t)^n$  (rule IV)

$X \sim \text{Geom}(p) \Rightarrow M_X(t) = \frac{pe^t}{1 - e^t(1-p)}$  if  $t < \ln(1/(1-p))$  (from def) (Hw)

$X \sim \text{Exp}(\lambda) \Rightarrow M_X(t) = \frac{\lambda}{\lambda - t}$  if  $t < \lambda$  (from def)

$Z \sim N(0, 1) \Rightarrow M_Z(t) = e^{\frac{t^2}{2}}$ ,  $X = \mu + \sigma Z$ ,  $M_X(t) = \dots$

Proof this:  $X \sim N(\mu, \sigma^2) \Rightarrow M_X(t) = e^{\mu t + \frac{1}{2}\sigma^2 t^2}$   $X \sim \text{Log}(c) \Rightarrow M_X(t) = e^{ct}$  Proof this

$$X_1 \sim N(\mu_1, \sigma_1^2), \text{ iid of } X_2 \sim N(\mu_2, \sigma_2^2)$$

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⑤ Lévy Continuity Thm

$$X_1 + X_2 \sim N(\mu_1 + \mu_2, \sigma_1^2 + \sigma_2^2)$$

Let  $X_1, X_2, \dots$  be a sequence of r.v.'s

$$\text{if } \lim_{n \rightarrow \infty} M_{X_n}(t) = M_Y(t) \Rightarrow \lim_{n \rightarrow \infty} F_{X_n}(x) = F_Y(x) \quad \forall x \quad \text{i.e.} \quad \lim_{n \rightarrow \infty} X_n \stackrel{d}{=} Y$$

if the r.m.f. looks more and more like the r.m.f. of another r.v., the limit is that other r.v.

Assume  $X_1, X_2, \dots, X_n$  i.i.d. same dist with finite  $\mu$ .

Recall  $\bar{X}_n \rightarrow \mu$  "the law of large #s (LLN)"

Note:  $\mu \sim \text{deg}(\mu)$

So if  $\lim_{n \rightarrow \infty} M_{\bar{X}_n}(t) = M_\mu(t) = e^{t\mu}$ , then this is proven  $x = e^{t\mu}$

$$M_{\bar{X}_n}(t) = M_{\frac{X_1 + \dots + X_n}{n}}(t) = M_{X_1 + \dots + X_n}\left(\frac{t}{n}\right) = \left(M_X\left(\frac{t}{n}\right)\right)^n = e^{n \ln\left(M_X\left(\frac{t}{n}\right)\right)} = e^{\frac{\ln(M_X(\frac{t}{n}))}{\frac{1}{n}}}$$

$$\lim_{n \rightarrow \infty} M_{\bar{X}_n}(t) = e^{\lim_{n \rightarrow \infty} \frac{\ln(M_X(\frac{t}{n}))}{\frac{1}{n}}} = e^{\lim_{d \rightarrow 0} \frac{\ln(M_X(td))}{d}} = e^{\lim_{d \rightarrow 0} \frac{t M_X'(td)}{M_X(td)}} = e^{t\mu} \quad \checkmark$$

$$\text{let } d = \frac{1}{n} \quad n \rightarrow \infty \Rightarrow d \rightarrow 0$$

L'Hôpital's Rule

Recall

$$M_X(0) = E[e^{t \cdot 0}] = E[1] = 1$$

$$M_X'(0) = E[X] = \mu$$

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$$\Rightarrow \mu_{\frac{1}{\sqrt{n}}}(t) = e^{t/\sqrt{n}} = \mu_n(t) \Rightarrow \bar{X} \Rightarrow \mu \text{ proof of CLT}$$

Still... why is ZMC? or special?

$X_1, \dots, X_n$  iid satisfying with mean  $\mu$  and s.d.  $\sigma$

Consider  $C_n := \frac{\bar{X}_n - \mu}{\frac{\sigma}{\sqrt{n}}}$   $C_n$  is the standardized  $\bar{X}_n$

$E(C_n) = 0$   
 $SE(C_n) = 1$  } why? Recall the problem...

Presup:

$$\frac{\bar{X}_n - \mu}{\frac{\sigma}{\sqrt{n}}} = \frac{\sum_{i=1}^n (X_i - \mu)}{\sigma}$$

$$= \frac{\sum_{i=1}^n (X_i - \mu)}{\sigma} \rightarrow \mu = \frac{n\mu}{n}$$

$$= \frac{\sum_{i=1}^n X_i - n\mu}{\sigma}$$

$$= \frac{(X_1 - \mu) + \dots + (X_n - \mu)}{\sigma \sqrt{n}}$$

$$= \frac{1}{\sqrt{n}} \left( \frac{X_1 - \mu}{\sigma} + \dots + \frac{X_n - \mu}{\sigma} \right)$$

$$= \frac{1}{\sqrt{n}} (Z_1 + \dots + Z_n)$$

let  $Z_i := \frac{X_i - \mu}{\sigma}$

Note:  $E(Z_i) = 0$ ,  $SE(Z_i) = 1$  why?

$$\mu_{C_n}(t) = \mu_{\frac{1}{\sqrt{n}}(Z_1 + \dots + Z_n)}(t) = \mu_{Z_1 + \dots + Z_n}\left(\frac{t}{\sqrt{n}}\right) = \left(\mu_Z\left(\frac{t}{\sqrt{n}}\right)\right)^n = e^{n \ln\left(\mu_Z\left(\frac{t}{\sqrt{n}}\right)\right)} = e^{\frac{\ln(\mu_Z(\frac{t}{\sqrt{n}}))}{\frac{1}{n}}} = e^{\frac{t^2 \ln(\mu_Z(\frac{t}{\sqrt{n}}))}{t^2/n}}$$

let  $u = \frac{1}{\sqrt{n}}$

$$M_{C_n} = e^{t^2 \frac{\ln(M_2(t))}{t^2}}$$

$$\lim_{n \rightarrow \infty} M_{C_n} = \lim_{u \rightarrow 0} e^{t^2 \frac{\ln(M_2(t))}{t^2}} = e^{t^2 \lim_{u \rightarrow 0} \frac{\ln(M_2(t))}{t^2}}$$

L'Hopital's  $\downarrow$   $\frac{M_2'(t)}{M_2(t)}$   $\downarrow$   $\frac{1}{2} \lim_{u \rightarrow 0} \frac{\frac{d}{dt} [M_2'(t)/M_2(t)]}{t}$

$$\frac{M_2'(t) M_2'(t) + M_2(t) M_2''(t)}{M_2(t)^2} = \frac{1}{2} \frac{M_2'(0)^2 + M_2(0) M_2''(0)}{M_2(0)^2} = \frac{1}{2}$$

$\sigma_1$

$$\Rightarrow \lim_{n \rightarrow \infty} M_{C_n}(t) = e^{\frac{t^2}{2}} \Rightarrow C_n \rightarrow N(0,1)$$

Central Limit Theorem !! (CLT)

$$X_1, X_2, \dots \stackrel{i.i.d.}{\sim} f(t, \sigma^2) \leftarrow \text{Common structure}$$

$$\Rightarrow C_n = \frac{\bar{X} - \mu}{\frac{\sigma}{\sqrt{n}}} \rightarrow N(0,1) \leftarrow \text{Super important!}$$

This is the mystery ... solved!

How to use CLT.

Obtaining  $n \rightarrow \infty$  is impractical

if  $n$  is large, CLT approximately works. There are three formulas you should know. If  $X_1, X_2, \dots, X_n \stackrel{iid}{\sim} \mathcal{N}(\mu, \sigma^2)$

$$\textcircled{1} \frac{\bar{X} - \mu}{\frac{\sigma}{\sqrt{n}}} \stackrel{d}{\approx} N(0, 1)$$

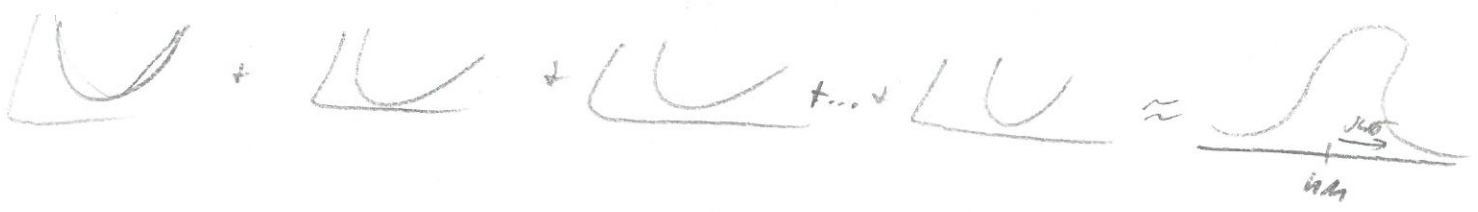
$$\begin{array}{l} \textcircled{2} \bar{X} \stackrel{d}{\approx} N\left(\mu, \left(\frac{\sigma}{\sqrt{n}}\right)^2\right) \quad * \\ \textcircled{3} T \stackrel{d}{\approx} N\left(\mu, \left(\frac{\sigma}{\sqrt{n}}\right)^2\right) \quad * \end{array} \quad \left. \vphantom{\begin{array}{l} \textcircled{2} \\ \textcircled{3} \end{array}} \right\} \text{most useful}$$

$$X \sim \text{Bernoulli}(p)$$

$$\underbrace{\text{|||||}} + \underbrace{\text{|||||}} = \underbrace{\text{|||||}}_{\text{already it begins!}}$$



$$X_1 + \dots + X_n \approx N(\mu, \sigma^2)$$



$$U = \sim ??$$

$$f(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} \approx e^{-x^2} \text{ makes tails small!}$$

normal better because means & errors

When is a certain data considered "normal"?

Why is it called the normal distr??

Complex  $X_1, \dots, X_n \sim \text{Geo}(\frac{1}{2})$

What's the prob the avg wait time is more than 2.75?  
sad!!

$$P(\bar{X} \geq 2.75) \approx P\left(\frac{\bar{X} - 0.5}{.258} \geq \frac{2.75 - 0.5}{.258}\right) = P(Z \geq 3) = .0044$$

$$\text{Roll } \bar{X} \approx N\left(\mu, \left(\frac{\sigma}{\sqrt{n}}\right)^2\right) \Rightarrow \bar{X} \approx N(0.5, .258^2)$$

$$\mu = p = \frac{1}{2}, \sigma = \sqrt{\frac{p(1-p)}{n}} = \sqrt{\frac{.5}{4}} = .258 \Rightarrow \frac{\sigma}{\sqrt{n}} \approx \frac{.258}{2} \approx .129$$