

Recall

11/27 - Lecture 20

$$\begin{cases} M_X(t) = E[e^{tx}] \\ \text{Mgf} \end{cases}$$

(I) $M_X(t) = M_Y(t) \Leftrightarrow X \stackrel{d}{=} Y$

(II) $E[X^k] = M_X^{(k)}(0)$

(III) $Y = aX + c \Leftrightarrow M_Y(t) = e^{tc} M_X(at)$

(IV) if X, Y independent $\Rightarrow M_{X+Y}(t) = M_X(t) M_Y(t)$

$X \sim \text{Bern}(p) \Rightarrow M_X(t) = 1 - p + pe^t$

$X \sim \text{Binomial}(n, p) \Rightarrow M_X(t) = (1 - p + pe^t)^n$

$X \sim \text{Exp}(\lambda) \Rightarrow M_X(t) = \frac{\lambda}{\lambda - t} \text{ if } t < \lambda$

$Z \sim N(0, 1) \Rightarrow M_Z(t) = e^{t^2/2}$

$X \sim N(\mu, \sigma^2)$

$X = \mu + \sigma Z \sim N(\mu, \sigma^2) \Rightarrow M_X(t) = e^{t\mu} M_Z(\sigma t) = e^{t\mu} e^{(\sigma t)^2/2} = e^{t\mu + \frac{\sigma^2 t^2}{2}}$

$X_1 \sim N(\mu_1, \sigma_1^2)$ independence of $X_2 \sim N(\mu_2, \sigma_2^2)$

$Y = X_1 + X_2 \sim ?$

$M_Y(t) = M_{X_1}(t) M_{X_2}(t)$

$= e^{t\mu_1 + \frac{\sigma_1^2 t^2}{2}} e^{t\mu_2 + \frac{\sigma_2^2 t^2}{2}}$

$= e^{(\mu_1 + \mu_2)t + \frac{(\sigma_1^2 + \sigma_2^2)t^2}{2}}$

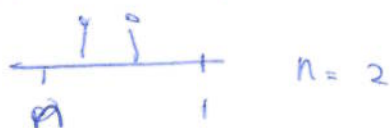
$\Rightarrow Y \sim N(\mu_1 + \mu_2, \sigma_1^2 + \sigma_2^2)$

(V) Levy's Continuity Theorem

Let x_1, x_2 be a sequence of r.v.'s

$$\text{if } \lim_{n \rightarrow \infty} \mu_{x_n}(t) = \mu_x(t) \Rightarrow \lim_{n \rightarrow \infty} F_{x_n}(x) = F_x(x)$$

convergence in distribution.



$\bar{x} \rightarrow \mu = E(x)$ ← as long as iid, \bar{x} will equal μ .

Law of large #'s

$\mathcal{H} - \text{Deg}(\mathcal{H})$

Assume

x_1, x_2, \dots iid with mean μ

$$\lim_{x_n} \mu_{x_n}(t) = e^{t\mu} \leftarrow \text{for the Deg}(\mu)$$

III

$$\Rightarrow \lim_{n \rightarrow \infty} \mu_{\frac{x_1 + \dots + x_n}{n}}(t) \stackrel{\text{III}}{=} \lim_{n \rightarrow \infty} \mu_{x_1 + \dots + x_n}\left(\frac{t}{n}\right)$$

$$= \lim_{n \rightarrow \infty} \mu_{x_1}\left(\frac{t}{n}\right) \stackrel{\text{IV}}{=} \mu_{x_2}\left(\frac{t}{n}\right) \dots \mu_{x_n}\left(\frac{t}{n}\right) =$$

$$= \lim_{n \rightarrow \infty} \left(\mu_x\left(\frac{t}{n}\right) \right)^n = \lim_{n \rightarrow \infty} e^{\ln(\mu_x(\frac{t}{n}))^n}$$

$$= \lim_{n \rightarrow \infty} e^{n \ln(\mu_x(\frac{t}{n}))} = \lim_{n \rightarrow \infty} e^{\frac{\ln(\mu_x(\frac{t}{n}))}{\frac{1}{n}}} = e^{\lim_{n \rightarrow \infty} \frac{\ln(\mu_x(\frac{t}{n}))}{\frac{1}{n}}}$$

↳

$$\text{let } v = \frac{1}{n}$$

$$\Rightarrow n \rightarrow \infty \Rightarrow v \Rightarrow 0$$

$$= e^{\lim_{v \rightarrow 0} \frac{\ln(\mu_x(vt))}{n}} =$$

↑
L'Hôpital Rule

$$= e^{\lim_{v \rightarrow 0} \frac{t \mu_x'(vt)}{\mu_x(vt)}} = e^{\frac{t(\mu_x'(0))}{\mu_x(0)}} = e^{t\mathcal{M}}$$

Note

$$\frac{d}{dx} [\ln(f(x))] = \frac{f'(x)}{f(x)}$$

$$\textcircled{\text{V}} \lim_{n \rightarrow \infty} \mu_{X_n}(t) = \mu_X(t) \Rightarrow X_n \rightarrow X$$

CDF's are limiting

$$\text{i.e. } \lim_{n \rightarrow \infty} F_{X_n}(x) = F_X(x) \quad \text{if } X$$

Assume

X_1, X_2, \dots iid with mean μ and variance σ^2

consider

standardized average

$$C_n = \frac{\bar{X} - \mu}{\frac{\sigma}{\sqrt{n}}}$$

$$\Rightarrow E[C_n] = 0$$

$$\Rightarrow SE[C_n] = 1$$

$$C_n = \frac{\bar{X} - \mu}{\frac{\sigma}{\sqrt{n}}} = \frac{\sqrt{n}(\bar{X} - \mu)}{\sigma} = \frac{\sqrt{n} \left(\frac{X_1 + X_2 + \dots + X_n}{n} - \mu \right)}{\sigma}$$

$$= \frac{\sqrt{n} \left(\frac{X_1 + \dots + X_n}{n} - \frac{\mu + \dots + \mu}{n} \right)}{\sigma}$$

$$= \frac{(X_1 - \mu) + \dots + (X_n - \mu)}{\sigma \sqrt{n}} = \frac{1}{\sqrt{n}} \left(\frac{X_1 - \mu}{\sigma} + \frac{X_2 - \mu}{\sigma} + \dots + \frac{X_n - \mu}{\sigma} \right)$$

$$\text{let } Z_i = \frac{X_i - \mu}{\sigma} \quad E[Z_i] = 0 \quad = \frac{1}{\sqrt{n}} (Z_1 + \dots + Z_n)$$

$$SE[Z_i] = \sigma$$

$$\mu_{L_n}(t) = \frac{\mu_{z_1 + \dots + z_n}(t)}{\sqrt{n}} \stackrel{\text{III}}{=} \mu_{z_1 + \dots + z_n}\left(\frac{t}{\sqrt{n}}\right)$$

$$\stackrel{\text{IV}}{=} \left(\mu_z\left(\frac{t}{\sqrt{n}}\right)\right)^n = e^{\ln\left(\left(\mu_z\left(\frac{t}{\sqrt{n}}\right)\right)^n\right)} = e^{n \ln \mu_z\left(\frac{t}{\sqrt{n}}\right)}$$

$$= \lim_{n \rightarrow \infty} \mu_{L_n}(t) = e^{\lim_{n \rightarrow \infty} n \ln \mu_z\left(\frac{t}{\sqrt{n}}\right)}$$

$$= \lim_{n \rightarrow \infty} \frac{\ln \mu_z\left(\frac{t}{\sqrt{n}}\right)}{\frac{1}{n}} \stackrel{0/0}{=} e^{t^2} \lim_{n \rightarrow \infty} \frac{\ln \mu_z\left(\frac{t}{\sqrt{n}}\right)}{\frac{t^2}{n}}$$

$$\text{let } u = \frac{t}{\sqrt{n}} \quad n \rightarrow \infty \Rightarrow u \rightarrow 0$$

$$= e^{t^2} \lim_{u \rightarrow 0} \frac{\ln \mu_z(u)}{u^2} = e^{t^2} \lim_{u \rightarrow 0} \frac{\mu_z'(u)}{2u} = e^{t^2} \lim_{u \rightarrow 0} \frac{\mu_z''(u)}{2}$$

↑
L'Hopital Rule

↑
L'Hopital Rule

Note: Look up quotient rule

$$\frac{d}{dx} \frac{f(x)}{g(x)} = \frac{f'(x)g(x) + g'(x)f(x)}{g(x)^2}$$

$$= e^{t^2} \lim_{u \rightarrow 0} \frac{\mu_z''(u) + \mu_z'(u)^2}{2\mu_z(u)^2} = e^{t^2}$$

$$= e^{t^2} \lim_{u \rightarrow 0} \frac{\mu_z(u) \mu_z''(u) - (\mu_z'(u))^2}{\mu_z(u)^2} = \frac{e^{t^2} \mu_z'(0) \mu_z''(0) - (\mu_z'(0))^2}{\mu_z(0)^2}$$

$$1 = \text{Var}[z] = E[z^2] - E[z]^2$$

$$\Rightarrow \frac{\bar{x} - \mu}{\frac{\sigma}{\sqrt{n}}} \Rightarrow N(0,1)$$

Central Limit theorem (CLT)

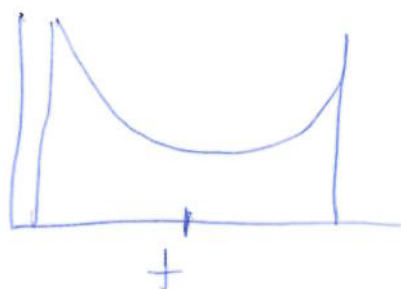
How to use CLT to solve problems

first note that $n \rightarrow \infty$ is impossible
so Γ_n never truly converges

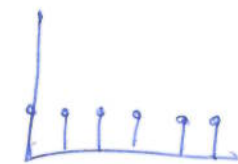
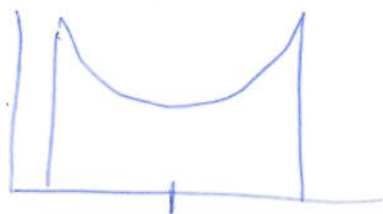
① If n is "large enough" then $\frac{\bar{x} - \mu}{\frac{\sigma}{\sqrt{n}}} \approx N(0,1)$

② $\bar{x} \approx N(\mu, (\frac{\sigma}{\sqrt{n}})^2)$

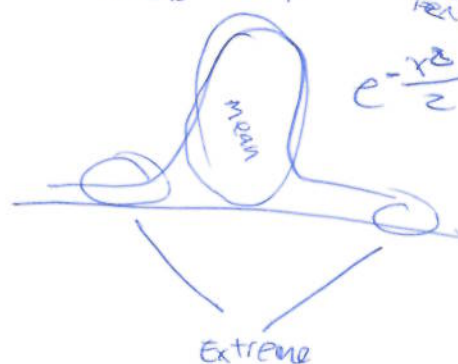
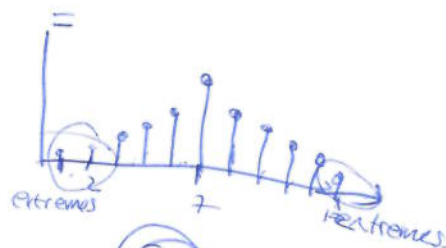
③ $T \approx N(n\mu, (n\sigma^2))$



fat tails



+



$$X_1 \dots X_{30} \approx \text{Geom}(\frac{1}{2}) = \mu = \frac{1}{\frac{1}{2}} = 2, \sigma = \frac{\sqrt{1-p}}{p} = \frac{\sqrt{\frac{1}{2}}}{\frac{1}{2}} \approx 1.414$$

What is the probability the avg wait time is more than 2.75?

$$P(\bar{X} \geq 2.75) = P\left(\frac{\bar{X} - 2}{\frac{1.414}{\sqrt{30}}} \geq \frac{2.75 - 2}{\frac{1.414}{\sqrt{30}}}\right) = P(Z \geq 3) \approx .0015$$

$n=30 \Rightarrow$ "large" = CLT

$$\bar{X} \approx N(\mu, (\frac{\sigma}{\sqrt{n}})^2) = N(2, (1.414)^2) = N(2, 0.258^2)$$

Take 100 steps with probability. Forward and backward being $\frac{1}{2}$

$$x \sim \begin{cases} +1 & \text{wp } \frac{1}{2} \\ -1 & \text{wp } \frac{1}{2} \end{cases}$$

What is the prob. you are more than 10 steps away from ~~at where you~~ starting position after

$$T = x_1 + \dots + x_{100} \stackrel{d}{\approx} \underbrace{N(n\mu, \sigma\sqrt{n})}_{\text{III}} \stackrel{100 \text{ steps}}{=} N(0, \sqrt{100}) = N(0, 10)$$

$$P(|T| \geq 10)$$