

# Nov 29 Lecture

How to use CLT if  $X_1, \dots, X_n \stackrel{iid}{\sim}$  and  $n$  is "large"?

1)  $\bar{X} \stackrel{d}{\sim} N(\mu, (\frac{\sigma}{\sqrt{n}})^2)$

3)  $T \stackrel{d}{\sim} N(n\mu, (n\sigma^2))$

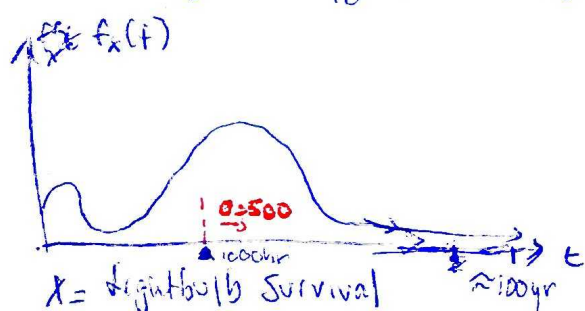
$X_1, \dots, X_{100} \stackrel{iid}{\sim} \begin{matrix} \leq 1 \text{ w.p. } \frac{1}{2} \\ \geq 1 \text{ w.p. } \frac{1}{2} \end{matrix} \Rightarrow \mu=0, \sigma^2=1 \Rightarrow \sigma=1$   
or equal to

What is the probability of being more than 10 steps away from the origin after 100 steps?

$T = X_1 + X_2 + \dots + X_n \approx N(n\mu, (n\sigma)^2) = N(0, 10^2)$

$P(T \geq 10 \text{ or } T \leq -10) = P(T \geq 10) + P(T \leq -10) =$   
 $= P(|T| \geq 10)$

$\rightarrow P(\frac{T-0}{10} \geq \frac{10-0}{10}) + P(\frac{T-0}{10} \leq \frac{-10-0}{10}) = P(Z \geq 1) + P(Z \leq -1) = 16\% + 16\% = 32\%$



If you get 500 bulbs. What is the prob the average lifetime is more than 1300 hrs?

$P(\bar{X} > 1300)$

$\bar{X} \stackrel{d}{\sim} N(\mu, (\frac{\sigma}{\sqrt{n}})^2) = N(1000, (\frac{500}{\sqrt{500}})^2) = N(1000, 70.7^2)$

by CLT

standardize

$\rightarrow P(\bar{X} > 1300) \approx P(\frac{\bar{X} - 1000}{70.7} > \frac{1300 - 1000}{70.7}) = P(Z > 4.24)$   
 $\approx 0$



Shipments are late 2% of the time.

In 10,000 orders, what is the probability more than 3% are late?

$$P(\bar{X} > 0.03) \approx P$$

$$X_1, X_2, \dots, X_{10000} \stackrel{iid}{\sim} \text{Bern}(0.02)$$

$$\bar{Y} \stackrel{d}{\approx}_{\text{by CLT}} N\left(\mu, \left(\frac{\sigma}{\sqrt{n}}\right)^2\right) = N\left(0.02, \left(\frac{0.14}{\sqrt{10000}}\right)^2\right) = N(0.02, 0.0014^2)$$

$$\Rightarrow \mu = 0.02$$

$$\sigma = \sqrt{0.02(1-0.02)} = 0.14$$

$$P(\bar{X} > 0.03) \approx P\left(\frac{\bar{X} - 0.02}{0.0014} > \frac{0.03 - 0.02}{0.0014}\right) = P(Z > 7.14) \approx 0$$

General

$$\bar{X} \stackrel{d}{\approx} N\left(\mu, \left(\frac{\sigma}{\sqrt{n}}\right)^2\right)$$

estimate

$$\hat{p} = \bar{X} = \frac{1+1+0+0+0}{5} = 0.4$$

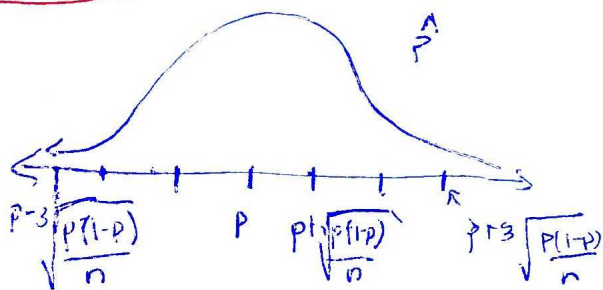
sample of proportion denoted by  $\hat{p}$

$$\hat{p} \stackrel{d}{\approx} N\left(p, \left(\sqrt{\frac{p(1-p)}{n}}\right)^2\right)$$

$$\hat{p} = \bar{X} = \frac{1+1+0+0+0}{5} = 0.4 \approx p$$

special case of  $\bar{X}$ , if  $\bar{X} \sim \text{Bern}$ , then

is one realization of  $\hat{p}$



$$p \pm n \sqrt{\frac{p(1-p)}{n}}$$

prob ↑  
stat ↓

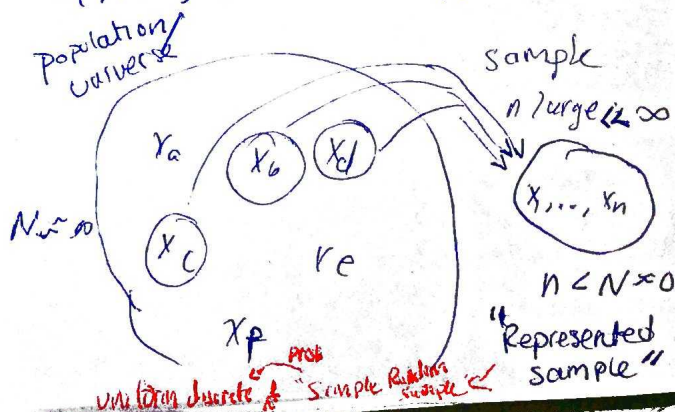
DONE WITH PROB, NOW Stats.

## Statistical Inference

3 goals:

- (I) Estimate best guess of  $p$  ( $\hat{p}$ )
- (II) Estimate range/window of  $p$  (confidence interval)
- (III) Test theories about  $p$  (hypothesis testing)

Need  $X_1, \dots, X_n \stackrel{iid}{\sim} \text{Bern}(p)$





$$P(p \in [\hat{p} \pm \sqrt{\frac{p(1-p)}{n}}]) = P(\hat{p} - \sqrt{\frac{p(1-p)}{n}} \leq p \leq \hat{p} + \sqrt{\frac{p(1-p)}{n}})$$

$$= P(-\sqrt{\frac{p(1-p)}{n}} \leq p - \hat{p} \leq \sqrt{\frac{p(1-p)}{n}})$$

$$= P(-1 \leq \frac{p - \hat{p}}{\sqrt{\frac{p(1-p)}{n}}} \leq 1) = P(-1 \leq -z \leq 1) = P(1 \geq z \geq -1)$$

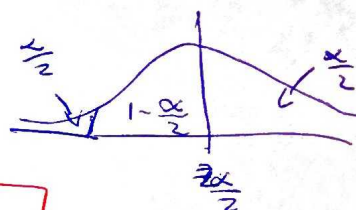
$$= P(z \in [-1, 1]) = .68$$

$$[\hat{p} \pm z_{\frac{\alpha}{2}} \sqrt{\frac{p(1-p)}{n}}]$$

$$z_{\frac{\alpha}{2}} := F_z^{-1}(1 - \frac{\alpha}{2}) \Rightarrow 1 - \frac{\alpha}{2} = \int_{-\infty}^{z_{\frac{\alpha}{2}}} f_z(z) dz$$

$$P(p \in [\hat{p} \pm z_{\frac{\alpha}{2}} \sqrt{\frac{p(1-p)}{n}}]) = P(z \in [-\frac{z_{\alpha}}{2}, \frac{z_{\alpha}}{2}])$$

$$= F(\frac{z_{\alpha}}{2}) - F(-\frac{z_{\alpha}}{2}) = (1 - \frac{\alpha}{2}) - (\frac{\alpha}{2}) = 1 - \alpha$$



$$97.5\% = \int_{-\infty}^{z_{\frac{\alpha}{2}} = z} f_z(z) dz$$

known.

$$[\hat{p} \pm z_{\frac{\alpha}{2}} \sqrt{\frac{p(1-p)}{n}}] \text{ WRONG, because } p \text{ is unknown}$$

When repeated ... gives you  $1 - \alpha$  "coverage" of  $p$ .

longer of debate

$$[\hat{p} \pm z_{\frac{\alpha}{2}} \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}] \rightsquigarrow [\hat{p} \pm z_{\frac{\alpha}{2}} \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}] := CI_{p, 1-\alpha}$$

z sided

1 - proportion confidence interval