

# The offsets between the galaxies and the dark matter in the galaxy clusters of the Illustris simulation

Karen Y. Ng,<sup>1</sup> Annalisa P. Pillepich,<sup>2</sup> D. Wittman,<sup>1</sup> William A. Dawson,<sup>3</sup>  
Lars Hernquist<sup>2</sup>

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## ABSTRACT

Assuming that dark matter has a zero self-interacting cross section, how likely is it for us to see the reported offset values between dark matter and galaxies from real data of merging galaxy clusters? This paper formulates a hypothesis test using cluster data in the cosmological simulation, the Illustris simulation to answer the question. In the process, we examined the accuracies of commonly used galaxy summary statistics, including the luminosity peak, number density peak, shrinking aperture peak, centroid and the brightest cluster galaxy (BCG), with broad applications to the optical studies of galaxy clusters. We found that the choice of summary statistic to affect the offset value significantly with the BCG and the luminosity peak giving the tightest 68-th percentile offsets levels. Out of the 15 reported offset from observed merging clusters that we examined, 13 of them are consistent with the Illustris offset levels to 2-sigma (95-th percentile) level. Although two of the reported offsets from inferred from luminosity peaks lie outside the 99-th percentile level, it is unclear if the large reported offset discrepancy is due to different ways of determining the smoothing kernel width for the luminosity map. We also identified that the long tail of the offset distribution of the BCG to be due to projected substructures. In general, galaxy summary statistics such as shrinking aperture, number density and centroid give a large bias of  $\sim 50 - 100$  kpc at the 68-th percentile level, even for clusters with only one dominant mass component. Excluding the BCG, the luminosity peak is the most robust to the bias from substructures, if a proper statistical procedure is used to determine the smoothing bandwidth for the luminosity map.

**Key words:** galaxy clusters, dark matter, statistics

## 1 INTRODUCTION

During the latest stage of structure formation, the universe gave birth to non-linear, hierarchical structures known as galaxy clusters. These clusters, made up of dark matter, galaxies and hot gas, are constantly accreting mass, merging and evolving with their environments. Bright galaxies that belong to a galaxy cluster or group, in particular, highlight the overdensities of the underlying dark matter (DM) distribution.

In these dense regions of the clusters, the rates of particle interactions can be enhanced, including the long-suspected self-interaction of DM particles (hereafter, SIDM). Many papers have used the offsets between the summary statistics of the DM density and the galaxy density to give constraints on the self-interaction cross section, i.e.  $\sigma_{\text{SIDM}}$ , of dark matter. A lot of observational studies focus on using merging galaxy clusters as they assume the high collisional velocity should increase the chance of detecting the effects of SIDM. By assuming galaxies being relatively collisionless  $\sigma_{\text{gal}} \approx 0$ , any offset of the DM population from the galaxy provides  $\sigma_{\text{SIDM}}$  relative to  $\sigma_{\text{gal}} \approx 0 \text{ cm}^2$ . These observational studies include Markevitch et al. (2004) and Bradač et al. (2006) reporting an offset of 25 kpc for the Bullet Cluster; Dawson (2013a) reporting

an offset of 129 kpc and 47 kpc for the southern and the northern subcluster respectively; Jee et al. (2015) reporting an offset of 190 kpc for MACSJ1752, and others that we list in detail in table 3. However, other studies using 129 X-ray selected relaxed galaxy groups, such as George et al. (2012) also report offsets of the same order of magnitude, at around 50 – 150 kpc.

Simulators have also joined hands to study the velocity dependence of the effects of SIDM. There are many staged simulations of mergers of galaxy clusters that focused on detecting the signal from SIDM. The ‘staging’ step of these simulations usually involve prescribing a parametric spatial distribution of galaxies (Randall et al. 2008, Kahlhoefer et al. 2014, Robertson et al. 2016), such as an NFW profile, and may not simulate the galaxy morphology, nor include dynamical friction. They try to show the level of offsets solely due to SIDM (Kahlhoefer et al. 2014) by initializing the galaxy-DM offset to be zero at the beginning of their simulations. Furthermore, these staged simulations commonly use a much higher number of galaxy particles than the realistic observable number of galaxies. Randall et al. (2008) found an offset of only 1.8 kpc in the staged merger simulation with  $\sigma_{\text{SIDM}} = 0 \text{ cm}^2/\text{g}$  using  $10^5$  galaxy particles. When zero impact parameter was initialized for

mergers, Kim and Peter et al. (in prep.), using 5.7k or 57k galaxy tracer particles, also show null galaxy-DM offset during most periods of their control staged simulation with  $\sigma_{\text{SIDM}} = 0 \text{ cm}^2/\text{g}$ . While we provide a more in-depth comparison with these staged simulation in the discussion, we argue these staged simulations do not sufficiently probe how statistical and observational uncertainties realistically contribute to the galaxy-DM offsets. As such, when the  $\sigma_{\text{SIDM}}$  is increased in the aforementioned staged simulations, the simulators can guarantee the offsets are maximally due to SIDM. When these simulations set the  $\sigma_{\text{SIDM}}$  to observationally motivated levels of  $< 3 \text{ cm}^2\text{g}^{-1}$ , different authors have consistently reported offset signals (50 kpc) smaller with some of the more extreme uncertainties estimated from individual observations (100 kpc). These simulations have raised questions about how strongly the galaxy-DM offsets can constrain the effects of SIDM. When Kahlhoefer et al. (2014) simulated SIDM with both low-momentum-transfer self-interaction and rare self-interactions of DM with high momentum transfer, they found the maximum offsets that are  $< 30 \text{ kpc}$  for  $\sigma_{\text{SIDM}}$  as high as  $1.6 \text{ cm}^2/\text{g}$ . The reported offset from Randall et al. (2008) for  $\sigma_{\text{SIDM}} = 1.24 \text{ cm}^2/\text{g}$  is only 53.9 kpc. While Kim and Peter et al. (in prep.) found a maximum offset  $< 50 \text{ kpc}$  for  $\sigma_{\text{SIDM}} = 3 \text{ cm}^2/\text{g}$ , and Robertson et al. (2016) also found a maximum offset  $\lesssim 40 \text{ kpc}$  from a simulation suite of a Bullet Cluster analog with  $\sigma_{\text{SIDM}} = 1 \text{ cm}^2/\text{g}$ .

An alternative explanation for most of the observed galaxy-DM offset signal is due to statistical and observational uncertainties. Galaxies are sparse samples of the underlying DM overdensities, it is possible that the summary statistic of the sparse sample to be different from those of the underlying distribution. It is not clear if there is any physical origin of the galaxy-DM offset in a CDM universe, but any statistical noise leading to an offset can influence this method of inferring  $\sigma_{\text{SIDM}}$ . Since the Illustris simulation assumes no SIDM, this study is complementary to staged simulations for understanding what can contribute to the offsets. Simply put, we perform a hypothesis test with the galaxy-DM offsets in the Illustris simulation directly corresponding to our null hypothesis  $\mathcal{H}_0$ , with:

$$\begin{cases} \text{the null hypothesis } \mathcal{H}_0 : \text{Cold Dark Matter (CDM)} \\ \text{the alternative hypothesis } \mathcal{H}_1 : \text{Self-interacting Dark Matter (SIDM)} \end{cases} \quad (1)$$

and we try to see if the observed offset data can be compatible with offsets derived from a CDM simulation.

This exercise is further complicated by the fact that there is no theoretical foundation showing which observable would be the most sensitive to each possible type of SIDM. In fact, Kahlhoefer et al. (2014) have argued that SIDM does not cause significant offsets between the galaxy and DM peaks, and only leads to an offset between the corresponding centroids within the dynamical timescale for relaxation ( $\sim$  several Gyr). Popular choices for computing the offsets often involve first inferring the summary statistic of the DM population and those of the galaxy population of a cluster independently before taking a difference. While there are well established procedures driven by lensing physics for inferring the DM spatial distribution, there is no standard procedure for mapping the sparse member galaxy distribution. We quantify the bias and uncertainty associated with the statistic for summarizing the member galaxy population.

Understanding the characteristics of different galaxy summary statistics of clusters is also important for probing the matter fluctuations in the universe. One such study is done by performing lensing analyses on the stacked images of many small galaxy groups and clusters. The derived cluster mass function can provide constraints

to cosmological parameters such as  $\sigma_8$ . For such studies, stacking on the ‘wrong’ centers is a commonly cited source of uncertainties (Johnston et al. 2007, Ford et al. 2014). By comparing the discrepancies of different galaxy summary statistic, we can find out what can help maximize the lensing signal and the possible cause(s) of miscentering.

In this paper, we 1) extract realistic observables from the Illustris simulation for comparison with observations, 2) explore the pros and cons of the different statistic for summarizing the member galaxy population of a galaxy cluster, 3) give estimates for the offsets between the summary statistics of the galaxy population and the DM population under  $\Lambda$ CDM cosmology, which we call

$$\Delta s \equiv s_{\text{gal}} - s_{\text{DM}}. \quad (2)$$

where  $s_{\text{gal}}$  and  $s_{\text{DM}}$  are the two-dimensional (2D) spatial locations of the summary statistic of the galaxy population, and the density peak of DM respectively. This gives an estimate of the baseline scatter of offsets without any SIDM. And finally we 4) examine the properties of the clusters that give outliers in the offset distribution and 5) investigate the correlations between the physical properties of a cluster and the projected observables such as  $\Delta s$ .

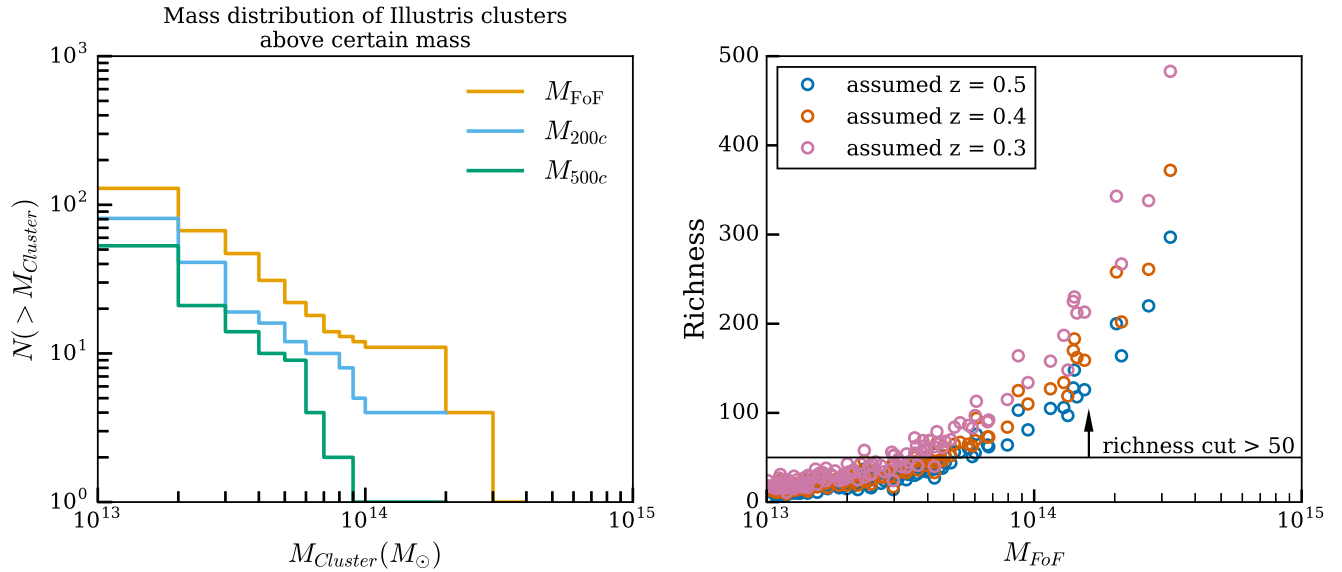
The organization of this paper is as follows: In section 2, we will describe the physical properties of the data of the Illustris simulation (Vogelsberger et al. 2014a, Genel et al. 2014), and the selection criteria that we have employed to ensure that the quantities that we examine resemble observables but without noise and systematics from observations. Then in section 3, we explain the methods for computing various summary statistics of the spatial distribution of galaxies how we prepare our dark matter spatial data to resemble convergence maps. We show the statistical performance of the different summary statistics before we show the main results in section 4. In the discussion in section 5, we list the implications of our results and compare it to other simulations and observations. We also show how one may make use of the population offset statistical distribution from the Illustris data to construct a test with a null hypothesis of  $\sigma_{\text{SIDM}} = 0$  and discuss the caveats.

Our analysis makes use of the same flat Lambda Cold Dark Matter ( $\Lambda$ CDM) cosmology as the Illustris simulation. The relevant cosmological parameters are  $\Omega_{\Lambda} = 0.7274$ ,  $\Omega_m = 0.2726$ , and  $H_0 = 70.4 \text{ km s}^{-1} \text{ Mpc}^{-1}$ .

## 2 THE ILLUSTRIS SIMULATION DATA

The Illustris simulation contains some of the most realistic, simulated galaxies in clusters to date, making it especially suitable for verifying the properties of galaxy clusters. We obtained our data from snapshot number 135 (cosmological  $z = 0$ ) of the Illustris-1 simulation. The Illustris-1 simulation has the highest particle resolution and has incorporated the most comprehensive baryonic physics among the different Illustris simulation suites. The sophisticated galaxy formation model in Illustris-1 includes star formation rate, and stellar evolution due to environmental effects of the intracluster medium, such as ram pressure stripping and strangulation and feedback from Active Galactic Nuclei (AGN) etc. (Genel et al. 2014). The physics of stellar evolution were solved using a moving mesh code AREPO (Springel 2010). The observable properties of galaxies were verified to be statistically consistent with the Sloan Digital Sky Survey (SDSS) data (Vogelsberger et al. 2014a).

As the stellar population in Illustris were evolved from the initial condition, these makes the spatial distribution of galaxies in the Illustris data more realistic than galaxies that are prescribed

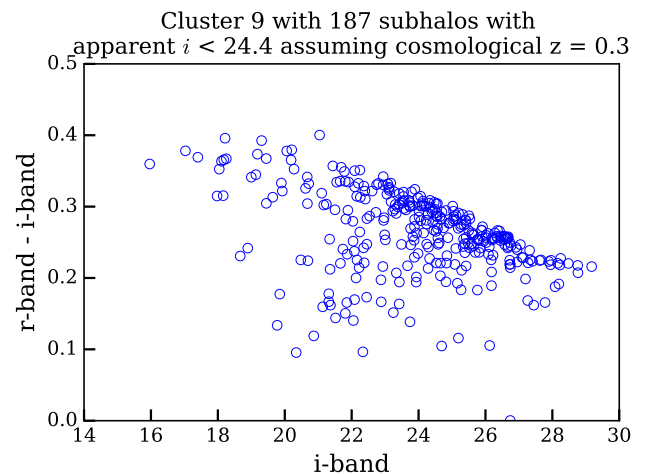


**Figure 1.** Left figure: Mass distribution of the group / cluster sized DM halos for different halo selection schemes. Mass estimates obtained by the FoF algorithm are labeled as  $M_{\text{FoF}}$ . Masses centered on the most bound particle within a radius those the average density is 200 or 500 times the critical density of the universe are labeled as  $M_{200c}$  and  $M_{500c}$  respectively. Right figure: Mass-richness relationship of galaxy clusters and groups with  $M_{\text{FoF}} > 10^{13} M_{\odot}$  assuming different cosmological redshifts of the observed clusters.

onto DM-only cosmological simulation data such as those used in Harvey et al. (2014). Gravitational effects in Illustris-1 have provided realistic dynamics and spatial distribution of subhalos. The simulated effects include tidal stripping, dynamical friction and merging. Since the profiles of the galaxies clusters were not provided in symmetrical, parametric forms, we can study how asymmetry in the cluster profile affects the estimate of our summary statistic. This data allows us to examine cluster galaxies in a realistic, yet noise-free way. The softening length of the DM particles is 1.4 kpc and those of the stellar particles is 0.7 kpc, both in constant comoving units (Genel et al. 2014).

The two sets of data catalogs in use are obtained through two types of halo finders. The catalog that maps particles to the halo of a certain cluster was created by the SUBFIND algorithm. The friends-of-friends (FoF) finder (Davis et al. 1985) was further used to identify the affinity of galaxy-sized halos to a galaxy-cluster. These galaxy-size halos are referred to as *subhalos* and they are the dark matter hosts of what we refer to as galaxies in Illustris-1. Vogelsberger et al. (2014b) also extracted the absolute magnitude of each subhalo in the SDSS bands of  $g$ ,  $r$ ,  $i$ ,  $z$  as part of the SUBFIND catalog using stellar population synthesis models.

For our analyses, we make use of galaxy clusters / groups with at least 50 member galaxies that are within a reasonable observational limit, i.e. apparent  $i \leq 24.4$  which is the limiting magnitude of the DEIMOS spectrometer on the Keck telescope when we assume a cosmological redshift of  $z = 0.3$  in the  $i$  band. The limiting magnitude of 24.2 in the F814W filter of the Hubble Space Telescope, and the limiting magnitude of the Canadian-Hawaii French Telescope of 24.5, are also close to our chosen limiting magnitude. There is relatively large statistical uncertainty if we try to analyze clusters with less than 50 member galaxies. As indicated by the right-hand panel of Fig. 1, a total of 43 clusters has survived this magnitude cut. These simulated galaxy clusters (or groups) have masses ranging from  $10^{13} M_{\odot}$  to  $10^{14} M_{\odot}$ .



**Figure 2.** Color-magnitude diagram of one of the galaxy clusters that is selected for analysis. This cluster is the 9th most massive. The apparent magnitude is calculated assuming that the cosmological redshift (distance) is  $z = 0.3$ . We can see a clear overdense region that corresponds to a red-sequence. The color-magnitude diagrams of the other clusters can be found in the Jupyter notebook at <https://goo.gl/TJmI6s>.

## 2.1 Cluster properties

### 2.1.1 Quantifying the dynamical states (relaxedness) of the galaxy clusters

Clusters undergo merger activities of a large range of physical scales and in the time scale of million of years. The dynamical history, or what we call “unrelaxedness”, cannot be directly computed from observations. We quantify the state of the cluster by providing several quantitative definitions of unrelaxedness and see how they correlate with  $\Delta s$ . Some possible definitions of unrelaxedness referred by the simulation community include:

- $\text{unrelaxedness}_0$ : the ratio of mass outside the dominant dark matter halo over the total mass of the galaxy cluster. The lower the ratio, the fewer substructures there are in the cluster.
- $\text{unrelaxedness}_1$ : the distance between the most bound particle from the center of mass as a function of  $R_{200c}$ , the three-dimensional (3D) radius in which the average density is 200 times the critical density of the universe. The smaller the distance, there are less asymmetric substructures.

To relate these simulation quantities to observation, we compute more observation-oriented quantities in the method section 3.0.2.

## 2.2 Selection of the field-of-view

We used SUBFIND for member particle identification of the DM and the FOF finder for subhalo identification. We understand that this choice of volume selection can be less restrictive than observational conditions. We make use of this volume selection scheme for baseline comparisons. Assuming a conservative line-of-sight (los) distance, i.e. cosmological redshift, with  $z = 0.3$ , the projected extent for most of the Illustris galaxy clusters and groups, fits inside the field of view of telescopes, such as the Subaru Suprime Camera, which covers a physical area of  $\sim 9 \text{ Mpc} \times 7 \text{ Mpc}$  (See <https://goo.gl/CIZNvM> for a Jupyter notebook showing the extent of the Dark Matter distribution of the most massive 129 clusters).

### 2.2.1 Spatial Projections

Unlike in staged simulation, picking out a particular projection for a cluster does not always make physical sense. For highly symmetrical clusters, most projections are similar. However, for mergers or asymmetrical clusters, there is no obvious choice for the plane of projection that can allow us to understand the cluster. Depending on the number of merging components, there may not be any simplistic merger axis that should be projected along the plane of the field of view.

We therefore compute observables based on even sampling of angular orientation as our line-of-sight. As the order of projecting the data and estimating the summary statistic is non-commutative, we first project the data before estimating any projected observable. The computation of even angular orientation is done by using HEALPY, which is a PYTHON wrapper for HEALPIX<sup>1</sup> (Gorski et al. 2005). Each line-of-sight centers on a HEALPIX pixel. The number of projections that we employed is 768 for each cluster. With around 768 projections, the offset distributions of each cluster start to converge to a stable distribution. Even though there are at least 2 identical projections for each cluster due to one possible line-of-sight from the front and one from the back, it does not affect any summary statistic. We do not remove the duplication as it breaks the rotational symmetry in the 2D plane when we try to compute the 2D population distribution of offsets.

## 2.3 Properties of the galaxies in Illustris clusters

Different galaxies have different masses, so they should not be considered with equal importance for peak identification, which requires summing the mass proxies of different galaxies. One of the most common weighting schemes employed for galaxy data is to weight by the luminosity in a particular band. For some of the

methods, we investigate the differences in peak identification with and without any luminosity weights. We pick the  $i$ -band magnitude associated with each subhalo as the weight. Since the  $i$ -band is one of the redder bands, the mass-to-light ratio is not skewed as much due to star formation activities. We further examined if the colors distribution of galaxies in Illustris-1 are similar to the observed color-magnitude diagrams for clusters. The Illustris cluster galaxies are realistic enough that it is easy to identify an overdense region of galaxies known as the red-sequence in the color-magnitude diagram such as Fig. 2. The red-sequence is prominent even if we use other colors formed by different combinations of the  $r, i, z$  bands.

## 3 METHODS

A common and the most precise way of summarizing the DM distribution in a galaxy cluster is by finding the lensing peaks (Medezinski et al. 2013, Markevitch et al. 2004, Zitrin et al. 2013). Additionally, the peak region is physically interesting due to the higher particle density and interaction rates. The most direct analogous statistic for summarizing the member galaxy population in a cluster is therefore, also the peak. Comparing the DM peak with the summary statistics of the galaxy population that are not the peak can have an *offset* purely due to the difference in the choice of the statistic. We will state the definition and implementation of the five commonly used point statistic or location for summarizing the member galaxy population in a galaxy cluster.

We avoid any manual methods for comparison purposes, scalability and reproducibility. Since all the methods listed in this paper are automated with the source code openly available, it is possible for future studies to reuse our code for comparisons. Another major advantage for automation is that it allows us to apply the same methods across the different snapshots of the (Illustris) simulations to examine the variability of  $\Delta s$  across time in future studies.

### 3.0.1 Computing the weighted centroid

We follow the usual definition of the weighted centroid:

$$\bar{\mathbf{x}}_w = \frac{\sum_i w_i \mathbf{x}_i}{\sum_i w_i}, \quad (3)$$

with  $\mathbf{x}_i$  being the positional vector of each subhalo and we use the  $i$ -band luminosity as the weight  $w_i$  for the  $i$ -th galaxy. Centroids can be biased by subcomponents from merging activities yet the centroid estimate does not provide explicit evidence for ongoing merger or accretion. These estimates are also sensitive to odd boundaries of the field of view.

### 3.0.2 Cross-validated Kernel Density Estimation (KDE) and the peak finder

Finding the exact peak of a set of data points involves computing the density estimate of the data points and sorting through the density estimates. A specific version of this density estimation process is known as histogramming. During the making of a histogram, each data point is given some weight using a tophat kernel and the weights are summed up at specific data locations (e.g.  $\mathbf{x}_i$ ). A histogram is not good for peak estimate for *sparse* data for two reasons: 1) the choice of laying down the bin boundaries affects the count in each bin, 2) the choice of bin width also affects the count in the bin. Only when the available number of data points for binning is large, the estimates of histograms and smoothed density estimates are approximately the

<sup>1</sup> HEALPix is currently hosted at <http://healpix.sourceforge.net>



**Table 1.** Selection criteria for stellar subhalos (member galaxies) for each cluster / group

Data	Selection strategy	Sensitivity	Relevant section
Field of view (FOV)	FoF halo finder	comparable to FOV of the Subaru Suprime camera	2.2
Observed filter	$i$ -band	consistent among the redder $r$ , $i$ , $z$ bands	2.3
Cluster richness	$i \leq 24.4$ and $z = 0.3$	sensitive to the assumed cosmological redshift of cluster and the assumed limiting magnitude of telescope	2
Two-dimensional projections	even HEALPix samples over half a sphere	discussed as results	2.2.1

same. The number of member galaxies ( $< 500$ ) is sparse enough for the uncertainty introduced by histogramming to bias our peak estimate. For the density estimate of galaxy luminosity, we adopt a Gaussian kernel. The exact choice of the functional form of the smoothing kernel does not dominate the density estimate as long as the chosen kernel is smooth (Feigelson & Babu 2014).

For computing the density estimate, the most important parameter of computing is the bandwidth of the smoothing kernel, which takes the form of a matrix in the 2D case. When the kernel width is too large (bottom left panel), the data is over-smoothed, resulting in a bias of the peak estimate. On the other hand, when the kernel width is too small, it results in high variances of the estimate and too many peaks due to noise. A good illustration can be seen in VanderPlas et al. 2012 from [http://www.astroml.org/book\\_figures/chapter6/fig\\_hist\\_to\\_kernel.html](http://www.astroml.org/book_figures/chapter6/fig_hist_to_kernel.html). The decision of having to balance between creating high bias or high variance estimates is also known as the bias-variance tradeoff. Any other smoothing procedures, including interpolation with splines, polynomials, and filter convolutions, also face the same tradeoff.

A well-known way to minimize the fitting error from the density estimate is through a data-based approach called cross-validation to obtain the optimal 2D smoothing bandwidth matrix ( $H$ ) of the 2D Gaussian kernel for the density estimate  $\hat{f}$ :

$$\hat{f}(\chi; H) = \frac{1}{n} \frac{1}{(2\pi)^{d/2} |H|^{1/2}} \sum_{i=1}^n w_i \exp((\chi - \mathbf{x}_i)^T H^{-1} (\chi - \mathbf{x}_i)), \quad (4)$$

where the dimensionality is  $d = 2$  for our projected quantities,  $\chi$  represents the uniform grid points for evaluation, and  $\mathbf{x}_i$  contains the spatial coordinates for each of the identified member galaxies that survived our brightness cut and  $w_i$  is again the  $i$ -band luminosity weights for each galaxy. The idea behind cross-validation is to leave a small fraction of data point out as the test set, and use the rest of the data points as the training set for computing the estimated density. Then it is possible to estimate and minimize the asymptotic mean-integrated squared error (AMISE) by searching for the best set of bandwidth matrix values, eliminating any free parameter.

Specifically, we made use of the smoothed-cross validation (Hall et al. 1992) bandwidth selector in the statistical package KS (Duong 2007) in the R statistical computing environment (R Core Team 2014). Among all the different R packages, KS is the only package capable of handling the magnitude weights of the data points while inferring the density estimates (Deng & Wickham 2011). Although the particular implementation of KDE has a computational runtime of  $O(n^2)$ , the number of cluster galaxies is small enough for this method to finish quickly ( $\lesssim 0.65$  second per projection per cluster).

The resulting KDE contains rich information about the spatial distribution of the clusters, and we focus on the peak regions. We employed both a first and second-order finite differencing algorithm to find the local maxima. The local maxima were then sorted

according to the KDE density in a descending fashion before we perform peak matching and compute the offset. The exact procedure is discussed in section 3.3.

For each projection of each cluster, we normalize the density of all luminosity peaks to those of the brightest peak. Luminosity peaks that sit on top of actual subclusters would then have a density comparable to those of the brightest peak. Then we sum the density of all the galaxy peaks for a cluster and call this value  $\nu$ . When the value of  $\nu$  much gets bigger than 1, it indicates the presence of projected substructure(s). Even though  $\nu$  is not expressed in terms of masses, it is a practical measurement for optical survey data those galaxy magnitudes, but not the spatial mass distribution from lensing estimates, are available.

### 3.0.3 Shrinking aperture estimates

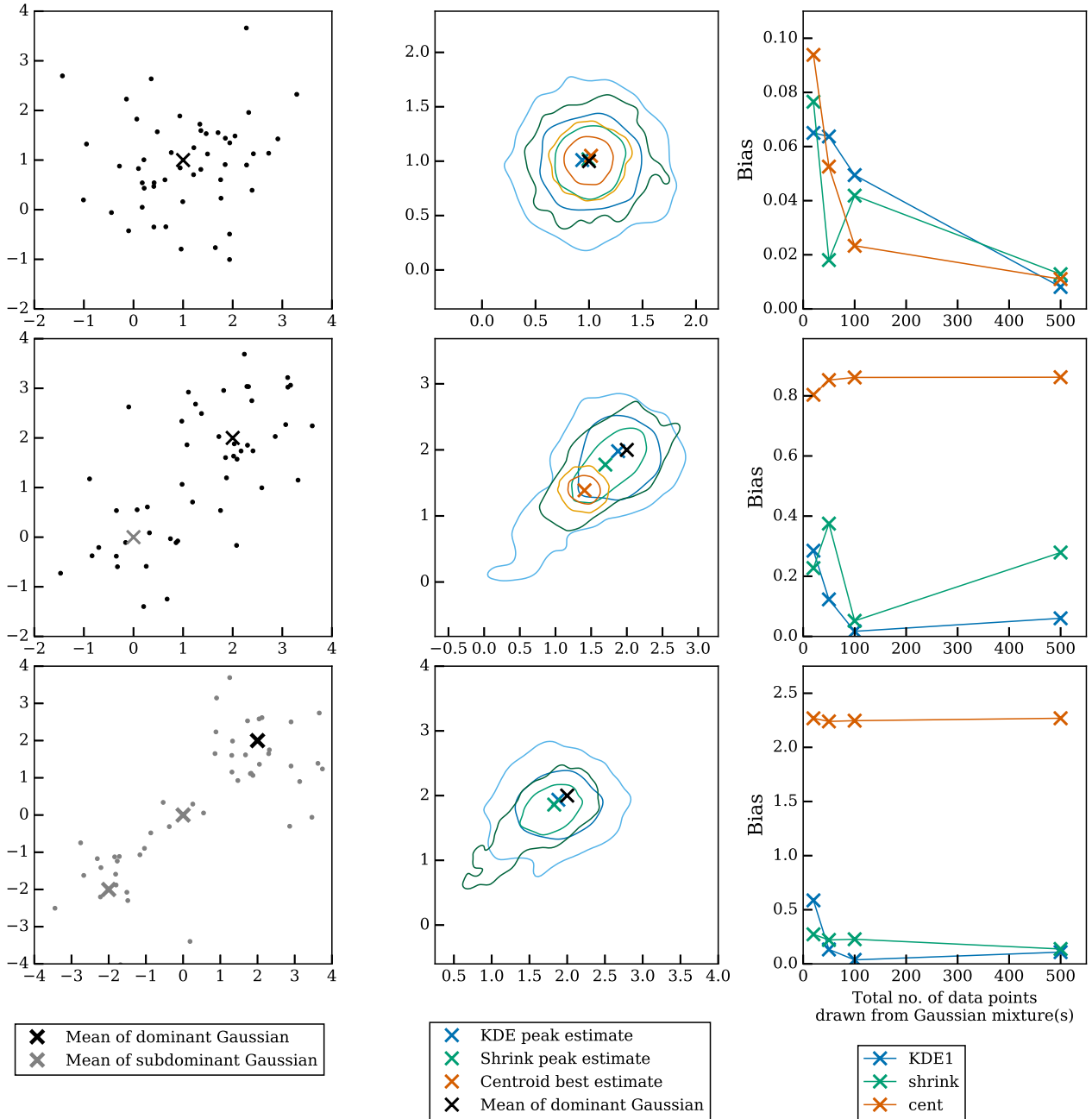
Another popular method among astronomers for finding the peak of a spatial distribution is what we call the shrinking aperture method. While we do not endorse this method, we test if the shrinking aperture method is able to reliably recover the peak of the luminosity map. This method is dependent on the initial diameter and the initial center location of the aperture. This method does not evaluate if the cluster is made up of several components. The estimate using the shrinking aperture algorithm can be biased by substructures. The only way to inform the algorithm about substructures would be to introduce another parameter to restrict the extent of the aperture, or to partition the data with another (statistical) algorithm. More to the point, the convergence of results of this method is unstable. We use a convergence criterion of having the aperture distance not change more than 2% between successive iterations as a reference. The actual implementation in PYTHON can be found at <https://goo.gl/nqxJl8> while the pseudo-code can be found in Appendix A.

### 3.0.4 Brightest Cluster Galaxies (BCG)

The BCGs are formed by the merger of many smaller galaxies. The galaxy-cannibalism makes BCGs typically brighter than the rest of the cluster galaxy population by several orders of magnitude. However, star formation can cause less massive galaxies to be brighter in the bluer photometric bands. To avoid star formation from biasing our algorithm for identifying the BCG, we find the brightest galaxies in redder bands i.e. the  $r$ ,  $i$ ,  $z$  bands and found that they give consistent results for all selected clusters. We used the  $i$ -band to pick the BCG for computing the plots and the final results.

## 3.1 Comparison of the methods from Gaussian mixture data

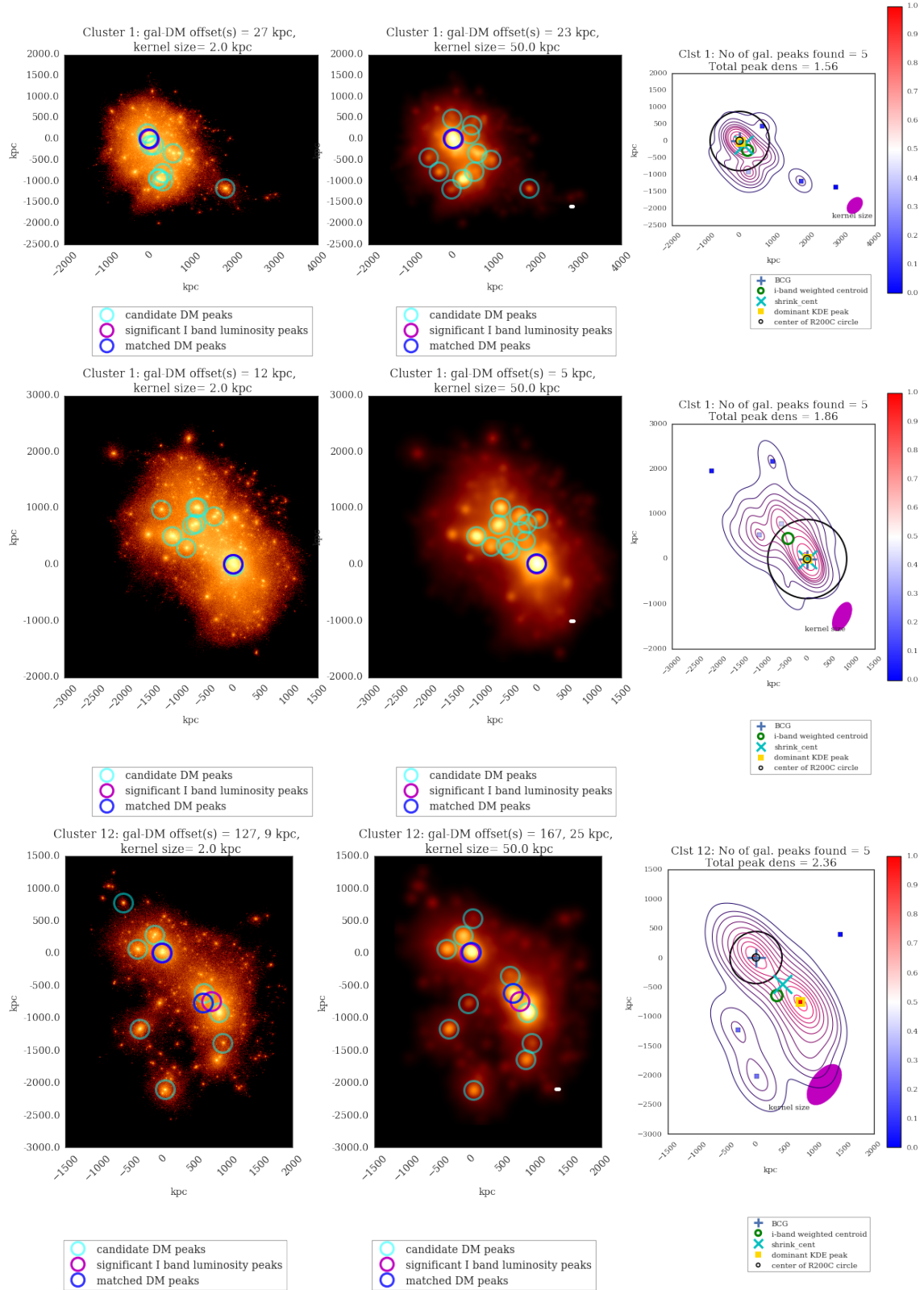
In order to examine the statistical properties of commonly used point-estimates of the distribution of the galaxy data, we test them on data drawn from Gaussian mixtures with known mean and variance. (See Fig. 3). The main factors that affect the performance of the methods



**Figure 3.** Comparison of peak finding performances of the shrinking aperture peak (shrink peak), KDE peak estimate, and centroid, by drawing data points (i.e. 20, 50, 100, 500) from known number of Gaussian mixtures. Panels from the top row contain data drawn from a single Gaussian mixture. The panels from the middle row contain data from two Gaussian mixtures with weight ratio = 7:3. The panels from the bottom row contain data drawn from three Gaussian mixtures with weight ratio = 55:35:10. The left column shows how 50 data points drawn from the fixed number of Gaussian mixtures look like. Due to the statistical nature of this exercise, we sampled the data and performed the analyses 50 times to create the 68% and 95% Monte Carlo confidence contours of the estimates in the zoomed-in view of the data in the middle column. The rightmost column shows how population bias vary as a function of the number of drawn data points from the Gaussian mixtures. From the middle and the rightmost column, we can tell that the cross-validated KDE peak estimate is the most accurate when there is more than one significant component.

are sensitive to the statistical fluctuations of the drawn data, e.g. the spatial distribution of the data, including 1) the density profile and 2) the location(s) of subdominant mixtures, and 3) the number of data points that we draw. It is also not enough to just compare the performance by applying each method for one realization of the data. We provide the 68% and the 95% confidence regions by applying the each method for many Monte Carlo realizations. In general, the

peaks identified from the KDE density is closer to the peak of the dominant mixture (more accurate) than both the weighted centroid method and the shrinking aperture method. For example, in the bottom middle panel of fig. 3, it is clear that the green contours that represents the confidence region for the shrinking aperture peak is biased due to the substructure, whereas the confidence region for the centroid is so biased that it is outside the field of view of



**Figure 4.** Visualization of clusters (each row is for the same projection of the same cluster). **Left column:** Projected density distribution of DM particle data is shown in orange, with the dense regions in yellow. The identified density peaks are indicated by colored circles. **Middle column:** The same DM projection after smoothing with a 50 kpc smoothing kernel (kernel size is indicated by a white dot on lower right of the panel). The thickness of the dot may be larger than 2 kpc for the plots on left hand column). **Right column:** Projected galaxy kernel density estimates (KDE) of the  $i$ -band luminosity map for the member galaxies of the same clusters. Each colored contour denotes a 10% drop in density mass starting from the highest level in red. Each of the magenta ellipse on the bottom right corner of each plot show the Gaussian kernel matrix  $H$  from eq. (4). The big black circle is centered on the most bound particle as identified by **SUBFIND** and the radius of the circle indicates the  $R_{200C}$ . The luminosity peaks (square markers) are colored by the relative density to the densest peak, the relative density is shown by the color bar. See <http://goo.gl/WiDijQ> and <http://goo.gl/89edcM> for the visualization of the selected clusters inside two Jupyter notebooks.

that panel. For right panel of fig. 3, we present how the population bias of each method shrinks as the number of data points increases. For the data generated with more than one mixture, KDE peak consistently show less population bias than the shrinking aperture method. The performance of the shrinking aperture method fluctuate and is unstable when the number of data points is increased.

### 3.2 Modeling the DM map in Illustris-1 and the lensing kernel

The most well established method of inferring the projected dark matter spatial distribution from observations is through gravitational lensing. It works by detecting subtle image distortions of background galaxies due to the foreground dark matter. The resolution of the inferred map therefore depends on the properties of the source galaxies that are being lensed, such as the projected number density, the intrinsic ellipticities and morphology etc. To achieve a sufficient signal-to-noise ratio for lensing, Hoag et al. (2016) has performed simulation for inferring the optimal bandwidth for a Gaussian smoothing kernel for the cluster MACSJ0416. In the strong lensing regime, Hoag et al. (2016) found a resolution of 11 arcseconds can best fit the MACSJ0416 data. A kernel bandwidth (this is the standard deviation) of 11 arcseconds translates to an angular diameter distance of 50 kpc assuming a cosmological redshift of  $z \approx 0.3$ . In order to match the resolution of lensing data, we also employed a smoothing kernel of a similar physical size of 50 kpc.

To compute the DM spatial distribution from our data, we first make histogram with  $2 \text{ kpc} \times 2 \text{ kpc}$  bin size which is slightly larger than the DM softening length of 1.4 kpc. After that, we use a (50, 50) kpc 2D radial Gaussian kernel to smooth the DM histogram Illustris DM particle data. There are resolution differences between the smoothed and unsmoothed DM maps. The unsmoothed histograms tend to show many more local maxima around the major density peaks (i.e. show high variance). The number of DM particles for each cluster is of the order of millions and densely packed in the region of interest. Physically, the smoothed histograms of the dark matter of each cluster is analogous to a convergence map from a lensing analysis.

### 3.3 Finding the offsets

It is possible to have several peak estimates from the KDE of the member galaxy population of a cluster. From the density estimate at each peak, we can sort the peaks according to their densities. We only match luminosity peaks that are at least 20% as dense as the brightest galaxy-luminosity peak to avoid computing the offsets of spurious substructures, such as the peaks due to small number of galaxies that are located far away from the main concentration of mass.

In general, there are many more DM peaks because there are many more dark subhalos than galaxies for each cluster and the resolution of the DM data is much higher. We write our peak matching algorithm to mimic how humans would find the closest DM peak to the galaxies peak. This is done using a k-dimensional tree (KD-Tree; in our case,  $k = 2$ ) using the densest  $n_{\text{DM}}$  number of DM peaks:

$$n_{\text{DM}} = \begin{cases} 3 \times (n_{\text{gal}} + 1) & \text{if } n_{\text{gal}} < 3 \\ 3 \times n_{\text{gal}} & \text{if } n_{\text{gal}} \geq 3. \end{cases} \quad (5)$$

where  $n_{\text{gal}}$  is the number of significant galaxy peak, and  $n_{\text{DM}}$  is the number of peaks that went into the construction of the KD-tree. When there are more than one dense galaxy peaks located far away

from one another, the top few densest DM peaks (subhalos) can be located around the same galaxy peak. i.e. there is no one-to-one matching between the luminosity of galaxies and the density of detected DM peaks. Matching purely based on density and luminosity leads to larger offsets. From inspection of figures similar to Fig. 4, using eq. (5) works well to match the appropriate peaks. To avoid peak matching from affecting the offset results, we will only use the offsets with the dominant DM peak for the final results. We show in the result section that most of the chosen DM peaks do not have significant deviation from the most gravitationally bound particles. After identifying the DM peaks, we compute the offsets between the matched DM peaks, and the following spatial estimates, including

- the most (gravitationally) bound particle
- the shrinking aperture peaks, the corresponding offset is  $\Delta s_{\text{shrink}}$ ,
- the number density peaks, the corresponding offset is  $\Delta s_{\text{num.dens}}$ ,
- the BCGs, the corresponding offset is  $\Delta s_{\text{BCG}}$  and
- the luminosity weighted centroid, the corresponding offset is  $\Delta s_{\text{BCG}}$ .

Since there can be more than one number density peak from the corresponding KDE map, we also use a KD-tree to location the closest number density peak to the identified DM peak.

### 3.4 Constructing the hypothesis test

After matching the peaks, we use the offsets as the basis of our hypothesis test. We compute the p-value as the narrowest density interval of simulated offsets that are below observed values of offsets in the literature. The distribution of offsets in the Illustris simulation represents the possible ways that offsets would be observed in a CDM universe. This gives us a rough estimation of the probability of seeing the offset from observations under the null hypothesis of CDM being true. We report the percentile for the offset distributions, and also the biweight statistical estimates such as the biweight location (analogous to the median) and the midvariance (robust standard deviation estimate). The biweight statistics are less susceptible to the effects of outliers (Beers et al. 1990). We compute the robust statistics using the implementation from Robitaille et al. (2013) as part of ASTROPY. For instance, the 95% interval is computed as the interval that encompasses 95% of total density (2.5% of density mass at each end of the tail is excluded). In case of degeneracy, the interval is also required to cover the location estimate for the distribution.

The representations of the distributions of  $\Delta s$  carry different information and allow different types of statistical tests. The most faithful representation of the offsets without any information loss is:

$$\Delta \mathbf{s} = (\mathbf{x}_{\text{gal}} - \mathbf{x}_{\text{DM}}, \mathbf{y}_{\text{gal}} - \mathbf{y}_{\text{DM}}). \quad (6)$$

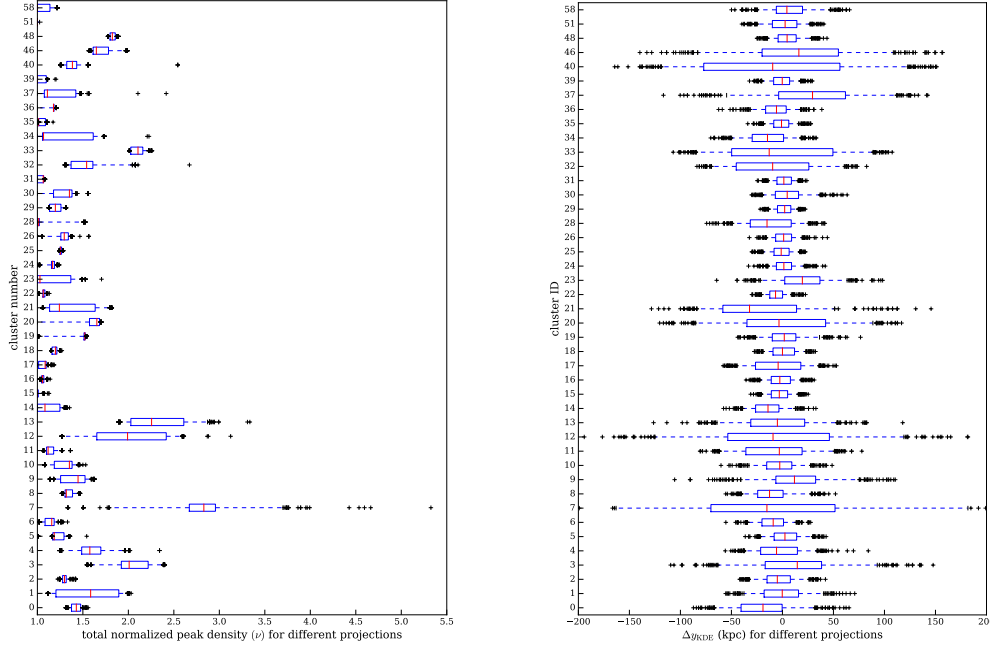
The PDF of  $\Delta \mathbf{s}$  in eq. 6 peaks at (0, 0) when there is no real offset. It is also possible to do directionality tests, such as Rayleigh z test, to see if the data points have a preference to land in a certain direction. However, when one takes the magnitude of  $\Delta \mathbf{s}$ , i.e.:

$$|\Delta \mathbf{s}| = \sqrt{(\mathbf{x}_{\text{gal}} - \mathbf{x}_{\text{DM}})^2 + (\mathbf{y}_{\text{gal}} - \mathbf{y}_{\text{DM}})^2}, \quad (7)$$

the resulting 1D distribution of  $|\Delta \mathbf{s}|$ , those support being  $[0, \infty)$ , will not peak at zero even if the original distribution of  $\Delta \mathbf{s}$  peaks at (0, 0). To illustrate how difficult it is to interpret the magnitude values  $|\Delta \mathbf{s}|$ , we can imagine the following transformation. The values drawn from:

$$\begin{pmatrix} \mathbf{x} \\ \mathbf{y} \end{pmatrix} \sim \mathcal{N} \left( \begin{pmatrix} 0 \\ 0 \end{pmatrix}, \begin{pmatrix} \sigma^2 & 0 \\ 0 & \sigma^2 \end{pmatrix} \right), \quad (8)$$





**Figure 5.** Left: A box plot showing the distribution of the total normalized peak density ( $\nu$ ) for each cluster. Clusters with only one dominant peak has  $\nu = 1.0$ , values bigger than 1 indicate density fraction contributed by other peaks. Right: A box plot showing the distribution of  $\Delta y_{\text{KDE}}$  for each cluster based on 768 projections. The offsets were computed between the closest DM peak to the brightest luminosity peak of each cluster. The red line of each box shows the median of the projections, the box encompasses the 25% and 75% percentile of the distribution while the whiskers mark the 5% and the 95% percentile. The other black crosses are data points with extreme values beyond the 5% and 95% percentile.

will result in  $|\Delta s|$  values that follow the Rayleigh distribution:

$$f(\Delta s|\sigma) = \Delta s/\sigma^2 \exp(-\Delta s^2/2\sigma^2) \quad (9)$$

those peak is at  $|\Delta s| = \sigma$ , the same standard deviation value of the 2D Gaussian. Note that this representation rules out any probability mass at  $|\Delta s| = 0$  by construction. The dependency of  $|\Delta s|$  on the parameters of the 2D distribution is even more complicated when the 2D distribution does not approximate a Gaussian or when there is more than one peak in the 2D space. The shifting of the peak location due to variable transformation is seen in the distribution of  $|\Delta s|$  recorded in table B3. For an asymmetrical 1D distribution due to taking the magnitude, finding the narrowest 68% and 95% interval is not a standard statistical procedure and can be prone to error.

On the other hand, the 1D distributions of offsets along a particular spatial axis, e.g.  $\Delta x$  and  $\Delta y$ , each with a support of  $\mathbb{R}$ , will not exhibit a discontinuity at zero. Any shift or asymmetry in the 2D peak location is still obvious. The distributions represented by  $\Delta x$  or  $\Delta y$  can be symmetrical so the highest density interval is easier to find. Since we have enough samples for there to be rotational symmetry for the distribution of  $(\Delta x, \Delta y)$ , we will show that it does not matter much if we picked  $\Delta x$  or  $\Delta y$  for the 1D representation. We compute the hypothesis test significance level with the offset  $\Delta x$  along one of the spatial axes. To report the statistics, we also make use of estimates that do not make any underlying assumption of the shape of the distributions that may skew the statistical parameter estimate. In fact, several studies have reported poor single 2D Gaussian fits to 2D offset data due to the long tails (Zitrin et al. 2012, Oguri et al. 2010).

In the following sections of this paper, we use  $\Delta s$  to represent the two-dimensional offsets,  $|\Delta s|$  for the magnitude of the offset as calculated according to the Euclidean distance, and  $\Delta x$  or  $\Delta y$  to denote the one-dimensional offset along one of the spatial dimen-

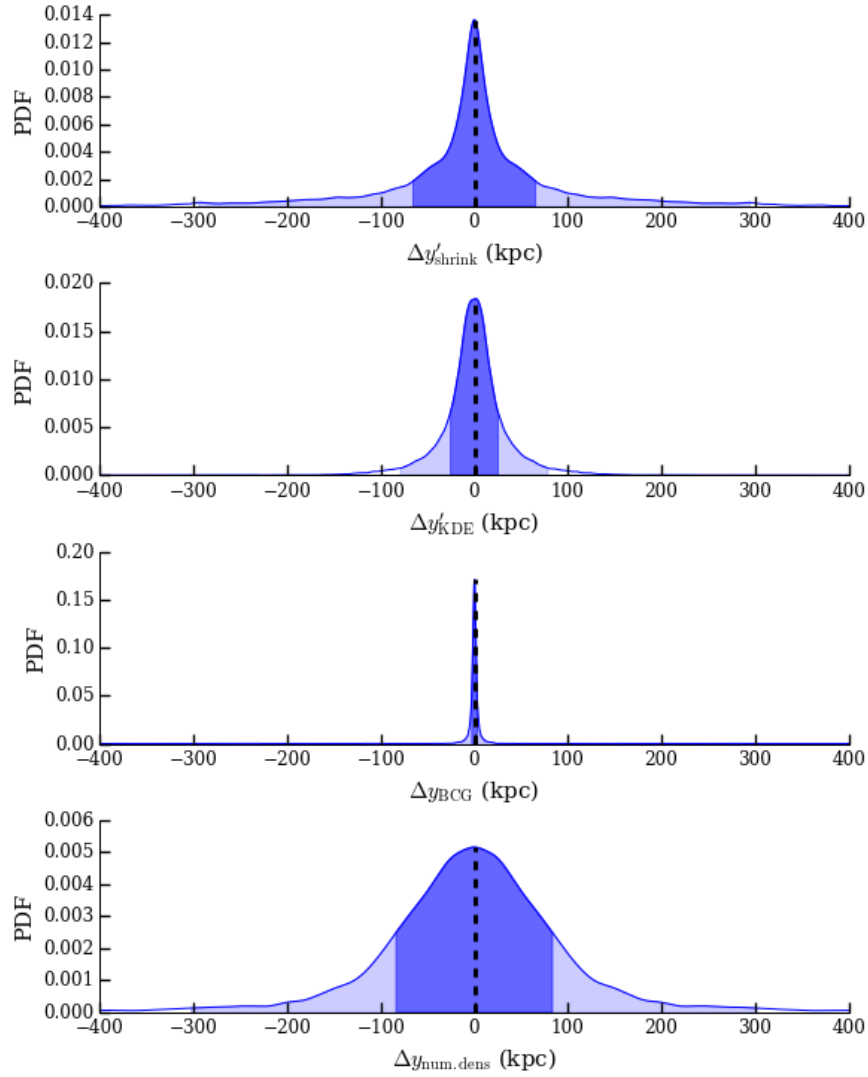
sions. To compare with observed data, we estimate the 1D spatial components of the offsets from the merging cluster observations from various sources. We make our best attempt to measure  $\Delta y_{\text{obs}}$ , the spatial component of the observed offset along the axis connecting the subclusters if subcomponents exist. In our observed samples, Abell 3827 is the only exception that has no subclusters but only four bright galaxy peaks in the central region. We also show  $\Delta x_{\text{obs}}$  components, the offset perpendicular to  $\Delta y_{\text{obs}}$  for comparison. For most of the observed offsets, we obtain the estimates from the contour plots and descriptions of the corresponding papers. For the offsets that are roughly in line with the axis connecting the two subclusters, we let  $\Delta y = |\Delta s|$ . If  $\Delta s_{\text{obs}}$  is not aligned along the line joining the subclusters, using  $|\Delta s_{\text{obs}}|$  instead of  $\Delta y_{\text{obs}}$  to come up with a p-value from the distribution of  $\Delta y$  will lead to a spurious increase in significance.

## 4 RESULTS

### 4.1 The dynamical states (relaxedness) of the clusters

Out of the  $43 \times 768 = 33024$  projections,  $\sim 45\%$  of the projections have one dominant luminosity peak and negligible substructures, with the total peak density of the projection being  $\nu \leq 1.2$ . Another  $\sim 50\%$  of the projections have more than one dominant luminosity peak with  $1.2 < \nu < 2.2$ . Fig. 5 that shows the spread of  $\nu$  estimate over different projections per cluster indicates most clusters show signs of other subdominant peaks.

Visually, the spread of the  $\nu$  distribution is indicated by the horizontal length of the blue box. the median of  $\nu$  per cluster are indicated by the red central vertical line inside each box in Fig. 5. Only 7 clusters (with ID = 15, 16, 17, 22, 31, 35, 51) out of 43 clusters have  $\nu \lesssim 1.2$  for most of the projections. Clusters with



**Figure 6.** The smoothed distribution of different offsets of 43 clusters with all 768 projections. The smoothing bandwidth is determined by Scott’s rule for visualization. For estimates where several peaks of galaxy data are possible, only the densest peak is matched to the DM peak for computing the offsets in this figure. The dark blue area indicates the 68% density interval while the light blue area shows the 95% density interval. The table summarizing the statistic of each distribution is available in table 2.

**Table 2.** Robust estimates and the distribution of offsets along the y-axis (This is different from the magnitude which has discontinuity at zero).

sample	offset (kpc)	location	lower 68%	lower 95%	lower 99%	upper 68%	upper 95%	upper 99%
all $\nu$	$\Delta y_{\text{BCG}}$	0	-3	-22	-496	3	456	1449
all $\nu$	$\Delta y'_{\text{KDE}}$	0	-25	-79	-127	25	79	126
all $\nu$	$\Delta y_{\text{num.dens}}$	0	-84	-303	-693	84	302	691
all $\nu$	$\Delta y'_{\text{shrink}}$	0	-65	-295	-652	65	295	655
$\nu < 1.2$	$\Delta y_{\text{BCG}}$	0	-3	-10	-19	2	9	19
$\nu < 1.2$	$\Delta y'_{\text{KDE}}$	0	-18	-48	-82	18	48	83
$\nu < 1.2$	$\Delta y_{\text{centroid}}$	0	-108	-255	-395	108	254	394
$\nu < 1.2$	$\Delta y_{\text{num.dens}}$	0	-73	-195	-303	73	195	302
$\nu < 1.2$	$\Delta y'_{\text{shrink}}$	0	-51	-187	-285	51	187	285
$1.2 < \nu < 2.2$	$\Delta y_{\text{BCG}}$	0	-3	-160	-684	4	807	1570
$1.2 < \nu < 2.2$	$\Delta y'_{\text{KDE}}$	0	-32	-89	-125	32	89	124
$1.2 < \nu < 2.2$	$\Delta y_{\text{centroid}}$	0	-262	-663	-905	262	663	904
$1.2 < \nu < 2.2$	$\Delta y_{\text{num.dens}}$	0	-87	-299	-739	87	298	738
$1.2 < \nu < 2.2$	$\Delta y'_{\text{shrink}}$	0	-85	-386	-777	85	386	779

**Table 3.** Observed offsets from clusters with reported evidence of mergers along line connecting two subclusters ( $\Delta y$ ) and the approximate perpendicular offset ( $\Delta x$ ). The table mainly contains clusters that have been used to constrain  $\sigma_{\text{SIDM}}$  using the reported offsets. Any approximate error estimates are the corresponding 68% lensing peak uncertainty in the figure(s) of the references, this is due to the lack of uncertainty estimates from the galaxy summary statistics from most literature. Error estimates are omitted when they are not reported by the authors. All p-value lower bounds are reported by matching to the corresponding method for estimating galaxy summary statistic in table 2.

Cluster	$\Delta y$ (kpc)	$\Delta x$ (kpc)	$ \Delta s $ (kpc)	galaxy peak	DM peak	p-value	subcluster	mass ( $10^{14} M_{\odot}$ )	reference
Bullet	9	-23	$25 \pm 29$	num. or lum.	SL & WL	$\gtrsim 0.32$	northwest	1.5	Randall et al. 2008
Baby Bullet	-40	0	$\sim 40 \pm \sim 50$	lum.	SL & WL	$\gtrsim 0.05$	northwest	2.6	Bradač et al. 2008:Fig.4
Baby Bullet	30	0	$\sim 30 \pm \sim 75$	lum.		$\gtrsim 0.32$	southeast	2.5	Bradač et al. 2008:Fig.4
Musketball	129	0	$129 \pm \sim 63$	num.	WL	$> 0.05$	southern	3.1	Dawson 2013a:Fig.4.7
Musketball	-47	0	$47 \pm \sim 50$	num.		$> 0.32$	northern	1.7	Dawson 2013a:Fig.4.7
Abell 3827	6	0	6	BCG	SL	$> 0.05$	central		Williams & Saha 2011
Abell 520	0	50	$\sim 50 \pm \sim 50$	lum.	WL	$\gtrsim 0.32$	blue	5.7	Clowe et al. 2012:Fig. 4
El Gordo	58	0	$\sim 58 \pm \sim 100$	lum.	WL	$> 0.05$	northwest	11	Jee et al. 2014:Fig.7,8
El Gordo	30	110	$115 \pm \sim 60$	num.		$> 0.32$	northwest		Jee et al. 2014:Fig.7,8
El Gordo	6	25	$\sim 26 \pm \sim 50$	lum.		$> 0.32$	northwest	7.9	Jee et al. 2014:Fig.7, 8
El Gordo	280	280	$400 \pm \sim 40$	num.		$> 0.05$	southeast		Jee et al. 2014:Fig.7, 8
Sausage	160	100	$\sim 190 \pm \sim 150$	num.	WL	$\gtrsim 0.05$	north	11.	Jee et al. 2015:Fig.10
Sausage	160	160	$\sim 190 \pm \sim 150$	num.		$\gtrsim 0.05$	south	9.8	Jee et al. 2015:Fig.10
Sausage	320	130	$\sim 340 \pm \sim 150$	lum.		$\gtrsim 0.01$	north	11.	Jee et al. 2015:Fig.10
Sausage	160	160	$\sim 230 \pm \sim 150$	lum.		$\lesssim 0.01$	south	9.8	Jee et al. 2015:Fig.10

num. is a short hand for the peak estimate from the number density map.

lum. is a short hand for the peak estimate from the luminosity density map, or KDE<sup>2</sup> in the method description.

SL is a short hand for strong lensing.

WL is a short hand for weak lensing.

median values of  $\nu > 2.2$  usually have multiple subclusters. Cluster with ID = 7, for instance, is made up of around 4 disconnected clusters that span several Mpc. There is also a strong correlation of  $\sim 0.8$  between  $\nu$  and each of the two unrelaxedness quantities defined in section 2.1.1. This shows that  $\nu$  can be a good indicator for the dynamical states of the cluster.

#### 4.2 Offset between the matched DM peaks and the corresponding most bound particle

There is no significant differences between the matched DM peaks and the most gravitationally bound particle (hereafter most bound particle). The median of the offset between the DM peak and the gravitationally bound particle is (0, 0) kpc. The 75-th percentile of the offsets are at  $(\pm 2, \pm 2)$  kpc. Most of the other offset values occur below  $(\pm 9, \pm 9)$ . Large offsets are only seen for clusters with  $\nu > 1.2$ . The densest DM peak in 3D where the most bound particle is located, does not necessarily correspond the densest projected peak in 2D in the presence of significant DM substructures.

#### 4.3 Offset between galaxy summary statistic and the most bound particle

As another sanity check, we computed the offsets between different galaxy summary statistic and the most bound particle. Interested reader can refer to table B2 for the different percentile and robust estimates of the distribution of offsets from the most bound particle. The ranking in terms of increasing distance to the most bound particle computed by different method is as follows:

- the BCG
- the densest peak of the luminosity map created by weighted the KDE
- the shrinking aperture center from the luminosity weighted galaxy data
- the densest peak of the number density map created by the un-weighted KDE

- the centroid estimate using luminosity weights, which is a proxy for the center of mass

In fact, most of the BCG offsets are very small except for two clusters with ID 13 and 33. Both clusters have all the values of  $\nu > 1.5$  over different projections. From the projected density map, we further confirm that both clusters have significant substructures. It is therefore possible for the most bound particle to have a similar gravitational potential level as another substructure where the BCG is located. In general, the offset distributions between the galaxy summary statistics and the most bound particle have approximately the same level of variance but more extreme outliers (at the 99%) than the offset distribution between the DM peaks and the corresponding galaxy summary statistics.

#### 4.4 Galaxy-DM Offset in Illustris

##### 4.4.1 The two-dimensional distribution and distribution of $\Delta y$

The 2D distribution of  $\Delta s$  from most methods peak at around zero ( $\lesssim 4$  kpc) with high rotational symmetry. The possible sources of offset asymmetry come from clusters with unusual configurations, those clusters with more distinct, spatially separated subcomponents. They are the clusters with higher offset variance over different projections. Their outlier offsets are in general more extreme than other clusters. Since there are a limited number of projections, clusters with such unusual configurations and offset levels can leave patterns outside the main concentration of 2D offset data, such as rings. This is due to projections that give similar views of the cluster. The luminosity weighted centroid offset, i.e. median( $\Delta x'_{\text{centroid}}$ ) = -37 kpc, has the highest asymmetry along x-axis for the peak. These offsets distributions are recorded in detail in table B3.

The method that gives the tightest offset is the BCG, consistent with what we found in the previous section. The 2D offset  $\Delta s_{\text{BCG}}$  has most of its density located near zero ( $\pm 3$  kpc) but contains outliers. Having outliers is possible as the DM peak is chosen as the closest DM peak to match the brightest luminosity peak in a

particular projection. See the bottom right panel of 4, the BCG coincides with the most bound particle. However, the luminosity peak of the cluster is located at the other mass substructure. When there are distantly separated subclusters of similar masses, the brightest projected luminosity peak may shift from one subcluster to another subcluster between different projections, while the BCG identification is unchanged between projections. The aforementioned bias from substructure can be seen when we compare the offset estimates between the relatively relaxed sample of  $\nu < 1.2$  and the unrelaxed samples  $1.2 < \nu < 2.2$ . The 99-th percentile increased drastically from  $\pm 19$  kpc to an asymmetrical extreme estimates of  $(-684, +1570)$  kpc. Again, these values are possible because there can be several DM peaks of similar density due to subclusters located far apart from one another. The finite number of projections, combined with the substructures, have caused the 95-th and 99-th percentile tails of  $\Delta y_{\text{BCG}}$  of both the full sample and the unrelaxed sample exhibit noticeable asymmetry, but not the relaxed samples.

It is noteworthy that the population spread of  $\Delta s$  computed by each method differ enough that one needs to know which offset method an astronomical study used for a fair comparison. For the full sample in table 2 and fig. 6, the offsets computed by the peak from the luminosity weighted KDE has the second smallest variance. The 68-th percentile of  $\Delta y'_{\text{KDE}}$  is at  $\pm 25$  kpc. Using shrinking aperture to estimate the peak location from the luminosity map increased the 68-th percentile of the offset to more than double those of  $\Delta y'_{\text{KDE}}$  at  $\pm 65$  kpc. The peak estimate from the number density map has even larger variance, with its 68-th percentile being  $\pm 84$  kpc.

Most of the percentile intervals of the unrelaxed samples  $1.2 < \nu < 2.2$ , when compared to the relaxed samples, are around a factor of 2 larger. Among the relatively relaxed samples, the variance of the inferred offsets from different methods still show significant discrepancies. The variance of the offset computed from the shrinking aperture method, the number density map, and the weighted centroid are still at least a factor of 1.5 larger than those computed using the luminosity-weighted KDE. In particular, the 68% percentile of the centroid method is  $\pm 108$  kpc. This is around one-fourth the typical core radius of massive clusters (Allen 1998). Our centroid estimates can be more extreme than observations because we do not restrict the field of view. The spread of the offsets inferred by each method affects their ability for constraining  $\sigma_{\text{SIDM}}$ . We will further elaborate on this point when we compare our results with staged simulations of SIDM in section 5.4.

#### 4.5 Offset projection uncertainty of each cluster

When we gather the offsets  $\Delta s'_{\text{KDE}}$  of the 768 projections for each cluster, we can find the offset uncertainty due to projection effects. The distributed are illustrated in the box plot of Fig. 5. The values of the biweight mid-variance of  $\Delta y'_{\text{KDE}}$  for half of the clusters are  $< 23$  kpc. Of the ten clusters (ID = 3, 7, 12, 20, 21, 32, 33, 37, 40 and 46) that have mid-variance  $> 40$  kpc, all of them have the median of  $\nu > 1.2$ .

#### 4.6 Correlations between different variables and the offsets

Here we investigate a list of physical quantities that have significant to little correlation with the offsets. We use the Pearson product-moment correlation coefficient to quantify linear relationship between the pairs of variables (aka Pearson's  $r$ , hereafter  $\rho$ ). We describe the significance of the correlation based on the p-value reported by SCIPY of seeing the level of correlation by chance assuming the pair of quantities has no correlation. If the p-value is greater

than 0.1, we call the correlation as insignificant. As a reference, the correlation of between the unrelaxedness criteria defined in section 2.1.1 for the 43 selected clusters is as high as 0.82.

Each of the two unrelaxedness criteria has significant a positive correlation of  $\sim 0.70$  with the maximum of  $\Delta s'_{\text{KDE}}$  of each cluster (hereafter  $\max(\Delta s'_{\text{KDE}})$ ). The offset  $\max(\Delta s'_{\text{KDE}})$  per cluster also show a high correlation of 0.77 with the median  $\nu$  per cluster (hereafter,  $\text{median}(\nu)$ ). The FoF mass of each cluster shows only a slight correlation of 0.28 with  $\text{median}(\nu)$ .

There is a weak correlation between the richness of the clusters with  $\max(\Delta s_{\text{KDE}})$  ( $\rho = 0.21$ ). This weak correlation may be due to the fact that the peak estimate is only affected strongly by a few bright galaxies near the peak. The richness is slightly more strongly correlated with the median of  $\nu$  with  $\rho = 0.33$ .

There is an insignificant negative correlation ( $\rho = -0.20$ ) between the different mass measured within a certain density threshold, such as  $M_{200C}$ ,  $M_{500C}$ , and  $\max(\Delta s_{\text{KDE}})$ . The quantities  $M_{200C}$  and  $M_{500C}$ , which are computed within a shell centered on the most bound particles, have symmetry assumptions that may not capture the total mass well if there are substructures. The FoF mass, which captures the total mass without any symmetry assumption, correlates positively and weakly ( $\rho = 0.13$ ) with  $\max(\Delta s_{\text{KDE}})$ .

Overall, other than substructures, the offsets do not seem to be affected by other physical attributes of the clusters.

## 5 DISCUSSION

### 5.1 How to best visualize the cluster galaxy data?

We inspected both the luminosity maps and the number density maps of the member galaxy populations. With the same selection of bright galaxies of apparent  $i$ -band  $< 24.4$  at  $z = 0.3$ , the luminosity maps in general resemble the DM maps more closely than the number density maps. A comparison of the projected DM map, the luminosity map and the number density map of 129 clusters can be found at <https://goo.gl/kZUWrg>, <https://goo.gl/R7VNi9> and <https://goo.gl/lmQUPd> respectively. We encourage our readers to see how it is possible to create scientifically accurate luminosity contours that resemble the DM distribution if the member galaxy data is of high completeness and purity. The KDE is more than a method for identifying the luminosity peaks. (As the high resolution figures need to be downloaded, the corresponding Jupyter notebooks may take some time and several refresh of the web page before they are rendered properly.)

In real observations, missing selection of member galaxies, or foreground objects can both affect the inference of the galaxy spatial distribution. The number density map can be less susceptible to bias from bright foreground objects. It is less clear about the effect of missing member galaxies for computing the luminosity map because there is a selection bias favoring bright member galaxies.

### 5.2 Comparison to merging cluster observations

The offset distributions in table 2 represent an estimation of the spread of  $\Delta s$  in a  $\Lambda$ CDM universe. While the compilation of the offset distribution suffers from the small number of high mass clusters in the Illustris simulation, the distribution probes projection uncertainties comprehensively. Possible discrepancies that affect our comparison can arise due to the high purity and completeness of the Illustris galaxy data. Most observations of massive galaxy clusters with a total mass of  $M_{\odot} \approx 10^{14}$  have richness  $\lesssim 300$ , while



the most massive Illustris cluster ( $M_{\text{FoF}} = 3.23 \times 10^{14} M_{\odot}$ ) has a richness of 483. Furthermore, the Illustris clusters are less massive than the observed clusters with huge offsets. The Illustris clusters may also provide sufficient samples of possible spatial configurations of observed clusters. Despite the possible differences, we try to do a fair comparison to observation by matching the method of offset inference.

Other complications can arise from how one choose to subset the offset estimate based on the cluster properties. The full sample in table 2 preserves the cluster mass-abundance relationship, i.e. each cluster has the same number of projections in the full sample. However, it also underestimates the offset spread because the full sample includes  $\sim 45\%$  of relatively relaxed projections that only have one primary luminosity component. These clusters would have been excluded for comparison with bimodal mergers. Subsetting with  $1.2 < \nu < 2.2$  picks out cluster projections that are in more similar dynamical states as the observed merging cluster. However, some simulated clusters may have more projections included in this sample than the other clusters. From inspection of the mass abundance relations in Fig. A1, we found that subsampling with  $1.2 < \nu < 2.2$  includes a higher proportion of projections from massive clusters ( $\sim 20\%$  more) than the full sample. This sampling should not introduce significant bias. The observed merging clusters are still more massive than the unrelaxed samples.

We compile table 3 with the corresponding lower bounds for the p-values using the  $\Delta y$  distributions of the unrelaxed samples ( $1.2 < \nu < 2.2$ ). The p-value lower bounds are reported base on the  $\Delta y$  distribution from the same methods that the observed values were computed. The majority of the p-values from table 3, 13 out of 15, are  $> 0.05$ , only 2 p-values are below 0.05. This means the observations are mainly consistent with the null hypothesis: It is possible to see offset values as extreme as reported by observations in a CDM universe. However, this does not mean that the CDM model is more probable than the SIDM model. We will discuss the full physical implication of this results in section 5.5 after comparing our results to the SIDM signal in section 5.4. We continue to discuss the possible use of table 2 and the p-values in tables 3.

To avoid any report from this study directly quoting a p-value, we did not combine the p-values from different observations. This is because the computation of p-value does not fully take the true uncertainties of the observations into account. Most of the quoted uncertainties from 3 are only from the lensing estimate, but not the error estimates of the galaxy summary statistic. Any future studies that wish to claim significance or rejection based on comparison to a simulation will need to carefully track down the contribution of uncertainties from each aspect of the cluster analysis. It may be reasonable to inversely weight offsets by the reported bootstrapped uncertainty  $|\Delta s_{\text{KDE}}|$  for each cluster.

As warned by the American Statistical Association in a statement (Wasserstein & Lazar 2016), p-value can be misinterpreted easily, so we must take caution in the procedure of computing and interpreting the p-values. For example, the high energy physics community deliberately set the detection threshold to be  $5\sigma$  to account for how systematic and bootstrapped uncertainties can be underestimated or unaccounted for. If the reliability of each observation is approximately the same, averaging the p-values can be one way of combining the p-values, but there is no consensus on the best practice of combining the p-values. The rule used to combine the p-values needs to be carefully designed according to the goal of the experiment. When one keeps on including more observations, there will be some observations with extreme p-values that greatly influences the overall p-value. It is necessary to adjust for the sample

size for such running experiments. There are discussions of how to set stopping rules (Demortier 2007) to account for the sample size to determine what p-value level should indicate meaningful discrepancies from the expectation of the null hypothesis *before* analyzing the data. For this analysis, there is a long list of choices that can be up for debate, such as, whether to count the sample size base on each offset value or each cluster, whether to count all the observations of the same cluster (e.g. both Williams & Saha 2011 and Massey et al. 2015 have studied Abell 3827), or whether to count both  $\Delta x$  and  $\Delta y$  or just the direction along the approximate merger axis, and the list goes on. Regardless, computing a combined p-value for all the listed observations alone does not help understand how to best constrain SIDM. We instead focus on discussing the implications of the large uncertainties for constraining SIDM in section 5.5.

### 5.3 Comparison of the offset results to studies of clusters and groups

Our results showing offset distribution are highly relevant to two other types of observational studies: The first type tries to map the DM distribution of galaxy clusters using lensing techniques. The second type focuses on finding the best way to stack galaxy groups for inferring cosmological parameters from the galaxy cluster mass function. In the following section, we will show that cosmological simulations generally show the BCGs has the tightest offsets from the DM peaks, which is consistent with our finding. Observational studies, however, in general do not find offset as tight as the simulations. We discuss some factors (other than SIDM) quoted in the literature that have shown to affect the observed offsets, including the lensing resolution and the choice of the summary statistic for the member galaxy population. In particular, in the comparison of the luminosity peak to other galaxy summary statistic, the luminosity peak shows promise to give the second least amount of bias, after the BCG.

To establish the baseline of the tightest  $\Delta s$ , we first discuss and compare the  $\Delta s_{\text{BCG}}$  constraints. Cui et al. (2016), using 184 galaxy clusters with  $M > 10^{14} M_{\odot}$  in an N-body and hydrodynamical cosmological simulation suite powered by GADGET-3, also identified the maximum smoothed particle hydrodynamic (SPH) density peak to summarize the galaxy population in their cluster samples. Cui et al. (2016) found the majority of offsets between BCGs and the most gravitationally bound particle to be mostly below  $10 h^{-1} \text{ kpc}$ . They also reported some extreme outliers spanning up to several hundred  $h^{-1} \text{ kpc}$  due to the disturbed cluster morphology. Our tight 68-th percentile of  $\Delta y_{\text{BCG}}$  at  $\pm 3 \text{ kpc}$  gives some confidence that we have identified most of the BCGs correctly in the Illustris simulation.

The distributions of  $\Delta s_{\text{BCG}}$  derived from simulations, in general, are less spread out than those computed in most observational studies. For example, Oguri et al. (2010) have analyzed 25 X-ray luminous massive galaxy clusters of the LoCuSS survey. Observations were performed with a large FOV ( $\sim 3 h^{-1} \text{ Mpc}$ ) using the Subaru Suprime Camera. By fitting elliptical NFW models to the weak lensing data, Oguri et al. (2010) showed a long tail distribution for  $\Delta s_{\text{BCG}}$ , which they fit with two 2D Gaussians. The first tighter 2D Gaussian had a standard deviation being  $90 h^{-1} \text{ kpc}$  for describing the offset for most clusters, the long tail spans around 1 Mpc was fit by a second 2D Gaussian with a standard deviation of  $420 h^{-1} \text{ kpc}$ . This second component in the tail region contains  $\sim 10\%$  of the clusters in the study and is consistent with the number of extreme outliers that we have. Note that by subsetting to only

consider relatively relaxed clusters, there is a tight maximum bound  $\pm 19$  kpc for  $\Delta y_{\text{BCG}}$  in our study.

One major source of uncertainties for computing  $\Delta s_{\text{BCG}}$  is the misidentification of BCG. To see the effects of misidentifications of BCGs, some studies have made use of N-body hydrodynamical cosmological simulations to compute the 2D distances between the BCG and the second most massive galaxies. Johnston et al. (2007) and Hilbert & White (2010) (using the Millenium simulation) found the one sigma level offsets for misidentified BCGs at  $380 h^{-1}$  kpc, and  $410 h^{-1}$  kpc respectively. This is consistent with the 95-th percentile of the unrelaxed and the full sample of the BCG in our study and also the tail of the  $|\Delta s_{\text{BCG}}|$  for Cui et al. (2016). If one wishes to use BCG with high confidence, it may be necessary to set a stringent standard of the morphological characteristics such as using the large half light radius for classifying a BCG.

Another source of uncertainty for  $\Delta s_{\text{BCG}}$  was from lensing. Dietrich et al. (2012) performed an analogous analysis of the work of Oguri et al. (2010) using the N-body Millenium Run (MR) simulation. They showed that a combination of shape noise and modeling choices alone can lead to hundred-kpc-level offsets between the most bound particle (a proxy of the BCG) and the lensing peak for cluster-sized DM halos ( $M > 10^{14} M_{\odot}$ ). Dietrich et al. (2012) ray-traced through the DM substructures in the MR simulation as mock lensing observation of 512 clusters. Without any smoothing and shape noise, Dietrich et al. (2012) showed 90% of the lensing peak and the most bound particle, which is a proxy for the BCG, agree to  $2.0 h^{-1}$  kpc (0.65 arcsec at  $z = 0.3$ ). With shape noise, even at the source galaxy density of space-based quality optical data of  $n = 80 \text{ arcmin}^{-1}$ , fitting NFW halos and using the center as the DM gave a distribution of  $|\Delta s_{\text{BCG}}|$  those mode is at around 9 arcsecs (this is approximately the 1 standard deviation estimate in 2D). Lowering the source galaxy density to  $30 \text{ arcmin}^{-2}$  increased the mode to 22 arcsec with a 95-th percentile at 85 arcsec ( $\sim 90$  kpc and  $\sim 400$  kpc at  $z = 0.3$ , comparable to Oguri et al. 2010). Smoothing, in the presence of shape noise, resulted in an offset distribution with the mode at 15 arcsec ( $\sim 60$  kpc at  $z = 0.3$ ). The offsets do not simply depend on the smoothing bandwidth, but also the number density of the source galaxies that are lensed. While the uncertainty from smoothing the DM map is relatively unimportant in the Illustris analysis due to the much higher resolution of the DM particles than the sparser source galaxies used for lensing, it highlights why the bootstrapped uncertainties from the observed DM peak need to be accounted for during the comparison between the Illustris results and the observations.

It is noteworthy from the work of Dietrich et al. (2012) that smoothing alone can cause the peak offset to shift from several arcsecs to around 1 arcmin. Any other morphological features from the smoothed DM map will be subject to uncertainty of similar order of magnitude as the peak estimate, but with a lower signal. It is therefore hard to use other morphological features, such as tail or shape of a peak, as suggested by Kahlhoefer et al. (2014) for constraining SIDM. There is a risk of seeing substructures that mimic the sought-after morphological patterns. After all, other cosmological experiments that have shown that it is possible to detect patterns once scientists look hard enough, such as the initials of Stephen Hawking in the seventh-year Wilkinson Microwave Anisotropy Probe (WMAP) Cosmic Microwave Background data (Bennett et al. 2011).

Now we turn to exploring complementary methods for summarizing the galaxy population of a cluster, and show that the luminosity peak is the second best choice than the BCG for summarizing the galaxy statistic. A unique BCG does not always exist for a cluster (or the subcluster). Bright galaxies in the dense region of the cluster

are the possible progenitors of the BCG and therefore a reasonable choice when there is no unique BCG. George et al. (2012), for example, examined 129 X-ray selected non-merging galaxy groups in the COSMOS field. They found that around 20% to 30% of groups have non-negligible discrepancies between different galaxy centroids. By stacking on a bright galaxy near the X-ray centroid, they found the resulting lensing strength is higher than the stacked lensing signal based on other galaxy centroids, including the BCG. For groups with clear BCG candidate, George et al. (2012) gave the range of offset between the BCG and the assumed halo center as  $\lesssim 75$  kpc. The KDE peaks from the luminosity maps of the Illustris samples show a much tighter offset to the lensing center than any other centroids that George et al. (2012) investigated. The weighted or unweighted centroid measurement from George et al. (2012) has a  $|\Delta s|$  with standard deviation at 50 - 150 kpc from the lensing center with long tails (of around several hundred kpc). In comparison, the median (26 kpc), mean (37 kpc), standard deviation (35 kpc) and 75-th percentile (49 kpc) of  $|\Delta s_{\text{KDE}}|$  from all the Illustris samples are below 50 kpc.

We attribute the small amount of population bias of  $\Delta s_{\text{KDE}}$  in our analysis due to cross-validation, a procedure that is commonly seen and well accepted in the top journals from the statistics and the machine learning communities. Not only does the algorithm help determine the eigenvalues, but also the optimal eigenvector direction of the bandwidth matrix. Most literature does not discuss how the bandwidth of the kernel for the smoothed maps are determined. It is unclear if such results will enjoy the same accuracy level as what we demonstrated. If scientists hand tune the smoothing bandwidth, it is hard to avoid setting the bandwidth to fit the preconception of how the density contours of the cluster should look like, and inadvertently biasing  $\Delta s$ .

#### 5.4 Comparison to staged simulations with SIDM

Staged simulations are controlled experiments for probing the contribution of SIDM to offsets in mergers of cluster components. The non-deterministic nature of particle interactions means that it is not easy to predict the offsets analytically. Furthermore, there is no consensus on what type of clusters might best show the effects of SIDM, which depend on the model of the SIDM. Currently, the studies of SIDM have been restricted to those with isotropic scattering and velocity-independent cross sections. The two main classes of SIDM that are studied include those with frequent, short-range interactions that can be broadly be modeled by an effective drag force, and those with rare, longer range interactions those effects are not well approximated by a drag-force (Kahlhoefer et al. 2014). In addition having a dependence on the merger configurations, the offset due to SIDM also has a time-dependence.

At SIDM cross section level favored by current literature  $\sigma_{\text{SIDM}} \lesssim 1 \text{ cm}^2 \text{ g}^{-1}$ , a list of SIDM simulation studies (Kim & Peter 2016, Robertson et al. 2016, Kahlhoefer et al. 2014, Randall et al. 2008) have reported that, during the offset is observable, offset generally scales linearly with  $\sigma_{\text{SIDM}}$ , up to an offset of 50 kpc for the range of investigated  $\sigma_{\text{SIDM}} = 3 \text{ cm}^2/\text{g}$ . For simulations of major mergers of two subclusters, it takes time for the self-interaction of DM to manifest and lag behind the galaxies. Kim & Peter (2016), for example, showed that this lag starts to be observable approximately after the subclusters reach apocenter. While the offset may persist when the subclusters are returning for a second collision, the magnitude of the offset can fluctuate over this period. In general, the offset is affected by the phase of major mergers that cannot be directly calculated from observations. Some of the best estimates

of merger phase can only give 1-sigma uncertainties up to 1 Gyr precision (Dawson 2013b, Ng et al. 2015).

This small level of offset due to SIDM favors the use of statistics with less uncertainty for computing SIDM constraints. The uncertainties computed from the number density peak, centroid and shrinking aperture method simply overwhelm the signal. A offset level of 50 kpc is within the one-sigma level of  $\Delta s_{\text{num.dens}}$ ,  $\Delta s_{\text{shrink}}$ , and  $\Delta s_{\text{centroid}}$ , and the two-sigma level of the  $\Delta s_{\text{KDE}}$ . Despite  $\Delta s_{\text{BCG}}$  has tight distribution for relaxed clusters, misidentification of the BCG may increase the tail of the distribution to render it unsuitable to be used for secure constraints. To sample the tight distribution of  $\Delta s_{\text{BCG}}$  at a higher resolution than provided here, a simulation at galactic scales with realistic baryonic feedback will be needed.

Our results also illustrate one main difficulty for inferring  $\sigma_{\text{SIDM}}$ , it is hard to propagate and characterize observational error for the computation of  $\sigma_{\text{SIDM}}$ . Kim & Peter (2016) have shown that the SIDM offsets depend on the time of the merger, impact parameters, collisional velocity and mass concentration. Yet, there is no analytical model that includes all these merger parameters in the calculation of  $\sigma_{\text{SIDM}}$ . The alternative way of estimating  $\sigma_{\text{SIDM}}$  is to compare the observed offset to staged SIDM simulations. However, SIDM simulations treating all the observed offset signal of a cluster to be contributed by SIDM can give a biased estimate. As we have demonstrated,  $\sigma_{\text{SIDM}} = 0$  does not mean zero offset for individual clusters. Only the population estimate of offsets in a CDM universe, i.e. the mean, gives zero offset. At this point, it is unclear that the analysis with a population of merging clusters, can provide an excess in population offset signal at a characteristic scale than the noise level, and give better estimates for  $\sigma_{\text{SIDM}}$  than the study of individual cluster.

### 5.5 Prospect of detecting SIDM with confidence based on our results

So far our analysis has focused on two comparisons: 1) comparing  $\Delta s$  from the Illustris simulation to the observed values and 2) comparing  $\Delta s$  to the pure SIDM offset signal from staged simulations of SIDM. These comparisons informed us that the SIDM signal ( $< 50$  kpc) produced by the most favored  $\sigma_{\text{SIDM}}$  is less consistent with the higher end of the observed offset values ( $\sim 100$  kpc) than the offset uncertainties in the Illustris CDM simulation (68-th percentile ranging from 3 to 200 kpc depending on the galaxy summary statistic). However, we must emphasize these comparisons are insufficient to rule out small SIDM cross section.

Realistic mock observations of SIDM (cosmological) simulations with small cross section may match observations in better or worse ways. It is also possible that the offset levels generated by clusters with small  $\sigma_{\text{SIDM}}$  to be indistinguishable from our results. SIDM with a small cross section only produces offset under restricted set of merger configurations (Kim & Peter 2016): massive clusters with relatively low merger velocities, small impact parameters, and large halo concentrations. Yet the offset may only last for below 5 Gyrs with fluctuating magnitudes. Most clusters, therefore, do not have much constraining power on SIDM. Even in cosmological simulation of SIDM, one would need a selective observational strategy before the SIDM have statistically significant offset signal when compared to the corresponding population of clusters in a CDM simulation. It is unclear that the Illustris simulation has enough volume to contain clusters with such restrictive requirements. There is also the practical concern that such a cosmo-

logical simulation with SIDM will have a forbidding computational cost.

### 5.6 Possible improvements to better understand clusters

The data in the Illustris simulation has proven to be an excellent source of understanding observational uncertainties of galaxy clusters. Further work can be done to make mock observations more realistic with some of the following improvements:

- adding foreground structures
- improving the peak detection threshold from the KDE luminosity map based on the flux per area at the peak location and compare that to observational limitations
- examining the effects of incompleteness removing some member galaxies randomly
- computing the velocity dispersion of the galaxies along different line-of-sight to see how the velocity dispersion correlates the different measures of unrelaxedness for different projections
- testing the selection of member galaxies using the red sequence

## 6 SUMMARY

We have shown that

- it is possible to see  $\Delta s$  as extreme as those in observed merging galaxy clusters assuming that  $\Lambda$ CDM is the true underlying physical model
- the contribution of statistical uncertainty to the galaxy-DM offsets for  $\Lambda$ CDM clusters is *not* negligible when compared to the reported levels of offset from staged simulations ( $\sim 50$  kpc).
- only the galaxy-DM offsets derived from BCG and the weighted peak derived from the cross-validated luminosity map have two-sigma uncertainty levels within the reported SIDM offset level ( $< 50$  kpc). Other methods have one-sigma uncertainty levels that overwhelm the SIDM offset signal.
- while the location estimates of the 2D spatial distribution of offsets and the 1D spatial distributions of the offsets ( $\Delta y$ ) in the Illustris simulation are approximately zero. The root-mean-square of the magnitude of the offsets,  $|\Delta s_{\text{KDE}}|'$  is  $\sim 30$  kpc from the Illustris sample. does not necessarily map to zero magnitude of offset. Any studies that use a linear  $\sigma_{\text{SIDM}}$ -offset relationship that starts with  $\Delta s = 0$  to  $\sigma_{\text{SIDM}} = 0$  can be biased. Since the uncertainty does not always bias the offset high the following scenario may be possible: zero magnitude of  $\Delta s$  from merging clusters can be consistent with a range of models with small  $\sigma_{\text{SIDM}}$  under realistic observational conditions, but this hypothesis cannot be probed by the Illustris data.
- it is unclear if using a population of merging galaxy clusters can generate an excess of population offset signal at any characteristic scale above the statistical and observational noise level as estimated by our study.
- the locations of the most gravitationally bound particle are consistent with the BCG for systems with little substructures.
- the identified BCG can have a large offset from the DM peak due to a combination of effects from substructures and projection.



- the BCG has the smallest offset to the dominant DM peak. The KDE peak of the luminosity map after careful cross-validation gives the second tightest offsets from the DM peak.
- a naive implementation of the shrinking aperture is easily affected by substructures even for clusters with one dominant component.
- with high completeness and purity of member galaxy data, the luminosity map produced by a cross-validated kernel density estimate resembles the DM spatial distribution more closely than the number density map of member galaxies. The significant peaks from the luminosity map can also be used to characterize the amount of substructures in the cluster.

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## REFERENCES

- Allen S. W., 1998, *MNRAS*, 296, 392  
 Beers T. C., Flynn K., Gebhardt K., 1990, *AJ*, 100, 32  
 Bennett C. L., et al., 2011, *Astrophys. J. Suppl. Ser.*, 192, 17  
 Bradač M., et al., 2006, *ApJ*, 652, 937  
 Bradač M., Allen S. W., Treu T., Ebeling H., Massey R., Morris R. G., von der Linden A., Applegate D., 2008, *ApJ*, 687, 959  
 Clowe D., Markevitch M., Bradač M., Gonzalez A. H., Chung S. M., Massey R., Zaritsky D., 2012, *ApJ*, 758, 128  
 Cui W., et al., 2016, *MNRAS*, 456, 2566  
 Davis M., Efstathiou G., Frenk C. S., White S. D. M., 1985, *ApJ*, 292, 371  
 Dawson W. A., 2013a, PhD thesis, University of California, Davis  
 Dawson W. A., 2013b, *ApJ*, 772, 131  
 Demortier L., 2007, CDF/MEMO/STATISTICS/PUBLIC/8662  
 Deng H., Wickham H., 2011, Technical report, Density estimation in R. had.co.nz  
 Dietrich J. P., Böhnert A., Lombardi M., Hilbert S., Hartlap J., 2012, *MNRAS*, 419, 3547  
 Duong T., 2007, *J. Stat. Softw.*, 21, 1  
 Feigelson E. D., Babu G. J., 2014, *Contemp. Phys.*, 55, 126  
 Ford J., et al., 2014, *MNRAS*, 447, 1304  
 Genel S., et al., 2014, *MNRAS*, 445, 175  
 George M. R., et al., 2012, *ApJ*, 757, 2  
 Gorski K. M., Hivon E., Banday A. J., Wandelt B. D., Hansen F. K., Reinecke M., Bartelmann M., 2005, *ApJ*, 622, 759  
 Hall P., Marron J. S., Park B. U., 1992, *Probab. Theory Relat. Fields*, 92, 1  
 Harvey D., et al., 2014, *MNRAS*, 441, 404  
 Hilbert S., White S. D. M., 2010, *MNRAS*, 404, 486  
 Hoag A., et al., 2016, arXiv Prepr., p. 48  
 Jee M. J., Hughes J. P., Menanteau F., Sifón C., Mandelbaum R., Barrientos L. F., Infante L., Ng K. Y., 2014, *ApJ*, 785, 20  
 Jee M. J., et al., 2015, *ApJ*, 802, 46  
 Johnston D. E., et al., 2007, arXiv Prepr.

- Kahlhoefer F., Schmidt-Hoberg K., Frandsen M. T., Sarkar S., 2014, *MNRAS*, 437, 2865  
 Markevitch M., Gonzalez A. H., Clowe D., Vikhlinin A., Forman W., Jones C., Murray S., Tucker W., 2004, *ApJ*, 606, 819  
 Massey R., et al., 2015, *MNRAS*, 449, 3393  
 Medezinski E., et al., 2013, *ApJ*, 777, 43  
 Ng K. Y., Dawson W. A., Wittman D., Jee M. J., Hughes J. P., Menanteau F., Sifón C., 2015, *MNRAS*, 453, 1531  
 Oguri M., Takada M., Okabe N., Smith G. P., 2010, *MNRAS*, 405, no  
 Perez F., Granger B. E., 2007, *Comput. Sci. Eng.*, 9, 21  
 R Core Team 2014, R: A Language and Environment for Statistical Computing. R Foundation for Statistical Computing, Vienna, Austria, <http://www.R-project.org/>  
 Randall S. W., Markevitch M., Clowe D., Gonzalez A. H., Bradač M., Bradac M., 2008, *ApJ*, 679, 1173  
 Robertson A., Massey R., Eke V., 2016, arXiv Prepr., p. 20  
 Robitaille T. P., et al., 2013, *A&A*, 558, A33  
 Springel V., 2010, *MNRAS*, 401, 791  
 VanderPlas J., Connolly A. J., Ivezić Z., Gray A., 2012, in 2012 Conf. Intell. Data Underst.. IEEE, pp 47–54, doi:10.1109/CIDU.2012.6382200, <http://ieeexplore.ieee.org/lpdocs/epic03/wrapper.htm?arnumber=6382200>  
 Vogelsberger M., et al., 2014a, *MNRAS*, 444, 1518  
 Vogelsberger M., et al., 2014b, *Nature*, 509, 177  
 Wasserstein R. L., Lazar N. A., 2016, *Am. Stat.*, 70, 129  
 Williams L. L. R., Saha P., 2011, *MNRAS*, 415, 448  
 Zitrin A., Bartelmann M., Umetsu K., Oguri M., Broadhurst T., 2012, *MNRAS*, 426, 2944  
 Zitrin A., Menanteau F., Hughes J. P., Coe D., Barrientos L. F., Infante L., Mandelbaum R., 2013, *ApJ*, 770, L15

## APPENDIX A: ALGORITHM OF THE SHRINKING APERTURE ESTIMATES

**Data:** subhalo that satisfy cuts as a galaxy

---

```

initial aperture centroid = weighted mean galaxy location in
each spatial dimension
distance array = euclidean distances between initial aperture
center and each galaxy location
aperture radius = 90th percentile of the weighted distance
array
while (newCenterDist - oldCenterDist) / oldCenterDist ≥
2e-2 do
    new data array = old data array within aperture
    newCenter = weighted mean value of new data along
    each spatial dimension
end

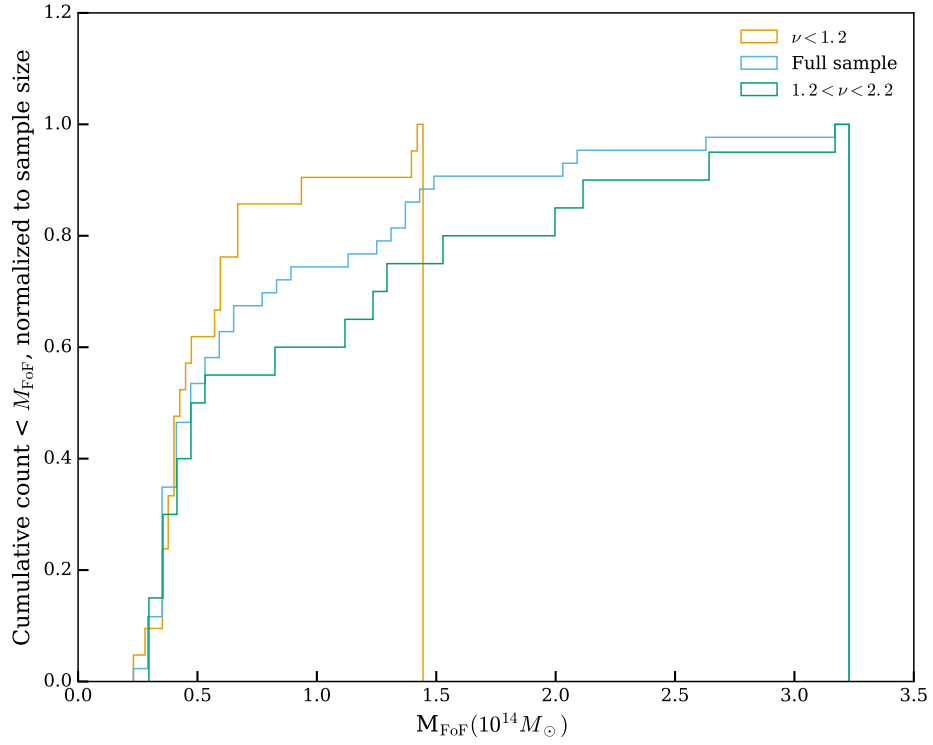
```

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**Algorithm 1:** Shrinking aperture algorithm with luminosity weights

## APPENDIX B: TABLE OF RESULTS





**Figure A1.** Cumulative distribution of clusters above a certain mass threshold for different samples. Each distribution is normalized to the sample size. The unrelaxed samples  $1.2 < \nu < 2.2$  If the subsets have the same cluster mass abundance as the full sample, the three plots should lie on top of one another.

**Table B1.** Properties of the clusters used in the analysis. Richness is computed based on  $i$ -band  $< 24.4$  assuming  $z = 0.3$ .

ID	richness	$M_{200C} (10^{14} M_{\odot})$	$M_{500C} (10^{14} M_{\odot})$	$M_{FzF} (10^{14} M_{\odot})$	unrelaxedness <sub>0</sub>	unrelaxedness <sub>1</sub>	midvar( $\Delta y_{KDE}$ ) (kpc)	max( $\Delta y_{KDE}$ ) (kpc)	median( $\nu$ )
0	483	1.64	1.09	3.23	29	33	31	65	1.43
1	338	1.57	0.62	2.68	20	16	25	71	1.59
2	267	1.53	0.87	2.12	17	3	18	42	1.30
3	343	0.82	0.56	2.03	37	59	44	148	2.01
4	213	1.19	0.66	1.54	21	4	24	84	1.58
5	212	0.90	0.56	1.44	20	27	16	43	1.19
6	225	0.96	0.60	1.40	18	7	15	28	1.16
7	230	0.31	0.17	1.41	54	280	101	379	2.83
8	148	0.83	0.54	1.34	24	26	20	52	1.32
9	187	0.79	0.50	1.29	23	12	33	111	1.45
10	158	0.73	0.53	1.15	19	8	19	49	1.35
11	134	0.57	0.33	0.95	20	9	36	78	1.12
12	164	0.20	0.09	0.87	64	142	77	218	1.99
13	115	0.22	0.14	0.79	63	143	38	118	2.26
14	90	0.45	0.29	0.67	15	8	17	33	1.08
15	92	0.51	0.35	0.68	11	3	11	25	1.00
16	113	0.40	0.23	0.61	19	4	13	31	1.06
17	97	0.42	0.18	0.60	21	8	27	53	1.09
18	83	0.45	0.31	0.59	15	8	14	32	1.20
19	86	0.26	0.19	0.57	30	68	18	77	1.52
20	84	0.15	0.11	0.50	60	122	54	117	1.65
21	89	0.26	0.12	0.53	23	8	47	146	1.24
22	70	0.42	0.30	0.49	14	7	10	23	1.06
23	68	0.25	0.17	0.47	30	25	26	98	1.03
24	66	0.33	0.26	0.44	14	14	11	42	1.17
25	79	0.23	0.15	0.43	23	25	11	22	1.25
26	61	0.26	0.18	0.45	28	40	11	44	1.30
28	69	0.30	0.16	0.41	22	12	26	42	1.01
29	62	0.30	0.20	0.42	16	14	9	22	1.20
30	59	0.18	0.14	0.40	42	78	17	63	1.35
31	57	0.29	0.21	0.40	14	15	10	24	1.06
32	56	0.18	0.13	0.38	35	23	43	83	1.54
33	69	0.19	0.10	0.38	49	54	60	108	2.11
34	63	0.21	0.14	0.39	23	20	22	33	1.07
35	69	0.29	0.22	0.41	12	3	11	28	1.01
36	72	0.24	0.16	0.36	21	22	16	39	1.18
37	63	0.21	0.16	0.36	25	23	51	142	1.11
39	55	0.27	0.18	0.36	11	3	12	29	1.00
40	54	0.18	0.10	0.33	44	69	81	151	1.39
46	52	0.08	0.06	0.30	57	73	59	157	1.65
48	53	0.12	0.08	0.30	40	104	13	44	1.82
51	56	0.19	0.13	0.29	12	5	17	41	1.00
58	58	0.14	0.09	0.23	29	10	21	66	1.00

**Table B2.** Summary statistic characterizing the offset distributions between the most bound particle and various summary statistics of the member galaxy population.

offset (kpc)	location	lower 68%	lower 95%	lower 99%	upper 68%	upper 95%	upper 99%
$\Delta y_{\text{BCG}}$	0	-2	-2	-252	2	528	1107
$\Delta y'_{\text{centroid}}$	0	-134	-491	-1176	134	491	1176
$\Delta y'_{\text{KDE}}$	0	-19	-82	-1182	19	82	1182
$\Delta y_{\text{num.dens}}$	0	-83	-302	-1114	83	302	1114
$\Delta y'_{\text{shrink}}$	0	-50	-288	-1025	50	288	1025

The offsets represented with the prime ' symbols are estimated using the luminosity weighted galaxy data.

**Table B3.** Summary statistic characterizing the offset distributions for between the DM peak and the estimated galaxy location. All 43 clusters and all 768 projections are used in this table. The highest density values were used for the computation when there were more than one peak value estimated from the KDE. There are different levels of asymmetry depending on how sparse a region is.

kpc	mean	std	min	25%	50%	75%	max
$ \Delta s_{\text{BCG}} $	69	294	0	2	3	7	2335
$\Delta x_{\text{BCG}}$	-14	226	-2331	-2	-0	1	2327
$\Delta y_{\text{BCG}}$	23	197	-1980	-2	0	2	2332
$ \Delta s'_{\text{centroid}} $	261	209	2	114	202	317	1103
$\Delta x'_{\text{centroid}}$	-42	224	-1022	-164	-37	66	1101
$\Delta y'_{\text{centroid}}$	0	244	-1102	-111	-0	111	1100
$ \Delta s'_{\text{shrink}} $	118	156	0	21	60	165	1454
$\Delta x'_{\text{shrink}}$	-7	131	-1089	-39	-3	23	969
$\Delta y'_{\text{shrink}}$	0	145	-1091	-32	0	32	1109
$ \Delta s'_{\text{KDE}} $	37	35	0	14	26	49	498
$\Delta x'_{\text{KDE}}$	-2	35	-330	-17	-2	12	386
$\Delta y'_{\text{KDE}}$	-0	37	-439	-15	0	15	440
$ \Delta s_{\text{num.KDE}} $	136	161	1	56	92	147	2126
$\Delta x_{\text{num.KDE}}$	-12	142	-1967	-55	-4	53	993
$\Delta y_{\text{num.KDE}}$	-0	155	-1415	-54	-0	54	1417

The offsets represented with the prime ' symbols are estimated using the luminosity weighted galaxy data.



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