Lecture 1: Causal Identification

POL-GA 1251 Quantitative Political Analysis II Prof. Cyrus Samii NYU Politics

January 23, 2022

Plan for the week

Mon:

- Discuss overview of class and syllabus.
- Explain what "causal identification" means.
- ► Introduce potential outcomes and causal graphs.

Wed:

- Randomized experiments.
- Explain estimation concepts (estimand, estimators, bias, consistency, efficiency).
- Explain *statistical inference* concepts (sampling distribution, randomization distribution, CLT, confidence intervals, *p*-value).

Where this class fits in

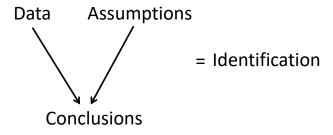
Model of quantitative research process:

- ► Theory motivates causal hypothesis or target of inference:
 - ► H: manipulating X results in (...) effect on Y.
- Hypothesis, statistical theory, and substantive theory motivate a research design:
 - ▶ Operationalize *X* and *Y*.
 - Define ways to get necessary variation in X and Y given constraints.
- Research design and statistical theory motivate analysis plan:
 - Optimal estimation strategy, given constraints.
 - Optimal testing strategy, given constraints.

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In this class we focus on causal identification. This is a specific application of the general idea of "identification," distinct from some other applications:

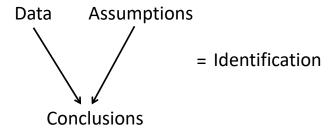
Alternative application of "identification" (I):

Suppose in the NYC mayoral primary someone said...

...they preferred Adams over Wiley, and

...they preferred Garcia over Adams.

Does this information (data) identify the person's preference ordering over these three candidates?



Alternative application of "identification" (II):

Suppose none of the coefficients below are equal to zero but the error terms (last ones) are iid mean zero draws. Which system of simultaneous equations identifies its coefficients?

$$x_{t} = \alpha_{1}^{a} + \alpha_{2}^{a} y_{t} + v_{t}^{a}$$

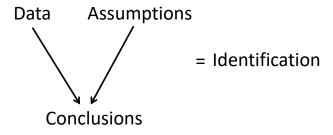
$$y_{t} = \beta_{1}^{a} + \beta_{2}^{a} x_{t} + \varepsilon_{t}^{a}$$

$$x_{t} = \alpha_{1}^{b} + \alpha_{2}^{b} y_{t} + \alpha_{3}^{b} w_{t} + \alpha_{4}^{b} v_{t} + v_{t}^{b}$$

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Angrist and Krueger (1999):

The combination of a clearly labeled source of identifying variation in a causal variable and the use of a particular econometric[/statistical] technique to exploit this information is what we call an identification strategy.

The Road Not Taken

Two roads diverged in a yellow wood, And sorry I could not travel both And be one traveler, long I stood And looked down one as far as I could To where it bent in the undergrowth;

Then took the other, as just as fair, And having perhaps the better claim, Because it was grassy and wanted wear; Though as for that the passing there Had worn them really about the same,

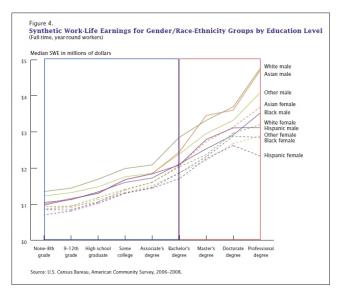
And both that morning equally lay In leaves no step had trodden black. Oh, I kept the first for another day! Yet knowing how way leads on to way, I doubted if I should ever come back.

I shall be telling this with a sigh Somewhere ages and ages hence: Two roads diverged in a wood, and I— I took the one less traveled by, And that has made all the difference. Modern frameworks for causal analysis:

- ▶ Potential outcomes (Neyman, 1923; Rubin, 1974, 1978).
- ► Causal graphs (Pearl, 2009).

Both rely on "counterfactual" logic.

Running Example: Effect of College on Earnings*



^{*}SWE = expected total earnings over 25-64.

A causal effect can be defined as a contrast between "potential outcomes."

	Pret	reatm alues X	nent	which treatment	Y									M	Missing data indicator							
					γ1				Υ ^T			м×			м¹				м ^T			
	х,		X _c		Y11		Y _d		Y_1^T		Y_d^T	$\mathbf{M}_{1}^{\mathbf{X}}$		M ^K C	м1		M_d^1		M_1^T		M_d^T	
Experimental units in population P										,												

Fig. 1. All values in a study of T treatments.

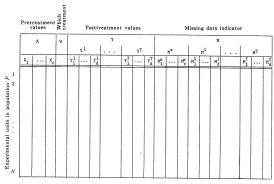


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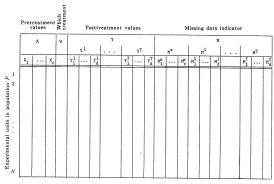


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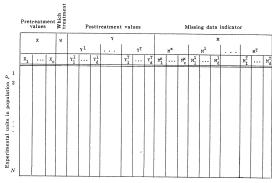


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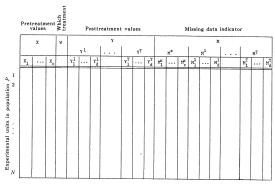


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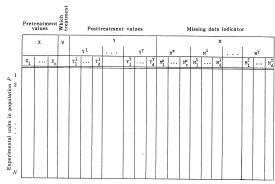


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- Missing data indicators, $M_{i,j}^{(k)}$.

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- ▶ Unit level causal effects for members of \mathscr{P} are fixed a priori, and they compare y_{wi} to $y_{\tilde{w}i}$ for $w \neq \tilde{w}$ and $i \in \mathscr{P}$.
- ▶ Population causal effects for compare aggregates of unit level causal effects for members of \mathscr{P} .
- ► Effects are defined in an "agnostic" or "non-parametric" way.
- Potential outcomes and covariates are fixed, treatments and response indicators stochastic.
- ► Effects are defined by letting only treatments vary, holding units fixed.
- ► Thus, causal effects are clearly defined for units that can conceivably receive different treatment values.
- ▶ A test for the above is "manipulation" (Holland, 1986).

Potential outcomes, causal effects, and manipulability

Holland (1986): "For causal inference, it is critical that each unit be potentially exposable to any one of the causes."

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Issues arise when trying to interpret things like race or gender. See VanderWeele and Robinson (2014) for a formal treatment of ways to interpret "race effects."

Potential outcomes and fundamental problem of causal inference

Recall, a unit level causal effect compares y_{wi} to $y_{\tilde{w}i}$ for $w \neq \tilde{w}$.

"Fundamental problem of causal inference" (Holland, 1986): For each i potential outcomes for all w exist, but we only observe the potential outcome for the treatment value that i receives.

- "Scientific solution": Use theory to determine when units are interchangeable.
- "Statistical solution": Study features of conditional distributions, such as averages.



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- Each draw is characterized by
 - \triangleright a covariate vector, X_i ,
 - ▶ potential outcomes that under SUTVA are characterized as Y_{di} for all $d \in \mathcal{D}$, as well as
 - treatment assignments, $D_i \in \mathcal{D}$.

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for which,

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For our running example, we have $D_i = 1$ if college, $D_i = 0$ if not. Outcome of interest is income. ρ is the average income benefit of college.

Consider simple difference in mean college grad incomes vs mean no college incomes:

$$E[Y_i|D_i=1] - E[Y_i|D_i=0] = E[Y_{1i}|D_i=1] - E[Y_{0i}|D_i=0]$$

▶ Then what does this difference equal?

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- ► E.g, consider a decomposition wrt ATT:

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$$= \underbrace{E[Y_{1i}|D_i=1] - E[Y_{0i}|D_i=1]}_{\text{Average treatment effect on the treated (ATT)}}$$

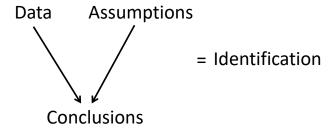
$$+ \underbrace{E[Y_{0i}|D_i=1] - E[Y_{0i}|D_i=0]}_{\text{Selection bias}}.$$

$$(2)$$

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$$\begin{split} E[Y_i|D_i = 1] - E[Y_i|D_i = 0] &= E[Y_{1i}|D_i = 1] - E[Y_{0i}|D_i = 0] \\ &= \underbrace{E[Y_{1i}|D_i = 1] - E[Y_{0i}|D_i = 1]}_{\text{Average treatment effect on the treated (ATT)}} \\ &+ \underbrace{E[Y_{0i}|D_i = 1] - E[Y_{0i}|D_i = 0]}_{\text{Selection bias}}. \end{split}$$

► Could do similar wrt to ATC or effect heterogeneity (cf. CCI).



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$$\underbrace{E[Y_{1i}|D_i=1] - E[Y_{0i}|D_i=1]}_{\text{ATT}} = E[Y_{1i} - Y_{0i}],$$

so the simple difference, (2), equals ρ .

Identifying assumption 2 (conditionally independent/unconfounded/strongly ignorable assignment):

$$D_i \perp \!\!\! \perp (Y_{1i}, Y_{0i}) | X_i \text{ and } 0 < Pr[D_i = 1 | X_i = x] < 1 \text{ for all } x \in \mathscr{X}$$
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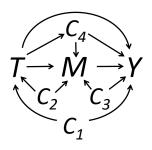
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Marginalization over \mathcal{X} , the support of X_i , yields,

$$\int_{\mathscr{X}} \rho(x) dF(x) = \rho.$$



Causal graphs

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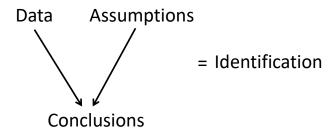
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▶ Difference between $E[Y|do(D_i = 1)] - E[Y|do(D_i = 0)]$ and $E[Y|D_i = 1] - E[Y|D_i = 0]$ is "backdoor paths" from D to Y.



► Random assignment ⇒ intervention graph *is* population graph:



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► CIA implies no backdoor path through *U*:



▶ Random assignment \Rightarrow intervention graph *is* population graph:



► CIA implies no backdoor path through *U*:



► Conditioning on *X* removes the other backdoor path:



► Marginalize over *x* to recover the intervention graph.

- ► These are examples of "closing" backdoor paths.
- ▶ Other operations, e.g., "opening" backdoor paths by conditioning on "colliders":





Looking forward to the rest of the class

- ▶ With respect to our example and cases that resemble it:
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- ► Time for questions.