

Moderation

Introduction to SEM with Lavaan



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Outline

Moderation Basics

Post Hoc Analysis

Latent Variable Interactions

- Products of Manifest Variables

- Products of Latent Variables

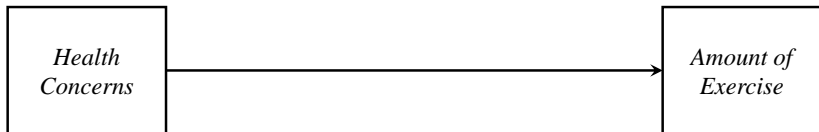
Multiple Moderation

Categorical Moderators



Refresher: Focal Effect Only

The *healthConcerns* → *exerciseAmount* relation is our *focal effect*



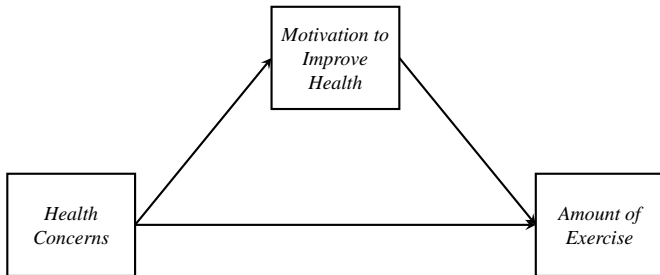
- Mediation, moderation, and conditional process analysis all attempt to describe the focal effect in more detail.
- We always begin by hypothesizing a focal effect.



Refresher: Mediation Hypothesis

A mediation analysis will attempt to describe how health concerns affect amount of exercise.

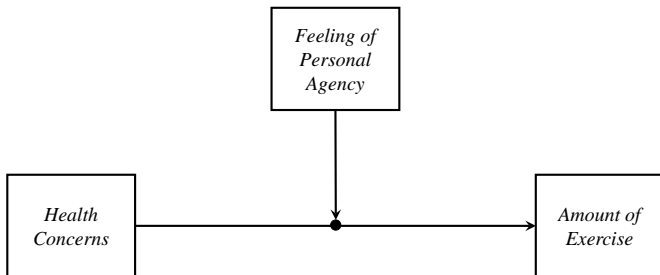
- The *how* is operationalized in terms of intermediary variables.
- Mediator: Motivation to improve health (*motivation*).



Refresher: Moderation Hypothesis

A moderation hypothesis will attempt to describe when health concerns affect amount of exercise.

- The *when* is operationalized in terms of interactions between the focal predictor and contextualizing variables
- Moderator: Sense of personal agency relating to physical health (*agency*).



Equations

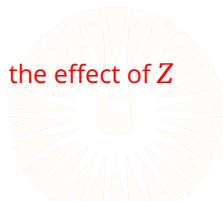
In additive MLR, we might have the following equation:

$$Y = \beta_0 + \beta_1 X + \beta_2 Z + \varepsilon$$

This additive equation assumes that X and Z are independent predictors of Y .

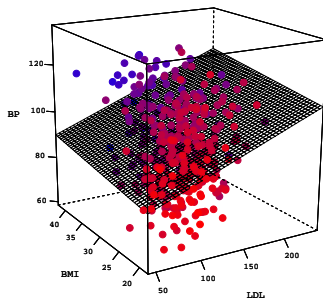
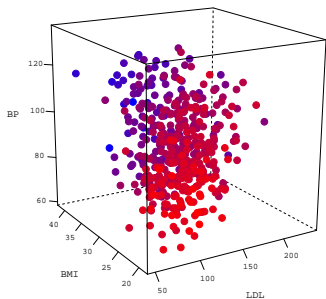
When X and Z are independent predictors, the following are true:

- X and Z *can* be correlated.
- β_1 and β_2 are *partial* regression coefficients.
- The effect of X on Y is the same at **all levels** of Z , and the effect of Z on Y is the same at **all levels** of X .



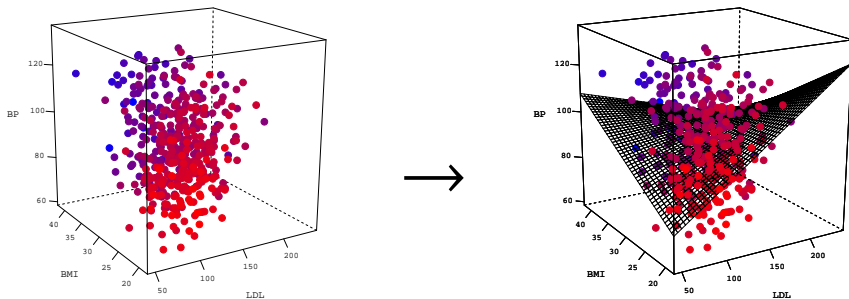
Additive Regression

The effect of X on Y is the same at **all levels** of Z .



Moderated Regression

The effect of X on Y varies **as a function** of Z .



Equations

The following derivation is adapted from Hayes (2022).

- When testing moderation, we hypothesize that the effect of X on Y varies as a function of Z .
- We can represent this concept with the following equation:

$$Y = \beta_0 + f(Z)X + \beta_2Z + \varepsilon \quad (1)$$



Equations

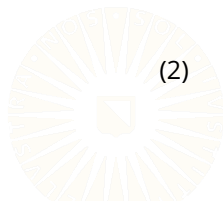
The following derivation is adapted from Hayes (2022).

- When testing moderation, we hypothesize that the effect of X on Y varies as a function of Z .
- We can represent this concept with the following equation:

$$Y = \beta_0 + f(Z)X + \beta_2Z + \varepsilon \quad (1)$$

- If we assume that Z linearly (and deterministically) affects the relationship between X and Y , then we can take:

$$f(Z) = \beta_1 + \beta_3Z \quad (2)$$



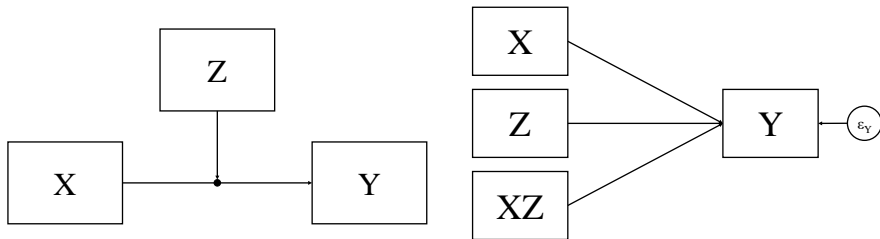
Equations

- Substituting Equation 2 into Equation 1 leads to:

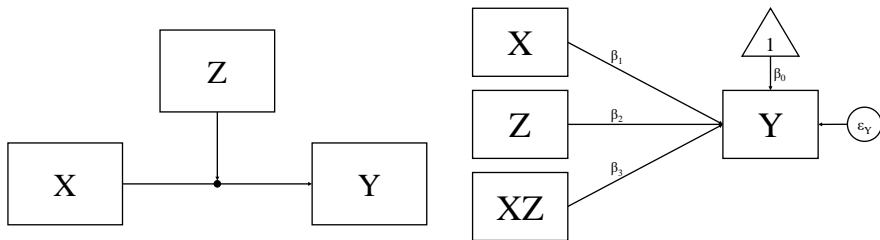
$$Y = \beta_0 + (\beta_1 + \beta_3 Z)X + \beta_2 Z + \varepsilon$$



Conceptual vs. Analytic Diagrams



Conceptual vs. Analytic Diagrams



Testing Moderation

Now, we have an estimable regression model that quantifies the linear moderation we hypothesized.

$$Y = \beta_0 + \beta_1 X + \beta_2 Z + \beta_3 XZ + \varepsilon$$

- To test for significant moderation, we simply need to test the significance of the interaction term, XZ .
 - Check if $\hat{\beta}_3$ is significantly different from zero.



Interpretation

Given the following equation:

$$Y = \hat{\beta}_0 + \hat{\beta}_1 X + \hat{\beta}_2 Z + \hat{\beta}_3 XZ + \hat{\varepsilon}$$

- $\hat{\beta}_3$ quantifies the effect of Z on the focal effect (the $X \rightarrow Y$ effect).
 - For a unit change in Z , $\hat{\beta}_3$ is the expected change in the effect of X on Y .
- $\hat{\beta}_1$ and $\hat{\beta}_2$ are *conditional effects*.
 - Interpreted where the other predictor is zero.
 - For a unit change in X , $\hat{\beta}_1$ is the expected change in Y , when $Z = 0$.
 - For a unit change in Z , $\hat{\beta}_2$ is the expected change in Y , when $X = 0$.

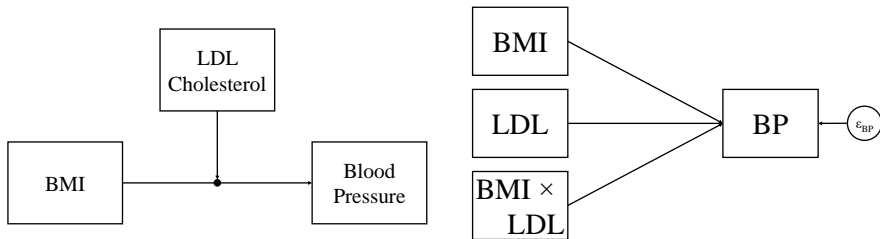
Example

Looking at the *diabetes* dataset.

- We suspect that patients' BMIs are predictive of their average blood pressure.
- We further suspect that this effect may be differentially expressed depending on the patients' LDL levels.



Diagrams



Example

```
dDat <- readRDS("../data/diabetes.rds")
```

```
## Focal Effect:
```

```
out0 <- lm(bp ~ bmi, data = dDat)
```

```
partSummary(out0, -c(1, 2))
```

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	61.9973	3.6659	16.91	<2e-16
bmi	1.2379	0.1371	9.03	<2e-16

Residual standard error: 12.72 on 440 degrees of freedom

Multiple R-squared: 0.1563, Adjusted R-squared: 0.1544

F-statistic: 81.54 on 1 and 440 DF, p-value: < 2.2e-16

Example

```
## Additive Model:
```

```
out1 <- lm(bp ~ bmi + ldl, data = dDat)  
partSummary(out1, -c(1, 2))
```

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	59.26577	3.91281	15.147	< 2e-16
bmi	1.16567	0.14156	8.235	2.08e-15
ldl	0.04016	0.02056	1.953	0.0515

Residual standard error: 12.68 on 439 degrees of freedom

Multiple R-squared: 0.1636, Adjusted R-squared: 0.1598

F-statistic: 42.94 on 2 and 439 DF, p-value: < 2.2e-16

Example

```
## Moderated Model:
```

```
out2 <- lm(bp ~ bmi * ldl, data = dDat)  
partSummary(out2, -c(1, 2))
```

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	14.480616	14.291677	1.013	0.311514
bmi	2.867825	0.541312	5.298	1.86e-07
ldl	0.448771	0.127160	3.529	0.000461
bmi:ldl	-0.015352	0.004716	-3.255	0.001221

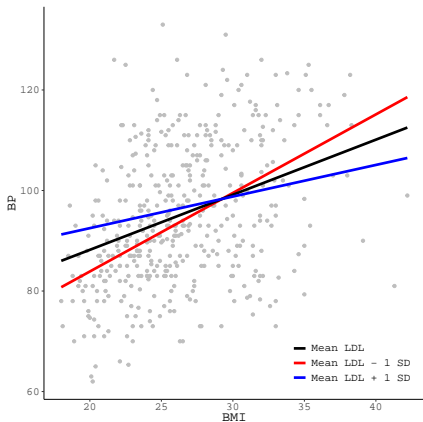
Residual standard error: 12.54 on 438 degrees of freedom

Multiple R-squared: 0.1834, Adjusted R-squared: 0.1778

F-statistic: 32.78 on 3 and 438 DF, p-value: < 2.2e-16

Visualizing the Interaction

We can get a better idea of the patterns of moderation by plotting the focal effect at conditional values of the moderator.



Example

Of course, we can fit the same model in **lavaan**.

```
library(lavaan)

## Specify the model:
mod <- 'bp ~ 1 + bmi + ldl + bmi:ldl'

## Estimate the model:
lavOut <- sem(mod, data = dDat)
```

Example

```
partSummary(lavOut, 7:9)
```

Regressions:

	Estimate	Std.Err	z-value	P(> z)
bp ~				
bmi	2.868	0.539	5.322	0.000
ldl	0.449	0.127	3.545	0.000
bmi:ldl	-0.015	0.005	-3.270	0.001

Intercepts:

	Estimate	Std.Err	z-value	P(> z)
.bp	14.481	14.227	1.018	0.309

Variances:

	Estimate	Std.Err	z-value	P(> z)
.bp	155.871	10.485	14.866	0.000

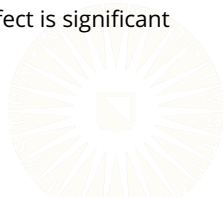
POST HOC ANALYSIS



Probing the Interaction

A significant estimate of β_3 tells us that the effect of X on Y depends on the level of Z , but not much more.

- The plot above gives a descriptive illustration of the pattern, but does not support statistical inference.
 - The three conditional effects we plotted look different, but we cannot say much about how they differ with only the plot and $\hat{\beta}_3$.
- This is the purpose of *probing* the interaction.
 - Try to isolate areas of Z 's distribution in which $X \rightarrow Y$ effect is significant and areas where it is not.



Probing the Interaction

The most popular method of probing interactions is to do a so-called *simple slopes* analysis.

- Pick-a-point approach
- Spotlight analysis

In simple slopes analysis, we test if the slopes of the conditional effects plotted above are significantly different from zero.

- To do so, we test the significance of *simple slopes*.



Simple Slopes

Recall the derivation of our moderated equation:

$$Y = \beta_0 + \beta_1 X + \beta_2 Z + \beta_3 XZ + \varepsilon$$

We can reverse the process by factoring out X and reordering terms:

$$Y = \beta_0 + (\beta_1 + \beta_3 Z)X + \beta_2 Z + \varepsilon$$

Where $f(Z) = \beta_1 + \beta_3 Z$ is the linear function that shows how the relationship between X and Y changes as a function of Z .

$f(Z)$ is the *simple slope*.

- By plugging different values of Z into $f(Z)$, we get the value of the conditional effect of X on Y at the chosen level of Z .

Significance Testing of Simple Slopes

The values of Z used to define the simple slopes are arbitrary.

- The most common choice is: $\{(\bar{Z} - SD_Z), \bar{Z}, (\bar{Z} + SD_Z)\}$
- You could also use interesting percentiles of Z 's distribution.

The standard error of a simple slope is given by:

$$SE_{f(Z)} = \sqrt{SE_{\beta_1}^2 + 2Z \cdot \text{cov}(\beta_1, \beta_3) + Z^2 SE_{\beta_3}^2}$$

So, you can test the significance of a simple slope by constructing a t-statistic or confidence interval using $\hat{f}(Z)$ and $SE_{f(Z)}$:

$$t = \frac{\hat{f}(Z)}{SE_{f(Z)}}, \quad CI = \hat{f}(Z) \pm t_{crit} \times SE_{f(Z)}$$



Example

We can use **semTools** routines to probe interaction in **lavaan** models.

- `probe2WayMC()`: simple slopes/intercepts analysis
- `plotProbe()`: simple slopes plots

```
library(semTools)

## Estimate and test simple slopes and simple intercepts:
ssOut <- probe2WayMC(lavOut,
  nameX = c("bmi", "ldl", "bmi:ldl"),
  nameY = "bp",
  modVar = "ldl",
  valProbe = quantile(dDat$ldl, c(0.25, 0.50, 0.75))
)
```

Example

```
## View the results:
```

```
ssOut
```

```
$SimpleIntcept
```

	ldl	est	se	z	pvalue
25%	96.05	57.585	4.017	14.334	0
50%	113.00	65.192	3.736	17.449	0
75%	134.50	74.840	4.944	15.139	0

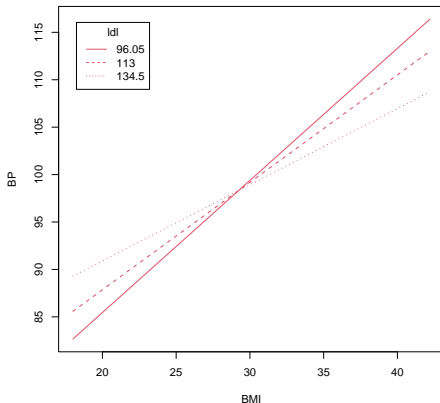
```
$SimpleSlope
```

	ldl	est	se	z	pvalue
25%	96.05	1.393	0.156	8.942	0
50%	113.00	1.133	0.140	8.107	0
75%	134.50	0.803	0.178	4.508	0

Example

```
## Plot the simple slopes:
```

```
plotProbe(ssOut, xlim = range(dDat$bmi), xlab = "BMI", ylab = "BP")
```



LATENT VARIABLE INTERACTIONS



Latent Variable Interactions

When we have two observed variables interacting to predict a latent variable, our job is easy:

1. Construct a product term from the focal and moderator variables.
2. Use the observed focal, moderator, and interaction variables to predict the latent DV.

If we want to model moderation when at least one of the predictors is latent, things get more difficult.

- For observed and discrete moderators, use multiple group modeling.
- For continuous and/or latent moderators, we need fancier methods.

Two basic approaches:

1. Methods based on products of manifest variables
2. Methods based on directly estimating the products of latent variables

Computing Interaction Indicators

The simplest approach is to create observed product terms and directly use those products as indicators of an interaction construct.

- Naively indicating an interaction construct with the raw product terms is probably sub-optimal.
- Collinearity among the interaction indicators and the raw items can cause estimation problems.
- We'd also like to interpret our final model holistically.

Two recommended approaches:

1. Orthogonalization through residual centering (Little, Bovaird, & Widaman, 2006).
2. Double mean centering (Lin, Wen, Marsh, & Lin, 2010).

Orthogonalization Procedure

Say we think that Z moderates the $X \rightarrow Y$ effect.

- X , Y , and Z are latent variables indicated by $\{x_1, x_2, x_3\}$, $\{y_1, y_2, y_3\}$, and $\{z_1, z_2, z_3\}$, respectively.



Orthogonalization Procedure

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- X , Y , and Z are latent variables indicated by $\{x_1, x_2, x_3\}$, $\{y_1, y_2, y_3\}$, and $\{z_1, z_2, z_3\}$, respectively.

We perform the orthogonalization procedure as follows:

1. Construct all possible product terms:
 $\{x_1z_1, x_1z_2, x_1z_3, x_2z_1, x_2z_2, x_2z_3, x_3z_1, x_3z_2, x_3z_3\}$.



Orthogonalization Procedure

Say we think that Z moderates the $X \rightarrow Y$ effect.

- X , Y , and Z are latent variables indicated by $\{x_1, x_2, x_3\}$, $\{y_1, y_2, y_3\}$, and $\{z_1, z_2, z_3\}$, respectively.

We perform the orthogonalization procedure as follows:

1. Construct all possible product terms:
 $\{x_1z_1, x_1z_2, x_1z_3, x_2z_1, x_2z_2, x_2z_3, x_3z_1, x_3z_2, x_3z_3\}$.
2. Regress each product term onto all observed indicators of X and Z :

$$\widehat{x_1z_1} = \alpha + \beta_1x_1 + \beta_2x_2 + \beta_3x_3 + \beta_4z_1 + \beta_5z_2 + \beta_6z_3$$

$$\widehat{x_2z_1} = \alpha + \beta_1x_1 + \beta_2x_2 + \beta_3x_3 + \beta_4z_1 + \beta_5z_2 + \beta_6z_3$$

$$\vdots$$

$$\widehat{x_3z_3} = \alpha + \beta_1x_1 + \beta_2x_2 + \beta_3x_3 + \beta_4z_1 + \beta_5z_2 + \beta_6z_3$$

Orthogonalization Procedure

3. Calculate each product term's residual:

$$\delta_{x_1z_1} = x_1z_1 - \widehat{x_1z_1}$$

$$\delta_{x_2z_1} = x_2z_1 - \widehat{x_2z_1}$$

$$\vdots$$

$$\delta_{x_3z_3} = x_3z_3 - \widehat{x_3z_3}$$



Orthogonalization Procedure

3. Calculate each product term's residual:

$$\delta_{x_1z_1} = x_1z_1 - \widehat{x_1z_1}$$

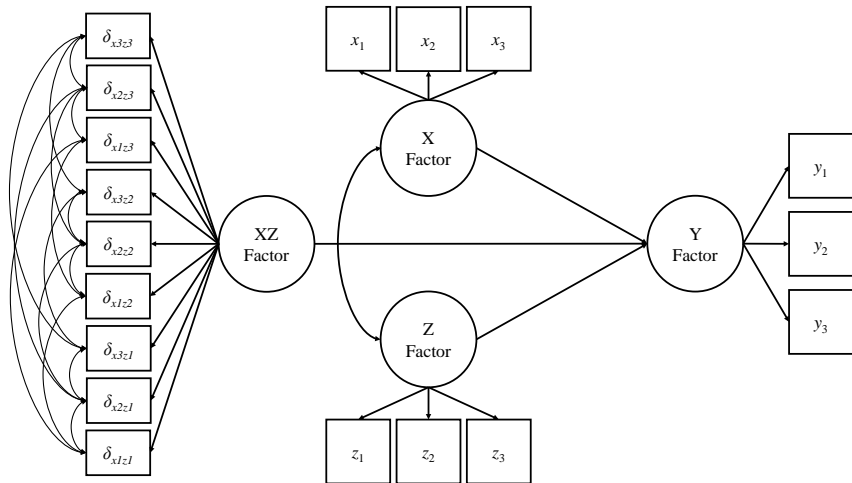
$$\delta_{x_2z_1} = x_2z_1 - \widehat{x_2z_1}$$

$$\vdots$$

$$\delta_{x_3z_3} = x_3z_3 - \widehat{x_3z_3}$$

4. Use these residuals to indicate a latent interaction construct.

Orthogonalization Diagram



Example

First, we check the measurement model.

```
## Read in some data:
dat1 <- readRDS("../data/lecture12Data.rds")

## Define the CFA model:
mod1 <- '
fX =~ x1 + x2 + x3
fZ =~ z1 + z2 + z3
fY =~ y1 + y2 + y3
'

## Estimate the model:
out1 <- cfa(mod1, data = dat1, std.lv = TRUE)
```

Example

```
partSummary(out1, 1:6)
```

```
lavaan 0.6-18 ended normally after 17 iterations
```

Estimator	ML
Optimization method	NLMINB
Number of model parameters	21
Number of observations	500

```
Model Test User Model:
```

Test statistic	41.021
Degrees of freedom	24
P-value (Chi-square)	0.017

```
Parameter Estimates:
```

Standard errors	Standard
Information	Expected
Information saturated (h1) model	Structured

Example

```
partSummary(out1, 7)
```

Latent Variables:

	Estimate	Std.Err	z-value	P(> z)
fX =~				
x1	0.671	0.044	15.407	0.000
x2	0.661	0.043	15.226	0.000
x3	0.702	0.045	15.481	0.000
fZ =~				
z1	0.738	0.048	15.343	0.000
z2	0.734	0.048	15.157	0.000
z3	0.718	0.046	15.601	0.000
fY =~				
y1	0.787	0.045	17.614	0.000
y2	0.729	0.045	16.325	0.000
y3	0.761	0.043	17.797	0.000

Example

```
partSummary(out1, 8)
```

Covariances:

	Estimate	Std.Err	z-value	P(> z)
fX ~~				
fZ	0.232	0.058	3.987	0.000
fY	0.827	0.033	25.310	0.000
fZ ~~				
fY	0.156	0.057	2.739	0.006

Example

```
partSummary(out1, 9)
```

Variances:

	Estimate	Std.Err	z-value	P(> z)
.x1	0.510	0.042	11.998	0.000
.x2	0.514	0.042	12.141	0.000
.x3	0.550	0.046	11.938	0.000
.z1	0.523	0.052	10.141	0.000
.z2	0.546	0.052	10.443	0.000
.z3	0.461	0.048	9.706	0.000
.y1	0.492	0.044	11.185	0.000
.y2	0.545	0.044	12.253	0.000
.y3	0.444	0.040	11.007	0.000
fX	1.000			
fZ	1.000			
fY	1.000			

Example

```
fitMeasures(out1, c("chisq", "df", "pvalue", "cfi", "tli", "rmsea", "srmr"))
```

chisq	df	pvalue	cfi	tli	rmsea	srmr
41.021	24.000	0.017	0.987	0.981	0.038	0.026

Example

Now, we'll fit the additive model as an SEM.

```
## Define the additive structural model:
mod2 <- '
fX =~ x1 + x2 + x3
fZ =~ z1 + z2 + z3
fY =~ y1 + y2 + y3

fY ~ fX + fZ
'

## Fit the model:
out2 <- sem(mod2, data = dat1, std.lv = TRUE)
```


Example

```
partSummary(out2, 1:6)
```

```
lavaan 0.6-18 ended normally after 22 iterations
```

Estimator	ML
Optimization method	NLMINB
Number of model parameters	21
Number of observations	500

```
Model Test User Model:
```

Test statistic	41.021
Degrees of freedom	24
P-value (Chi-square)	0.017

```
Parameter Estimates:
```

Standard errors	Standard
Information	Expected
Information saturated (h1) model	Structured

Example

```
partSummary(out2, 7)
```

Latent Variables:

	Estimate	Std.Err	z-value	P(> z)
fX =~				
x1	0.671	0.044	15.407	0.000
x2	0.661	0.043	15.226	0.000
x3	0.702	0.045	15.481	0.000
fZ =~				
z1	0.738	0.048	15.343	0.000
z2	0.734	0.048	15.157	0.000
z3	0.718	0.046	15.601	0.000
fY =~				
y1	0.442	0.044	10.079	0.000
y2	0.409	0.041	9.877	0.000
y3	0.427	0.042	10.099	0.000

Example

```
partSummary(out2, 8:9)
```

Regressions:

	Estimate	Std.Err	z-value	P(> z)
fY ~				
fX	1.488	0.190	7.820	0.000
fZ	-0.066	0.090	-0.732	0.464

Covariances:

	Estimate	Std.Err	z-value	P(> z)
fX ~~				
fZ	0.232	0.058	3.987	0.000

Example

```
partSummary(out2, 10)
```

Variances:

	Estimate	Std.Err	z-value	P(> z)
.x1	0.510	0.042	11.998	0.000
.x2	0.514	0.042	12.141	0.000
.x3	0.550	0.046	11.938	0.000
.z1	0.523	0.052	10.141	0.000
.z2	0.546	0.052	10.443	0.000
.z3	0.461	0.048	9.706	0.000
.y1	0.492	0.044	11.185	0.000
.y2	0.545	0.044	12.253	0.000
.y3	0.444	0.040	11.007	0.000
fX	1.000			
fZ	1.000			
.fY	1.000			

Example

```
fitMeasures(out2, c("chisq", "df", "pvalue", "cfi", "tli", "rmsea", "srmr"))
```

chisq	df	pvalue	cfi	tli	rmsea	srmr
41.021	24.000	0.017	0.987	0.981	0.038	0.026

Example

```
library(dplyr) # For data processing routines

## Set aside a copy of the predictor data for later use:
preds <- select(dat1, matches("x\\d|z\\d")) %>% as.matrix()

## Construct product terms:
products <- mutate(dat1,
  x1z1 = x1 * z1,
  x1z2 = x1 * z2,
  x1z3 = x1 * z3,

  x2z1 = x2 * z1,
  x2z2 = x2 * z2,
  x2z3 = x2 * z3,

  x3z1 = x3 * z1,
  x3z2 = x3 * z2,
  x3z3 = x3 * z3,
  .keep = "none")
```

Example

```
## Residualize the product terms:
products <- sapply(products,
                    function(y, x) lm(y ~ x)$resid,
                    x = preds)

## Join data pieces:
dat2 <- data.frame(dat1, products)
```

Example

```
mod3 <- '  
fX  =~ x1 + x2 + x3  
fZ  =~ z1 + z2 + z3  
fY  =~ y1 + y2 + y3  
fXZ =~ x1z1 + x1z2 + x1z3 + x2z1 + x2z2 + x2z3 + x3z1 + x3z2 + x3z3  
  
fY ~ fX + fZ + fXZ  
  
fX ~~ 0*fXZ  
fZ ~~ 0*fXZ  
  
x1z1 ~~ x1z2 + x1z3 + x2z1 + x3z1  
x1z2 ~~ x1z3 + x2z2 + x3z2  
x1z3 ~~ x2z3 + x3z3  
  
x2z1 ~~ x2z2 + x2z3 + x3z1  
x2z2 ~~ x2z3 + x3z2  
x2z3 ~~ x3z3  
  
x3z1 ~~ x3z2 + x3z3  
x3z2 ~~ x3z3  
'
```


Example

```
## Estimate the model:
```

```
out3 <- sem(mod3, data = dat2, std.lv = TRUE, meanstructure = TRUE)
```

```
partSummary(out3, 7, 1:14)
```

Latent Variables:

	Estimate	Std.Err	z-value	P(> z)
fX =~				
x1	0.670	0.043	15.424	0.000
x2	0.660	0.043	15.256	0.000
x3	0.704	0.045	15.569	0.000
fZ =~				
z1	0.738	0.048	15.342	0.000
z2	0.734	0.048	15.156	0.000
z3	0.718	0.046	15.602	0.000
fY =~				
y1	0.396	0.046	8.545	0.000
y2	0.369	0.044	8.441	0.000
y3	0.383	0.045	8.558	0.000

Example

```
partSummary(out3, 7, c(1, 2, 15:24))
```

Latent Variables:

	Estimate	Std.Err	z-value	P(> z)
fxZ =~				
x1z1	0.361	0.053	6.833	0.000
x1z2	0.427	0.056	7.615	0.000
x1z3	0.432	0.053	8.190	0.000
x2z1	0.558	0.056	9.914	0.000
x2z2	0.616	0.062	10.008	0.000
x2z3	0.520	0.057	9.153	0.000
x3z1	0.516	0.059	8.805	0.000
x3z2	0.626	0.063	10.007	0.000
x3z3	0.521	0.058	8.936	0.000

Example

```
partSummary(out3, 8:9, -(14:39))
```

Regressions:

	Estimate	Std.Err	z-value	P(> z)
fY ~				
fX	1.658	0.239	6.930	0.000
fZ	-0.074	0.099	-0.750	0.453
fXZ	0.488	0.120	4.049	0.000

Covariances:

	Estimate	Std.Err	z-value	P(> z)
fX ~~				
fXZ	0.000			
fZ ~~				
fXZ	0.000			
fX ~~				
fZ	0.232	0.058	3.987	0.000

Example

```
fitMeasures(out3, c("chisq", "df", "pvalue", "cfi", "tli", "rmsea", "srmr"))
```

chisq	df	pvalue	cfi	tli	rmsea	srmr
74.899	113.000	0.998	1.000	1.015	0.000	0.019

```
## Test simple slopes:
```

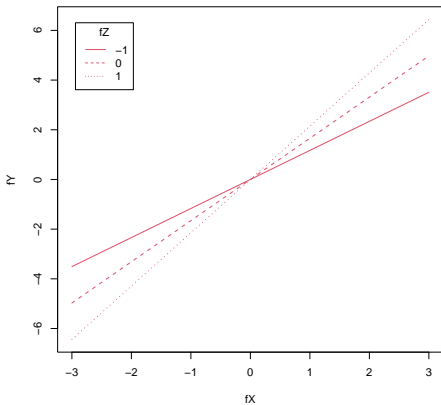
```
probeOut3 <- probe2WayRC(fit      = out3,  
                          nameX    = c("fX", "fZ", "fXZ"),  
                          nameY    = "fY",  
                          modVar   = "fZ",  
                          valProbe = c(-1, 0, 1)  
                          )
```

```
probeOut3$SimpleSlope
```

	fZ	est	se	z	pvalue
1	-1	1.170	0.207	5.653	0
2	0	1.658	0.238	6.974	0
3	1	2.145	0.310	6.923	0

Example

```
plotProbe(probeOut3, xlim = c(-3, 3), xlab = "fX", ylab = "fY")
```



Double Mean Centering Procedure

We can also define the interaction factor with double mean centering.

1. Mean center every indicator of X and Z :

$$x_1^c = x_1 - \bar{x}_1$$

$$\vdots$$

$$z_1^c = z_1 - \bar{z}_1$$

$$\vdots$$


Double Mean Centering Procedure

We can also define the interaction factor with double mean centering.

1. Mean center every indicator of X and Z :

$$x_1^c = x_1 - \bar{x}_1$$

$$\vdots$$

$$z_1^c = z_1 - \bar{z}_1$$

$$\vdots$$

2. Use the centered indicators to construct all possible product terms:
 $\{x_1^c z_1^c, x_1^c z_2^c, x_1^c z_3^c, x_2^c z_1^c, x_2^c z_2^c, x_2^c z_3^c, x_3^c z_1^c, x_3^c z_2^c, x_3^c z_3^c\}.$

Double Mean Centering Procedure

3. Mean center each product term:

$$(x_1 z_1)^c = x_1^c z_1^c - \overline{x_1^c z_1^c}$$

$$(x_1 z_2)^c = x_1^c z_2^c - \overline{x_1^c z_2^c}$$

$$\vdots$$

$$(x_3 z_3)^c = x_3^c z_3^c - \overline{x_3^c z_3^c}$$



Double Mean Centering Procedure

3. Mean center each product term:

$$(x_1z_1)^c = x_1^cz_1^c - \overline{x_1^cz_1^c}$$

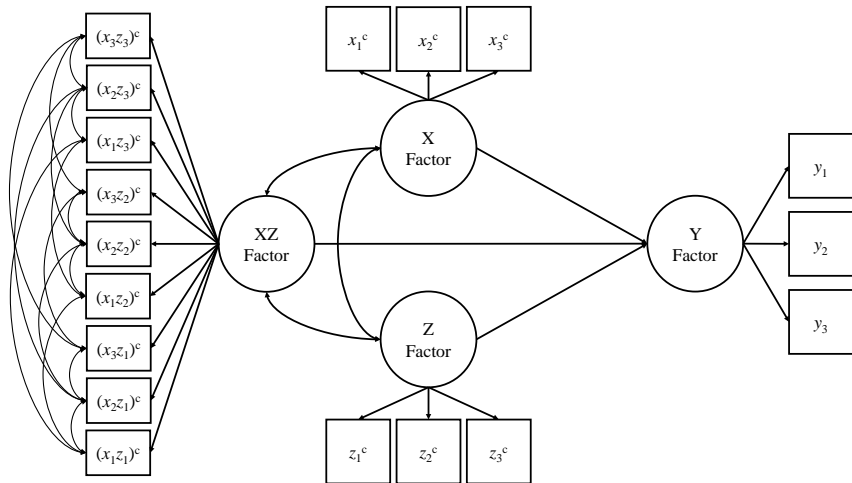
$$(x_1z_2)^c = x_1^cz_2^c - \overline{x_1^cz_2^c}$$

$$\vdots$$

$$(x_3z_3)^c = x_3^cz_3^c - \overline{x_3^cz_3^c}$$

4. Use the mean centered indicators of X and Z , and the “double mean centered” product terms to specify the latent interaction model.

Double Mean Centering Diagram



Example

```
## Mean-center the predictor variables:
```

```
preds <- scale(preds, scale = FALSE) %>% as.data.frame()
```

```
## Construct and mean-center the product terms:
```

```
products <- mutate(preds,  
  x1z1 = x1 * z1,  
  x1z2 = x1 * z2,  
  x1z3 = x1 * z3,  
  
  x2z1 = x2 * z1,  
  x2z2 = x2 * z2,  
  x2z3 = x2 * z3,  
  
  x3z1 = x3 * z1,  
  x3z2 = x3 * z2,  
  x3z3 = x3 * z3,  
  .keep = "none") %>%  
  scale(scale = FALSE)
```

```
## Join the data pieces:
```

```
dat3 <- select(dat1, matches("y\\d")) %>% data.frame(preds, products)
```

Example

```
mod4 <- '  
fX =~ x1 + x2 + x3  
fZ =~ z1 + z2 + z3  
fY =~ y1 + y2 + y3  
fXZ =~ x1z1 + x1z2 + x1z3 + x2z1 + x2z2 + x2z3 + x3z1 + x3z2 + x3z3  
  
fY ~ fX + fZ + fXZ  
  
x1z1 ~~ x1z2 + x1z3 + x2z1 + x3z1  
x1z2 ~~ x1z3 + x2z2 + x3z2  
x1z3 ~~ x2z3 + x3z3  
  
x2z1 ~~ x2z2 + x2z3 + x3z1  
x2z2 ~~ x2z3 + x3z2  
x2z3 ~~ x3z3  
  
x3z1 ~~ x3z2 + x3z3  
x3z2 ~~ x3z3  
'
```

Example

```
## Estimate the model:  
out4 <- sem(mod4, data = dat3, std.lv = TRUE)  
  
partSummary(out4, 7, 1:14)
```

Latent Variables:

	Estimate	Std.Err	z-value	P(> z)
fX =~				
x1	0.673	0.043	15.555	0.000
x2	0.659	0.043	15.260	0.000
x3	0.702	0.045	15.569	0.000
fZ =~				
z1	0.738	0.048	15.360	0.000
z2	0.734	0.048	15.154	0.000
z3	0.718	0.046	15.597	0.000
fY =~				
y1	0.386	0.048	8.009	0.000
y2	0.359	0.045	7.925	0.000
y3	0.373	0.047	8.018	0.000

Example

```
partSummary(out4, 7, c(1, 2, 15:24))
```

Latent Variables:

	Estimate	Std.Err	z-value	P(> z)
fxZ =~				
x1z1	0.367	0.053	6.902	0.000
x1z2	0.434	0.056	7.715	0.000
x1z3	0.441	0.053	8.300	0.000
x2z1	0.550	0.056	9.788	0.000
x2z2	0.616	0.062	9.970	0.000
x2z3	0.519	0.057	9.115	0.000
x3z1	0.504	0.059	8.604	0.000
x3z2	0.628	0.063	10.039	0.000
x3z3	0.535	0.059	9.128	0.000

Example

```
partSummary(out4, 8:9, -(10:35))
```

Regressions:

	Estimate	Std.Err	z-value	P(> z)
fY ~				
fX	1.757	0.270	6.515	0.000
fZ	-0.111	0.105	-1.062	0.288
fXZ	0.557	0.141	3.962	0.000

Covariances:

	Estimate	Std.Err	z-value	P(> z)
fX ~~				
fZ	0.232	0.058	3.987	0.000
fXZ	-0.087	0.067	-1.297	0.195
fZ ~~				
fXZ	0.040	0.066	0.613	0.540

Example

```
## Check model fit:
fitMeasures(out4, c("chisq", "df", "pvalue", "cfi", "tli", "rmsea", "srmr"))
```

chisq	df	pvalue	cfi	tli	rmsea	srmr
134.186	111.000	0.066	0.993	0.991	0.020	0.030

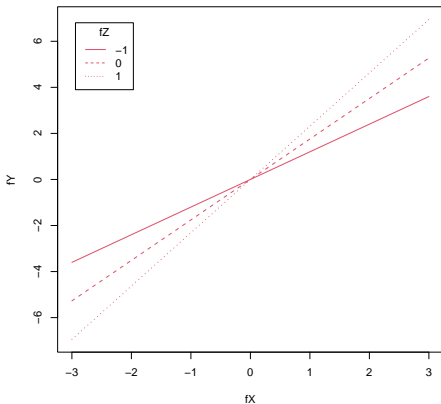
```
## Test simple slopes:
probeOut4 <- probe2WayMC(fit      = out4,
                        nameX     = c("fX", "fZ", "fXZ"),
                        nameY     = "fY",
                        modVar    = "fZ",
                        valProbe  = c(-1, 0, 1)
                        )

probeOut4$SimpleSlope
```

	fZ	est	se	z	pvalue
1	-1	1.200	0.210	5.722	0
2	0	1.757	0.270	6.515	0
3	1	2.314	0.376	6.163	0

Example

```
plotProbe(probeOut4, xlim = c(-3, 3), xlab = "fX", ylab = "fY")
```



Orthogonalization vs. Double Mean Centering

Orthogonalization and double mean centering tend to behave comparably, but each has its own strengths:

- When X and Z are bivariate normally distributed, both methods produce the same results.
- As X and/or Z stray from normality, orthogonalization produces biased estimates of the interaction effect, but double mean centering does not.
- Orthogonalization ensures that the latent XZ is perfectly independent of X and Z .
 - The X and Z parameters can be directly interpreted, without any conditioning

Example

We can also use the `indProd()` function from **semTools** to create the product indicators.

```
## Use semTools to orthogonalize:
dat2.2 <- indProd(data      = dat1,
                  var1      = c("x1", "x2", "x3"),
                  var2      = c("z1", "z2", "z3"),
                  match     = FALSE,
                  meanC     = FALSE,
                  doubleMC  = FALSE,
                  residualC = TRUE,
                  namesProd = colnames(products)
                  )
```

```
## Compare to our manual results:
all.equal(dat2[colnames(products)],
          dat2.2[colnames(products)],
          check.attributes = FALSE)
```

```
[1] TRUE
```

Example

```
## Use semTools to double mean center:
dat3.2 <- indProd(data      = dat1,
                  var1      = c("x1", "x2", "x3"),
                  var2      = c("z1", "z2", "z3"),
                  match     = FALSE,
                  meanC     = TRUE,
                  doubleMC  = TRUE,
                  residualC = FALSE,
                  namesProd = colnames(products)
                  )

all.equal(dat3[colnames(products)],
          dat3.2[colnames(products)],
          check.attributes = FALSE)

[1] TRUE
```

Estimating Products of Latent Variables

In theory, we can directly estimate a latent variable defined as the product of two (or more) other latent variables.

- We cannot use ordinary ML estimation methods.

Three approaches seem sensible.

1. Latent moderated structural equations (LMS; Klein & Moosbrugger, 2000; Klein, Moosbrugger, Schermelleh-Engel, & Frank, 1997).
2. Structural-after-measurement (SAM) estimation (Rosseel & Loh, 2022).
3. Bayesian SEM.

Only the SAM approach is currently implemented in **lavaan**.

- The current implementation cannot probably handle latent variable interactions.

Estimating Products of Latent Variables

We can directly estimate the interaction between two latent variables with the *latent moderated structural equations* (LMS) method.

- Introduced by Klein et al. (1997) and formalized by Klein and Moosbrugger (2000)
- Uses numerical integration to estimate the unobserved latent interaction term
- Traditionally, only available in Mplus (via the `Xwith` command)
- Now available in R through **modsem** (Solem Slupphaug, 2024)

LMS Example

```
## Load the modsem package:
library(modsem)

## Define lavaan syntax for our model:
mod <- "
fX =~ x1 + x2 + x3
fZ =~ z1 + z2 + z3
fY =~ y1 + y2 + y3

fY ~ fX + fZ + fX:fZ
"

## Estimate the model:
out <- modsem(mod, data = dat1, method = "lms")
```

LMS Example

```
partSummary(out, 1:4)
```

```
----Model Summary-----
```

```
Number of iterations: 29
```

```
Final loglikelihood: -5368.489
```

```
Comparative fit to H0 (no interaction effect)
```

```
Loglikelihood change = 419.70
```

```
D(1) = -839.39, p = <2e-16
```

```
R-squared:
```

```
  fY = 0.764
```

```
R-squared Null-Model (H0):
```

```
  fY = 0.684
```

```
R-squared Change:
```

```
  fY = 0.080
```


LMS Example

```
partSummary(out, 5:6)
```

```
-----Estimates-----
```

Loadings:

fX:

variable	est	std.error	t.value	p.value	ci.lower	ci.upper
x1	1.000	NA	NA	NA	NA	NA
x2	0.991	0.068	14.633	< 2e-16	0.858	1.123
x3	1.049	0.071	14.774	< 2e-16	0.910	1.188

fZ:

variable	est	std.error	t.value	p.value	ci.lower	ci.upper
z1	1.000	NA	NA	NA	NA	NA
z2	0.994	0.083	11.935	< 2e-16	0.831	1.157
z3	0.961	0.079	12.093	< 2e-16	0.805	1.116

fY:

variable	est	std.error	t.value	p.value	ci.lower	ci.upper
y1	1.000	NA	NA	NA	NA	NA
y2	0.930	0.062	14.907	< 2e-16	0.808	1.053
y3	0.963	0.060	15.921	< 2e-16	0.845	1.082

LMS Example

```
partSummary(out, 7:8)
```

Regressions:

fY:

variable	est	std.error	t.value	p.value	ci.lower	ci.upper
fX	1.002	0.074	13.578	< 2e-16	0.857	1.147
fZ	-0.053	0.041	-1.288	0.19789	-0.133	0.028
fX:fZ	0.398	0.065	6.101	1.05e-09	0.270	0.526

Intercepts:

variable	est	std.error	t.value	p.value	ci.lower	ci.upper
x1	-0.010	0.032	-0.323	0.74669	-0.073	0.052
x2	-0.033	0.032	-1.019	0.30824	-0.095	0.030
x3	-0.027	0.033	-0.798	0.42474	-0.092	0.039
z1	0.036	0.046	0.780	0.43536	-0.054	0.125
z2	0.041	0.046	0.883	0.37711	-0.050	0.131
z3	0.028	0.044	0.653	0.51371	-0.057	0.114
y1	-0.022	0.037	-0.603	0.54619	-0.094	0.050
y2	0.014	0.037	0.388	0.69811	-0.059	0.088
y3	0.025	0.035	0.709	0.47817	-0.044	0.094

LMS Example

```
partSummary(out, 8)
```

Intercepts:

variable	est	std.error	t.value	p.value	ci.lower	ci.upper
x1	-0.010	0.032	-0.323	0.74669	-0.073	0.052
x2	-0.033	0.032	-1.019	0.30824	-0.095	0.030
x3	-0.027	0.033	-0.798	0.42474	-0.092	0.039
z1	0.036	0.046	0.780	0.43536	-0.054	0.125
z2	0.041	0.046	0.883	0.37711	-0.050	0.131
z3	0.028	0.044	0.653	0.51371	-0.057	0.114
y1	-0.022	0.037	-0.603	0.54619	-0.094	0.050
y2	0.014	0.037	0.388	0.69811	-0.059	0.088
y3	0.025	0.035	0.709	0.47817	-0.044	0.094

LMS Example

```
partSummary(out, 9)
```

Variances:

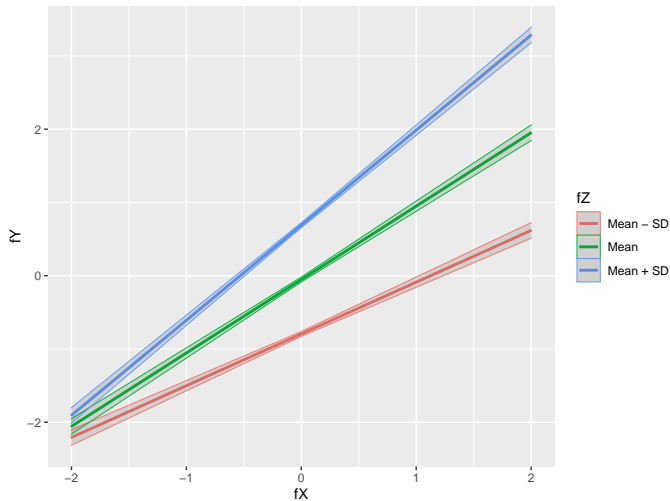
variable	est	std.error	t.value	p.value	ci.lower	ci.upper
x1	0.513	0.032	15.811	< 2e-16	0.449	0.576
x2	0.512	0.032	15.811	< 2e-16	0.449	0.576
x3	0.552	0.035	15.811	< 2e-16	0.484	0.621
z1	0.516	0.050	10.309	< 2e-16	0.418	0.614
z2	0.541	0.051	10.579	< 2e-16	0.441	0.641
z3	0.469	0.046	10.206	< 2e-16	0.379	0.559
y1	0.492	0.039	12.510	< 2e-16	0.415	0.570
y2	0.540	0.041	13.239	< 2e-16	0.460	0.620
y3	0.448	0.036	12.351	< 2e-16	0.377	0.519
fX	0.447	0.032	13.863	< 2e-16	0.383	0.510
fZ	0.112	0.038	2.963	0.00305	0.038	0.186
fZ	0.551	0.047	11.820	< 2e-16	0.459	0.642
fY	0.147	0.025	5.935	2.94e-09	0.099	0.196

LMS Example

We can also visualize the interaction.

```
sdZ <- out$parTable %>%  
  filter(lhs == "fZ", op == "~~", rhs == "fZ") %>%  
  extract2("est") %>%  
  sqrt()  
  
sdX <- out$parTable %>%  
  filter(lhs == "fX", op == "~~", rhs == "fX") %>%  
  extract2("est") %>%  
  sqrt()  
  
plot_interaction(model = out,  
  y = "fY",  
  x = "fX",  
  z = "fZ",  
  xz = "fX:fZ",  
  vals_x = c(-3 * sdX, 0, 3 * sdX),  
  vals_z = c(-sdZ, 0, sdZ))  
) + scale_color_discrete(labels = c("Mean - SD", "Mean", "Mean + SD"))
```

LMS Example



Estimating Products of Latent Variables

LMS Strengths:

- Tends to perform the best out of all available methods
 - Though, SAM may do better
- No need to pre-process the data by manually computing product terms
- Pretty easy to implement in Mplus or **modsem**.

LMS Weaknesses:

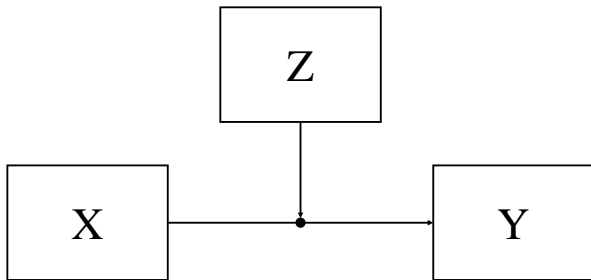
- The **modsem** implementation is quite limited, for now.
- Numerical integration is very slow and precludes calculation of most fit indices
- LMS does not work with categorical observed moderators

MULTIPLE MODERATION



Starting Point

So far, we've been looking at this type of model:



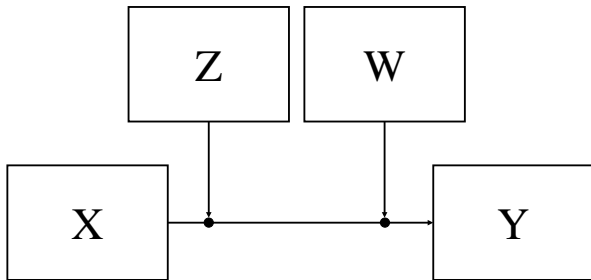
We've had one focal variable and one moderator.

- We've been asking questions about how the focal effect changes as a function of the moderator.
- There's no reason we need to restrict ourselves to a single moderator.

Multiple Moderation

Maybe we suspect that the focal effect changes as a function of two other variables.

- We could fit this type of model:



Now, the focal effect of X on Y changes as a function of both Z and W .

Multiple Moderation

The preceding diagram implies the following formula:

$$Y = \beta_0 + f(Z, W)X + \beta_2Z + \beta_3W + e,$$

Taking $f(Z, W)$ to be the following simple slope:

$$f(Z, W) = \beta_1 + \beta_4Z + \beta_5W$$

Produces the following analytic equation:

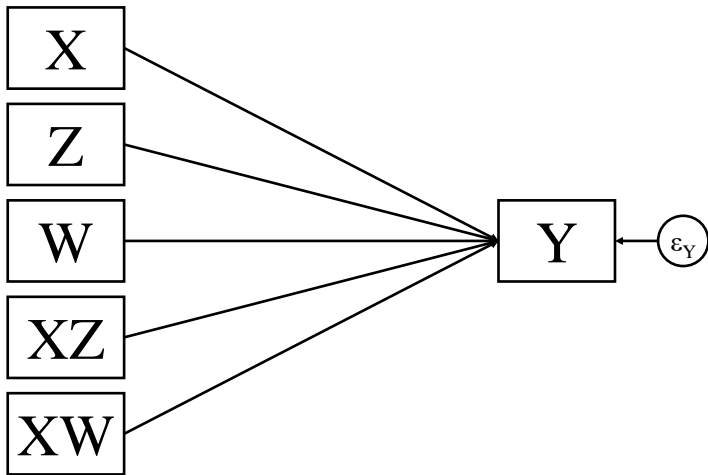
$$Y = \beta_0 + \beta_1X + \beta_2Z + \beta_3W + \beta_4XZ + \beta_5XW + e$$

We can fit this model in any regression (or path modeling) software.

- We can test for significant moderating effects of Z and W by testing for non-zero β_4 and β_5 , respectively.

Multiple Moderation

Our analytic diagram is predictably extended:

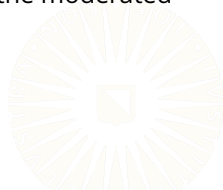


Moderated Moderation

The additive two-way interaction model is more flexible than the simple single-moderator model, but it still imposes constraints.

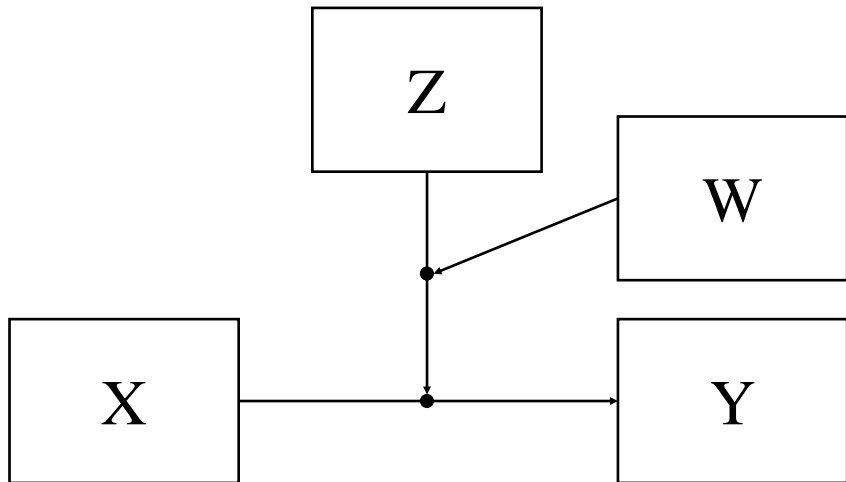
- The moderating effect of Z (or W) on the $X \rightarrow Y$ relation is assumed to be constant across levels of W (or Z).
- I.e., the moderation is not moderated

We can relax this constraint by modeling moderation of the moderated effect using a three-way interaction.



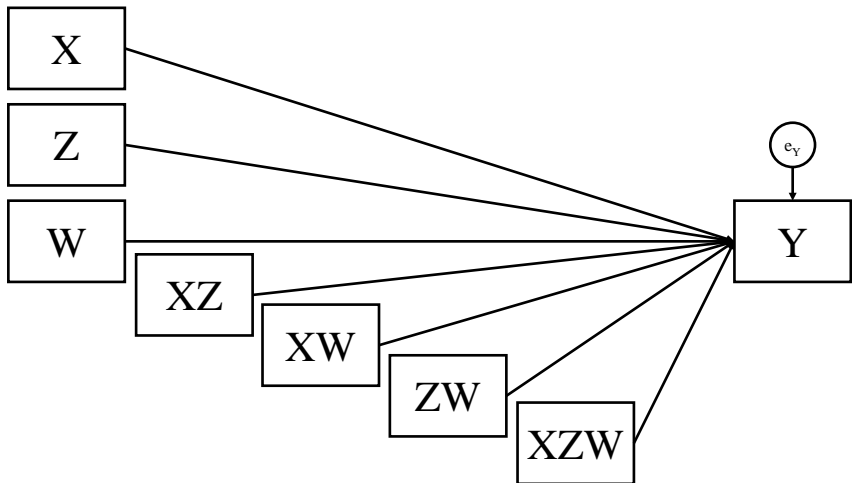
Moderated Moderation

Moderated moderation implies the following conceptual diagram:



Moderated Moderation

The preceding conceptual diagram implies this analytic diagram:



Moderated Moderation

The preceding diagram represents the following equation:

$$Y = \beta_0 + \beta_1 X + \beta_2 Z + \beta_3 W + \beta_4 XZ + \beta_5 XW + \beta_6 ZW + \beta_7 XZW + e$$

Which can be restructured into:

$$\begin{aligned} Y &= \beta_0 + (\beta_1 + \beta_4 Z + \beta_5 W + \beta_7 ZW)X + \beta_2 Z + \beta_3 W + \beta_6 ZW + e \\ &= \beta_0 + g(Z, W)X + \beta_2 Z + \beta_3 W + \beta_6 ZW + e \end{aligned}$$

With moderated moderation, the simple slope is given by:

$$g(Z, W) = \beta_1 + \beta_4 Z + \beta_5 W + \beta_7 ZW$$

Which has the same structure as a single moderator model.

Example

```
## Read in the BFI data:
dat1 <- readRDS("../data/bfiData1.rds")

## Three-way interaction model:
mod <- '
agree ~ 1 + open + conc + neuro + open:conc + open:neuro + conc:neuro + ocn
'

## Create the 3-way interaction via a pipeline and estimate the model:
out <- dat1 %>%
  mutate(ocn = open * conc * neuro) %>%
  sem(mod, data = .)
```

Example

```
partSummary(out, 7:9)
```

Regressions:

	Estimate	Std.Err	z-value	P(> z)
agree ~				
open	1.279	0.257	4.975	0.000
conc	1.208	0.265	4.557	0.000
neuro	0.738	0.322	2.292	0.022
open:conc	-0.297	0.069	-4.292	0.000
open:neuro	-0.216	0.081	-2.676	0.007
conc:neuro	-0.256	0.082	-3.114	0.002
ocn	0.065	0.020	3.230	0.001

Intercepts:

	Estimate	Std.Err	z-value	P(> z)
.agree	-0.587	0.965	-0.609	0.543

Variances:

	Estimate	Std.Err	z-value	P(> z)
.agree	0.481	0.013	35.721	0.000

Example

```
## Compute simple slopes by conditioning on both moderators:
ssOut <- probe3WayMC(out,
  nameX = c("open", "conc", "neuro",
            "open:conc", "open:neuro", "conc:neuro",
            "ocn"),
  nameY = "agree",
  modVar = c("conc", "neuro"),
  valProbe1 = quantile(dat1$conc, c(0.25, 0.5, 0.75)),
  valProbe2 = quantile(dat1$neuro, c(0.25, 0.5, 0.75))
)
```

Example

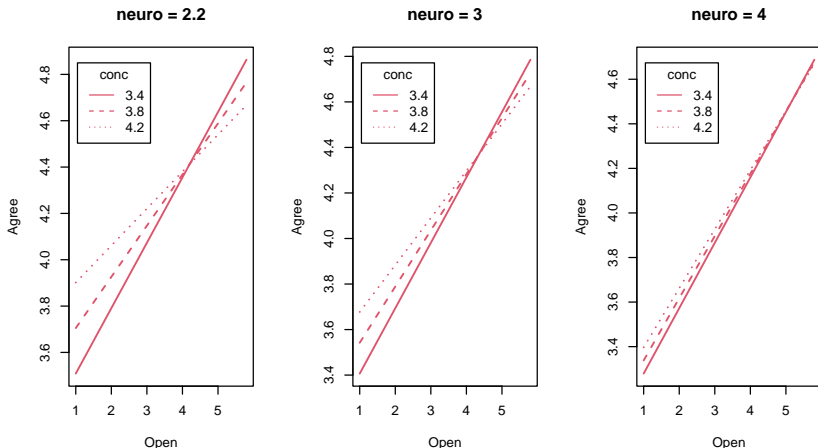
```
## View simple slopes:
```

```
ssOut$SimpleSlope
```

	conc	neuro	est	se	z	pvalue
1	3.4	2.2	0.282	0.034	8.316	0
2	3.8	2.2	0.221	0.035	6.400	0
3	4.2	2.2	0.160	0.042	3.802	0
4	3.4	3.0	0.287	0.031	9.346	0
5	3.8	3.0	0.247	0.027	8.980	0
6	4.2	3.0	0.206	0.032	6.379	0
7	3.4	4.0	0.293	0.041	7.179	0
8	3.8	4.0	0.279	0.032	8.607	0
9	4.2	4.0	0.265	0.033	8.032	0

Example

```
plotProbe(ssOut, xlim = range(dat1$open), xlab = "Open", ylab = "Agree")
```



CATEGORICAL MODERATORS



Categorical Moderators

Categorical moderators encode *group-specific* effects.

- E.g., if we include sex as a moderator, we are modeling separate focal effects for males and females.

Given a set of codes representing our moderator, we specify the interactions as before:

$$Y_{total} = \beta_0 + \beta_1 X_{inten} + \beta_2 Z_{male} + \beta_3 X_{inten} Z_{male} + \varepsilon$$

$$Y_{total} = \beta_0 + \beta_1 X_{inten} + \beta_2 Z_{lo} + \beta_3 Z_{mid} + \beta_4 Z_{hi} \\ + \beta_5 X_{inten} Z_{lo} + \beta_6 X_{inten} Z_{mid} + \beta_7 X_{inten} Z_{hi} + \varepsilon$$

Example

```
## Load data:
socSup <- readRDS("../data/social_support.rds")

## Focal effect model:
mod <- 'bdi ~ 1 + tanSat'

## Fit the model and summarize the results:
sem(mod, data = socSup) %>% partSummary(7:9)
```

Regressions:

	Estimate	Std.Err	z-value	P(> z)
bdi ~				
tanSat	-0.810	0.309	-2.621	0.009

Intercepts:

	Estimate	Std.Err	z-value	P(> z)
.bdi	24.409	5.294	4.611	0.000

Variances:

	Estimate	Std.Err	z-value	P(> z)
.bdi	84.261	12.226	6.892	0.000

Example

```
## Moderated model:
mod <- 'bdi ~ 1 + tanSat + male + tanSat:male'

## Dummy code sex in a pipeline and estimate the model:
out <- socSup %>%
  mutate(male = as.numeric(sex == "male")) %>%
  sem(mod, data = .)
```

Example

```
partSummary(out, 7:9)
```

Regressions:

	Estimate	Std.Err	z-value	P(> z)
bdi ~				
tanSat	-0.577	0.354	-1.632	0.103
male	14.367	11.946	1.203	0.229
tanSat: male	-0.948	0.702	-1.350	0.177

Intercepts:

	Estimate	Std.Err	z-value	P(> z)
.bdi	20.848	6.079	3.429	0.001

Variances:

	Estimate	Std.Err	z-value	P(> z)
.bdi	82.261	11.936	6.892	0.000

Example

```
## Test simple slopes and intercepts:
```

```
ssOut <- probe2WayMC(out,  
                      nameX    = c("tanSat", "male", "tanSat:male"),  
                      nameY    = "bdi",  
                      modVar   = "male",  
                      valProb  = 0:1)
```

```
ssOut
```

```
$SimpleIntcept
```

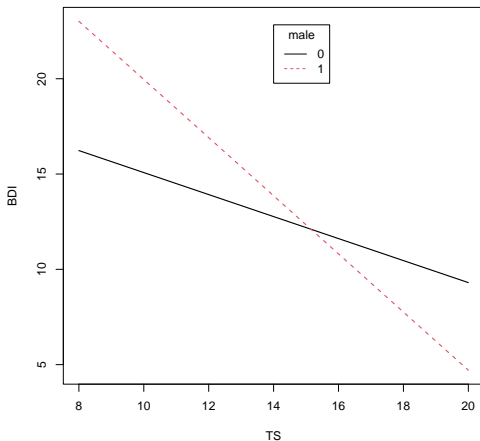
	male	est	se	z	pvalue
1	0	20.848	6.079	3.429	0.001
2	1	35.215	10.283	3.425	0.001

```
$SimpleSlope
```

	male	est	se	z	pvalue
1	0	-0.577	0.354	-1.632	0.103
2	1	-1.525	0.607	-2.513	0.012

Example

```
plotProbe(ssOut, xlim = range(socSup$stanSat), xlab = "TS", ylab = "BDI")
```



Visualizing Categorical Moderation

$$\hat{Y}_{BDI} = 20.85 - 0.58X_{tsat} + 14.37Z_{male} - 0.95X_{tsat}Z_{male}$$

$$\hat{Y}_{BDI} = 28.10 - 1.00X_{tsat} - 1.05Z_{male}$$

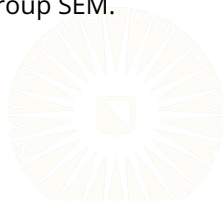
```
Error in 'geom_jitter()':  
! Problem while converting geom to  
grob.  
i Error occurred in the 1st layer.  
Caused by error:  
! Unknown colour name: male
```

```
Error in 'geom_jitter()':  
! Problem while converting geom to  
grob.  
i Error occurred in the 1st layer.  
Caused by error:  
! Unknown colour name: male
```

Moderation via Multiple Group SEM

When our moderator is a categorical variable, we can use multiple group CFA/SEM to test for moderation.

- Categorical moderators define groups.
- Significant moderation with categorical moderators implies between-group differences in the focal effect.
- We can directly test these hypotheses with multiple group SEM.



Example

```
## Read the data and subset to only high school and college graduates:
dat2 <- readRDS("../data/bfiData2.rds") %>%
  filter(educ %in% c("highSchool", "college"))

## Specify the (configurally invariance) measurement model:
mod0 <- '
agree =~ A1 + A2 + A3 + A4 + A5
open  =~ O1 + O2 + O3 + O4 + O5
'

## Estimate the unrestricted model:
out0 <- cfa(mod0, data = dat2, std.lv = TRUE, group = "educ")
```

Example

```
## Define the weakly invariant model:
mod1 <- measEq.syntax(configural.model = out0,
                      group = "educ",
                      group.equal = "loadings") %>%
  as.character()

## Define the strongly invariant model:
mod2 <- measEq.syntax(configural.model = out0,
                      group = "educ",
                      group.equal = c("loadings", "intercepts")
                      ) %>%
  as.character()

## Estimate the models:
out1 <- cfa(mod1, data = dat2, group = "educ")
out2 <- cfa(mod2, data = dat2, group = "educ")

## Test measurement invariance:
compareFit(out0, out1, out2) %>% summary()
```


Example

Nested Model Comparison

Chi-Squared Difference Test

	Df	AIC	BIC	Chisq	Chisq diff	RMSEA	Df diff	Pr(>Chisq)
out0	68	75336	75694	418.25				
out1	76	75358	75669	455.70	37.451	0.055855	8	9.505e-06 ***
out2	84	75461	75726	575.15	119.446	0.108654	8	< 2.2e-16 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Model Fit Indices

	chisq	df	pvalue	rmsea	cfi	tli	srmr	aic	bic
out0	418.253†	68	.000	.066	.906†	.875	.046†	75336.292†	75693.810
out1	455.705	76	.000	.065†	.898	.879†	.050	75357.744	75669.130†
out2	575.151	84	.000	.070	.868	.858	.056	75461.190	75726.445

Differences in Fit Indices

	df	rmsea	cfi	tli	srmr	aic	bic
out1 - out0	8	-0.001	-0.008	0.004	0.004	21.451	-24.680
out2 - out1	8	0.005	-0.030	-0.021	0.006	103.446	57.315

Example

```
## Specify a structural model:
mod3 <- '
agree =~ A1 + A2 + A3 + A4 + A5
open  =~ O1 + O2 + O3 + O4 + O5

agree ~ open
'

## Estimate the model with strong invariance constraints:
out3 <- sem(mod3,
             dat = dat2,
             std.lv = TRUE,
             group = "educ",
             group.equal = c("loadings", "intercepts")
             )
```

Example

```
## Check the group-specific slopes:
```

```
partSummary(out3, c(8, 10, 14, 16))
```

Group 1 [highSchool]:

Regressions:

	Estimate	Std.Err	z-value	P(> z)
agree ~				
open	-0.321	0.040	-7.957	0.000

Group 2 [college]:

Regressions:

	Estimate	Std.Err	z-value	P(> z)
agree ~				
open	-0.203	0.051	-3.972	0.000

Example

```
## Specify the restricted model:
mod4 <- '
agree =~ A1 + A2 + A3 + A4 + A5
open  =~ O1 + O2 + O3 + O4 + O5

agree ~ c(beta, beta) * open
'

## Estimate the model:
out4 <- sem(mod4,
            dat = dat2,
            std.lv = TRUE,
            group = "educ",
            group.equal = c("loadings", "intercepts")
            )
```

Example

```
## Check the slopes:
```

```
partSummary(out4, c(8, 10, 14, 16))
```

Group 1 [highSchool]:

Regressions:

		Estimate	Std.Err	z-value	P(> z)
agree ~					
open	(beta)	-0.278	0.032	-8.621	0.000

Group 2 [college]:

Regressions:

		Estimate	Std.Err	z-value	P(> z)
agree ~					
open	(beta)	-0.278	0.032	-8.621	0.000

Example

```
## Do a chi-squared difference test for moderation:
```

```
anova(out3, out4)
```

Chi-Squared Difference Test

	Df	AIC	BIC	Chisq	Chisq diff	RMSEA	Df diff	Pr(>Chisq)
out3	84	75461	75726	575.15				
out4	85	75463	75722	578.59	3.435	0.045426	1	0.06383 .

Signif. codes:	0	'***'	0.001	'**'	0.01	'*'	0.05	'.' 0.1 ' ' 1

Example

```
## Do a similar test via OLS regression:
dat1 %>%
  filter(educ %in% c("highSchool", "college")) %$%
  lm(agree ~ open * educ) %>%
  partSummary(-(1:2))
```

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	3.24965	0.12321	26.376	<2e-16
open	0.27115	0.03134	8.652	<2e-16
educcollege	0.03975	0.22849	0.174	0.862
open:educcollege	-0.05654	0.05856	-0.965	0.334

Residual standard error: 0.6972 on 2356 degrees of freedom

Multiple R-squared: 0.05314, Adjusted R-squared: 0.05194

F-statistic: 44.08 on 3 and 2356 DF, p-value: < 2.2e-16

Probing Multiple Group Moderation

Testing moderation with multiple group SEM has several advantages.

- Remove measurement error from the estimates
- Test for factorial invariance
- All simple effects are directly estimated in the unrestricted model



Simple Slopes & Intercepts

Group 1 [highSchool]:

Regressions:

	Estimate	Std.Err	z-value	P(> z)
agree ~				
open	-0.321	0.040	-7.957	0.000

Intercepts:

Estimate	Std.Err	z-value	P(> z)
----------	---------	---------	---------

Regressions:

	Estimate	Std.Err	z-value	P(> z)
agree ~				
open	-0.203	0.051	-3.972	0.000

Intercepts:

		Estimate	Std.Err	z-value	P(> z)
.A1	(.24.)	2.292	0.030	76.639	0.000
.A2	(.25.)	4.862	0.027	180.210	0.000

Simple Slopes Visualized

We can visualize the simple slopes by plotting the factor scores.

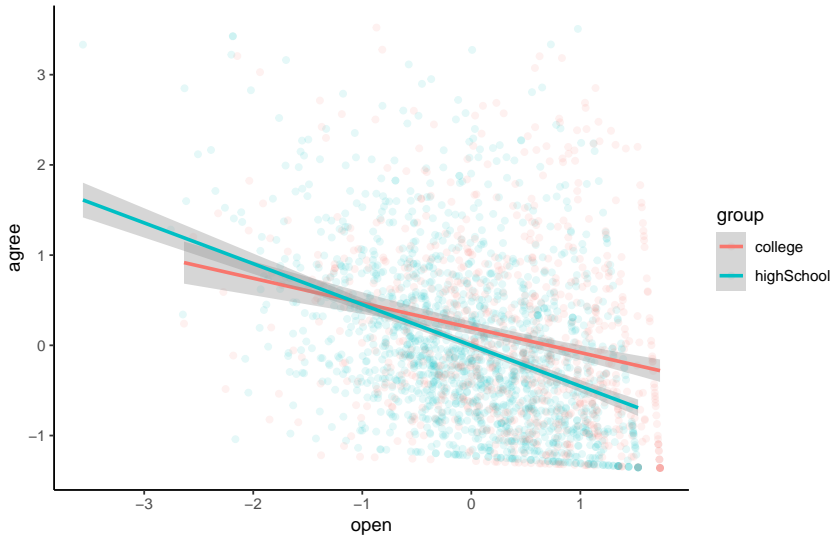
```
library(ggplot2)

## Generate factor scores:
tmp <- predict(out3)

## Stack factor scores into a "tidy" dataset:
pData <- data.frame(do.call(rbind, tmp),
                    group = rep(names(tmp), sapply(tmp, nrow))
                    )

## Create a simple slopes plot:
ssPlot <- ggplot(pData, aes(open, agree, color = group)) +
  geom_point(alpha = 0.1) +
  geom_smooth(method = "lm") +
  theme_classic()
```

Simple Slopes Visualized



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