### **More on functional forms**

EC 339

Marcio Santetti Fall 2022

# Motivation

### New functional forms

There is more to OLS than linear-in-variables models or log-transformed models.

But do these models **preserve** OLS *Classical Assumptions*?

- They do!
- But under what conditions?

As long as the model remains linear in parameters, everything is fine.

### New functional forms

- 1. Regression through the origin
- 2. Regression with quadratic terms
- **3**. **Inverse** forms
- 4. Interaction terms
- **5**. **Binary** (*dummy*) variables

Regression through the origin

# Regression through the origin

It is used whenever we need to impose the **restriction** that, when x=0, the expected value of y is also zero.

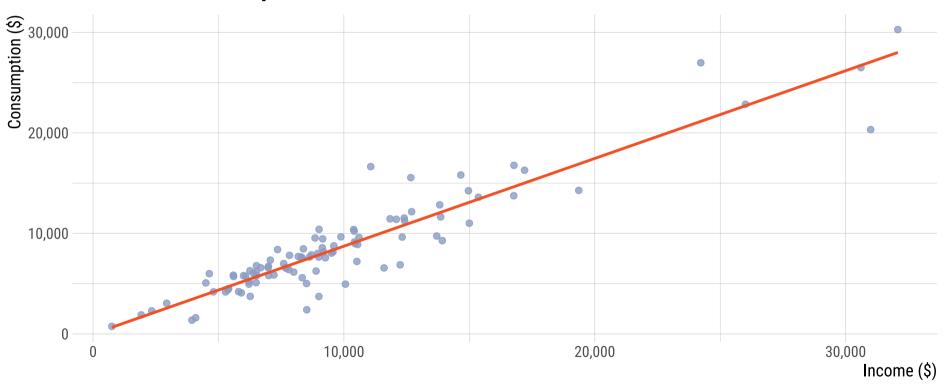
It should be applied **only** when theory recommends to do so.

$$y_i = \beta_1 x_{1i} + u_i$$

# Regression through the origin

$$Cons_i = \beta_1 Inc_i + u_i$$

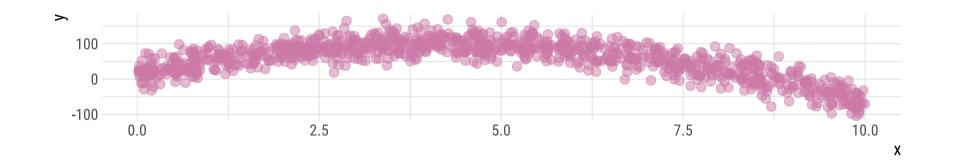
#### **Income vs. Consumption**

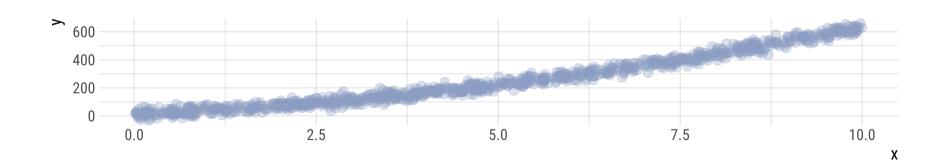


Many times, the effect of a variable  $x_i$  on y also depends on the **level** of that independent variable.

We can also apply quadratic terms when the effect of  $x_i$  on y changes after a given threshold.

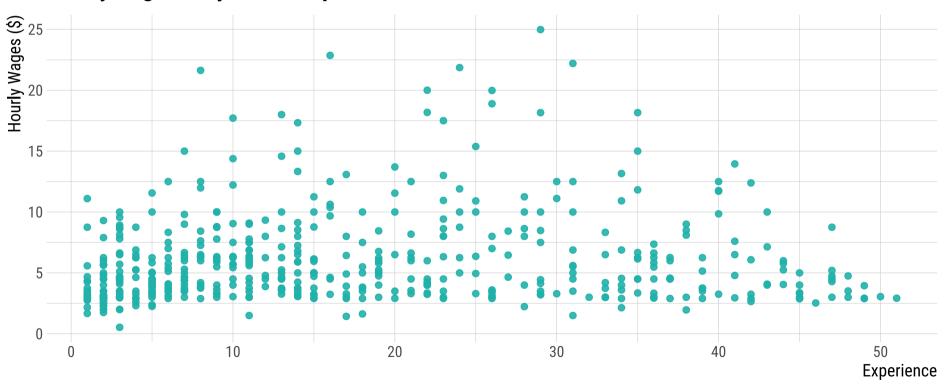
$$y_i = eta_0 + eta_1 x_{1i} + eta_2 (x_{1i})^2 + \dots + eta_k x_{ki} + u_i$$





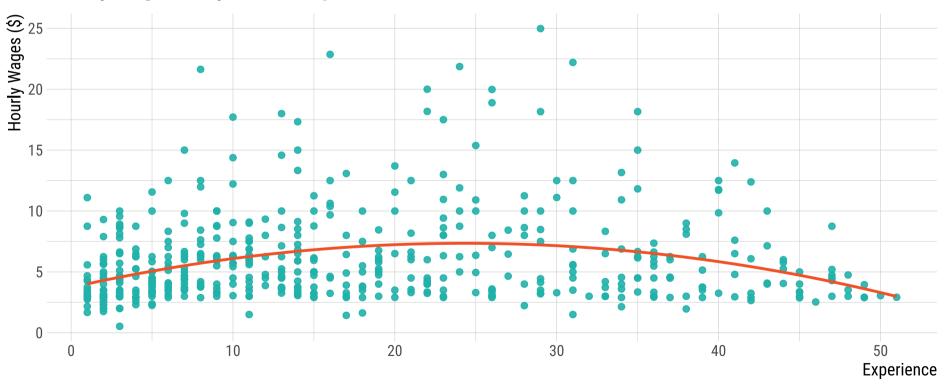
$$wage_i = eta_0 + eta_1 exper_i + eta_2 exper_i^2 + u_i$$

### Hourly wages vs. years of experience



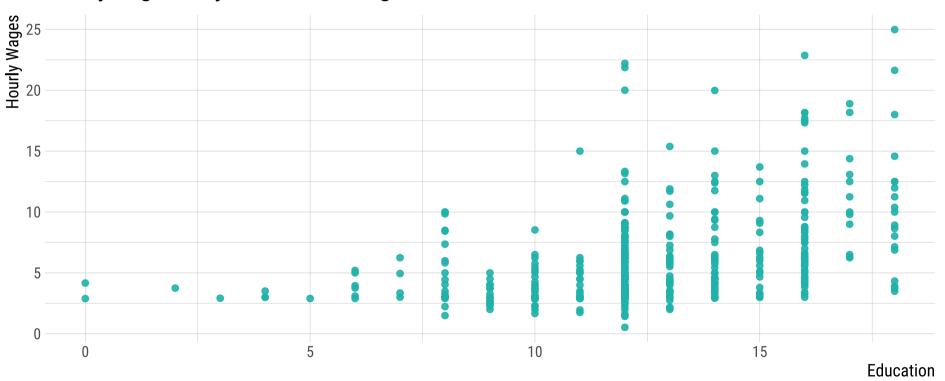
$$wage_i = eta_0 + eta_1 exper_i + eta_2 exper_i^2 + u_i$$

### Hourly wages vs. years of experience



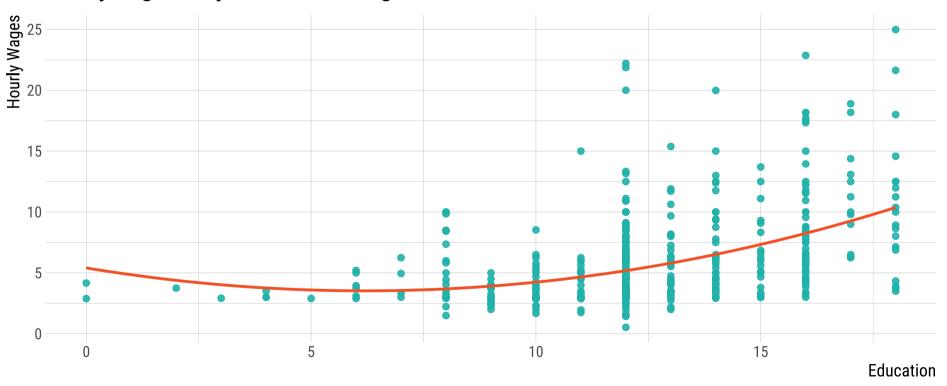
$$wage_i = eta_0 + eta_1 educ_i + eta_2 educ_i^2 + u_i$$

### Hourly wages vs. years of schooling



$$wage_i = eta_0 + eta_1 educ_i + eta_2 educ_i^2 + u_i$$

### Hourly wages vs. years of schooling



#### **Interpretation**

$$y_i = eta_0 + eta_1 x_{1i} + eta_2 x_{1i}^2 + u_i$$

$$rac{\partial \ y}{\partial \ x_1} = eta_1 + 2 \ \cdot \ eta_2 \ \cdot \ x_1$$

$$wage_i = eta_0 + eta_1 educ_i + eta_2 educ_i^2 + u_i$$

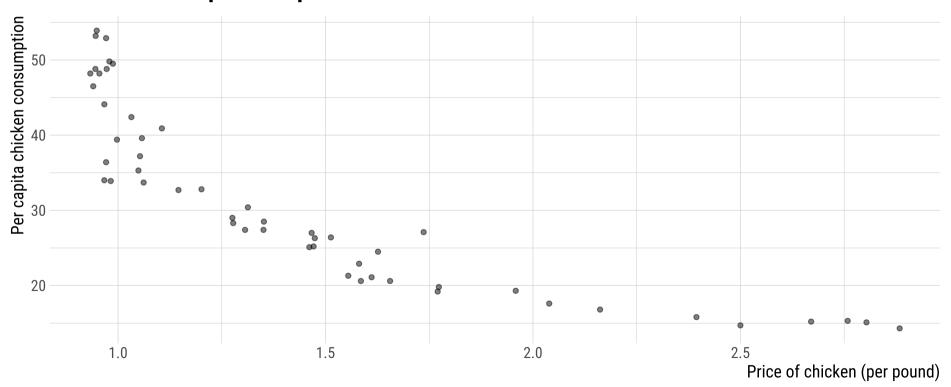
$$rac{\partial \ wage}{\partial \ educ} = eta_1 + 2 \ \cdot \ eta_2 \ \cdot \ educ$$

Inverse forms are used whenever the effect of an independent variable on  $y_i$  is expected to approach **zero** as its value approaches **infinity**.

As always, but especially important to this category, **economic theory** should *strongly recommend* the use of such functional form.

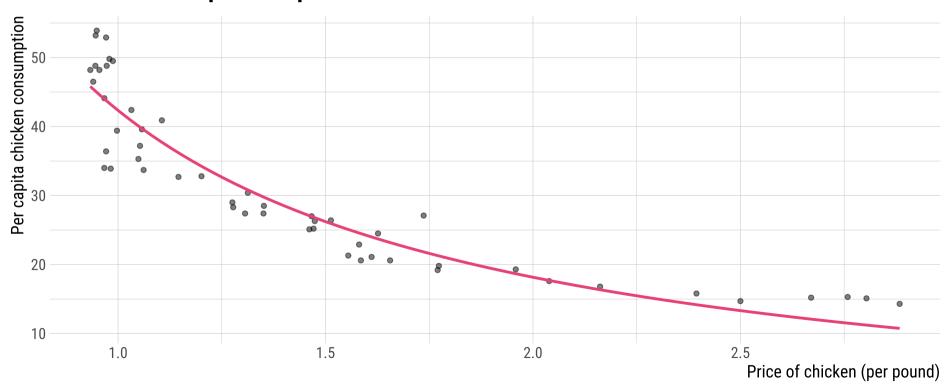
$$qchicken_i = eta_0 + eta_1 rac{1}{pchicken_i} + u_i$$

### Chicken consumption vs. price of chicken



$$qchicken_i = eta_0 + eta_1 rac{1}{pchicken_i} + u_i$$

### Chicken consumption vs. price of chicken



#### **Interpretation**

$$y_i = eta_0 + eta_1 rac{1}{x_{1i}} + u_i \ rac{\partial \ y}{\partial \ x_1} = rac{-eta_1}{x_1^2}$$

$$egin{aligned} qchicken_i &= eta_0 + eta_1 rac{1}{pchicken_i} + u_i \ & rac{\partial \ qchicken}{\partial \ pchicken} = rac{-eta_1}{pchicken^2} \end{aligned}$$

Whenever the effect of one variable on y depends on the **level of another variable**, the best **modeling** strategy is to use *interaction terms*.

For example, do we believe that an individual's wage depends on their education?

• If so, is this effect the **same** or **different** for two individuals with, e.g., a *college* degree, but with different years of experience on the job market?

Then, we represent a model by

$$wage_i = \beta_0 + \beta_1 educ_i + \beta_2 exper_i + \beta_3 educ_i \cdot exper_i + u_i$$

In more general terms, regression estimates  $(\hat{\beta}_i)$  describe **average effects**.

Some of these average effects may "hide" **heterogeneous effects** that differ by **group** or by the **level of another variable**.

Interaction terms help us in modeling such heterogeneous effects.

• For instance, it is plausible to consider that returns on education will differ by gender, race, region, etc.

#### **Interpretation**

$$egin{align} y_i &= eta_0 + eta_1 x_{1i} + eta_2 x_{2i} + eta_3 x_{1i} x_{2i} + u_i \ & rac{\partial \ y}{\partial \ x_1} = eta_1 + eta_3 \ \cdot \ x_2 \ \end{aligned}$$

Categorical variables are used to translate qualitative information into numbers.

• For instance, race, gender, being employed or not, enrolled in EC 339 or not, etc.

The easiest way to work with qualitative information is by using binary (dummy) variables.

For example,

$$y_i = \beta_0 + \beta_1 D_i + u_i$$

where  $D_i = 1$  if the criterion is fulfilled, and  $D_i = 0$  otherwise.

When **interpreting** regression coefficients associated with *dummy* variables, the *intercept*'s interpretation changes slightly.

Moreover, the slope coefficient on  $D_i$  is not interpreted in the same way we are used to.

Consider:

$$interviews_i = eta_0 + eta_1 graduate_i + u_i$$

#### where

- $interviews_i$  is the number of interviews a candidate is called for in a given period;
- $graduate_i$  equals 1 if she has graduated from college, and 0 otherwise.

$$interviews_i = eta_0 + eta_1 graduate_i + u_i$$

#### For this model,

- $\beta_0$  is the expected number of interviews when  $graduate_i = 0$  (non-graduates);
- $\beta_1$  is the expected **difference** in interview calls between graduates and non-graduates;
- And  $\beta_0 + \beta_1$  is the expected number of interviews for graduates (when  $graduate_i = 1$ ).

• In this case, non-graduates are the **reference group**.

$$interviews_i = eta_0 + eta_1 graduate_i + u_i$$

The model above is an example of an **intercept** dummy variable.

• We only have different **intercepts** when comparing two groups, but **slopes** are the same.

In order to allow for different **slopes**, we appeal to interaction terms involving categorical variables

• i.e., **slope** dummy variables.

## Log-Level Model

Important! If you have a **log-linear** model with a *binary* variable, the interpretation of the coefficient on that variable **changes**.

$$\log(y_i) = eta_0 + eta_1 D_i + u_i$$

with *D* being a *dummy* variable.

Interpretation of  $\beta_1$ :

- ullet When D=1, y will increase by  $100 imes \left(e^{eta_1-1}
  ight)$  percent.
- ullet When D=0, y will decrease by  $100 imes \left(e^{-eta_1-1}
  ight)$  percent.

## Log-Level Example

Binary explanatory variable: inlf

- inlf = 1 if the  $i^{th}$  individual is in the labor force.
- inlf = 0 if the  $i^{th}$  individual is not in the labor force.

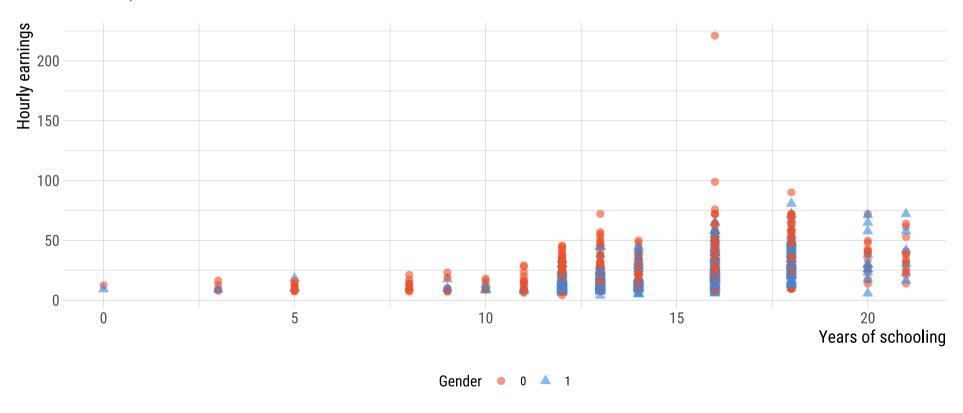
$$\widehat{log(slee}p_i) = 8.08 - 0.00365 \ inlf_i$$

- How do we interpret the coefficient on inlf?
  - Labor force participants sleep 36.65% less than non-participants.
  - Individuals that are not in the labor force sleep 36.92%% more than participants.

# Slope dummy variables

### Hourly wages vs. years of education (by gender)

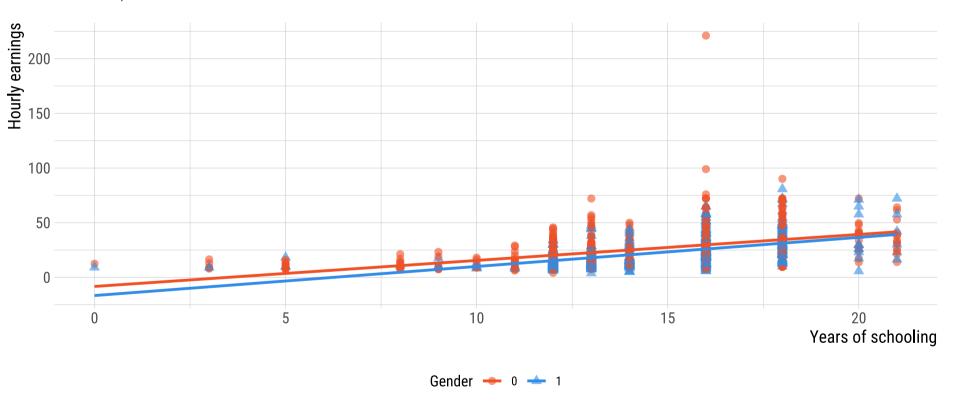
Female=1, Non-female=0



# Slope dummy variables

### Hourly wages vs. years of education (by gender)

Female=1, Non-female=0



## Slope dummy variables

#### **Interpretation**

$$egin{aligned} y_i &= eta_0 + eta_1 x_{1i} + eta_2 D_i + eta_3 D_i x_{1i} + u_i \ & rac{\partial \ y}{\partial \ x_1} = eta_1 + eta_3 \ \cdot \ D \ & rac{\partial \ y}{\partial \ D} = eta_2 + eta_3 \ \cdot \ x_1 \end{aligned}$$

$$egin{aligned} wage_i &= eta_0 + eta_1 educ_i + eta_2 female_i + eta_3 educ_i \cdot female_i + u_i \ & rac{\partial \ wage}{\partial \ educ} = eta_1 + eta_3 \ \cdot \ female \ & rac{\partial \ wage}{\partial \ female} = eta_2 + eta_3 \ \cdot \ educ \end{aligned}$$

Next time: Functional forms in practice