

Estimating the seal pup abundance in the Greenland Sea with Bayesian hierarchical modeling

Martin Jullum (NR)

Joint with Thordis Thorarinsdottir (NR)
and Fabian Bachl (Uni. of Edinburgh)

Oslo, 15.06.17



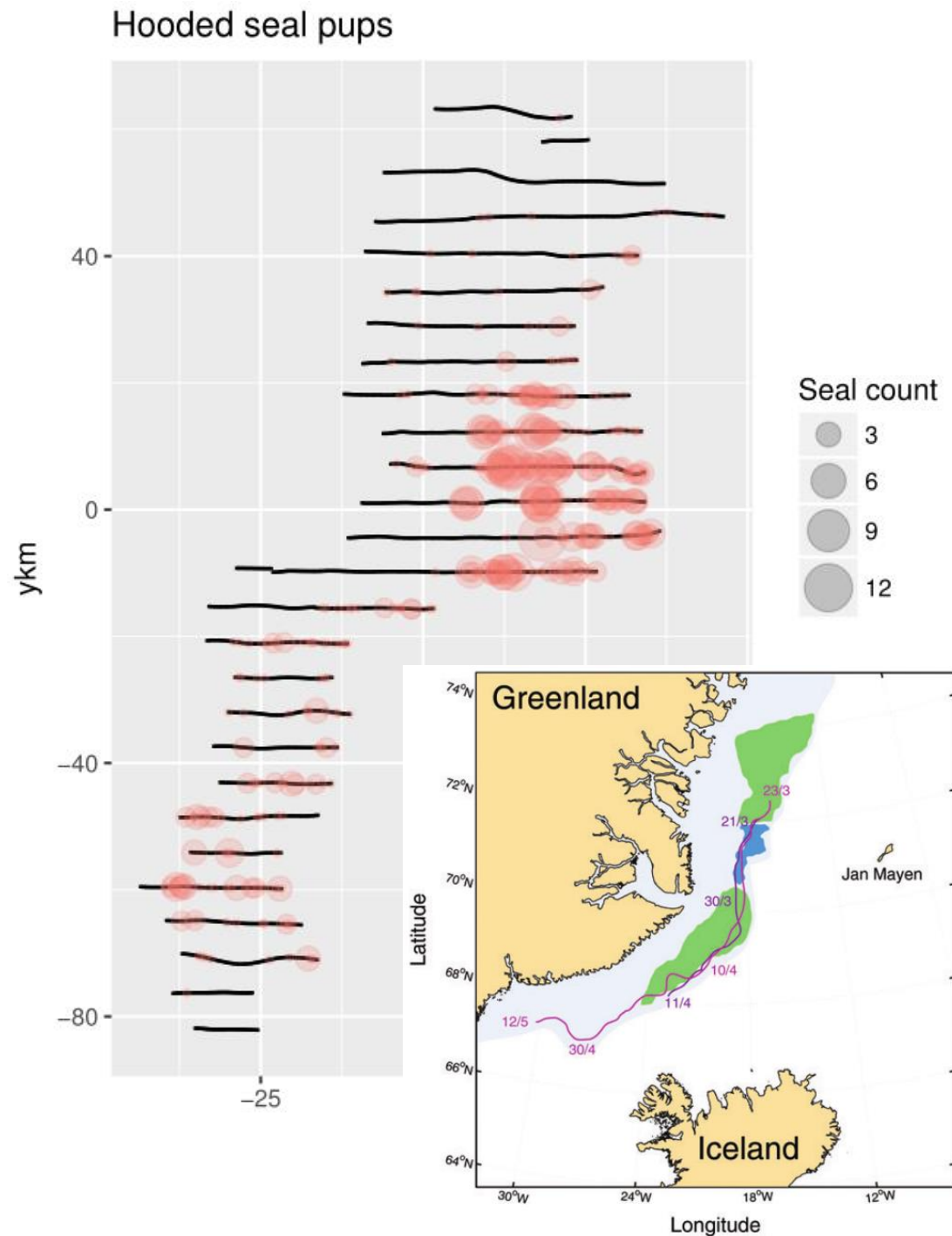
Problem

- ▶ Ultimate goal: Monitor seal abundance in the North Atlantic
- ▶ Well established dynamic abundance model for seals
 - Key component is **estimate + uncertainty of the number of seal pups**
 - Existing methods
 - Very basic ad-hoc scaling method
 - Spatial GAM (splines) model
- ▶ **Our task: Propose method to estimate the total number of pups with uncertainty + validate it!**



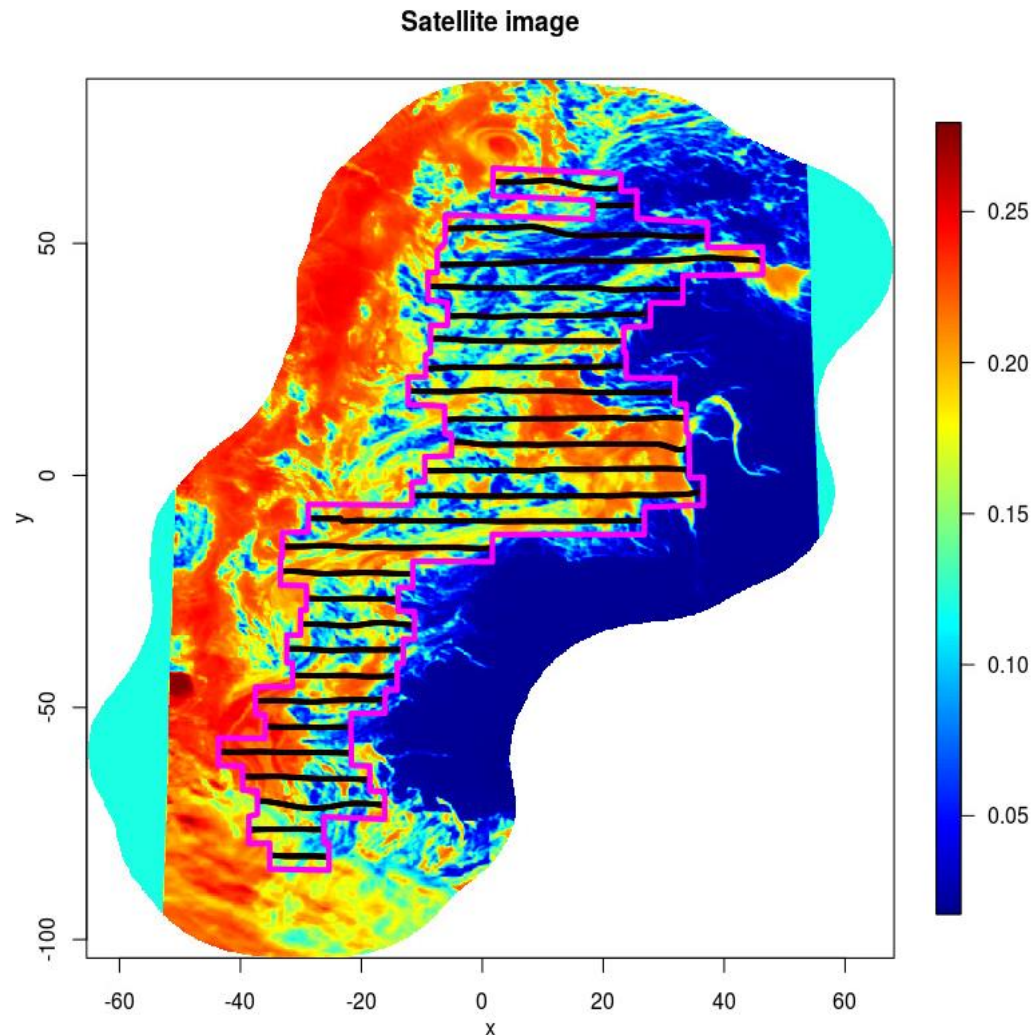
Data

- ▶ From an aerial photo survey conducted east of Greenland in 2012
- ▶ Number of pups in 2792 photos (A) in 27 transects sparsely covering the seal domain



Data

- ▶ From an aerial photo survey conducted east of Greenland in 2012
- ▶ Number of pups in 2792 photos (A) in 27 transects sparsely covering the seal domain
- ▶ Additional info: Quantified satellite image to indicate ice thickness
- ▶ Seal domain Ω shown in pink



A model for the seal pup appearance

- Model the spatial distribution of the seal pups with a Log-Gaussian Cox Process (LGCP)
 - Gaussian latent field Z
 - Point pattern $Y|Z \sim \text{PoissonProcess}(\lambda(s) = \exp(Z(s)))$
 - LGCP property: Given Z , counts $N(B)$ in disjoint Borel sets B indep. and distributed as $\text{Poisson}(\lambda = \int_B \exp(Z(s)) ds)$
 - LGCP Log-likelihood

$$|A| - \int_A \exp(Z(s)) ds + \sum_{i=1}^n Z(s_i),$$

Discretizing the LGCP-model

- ▶ Data are aggregated counts per photo
- ▶ Solution: Discretize the LGCP-model to the set where our data lives
- ▶ Let N_1, \dots, N_n be the counts in the $n = 2792$ photos, covering the space A_1, \dots, A_n
- ▶ Discretize the LGCP-model to

$$p(N_1, \dots, N_n | Z) = \prod_{i=1}^n \text{Poisson}(k = N_i, \lambda = \int_{A_i} Z(s) \, ds),$$

with $\text{Poisson}(k, \lambda) = \lambda^k \exp(-\lambda) / k!$

INLA and SPDE-INLA

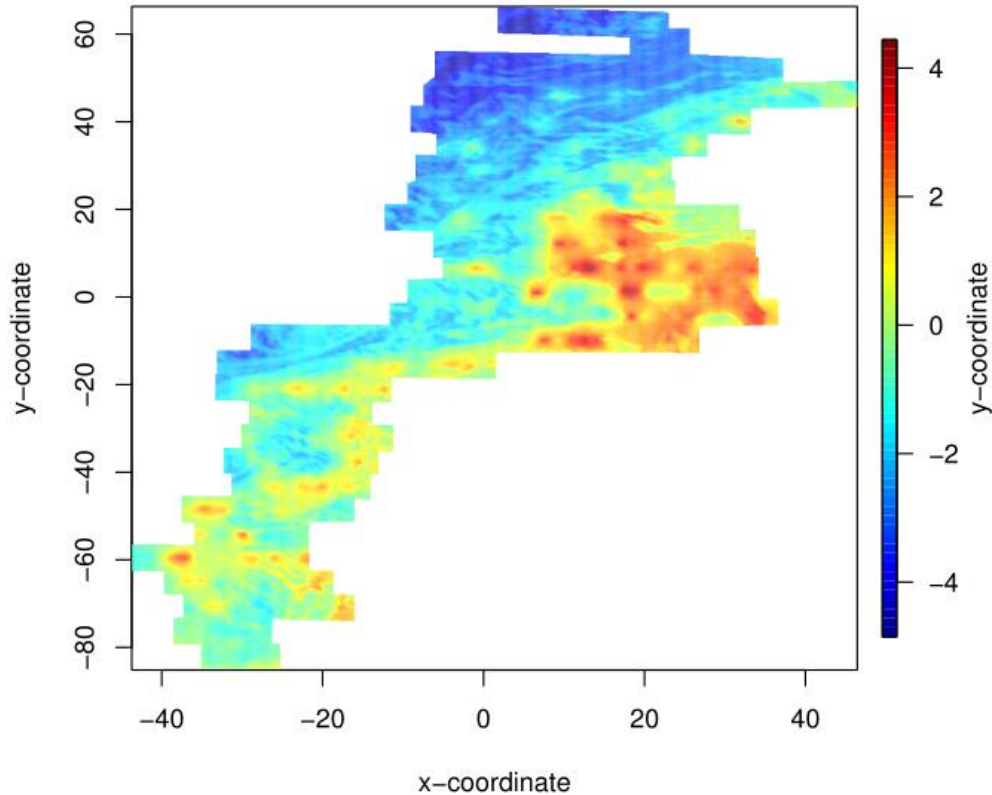
- ▶ Integrated Nested Laplace Approximation (INLA)
 - Computational feasible approximate Bayesian inference for Gaussian latent models with discrete latent field (GMRF)
 - Based on extensive Laplace approx. and numerical optimization.
- ▶ Spatial Partial Differential Equation (SPDE) approach
 - Makes INLA applicable to Gaussian latent models with continuous fields
 - Triangulates continuous latent field which translates to certain GMRF by formulation through solution to a SPDE

Modeling approach

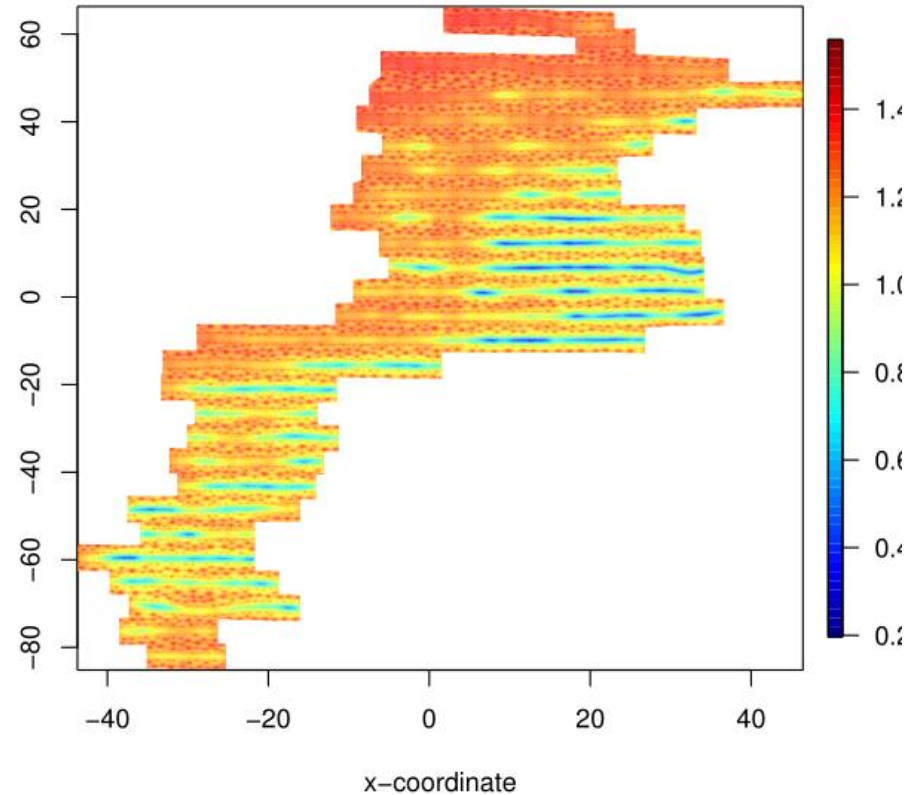
- ▶ Latent field $Z(s) = \alpha + \beta^t x_s + g(s)$, x_s satellite information, $g(s)$ zero-mean Gaussian field with a Matern covariance structure
- ▶ Bayesian approach with vague priors on all parameters
- ▶ The Bayesian solution to our problem is the «**posterior predictive distribution**» of seal pup counts in the seal domain $p(N(\Omega)|Y)$
 - Easy to compute with samples from $p(Z|Y)$
- ▶ Use SPDE-INLA to fit the model and perform the posterior sampling

Results our approach

Mean of latent field



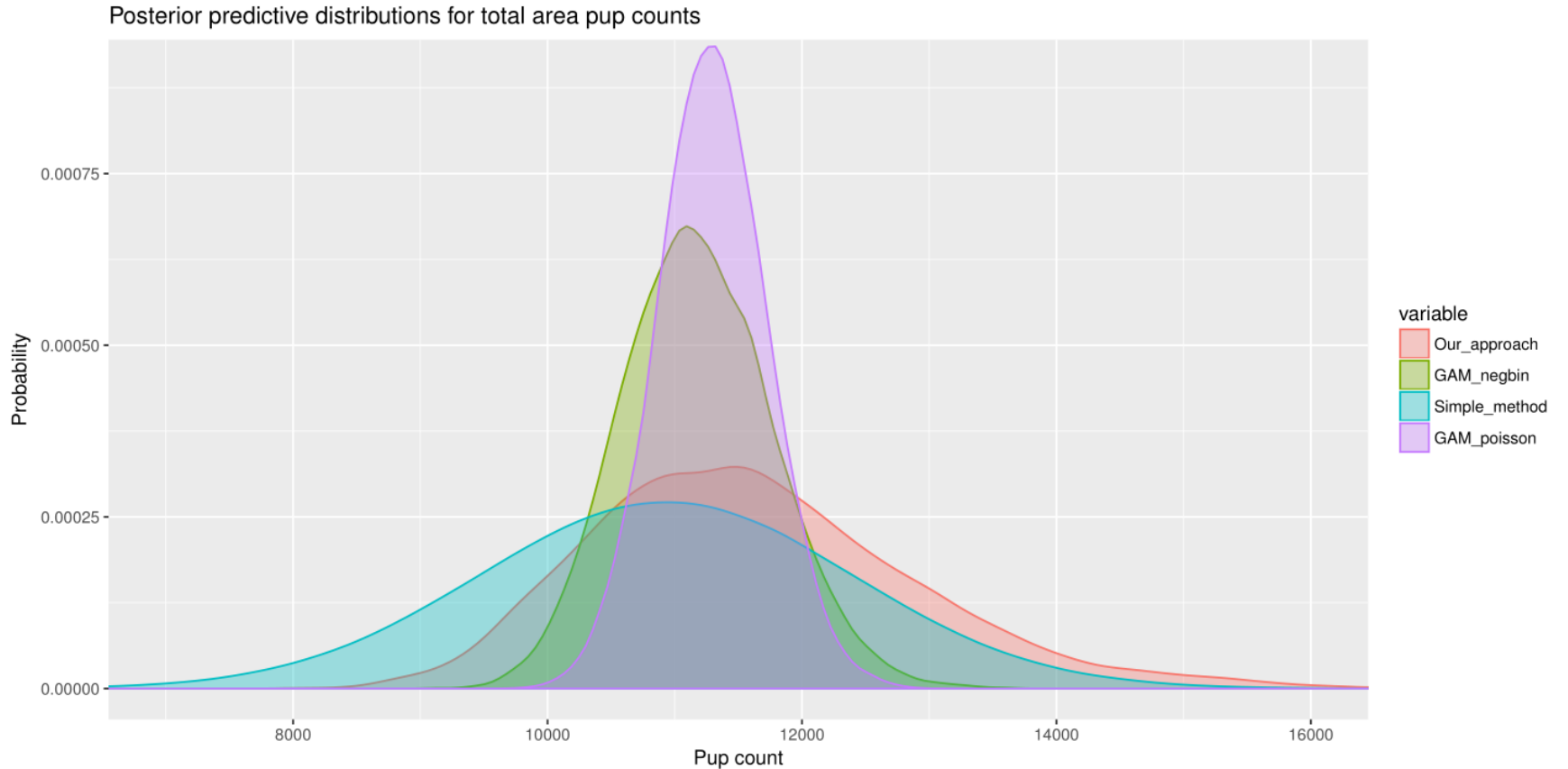
Sd of latent field



Method comparison



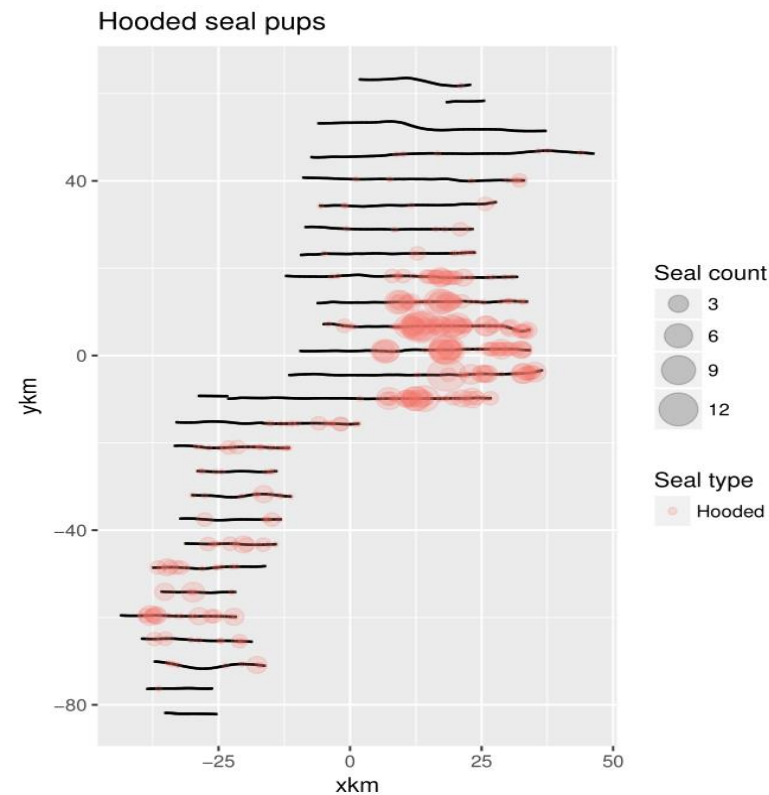
Method comparison



Result validation

- ▶ 2 different CV schemes
 - Leave out random photos
 - Leave out all photos on transect
 - Evaluate posterior predictive distribution both per photo and per transect
- ▶ Evaluation criterion

$$CRPS(F, y) = \int_{-\infty}^{\infty} (F(x) - \mathbf{1}(x - y))^2 dx$$
$$\text{logscore}(f, y) = \log(f(y))$$

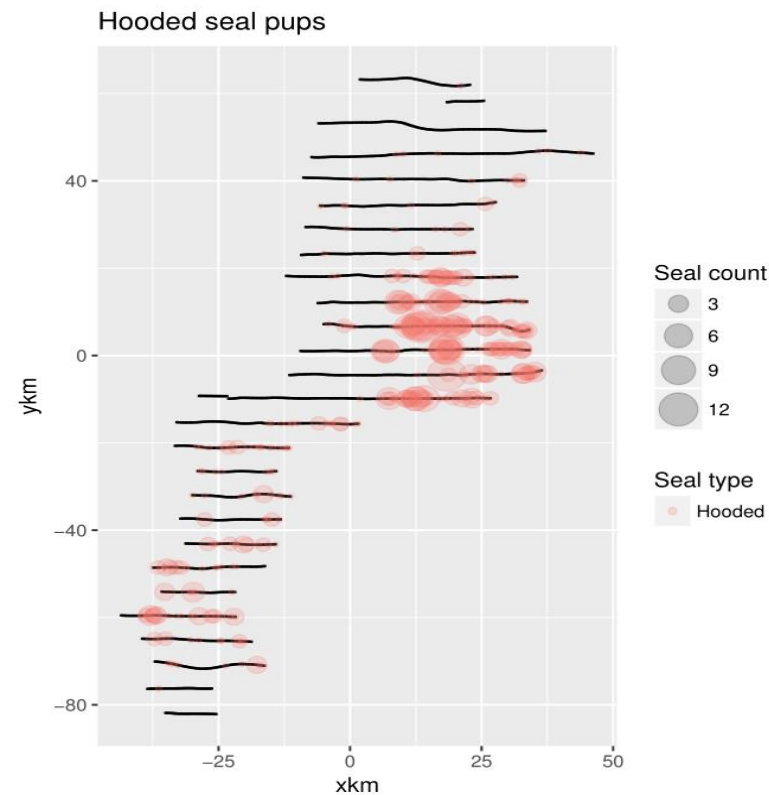


Result validation

- ▶ 2 different CV schemes
 - Leave out random photos
 - Leave out all photos on transect
 - Evaluate posterior predictive distribution both per photo and per transect
- ▶ Evaluation criterion

$$CRPS(F, y) = \int_{-\infty}^{\infty} (F(x) - \mathbf{1}(x - y))^2 dx$$
$$\text{logscore}(f, y) = \log(f(y))$$

- ▶ Long story short
 - ▶ Our method is significantly better on photo level
 - ▶ We do as well as the GAM approaches (better than the simple method) on transect level



Result validation

PHOTO LEVEL

| | Random 10-fold CV CRPS | Leave-out full transect CRPS |
|----------------------|---------------------------------|---------------------------------|
| Our approach | <i>0.18 (0.16, 0.19)</i> | <i>0.22 (0.20, 0.25)</i> |
| GAM_negbin | 0.21 (0.19, 0.23) | <i>0.22 (0.20, 0.24)</i> |
| GAM_poisson | 0.22 (0.20, 0.24) | 0.24 (0.22, 0.26) |
| Simple_method | 0.26 (0.24, 0.28) | 0.26 (0.24, 0.29) |

AGGREGATE/TRANSECT LEVEL

| | Random 10-fold CV CRPS | Leave-out full transect CRPS |
|----------------------|---------------------------|---------------------------------|
| Our approach | 5.43 (4.04, 6.99) | 9.91 (5.99, 14.80) |
| GAM_negbin | 5.93 (4.95, 7.00) | <i>9.37 (5.66, 13.63)</i> |
| GAM_poisson | 5.90 (4.49, 7.42) | 10.14 (5.86, 15.09) |
| Simple_method | <i>4.83 (3.27, 6.66)</i> | 15.57 (11.77, 19.68) |

Alternative model

- ▶ Arnt-Børre + Tor Arne + others (2009)
 - Let $Z_0(s) = f_{GAM}(s)$, with $f_{GAM}(s)$ a (spatial) smooth spline.
 - $\mu_i = |A_i| \exp(Z_0(s_i^*))$ for s_i^* the mid-point in cell A_i
 - Fit the counts per photo as a negative binomial regression with constant shape κ and $|A_i|$ as offset
 - Frequentist approach
 - Smoothness of $f_{GAM}(s)$ chosen through generalized CV
- We test this formulation, also with satellite data and Poisson distributions
- Use sampling to produce predictive distribution of total pup counts for comparison