

An approximate Bayesian inversion framework based on local Gaussian likelihoods

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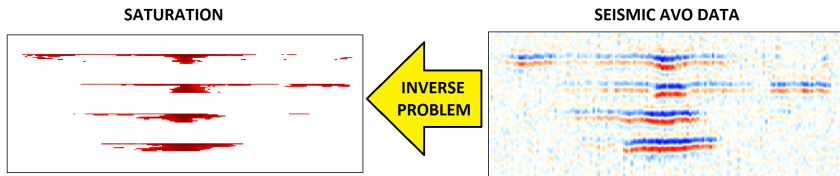
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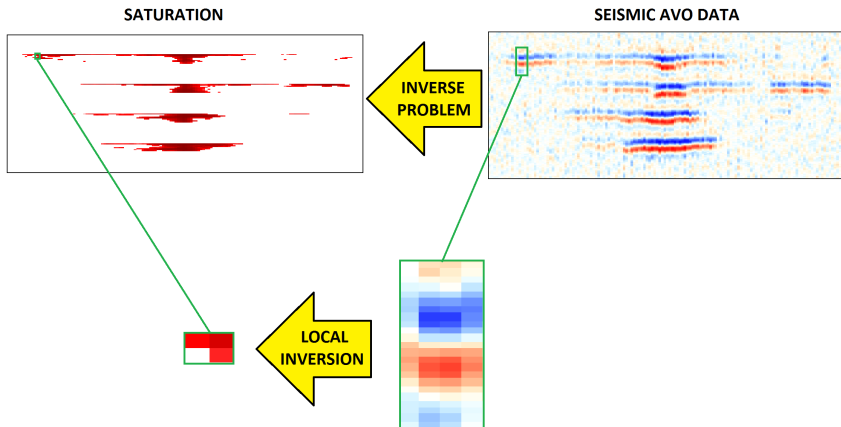
September 10, 2015

The EAGE logo consists of the word "EAGE" in white, serif, uppercase letters, centered within a solid purple square.The logo for the 2015 Petroleum Geostatistics conference. It features the word "PETROLEUM" in a bold, sans-serif font, with a stylized vertical bar to its left. To the right of "PETROLEUM" is a purple square containing the year "2015". Below "PETROLEUM" is the word "GEOSTATISTICS" in a similar bold, sans-serif font, with a horizontal line separating the two words.

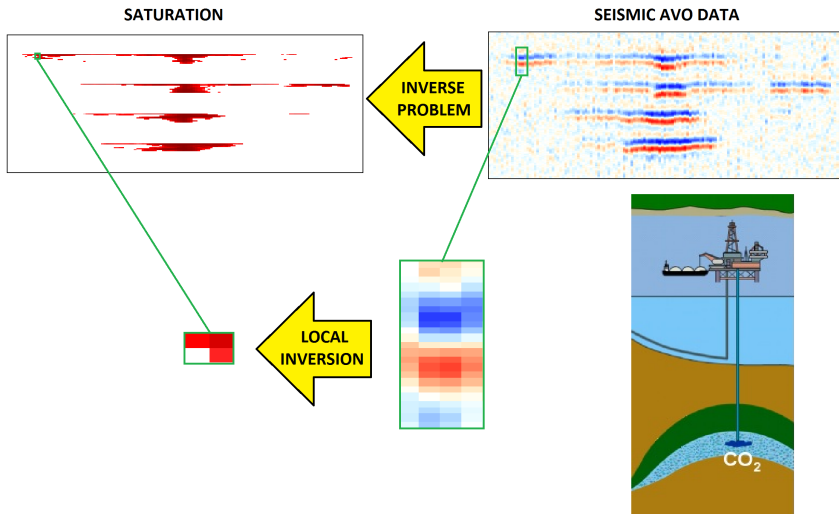
Outline



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Problem setup

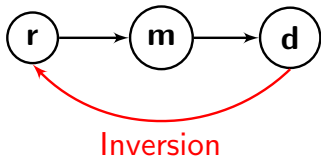
Forward model



- **r**: Rock properties (e.g. saturation, porosity, permeability, lithology, etc.)
- **m**: Geophysical properties (= elastic parameters \mathbf{V}_p , \mathbf{V}_s and ρ)
- **d**: Geophysical data (= Seismic AVO data)

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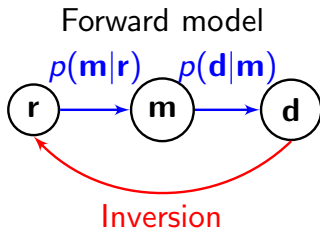
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Inverse problem: Which values of **r** generated the observed **d**?

Problem setup

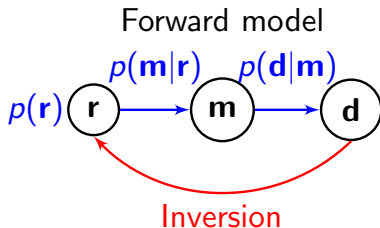


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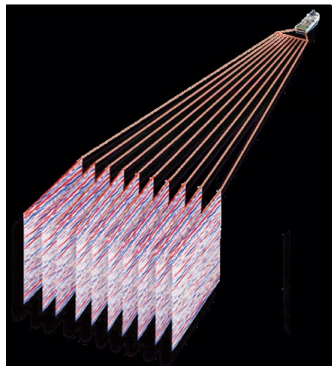
Inverse problem: Which values of **r** generated the observed **d**?

- Statistical model setup
- The Bayesian approach is natural
 - Inversion \Leftrightarrow consult posterior distribution $p(\mathbf{r}|\mathbf{d})$

Computing posterior

Posterior distribution: $p(\mathbf{r}|\mathbf{d}) \propto \int p(\mathbf{d}|\mathbf{m})p(\mathbf{m}|\mathbf{r})p(\mathbf{r})d\mathbf{m}$

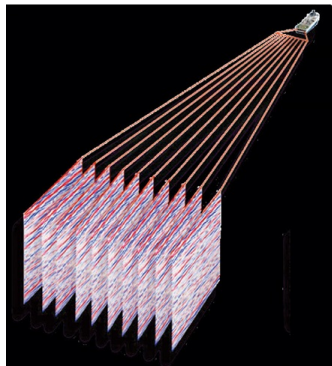
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- High dimensional problem
- Enormous amount of highly correlated data \mathbf{d}
- Complex dependency structures
- Analytical expression for posterior seldom available
- Full Gaussian linear inversion (Buland and Omre (2003)) and similar approaches may not be valid
- MCMC can be very time consuming



Key elements of our approach

- Divide the global inversion problem into several local inversions
 - Approximate marginal posterior $p(\mathbf{r}_i|\mathbf{d})$ for each cell i in the gridded region, rather than the global $p(\mathbf{r}|\mathbf{d})$ for the full region

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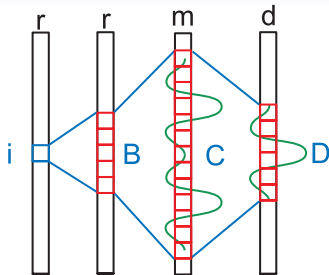
Assumptions:

- $p(\mathbf{d}|\mathbf{m}) \sim N(\mathbf{G}\mathbf{m}, \Sigma)$, i.e. $\mathbf{d} = \mathbf{G}\mathbf{m} + \varepsilon$, $\varepsilon \sim N(0, \Sigma)$
- We may sample from $p(\mathbf{m}|\mathbf{r})$ and $p(\mathbf{r})$

Local variables

Inversion for cell i : $p(\mathbf{r}_i|\mathbf{d})$

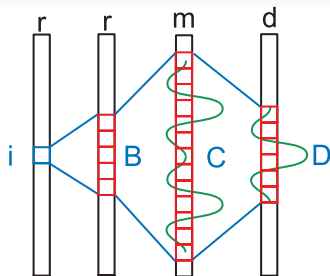
- 3 neighborhoods of cells:
 $B = B(i)$, $C = C(i)$,
 $D = D(i)$
- These have corresponding local variables: \mathbf{r}_B , \mathbf{m}_C , \mathbf{d}_D



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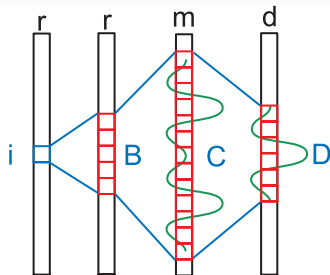
- Rely on the following relations

$$p(\mathbf{r}_i|\mathbf{d}) \approx p(\mathbf{r}_i|\mathbf{d}_D) = \int p(\mathbf{r}_B|\mathbf{d}_D) d\mathbf{r}_{B-i} \propto \int p(\mathbf{d}_D|\mathbf{r}_B)p(\mathbf{r}_B) d\mathbf{r}_{B-i}.$$

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- We shall approximate $p(\mathbf{d}_D|\mathbf{r}_B)$

Local Gaussian likelihood approximation

Need to approximate the local likelihood

$$p(\mathbf{d}_D | \mathbf{r}_B) = \int p(\mathbf{d}_D | \mathbf{m}_C) p(\mathbf{m}_C | \mathbf{r}_B) d\mathbf{m}_C$$

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- $p(\mathbf{d} | \mathbf{m}) \sim N(G\mathbf{m}, \Sigma) \Rightarrow p(\mathbf{d}_D | \mathbf{m}_C) \approx p^*(\mathbf{d}_D | \mathbf{m}_C) \sim N(G_0\mathbf{m}_C, \Sigma_0)$

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- $p(\mathbf{m}_C|\mathbf{r}_B) \approx p^*(\mathbf{m}_C|\mathbf{r}_B) \sim N(\boldsymbol{\mu}, \boldsymbol{\Gamma})$, where

$$\boldsymbol{\mu} = \boldsymbol{\mu}(\mathbf{r}_B) = E^*[\mathbf{m}_C|\mathbf{r}_B] \text{ and } \boldsymbol{\Gamma} = \boldsymbol{\Gamma}(\mathbf{r}_B) = \text{Cov}^*[\mathbf{m}_C|\mathbf{r}_B]$$

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- Sample large number of 'pairs' $(\mathbf{m}_C, \mathbf{r}_B)$ from joint distribution
- Fit $\boldsymbol{\mu}$ -function by nonlinear regression
- Use residuals to fit different $\boldsymbol{\Gamma}$ -functions for different \mathbf{r}_B

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- Fit $\boldsymbol{\mu}$ -function by nonlinear regression
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- Final approximation:

$$\begin{aligned} p(\mathbf{d}_D|\mathbf{r}_B) &\approx \int p^*(\mathbf{d}_D|\mathbf{m}_C)p^*(\mathbf{m}_C|\mathbf{r}_B) d\mathbf{m}_C \\ &= p^*(\mathbf{d}_D|\mathbf{r}_B) \sim N(\mathbf{G}_0\boldsymbol{\mu}, \Sigma_0 + \mathbf{G}_0\boldsymbol{\Gamma}\mathbf{G}_0^t) \end{aligned}$$

Approximate inversion quantities

Approximate local posterior:

$$p(\mathbf{r}_i|\mathbf{d}_D) \approx p^*(\mathbf{r}_i|\mathbf{d}_D) \propto \int p^*(\mathbf{d}_D|\mathbf{r}_B)p(\mathbf{r}_B) d\mathbf{r}_{B-i}$$

- Inversion: Compute inversion quantities like the posterior mean, median, sd or credibility intervals

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Weighted Monte Carlo:

- Sample from $p(\mathbf{r}_B)$
- Weight the sampled \mathbf{r}_i by $p^*(\mathbf{d}_D|\mathbf{r}_B)$ and normalize the weights
- Compute appropriate weighted averages over r_i to approximate inversion quantities

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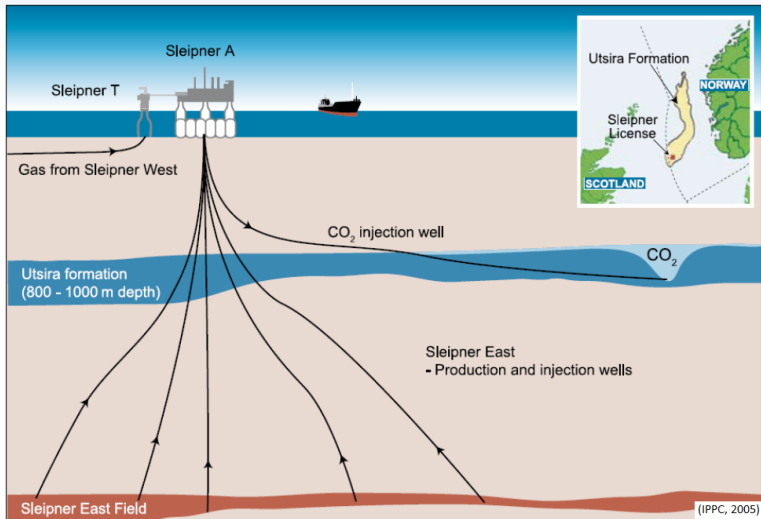
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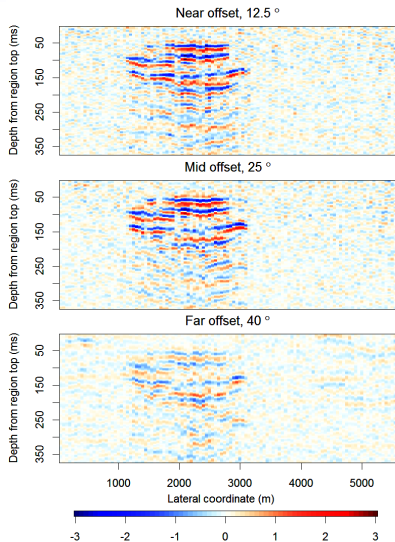
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- Under stationarity: Re-use the general form of $p^*(\mathbf{d}_D | \mathbf{r}_B)$ and the samples from $p(\mathbf{r}_B)$ and repeat the Weighted Monte Carlo procedure for all i in the region

Illustration: CO₂ monitoring \Rightarrow map saturation



Real case

Seismic amplitude differences

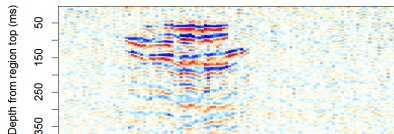


d = Seismic AVO difference data
m = change in log of ($\mathbf{V}_p, \mathbf{V}_s, \rho$)
r = saturation change

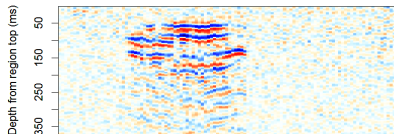
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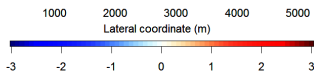
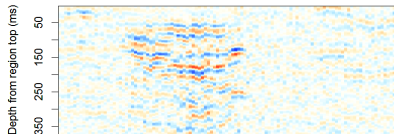
Near offset, 12.5 °



Mid offset, 25 °

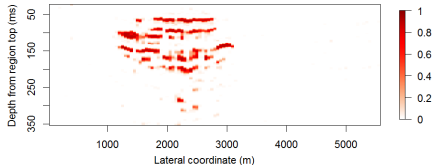


Far offset, 40 °



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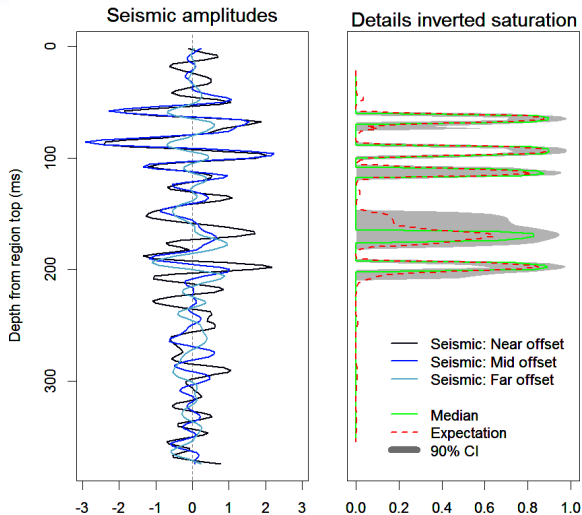
Approximate posterior mean saturation



18 000 cells inverted < 30 min
 on 4-cored Windows laptop

Real case

Vertical trace



Method comparison based on inversion of synthetic data

Regular Gaussian inversion: $p^*(\mathbf{m}|\mathbf{d}) \propto p(\mathbf{d}|\mathbf{m})p^*(\mathbf{m})$, Buland and Omre (2003)

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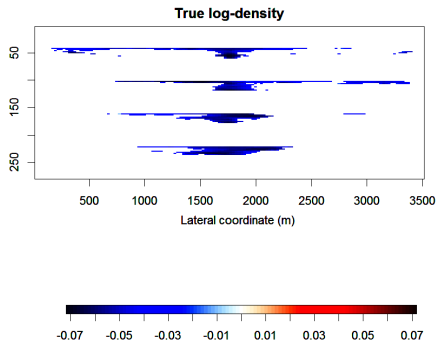
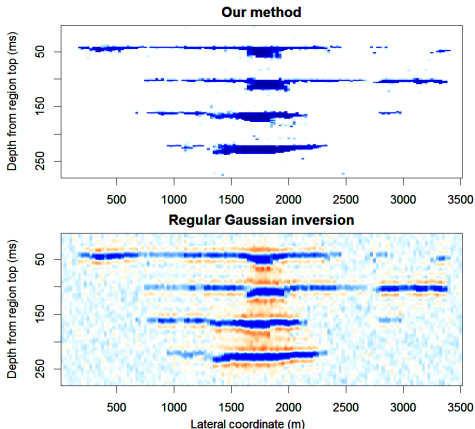
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Conclusion

Our method summarized:

- Split the global inversion problem into lots of cell-wise inversions
 - Use only information in variables spatially close to the cell to invert
 - Clever local-Gaussian approximations and weighted Monte Carlo
-
- Conceptually easy to extend the method
 - Well suited for parallelization