Module 10: Causal Mechanisms

Fall 2021

Matthew Blackwell

Gov 2003 (Harvard)

1/ Causal Mechanisms

• Theory \implies (or \equiv) causal effects

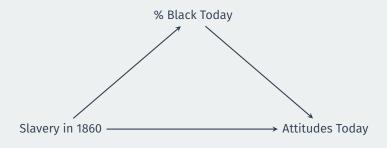
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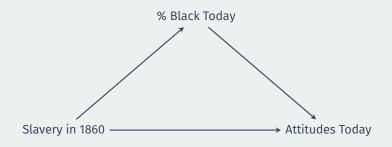
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- · How to adjudicate between theories that predict the same ATE?

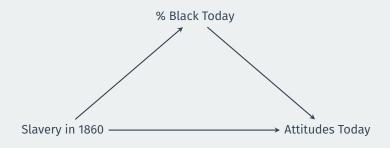
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- How to adjudicate between theories that predict the same ATE?
- Put differently: what mechanism drives a particular causal effect?



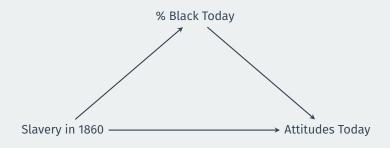
• Effect of antebellum slavery on modern white attitudes:



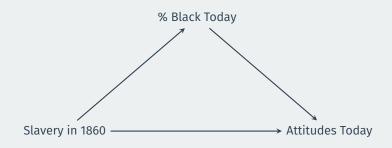
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- Two possible mechanisms with very different implications:
 - · Historical persistence of attitudes via intergenerational transfer.
 - Or is this effect due to demographic persistence? (More African Americans in former enslaved areas today → whites threatened today)

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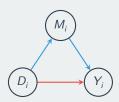
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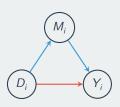
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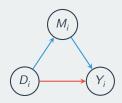
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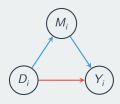
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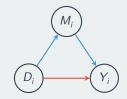
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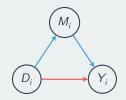


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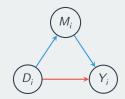
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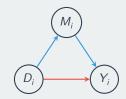
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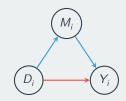
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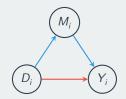
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- Consistency: $M_i = M_i(D_i)$ and $Y_i(D_i, M_i(D_i))$.

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 - · Not just the fundamental problem of CI.

2/ Estimands

• Definition for each $m \in \mathcal{M}$:

Individual: $\xi_i(m) = Y_i(1, m) - Y_i(0, m)$

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 - \rightsquigarrow can be used to establish existence of unmediated path from $D \rightarrow Y$.
- Can capture **interactions** if $\overline{\zeta}_i(m) \neq \overline{\zeta}_i(m')$

• Definition of the natural indirect effect (NIE):

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- If D_i doesn't affect M_i , so that $M_i(1) = M_i(0)$, then $\delta_i = 0$.

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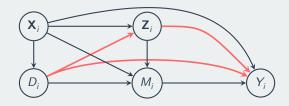
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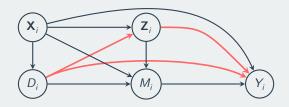
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- Total effect decomposition:

$$\tau_i = Y_i(1, M_i(1)) - Y_i(0, M_i(0)) = \underbrace{\delta_i(d)}_{\text{NIE}} + \underbrace{\zeta_i(1-d)}_{\text{NDE}}$$

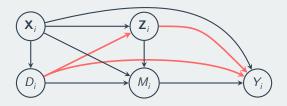
3/ Identification



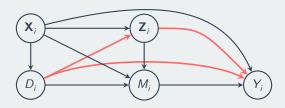
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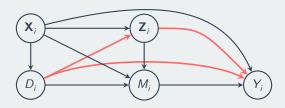


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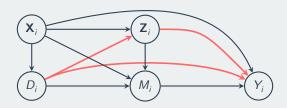
$$\begin{aligned} \{Y_i(d',m),M_i(d)\} \perp \!\!\! \perp D_i \mid \mathbf{X}_i = \mathbf{x} \\ Y_i(d,m) \perp \!\!\! \perp M_i \mid \mathbf{X}_i = \mathbf{x},D_i = d, \mathbf{Z}_i = \mathbf{z} \end{aligned}$$



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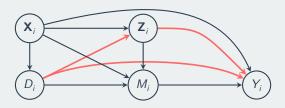
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- Interpretation: two "selection-on-observables" assumptions.
 - D_i randomly assigned conditional on **X**_i.
 - M_i randomly assigned conditional on \mathbf{X}_i , D_i , and \mathbf{Z}_i .

Post-treatment bias if we just condition on Z_i:

$$\begin{split} \overline{\xi}(m) \neq \sum_{\mathbf{x}, \mathbf{z}} \left\{ \mathbb{E}[Y_i \mid D_i = 1, M_i = m, \mathbf{X}_i = \mathbf{x}, \mathbf{Z}_i = \mathbf{z}] \right. \\ &- \mathbb{E}[Y_i \mid D_i = 0, M_i = m, \mathbf{X}_i = \mathbf{x}, \mathbf{Z}_i = \mathbf{z}] \right\} \mathbb{P}(\mathbf{X}_i = \mathbf{x}, \mathbf{Z}_i = \mathbf{z}) \end{split}$$

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• Ignores that **Z**_i depends on D_i!

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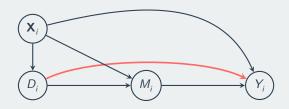
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• g-formula (Robins) generalizes to any number of treatments

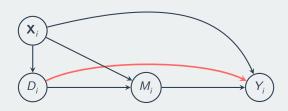
Identification for mediation



· Sequential ignorability (Imai et al):

$$\begin{aligned} &\{Y_i(d',m),M_i(d)\} \perp \!\!\! \perp D_i \mid \mathbf{X}_i = \mathbf{x} \\ &Y_i(d,m) \perp \!\!\! \perp M_i \mid \mathbf{X}_i = \mathbf{x},D_i = d \end{aligned}$$

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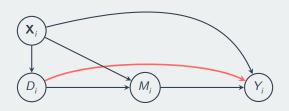


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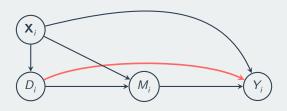
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 - Assumes away post-treatment bias conditioning on M_i

Identifying (in)direct effects

ANIE under binary treatment/mediator:

$$\begin{split} \bar{\delta}(d) &= \sum_{\mathbf{x}} \left(\{ \underbrace{\mathbb{P}[M_i = 1 \mid D_i = 1, \mathbf{X}_i = \mathbf{x}] - \mathbb{P}[M_i = 1 \mid D_i = 0, \mathbf{X}_i = \mathbf{x}]}_{\text{effect of } D_i \text{ on } M_i} \right. \\ &\times \{ \underbrace{\mathbb{E}[Y_i \mid M_i = 1, D_i = d, \mathbf{X}_i = \mathbf{x}] - \mathbb{E}[Y_i \mid M_i = 0, D_i = d, \mathbf{X}_i = \mathbf{x}]}_{\text{effect of } M_i \text{ on } Y_i} \\ &\times \mathbb{P}(\mathbf{X}_i = \mathbf{x}) \end{split}$$

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Multiply paths given X_i and aggregate intuitive given DAG:



(In)direct effects with non-binary mediators

• Let's say that the mediator has J categories:

$$\begin{split} \overline{\delta}(d) &= \sum_{\mathbf{x}} \Big(\sum_{m=0}^{J-1} \mathbb{E}[Y_i \mid M_i = m, D_i = d, \mathbf{X}_i = \mathbf{x}] \\ &\times \big\{ \mathbb{P}[M_i = m \mid D_i = 1, \mathbf{X}_i = \mathbf{x}] - \mathbb{P}[M_i = m \mid D_i = 0, \mathbf{X}_i = \mathbf{x}] \big\} \Big) \\ &\times \mathbb{P}(\mathbf{X}_i = \mathbf{x}) \end{split}$$

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The ANDE is the following:

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Effect of D_i for a fixed m averaged over the distribution of M_i when
D_i = d.

 Robins proposed a different identification strategy, based on a no-interactions assumption:

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- → ACDE = ANDE.
- Strong assumption because it has to hold at the individual level (like monotonicity for IV).

4/ Linear Structural Equation Models

$$M_i(d) = \alpha_0 + \alpha_1 d + \eta_i$$

$$Y_i(d, m) = \beta_0 + \beta_1 d + \beta_2 m + \varepsilon_i$$

• Let's say that we have a linear, structural model for all variables:

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Linear models and mediation

 It's clear that we can write the total effect of the treatment in the following way:

$$\begin{split} Y_i(1,M_i(1)) - Y_i(0,M_i(0)) = & \beta_0 + \beta_1 + \beta_2(\alpha_0 + \alpha_1 + \eta_i) + \varepsilon_i \\ & - \beta_0 - \beta_2(\alpha_0 + \eta_i) - \varepsilon_i \\ = & \beta_1 + \beta_2 \cdot \alpha_1 \end{split}$$

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· What about the indirect effect:

$$\begin{split} Y_i(0,M_i(1)) - Y_i(0,M_i(0)) = & \beta_0 + \beta_2(\alpha_0 + \alpha_1 + \eta_i) + \varepsilon_i \\ & - \beta_0 - \beta_2(\alpha_0 + \eta_i) - \varepsilon_i \\ = & \beta_2 \cdot \alpha_1 \end{split}$$

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- Indirect effect as the product: $\widehat{ANIE} = \hat{\alpha}_1 \hat{\beta}_2$.

Interactions

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 We could incorporate an interaction into the model here to allow for the indirect effect to vary.

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$$\mathbb{V}[\widehat{\textit{ANIE}}] \approx \hat{\alpha}_1^2 \mathbb{V}[\hat{\beta}_2] + \hat{\beta}_2^2 \mathbb{V}[\hat{\alpha}_1]$$

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 We can use this formula to estimate standard errors for the indirect effects.

5/ Nonparametric Estimation

Nonparametric estimation

Nonparametric estimation

- If the number of categories in M_i, D_i, and X_i are small, use plug-in estimator for the CEF of Y_i:

$$\widehat{\mathbb{E}}[Y_i \mid M_i = m, D_i = d, \mathbf{X}_i = \mathbf{x}] = \frac{\sum_i Y_i \mathbb{1}\{M_i = m, D_i = d, \mathbf{X}_i = \mathbf{x}\}}{\sum_i \mathbb{1}\{M_i = m, D_i = d, \mathbf{X}_i = \mathbf{x}\}}$$

Nonparametric estimation

- LSEMs require strong modeling assumptions → what about nonparametrics?
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• Same for M_i :

$$\widehat{\mathbb{P}}[M_i = m \mid D_i = d, \mathbf{X}_i = \mathbf{x}] = \frac{\sum_i \mathbb{1}\{M_i = m, D_i = d, \mathbf{X}_i = \mathbf{x}\}}{\sum_i \mathbb{1}\{D_i = d, \mathbf{X}_i = \mathbf{x}\}}$$

 If the number of categories is large, then we can use nonparametric regressions for the outcome and the mediator.

$$\mu_{dm}(\mathbf{x}) = \mathbb{E}[Y_i \mid M_i = m, D_i = d, \mathbf{X}_i = \mathbf{x}]$$

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- Need to be careful with the curse of dimensionality in X_i. Use good nonparametric strategies (cross-validation, etc)

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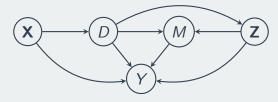
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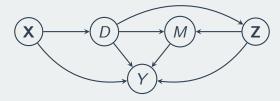
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- Obviously, this is a much harder problem. In this case, we actually can use Monte Carlo simulation to take the integral.
- Modeling M_i probably appropriate here.

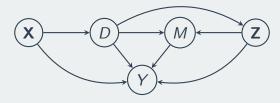
6/ Estimating Controlled Direct Effects



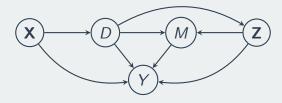
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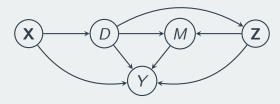


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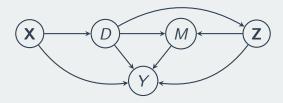
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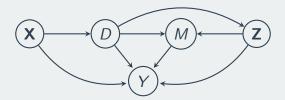
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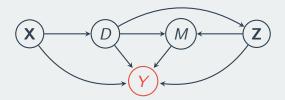
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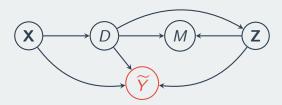
- γ_1 is not the CDE (posttreatment bias)
- γ_2 is the effect of M_i on Y_i (if model is correct)



• Create a blipped down (or demediated) outcome: $\widetilde{Y}_i = Y_i - \widehat{\gamma}_2 M_i$

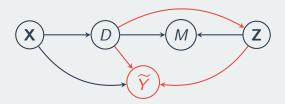


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· Relies on correct modeling of the outcome!

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- ATE ACDE ≠ an indirect effect, but still can tell us something about mechanisms.

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