Module 7b: Weighting Estimators

Fall 2021

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Gov 2003 (Harvard)

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- $K_M(i)$ is the number of times i is used as a match.
- Weighting estimators choose the weights directly to reduce imbalance.

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• Key idea: reweight sample to be representative of population.

• Applying HT potential outcomes: weight by inverse propensity score.

$$\widehat{\mathsf{ATE}} = \widehat{\tau}_{ipw} = \frac{1}{n} \sum_{i=1}^{n} \left(\frac{D_i Y_i}{\widehat{\pi}(\mathbf{X}_i)} - \frac{(1 - D_i) Y_i}{1 - \widehat{\pi}(\mathbf{X}_i)} \right)$$

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- Similar expression for ATT:

$$\widehat{\mathsf{ATT}} = \widehat{\tau}_{ipw,t} = \frac{1}{n_1} \sum_{i=1}^n \left(D_i Y_i - \frac{\widehat{\pi}(\mathbf{X}_i) (1 - D_i) Y_i}{1 - \widehat{\pi}(\mathbf{X}_i)} \right)$$

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 - · A kind of "continuous" version of matching with replacement.

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· Practically, weighted least squares gives automatic normalization:

$$(\hat{\alpha}_{\text{wls}}, \widehat{\tau}_{\text{wls}}) = \operatorname*{arg\,min}_{\alpha,\tau} \sum_{i=1}^{n} \left(\frac{D_{i}}{\widehat{\pi}(\mathbf{X}_{i})} + \frac{1-D_{i}}{1-\widehat{\pi}(\mathbf{X}_{i})} \right) \left(Y_{i} - \alpha - \tau \, D_{i} \right)^{2}$$

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 - Moment conditions for the propensity score model with parameters θ :

$$\mathbb{E}\left[\underbrace{\left(\frac{D_i}{\pi_{\theta}(\mathbf{X}_i)} - \frac{1 - D_i}{1 - \pi_{\theta}(\mathbf{X}_i)}\right) \frac{\partial \pi_{\theta}(\mathbf{X}_i)}{\partial \theta}}_{\text{score for treatment model}}\right] = 0$$

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· Moment conditions for weighting estimators:

$$\begin{aligned} & \text{HT: } \mathbb{E}\left[\frac{D_iY_i}{\pi_{\theta}(\mathbf{X}_i)} - \mathbb{E}[Y_i(1)]\right] = \mathbb{E}\left[\frac{(1-D_i)Y_i}{1-\pi_{\theta}(\mathbf{X}_i)} - \mathbb{E}[Y_i(0)]\right] = 0 \\ & \text{Hajek: } \mathbb{E}\left[\frac{D_i(Y_i - \mathbb{E}[Y_i(1)])}{\pi_{\theta}(\mathbf{X}_i)}\right] = \mathbb{E}\left[\frac{(1-D_i)(Y_i - \mathbb{E}[Y_i(0)])}{1-\pi_{\theta}(\mathbf{X}_i)}\right] = 0 \end{aligned}$$

- If $\widehat{\pi}(\mathbf{X}_i)$ is estimated, how to estimate $\mathbb{V}[\widehat{\tau}_{\mathsf{ipw}}]$ or $\mathbb{V}[\widehat{\tau}_h]$?
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 Replace with sample versions and use delta method to get asymptotic variance.

Estimated versus known pscores

```
ht.est <- function(y, d, w) {</pre>
  n <- length(v)</pre>
  (1/n) * sum((v * d * w) - (v * (1-d) * w))
n <- 200
x \leftarrow rbinom(n, size = 1, prob = 0.5)
dprobs <- 0.5*x + 0.4*(1-x)
d <- rbinom(n, size = 1, prob = dprobs)</pre>
y < -5 * d - 10 * x + rnorm(n, sd = 5)
true.w <- ifelse(d == 1, 1/dprobs, 1/(1-dprobs))</pre>
pprobs <- predict(glm(d ~ x))
est.w <- ifelse(d == 1, 1/pprobs, 1/(1 - pprobs))
ht.est(y, d, est.w)
```

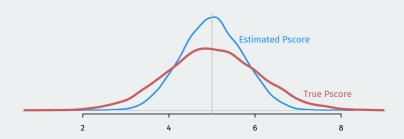
```
## [1] 5.22
ht.est(y, d, true.w)
```

```
## [1] 5.56
```

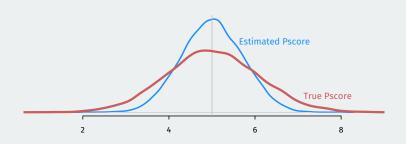
Sampling distribution of the HT estimators

```
sims <- 10000
true.holder <- rep(NA, sims)</pre>
est.holder <- rep(NA, sims)
for (i in 1:sims) {
  x \leftarrow rbinom(n, size = 1, prob = 0.5)
  dprobs <-0.5*x + 0.4*(1-x)
  d <- rbinom(n, size = 1, prob = dprobs)</pre>
  y < -5 * d - 10 * x + rnorm(n, sd = 5)
  true.w <- ifelse(d == 1, 1/dprobs, 1/(1-dprobs))
  pprobs <- predict(glm(d ~ x))</pre>
  est.w <- ifelse(d == 1, 1/pprobs, 1/(1 - pprobs))
  est.holder[i] <- ht.est(y, d, est.w)</pre>
  true.holder[i] <- ht.est(y, d, true.w)</pre>
```

Sampling distribution of the HT estimators



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var(est.holder)

[1] 0.506

var(true.holder)

[1] 1.15

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- True PS only adjusts for the **expected** differences between samples.
- · Only true if propensity score model is correctly specified!!

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Augmented IPW estimator combines regression and weighting:

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 - Can allow each model to be more flexible without hurting asymptotics.
- **Efficient**: lowest asymptotic variance among consistent estimators when PS model is correct.

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