Module 6(b): Two Stage Least Squares

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Gov 2003 (Harvard)

1/ Basic two-stage least squares

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- This implies the following CEF form for Y_i conditional on Z_i :

$$\mathbb{E}[Y_i \mid Z_i] = \alpha + \tau \, \mathbb{E}[D_i \mid Z_i] = \alpha + \tau \cdot (\gamma Z_i)$$

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- If the CEF is linear, we have this simple relationship slopes:

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$$\tau = \frac{\mathsf{cov}(Y_i, \gamma Z_i)}{\mathbb{V}[\gamma Z_i]} = \frac{\mathsf{cov}(Y_i, Z_i)}{\gamma \mathbb{V}[Z_i]} = \frac{\mathsf{cov}(Y_i, Z_i)}{\mathsf{cov}(D_i, Z_i)}$$

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- · TSLS estimator:
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 - Estimate $\widehat{\tau}_{2SLS}$ as the slope of a regression of Y_i on $\widehat{\gamma}Z_i$
 - Under this model, $\widehat{ au}_{2SLS} \overset{p}{ o} au$ (but don't use SEs from second stage)

• Under binary treatment/instrument, TSLS estimand is the LATE:

$$\tau = \frac{\mathsf{cov}(Y_i, Z_i)}{\mathsf{cov}(D_i, Z_i)} = \frac{\mathbb{E}[Y_i \mid Z_i = 1] - \mathbb{E}[Y_i \mid Z_i = 0]}{\mathbb{E}[D_i \mid Z_i = 1] - \mathbb{E}[Y_i \mid D_i = 0]} = \frac{\mathsf{ITT}_Y}{\mathsf{ITT}_D} = \tau_{\mathsf{LATE}}$$

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- Otherwise, au is an odd weighted function of causal effects and $au
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- Example completely irrelevant instrument:

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- · We can write the bias of the Wald estimator as:

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- Inconsistent and asymptotically heavy tails (bc of Cauchy)
 - When $Z \rightarrow D$ effect is small but non-zero we see similar behavior.

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- · Can invert (analytically!) to get confidence intervals

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 - If instrument can only increase by 1 dose, then simplifies to weighted average of principal strata effects.

2/ General two-stage least squares

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• Linear model for each i:

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Nasty Matrix Algebra

• Projection matrix projects values of \mathbf{X}_i onto \mathbf{Z}_i :

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• And plug in the sample values (the *n* cancels out):

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- In-sample projection matrix produces fitted values: $\widehat{\mathbb{X}}=\mathbb{Z}(\mathbb{Z}'\mathbb{Z})^{-1}\mathbb{Z}'\mathbb{X}$
 - Fitted values of regression of \mathbb{X} on \mathbb{Z} .
 - Matrix party trick: $\mathbb{X}'\mathbb{Z}/n = (1/n)\sum_{i=1}^{n} \mathbf{X}_{i}\mathbf{Z}'_{i} \stackrel{p}{\to} \mathbb{E}[\mathbf{X}_{i}\mathbf{Z}'_{i}].$
- Take the population formula for the parameters:

$$\boldsymbol{\beta} = (\mathbb{E}[\tilde{\mathbf{X}}_i \mathbf{X}_i'])^{-1} \mathbb{E}[\tilde{\mathbf{X}}_i Y_i]$$

• And plug in the sample values (the *n* cancels out):

$$\widehat{\boldsymbol{\beta}}_{2SLS} = (\widehat{\mathbb{X}}'\mathbb{X})^{-1}\widehat{\mathbb{X}}'\mathbf{y} \overset{p}{\rightarrow} \boldsymbol{\beta}$$

• This is how R/Stata estimates the 2SLS parameters

· We can write the centered, normalized TSLS estimator as:

$$\sqrt{n}(\hat{\beta}_{2SLS} - \beta) = \underbrace{\left(n^{-1}\sum_{i}\widehat{\mathbf{X}}_{i}\widehat{\mathbf{X}}_{i}'\right)^{-1}}_{\stackrel{\rho}{\rightarrow} (\mathbb{E}[\widehat{\mathbf{X}}_{i}\widehat{\mathbf{X}}_{i}'])^{-1}} \underbrace{\left(n^{-1/2}\sum_{i}\widehat{\mathbf{X}}_{i}\varepsilon_{i}\right)}_{\stackrel{d}{\rightarrow} N(0,\mathbb{E}[\widehat{\mathbf{X}}_{i}'\varepsilon_{i}'\varepsilon_{i}\widehat{\mathbf{X}}_{i}])}$$

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• Thus, we have that $\sqrt{n}(\hat{\beta}_{2SLS} - \beta)$ has asymptotic variance:

$$(\mathbb{E}[\widehat{\mathbf{X}}_i\widehat{\mathbf{X}}_i'])^{-1}\mathbb{E}[\widehat{\mathbf{X}}_i'\varepsilon_i'\varepsilon_i\widehat{\mathbf{X}}_i](\mathbb{E}[\widehat{\mathbf{X}}_i\widehat{\mathbf{X}}_i'])^{-1}$$

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• Robust 2SLS variance estimator with residuals $\hat{u}_i = Y_i - \mathbf{X}_i'\hat{\boldsymbol{\beta}}$:

$$\widehat{\mathrm{var}}(\widehat{\boldsymbol{\beta}}_{\mathrm{2SLS}}) = (\widehat{\mathbb{X}}'\widehat{\mathbb{X}})^{-1} \Big(\sum_{i} \widehat{u}_{i}^{2} \widehat{\mathbf{X}}_{i} \widehat{\mathbf{X}}_{i}' \Big) (\widehat{\mathbb{X}}'\widehat{\mathbb{X}})^{-1}$$

We can write the centered, normalized TSLS estimator as:

$$\sqrt{n}(\widehat{\beta}_{2SLS} - \beta) = \underbrace{\left(n^{-1}\sum_{i}\widehat{\mathbf{X}}_{i}\widehat{\mathbf{X}}_{i}'\right)^{-1}}_{\stackrel{P}{\rightarrow} (\mathbb{E}[\widehat{\mathbf{X}}_{i}\widehat{\mathbf{X}}_{i}'])^{-1}} \underbrace{\left(n^{-1/2}\sum_{i}\widehat{\mathbf{X}}_{i}\varepsilon_{i}\right)}_{\stackrel{d}{\rightarrow} N(0,\mathbb{E}[\widehat{\mathbf{X}}_{i}'\varepsilon_{i}'\varepsilon_{i}\widehat{\mathbf{X}}_{i}])}$$

• Thus, we have that $\sqrt{n}(\hat{\beta}_{2SLS} - \beta)$ has asymptotic variance:

$$(\mathbb{E}[\widehat{\mathbf{X}}_i\widehat{\mathbf{X}}_i'])^{-1}\mathbb{E}[\widehat{\mathbf{X}}_i'\varepsilon_i'\varepsilon_i\widehat{\mathbf{X}}_i](\mathbb{E}[\widehat{\mathbf{X}}_i\widehat{\mathbf{X}}_i'])^{-1}$$

• Robust 2SLS variance estimator with residuals $\hat{u}_i = Y_i - \mathbf{X}_i'\hat{\boldsymbol{\beta}}$:

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· HC2, clutering, and autocorrelation versions exist