

# Module 6: Noncompliance and Instrumental Variables

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Matthew Blackwell

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# Where are we? Where are we going?

- We've covered randomized experiments (no confounding).
- We've covered selection on observables (no unmeasured confounding).
- What if there is unmeasured confounding? What can we do?
- First approach we'll explore: instrumental variables.
  - First: motivate IV through experiments and noncompliance.
  - Then: how does this relate to classical econometric methods like TSLS?

# 1/ Randomized experiments with noncompliance

# Noncompliance

- GOTV experiment with door-to-door canvassing.
- Households are randomized so treatment assignment is unconfounded.
  - $Z_i = 1$  for assigned to treatment (canvassing attempted),
  - $Z_i = 0$  for assigned to control (no canvassing attempted).
- **Noncompliance:** units don't follow treatment assignment.
  - Units assigned to treatment take control or vice versa.
  - $D_i = 1$  for actually took treatment (heard canvasser message).
  - $D_i = 0$  for actually took control (didn't answer the door).
  - Full compliance means  $Z_i = D_i$  for all  $i$

# How to handle noncompliance

- Two approaches common seen in applied studies.
- **Intent-to-treat** analysis (ITT): just use randomization.
  - Use  $Z_i$  as the treatment and analyze as a typical experiment.
  - Downside: can't learn about the effect of actually being canvassed.
- **As-treated** analysis: just use treatment uptake.
  - Act as if  $D_i$  was randomly assigned.
  - Not valid if uptake is **correlated** with the outcome.
  - $\rightsquigarrow$  unmeasured confounding between  $D_i$  and  $Y_i$  and bias.
- Alternative: leverage latent strata of **compliance types**

# Setup

- Treatment assignment,  $Z_i \in \{0, 1\}$ , treatment uptake  $D_i \in \{0, 1\}$
- Treatment uptake now affected by assignment:  $D_i(z)$ 
  - $D_i(1) = 1$  if assigned to canvassing, I **would** open my door.
  - $D_i(1) = 0$  if assigned to canvassing, I **would not** open my door.
  - Noncompliance means  $D_i(z) \neq z$  for some  $i$ .
- Consistency for the observed treatment as usual:

$$D_i = Z_i D_i(1) + (1 - Z_i) D_i(0)$$

- Canvassing is an example of **one-sided noncompliance**.
  - People might refuse treatment when offered ( $D_i(1) = 0$ )
  - But no one receives treatment if in control ( $D_i(0) = 0, \forall i$ )
  - **Two-sided noncompliance** is when you can refuse to comply with treatment **or** control.

# Potential outcomes

- Outcomes might depend on assignment and uptake:  $Y_i(z, d)$ .
  - $Y_i(1, 1)$ : would I vote if I were assigned to canvassing and received it?
- Can only observe two potential outcomes:  $Y_i(1, D_i(1))$  and  $Y_i(0, D_i(0))$ .
  - $Y_i(1, D_i(1))$ : potential outcome when assigned canvassing and whatever uptake occurs for unit  $i$  when assigned to canvassing.
  - $Y_i(1, 1 - D_i(1))$  not possible to ever observe (cross-world or a prior counterfactual)
- Consistency assumption:  $Y_i = Y_i(Z_i, D_i(Z_i))$

# Some notation

- Let's use 0/1 subscripts for assignment and t/c subscripts for uptake:

$$n_1 = \sum_{i=1}^n Z_i \quad n_0 = \sum_{i=1}^n 1 - Z_i \quad n_t = \sum_{i=1}^n D_i \quad n_c = \sum_{i=1}^n 1 - D_i$$

- Average outcomes and uptake in each assignment group:

$$\begin{aligned} \bar{Y}_1 &= \frac{1}{n_1} \sum_{i=1}^n Z_i Y_i & \bar{Y}_0 &= \frac{1}{n_0} \sum_{i=1}^n (1 - Z_i) Y_i \\ \bar{D}_1 &= \frac{1}{n_1} \sum_{i=1}^n Z_i D_i & \bar{D}_0 &= \frac{1}{n_0} \sum_{i=1}^n (1 - Z_i) D_i \end{aligned}$$

- Assumption 1: **randomization**  $[\{Y_i(d, z), \forall d, z\}, D_i(1), D_i(0)] \perp\!\!\!\perp Z_i$ 
  - For observational uses of IV, might condition on some  $\mathbf{X}_i$ .



# ITT effects

- **Intent-to-treat** (ITT) effects are just the ATEs of  $Z_i$

$$\text{ITT}_D = \frac{1}{n} \sum_{i=1}^n D_i(1) - D_i(0) \qquad \text{ITT}_Y = \frac{1}{n} \sum_{i=1}^n Y_i(1, D_i(1)) - Y_i(0, D_i(0))$$

- SATE of assignment on treatment uptake and the outcome.
- If noncompliance is one-sided, then  $\text{ITT}_D \geq 0$
- Standard estimators for these quantities:

$$\widehat{\text{ITT}}_D = \bar{D}_1 - \bar{D}_0 \qquad \widehat{\text{ITT}}_Y = \bar{Y}_1 - \bar{Y}_0$$

- Under randomization of  $Z_i$ , everything just like Neyman approach.
  - Variances, tests, CIs all standard.
- Problem:  $\text{ITT}_Y$  is a combination of true effect of  $D_i$  and noncompliance.
  - Effect of  $D_i$  is maybe more externally valid than  $Z_i$ .

## **2/** Compliance types

# Compliance status

- We can stratify units by their **compliance type**.
  - Compliance type is how they would respond to treatment assignment.
  - Basically it's the value of  $(D_i(0), D_i(1))$  for any unit.
- Under one-sided noncompliance, there are two types:
  - **Compliers** with  $D_i(1) = 1$  and **noncompliers** with  $D_i(1) = 0$ .
  - Compliers answer the door when assigned to canvassing
  - Noncompliers don't answer the door when assigned to canvassing
  - Everyone has  $D_i(0) = 0$ , so no noncompliance there.
- Compliance is a function of potential outcomes so it is **pretreatment!**
  - $\rightsquigarrow$  treatment assignment independent of  $C_i$

# ITTs among the compliance groups

- Compliance type indicator  $C_i \in \{\text{co}, \text{nc}\}$ .
  - Number of compliers:  $n_{\text{co}} = \sum_{i=1} \mathbf{1}(C_i = \text{co})$ .
  - Proportion of compliers:  $\pi_{\text{co}} = n_{\text{co}}/n$
  - Same for noncompliers:  $n_{\text{nc}}$  and  $\pi_{\text{nc}}$
- ITT on uptake directly related to compliance type:

$$\text{ITT}_{D,\text{co}} = \frac{1}{n_{\text{co}}} \sum_{i: C_i = \text{co}} D_i(1) - D_i(0) = 1$$

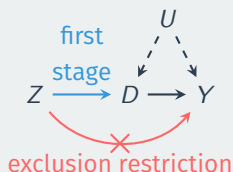
$$\text{ITT}_{D,\text{nc}} = \frac{1}{n_{\text{nc}}} \sum_{i: C_i = \text{nc}} D_i(1) - D_i(0) = 0$$

- Intuition: no effect of assignment on uptake for noncompliers!
- Implies overall ITT on uptake is equal to the **proportion of compliers**

$$\text{ITT}_D = \pi_{\text{co}} \text{ITT}_{D,\text{co}} + \pi_{\text{nc}} \text{ITT}_{D,\text{nc}} = \pi_{\text{co}}$$

## **3/** Instrumental variables

# Exclusion restriction



- Assumption 2: **first-stage**  $ITT_D = \pi_{co} \neq 0$ 
  - At least one person complies with treatment.
- Assumption 3: **exclusion restriction**  $Z_i$  only affects  $Y_i$  through  $D_i$ 
  - $Y_i(z, d) = Y_i(z', d)$  for all  $z, z'$  and  $d$ .
  - Assignment to canvassing only affects turnout through actual canvassing.
  - Not a testable assumption and can't be guaranteed by design.
- Implies that potential outcomes only a function of  $D_i$ :

$$Y_i(1) = Y_i(D_i = 1) = Y_i(Z_i = 1, D_i = 1) = Y_i(Z_i = 1, D_i = 1)$$

$$Y_i(0) = Y_i(D_i = 0) = Y_i(Z_i = 1, D_i = 0) = Y_i(Z_i = 1, D_i = 0)$$

# Outcome ITTs and compliance types

- We can define the ITTs on the outcome by compliance type as well.
  - $ITT_{Y,co}$  effect of assignment on outcome among compliers.
  - $ITT_{Y,nc}$  effect of assignment on outcome among noncompliers.
  - Only  $ITT_{Y,co}$  actually picks up an effect of  $D_i$
- Exclusion restriction has implications for these:
  - Implies that  $ITT_{Y,nc} = 0$ : if  $D_i$  doesn't change,  $Y_i$  can't change.
  - Implies that  $ITT_{Y,co}$  is due entirely to treatment uptake.
- Allows us to connect the ITT on the outcome to compliance groups:

$$ITT_Y = \pi_{co} ITT_{Y,co} + \pi_{nc} ITT_{Y,nc} = ITT_D ITT_{Y,co}$$

- Under the exclusion restriction,  $ITT_{Y,co}$  is the effect of treatment receipt:

$$\begin{aligned} ITT_{Y,co} &= \frac{1}{n_{co}} \sum_{i: C_i=co} Y_i(1, D_i(1)) - Y_i(0, D_i(0)) \\ &= \frac{1}{n_{co}} \sum_{i: C_i=co} Y_i(D_i = 1) - Y_i(D_i = 0) = \tau_{LATE} \end{aligned}$$

- This quantity is the **local ATE** (LATE), local to compliers.
  - It's a conditional ATE, where we condition on being a complier.
  - Also called the **complier average causal effect** (CACE).
- LATE Theorem** under one-sided noncompliance, exclusion restriction, first-stage, and randomization:

$$\tau_{LATE} = ITT_{Y,co} = \frac{ITT_Y}{ITT_D}$$



# Wald estimator

- **Wald** or **instrumental variables estimator** for the LATE:

$$\hat{\tau}_{iv} = \frac{\widehat{ITT}_Y}{\widehat{ITT}_D}$$

- Ratio of the two unbiased ITT estimators.
- Not unbiased, but it is **consistent** for  $\tau_{LATE}$ .
- Equivalent to the **two-stage least squares** estimator:
  - Regress  $D_i$  on  $Z_i$  and get fitted values  $\widehat{D}_i$
  - Regress  $Y_i$  on  $\widehat{D}_i$
- We can use the delta method to find the (superpopulation) variance:

$$\mathbb{V}[\hat{\tau}_{iv}] = \frac{1}{\widehat{ITT}_D^2} \mathbb{V}[\widehat{ITT}_Y] + \frac{\widehat{ITT}_Y^2}{\widehat{ITT}_D^4} \mathbb{V}[\widehat{ITT}_D] - 2 \frac{\widehat{ITT}_Y}{\widehat{ITT}_D^3} \text{cov}[\widehat{ITT}_Y, \widehat{ITT}_D]$$

## 4/ Two-sided noncompliance

# Two-sided noncompliance

- Two-sided noncompliance: those in control can select into treatment.
- **Encouragement design:** randomly assign an encouragement of some treatment.
  - Some may refuse encouragement and opt to not take treatment.
  - Some may take treatment even without encouragement.
- $Z_i$  is the encouragement and  $D_i$  is the treatment.
- No change in estimation, just different identification assumptions.

# Compliance types

- Four compliance types (or **principal strata**) in this setting:

- Complier  $D_i(1) = 1$  and  $D_i(0) = 0$
- Always-taker  $D_i(1) = D_i(0) = 1$
- Never-taker  $D_i(1) = D_i(0) = 0$
- Defier  $D_i(1) = 0$  and  $D_i(0) = 1$

- Connections between observed data and compliance types:

	$Z_i = 0$	$Z_i = 1$
$D_i = 0$	Never-taker or Complier	Never-taker or Defier
$D_i = 1$	Always-taker or Defier	Always-taker or Complier

- Let  $\pi_{co}$ ,  $\pi_{at}$ ,  $\pi_{nt}$ , and  $\pi_{de}$  be the proportions of each type.
- ITT effects on  $D_i$  are more murky:  $ITT_D = \pi_{co} - \pi_{de}$ 
  - Defiers really make things messy!

# Instrumental variables assumptions

- Canonical IV assumptions for  $Z_i$  to be a valid instrument:
  1. Randomization of  $Z_i$
  2. Presence of some compliers  $\pi_{co} \neq 0$  (first-stage)
  3. Exclusion restriction  $Y_i(z, d) = Y_i(z', d)$
  4. **Monotonicity**:  $D_i(1) \geq D_i(0)$  for all  $i$  (no defiers)
- Implies ITT effect on treatment equals proportion compliers:  $ITT_D = \pi_{co}$
- Implies ITT for the outcome has the same interpretation:

$$\begin{aligned} ITT_Y &= ITT_{Y,co} \pi_{co} + \underbrace{ITT_{Y,at}}_{=0 \text{ (ER)}} \pi_{at} + \underbrace{ITT_{Y,nt}}_{=0 \text{ (ER)}} \pi_{nt} + ITT_{Y,de} \underbrace{\pi_{de}}_{=0 \text{ (mono)}} \\ &= ITT_{co} \pi_{co} \end{aligned}$$

- $\rightsquigarrow$  same identification result:  $\tau_{LATE} = ITT_Y / ITT_D$

# Is the LATE useful?

- The LATE is a unknown subset of the data.
  - Treated units are a mix of always takers and compliers.
  - Control units are a mix of never takers and compliers.
- Without further assumptions,  $\tau_{\text{LATE}} \neq \tau$ .
- Complier group depends on the instrument  $\rightsquigarrow$  different IVs will lead to different identified estimands.
- But we cannot do any better in terms of point estimation without more assumptions.
  - Alternative: bound the ATE?