Module 4: Linear Regression and Randomized Experiments

Fall 2021

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Gov 2003 (Harvard)

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 - **Simplicity**: known tool that is already very common.
 - Increased precision: we may want to add covariates for more precise effect estimates.

1/ Regression with no covariates

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- \rightsquigarrow standard Neyman analysis for unbiasedness, sampling variance.
- Generalizes to discrete treatments with > 2 levels.

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 - Intercept $\alpha = \mathbb{E}[Y_i(0)]$ average control outcome.
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 - Error is deviation for control PO + treatment effect heterogeneity.

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- Randomization + consistency → linear model.
 - Does not imply homoskedasticity or normal errors, though!

Homoskedasticity

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- Under homoskedasticity, variance of the OLS estimator is:

$$\mathbb{V}[\widehat{ au}_{\mathsf{ols}} \mid \mathbf{D}] = rac{\sigma^2}{\sum_{i=1}^n \left(D_i - \overline{D}
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• "Standard" variance estimator under homoskedasticity:

$$\hat{\mathbb{V}}_{const} = \frac{\frac{1}{n-2}\sum_{i=1}^n \hat{\mathcal{E}}_i^2}{\sum_{i=1}^n (D_i - \overline{D})^2} = \frac{\frac{1}{n-2}\sum_{i=1}^n (Y_i - \hat{\alpha}_{\text{ols}} - \widehat{\tau}_{\text{ols}} D_i)^2}{\sum_{i=1}^n (D_i - \overline{D})^2}$$

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• We can rewrite this as a function of the **pooled** variance $\widehat{\sigma}_{Y|D}^2$:

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 - Homoskedasticity holds: $\sigma_1^2 = \sigma_0^2$
 - Design is balanced: $n_1 = n_0$

• Eicker-Huber-White (EHW) robust/sandwich variance estimator:

$$\begin{split} \widehat{\mathbb{V}}_{\mathsf{EHW}} &= \left(\sum_{i=1}^{n} \mathbf{X}_{i} \mathbf{X}_{i}'\right)^{-1} \left(\sum_{i=1}^{n} \widehat{\varepsilon}_{i}^{2} \mathbf{X}_{i} \mathbf{X}_{i}'\right) \left(\sum_{i=1}^{n} \mathbf{X}_{i} \mathbf{X}_{i}'\right)^{-1} \\ &= \left(\mathbb{X}' \mathbb{X}\right)^{-1} \left(\sum_{i=1}^{n} \widehat{\varepsilon}_{i}^{2} \mathbf{X}_{i} \mathbf{X}_{i}'\right) \left(\mathbb{X}' \mathbb{X}\right)^{-1} \quad \text{where} \quad \mathbb{X} = \begin{bmatrix} 1 & \mathbf{D} \end{bmatrix} \end{split}$$

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• Recall the PATE-targeted variance of the difference-in-means:

$$\mathbb{V}(\widehat{\tau}_{\mathsf{diff}}) = \frac{\sigma_0^2}{n_0} + \frac{\sigma_1^2}{n_1} = \frac{\mathbb{V}[Y_i(0)]}{n_0} + \frac{\mathbb{V}[Y_i(1)]}{n_1}$$

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- To see this, we can derive $\hat{\mathbb{V}}_{\text{EHW}}$ under our case:

$$\hat{\mathbb{V}}_{\text{EHW}} = \frac{\tilde{\sigma}_1^2}{n_1} + \frac{\tilde{\sigma}_0^2}{n_0}, \quad \text{where} \quad \tilde{\sigma}_d^2 = \frac{1}{n_d} \sum_{i:D_i = d} \left(Y_i - \overline{Y}_d \right)^2$$

• Eicker-Huber-White (EHW) robust/sandwich variance estimator:

$$\begin{split} \hat{\mathbb{V}}_{\mathsf{EHW}} &= \left(\sum_{i=1}^{n} \mathbf{X}_{i} \mathbf{X}_{i}'\right)^{-1} \left(\sum_{i=1}^{n} \hat{\varepsilon}_{i}^{2} \mathbf{X}_{i} \mathbf{X}_{i}'\right) \left(\sum_{i=1}^{n} \mathbf{X}_{i} \mathbf{X}_{i}'\right)^{-1} \\ &= \left(\mathbb{X}'\mathbb{X}\right)^{-1} \left(\sum_{i=1}^{n} \hat{\varepsilon}_{i}^{2} \mathbf{X}_{i} \mathbf{X}_{i}'\right) \left(\mathbb{X}'\mathbb{X}\right)^{-1} \quad \text{where} \quad \mathbb{X} = \begin{bmatrix} 1 & \mathbf{D} \end{bmatrix} \end{split}$$

• Recall the PATE-targeted variance of the difference-in-means:

$$\mathbb{V}(\widehat{\tau}_{\mathsf{diff}}) = \frac{\sigma_0^2}{n_0} + \frac{\sigma_1^2}{n_1} = \frac{\mathbb{V}[Y_i(0)]}{n_0} + \frac{\mathbb{V}[Y_i(1)]}{n_1}$$

• To see this, we can derive \hat{V}_{EHW} under our case:

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 ~ simple OLS + HC2 = unbiased point and variance estimator.

2/ Linear regression with covariates

$$(\widehat{\tau}_{\mathrm{adj}}, \widehat{\alpha}_{\mathrm{adj}}, \widehat{\beta}_{\mathrm{adj}}) = \operatorname*{arg\,min}_{\tau,\alpha,\beta} \sum_{i=1}^{n} \left(Y_i - \alpha - \tau \, D_i - \widetilde{\mathbf{X}}_i' \beta \right)^2$$

· What if we add covariates to our regression model?

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- Estimation: EHW robust variance estimators are consistent or asymptotically conservative for V[\(\hat{\tau}_{adi}\)]

• OLS estimator from fully interacted model, $\widehat{\tau}_{\text{inter}}$:

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- Always use robust/HC2 variance estimators unless you have good reasons.

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 - · Latter two allow for additional covariates to be added.

3/ Cluster randomized experiments

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- · Quantity of interest still at individual level:

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- · Neyman-style conservative variance:

$$\mathbb{V}[\widehat{\tau}_{\operatorname{cl}} \mid \mathbf{0}] \leq \frac{\mathbb{V}[\overline{Y}_k(1)]}{J_1} + \frac{\mathbb{V}[\overline{Y}_k(0)]}{J_0} \quad \text{where for } d = 0, 1 \quad \overline{Y}_k(d) = \frac{1}{m} \sum_{m=1}^m Y_{ik}(d)$$

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