

Bayesian Nonparametrics

some research vignettes

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**MLSS Tübingen Lectures
2013**

Probabilistic Modelling

- A model describes data that one could observe from a system
- If we use the mathematics of probability theory to express all forms of uncertainty and noise associated with our model...
- ...then *inverse probability* (i.e. Bayes rule) allows us to infer unknown quantities, adapt our models, make predictions and learn from data.

Why Bayesian Nonparametrics...?

- **Why Bayesian?**

Simplicity (of the framework)

- **Why nonparametrics?**

Complexity (of real world phenomena)

Parametric vs Nonparametric Models

- *Parametric models* assume some finite set of parameters θ . Given the parameters, future predictions, x , are independent of the observed data, \mathcal{D} :

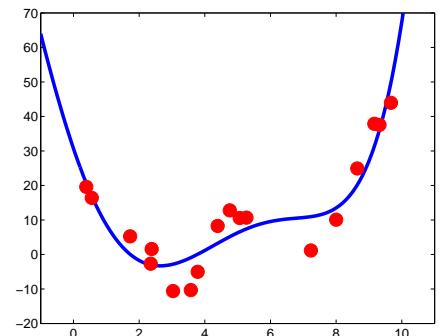
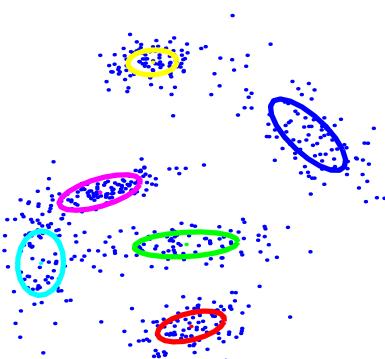
$$P(x|\theta, \mathcal{D}) = P(x|\theta)$$

therefore θ capture everything there is to know about the data.

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- So the complexity of the model is bounded even if the amount of data is unbounded. This makes them not very flexible.
 - *Non-parametric models* assume that the data distribution cannot be defined in terms of such a finite set of parameters. But they can often be defined by assuming an *infinite dimensional* θ . Usually we think of θ as a *function*.
 - The amount of information that θ can capture about the data \mathcal{D} can grow as the amount of data grows. This makes them more flexible.
-

Why nonparametrics?

- flexibility
- better predictive performance
- more realistic



Almost all successful methods in machine learning are essentially nonparametric¹:

- kernel methods / SVM / GP
- deep networks / large neural networks
- k-nearest neighbors, ...

¹or highly scalable!

Examples of non-parametric models

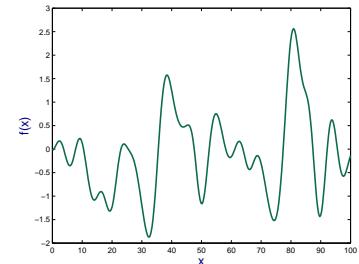
Bayesian nonparametrics has many uses.

Parametric	Non-parametric	Process	Application
polynomial regression	Gaussian processes	GP	function approx.
logistic regression	Gaussian process classifiers	GP	classification
mixture models, k-means	Dirichlet process mixtures	DP / CRP	clustering
hidden Markov models	infinite HMMs	HDP	time series
factor analysis/pPCA/PMF	infinite latent factor models	BP / IBP	feature discovery
...			

Gaussian and Dirichlet Processes

- Gaussian processes define a distribution on functions

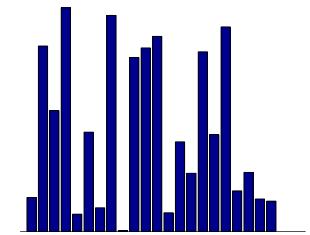
$$f \sim \text{GP}(\cdot | \mu, c)$$



where μ is the mean function and c is the covariance function.
We can think of GPs as “infinite-dimensional” Gaussians

- Dirichlet processes define a distribution on distributions

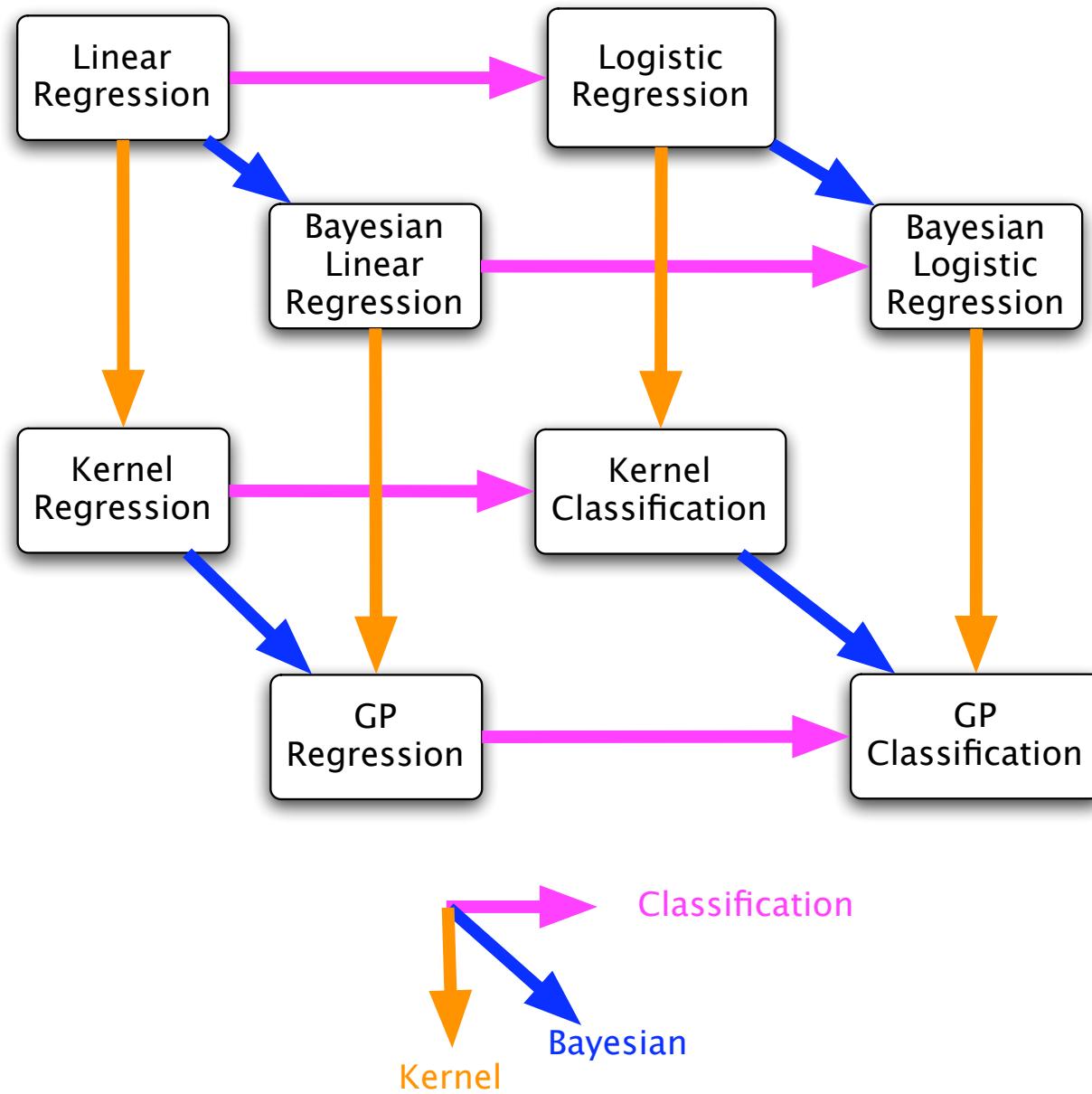
$$G \sim \text{DP}(\cdot | G_0, \alpha)$$



where $\alpha > 0$ is a concentration parameter, and G_0 is the base measure.
We can think of DPs as “infinite-dimensional” Dirichlet distributions.

Note that both f and G are infinite dimensional objects.

A picture



Dirichlet Distribution

The **Dirichlet distribution** is a distribution on the K -dim probability simplex.

Let \mathbf{p} be a K -dimensional vector s.t. $\forall j : p_j \geq 0$ and $\sum_{j=1}^K p_j = 1$

$$P(\mathbf{p}|\boldsymbol{\alpha}) = \text{Dir}(\alpha_1, \dots, \alpha_K) \stackrel{\text{def}}{=} \frac{\Gamma(\sum_j \alpha_j)}{\prod_j \Gamma(\alpha_j)} \prod_{j=1}^K p_j^{\alpha_j - 1}$$

where the first term is a normalization constant² and $E(p_j) = \alpha_j / (\sum_k \alpha_k)$

The Dirichlet is conjugate to the multinomial distribution. Let

$$c|\mathbf{p} \sim \text{Multinomial}(\cdot|\mathbf{p})$$

That is, $P(c=j|\mathbf{p}) = p_j$. Then the posterior is also Dirichlet:

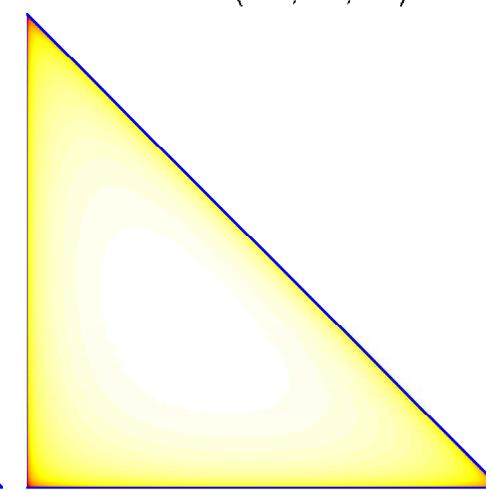
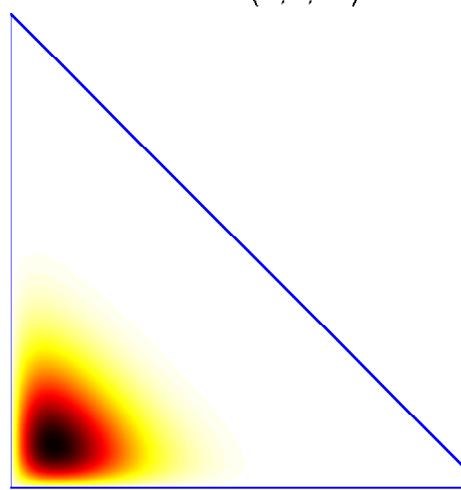
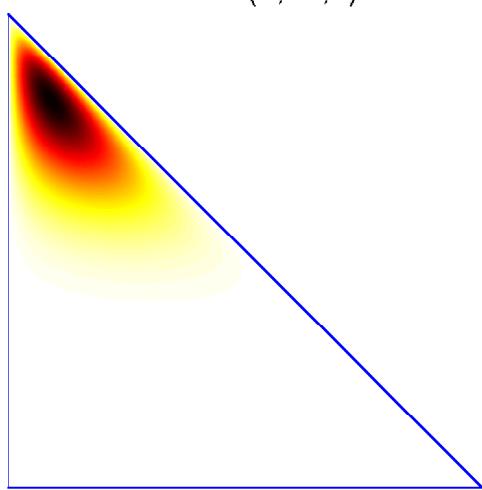
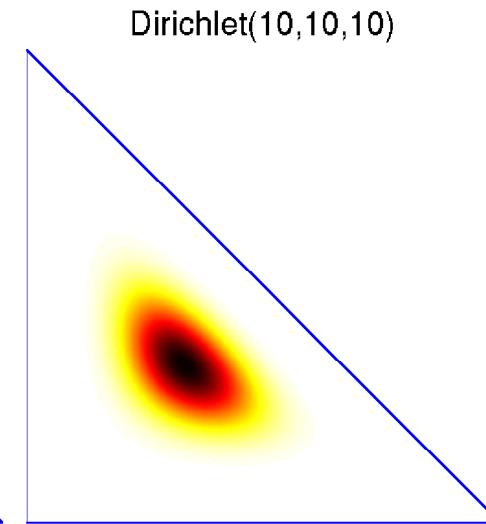
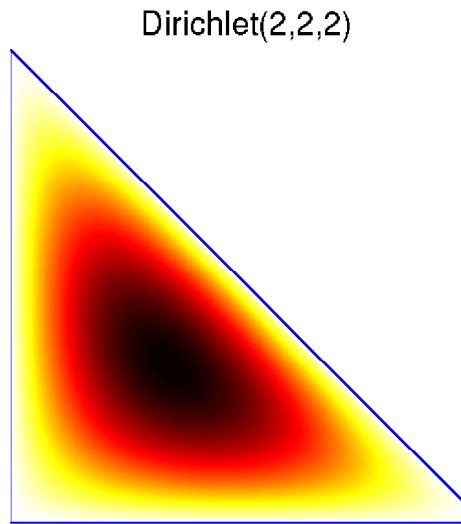
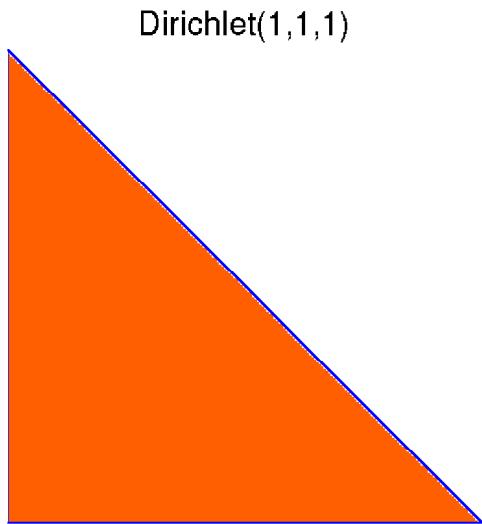
$$P(\mathbf{p}|c=j, \boldsymbol{\alpha}) = \frac{P(c=j|\mathbf{p})P(\mathbf{p}|\boldsymbol{\alpha})}{P(c=j|\boldsymbol{\alpha})} = \text{Dir}(\boldsymbol{\alpha}')$$

where $\alpha'_j = \alpha_j + 1$, and $\forall \ell \neq j : \alpha'_{\ell} = \alpha_{\ell}$

² $\Gamma(x) = (x-1)\Gamma(x-1) = \int_0^\infty t^{x-1} e^{-t} dt$. For integer n , $\Gamma(n) = (n-1)!$

Dirichlet Distributions

Examples of Dirichlet distributions over $\theta = (\theta_1, \theta_2, \theta_3)$ which can be plotted in 2D since $\theta_3 = 1 - \theta_1 - \theta_2$:



Dirichlet Process

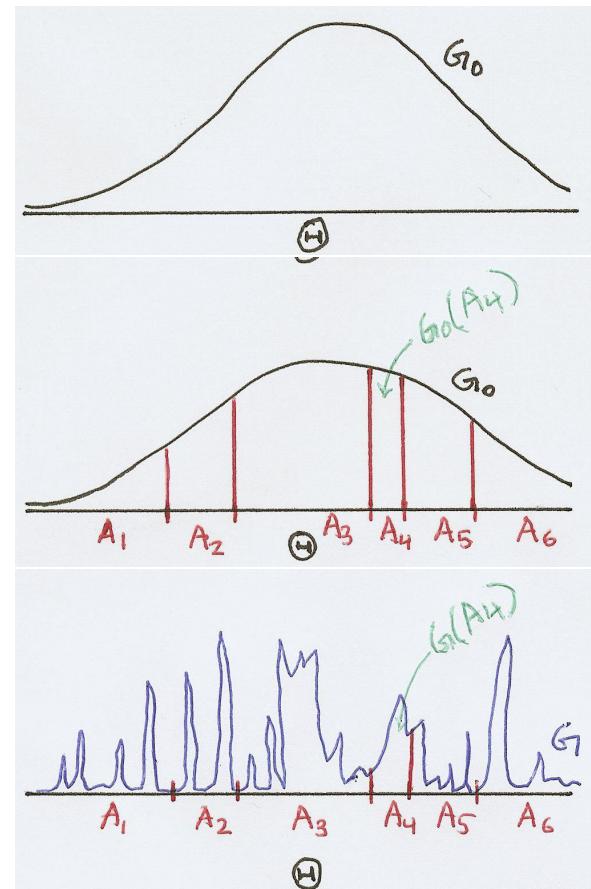
Let Θ be a measurable space, G_0 be a probability measure on Θ , and α a positive real number.

For all (A_1, \dots, A_K) finite partitions of Θ ,

$$G \sim \text{DP}(\cdot | G_0, \alpha)$$

means that

$$(G(A_1), \dots, G(A_K)) \sim \text{Dir}(\alpha G_0(A_1), \dots, \alpha G_0(A_K))$$



(Ferguson, 1973)

Dirichlet Process

$$G \sim \text{DP}(\cdot | G_0, \alpha)$$

OK, but what does it look like?

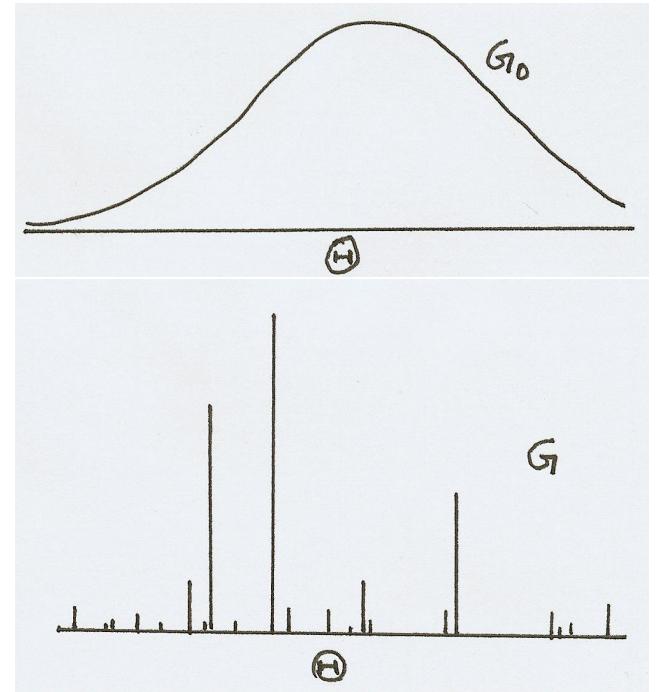
Samples from a DP are **discrete with probability one**:

$$G(\theta) = \sum_{k=1}^{\infty} \pi_k \delta_{\theta_k}(\theta)$$

where $\delta_{\theta_k}(\cdot)$ is a Dirac delta at θ_k , and $\theta_k \sim G_0(\cdot)$.

Note: $E(G) = G_0$

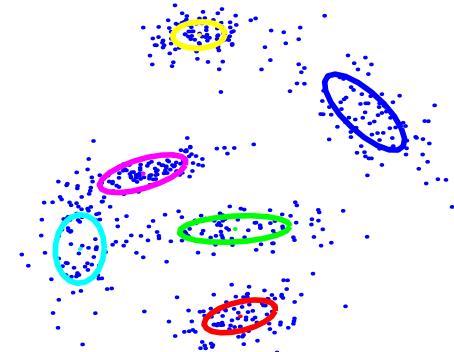
As $\alpha \rightarrow \infty$, G looks more “like” G_0 .



Clustering

Infinite mixture models

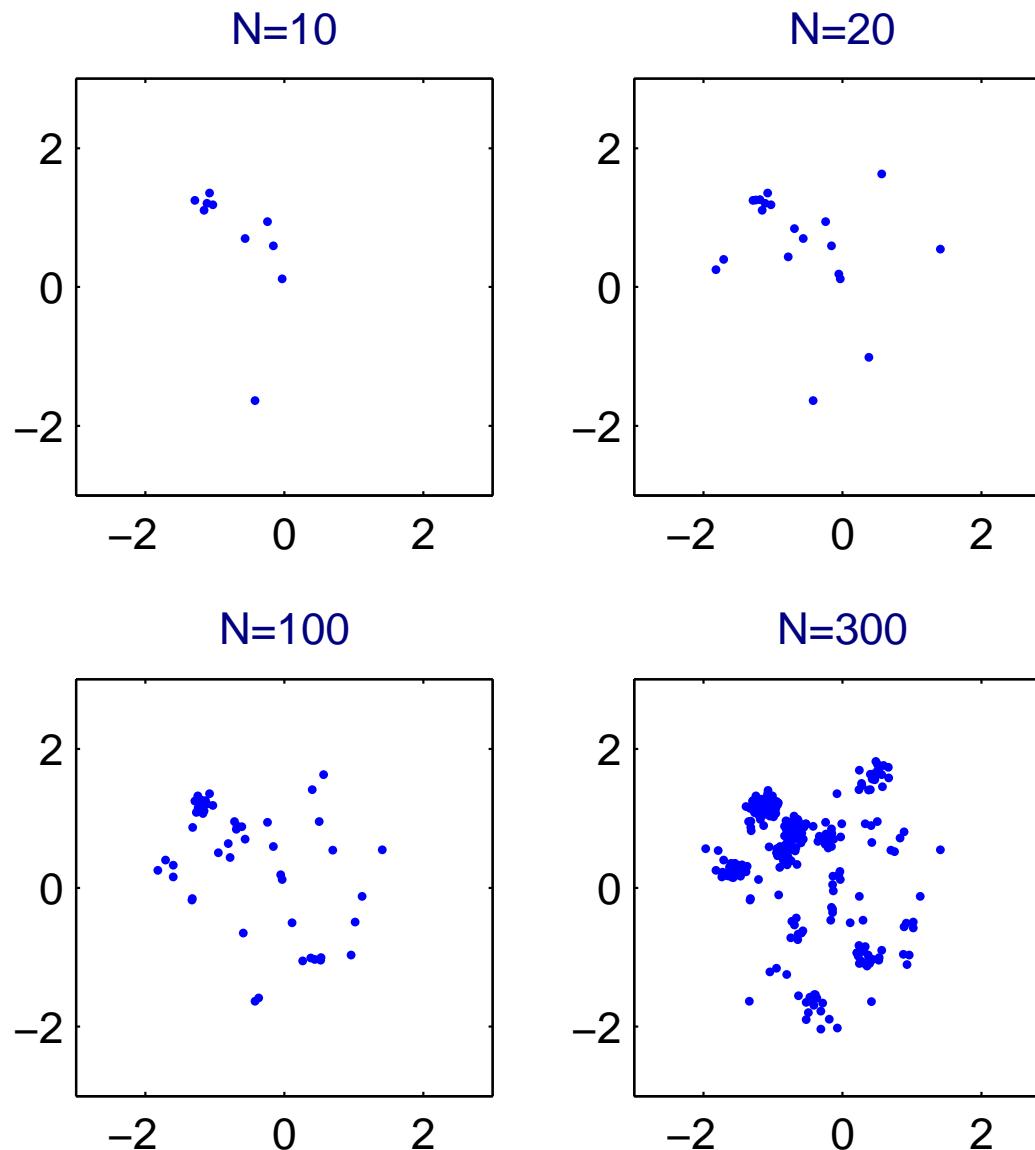
(e.g. Dirichlet Process Mixtures)



Why?

- You might not believe a priori that your data comes from a finite number of mixture components (e.g. strangely shaped clusters; heavy tails; structure at many resolutions)
- Inflexible models (e.g. a mixture of 6 Gaussians) can yield unreasonable inferences and predictions.
- For many kinds of data, the number of clusters might grow over time: clusters of news stories or emails, classes of objects, etc.
- You might want your method to automatically infer the number of clusters in the data.

Samples from a Dirichlet Process Mixture of Gaussians



Notice that more structure (clusters) appear as you draw more points.
(figure inspired by Neal)

Infinite mixture models

$$p(x) = \sum_{k=1}^K \pi_k p_k(x)$$

How?

- Start from a finite mixture model with K components and take the limit as number of components $K \rightarrow \infty$
- If you use symmetric Dirichlet priors on the $\{\pi_k\}$ you get a Dirichlet Process Mixture; the distribution over partitions is given by a Chinese Restaurant Process.
- But you have infinitely many parameters!
- Rather than optimize the parameters (ML, MAP), you integrate them out (Bayes) using, e.g:
 - MCMC sampling (Escobar & West 1995; Neal 2000; Rasmussen 2000)...
 - expectation propagation (EP; Minka and Ghahramani, 2003)
 - variational methods (Blei and Jordan, 2005)
 - Bayesian hierarchical clustering (Heller and Ghahramani, 2005)

Relationship between DPs and CRPs

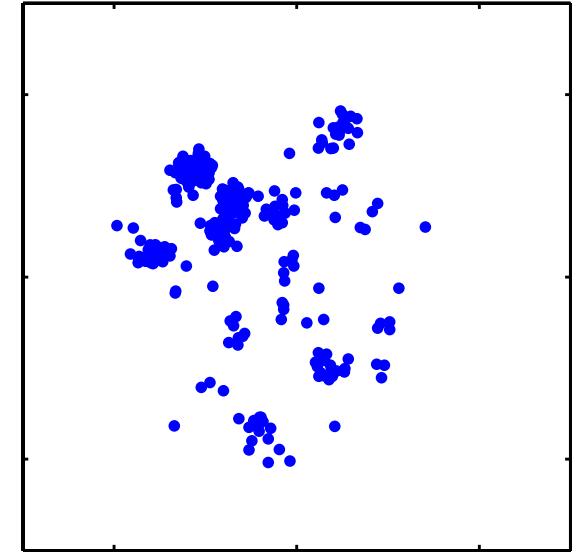
- DP is a **distribution on distributions**
- DP results in discrete distributions, so if you draw n points you are likely to get repeated values
- A DP induces a **partitioning** of the n points
e.g. $(1\ 3\ 4)\ (2\ 5) \Leftrightarrow \theta_1 = \theta_3 = \theta_4 \neq \theta_2 = \theta_5$
- Chinese Restaurant Process (CRP) defines the corresponding **distribution on partitions**
- Although the CRP is a sequential process, the distribution on $\theta_1, \dots, \theta_n$ is exchangeable (i.e. invariant to permuting the indices of the θ s): e.g.

$$P(\theta_1, \theta_2, \theta_3, \theta_4) = P(\theta_2, \theta_4, \theta_3, \theta_1)$$

Dirichlet Processes: Big Picture

There are many ways to derive the Dirichlet Process:

- Dirichlet distribution
- Urn model
- Chinese restaurant process
- Stick breaking
- Gamma process



DP: distribution on distributions

Dirichlet process mixture (DPM): a mixture model with infinitely many components where parameters of each component are drawn from a DP. Useful for **clustering**; assignments of points to clusters follows a CRP.

Bayesian nonparametrics for structured data

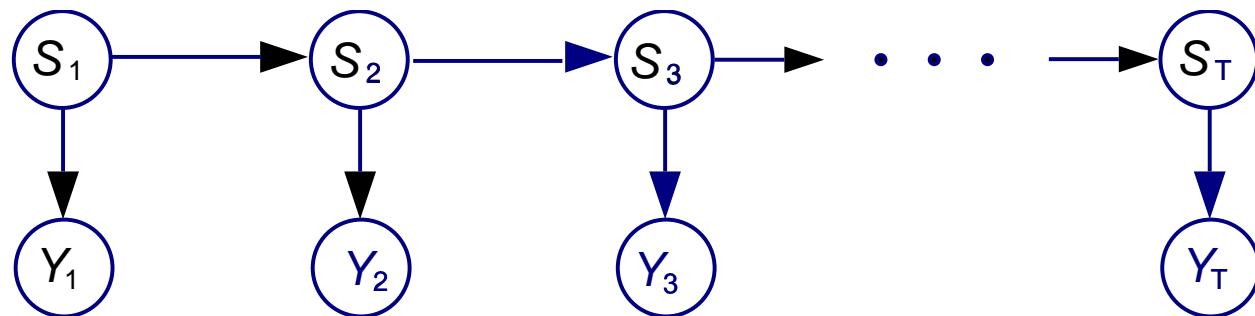
Bayesian nonparametrics applied to models of other structured objects:

- Clusters [and hierarchies]
- Time Series
- Sparse Matrices and Feature Allocation Models
- Networks

Times Series

Hidden Markov Models

Hidden Markov models (HMMs) are widely used sequence models for speech recognition, bioinformatics, biophysics, text modelling, video monitoring, etc.

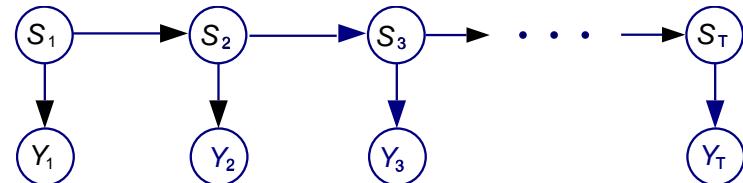


In an HMM, the sequence of observations y_1, \dots, y_T is modelled by assuming that it was generated by a sequence of discrete hidden states s_1, \dots, s_T with Markovian dynamics.

If the HMM has K states ($s_t \in \{1, \dots, K\}$) the transition matrix has $K \times K$ elements.

HMMs can be thought of as *time-dependent mixture models*.

Infinite hidden Markov models (iHMMs)

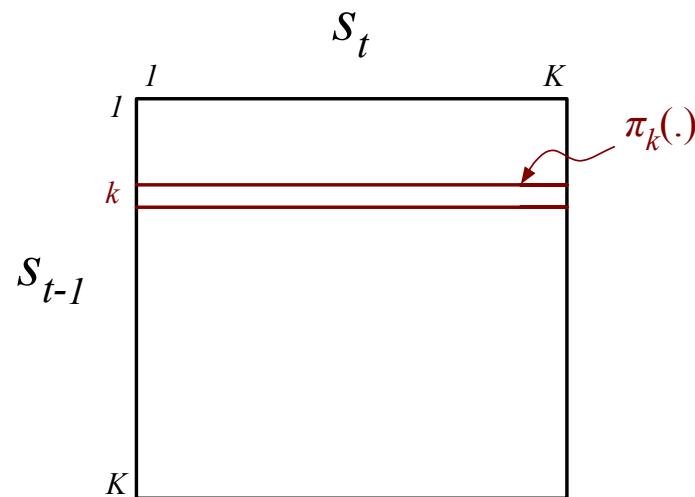


HMMs can be thought of as *time-dependent mixture models*.

In an HMM with K states, the transition matrix has $K \times K$ elements.

We want to let $K \rightarrow \infty$

Infinite HMMs can be derived from the HDP.



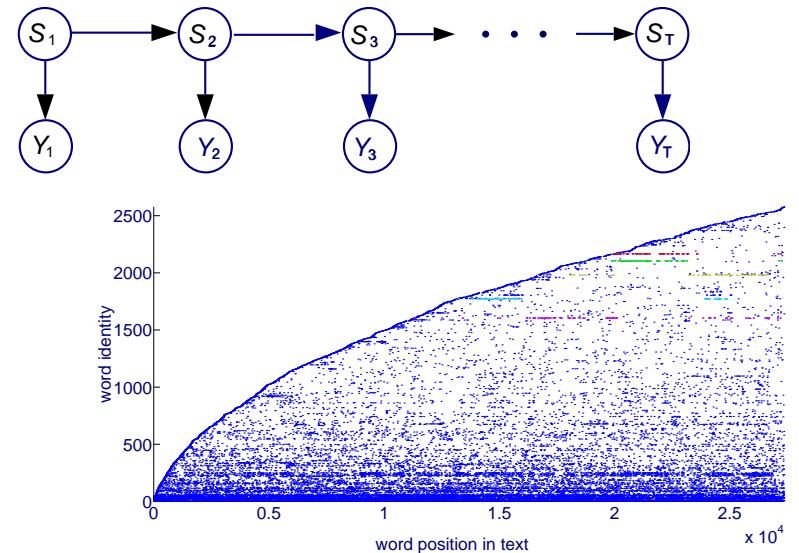
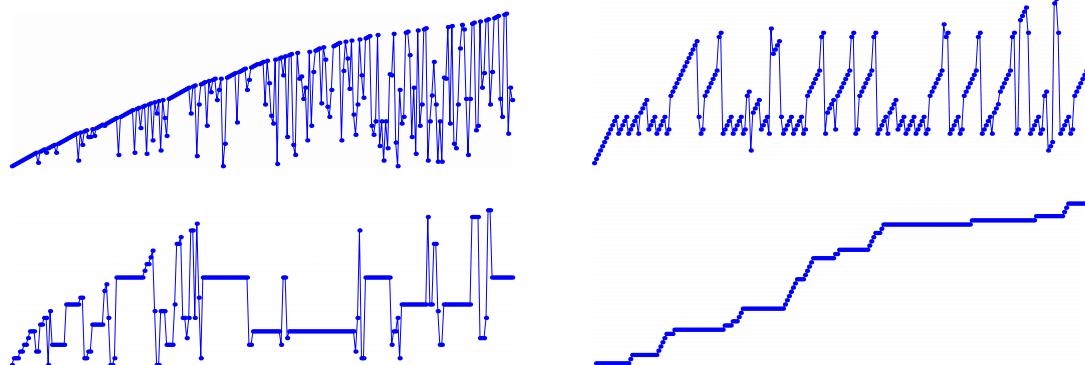
$\beta \gamma$	\sim	Stick($\cdot \gamma$)	(base distribution over states)
$\pi_k \alpha, \beta$	\sim	DP($\cdot \alpha, \beta$)	(transition parameters for state $k = 1, \dots$)
$\theta_k H$	\sim	$H(\cdot)$	(emission parameters for state $k = 1, \dots$)
$s_t s_{t-1}, (\pi_k)_{k=1}^{\infty}$	\sim	$\pi_{s_{t-1}}(\cdot)$	(transition)
$y_t s_t, (\theta_k)_{k=1}^{\infty}$	\sim	$p(\cdot \theta_{s_t})$	(emission)

(Beal, Ghahramani, and Rasmussen, 2002) (Teh et al. 2005)

Infinite hidden Markov models (iHMMs)

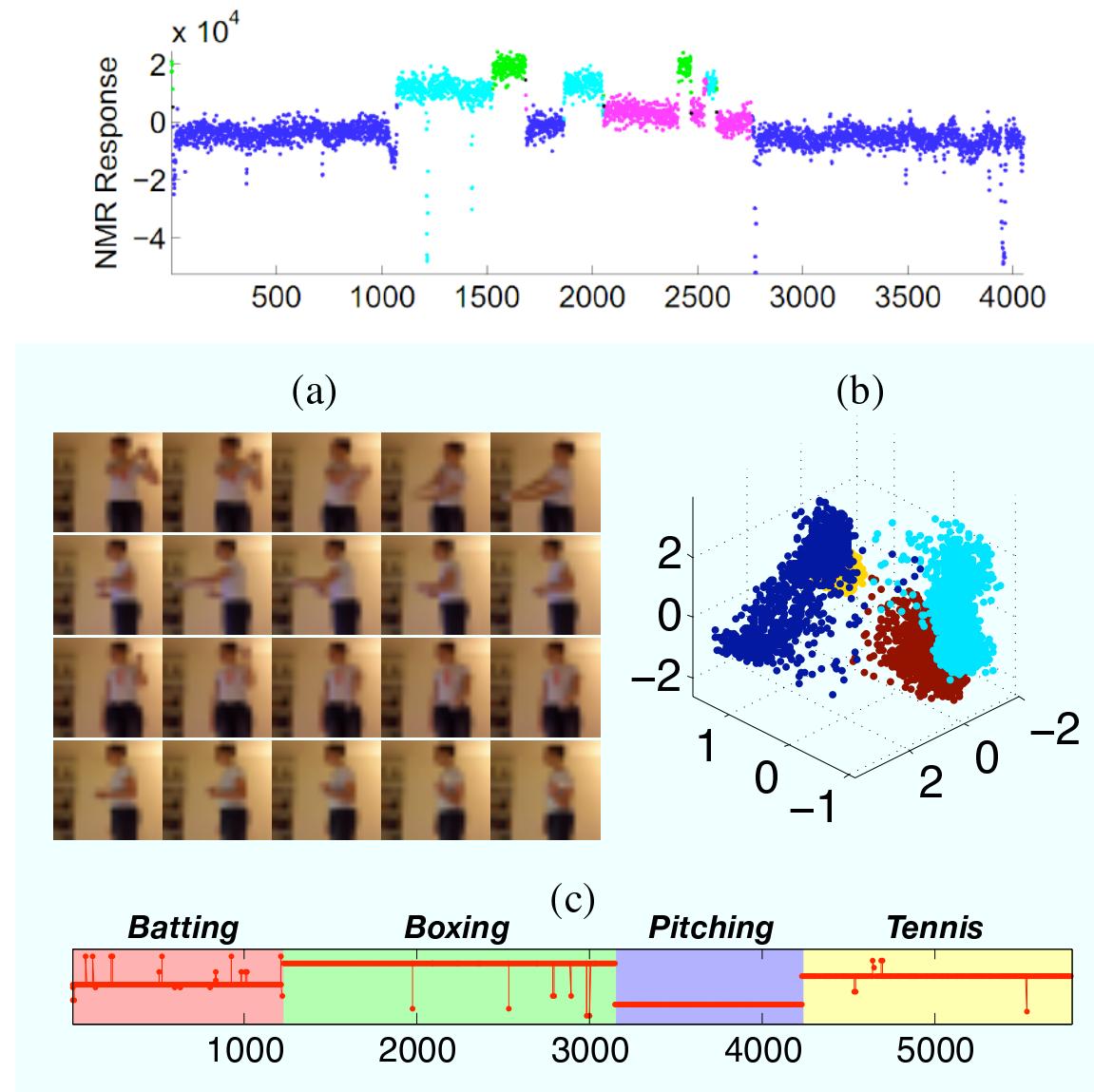
Let the number of hidden states $K \rightarrow \infty$.

Here are some typical state trajectories for an iHMM. Note that the number of states visited grows with T .



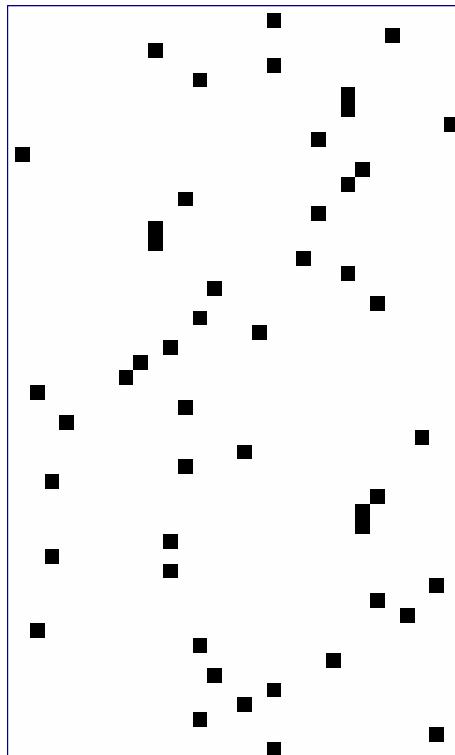
- Introduced in (Beal, Ghahramani and Rasmussen, 2002).
- Teh, Jordan, Beal and Blei (2005) showed that iHMMs can be derived from hierarchical Dirichlet processes, and provided a more efficient Gibbs sampler.
- We have recently derived a much more efficient sampler based on Dynamic Programming (Van Gael, Saatci, Teh, and Ghahramani, 2008). <http://mloss.org/software/view/205/>
- And we have parallel (.NET) and distributed (Hadoop) implementations (Bratières, Van Gael, Vlachos and Ghahramani, 2010).

Infinite HMM: Changepoint detection and video segmentation



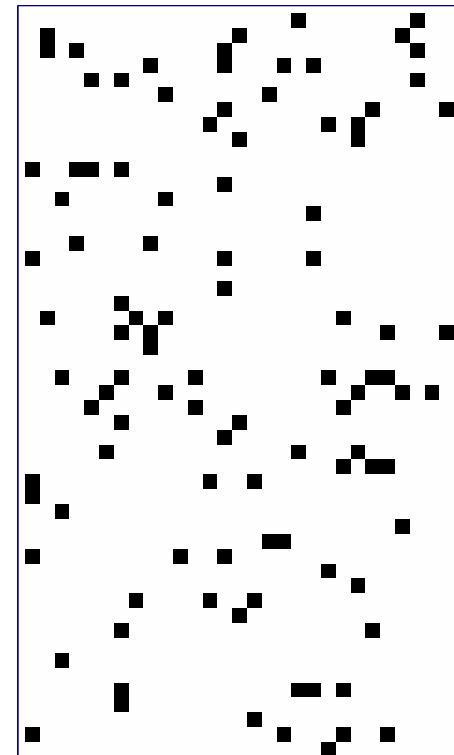
Sparse Matrices and Feature Models

Latent Cluster vs Feature Models



clustering

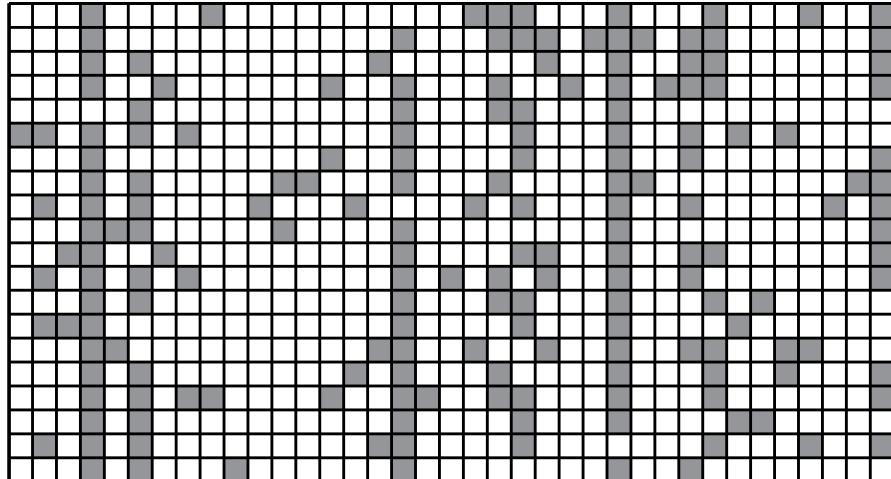
Chinese Restaurant Process (CRP)
Dirichlet process



feature allocation

Indian Buffet Process (IBP)
Beta process

Sparse binary matrices and latent feature models



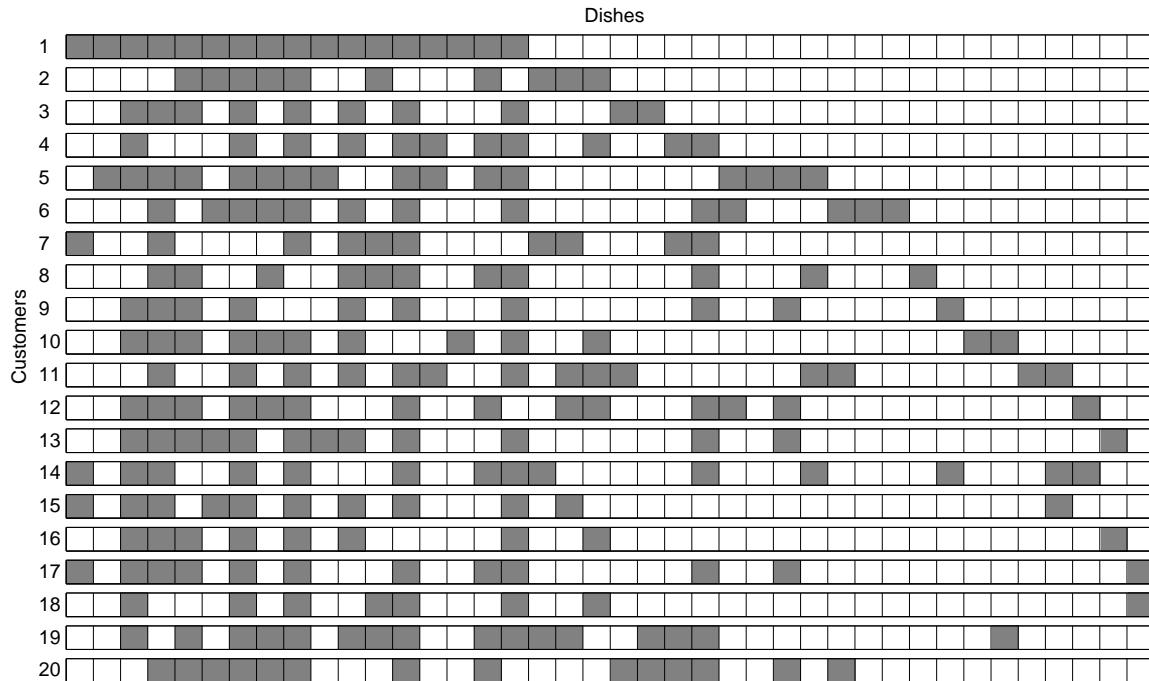
$z_{nk} = 1$ means object n has feature k :

$$z_{nk} \sim \text{Bernoulli}(\theta_k)$$

$$\theta_k \sim \text{Beta}(\alpha/K, 1)$$

- Note that $P(z_{nk} = 1|\alpha) = E(\theta_k) = \frac{\alpha/K}{\alpha/K+1}$, so as K grows larger the matrix gets **sparser**.
- So if \mathbf{Z} is $N \times K$, the expected number of nonzero entries is $N\alpha/(1+\alpha/K) < N\alpha$.
- Even in the $K \rightarrow \infty$ limit, the matrix is expected to have a finite number of non-zero entries.
- $K \rightarrow \infty$ results in an Indian buffet process (IBP)

Indian buffet process



“Many Indian restaurants offer lunchtime buffets with an apparently infinite number of dishes”



- First customer starts at the left of the buffet, and takes a serving from each dish, stopping after a Poisson(α) number of dishes as his plate becomes overburdened.
- The n^{th} customer moves along the buffet, sampling dishes in proportion to their popularity, serving himself dish k with probability m_k/n , and trying a Poisson(α/n) number of new dishes.
- The customer-dish matrix, Z , is a draw from the IBP.

(w/ Tom Griffiths 2006; 2011)

Properties of the Indian buffet process

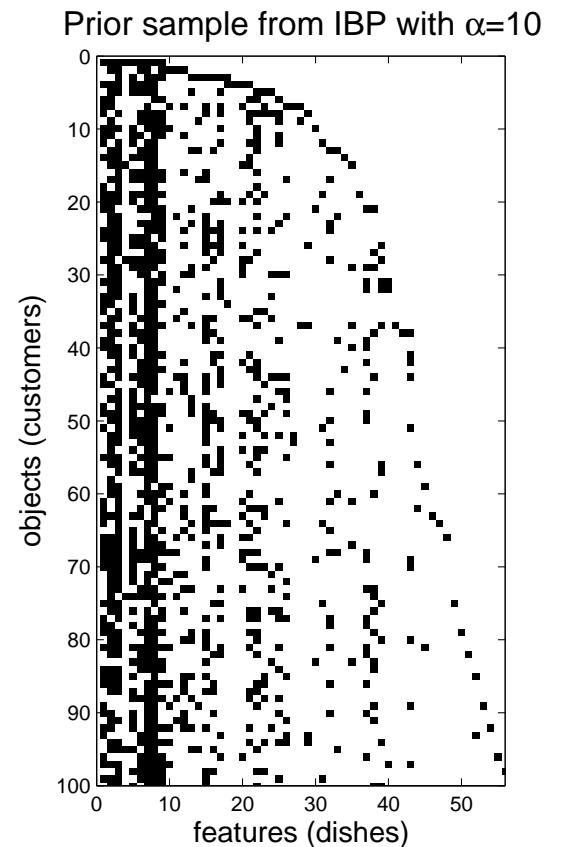
$$P([\mathbf{Z}]|\alpha) = \exp \left\{ -\alpha H_N \right\} \frac{\alpha^{K_+}}{\prod_{h>0} K_h!} \prod_{k \leq K_+} \frac{(N-m_k)!(m_k-1)!}{N!}$$

Shown in (Griffiths and Ghahramani 2006, 2011):

- It is infinitely exchangeable.
- The number of ones in each row is $\text{Poisson}(\alpha)$
- The expected total number of ones is αN .
- The number of nonzero columns grows as $O(\alpha \log N)$.

Additional properties:

- Has a stick-breaking representation (Teh, et al 2007)
- Has as its de Finetti mixing distribution the Beta process (Thibaux and Jordan 2007)
- More flexible two and three parameter versions exist (w/ Griffiths & Sollich 2007; Teh and Görür 2010)



Modelling Data with Indian Buffet Processes

Latent variable model: let \mathbf{X} be the $N \times D$ matrix of observed data, and \mathbf{Z} be the $N \times K$ matrix of sparse binary latent features

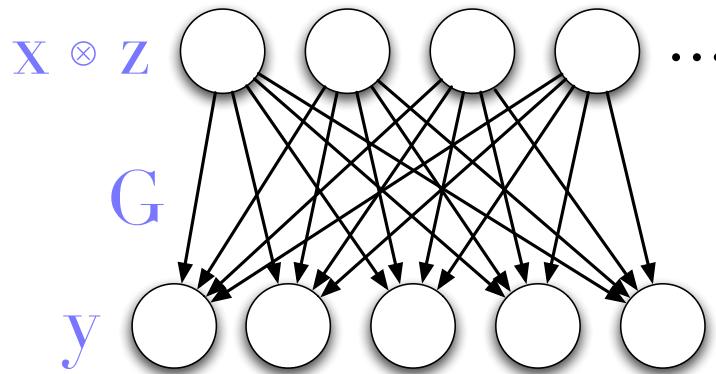
$$P(\mathbf{X}, \mathbf{Z} | \alpha) = P(\mathbf{X} | \mathbf{Z}) P(\mathbf{Z} | \alpha)$$

By combining the **IBP** with different likelihood functions we can get different kinds of models:

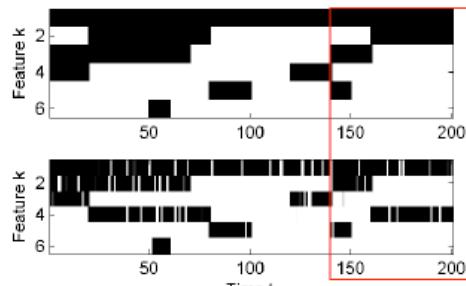
- Models for graph structures (w/ Wood, Griffiths, 2006; w/ Adams and Wallach, 2010)
- Models for protein complexes (w/ Chu, Wild, 2006)
- Models for choice behaviour (Görür & Rasmussen, 2006)
- Models for users in collaborative filtering (w/ Meeds, Roweis, Neal, 2007)
- Sparse latent trait, pPCA and ICA models (w/ Knowles, 2007, 2011)
- Models for overlapping clusters (w/ Heller, 2007)

Infinite Independent Components Analysis

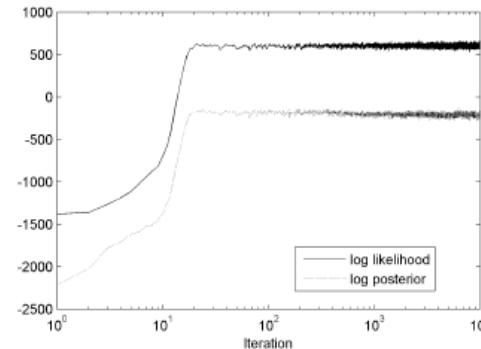
Model: $\mathbf{Y} = \mathbf{G}(\mathbf{Z} \otimes \mathbf{X}) + \mathbf{E}$



where \mathbf{Y} is the data matrix, \mathbf{G} is the mixing matrix $\mathbf{Z} \sim \text{IBP}(\alpha, \beta)$ is a mask matrix, \mathbf{X} is heavy tailed sources and \mathbf{E} is Gaussian noise.



(a) Top: True \mathbf{Z} . Bottom: Inferred \mathbf{Z} .
Red box denotes test data.



(b) Plot of the log likelihood and posterior for the duration of the iICA₂ run.

Fig. 1. True and inferred \mathbf{Z} and algorithm convergence.

(w/ David Knowles, 2007, 2011)

Nonparametric Binary Matrix Factorization

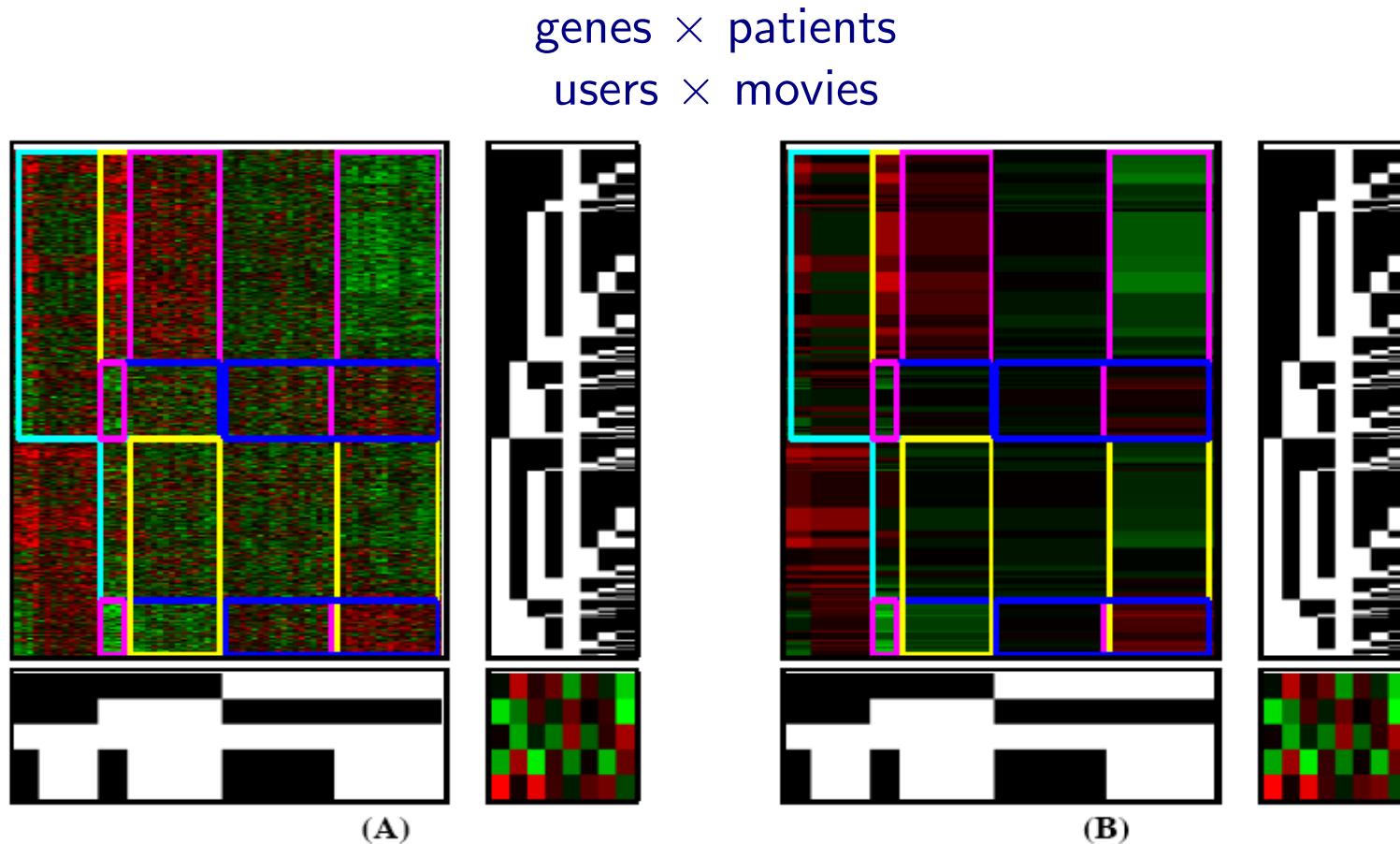
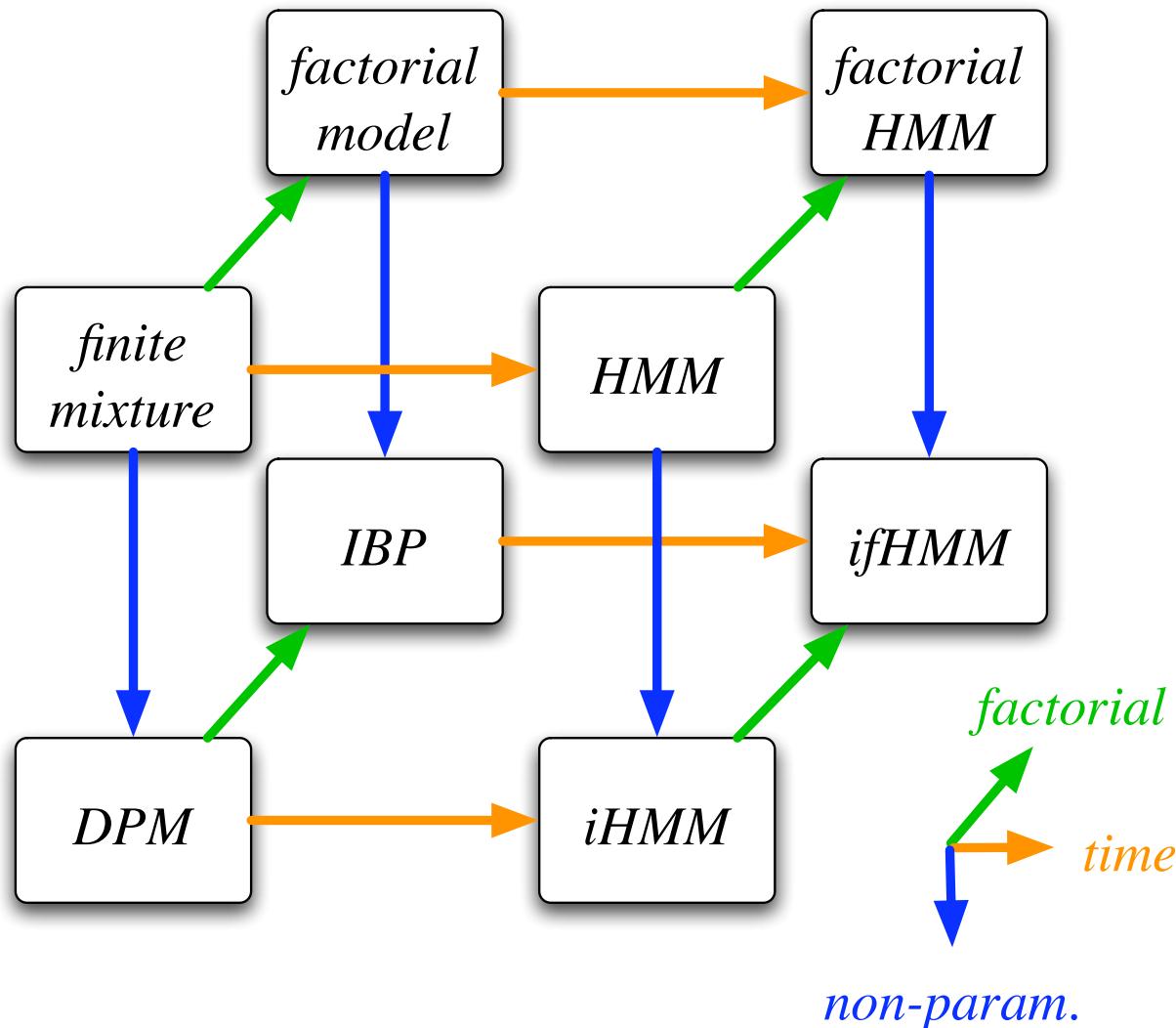


Figure 5: Gene expression results. (A) The top-left is \mathbf{X} sorted according to contiguous features in the final \mathbf{U} and \mathbf{V} in the Markov chain. The bottom-left is \mathbf{V}^\top and the top-right is \mathbf{U} . The bottom-right is \mathbf{W} . (B) The same as (A), but the expected value of \mathbf{X} , $\hat{\mathbf{X}} = \mathbf{U}\mathbf{W}\mathbf{V}^\top$. We have highlighted regions that have both u_{ik} and v_{jl} on. For clarity, we have only shown the (at most) two largest contiguous regions for each feature pair.

A Picture: Relations between some models



Posterior Inference in IBP Models

$$P(\mathbf{Z}, \alpha | \mathbf{X}) \propto P(\mathbf{X} | \mathbf{Z}) P(\mathbf{Z} | \alpha) P(\alpha)$$

Gibbs sampling: $P(z_{nk} = 1 | \mathbf{Z}_{-(nk)}, \mathbf{X}, \alpha) \propto P(z_{nk} = 1 | \mathbf{Z}_{-(nk)}, \alpha) P(\mathbf{X} | \mathbf{Z})$

- If $m_{-n,k} > 0$, $P(z_{nk} = 1 | \mathbf{z}_{-n,k}) = \frac{m_{-n,k}}{N}$
- For infinitely many k such that $m_{-n,k} = 0$: Metropolis steps with truncation* to sample from the number of new features for each object.
- If α has a Gamma prior then the posterior is also Gamma \rightarrow Gibbs sample.

Conjugate sampler: assumes that $P(\mathbf{X} | \mathbf{Z})$ can be computed.

Slice sampler: works for non-conjugate case, is not approximate, and has an *adaptive truncation level* using an IBP stick-breaking construction (Teh, et al 2007) see also (Adams et al 2010).

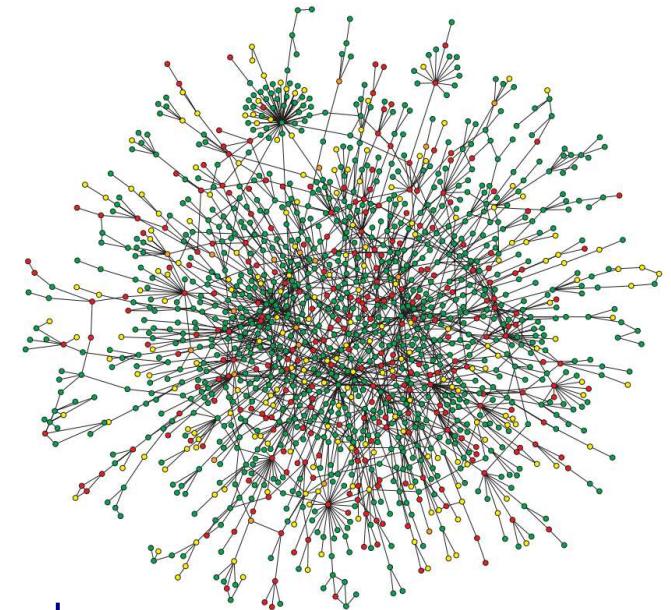
Deterministic Inference: variational inference (Doshi et al 2009a) parallel inference (Doshi et al 2009b), beam-search MAP (Rai and Daume 2011), power-EP (Ding et al 2010), submodular MAP (w/ Reed, 2013)

Some Case Studies

- Networks and relational data
- Scaling
- Discovering structure in Gaussian process kernels

Modelling Networks

We are interested in modelling networks.



Biological networks: protein-protein interaction networks

Social networks: friendship networks; co-authorship networks

We wish to have models that will be able to

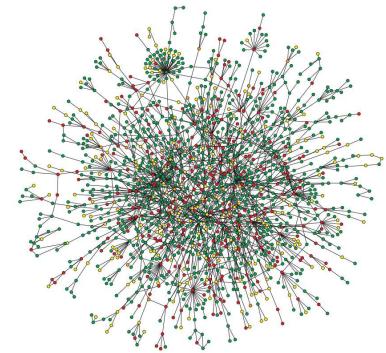
- predict missing links,
- infer latent properties or classes of the objects,
- generalise learned properties from smaller observed networks to larger networks.

Figure from Barabasi and Oltvai 2004: A protein-protein interaction network of budding yeast

Networks and Relational Data

Networks are just a way of representing certain kinds of *relational data*:

- friend(John,Mary)
- buy(Jack,iPhone)
- rate(Fred,Titanic,5)
- cite(PaperA,PaperB)
- author(PaperA,John)
- regulate(TranscriptionFactorA,GeneB) ...

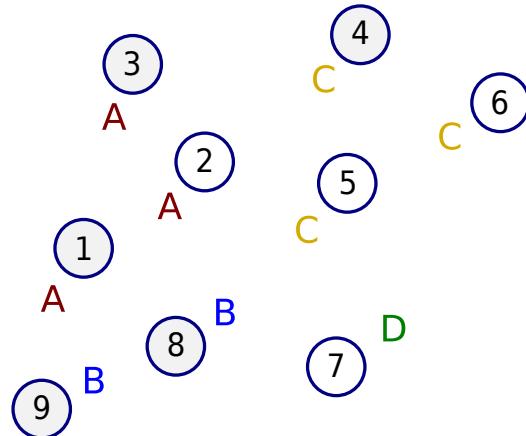


Relational data are ubiquitous; we need general models for such data.

There are deep and interesting connections between network modelling \leftrightarrow matrix factorization \leftrightarrow exchangeable arrays \leftrightarrow relational data.

Nonparametric Latent Class Models

Infinite Relational Model (Kemp et al 2006)



Each node v_i has a hidden class $c_i \in \{1, \dots, \infty\}$

For all i :

$$c_i | c_1, \dots, c_{i-1} \sim \text{CRP}(\alpha)$$

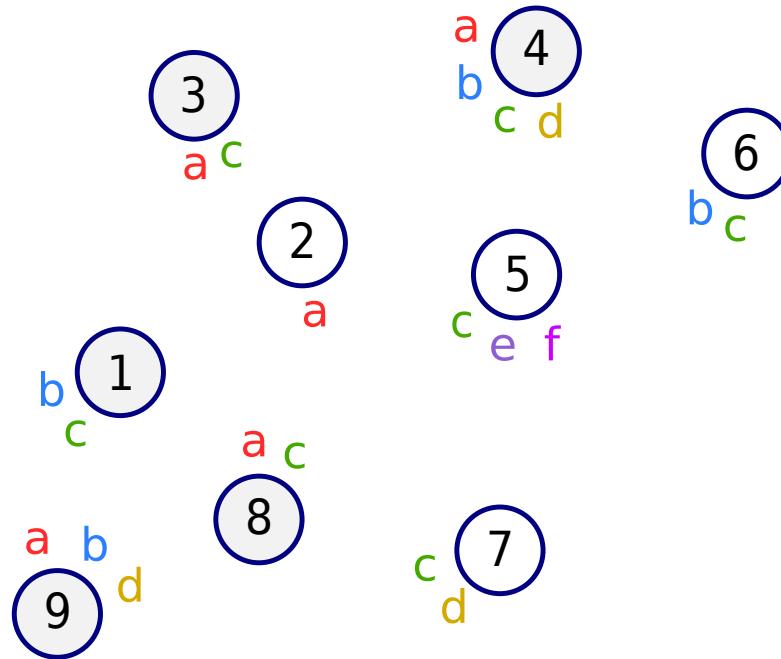
Probability of a link between two nodes v_i and v_j depends on their classes:

$$P(y_{ij} = 1 | c_i = k, c_j = \ell) = \rho_{k\ell}$$

Note that ρ is an infinitely large matrix, but if we give each element a beta prior we can integrate it out.

Inference done via MCMC. Fairly straightforward to implement.

Latent Feature Models



- Each node possesses some number of latent features.
- Alternatively we can think of this model as capturing *overlapping clusters or communities*
- The link probability depends on the latent features of the two nodes.
- The model should be able to accommodate a potentially unbounded (infinite) number of latent features (e.g. (Miller, Griffiths and Jordan 2010) use an IBP).

Exchangeable Sequences

Exchangeable sequence:

A sequence is exchangeable if its joint distribution is invariant under arbitrary permutation of the indices:

$$(X_1, X_2, \dots) \stackrel{d}{=} (X_{\pi(1)}, X_{\pi(2)}, \dots) \quad \forall \pi \in S_\infty.$$

de Finetti's Theorem:

$(X_i)_{i \in N}$ is exchangeable if and only if there exists a random probability measure Θ on X such that $X_1, X_2, \dots | \Theta \sim \text{iid } \Theta$

Interpretation:

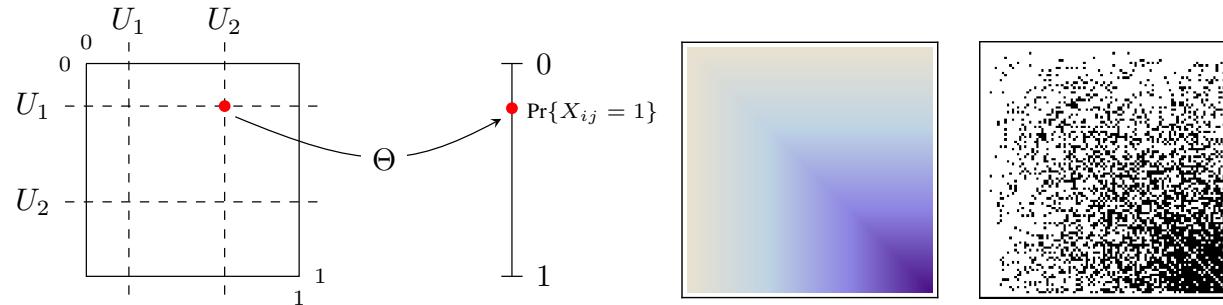
Any probabilistic model of data which assumes that the order of the data does not matter, can be expressed as a Bayesian mixture of iid models. Note that Θ may in general need to be infinite dimensional (i.e. *nonparameteric*).

Exchangeable Arrays

Exchangeable arrays: An array $X = (X_{ij})_{i,j \in \mathbb{N}}$ is called an exchangeable array if $(X_{ij}) \stackrel{d}{=} (X_{\pi(i)\pi(j)})$ for every $\pi \in S_\infty$.

Aldous-Hoover Theorem:

A random matrix (X_{ij}) is exchangeable if and only if there is a random (measurable) function $F : [0, 1]^3 \rightarrow X$ such that $(X_{ij}) \stackrel{d}{=} (F(U_i, U_j, U_{ij}))$ for every collection $(U_i)_{i \in \mathbb{N}}$ and $(U_{ij})_{i \leq j \in \mathbb{N}}$ of i.i.d. Uniform[0, 1] random variables, where $U_{ji} = U_{ij}$ for $j < i \in \mathbb{N}$.

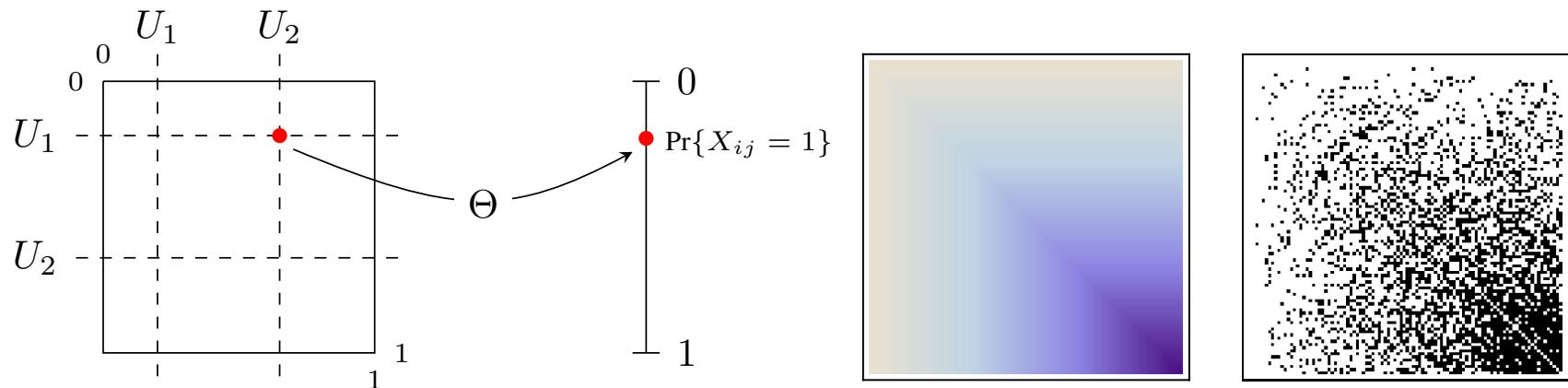


Interpretation:

Any model of matrices, arrays (or graphs) where the order of rows and columns (nodes) is irrelevant can be expressed by assuming *latent variables* associated with each row and column, and a *random function* mapping these latent variables to the observations.

Random Function Model

We develop a nonparametric probabilistic model for simple arrays and graphs that makes explicit the Aldous Hoover representation:



$$\Theta \sim \text{GP}(0, \kappa)$$

$$U_1, U_2, \dots \stackrel{\text{iid}}{\sim} \text{Uniform}[0, 1]$$

$$W_{ij} = \Theta(U_i, U_j)$$

$$X_{ij} \sim P[\cdot | W_{ij}]$$

(w/ James Lloyd, Dan Roy, Peter Orbanz, NIPS 2012)

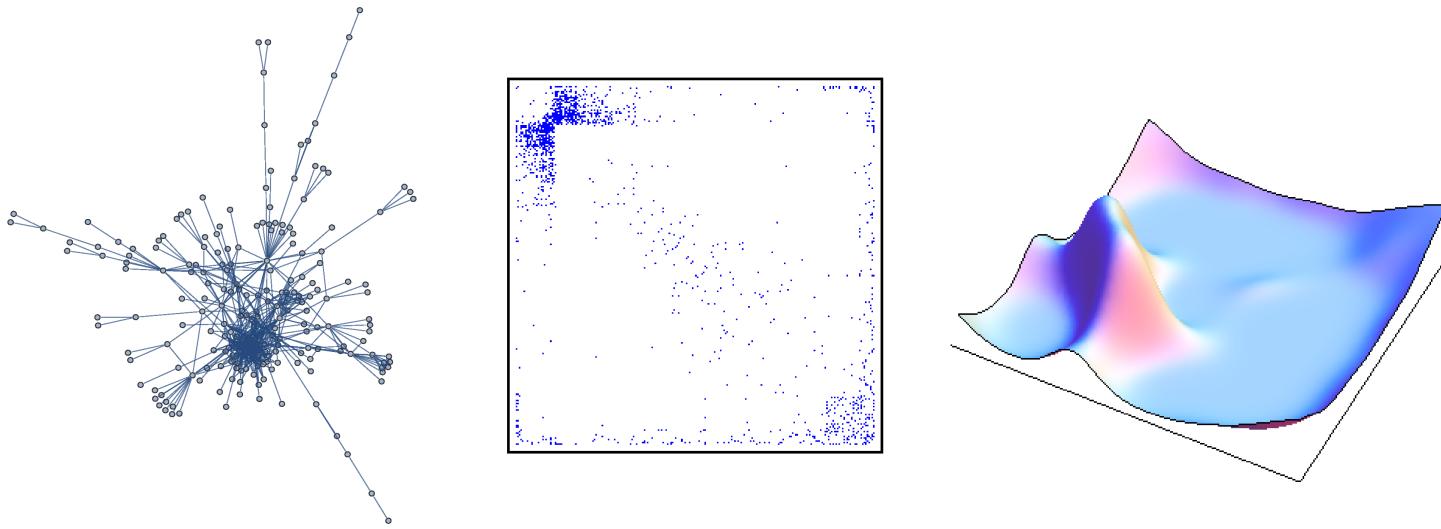


Random Function Model

The random function model can be related to a number of existing models for matrices, arrays/tensors, and graphs.

Graph data	
Random function model	$\Theta \sim \mathcal{GP}(0, \kappa)$
Latent class	$W_{ij} = m_{U_i U_j}$ where $U_i \in \{1, \dots, K\}$
IRM	$W_{ij} = m_{U_i U_j}$ where $U_i \in \{1, \dots, \infty\}$
Latent distance	$W_{ij} = - U_i - U_j $
Eigenmodel	$W_{ij} = U_i' \Lambda U_j$
LFRM	$W_{ij} = U_i' \Lambda U_j$ where $U_i \in \{0, 1\}^\infty$
ILA	$W_{ij} = \sum_d \mathbb{I}_{U_{id}} \mathbb{I}_{U_{jd}} \Lambda_{U_{id} U_{jd}}^{(d)}$ where $U_i \in \{0, \dots, \infty\}^\infty$
SMGB	$\Theta \sim \mathcal{GP}(0, \kappa_1 \otimes \kappa_2)$
Real-valued array data	
Random function model	$\Theta \sim \mathcal{GP}(0, \kappa)$
Mondrian process based	$\Theta = \text{piece-wise constant random function}$
PMF	$W_{ij} = U_i' V_j$
GPLVM	$\Theta \sim \mathcal{GP}(0, \kappa \otimes \delta)$

Random Function Model: Results



AUC results

Data set Latent dimensions	High school			NIPS			Protein		
	1	2	3	1	2	3	1	2	3
PMF	0.747	0.792	0.792	0.729	0.789	0.820	0.787	0.810	0.841
Eigenmodel	0.742	0.806	0.806	0.789	0.818	0.845	0.805	0.866	0.882
GPLVM	0.744	0.775	0.782	0.888	0.876	0.883	0.877	0.883	0.873
RFM	0.815	0.827	0.820	0.907	0.914	0.919	0.903	0.910	0.912

Scalable approximate inference

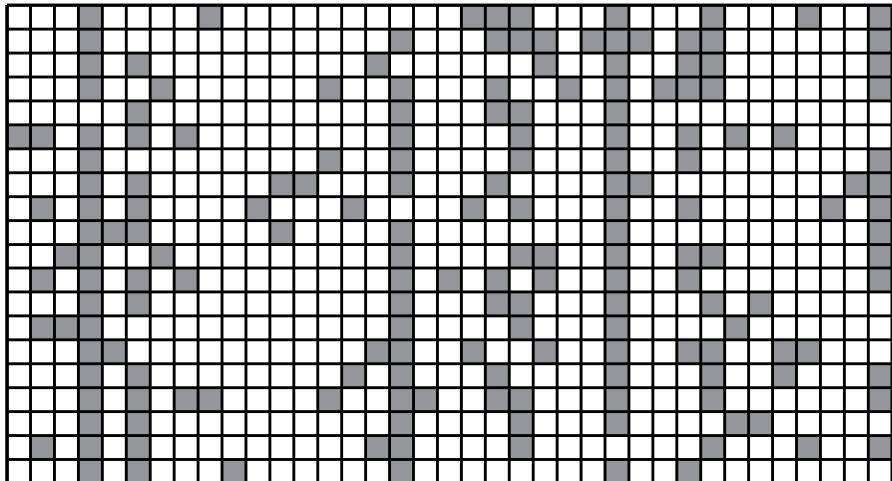
Scalable approximate inference

$$\begin{aligned} P(x|\mathcal{D}, m) &= \int P(x|\theta, \mathcal{D}, m)P(\theta|\mathcal{D}, m)d\theta \\ P(\mathcal{D}|m) &= \int P(\mathcal{D}|\theta, m)P(\theta|m)d\theta \end{aligned}$$

How do we compute these integrals in practice?

- Laplace Approximation
- Bayesian Information Criterion (BIC)
- Variational Bayesian approximations
- Expectation Propagation (and loopy belief propagation)
- Markov chain Monte Carlo
- Sequential Monte Carlo
- ...

Review of IBPs: Sparse binary matrices and feature allocation models



$z_{nk} = 1$ means object n has feature k :

$$z_{nk} \sim \text{Bernoulli}(\theta_k)$$

$$\theta_k \sim \text{Beta}(\alpha/K, 1)$$

- Note that $P(z_{nk} = 1|\alpha) = E(\theta_k) = \frac{\alpha/K}{\alpha/K+1}$, so as K grows larger the matrix gets **sparser**.
- So if \mathbf{Z} is $N \times K$, the expected number of nonzero entries is $N\alpha/(1+\alpha/K) < N\alpha$.
- Even in the $K \rightarrow \infty$ limit, the matrix is expected to have a finite number of non-zero entries.
- $K \rightarrow \infty$ results in an Indian buffet process (IBP)

Submodular MAP inference in IBPs

Approach: maximise over Z while maintaining a variational approximation to the integral over the other model parameters

Yields a *submodular maximisation* algorithm that can be maximized with a simple greedy algorithm with a $(1/3)$ -approximability guarantee

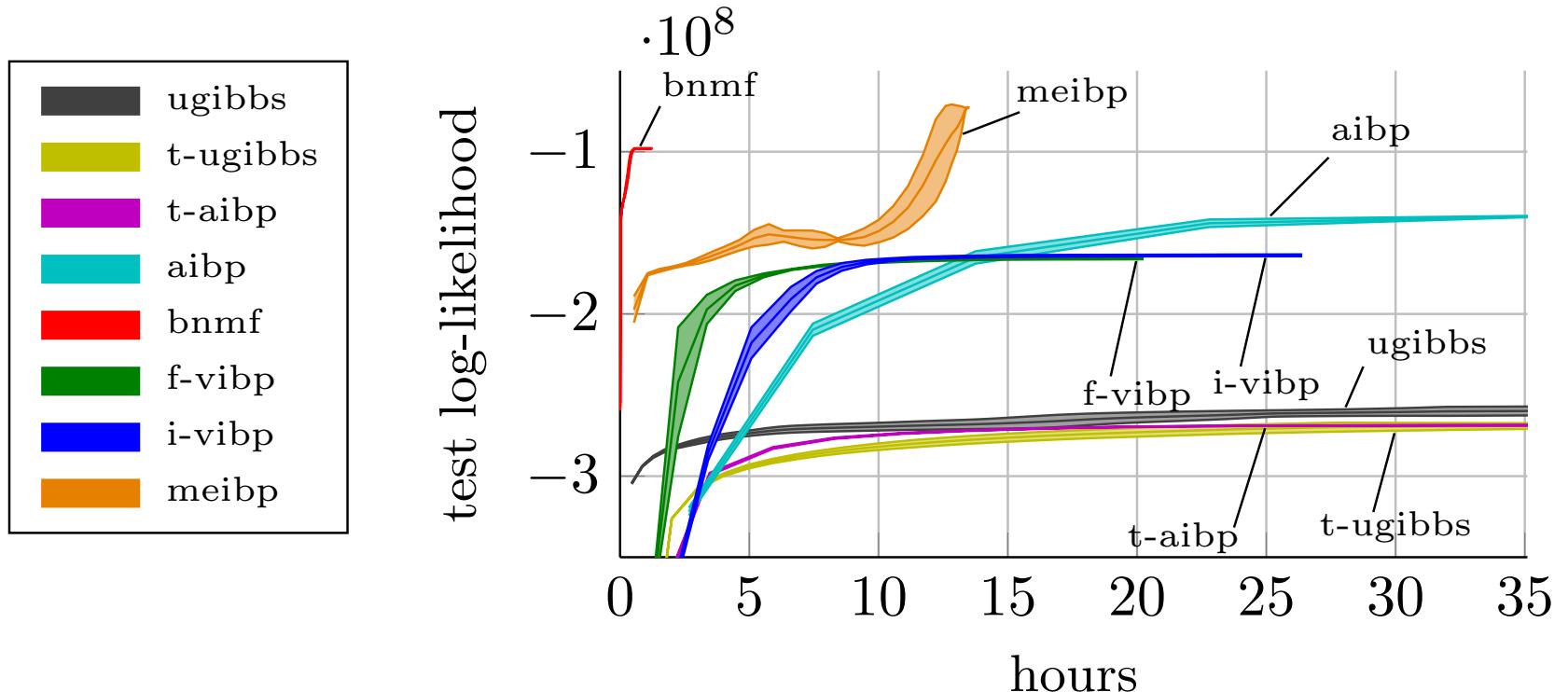
Submodularity: (diminishing returns) for $A \subseteq B$ and f a set function:

$$f(A \cup \{e\}) - f(A) \geq f(B \cup \{e\}) - f(B)$$

(w/ Colorado Reed, ICML 2013)



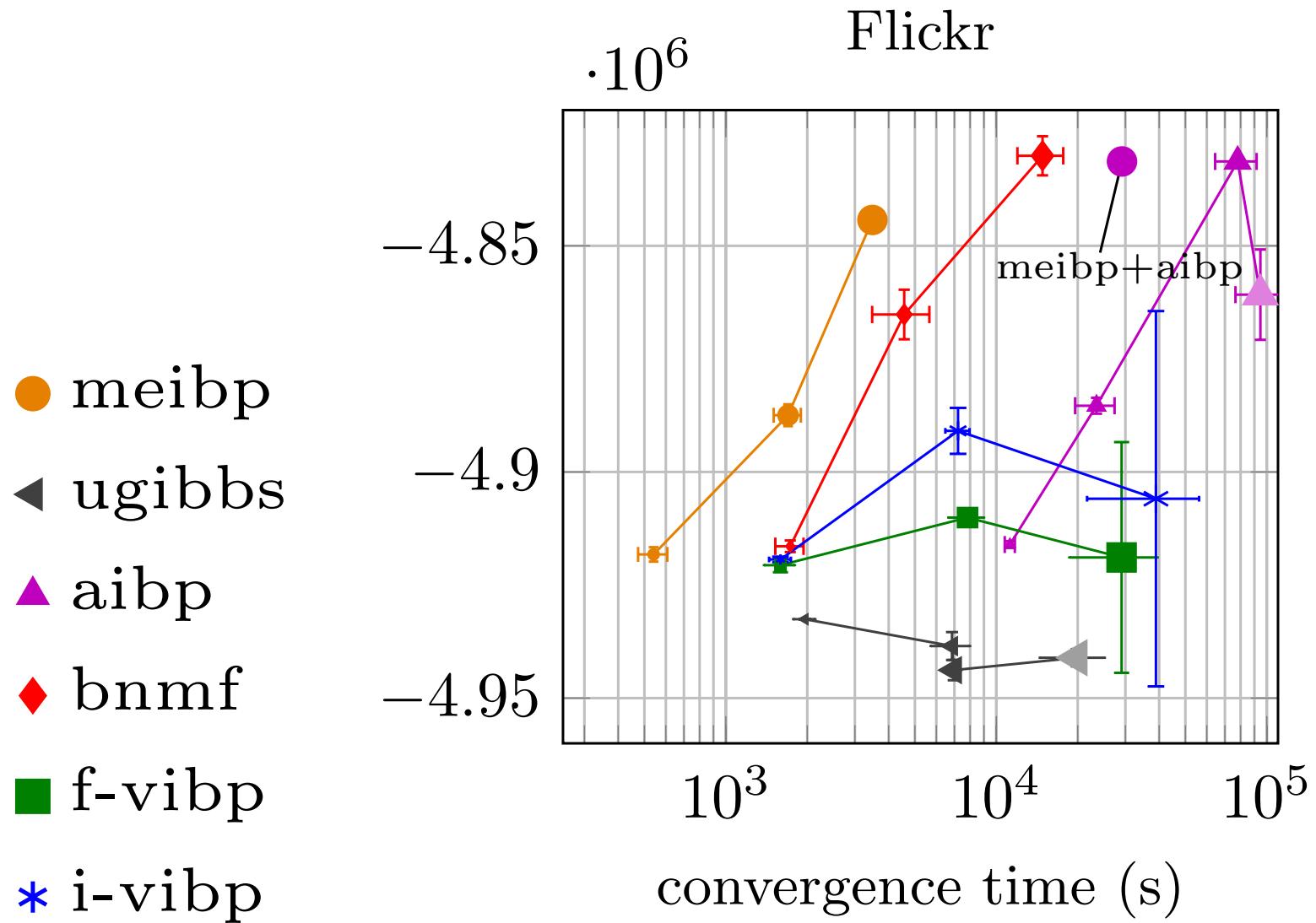
Submodular MAP inference in IBPs: Results



Synthetic datasets: $N = 10^5$ data points, $D = 10^3$ dimensions, $K_{\max} = 50$ features.

- meibp:** our new method based on submodular maximisation (w/ Reed, 2013)
- aibp:** accelerated IBP (w/ Doshi-Velez, 2009)
- ugibbs:** uncollapsed Gibbs sampling (w/ Doshi-Velez, 2009)
- vibp:** variational IBP (Doshi-Velez et al, 2009)
- bnmf:** iterated conditional modes for NMF (Schmidt et al 2009)
- bs-ipb:** beam-search (didn't complete one iteration): (Rai and Daume, 2011)
- inmf:** power-EP for non-negative IBP (didn't complete one iteration) (Ding et al, 2010)

Submodular MAP inference in IBPs: Results



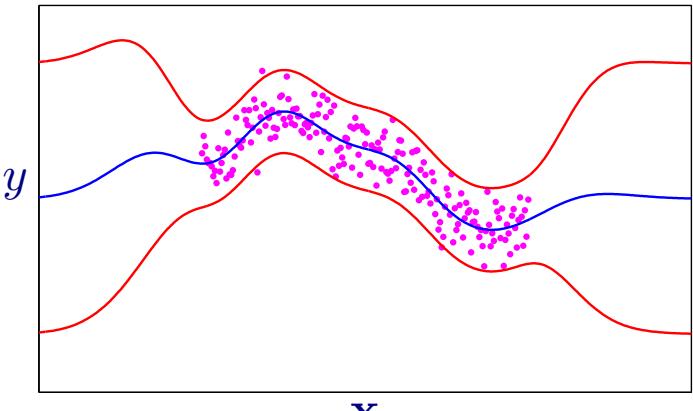
binary image-tag indicators from Flickr - dataset of 25000 images by 1500 indicators.

Structure Discovery for Gaussian Process Kernels

Nonlinear regression and Gaussian processes

Consider the problem of **nonlinear regression**:

You want to learn a function f with **error bars** from data $\mathcal{D} = \{\mathbf{X}, \mathbf{y}\}$



A **Gaussian process** defines a distribution on functions $p(f)$ which can be used for Bayesian regression:

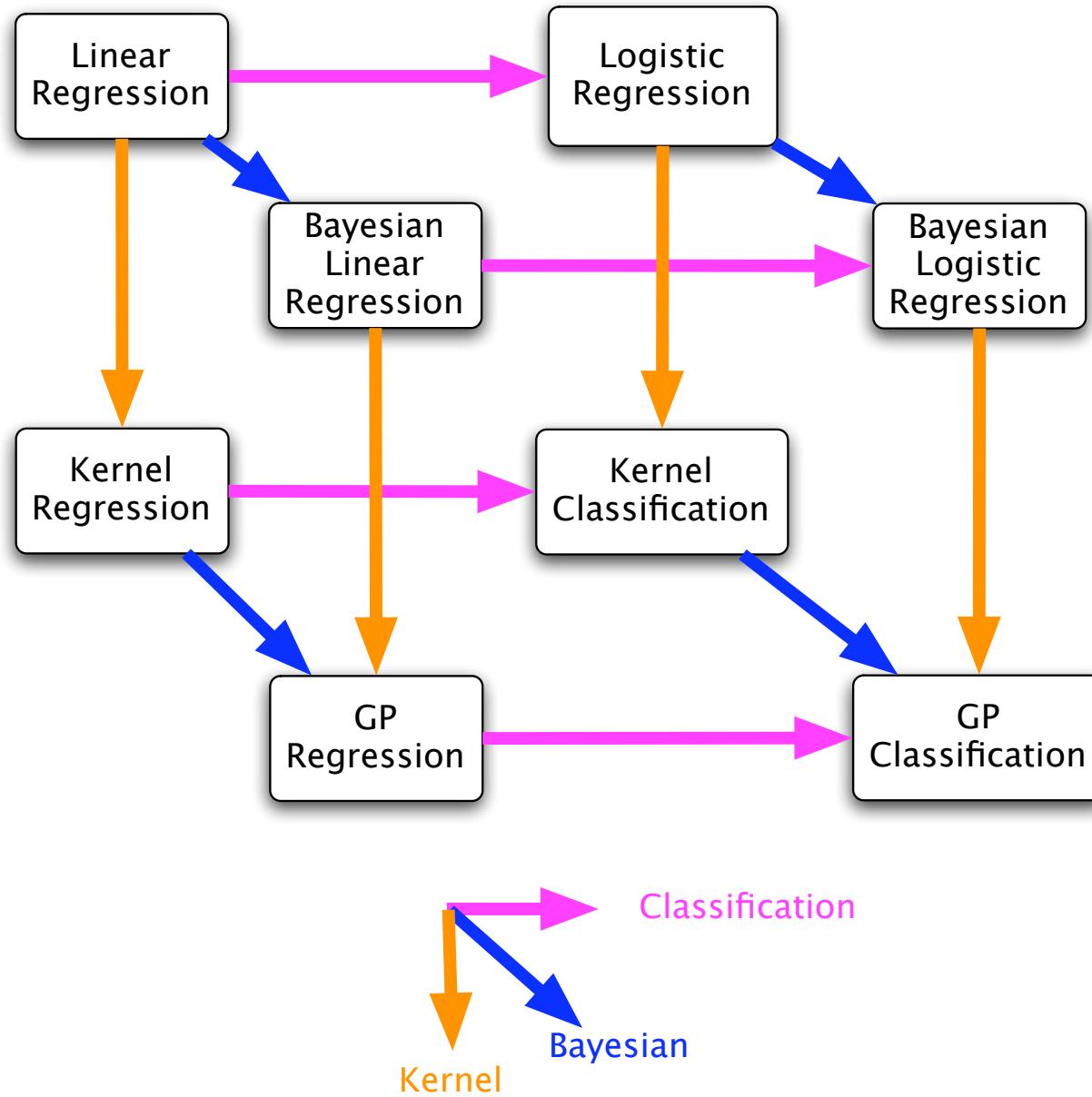
$$p(f|\mathcal{D}) = \frac{p(f)p(\mathcal{D}|f)}{p(\mathcal{D})}$$

Let $\mathbf{f} = (f(x_1), f(x_2), \dots, f(x_n))$ be an n -dimensional vector of function values evaluated at n points $x_i \in \mathcal{X}$. Note, \mathbf{f} is a random variable.

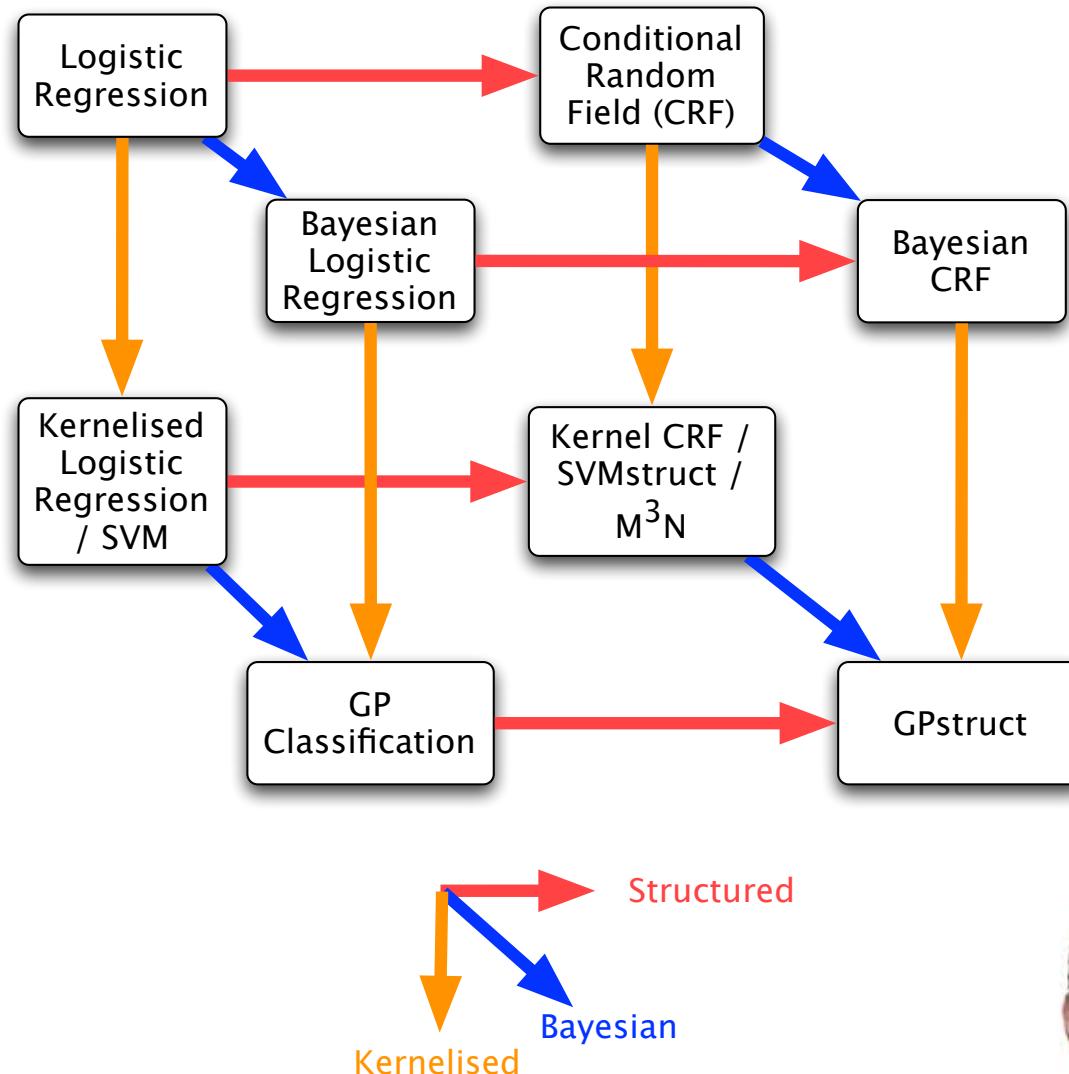
Definition: $p(f)$ is a **Gaussian process** if for any finite subset $\{x_1, \dots, x_n\} \subset \mathcal{X}$, the marginal distribution over that subset $p(\mathbf{f})$ is multivariate Gaussian.

$$f \sim \text{GP}(\mu, K)$$

Bayesian kernelised regression and classification



Bayesian kernelised structured prediction



(w/ Bratieres and Quadrianto, arXiv 2013)



How do we learn the kernel (covariance function)?

$$f \sim \text{GP}(\mu, K)$$

- Usual approach: parametrise the kernel with a few hyperparameters and optimise or infer these. An example covariance function:

$$K(x_i, x_j) = v_0 \exp \left\{ - \left(\frac{|x_i - x_j|}{r} \right)^\alpha \right\} + v_1 + v_2 \delta_{ij}$$

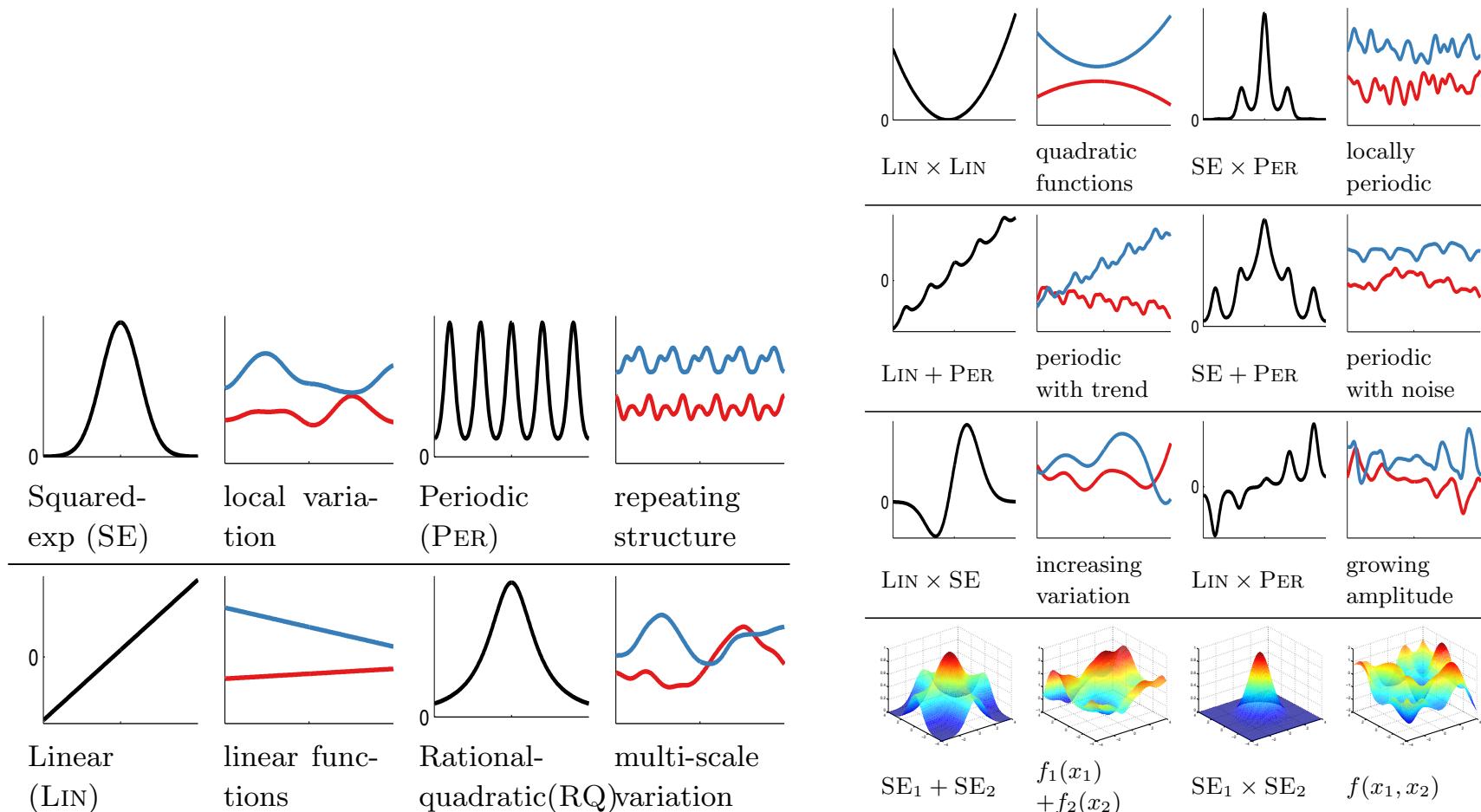
with parameters $(v_0, v_1, v_2, r, \alpha)$. These kernel parameters are **interpretable** and can be learned from data:

v_0	signal variance
v_1	variance of bias
v_2	noise variance
r	lengthscale
α	roughness

- Structure discovery for the kernel by searching over a grammar of kernels

Kernel Composition

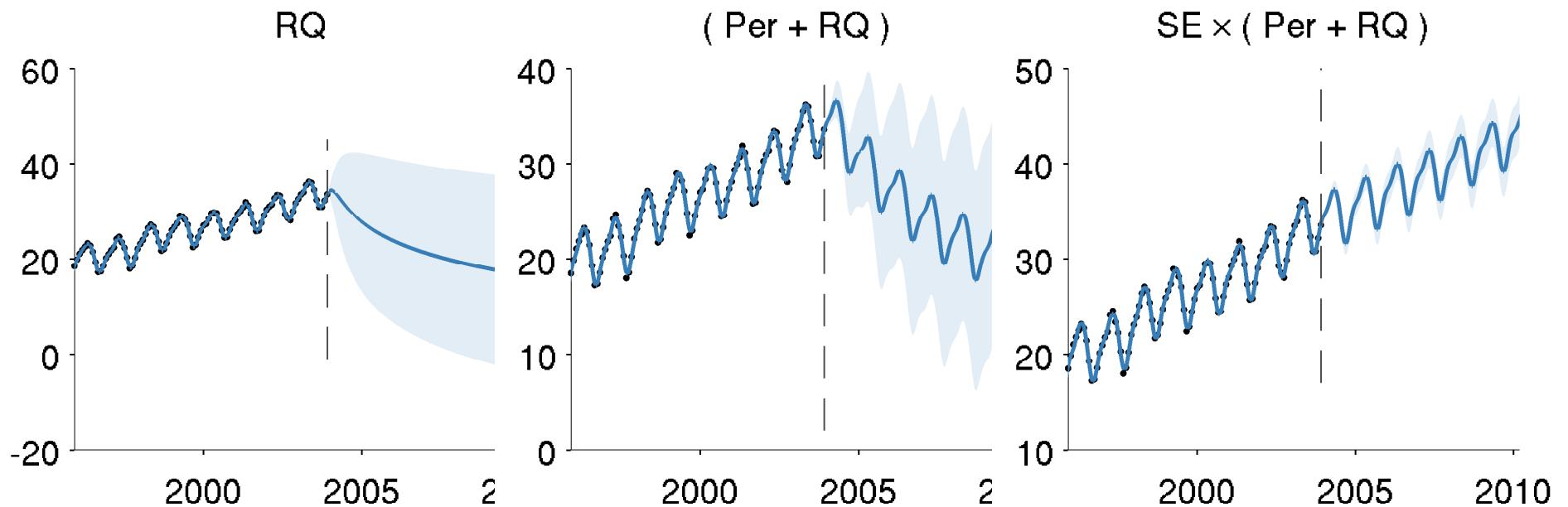
By taking a few simple **base kernels** and two **composition rules**, kernel addition and multiplication, we can span a rich space of structured kernels.



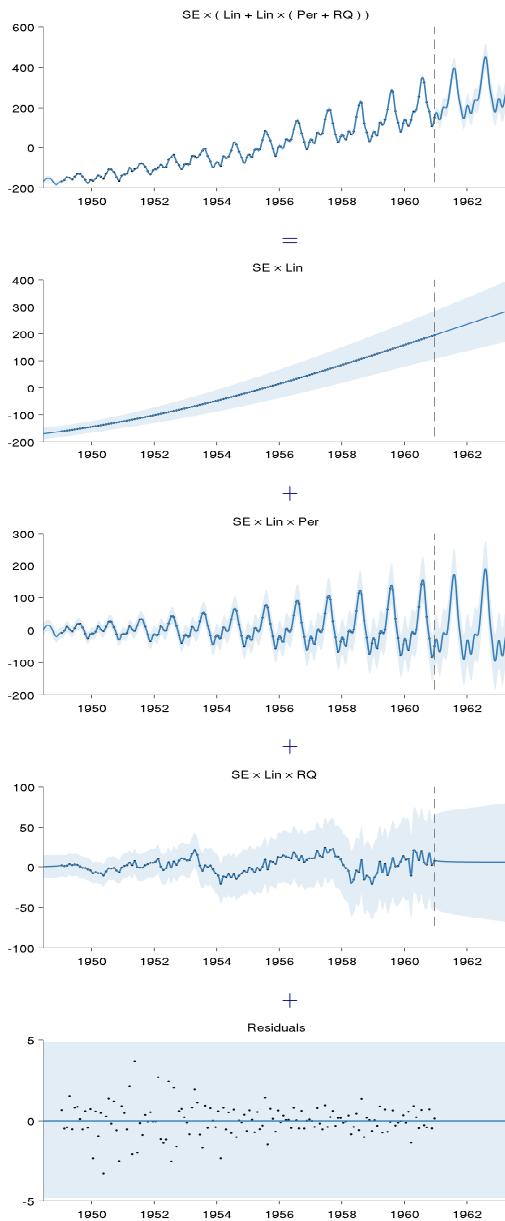
(w/ Duvenaud, Lloyd, Grosse, and Tenenbaum, ICML 2013)

(Wilson and Adams, ICML 2013)

Kernel Composition: Mauna Loa CO₂ Keeling Curve



Kernel Composition: Airline passenger prediction



Kernel Composition: results

Method	bach	Mean Squared Error (MSE)					bach	Negative Log-Likelihood				
		concrete	puma	servo	housing	concrete		puma	servo	housing		
Linear Regression	1.031	0.404	0.641	0.523	0.289	2.430	1.403	1.881	1.678	1.052		
GAM	1.259	0.149	0.598	0.281	0.161	1.708	0.467	1.195	0.800	0.457		
HKL	0.199	0.147	0.346	0.199	0.151	-	-	-	-	-		
GP SE-ARD	0.045	0.157	0.317	0.126	0.092	-0.131	0.398	0.843	0.429	0.207		
GP Additive	0.045	0.089	0.316	0.110	0.102	-0.131	0.114	0.841	0.309	0.194		
Structure Search	0.044	0.087	0.315	0.102	0.082	-0.141	0.065	0.840	0.265	0.059		

Summary of kernel structure discovery

Structure learning on kernels can be useful both for good automated prediction, and for interpretable results.

...towards and **automated statistician**.

Summary

- Bayesian machine learning is simply the application of probability theory to learning from data.
- Nonparametrics is needed to capture complexity of real data.
- **New models for complex structured data:**
 - Random Function Model for exchangeable arrays and relations
- **Fast scalable inference:**
 - submodular maximisation in IBPs
- **Automatic structure discovery:**
 - kernel structure discovery in Gaussian processes

Thanks to



Konstantina Palla



David Knowles



Colorado Reed



Andrew Wilson



James Lloyd



Peter Orbanz



Dan Roy



David Duvenaud

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postdocs available! to be advertised on ml-news@googlegroups.com

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Appendix

Dirichlet Process: Conjugacy

$$G \sim \text{DP}(\cdot | G_0, \alpha)$$

If the prior on G is a DP:

$$P(G) = \text{DP}(G | G_0, \alpha)$$

...and you observe θ ...

$$P(\theta | G) = G(\theta)$$

...then the posterior is also a DP:

$$P(G | \theta) = \frac{P(\theta | G)P(G)}{P(\theta)} = \text{DP} \left(\frac{\alpha}{\alpha+1} G_0 + \frac{1}{\alpha+1} \delta_\theta, \alpha+1 \right)$$

Generalization for n observations:

$$P(G | \theta_1, \dots, \theta_n) = \text{DP} \left(\frac{\alpha}{\alpha+n} G_0 + \frac{1}{\alpha+n} \sum_{i=1}^n \delta_{\theta_i}, \alpha+n \right)$$

Analogous to Dirichlet being conjugate to multinomial observations.

Dirichlet Process

Blackwell and MacQueen's (1973) urn representation

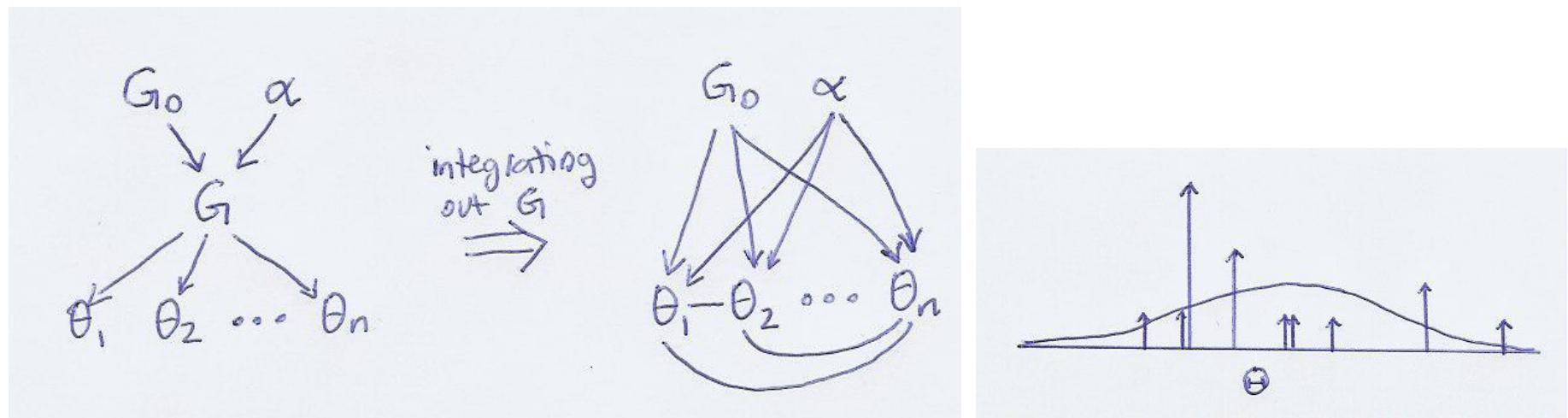
$$G \sim \text{DP}(\cdot | G_0, \alpha) \quad \text{and} \quad \theta | G \sim G(\cdot)$$

Then

$$\theta_n | \theta_1, \dots, \theta_{n-1}, G_0, \alpha \sim \frac{\alpha}{n-1+\alpha} G_0(\cdot) + \frac{1}{n-1+\alpha} \sum_{j=1}^{n-1} \delta_{\theta_j}(\cdot)$$

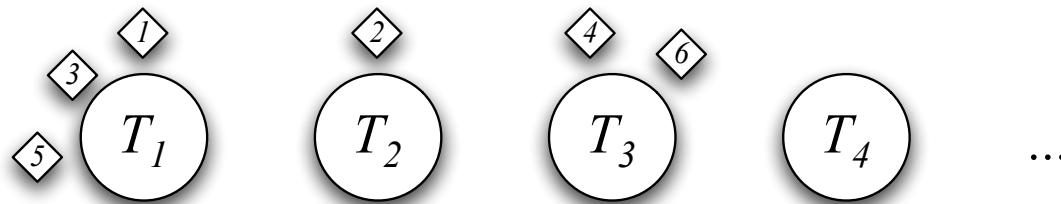
$$P(\theta_n | \theta_1, \dots, \theta_{n-1}, G_0, \alpha) \propto \int dG \prod_{j=1}^n P(\theta_j | G) P(G | G_0, \alpha)$$

The model exhibits a “clustering effect”.



Chinese Restaurant Process

The CRP generates samples from the distribution on partitions induced by a DPM.



Generating from a CRP:

customer 1 enters the restaurant and sits at table 1.

$$K = 1, n = 1, n_1 = 1$$

for $n = 2, \dots,$

customer n sits at table $\begin{cases} k & \text{with prob } \frac{n_k}{n-1+\alpha} \text{ for } k = 1 \dots K \\ K+1 & \text{with prob } \frac{\alpha}{n-1+\alpha} \end{cases}$ (new table)

if new table was chosen **then** $K \leftarrow K + 1$ **endif**

endfor

“Rich get richer” property.

(Aldous 1985; Pitman 2002)