### Intro to generalized additive models in R

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Material available at:

https://github.com/mfasiolo/workshop\_WARSAW19

#### These slides cover:

- What is an additive model?
- Introducing smooth effects
- 3 Introducing random effects
- 4 Diagnostics and model selection tools
- GAM modelling in mgcv

### Regression setting:

- y is our response or dependent variable
- x is a vector of covariates or independent variables

In **distributional regression** we want a good model for Dist(y|x).

Model is  $\mathrm{Dist}_m\{y|\theta_1(\mathbf{x}),\theta_2,\ldots,\theta_q\}$ , where  $\theta_1,\ldots,\theta_q$  are param.

We assume that  $\theta_2, \dots, \theta_q$  do not depend on  $\mathbf{x}$ .

### Gaussian additive model:

$$y|\mathbf{x} \sim N\{y|\mu(\mathbf{x}), \sigma^2\},$$

where

$$\mu(\mathbf{x}) = \mathbb{E}(y|\mathbf{x}) = \sum_{j=1}^{m} f_j(\mathbf{x}),$$

and

$$\sigma^2 = \mathsf{Var}(y).$$

 $f_i$ 's can be fixed, random or smooth effects.

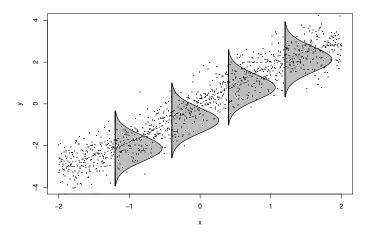


Figure: Gaussian model with variable mean. In mgcv: gam(y~s(x), family=gaussian).

Generalized additive model (GAM) (Hastie and Tibshirani, 1990):

$$y|\mathbf{x} \sim \mathsf{Distr}\{y|\theta_1 = \mu(\mathbf{x}), \theta_2, \dots, \theta_p\},$$

where

$$\mathbb{E}(y|\mathbf{x}) = \mu(\mathbf{x}) = g^{-1} \Big\{ \sum_{i=1}^m f_i(\mathbf{x}) \Big\},\,$$

and g is the link function.

Poisson GAM:

- $y|\mathbf{x} \sim \mathsf{Pois}\{y|\mu(\mathbf{x})\}$
- $\mathbb{E}(y|\mathbf{x}) = \mathsf{Var}(y|\mathbf{x}) = \exp\left\{\sum_{j=1}^m f_j(\mathbf{x})\right\}$
- $g = \log \text{ assures } \mu(\mathbf{x}) > 0$

Here  $\mathbb{E}(y|\mathbf{x})$  and  $Var(y|\mathbf{x})$  is implied by model...

... or we can have extra parameters for scale and shape.

#### Scaled Student's t GAM:

- $y|\mathbf{x} \sim \mathsf{ScaledStud}\{y|\mu(\mathbf{x}), \sigma, \nu\}$  (i.e.  $(y-\mu)/\sigma \sim \mathsf{Stud}(y|\nu)$ )
- $\mathbb{E}(y|\mathbf{x}) = \mu(\mathbf{x}) = \sum_{j=1}^{m} f_j(\mathbf{x})$
- ullet  $\sigma$  is scale parameter
- ullet u is shape parameter (degrees of freedom)
- $Var(y|\mathbf{x}) = \sigma^2 \frac{\nu}{\nu 2}$

Later we'll let all parameters be functions of x, eg:

•  $y|\mathbf{x} \sim N\{y|\mu(\mathbf{x}), \sigma(\mathbf{x})\}$ 

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Consider additive model

$$\mathbb{E}(y|\mathbf{x}) = \mu(\mathbf{x}) = g^{-1}\Big\{f_1(\mathbf{x}) + f_2(\mathbf{x}) + f_3(\mathbf{x})\Big\},\,$$

where

• 
$$f_1(\mathbf{x}) = \beta_0 + \beta_1 x_1 + \beta_2 x_1^2$$

• 
$$f_2(\mathbf{x}) = \begin{cases} 0 & \text{if } x_2 = \text{FALSE} \\ \beta_4 & \text{if } x_2 = \text{TRUE} \end{cases}$$

•  $f_3(\mathbf{x}) = f_3(x_3)$  is a non-linear smooth function.

Smooth effects built using spline bases

$$f_3(x_3) = \sum_{k=1}^r \beta_k b_k(x_3)$$

where  $\beta_k$  are unknown coeff and  $b_k(x_3)$  are known spline basis functions.

### Example: B-splines

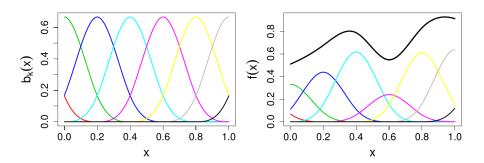
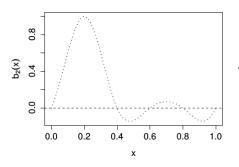
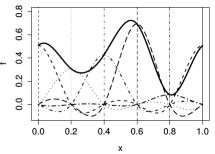


Figure: B-spline basis (left) and smooth (right).

# Types of smooths

$$s(x, bs = "cc")$$





Cyclic cubic regression splines make so that

- $f(x_{min}) = f(x_{max})$
- $f'(x_{min}) = f'(x_{max})$
- $f''(x_{min}) = f''(x_{max})$

### Example: Thin plate regression splines (TPRS)

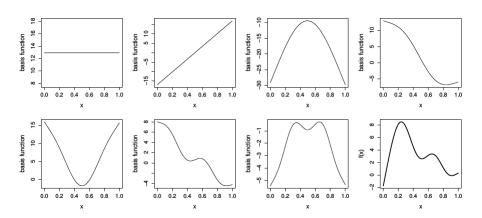


Figure: Rank 7 TPRS basis. Image from Wood (2017).

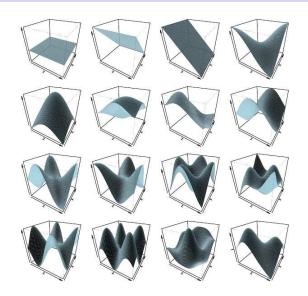


Figure: Rank 17 2D TPRS basis. Courtesy of Simon Wood.

In general

$$f(\mathbf{x}) = \sum_{k=1}^{r} \beta_k b_k(\mathbf{x}).$$

To determine complexity of  $f(\mathbf{x})$ :

- the basis rank r is large enough for sufficient flexibility
- ullet a complexity penalty on eta controls the wiggliness of the effects

# GAM model fitting

 $\hat{oldsymbol{eta}}$  is the maximizer of **penalized** log-likelihood

$$\hat{\boldsymbol{\beta}} = \operatorname*{argmax}_{\boldsymbol{\beta}} \operatorname{PenLogLik}(\boldsymbol{\beta}|\boldsymbol{\gamma}) = \operatorname*{argmax}_{\boldsymbol{\beta}} \big\{ \underbrace{\widetilde{\log p(\mathbf{y}|\boldsymbol{\beta})}}_{\text{genalize complexity}} - \underbrace{\operatorname{Pen}(\boldsymbol{\beta}|\boldsymbol{\gamma})}_{\text{penalize complexity}} \big\}$$

#### where:

- $\log p(\mathbf{y}|\beta) = \sum_{i} \log p(y_i|\beta)$  is log-likelihood (i.i.d. case)
- ullet Pen $(eta|\gamma)$  penalizes the complexity of the  $f_j$ 's
- $\gamma > 0$  smoothing parameters ( $\uparrow \gamma \uparrow$ smoothness)

# GAM model fitting

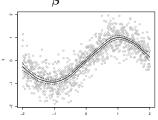
mgcv uses a hierarchical fitting framework:

lacksquare Select  $\gamma$  to determine smoothness

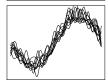
$$\hat{\gamma} = \underset{\gamma}{\operatorname{argmax}} \ \mathsf{LAML}(\gamma).$$

2) For fixed  $\gamma$ , estimate  $\beta$  to determine actual fit

$$\hat{oldsymbol{eta}} = \mathop{\mathsf{argmax}}_{oldsymbol{eta}} \mathop{\mathsf{PenLogLik}}(oldsymbol{eta}|oldsymbol{\gamma}).$$







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Suppose we have data on bone mineral density (bmd) as a function of age.

We have m subjects and n data pairs per subject

- subj 1:  $\{bmd_{11}, age_{11}\}, \dots, \{bmd_{n1}, age_{n1}\}$
- subj j:  $\{bmd_{1j}, age_{1j}\}, \ldots, \{bmd_{nj}, age_{nj}\}$
- subj m:  $\{bmd_{1m}, age_{1m}\}, \dots, \{bmd_{nm}, age_{nm}\}$

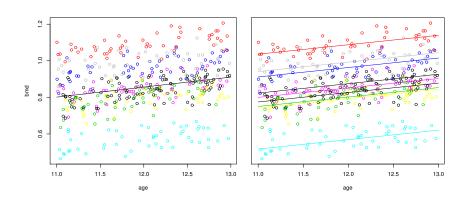
Standard linear model ignores individual differences

$$\mathbb{E}(\textit{bmd}|\textit{age}_{\textit{ij}}) = \mu(\textit{age}_{\textit{ij}}) = \alpha + \beta \textit{age}_{\textit{ij}}.$$

We can include random intercept per subject

$$\mu(\mathsf{age}_{ij}) = \alpha + \beta \mathsf{age}_{ij} + \mathsf{a}_j,$$

where 
$$\mathbf{a} = \{\mathsf{a}_1, \dots, \mathsf{a}_m\} \sim \mathcal{N}(\mathbf{0}, \mathbf{\Sigma}).$$



We can also include random slopes

$$\mu(age_{ij}) = \alpha + (\beta + b_i)age_{ij} + a_i$$

where  $\mathbf{a} \sim \mathcal{N}(\mathbf{0}, \boldsymbol{\Sigma}_{\mathbf{a}})$  and  $\mathbf{b} \sim \mathcal{N}(\mathbf{0}, \boldsymbol{\Sigma}_{\mathbf{b}})$ .

In mgcv random effect are specified as:

In simplest case  $\mathbf{\Sigma_a} = \gamma_{\mathbf{a}}\mathbf{I}$  and  $\mathbf{\Sigma_b} = \gamma_{\mathbf{b}}\mathbf{I}$ , that is

$$\Sigma_{\mathbf{a}} = \begin{bmatrix} \gamma_{\mathbf{a}} & 0 & 0 & \dots & 0 \\ 0 & \gamma_{\mathbf{a}} & 0 & \dots & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & 0 & \dots & \gamma_{\mathbf{a}} \end{bmatrix}$$

Variances  $\gamma_a$  and  $\gamma_b$  must be estimated (later I'll explain how).

Using the gam function in mgcv we can create and estimate simple random effects:

For more complex random effects (e.g. introducing correlation between random slopes and intercepts) we can use:

- gamm function estimates generalized additive mixed models using nlme
- gamm4 package based on lme4 package

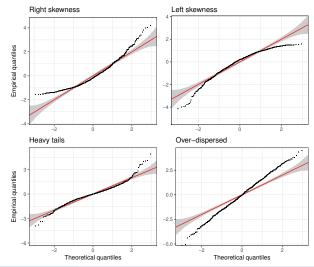
Here we will use only simple effects.

#### These slides cover:

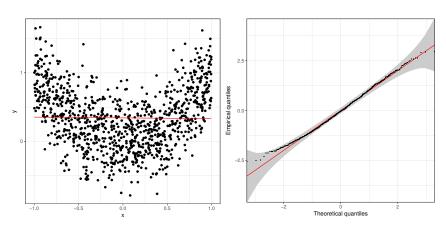
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In the first hands-on session we'll use few basic diagnostics.

### QQ-plots



Useful for choosing model  $\operatorname{Dist}_m(y|\mathbf{x})$  (e.g. Poisson vs Neg. Binom.) Less useful for finding omitted variables and non-linearities.



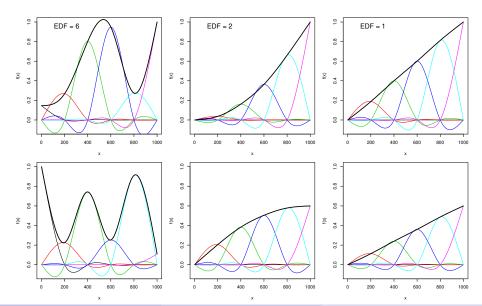
Recall structure of smooth effects:

$$f(\mathbf{x}) = \sum_{j=1}^k \beta_j b_j(\mathbf{x}).$$

where  $\beta$  shrunk toward smoothness by penalty.

Effective number of parameters we are using is < k.

Approximation is **Effective Degrees of Freedom** (EDF) < k.



By default k = 10 but this is arbitrary.

Exact choice of k not important, but it must not be too low.

Checking whether k is too low:

- look at conditional residuals checks
- ② look at output of gam.check(fit):

```
## k' edf k-index p-value

## s(wM) 9.00 8.60 0.91 <2e-16 ***

## s(wM_s95) 9.00 8.13 1.02 0.76

## s(Posan) 8.00 2.66 1.04 0.97
```

 $\odot$  increase k and see if a **model selection criterion** improves

#### Model selection

General criterion is approximate Akaike Information Criterion (AIC):

$$AIC = \underbrace{-2 \log p(\mathbf{y}|\hat{\beta})}_{\text{goodness of fit}} + \underbrace{2\tau}_{\text{model complexity}}$$

where  $\tau$  is EDF.

If  $AIC_{m1} < AIC_{m2}$  choose model 1.

To select which effects to include we can also look at p-values:

```
summary(fit)
## Estimate Std. Error t value Pr(>|t|)
## (Intercept) 267.2004 75.4197 3.543 0.000405 ***
## Fl 6.2854 1.0457 6.010 2.20e-09 ***
## loc2 79.8459 80.4130 0.993 0.320858
## loc3 -71.2728 86.1725 -0.827 0.408284
```

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### GAMs in mgcv

The exercises will be based on the mgcv package for GAM modelling.

mgcv is a recommended R package, included in R by default.

It contains methods for:

- creating GAM models
- fitting them
- visualizing and summarizing model output

## GAMs in mgcv

There are alternatives to mgcv, such as:

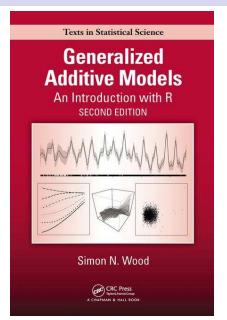
- mboost (Hothorn et al., 2010)
- gamlss (Rigby and Stasinopoulos, 2005)
- brms (Bürkner et al., 2017)
- BayesX (Brezger et al., 2003)
- INLA (Rue et al., 2009)

Each offers much flexibility (e.g. smooth effects types and distributions).

Strong points of mgcv's methods:

- Iittle tuning needed (automatic smoothing parameters selection)
- 2 fast and stable numerical implementation

## Further reading



### References I

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