This is a title Henry Linder mhlinder@gmail.com October 31, 2016

A citation: Gelman, Carlin, Stern, & Rubin (2014)

## A classical (frequentist) approach

Consider a simple linear regression of tree height  $y_i$  on age  $x_i$ 

$$y_i = \beta_0 + \beta_1 x_i + \epsilon_i \tag{1}$$

where  $\epsilon_i$  are independent  $N(0, \sigma^2)$  errors. For how it will pertain to the Bayesian analysis, note two things: first, we can equivalently write

$$y_i \sim N(\beta_0 + \beta_1 x_i, \sigma^2) \tag{2}$$

and second, this implies that the likelihood function for the data  $\mathbf{y} = (y_1, \dots, y_n)'$  is  $f(\mathbf{y}|\theta, \mathbf{x})$  where the data  $\mathbf{x} = (x_1, \dots, x_n)'$  is assumed known and the parameter vector is  $\theta = (\beta_0, \beta_1)'$ .

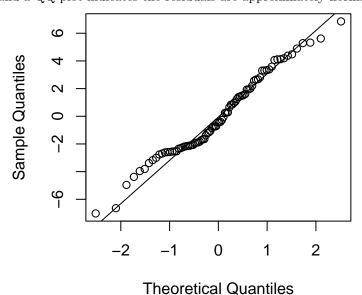
A regression fit is given by

```
m <- lm(height ~ age, data = Loblolly)
summary(m)</pre>
```

```
##
## Call:
## lm(formula = height ~ age, data = Loblolly)
##
## Residuals:
##
                1Q Median
                                        Max
## -7.0207 -2.1672 -0.4391 2.0539
                                    6.8545
##
## Coefficients:
               Estimate Std. Error t value Pr(>|t|)
## (Intercept) -1.31240
                           0.62183
                                    -2.111
                                              0.0379 *
                2.59052
                           0.04094
                                    63.272
                                              <2e-16 ***
## ---
                   0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
## Signif. codes:
##
## Residual standard error: 2.947 on 82 degrees of freedom
```

```
## Multiple R-squared: 0.9799, Adjusted R-squared: 0.9797 ## F-statistic: 4003 on 1 and 82 DF, p-value: < 2.2e-16
```

Note, in particular, that the regression coefficient  $\beta_1$  is highly significant, the  $R^2$  is high, and a QQ plot indicates the residuals are approximately normal:



## A Bayesian approach

A conventional, convenient, and conjugate choice in Bayesian regression gives independent normal priors to  $\beta_0$  and  $\beta_1$ . For simplicity, let  $\beta_0 \sim N(0,1)$  and  $\beta_1 \sim N(0,1)$ . A Bayesian analysis revolves around the fundamental relationship between the prior distribution  $p(\theta)$  and the posterior distribution  $p(\theta|\mathbf{y})$  of the parameter vector  $\theta = (\beta_0, \beta_1)'$ :

$$p(\theta|\mathbf{y}) \propto f(\mathbf{y}|\theta, \mathbf{x})p(\theta)$$
 (3)

where  $f(\mathbf{y}|\theta, \mathbf{x})$  is the likelihood of the observed data, and we take  $\mathbf{x}$  as given. In the present case, we can write the likelihood as

$$f(\mathbf{y}|\theta, \mathbf{x}) = \prod_{i=1}^{n} f(y_i|\theta, x_i) \propto \exp\left[-\frac{1}{2\sigma^2} \sum_{i=1}^{n} (y_i - \beta_0 - \beta_1 x_i)^2\right]$$
(4)

where we omit the constant of proportionality for clarity.

Too, the prior for  $\theta$  is

$$p(\theta) = p(\beta_0)p(\beta_1) \propto \exp\left[-\frac{1}{2}\left(\beta_0^2 + \beta_1^2\right)\right]$$
 (5)

So, we can find the posterior distribution as

$$p(\theta|\mathbf{y}) \propto p(\beta_0|\mathbf{y})p(\beta_1|\mathbf{y})$$
 (6)

## References

Gelman, A., Carlin, J. B., Stern, H. S., & Rubin, D. B. (2014). *Bayesian data analysis* (Vol. 2). Chapman & Hall/CRC Boca Raton, FL, USA.