

Causal Analysis: A Quick Intro

Part 2: Introduction to Machine Learning for Causal Analysis using
Observational Data

CAUSAL ANALYSIS USING DATA FROM OBSERVATIONAL STUDIES

- > We want to estimate the *causal effect* of treatment or social exposure T on outcome Y
 - > Causal effect is policy-relevant: what benefits accrue if we intervene to change T ?
 - > Treatment must be *modifiable* for this to make sense – otherwise, what's the point??
- > We have data from an **observational** study where T and Y are measured
 - > How were the individual units in the data set collected?
 - > Which population were these units drawn from?
 - > Temporal ordering: are we sure treatment was determined before outcome? **If not, game over!**

REGRESSION ESTIMATION

> Linear regression is workhorse for effect estimation

> For subject i , we observe their treatment t_i and outcome y_i and fit the model

$$y_i = a + bt_i + \epsilon_i$$

where we focus on binary treatment

$$t_i = \begin{cases} 1 & \text{if } i \text{ received treatment} \\ 0 & \text{control} \end{cases}$$

> Coefficient b is the **difference between the mean outcomes in the treatment and control groups**

> Usually estimate using ordinary least squares or maximum likelihood

Note: Can elaborate regression model if treatment is continuous

e.g. Add t_i^2, t_i^3, t_i^4 , etc. terms (curvilinear) or use dummy variables (stepwise linear) to capture more complex relationships in a limited way

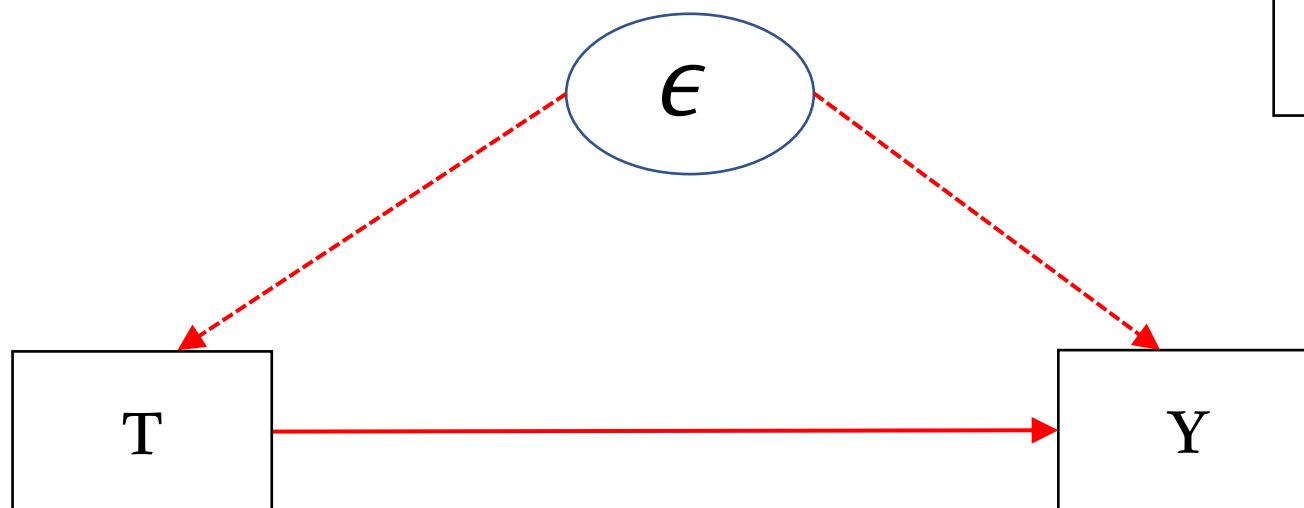
CONFOUNDING IN OBSERVATIONAL STUDIES

- > Regression coefficient b is a measure of association between T and Y
 - > Would equal *causal effect* if RCT data (randomised controlled trial) where T was randomized
- > But T not randomised: treatment selected **in a way that could depend (indirectly)** on Y
 - > Same ‘type’ of person who chooses treatment is also the sort who has high outcome (& vice versa)
 - > Banks give loans to people more likely to successfully pay off their loans
 - > Children from wealthier families more likely to attend private school and have better post-school outcomes
- > Would have done better anyway: association *confounds* this with the true effect of treatment

1. Graph for association



2. Causal graph



$$\hat{b} \neq ATE$$

ROLE OF BASELINE VARIABLES

- > Suppose study measures many other variables $X = (X_1, X_2, \dots, X_p)$
- > Throw away those we know happened after the treatment was chosen
 - > Not a baseline variable if so!
 - > We need to be sure we have a quasi-experimental study
- > Distribution of X generally different for treated and untreated in observational study

RANDOMISED CONTROLLED TRIAL

$$\hat{b} = ATE$$

Baseline
variables

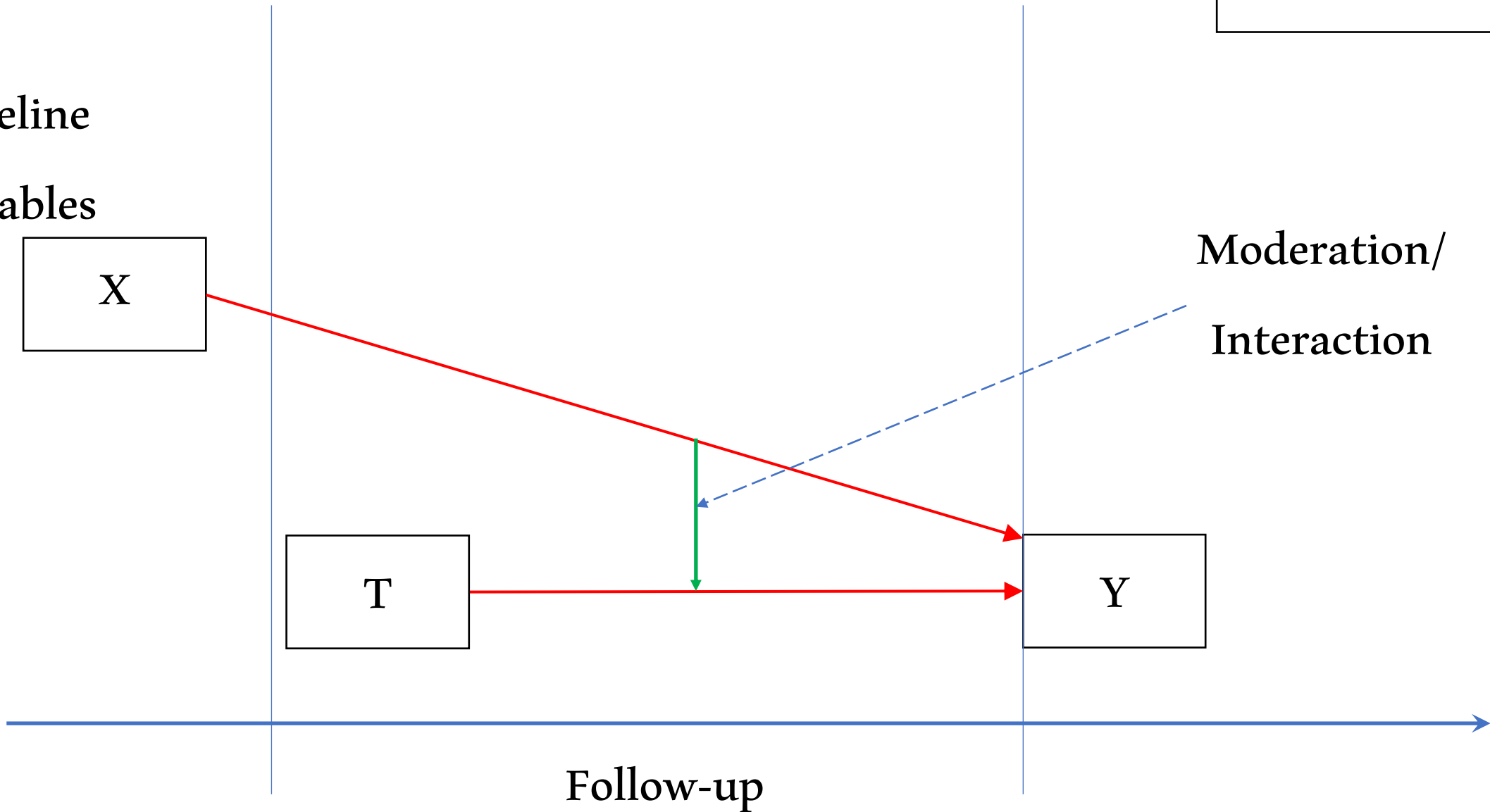
X

T

Y

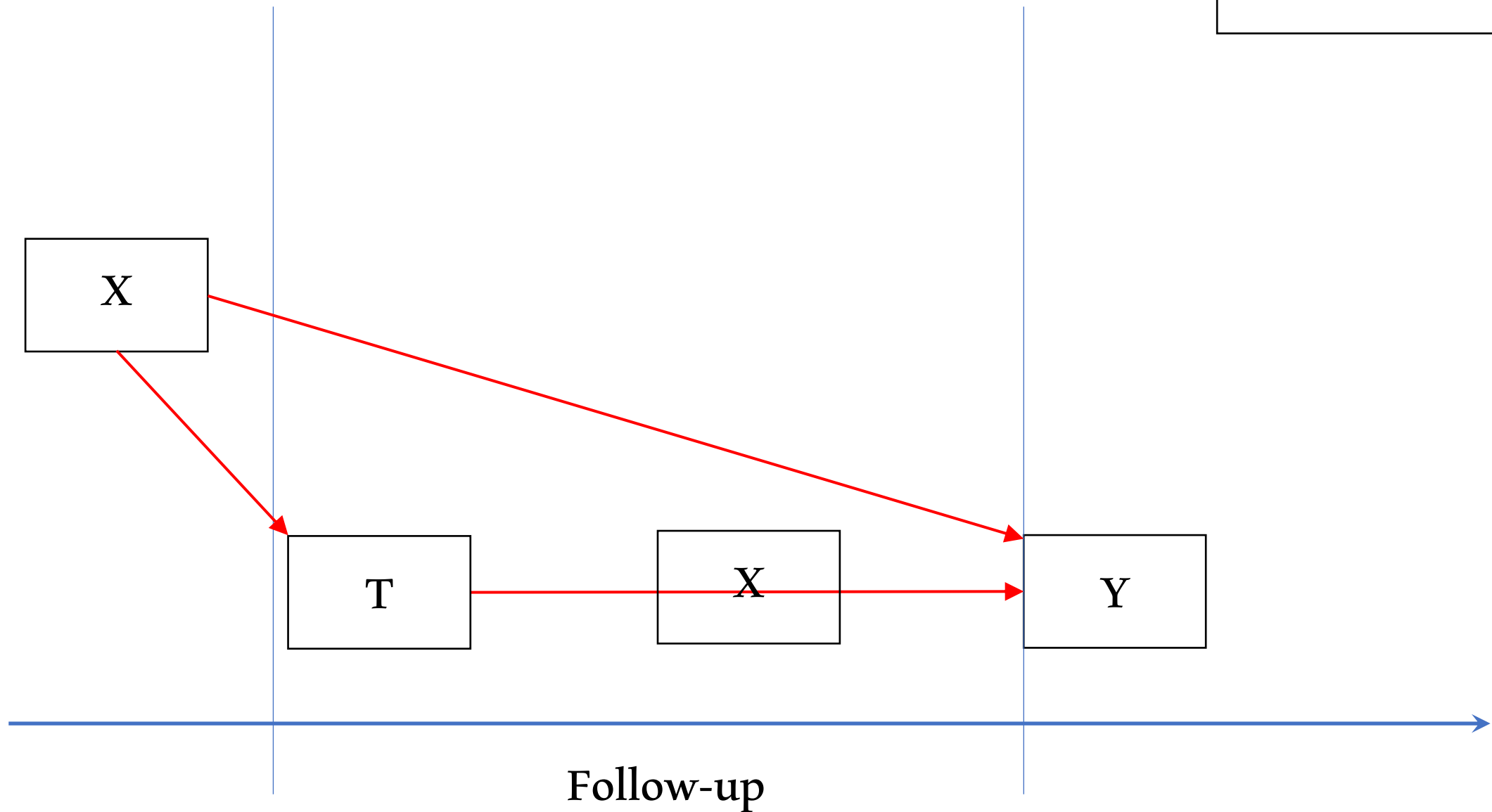
Moderation/
Interaction

Follow-up



OBSERVATIONAL STUDIES AS QUASI-EXPERIMENTS

$$\hat{b} \neq ATE$$



CAUSAL EFFECT: THE AVERAGE TREATMENT EFFECT (ATE)

> Potential outcomes

> For subject i , we observe their treatment t_i and outcome y_i and fit the model

$$y_i^0, y_i^1$$

where we observe only one

$$y_i = \begin{cases} y_i^1 & \text{if } i \text{ received treatment} \\ y_i^0 & \text{control} \end{cases}$$

> The average/mean of $y_i^1 - y_i^0$ across everyone in the *target population*

$$\text{ATE} = \frac{1}{N} \sum_i (y_i^1 - y_i^0) = E[y_i^1 - y_i^0]$$

Notes: If treatment T is polytomous or continuous then y_i^t is a set of values so need model for effect of treatment as many different ways of measuring treatment effect

Implicitly assume stable unit treatment value assumption (SUTVA): potential outcomes don't depend on what other units get

IGNORABLE SELECTION

Independent/Uncorrelated



- > Treatment selection is (strongly) ignorable if

$$\begin{pmatrix} y_i^1 \\ y_i^0 \end{pmatrix} \perp\!\!\!\perp t_i \mid x_i$$

- > Differences between treated and untreated among those subjects characterized by same X are **random**
- > Referred to as **no unobserved confounding** or **no omitted variables**
- > The challenges now are
 - > Verifying this is true [clue: *We can't! But must do what we can to mitigate confounding*]
 - > Adjusting the estimate of b to account for these effects [Today's focus!]

Other approaches needed if there is unobserved confounding (e.g. instrumental variables) but beyond scope

Weakly ignorable $y_i^0 \perp\!\!\!\perp t_i \mid x_i$ --- generally estimate ATE *among the treated*: $ATT = E[y_i^1 - y_i^0 \mid t_i = 1]$

CONDITIONAL AVERAGE TREATMENT EFFECT

- > Introducing X also introduces the conditional average treatment effect (CATE)

$$\text{CATE}(x_i) = E[y_i^1 - y_i^0 | x_i]$$

- > ATE is simply the average value of $\text{CATE}(x_i)$ in the population:

$$\text{ATE} = E\{\text{CATE}(x_i)\}$$

- > Also written as difference between conditional means:

$$\text{CATE}(x_i) = E[y_i^1 | x_i] - E[y_i^0 | x_i] = \mu_1(x_i) - \mu_0(x_i)$$

- > We exploit this later on...
-

REGRESSION ADJUSTMENT

- > Include X variables in the regression model

$$y_i = a + bt_i + cx_i + \epsilon_i$$

where $cx_i = c_1x_{1i} + \dots + c_px_{pi}$ is linear combination of the confounding variables

- > Regression works if the mean of untreated potential outcomes is linear, that is,

$$\mu_0(x_i) = E[\mathcal{Y}_i^0 | x_i] = a + cx_i$$

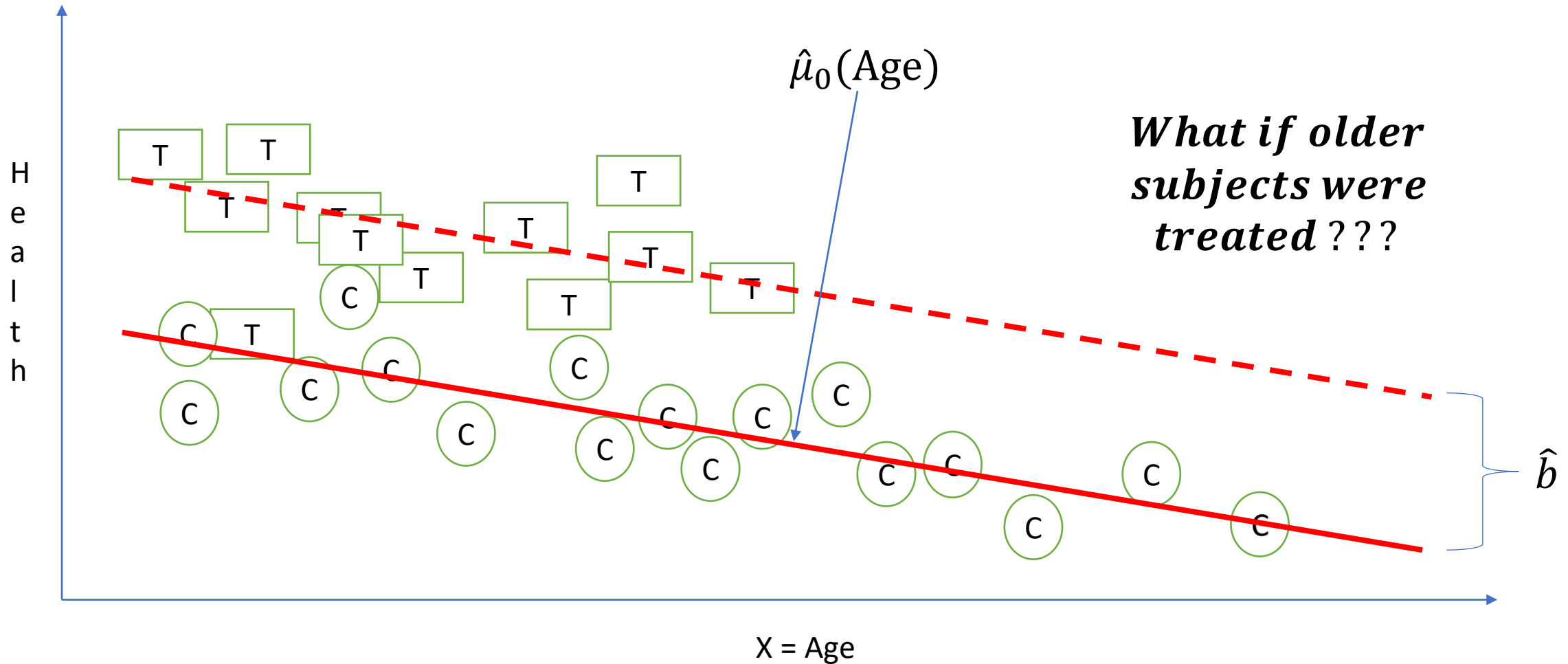
- > ... and the causal effects are **homogeneous**

- > This means that $\text{CATE}(x_i)$ is a constant that does not depend on x_i

- > Simply add b to get the treated mean: $\mu_1(x_i) = \mu_0(x_i) + b$

- > ‘Extrapolates’ whenever there is *no overlap* in the data (see next slide)

COVARIATE (NO) OVERLAP



INVERSE PROBABILITY WEIGHTING

- > Specify same model for treatment (it excludes X) as if data were from RCT

$$y_i = a + bt_i + \varepsilon_i$$

- > Handle selection by estimating *selection propensities*

$$\Pr[t_i = 1|x_i] = e(x_i)$$

- > No longer assume homogeneous effects

- > $e(x) = 0$ or 1 implies **no overlap** - makes clear we cannot estimate $\text{CATE}(x)$

- > Balance the sample by fitting a weighted regression with weights

$$w_i = \frac{t_i}{\hat{e}(x_i)} + \frac{1 - t_i}{1 - \hat{e}(x_i)} = \begin{cases} 1/\hat{e}(x_i) & \text{for treated} \\ 1/1 - \hat{e}(x_i) & \text{for untreated} \end{cases}$$

Estimate using logistic regression

Note. Ignorable assumption needed to ensure that $\Pr[t_i = 1|y_i^0, y_i^1, x_i] = \Pr[t_i = 1|x_i]$

Selection propensities can also play a key role for matching estimators (to estimate ‘counterfactual’ $y_i^{1-t_i}$ to match ‘factual’ y_i)

(SUPERVISED) MACHINE LEARNING (RECAP)

- > Algorithms that learn the true relationship between *input variables* and *output variables*
 - > Set up to accurately predict outputs/outcomes
 - > We call different ML algorithms *base learners* or just *learners*
 - > Earlier discussed regression classification, decision trees, random forests
- > Differences
 - > Move away from parametric models $Y = f(X; \theta)$, just learn rule $f: X \rightarrow Y$
 - > Move from statistical model selection to *train*, *validate* and *test* (incl. setting *meta-parameters*)
 - > Results in predicted outcomes rather than parameter estimates

Note. Even regression classification, linear model simply device for prediction, do not need to believe it is true

META-ALGORITHMS FOR ESTIMATING ATE: S-, T- AND X-LEARNERS

- > Use the power of ML to estimate causal effects more accurately
- > Learn $\mu_t(x_i)$ or $e(x_i)$ or both (depends on which estimator you choose)
- > Different learning strategies for ATE: S – single, T – two, X – hybrid strategy
- > Then plug in learnt $\mu_t(x_i)$ or $e(x_i)$ to a valid estimators of ATE:

$$\widehat{\text{ATE}}_{\text{PredDiff}} = \frac{1}{n} \sum_i \{\hat{\mu}_1(x_i) - \hat{\mu}_0(x_i)\}$$

$$\widehat{\text{ATE}}_{\text{IPW}} = \frac{1}{n} \sum_i \left\{ \frac{t_i}{\hat{e}(x_i)} - \frac{1 - t_i}{1 - \hat{e}(x_i)} \right\} y_i = \frac{1}{n} \sum_i w_i y_i$$

$$\widehat{\text{ATE}}_{\text{Pred-IPW}} = \frac{1}{n} \sum_i w_i \hat{\mu}_{t_i}(x_i)$$

DOUBLY ROBUST ESTIMATION

- > Theoretically,* the most accurate (low bias, small variance) estimator of all
- > Combines strengths of $\hat{\mu}_t(x_i)$ or $\hat{e}(x_i)$ (although $\widehat{ATE}_{\text{Pred-IPW}}$ does this too)
- > Also allows for biased estimation of either $\hat{\mu}_t(x_i)$ or $\hat{e}(x_i)$ but not both (no free lunch!)
- > Simply combine three estimators from previous slides:

$$\widehat{ATE}_{\text{DR}} = \widehat{ATE}_{\text{PredDiff}} + \widehat{ATE}_{\text{IPW}} - \widehat{ATE}_{\text{Pred-IPW}}$$

* Theoretically - simply means in large samples; its performance in small-to-medium sized samples is less clear-cut

S-LEARNERS

We'll use random forest regression

- > Learn single structural model from **all available data**

$$s(t, x) = E[y_i | t_i = t, x_i = x] + \text{error}$$

- > Then estimate

$$\hat{\mu}_0(x_i) = s(0, x_i) \text{ and } \hat{\mu}_1(x_i) = s(1, x_i)$$

- > Compared with linear regression:

- > Allows $\mu_0(x_i) = E[y_i^0 | x_i]$ to be non-linear
- > and heterogenous treatment effects

T-LEARNERS

> T: Two stages, one for treated, one for untreated:

> From treated units, learn $s_1(x_i) = E[y_i | t_i = 1, x_i] + \text{error}$

> From control units, learn $s_0(x_i) = E[y_i | t_i = 0, x_i] + \text{error}$

> Then

$$\hat{\mu}_0(x_i) = s_0(x_i) \text{ and } \hat{\mu}_1(x_i) = s_1(x_i)$$

Random forest regression



META-ALGORITHMS: X-LEARNER

- > Combines learner predictions with observed data:
- > As with T-learner, learn $s_1(x_i)$ from treated units and $s_0(x_i)$ from untreated units
- > Also learn selection propensity $e(x)$ using suitable base learner
- > Calculate ‘imputed’ individual treatment effects

$$D_i = \begin{cases} D_i^1 := y_i - s_0(x_i) & \text{if } i \text{ is treated} \\ D_i^0 := s_1(x_i) - y_i & \text{if } i \text{ is untreated} \end{cases}$$

- > Apply base learner to untreated and treated groups:
 - a) Learn $\hat{\tau}_0(x) = E[D_i^0 | x_i = x]$ and $\hat{\tau}_1(x) = E[D_i^1 | x_i = x]$
 - b) Calculate $\widehat{\text{CATE}}(x) = \hat{\tau}_0(x)(1 - \hat{e}(x)) + \hat{\tau}_1(x)\hat{e}(x)$
- > Estimate of ATE is simply sample average of $\widehat{\text{CATE}}(x)$

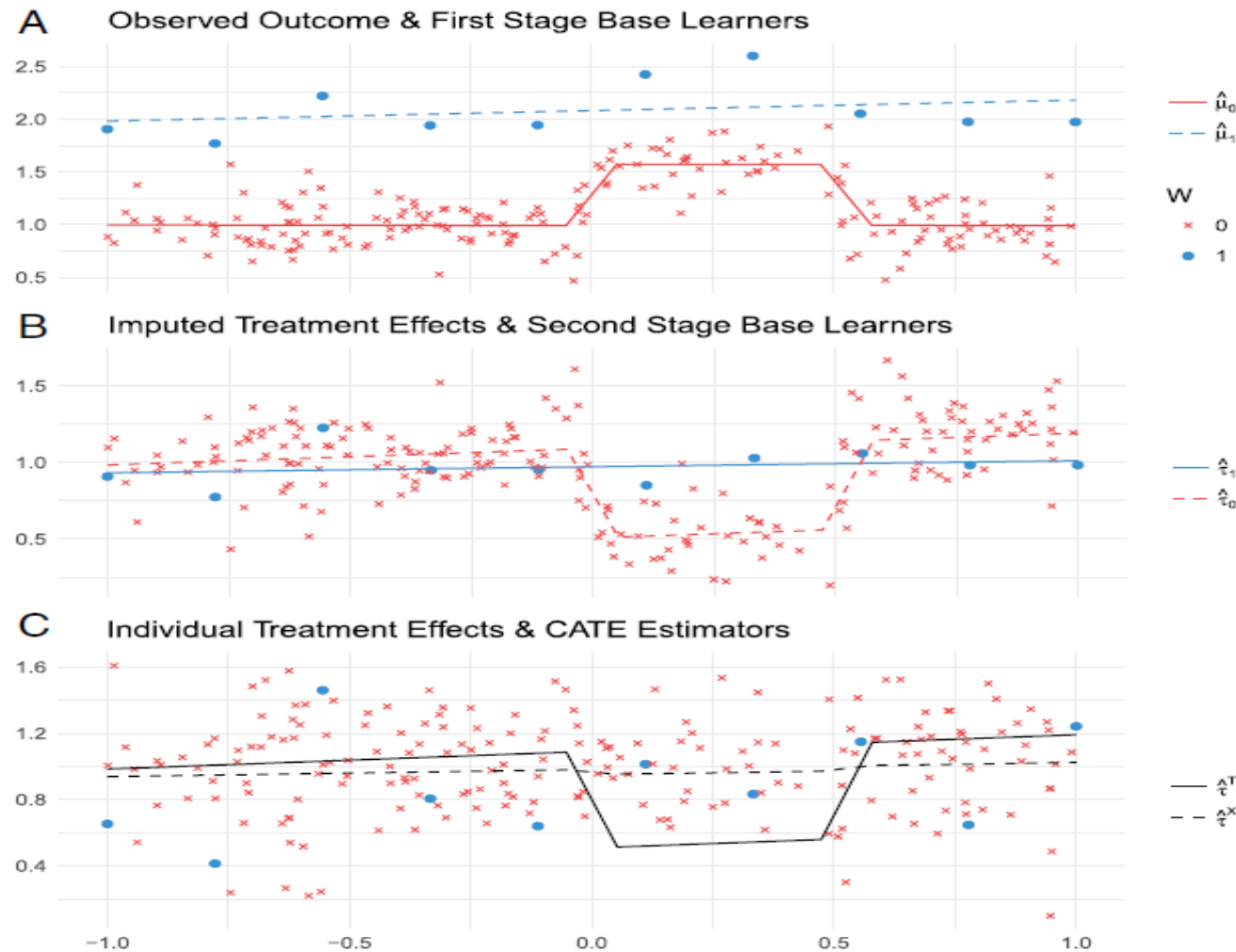


Fig. 1. Intuition behind the X-learner with an unbalanced design. (A) Observed outcome and first-stage base learners. (B) Imputed treatment effects and second-stage base learners. (C) ITEs and CATE estimators.

From Kuntzel et al. (2019) Metalearners for estimating heterogeneous treatment effects using machine learning.

SOME REFERENCES AND FURTHER READING

- **Wager and Athey (2018)** [identifying heterogenous treatment effects with random forests]
<https://doi.org/10.1080/01621459.2017.1319839>
- **Econ-ML repository** [research papers on ML in economics – there's a lot of work going on!]
<http://econ-neural.net/>
- **Hernan and Robins (2020)** [exhaustive book on causal analysis]
https://cdn1.sph.harvard.edu/wp-content/uploads/sites/1268/2020/02/ci_hernanrobins_21feb20.pdf
- **Kuntzel et al (2019)** [X-learners vs S- and T-learners]
<https://doi.org/10.1073/pnas.1804597116>
- **Xu et al (2020)** [where the computer scientists are headed...]
<https://arxiv.org/abs/2006.16789>

Now to the practical bit

But first, any questions...?