Lab 4

Multivariate Statistics with R

packages needed: lme4, arm, lattice (lattice comes with R; it needs to be loaded using library() but there is no need to install it)

Examples in previous lab sessions have highlighted changes over time (vocabulary study, arthritis study, fertility and climate data). Now we are going to focus on an alternative setup for repeated measurements where individuals are studied under a series of related conditions. We will use experiments from psycholinguistics (Lab 4) and cognitive psychology (Lab 5) as examples.

Nested and crossed effects

An important advantage of the lme4 package in R is that it offers mixed-effects modelling with crossed random effects. For illustration, we will use a (constructed) dataset taken from Raaijmakers, Schrijnemakers, and Gremmen (1999) (see their Table 2). A lexical decision task was performed where subjects had to decide whether each stimulus was a word. The interval between the presentation of a distractor stimulus and the target word was varied to see if this affected reaction time in the classification task. This interval is referred to as stimulus onset asynchrony (SOA). Eight words were shown to each participant under either a long or short SOA. Load the data.

Load data and examine design

```
# Load lexical decision task data
ldt <- read.table("raai.dat", header = T)</pre>
```

Question 1: Use table() to explore the design of the experiment. How many times does each participant see each stimulus item W1-W8? Are the ways of the data (subject, SOA, and item) nested or crossed?

Tip:

• Pass one or two columns from a data to table() to build a contingency table.

```
# Examine experimental design
table(ldt$SOA)
table(ldt[, c('Subject', 'Item')])
table(ldt[, c('Subject', 'SOA')])
table(ldt[, c('Item', 'SOA')])
```

• Nested means a group at one level of the hierarchy only appears in one of the higher-order levels. Crossed designs mean that each of the lower-order group is observed within each of the higher-order groups.

Variance component models

Because the stimulus words can be seen as samples from a larger population of possible stimuli, it is useful to model them as random effects along with the subjects. We start with a model that estimates the effect of SOA and reaction time while accounting for differences between subjects.

¹This data is also discussed in Baayen (2008) (Section 7.2.1, note that the wording and data frames differ in different (pre)prints of the book) as well as Baayen *et al.* (2008, p. 401).

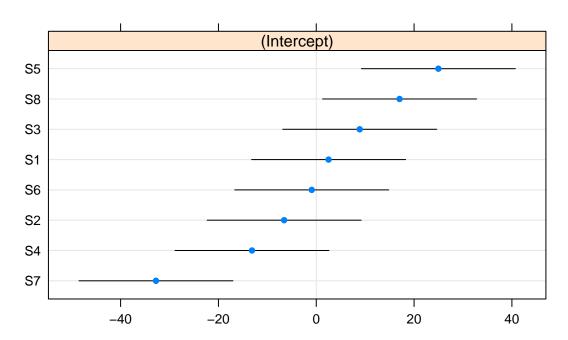
```
library(lme4)

# Model with subject as a random effect
ldt.M1 <- lmer(RT ~ 1 + SOA + (1 | Subject), data = ldt)</pre>
```

The subject intercepts from the model can be visualized with a "caterpillar plot" for conditional modes (lecture 3, slide 56; lecture 4, slide 33):

```
# Dot plot of random effects for subject
print(dotplot(ranef(ldt.M1, condVar = T))[['Subject']])
```

Subject



The condVar = T requests the standard errors for each subject's parameter estimate. The name of the group to be plotted goes between the [[' ']]. (In lme4 versions < 1.0 the condVar argument was named postVar.)

Question 2: What is the difference in reaction times between the long and short SOA condition?

- tells us that SOA is a factor with 2 levels "long" and "short" (long is the reference)
- tells us that reaction times for the short SOA are 22.4 ms longer than for the long SOA

Question 3: Model ldt.M1 fits an intercept for each subject as a random effect. Build another model, called ldt.M2, that also fits an intercept for each item as a random effect.

Tip: Note that the determination of nested or crossed random effects is done for you by lmer().

Question 4: Does model 2 offer a better fit compared to the likelihood ratio test?

Question 5: Which model has the better fit according to AIC?

Question 6: How does the estimate of the difference in RT between the long and short SOA from model 2 compare to model 1?

Between subject variability in effect of SOA

Like fitting a different slope for each subject to represent change over time in a longitudinal study, individual differences in treatment effects or experimental conditions can also be examined with linear mixed-effects models. For example, do subjects differ in how quickly they react under short versus long SOA?

```
# Varying slope of SOA for Subject
ldt.M3 <- lmer(RT ~ 1 + SOA + (1 + SOA | Subject) + (1 | Item), data = ldt)</pre>
```

Question 7: Do subjects differ in how changes in SOA affect their reaction times?

Question 8: Visualise these individual differences. Technically, produce a caterpillar plot of the by-subject random slopes for SOA.

Question 9: Based on this incremental model building, what is your conclusion with regard to the SOA effect?

Tip: keyword is *significance* of fixed effects.

Random-effects structure & zero-correlation parameter model

It has been suggested that linear mixed-effects models generalize best when they include the maximal random effects structure justified by the design (Barr, Levy, Scheepers, & Tily, 2013), see Lecture 4 slide 37 onwards.

Question 10: What is the maximal random effects structure justified by the design for the lexical decision data in the ldt data frame?

Tip: Random slopes are only appropriate for within-subjects and within-items factors solution.

Let's take another look at ldt.M3. The by-subject random effects for slopes and intercepts are paired observations. Therefore the model specification that we used here allows for these two random variables (intercept and slope) to be correlated. The estimate of this correlation is a parameter of the mixed effects model.

Question 11: How large is this correlation?

Tip: You will find it in the 'summary' output of model ldt.M3.

A question that arises is whether there is reliable evidence that this correlation parameter is different from zero. We can test this by specifying a zero-correlation parameter model, *i.e.*, we are forcing the correlation between by-subject random intercept and slope to be zero (See slides 46-48 in Lecture 4, from slide 47).

The syntax for uncorrelated random intercept and slope is

```
1 + A + (1 | subject) + (0 + A | subject) , or, more conveniently,  A + (A \mid \mid subject)
```

Question 12: Change the random-effects structure of ldt.M3 such that it turns into a zero-correlation parameter model; call this new model ldt.M3.zcp.

- a) Use the above syntax
- b) Well, if you use above syntax, there is still a correlation in the model output, how annoying. As I mentioned in the lecture, there are (currently) some constraints on using this syntax. A detailed account is provided here. Specifically, 'SOA' is a categorical variable and therefore coded as a factor. The solution to the problem is to turn 'SOA' into a numeric variable. If you don't know how to convert a factor to a numeric variable, ask Google. Then try again to specify a zero-correlation parameter model.

Question 13: Perform a likelihood ratio test to check whether forcing the correlation parameter to be zero significantly decreases the goodness of fit.

Tip: The critical model for this comparison are models ldt.M3.zcp and ldt.M3

Question 14: Match the R output obtained when answering Question 13 to a sentence of the following form: According to the likelihood ratio test, ldt.M3.zcp offers an [equal/worse/better] fit than ldt.M3, $\chi^2(\Delta df) = ??.??, p = .???$.

Tip:

- χ^2 is the Chi square value equal to the difference of the model deviances (which are equal to $-2 \times \log$ -likelihood, AKA -2LL)
- Δdf is difference in degrees of freedom (i.e., numbers of parameters) between the two models

References

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Baayen, R. H., Davidson, D. J., & Bates, D. M. (2008). Mixed-effects modeling with crossed random effects for subjects and items. *Journal of Memory and Language*, 59(4), 390-412. doi:10.1016/j.jml.2007.12.005

Barr, D. J., Levy, R., Scheepers, C., & Tily, H. J. (2013). Random effects structure for confirmatory hypothesis testing: Keep it maximal. *Journal of Memory and Language*, 68(3), 255-278. doi:10.1016/j.jml.2012.11.001

Raaijmakers, J. G. W., Schrijnemakers, J. M. C., & Gremmen, F. (1999). How to deal with "The

language-as-fixed-effect fallacy": Common misconceptions and alternative solutions. $Journal\ of\ Memory$

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