

# MATH 308 Assignment 16:

## Exercises 6.4

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**1**

We know that  $X \sim \mathcal{B}(n, p) \implies f_X(x) = \binom{n}{x} p^x (1-p)^{n-x}$ . Thus, the probability of obtaining the value  $X$  can be written as a function of the parameter  $p$ :

$$\begin{aligned} L(p) &= \binom{n}{X} p^X (1-p)^{n-X} \\ \implies L' / \binom{n}{X} &= X p^{X-1} (1-p)^{n-X} \\ &\quad - p^X (n-X) (1-p)^{n-X-1} \end{aligned}$$

We can then obtain the MLE  $\hat{p}$  by setting  $L' = 0$ :

$$\begin{aligned} X p^{X-1} (1-p)^{n-X} &= p^X (n-X) (1-p)^{n-X-1} \\ \implies \hat{p}(n-X) &= X(1-\hat{p}) \\ \implies \hat{p}n - \hat{p}X &= X - X\hat{p} \\ \implies \hat{p} &= X/n \quad \square \end{aligned}$$

**2**

$X \sim \mathcal{P}(\lambda) \implies f_X = \frac{\lambda^x e^{-\lambda}}{x!}$ . So, the likelihood of obtaining the sample  $\{x_1, x_2, \dots, x_n\}$  is  $L(\lambda) = \prod_{i=1}^n f_X(x_i)$ , since the elements of the sample are i.i.d.

$$\begin{aligned} L &= \prod_{i=1}^n \frac{\lambda^{x_i} e^{-\lambda}}{x_i!} \\ &= \frac{\lambda^{\sum x_i} e^{-n\lambda}}{\prod (x_i!)} = \lambda^{\sum x_i} e^{-n\lambda} / \mathcal{C} \end{aligned}$$

$$\begin{aligned} \implies LC &= \lambda^{\sum x_i} e^{-n\lambda} \\ \implies \log L + \log \mathcal{C} &= \sum x_i \log \lambda - n\lambda \\ \implies L' / L &= \sum x_i / \lambda - n \\ \implies 0 &= \sum x_i / \hat{\lambda} - n \\ \implies \hat{\lambda} &= \sum x_i / n = \bar{x} \quad \square \end{aligned}$$

**4**

As in the previous problem,

$$\begin{aligned} L(\theta) &= \prod_{i=1}^n \frac{x_i^3 e^{-x_i/\theta}}{6\theta^4} \\ &= \frac{(\prod x_i)^3 e^{-\sum x_i/\theta}}{6^n \theta^{4n}} \\ &= \mathcal{C} e^{-\sum x_i/\theta} \theta^{-4n} \\ \implies \log L - \log \mathcal{C} &= -\sum x_i / \theta - 4n \log \theta \\ \implies L' / L &= \sum x_i / \theta^2 - 4n / \theta \\ \implies 0 &= \frac{\sum x_i}{\hat{\theta}} - 4n \\ \implies \hat{\theta} &= \frac{\sum x_i}{4n} = \bar{x} / 4 \end{aligned}$$

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a)

$$\begin{aligned}
L(\mu) &= \prod_{i=1}^n \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(x_i - \mu)^2}{2\sigma^2}} \\
&= \mathcal{C} \prod_{i=1}^n e^{-\frac{x_i^2 + \mu^2 - 2x_i\mu}{2\sigma^2}} \\
&= \mathcal{C} \prod_{i=1}^n (e^{\frac{x_i\mu}{\sigma^2}} / e^{\frac{\mu^2}{2\sigma^2}}) \\
&= \mathcal{C} e^{\frac{\mu\Sigma x}{\sigma^2}} / e^{\frac{n\mu^2}{2\sigma^2}} \\
&= \mathcal{C} e^{\frac{2\mu\Sigma x - n\mu^2}{2\sigma^2}} \\
\Rightarrow \log L - \log \mathcal{C} &= \frac{2\mu\Sigma x - n\mu^2}{2\sigma^2} \\
\Rightarrow L'/L &= \frac{2\Sigma x - 2n\mu}{2\sigma^2} \\
\Rightarrow 0 &= \Sigma x - n\hat{\mu} \\
\Rightarrow \hat{\mu} &= \Sigma x / n = \bar{x}
\end{aligned}$$

b)

$$\begin{aligned}
L(\sigma) &= \prod_{i=1}^n \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(x_i - \mu)^2}{2\sigma^2}} \\
&= \frac{\mathcal{C}}{\sigma^n} \prod_{i=1}^n e^{\frac{\Sigma d_i^2}{2\sigma^2}} \\
\Rightarrow L/\mathcal{C} &= e^{\frac{\Sigma d_i^2}{2\sigma^2}} / \sigma^n \\
\Rightarrow \log L - \log \mathcal{C} &= -\frac{\Sigma d_i^2}{2\sigma^2} - n \log \sigma \\
\Rightarrow L'/L &= \Sigma d_i^2 / \sigma^3 - n/\sigma \\
\Rightarrow 0 &= \Sigma d_i^2 / \hat{\sigma}^2 - n
\end{aligned}$$

$$\begin{aligned}
\Rightarrow \hat{\sigma}^2 &= \Sigma d_i^2 / n \\
&= \Sigma (X_i - \mu)^2 / n \\
\Rightarrow \hat{\sigma} &= \sqrt{\Sigma (X_i - \mu)^2 / n}
\end{aligned}$$

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$$\begin{aligned}
L(\mu) &= \frac{1}{\sigma_X\sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma_X^2}} \frac{1}{\sigma_Y\sqrt{2\pi}} e^{-\frac{(y-1.3\mu)^2}{2\sigma_Y^2}} \\
&= \mathcal{C} e^{-\frac{(x-\mu)^2}{2\sigma_X^2} - \frac{(y-1.3\mu)^2}{2\sigma_Y^2}} \\
\Rightarrow \log L - \log \mathcal{C} &= -\frac{(x-\mu)^2}{2\sigma_X^2} - \frac{(y-1.3\mu)^2}{2\sigma_Y^2} \\
\Rightarrow L'/L &= \frac{(x-\mu)}{\sigma_X^2} + \frac{(y-1.3\mu)}{\sigma_Y^2} \\
\Rightarrow 0 &= \frac{95 - \hat{\mu}}{15^2} + \frac{130 - 1.3\hat{\mu}}{20^2} \\
\Rightarrow \hat{\mu} &\approx 97.1
\end{aligned}$$

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$$\begin{aligned}
L(r, \lambda) &= \prod_{i=1}^n \frac{\lambda^r}{\Gamma(r)} x_i^{r-1} e^{-\lambda x_i} \\
&= \frac{\lambda^{nr}}{\Gamma^n(r)} (\Pi x)^{r-1} e^{-\lambda \Sigma x} \\
\log L &= nr \log \lambda + (r-1) \log(\Pi x) \\
&\quad - \lambda \Sigma x - n \log \Gamma(r) \\
\Rightarrow \frac{1}{L} \frac{\partial L}{\partial r} &= n \log \lambda + \log(\Pi x) - n\psi(r) \\
\Rightarrow n\psi(\hat{r}) &= n \log \hat{\lambda} + \log(\Pi x) \quad (1) \\
\text{And } \frac{1}{L} \frac{\partial L}{\partial \lambda} &= nr/\lambda - \Sigma x \\
\Rightarrow \hat{\lambda} &= \frac{n\hat{r}}{\Sigma x} \quad (2)
\end{aligned}$$

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$$\begin{aligned}
X &= \{2, 3, 5, 9, 10\} \\
f(x) &= \frac{1}{\beta - \alpha + 1}, \alpha \leq x \leq \beta \\
\Rightarrow \mathbb{E}(X^k) &= \sum_{x=\alpha}^{\beta} \frac{x^k}{\beta - \alpha + 1} \\
\Rightarrow \bar{X} &= \frac{\sum_{x=\hat{\alpha}}^{\hat{\beta}} x}{\hat{\beta} - \hat{\alpha} + 1} \\
&= (\hat{\beta} + \hat{\alpha})/2
\end{aligned} \tag{1}$$

$$\begin{aligned}
\text{And } \bar{X^2} &= \frac{\sum_{x=\hat{\alpha}}^{\hat{\beta}} x^2}{\hat{\beta} - \hat{\alpha} + 1} \\
&= (\hat{\beta} - \hat{\alpha} + 2\hat{\alpha}^2\hat{\beta}^2 + 2\hat{\beta}^2 + 2\hat{\alpha}^2)/6
\end{aligned} \tag{2}$$

Where  $\bar{X} = \Sigma X/n = 5.8$  and  $\bar{X^2} = \Sigma X^2/n = 43.8$ . Solving 1 and 2, we get:

$$\begin{aligned}
\hat{\alpha} &= \bar{X} - \frac{\sqrt{12\bar{X^2} + 12\bar{X}^2} - 1}{2} \approx 1 \\
\hat{\beta} &= \bar{X} + \frac{\sqrt{12\bar{X^2} + 12\bar{X}^2} - 1}{2} \approx 11
\end{aligned}$$

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a)

$$\begin{aligned}
\mathbb{E}(X) &= r/\lambda \\
\Rightarrow \bar{X} &= \hat{r}/\hat{\lambda} \\
\mathbb{E}(X^2) &= \text{Var}(X) + \mathbb{E}(X)^2 \\
&= r/\lambda^2 + r^2/\lambda^2 \\
&= r(1+r)/\lambda \\
\Rightarrow \bar{X^2} &= \hat{r}(1+\hat{r})/\hat{\lambda}^2
\end{aligned} \tag{1}$$

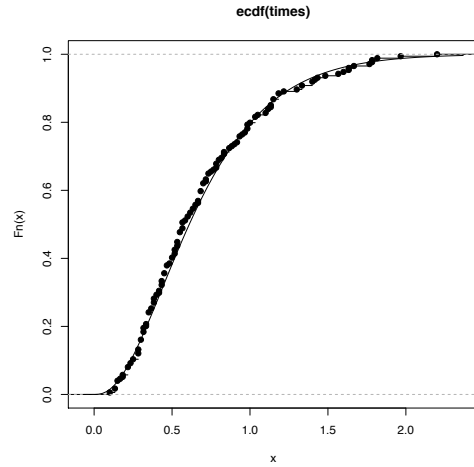
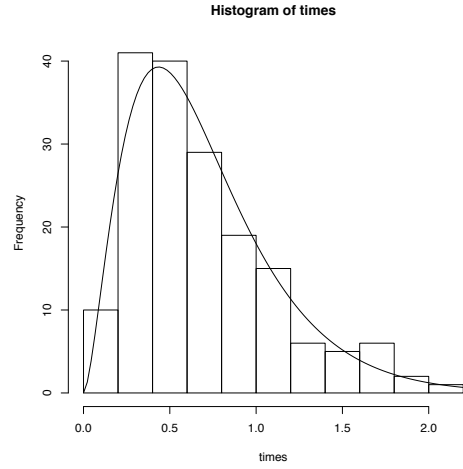
Solving 1 and 2,

$$\begin{aligned}
r &= \frac{\bar{X^2}}{\bar{X^2} - \bar{X}^2} \approx 2.67 \\
\lambda &= \frac{\bar{X}}{\bar{X^2} - \bar{X}^2} \approx 3.84
\end{aligned}$$

b)

We run a test on the null hypothesis  $H_0$  that the data are  $\sim \Gamma(\hat{r}, \hat{\lambda})$ , against the alternative hypothesis  $H_A$  that they are not. Running the goodness-of-fit test yields a  $p$ -value of 0.78, allowing us to strongly favour  $H_0$ .

c)



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$$\begin{aligned}\mathbb{E}(X) &= \mathbb{E}\left(\sum_{i=1}^n a_i X_i\right) \\ &= \sum_{i=1}^n a_i \mathbb{E}(X_i) \\ &= \sum_{i=1}^n a_i \mu \\ &= \mu \Sigma a\end{aligned}$$

For an unbiased estimator of  $\mu$ ,

$$\begin{aligned}\mathbb{E}(X) &= \mu \\ \implies \mu \Sigma a &= \mu' \\ \implies \Sigma a &= 1\end{aligned}$$

27

a)

$$\begin{aligned}\hat{\sigma}^2 &= \frac{1}{n} \Sigma (x_i - \bar{x})^2 \\ &= \frac{\sigma^2}{\sigma^2} \frac{n-1}{n} \frac{1}{n-1} \Sigma (x_i - \bar{x})^2 \\ &= \frac{\sigma^2}{n} \frac{(n-1)S^2}{\sigma^2}\end{aligned}$$

Where  $\frac{(n-1)S^2}{\sigma^2} \sim \chi_{n-1}^2$ .

$$\begin{aligned}\mathbb{E}(\hat{\sigma}^2) &= \mathbb{E}\left(\frac{\sigma^2}{n} \frac{(n-1)S^2}{\sigma^2}\right) \\ &= \frac{\sigma^2}{n} \mathbb{E}\left(\frac{(n-1)S^2}{\sigma^2}\right) \\ &= \frac{\sigma^2}{n} (n-1) = \sigma^2(1 - 1/n) \\ \implies \text{Bias}(\hat{\sigma}^2) &= \mathbb{E}(\hat{\sigma}^2) - \sigma^2 \\ &= \sigma^2(1 - 1/n - 1) = -\sigma^2/n\end{aligned}$$

b)

$$\begin{aligned}\text{Var } \hat{\sigma}^2 &= \frac{\sigma^4}{n^2} \text{Var } \frac{(n-1)S^2}{\sigma^2} \\ &= \frac{\sigma^4}{n^2} 2(n-1) = 2\sigma^4(n-1)/n^2\end{aligned}$$

c)

$$\begin{aligned}\text{MSE} &= \text{Var}(\hat{\sigma}^2) + \text{Bias}^2(\hat{\sigma}^2) \\ &= \frac{2\sigma^4(n-1)}{n^2} + \frac{\sigma^4}{n^2} \\ &= \sigma^4(2n-1)/n^2\end{aligned}$$

34

a)

$$\begin{aligned}f(x) &= 6x^5/\theta^6 \\ \implies F(x) &= \int_0^x 6x^5/\theta^6 dt = x^6/\theta^6 \\ f_{X,\max} &= f_{(n)}(x) \\ &= nf(x)F^{n-1}(x) \\ &= n \frac{6x^5}{\theta^6} \frac{x^{6n-6}}{\theta^{6n-6}} \\ &= 6nx^{6n-1}/\theta^{6n}\end{aligned}$$

b)

$$\begin{aligned}\mathbb{E}(X_{\max}) &= \int_0^\theta x \frac{6nx^{6n-1}}{\theta^{6n}} dx \\ &= \frac{6n}{\theta^{6n}} \frac{x^{6n+1}}{6n+1} \Big|_0^\theta \\ &= \frac{6n}{6n+1} \theta\end{aligned}$$

c)

$$\begin{aligned}\text{Bias} &= \mathbb{E}(X_{\max}) - \theta \\ &= \theta \left( \frac{6n}{6n+1} - 1 \right) \\ &= \frac{-\theta}{6n+1}\end{aligned}$$

d)

$$\begin{aligned}\mathbb{E}(X_{\max}^2) &= \int_0^\theta x^2 \frac{6nx^{6n-1}}{\theta^{6n}} dx \\ &= 6n\theta^2/(6n+2) \\ \text{Var} &= EX_{\max}^2 - E^2 X_{\max} \\ &= \frac{3\theta^2 n}{(3n+1)(6n+1)^2} \\ \text{MSE} = \text{Var} + \text{Bias}^2 &= \frac{\theta^2}{18n^2 + 9n + 1}\end{aligned}$$

**36**

a)

$$\begin{aligned}\mathbb{E}(W) &= \mathbb{E}(a\bar{X}) + \mathbb{E}((1-a)\bar{Y}) \\ &= aE\bar{X} + (1-a)E\bar{Y} \\ &= a\mu + (1-a)\mu = 1\end{aligned}$$

b)

$$\begin{aligned}\text{Var } W &= \text{Var}(a\bar{X} + (1-a)\bar{Y}) \\ &= a^2 \text{Var } \bar{X} + (1-a)^2 \text{Var } \bar{Y} \\ &= a^2 \sigma_1^2/n + (1-a)^2 \sigma_2^2/m \\ \implies \text{Var}'(W) &= 2a\sigma_1^2/n - 2(1-a)\sigma_2^2/m \\ \implies \hat{a}\sigma_1^2/n &= \sigma_2^2/m - \hat{a}\sigma_2^2/m \\ \implies \hat{a} &= \frac{\sigma_2^2/m}{\sigma_1^2/n + \sigma_2^2/m}\end{aligned}$$

# MATH 308 Assignment 17: Exercises 7.6

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1

a)

The actual mean of the distribution is not random; thus the probability of it being in any interval is necessarily either zero or one, and can never be an intermediate value such as 95%.

b)

This is a correct interpretation.

c)

No, because each sample will generate different 95% confidence intervals. However, we can be sure that 95% of these will contain the true mean.

d)

This is a correct interpretation.

e)

No, by same reasoning as (c).

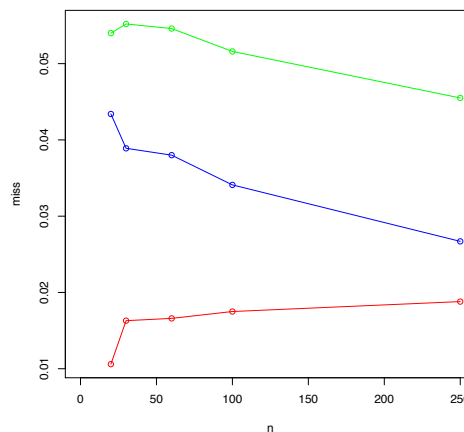
2

$1 - \alpha = 0.95 \implies 1 - \frac{\alpha}{2} = 0.975$ . The 95% confidence interval is given by  $\bar{X} \pm z_{0.975}\sigma/\sqrt{n} \approx 538 \pm 39 = [499, 577]$ .

7

$1 - \alpha = 0.95 \implies \alpha = 0.05$ . The 95% lower  $t$  confidence bound is  $\bar{X} - t_{n-1,0.05}S/\sqrt{n} \approx 118$ . This means that we can be 95% confident that the typical battery will last more than 122 hours.

8



We can see that the rate of the CI being too low (blue) somewhat decreases, that of the CI being too high (red) somewhat increases, and the overall miss rate (green) somewhat decreases. However, it must be noted that the plot is highly zoomed in and the actual change between any two endpoints is less than 1%.

10

$$\begin{array}{ll} n = 43 & m = 36 \\ \bar{X} = 5.8 & \bar{Y} = 1.9 \\ S_1 = 8.6 & S_2 = 4.2 \end{array}$$

$$\begin{aligned} \nu &\equiv \frac{(S_1^2/n + S_2^2/m)^2}{\frac{(S_1^2/n)^2}{n-1} + \frac{(S_2^2/m)^2}{m-1}} \approx 63 \\ T &\equiv \frac{(\bar{X} - \bar{Y}) - (\mu_X - \mu_Y)}{\sqrt{S_1^2/n + S_2^2/m}} \\ &\equiv \frac{\Delta^* - \Delta}{S} \approx T(\nu) \\ 1 - \alpha = 0.95 &\implies 1 - \alpha/2 = 0.975 \\ &\implies t \equiv t_{\nu, 1-\alpha/2} \approx 2.0 \end{aligned}$$

The CI is therefore  $\Delta^* \pm tS \approx 3.9 \pm 3.0 = (0.9, 6.9)$ . This interval tells us that we can be 95% confident that a severely obese patient on a low-carbohydrate diet will loose between 0.9 and 6.9 kilograms more than one on a low-fat diet.

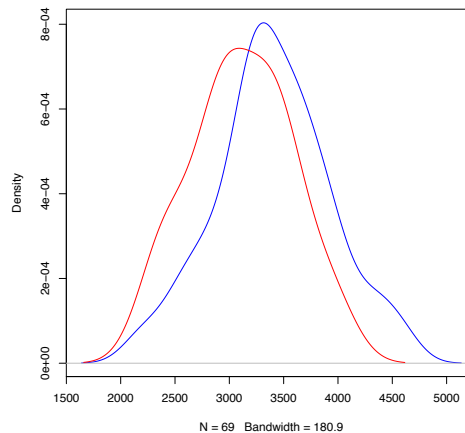
12

a)

Smoker data in red. Non-smoker data in blue. The plot points to greater birth weight for non-smoking mothers.

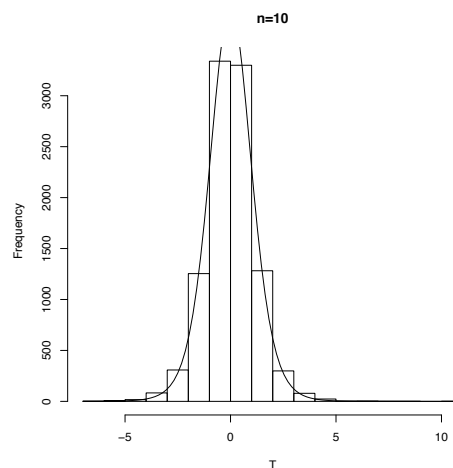
b)

95% CI is (-47,616). This interval tells us that we cannot eliminate the possibility that the difference in means is zero.

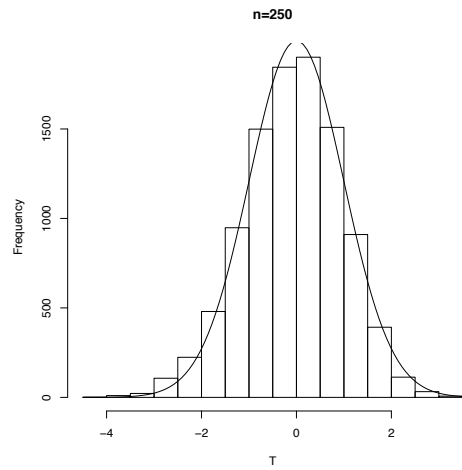
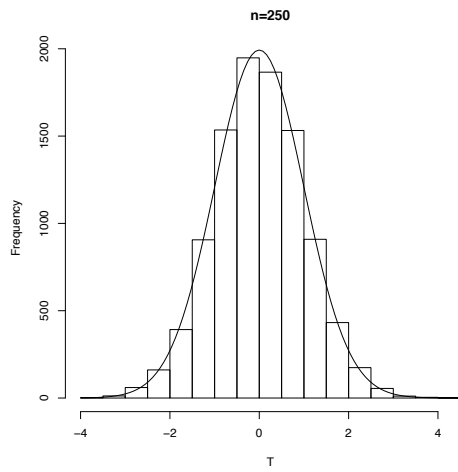
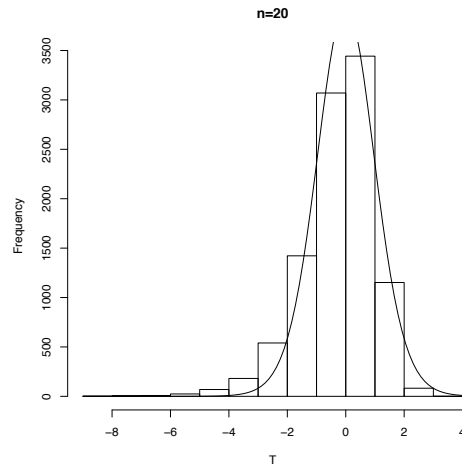
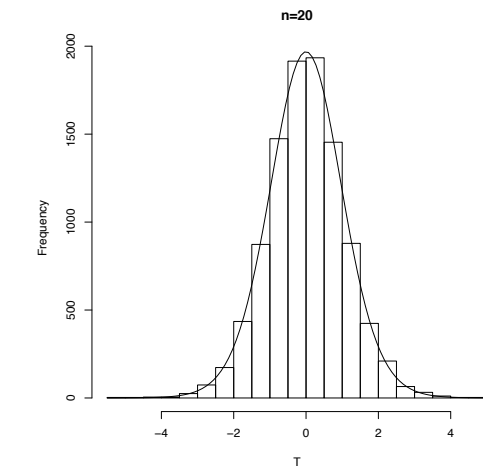


17

a) Uniform (0,1) Distribution



## b) Exponential (1) Distribution



As with the previous distribution, the statistic closely follows the  $t$ -distribution.

It can be seen that the statistic distribution is very close to the  $t$ -distribution, even for small values of  $n$ .



## 20

a)

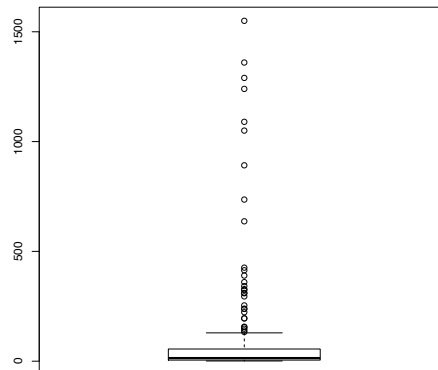
[44%, 50%].

b)

[53%, 60%]. The intervals do not overlap. We can conclude with 95% confidence that a man is more likely to have voted for Bush than Women.

c)

[-14%, -5%]. This tells us with 95% confidence that between 5% and 14% fewer women than men voted for Bush.



The data is concentrated around zero with several large outliers.

b)

$1 - \alpha = 0.95 \implies 1 - \alpha/2 = 0.975 \implies t \equiv t_{268, 1-\alpha/2} \approx 1.97$ . Therefore 95% CI is  $\bar{X} \pm tS/\sqrt{n} \approx 78.1 \pm 25.2 = [52.9, 103.3]$ .

c)

Bootstrap percentile confidence interval from Assignment 15, Problem 10: [55.0, 104.61]. Bootstrap  $t$  confidence interval: [57.0, 111.7]. We report the bootstrap  $t$  confidence interval due to the large sample size and high skewness of the underlying data.

## 24

-16%. This tells us that we can be 95% confident that at least 16% fewer patients on aspirin will have one or more adenomas than those on a placebo.

## 25

a)

Min.	1st Qu.	Median	Mean	3rd Qu.	Max.
1.00	5.00	14.20	78.08	55.50	1550.00

## 31

From Thm 7.1,

$$T = \frac{\bar{X} - \bar{Y} - (\mu_1 - \mu_2)}{S_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}} \sim T(n_1 + n_2 - 2)$$

Let  $\Delta^* \equiv \bar{X} - \bar{Y}$ ,  $\Delta \equiv \mu_1 - \mu_2$ , and  $S \equiv S_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}$ .

$$\begin{aligned} \Rightarrow 1 - \alpha &= P\left(-q \leq \frac{\Delta^* - \Delta}{S} \leq q\right) \\ &= P(\Delta^* - qS < \Delta < \Delta^* + qS) \end{aligned}$$

Therefore the CI is  $[\Delta^* - qS, \Delta^* + qS] = [\bar{X} - \bar{Y} - qS_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}, \bar{X} - \bar{Y} + qS_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}]$   $\square$

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$$\begin{aligned} 1 - \alpha &= P(-q_{\alpha/2} \leq 2\lambda X \leq q_{1-\alpha/2}) \\ &= P\left(\frac{-q_1}{2X} \leq \lambda \leq \frac{q_2}{2X}\right) \end{aligned}$$

The CI is therefore  $\approx [0.242/X, 5.57/X]$ .

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$$\begin{aligned} \frac{(n-1)S^2}{\sigma^2} &\sim \chi_{n-1}^2 \\ \Rightarrow 1 - \alpha &= P\left(q_1 \leq \frac{(n-1)S^2}{\sigma^2} \leq q_2\right) \\ &= P\left(\frac{q_1}{(n-1)S^2} \leq \frac{1}{\sigma^2} \leq \frac{q_2}{(n-1)S^2}\right) \\ &= P\left(\frac{(n-1)S^2}{q_1} \geq \sigma^2 \geq \frac{(n-1)S^2}{q_2}\right) \\ \Rightarrow \text{CI} &= [(n-1)S^2/q_2, (n-1)S^2/q_1] \quad \square \end{aligned}$$