

Model fit

HL Chapter 5 – part 2

Questions

- Does the model fit overall?
→ Summary goodness-of-fit tests ✓



- Are there any individual observations that don't fit?
→ Logistic regression diagnostics



Logistic regression diagnostics

Leverage

- How different from the other covariate pattern is this covariate pattern?

Change in Pearson χ^2 or Deviance

- How much do Pearson χ^2 and Deviance test statistics decrease if this covariate pattern is deleted
- I.e., is there any evidence of improved model fit if this covariate pattern is deleted?

Change in coefficients

- How much does deleting this covariate pattern affect the model coefficients?

Leverage

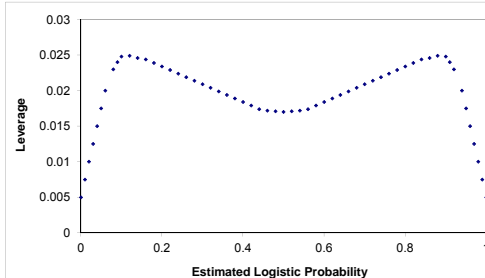
h

- How different from the other covariate patterns is this covariate pattern?

Leverage depends on $\hat{\pi}$

- $0.1 < \hat{\pi} < 0.9$: The greater h the more “unusual” the covariate pattern
- $\hat{\pi} \leq 0.1$ or $\hat{\pi} \geq 0.9$: h will be small even if the covariate pattern is “unusual”
- Look for deviations from the expected pattern (observations outside the “cloud”)

Leverage – expected pattern (sort of)



Change in Pearson chi-square or Deviance

ΔX^2 or ΔD

- How much smaller are the Pearson X^2 and Deviance goodness-of-fit test statistics if this covariate pattern is deleted?
- I.e., does model fit improve if the covariate pattern is deleted?

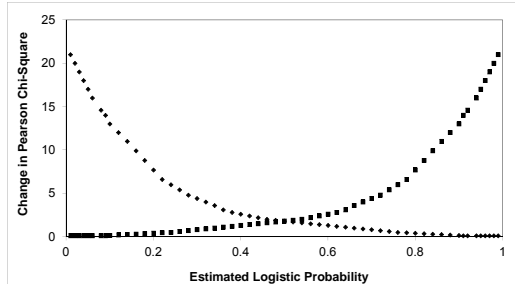
ΔX^2 and ΔD depend on $\hat{\pi}$

- Excellent fit most likely for
 - $\hat{\pi} < 0.1$ ($\hat{\pi}$ very close to $y=0$)
 - $\hat{\pi} > 0.9$ ($\hat{\pi}$ very close to $y=1$)
- Terrible fit most likely for
 - $\hat{\pi} < 0.1$ ($\hat{\pi}$ very different from $y=1$)
 - $\hat{\pi} > 0.9$ ($\hat{\pi}$ very different from $y=0$)
- Expect smallest and greatest ΔX^2 and ΔD for $\hat{\pi} < 0.1$ and $\hat{\pi} > 0.9$

ΔX^2 and ΔD depend on $\hat{\pi}$

- $0.3 < \hat{\pi} < 0.7$: ΔX^2 & ΔD likely to be fairly small
- Look for deviations from the expected pattern (observations outside the “cloud”), e.g.,
 - High ΔX^2 & ΔD for $\hat{\pi} > 0.3$ or $\hat{\pi} < 0.7$
 - Very high ΔX^2 & ΔD for $\hat{\pi} < 0.1$ or for $\hat{\pi} > 0.9$

ΔX^2 and ΔD - Expected pattern (sort of)



Change in coefficients

$\Delta \hat{\beta}$

- How much does deleting this covariate pattern affect the model coefficients?
- $\Delta \hat{\beta}$ combines leverage and ΔX^2

$\Delta \hat{\beta}$ depends on $\hat{\pi}$

- Expect greatest $\Delta \hat{\beta}$ values when neither leverage nor ΔX^2 are very small
- I.e., $\Delta \hat{\beta}$ is expected to be highest for $\hat{\pi}$ between 0.1 and 0.3 and for $\hat{\pi}$ between 0.7 and 0.9
- Look for deviations from the expected pattern (observations outside the “cloud”)

Change in specific coefficients

Note:

- $\Delta\hat{\beta}$ assesses the effect on all coefficients
- If $\Delta\hat{\beta}$ is large, it may be worth checking which coefficients are most affected by the deletion of the covariate pattern

If the model doesn't fit ...

- Try rebuilding the model
 - Continuous covariates may have been modeled in the wrong scale
 - Important risk factors, confounders or effect modifiers may have been missed
- Standard logistic regression model may not work for small or large $\hat{\pi}$
 - Try model with the extra parameters that allow for the tails to vary (from Stukel test)

If the model doesn't fit ...

- Logistic regression model may not work
 - Try a different regression model
- If nothing works, one or more crucial covariates may not have been measured
 - ☹️
- However, lack of fit does not necessarily mean that the OR estimates are way off or that the model won't predict well

Example – GLOW500 data set

Create covariate patterns

```
proc sort data=glow500;
  by priorfrac momfrac armassist raterisk2 height age;
run;

proc means n sum noprint data=glow500;
  by priorfrac momfrac armassist raterisk2 height age;
  var fracture; output out=jdat n=m_j sum=y_j;
run;
```

Number of observations in covariate pattern j

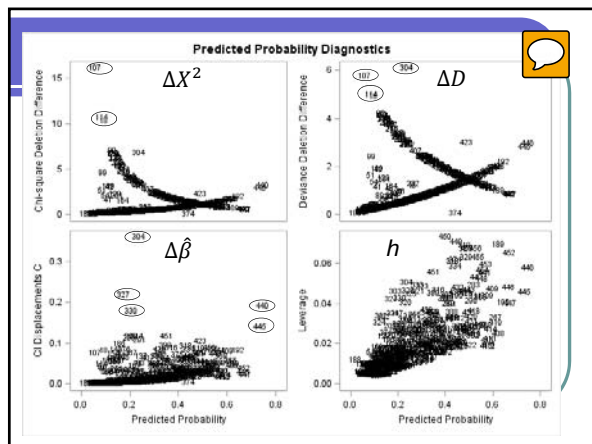
Number of observations with outcome in covariate pattern j

Run proc logistic using m_j and y_j Create graphs, save diagnostics

```
Plot h,  $\Delta X^2$ ,  $\Delta D$  and  $\Delta\hat{\beta}$  vs.  $\pi_{\text{hat}}$ 

proc logistic descending data=jdat plots(only label)=(phat);
  model y_j/m_j = priorfrac momfrac armassist raterisk2 height
    age priorfrac*age momfrac*armassist;
  output out=diag
    h=h difchisq=difchisq difdev=difdev c=db p=pihat;
run;

Save h,  $\Delta X^2$ ,  $\Delta D$  and  $\Delta\hat{\beta}$ 
```



Evaluate outliers

```
data diag; set diag; (j=_N_); run;
```

Add covariate
pattern numbers
to data set

```
proc print data=diag noobs;
```

```
where j in (107,304,327,330,440,445);
```

```
var j priorfrac momfrac armassist raterisk2 height age  
y_j m_j pihat h difchisq difdev db;  
run;
```

Outlier table

| j | PRIORFRAC | MOMFRAC | ARMASIST | RATERISK2 | HEIGHT | AGE | y_j | m_j | pihat | h | difchisq | difdev | db |
|-----|-----------|---------|----------|-----------|--------|-----|-----|-----|-------|-------|----------|--------|-------|
| 107 | 0 | 0 | 0 | 0 | 166 | 57 | 1 | 1 | 0.059 | 0.005 | 16.099 | 5.744 | 0.075 |
| 304 | 0 | 1 | 0 | 0 | 168 | 65 | 2 | 2 | 0.238 | 0.051 | 6.738 | 6.080 | 0.359 |
| 327 | 0 | 1 | 1 | 0 | 175 | 75 | 1 | 1 | 0.175 | 0.042 | 4.934 | 3.698 | 0.217 |
| 330 | 0 | 1 | 1 | 1 | 162 | 60 | 1 | 1 | 0.208 | 0.043 | 3.973 | 3.308 | 0.177 |
| 440 | 1 | 1 | 0 | 0 | 142 | 70 | 0 | 1 | 0.747 | 0.057 | 3.131 | 2.928 | 0.191 |
| 445 | 1 | 1 | 0 | 1 | 153 | 63 | 0 | 1 | 0.735 | 0.046 | 2.913 | 2.793 | 0.139 |

Why is 107 an outliers?

- Low risk profile
- Therefore, low pihat
- But FRACTURE=1 (y_j and m_j are 1)
- Covariate pattern may be unusual and not represented well by the model
- None of the covariate values are unreasonable

Why is 304 an outlier?

- Fairly low risk profile
- Therefore, lowish pihat
- But both observations in the covariate pattern have FRACTURE=1 (y_j and m_j are 2)
- Covariate pattern may be unusual and not represented well by the model
- None of the covariate values are unreasonable

Why are 327 and 330 outliers?

- PRIORFRAC seems to greatly influence pihat
- Highish risk profile but PRIORFRAC=0
- Therefore, lowish pihat
- But both covariate patterns have FRACTURE=1 (y_j and m_j are 1 in both)
- Covariate pattern may be unusual and not represented well by the model
- None of the covariate values are unreasonable

Why are 440 and 445 outliers?

- PRIORFRAC seems to greatly influence pihat
- High(ish) risk profile and PRIORFRAC=1
- Therefore, highish pihat
- But both covariate patterns have FRACTURE=0 ($y_j = 0$ and $m_j = 1$ in both)
- Covariate pattern may be unusual and not represented well by the model
- None of the covariate values are unreasonable

Determine the effect of outliers on the model

- Delete outlier (one at a time or in groups)
- Rerun logistic regression model
- Compare ORs and p-values to those in the model based on all data
- Decide what to do with the outlier(s)

Model without covariate pattern 107

Compare to proc logistic for final model in Chapter4_4Results.pdf

```
proc logistic descending data=diag;
  where j ne 107;
```

```
model y_j/m_j = priorfrac momfrac armassist raterisk2 height
  age priorfrac*age momfrac*armassist;
```

```
contrast 'raterisk greater vs. same/less' raterisk2 1/estimate=exp;
contrast 'height increase of 10 cm' height 10/estimate=exp;
```

Model w/o covariate pattern 107, cont.

```
contrast 'prior fracture yes vs. no at age 55' priorfrac 1 priorfrac*age 55
/estimate=exp;
contrast 'prior fracture yes vs. no at age 60' priorfrac 1 priorfrac*age 60
/estimate=exp;
< Etc. >
contrast 'prior fracture yes vs. no at age 90' priorfrac 1 priorfrac*age 90
/estimate=exp;
```

Model w/o covariate pattern 107, cont.

```
contrast 'age+10 at priorfrac=1' age 10 priorfrac*age 10/estimate=exp;
contrast 'age+10 at priorfrac=0' age 10 priorfrac*age 0/estimate=exp;
```

```
contrast 'momfrac yes vs. no at armassist=1' momfrac 1
  momfrac*armassist 1/estimate=exp;
contrast 'momfrac yes vs. no at armassist=0' momfrac 1
  momfrac*armassist 0/estimate=exp;
contrast 'armassist yes vs. no at momfrac=1' armassist 1
  momfrac*armassist 1/estimate=exp;
contrast 'armassist yes vs. no at momfrac=0' armassist 1
  momfrac*armassist 0/estimate=exp;
```

```
run;
```

ORs before and after deletion of outliers

| Contrast | OR0 | OR107 | OR304 | OR327 | OR330 | OR440 | OR445 |
|-------------------------------------|------|-------|-------|-------|-------|-------|-------|
| raterisk greater vs. same/less | 1.60 | 1.63 | 1.65 | 1.64 | 1.57 | 1.56 | 1.63 |
| height increase of 10 cm | 0.63 | 0.62 | 0.61 | 0.60 | 0.63 | 0.59 | 0.61 |
| prior fracture yes vs. no at age 55 | 4.82 | 5.09 | 4.95 | 4.69 | 5.00 | 5.07 | 5.27 |
| prior fracture yes vs. no at age 60 | 3.65 | 3.80 | 3.74 | 3.60 | 3.77 | 3.83 | 3.93 |
| prior fracture yes vs. no at age 65 | 2.77 | 2.84 | 2.83 | 2.76 | 2.84 | 2.89 | 2.93 |
| prior fracture yes vs. no at age 70 | 2.10 | 2.12 | 2.14 | 2.12 | 2.14 | 2.18 | 2.18 |
| prior fracture yes vs. no at age 75 | 1.59 | 1.59 | 1.61 | 1.63 | 1.62 | 1.64 | 1.62 |
| prior fracture yes vs. no at age 80 | 1.21 | 1.19 | 1.22 | 1.25 | 1.22 | 1.24 | 1.21 |
| prior fracture yes vs. no at age 85 | 0.92 | 0.89 | 0.92 | 0.96 | 0.92 | 0.94 | 0.90 |
| prior fracture yes vs. no at age 90 | 0.70 | 0.66 | 0.70 | 0.74 | 0.69 | 0.71 | 0.67 |
| age+10 at priorfrac=1 | 1.02 | 1.02 | 1.02 | 1.03 | 1.03 | 1.01 | 0.99 |
| age+10 at priorfrac=0 | 1.77 | 1.83 | 1.78 | 1.75 | 1.82 | 1.77 | 1.78 |
| momfrac yes vs. no at armassist=1 | 0.97 | 0.96 | 0.96 | 0.81 | 0.82 | 0.97 | 0.97 |
| momfrac yes vs. no at armassist=0 | 3.48 | 3.62 | 2.87 | 3.52 | 3.51 | 3.91 | 3.85 |
| armassist yes vs. no at momfrac=1 | 0.53 | 0.52 | 0.64 | 0.44 | 0.44 | 0.48 | 0.48 |
| armassist yes vs. no at momfrac=0 | 1.90 | 1.95 | 1.90 | 1.92 | 1.89 | 1.92 | 1.90 |

P-values before and after deletion of outliers

| Contrast | p0 | p107 | p304 | p327 | p330 | p440 | p445 |
|-------------------------------------|------|------|------|------|------|------|------|
| raterisk greater vs. same/less | 0.05 | 0.04 | 0.04 | 0.04 | 0.06 | 0.06 | 0.04 |
| height increase of 10 cm | 0.01 | 0.01 | 0.01 | 0.01 | 0.01 | 0.01 | 0.01 |
| prior fracture yes vs. no at age 55 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 |
| prior fracture yes vs. no at age 60 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 |
| prior fracture yes vs. no at age 65 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 |
| prior fracture yes vs. no at age 70 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 |
| prior fracture yes vs. no at age 75 | 0.07 | 0.07 | 0.06 | 0.06 | 0.06 | 0.05 | 0.06 |
| prior fracture yes vs. no at age 80 | 0.55 | 0.59 | 0.54 | 0.49 | 0.54 | 0.50 | 0.55 |
| prior fracture yes vs. no at age 85 | 0.84 | 0.77 | 0.85 | 0.92 | 0.84 | 0.88 | 0.80 |
| prior fracture yes vs. no at age 90 | 0.49 | 0.44 | 0.49 | 0.56 | 0.49 | 0.51 | 0.45 |
| age+10 at priorfrac=1 | 0.92 | 0.92 | 0.93 | 0.87 | 0.87 | 0.97 | 0.96 |
| age+10 at priorfrac=0 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 |
| momfrac yes vs. no at armassist=1 | 0.94 | 0.93 | 0.93 | 0.67 | 0.69 | 0.95 | 0.96 |
| momfrac yes vs. no at armassist=0 | 0.00 | 0.00 | 0.01 | 0.00 | 0.00 | 0.00 | 0.00 |
| armassist yes vs. no at momfrac=1 | 0.27 | 0.26 | 0.44 | 0.17 | 0.17 | 0.20 | 0.21 |
| armassist yes vs. no at momfrac=0 | 0.01 | 0.01 | 0.01 | 0.01 | 0.01 | 0.01 | 0.01 |

Decision

- May want to rerun Stukel test (see Chapter 5, part 1) after deleting outliers to see if the tails assumption is now satisfied
- But...deleting outliers results in little change in ORs and p-values
- Would probably keep all outliers in the data set

What if your model had poor overall goodness-of-fit?

- May want to rerun goodness-of-fit tests (see Chapter 5, part 1) after deleting outliers to see if the model fits better
- Not necessary in this example

Assessment of fit via external validation

- A model always performs better on the developmental data set

Idea

- If the data set is large enough, exclude a subsample
- Or, if possible, collect additional data
- Build the model on the original data
- Test model fit on the subsample or the additional data
- Treat the coefficients as fixed constants rather than estimated values
- Or: Try bootstrapping as described by Austin et al.